

ARE 210: Oct. 27, 2021 Section Notes

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Setup: Suppose that $\{X_i\}_{i=1}^n$ are a sequence of iid Poisson random variables with $\mathbb{P}(X_i = k) = \frac{\exp(-\theta)\theta^k}{k!}$ for $\theta > 0$.

- f) Compute the MLE of θ and denote it $\hat{\theta}$.
- g) What does $\hat{\theta}$ converge to in probability?
- h) Derive the limiting distribution of (a suitably normalized version of) $\hat{\theta}$
- i) Define the estimator $\hat{\delta} = \exp(-\hat{\theta})$. Is $\hat{\delta}$ the MLE of $\exp(-\theta)$? Justify your answer.

Solution:

Strategy: For part f we'll simply write down the sample likelihood and find the θ that maximizes it. Part g will follow from a law of large numbers. Similarly, part h will follow from a central limit theorem. For part i, rather than deriving the sample likelihood for this new estimand we'll invoke the invariance property of MLE to say $\hat{\delta}$ is indeed the MLE of $\exp(-\theta)$.

Step 1: Write down sample likelihood

Since $\mathbb{P}(X_i = k) = \frac{\exp(-\theta)\theta^k}{k!}$ and $\{X_i\}_{i=1}^\infty$ are independent we have

$$\begin{aligned}\mathcal{L}(\{x_i\}_{i=1}^\infty; \theta) &= \prod_{i=1}^n \frac{\exp(-\theta)\theta^{x_i}}{x_i!} \\ &= \frac{\exp(-n\theta)\theta^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!}\end{aligned}$$

Step 2: Optimize

Taking the log, then derivative of both sides and settting equal to zero yields

$$\begin{aligned}
\log(\mathcal{L}(\{x_i\}_{i=1}^{\infty}; \theta)) &= -n\theta + \log(\theta) \sum_{i=1}^n x_i - \sum_{i=1}^n \log(x_i!) \\
\frac{\partial \log(\mathcal{L}(\{x_i\}_{i=1}^{\infty}; \hat{\theta}))}{\partial \hat{\theta}} &= -n + \frac{1}{\hat{\theta}} \sum_{i=1}^n x_i = 0 \\
&\Rightarrow \frac{1}{\hat{\theta}} \sum_{i=1}^n x_i = n \\
&\Rightarrow \frac{1}{n} \sum_{i=1}^n x_i = \hat{\theta}
\end{aligned}$$

So the MLE is the sample mean.

Step 3: WLLN To find the probability limit of $\hat{\theta}$, we simply need to apply the WLLN. Since $\mathbb{E}[X_i] = \theta$, we have $\hat{\theta} \xrightarrow[n \rightarrow \infty]{P} \theta$

Step 4: CLT To find the limiting distribution, we'll want to invoke the CLT. Since $\hat{\theta} = \frac{1}{n} \sum_{i=1}^n x_i$, we have

$$\sqrt{n}(\hat{\theta} - \theta) \xrightarrow[n \rightarrow \infty]{d} N(0, \theta)$$

Where $Var(X_i) = \theta$ follows from the Poisson distribution.

Step 5: Invariance Property

Part i follows directly from the invariance property of MLE. To see why, note that $\log(\mathcal{L}(\{x_i\}_{i=1}^{\infty}; -\log(\delta))) = \log(\mathcal{L}(\{x_i\}_{i=1}^{\infty}; \theta))$. So $\hat{\theta}$ is the MLE of $-\log(\delta)$ and since $-\log(\cdot)$ is a one-to-one function it must be that $\hat{\delta}$ is the MLE of δ .