# Lecture 1: Case Control Association Testing & Association Testing with Quantitative Traits

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Summer Institute in Statistical Genetics 2022

#### Introduction

- Association mapping is now routinely being used to identify loci that are involved with complex traits.
- Technological advances have made it feasible to perform association studies on a genome-wide basis with millions of markers in a single study.
- We consider testing a genetic marker for association with a disease (e.g. 1/0, affected/unaffected, dead/alive) or a quantitative trait (e.g. height, BMI) in a sample of unrelated subjects.
- Vast amounts of literature on these topics!

#### **Case-Control Association Testing**

- Allelic Association Tests
  - Allele is treated as the sampling unit
  - Typically make an assumption of Hardy-Weinberg equilibrium (HWE) - alleles within an individual are conditionally independent given the disease status
  - e.g. Pearson's  $\chi^2$
- Genotypic Association Tests
  - Individual is the sampling unit
  - Does not assume HWE
  - e.g. Logistic regression

#### Pearson's $\chi^2$ Test for Allelic Association

- ➤ This test looks for deviations from independence between the trait and allele.
- Consider a single marker with 2 allelic types (e.g., a SNP) labeled "C" and "T".
- Let  $N_{ca}$  be the number of cases and  $N_{co}$  be the number of controls on which we have genotype data.

#### Pearson's $\chi^2$ Test for Allelic Association

ightharpoonup Below is a 2×2 contingency table for trait and allelic type

	Cases	Controls	Total
Allele C	n <sub>C</sub> <sup>ca</sup>	$n_C^{co}$	$N_C$
Allele T	$n_T^{ca}$	$n_T^{co}$	$N_T$
Total	$2N_{ca}$	2 <i>N<sub>co</sub></i>	2 <b>N</b>

- ▶  $n_C^{ca}$  is the number of C alleles in the cases and  $n_C^{ca} = 2 \times$  the number of homozygous CC cases + the number of heterozygous CT cases
- Hypotheses
  - ► H<sub>0</sub>: there is no association between the row variable and column variable
  - $ightharpoonup H_a$ : there is an association between the two variables

#### Pearson's $\chi^2$ Test for Allelic Association

▶ Can use Pearson's  $\chi^2$  test for independence. The statistic is:

$$X^{2} = \sum_{\text{all cells}} \frac{(\text{Observed cell} - \text{Expected cell})^{2}}{\text{Expected cell}}$$

What is the the expected cell number under  $H_0$ ? For each cell, we have

$$\mathsf{Expected} \ \mathsf{Cell} \ \mathsf{Count} = \frac{\mathsf{row} \ \mathsf{total} \ \times \ \mathsf{col} \ \mathsf{total}}{\mathsf{total} \ \mathsf{count}}$$

▶ Under  $H_0$ , the  $X^2$  test statistic has an approximate  $\chi^2$  distribution with (r-1)(c-1) = (2-1)(2-1) = 1 degree of freedom

► From Phasukijwattana et al. (2010), Leber Hereditary Optic Neuropathy (LHON) disease and genotypes for marker rs6767450:

	CC	СТ	TT
Cases	6	8	75
Controls	10	66	163

Corresponding 2 × 2 contingency table with allelic type instead of genotype

	Allele C	Allele T
Cases	20	158
Controls	86	392

Should we reject the hypothesis that allelic type is independent of disease status?

	Allele C	Allele T	Total
Cases	20	158	178
Controls	86	392	478
Total	106	550	656

- ► Intuition for the test: Suppose H<sub>0</sub> is true, so allelic type and case-control status are independent, then what counts would we expect?
  - the expected number of case alleles that are of type C is:

$$n_C^{ca} = \{\#cases\} \times P(\text{Allelic type is C} \mid \text{Allele is from a Case})$$

$$= \{\#cases\} \times P(\text{Allelic type is C}) \quad \text{(independence)}$$

$$= 178 \times \left(\frac{106}{656}\right) = 28.7622$$

▶ Fill in the remaining cells for the expected counts

	Allele C	Allele T	Total
Cases	28.7622		178
Controls	•		478
Total	106	550	656

Fill in the remaining cells for the expected counts

	Allele C	Allele T	Total
Cases	28.7622	149.2378	178
Controls	77.2378	400.7622	478
Total	106	550	656

Fill in the remaining cells for the expected counts

	Allele C	Allele T	Total
Cases	28.7622	149.2378	178
Controls	77.2378	400.7622	478
Total	106	550	656

► Calculate the X² statistic

$$X^{2} = \frac{(20 - 28.7622)^{2}}{28.7622} + \dots + \frac{(392 - 400.7622)^{2}}{400.7622} = 4.369$$

Fill in the remaining cells for the expected counts

	Allele C	Allele T	Total
Cases	28.7622	149.2378	178
Controls	77.2378	400.7622	478
Total	106	550	656

► Calculate the X<sup>2</sup> statistic

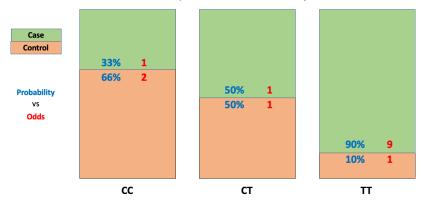
$$X^{2} = \frac{(20 - 28.7622)^{2}}{28.7622} + \dots + \frac{(392 - 400.7622)^{2}}{400.7622} = 4.369$$

▶ What is the *p*-value?

$$P(\chi_1^2 \ge 4.369) = .037$$

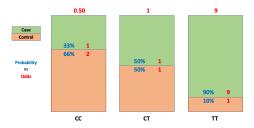
## Odds Ratios (ORs) for Genotypes

▶ What are **odds**? An expression of **relative probabilities**...



lacktriangle Odds of disease in an individual with the CC genotype =50%

### Odds Ratios (ORs) for Genotypes



Typically choose a reference genotype, e.g. CC

$$OR_{TT} = \frac{\text{odds of disease with the TT genotype}}{\text{odds of disease with the CC genotype}} = \frac{9}{0.50} = 18$$

- $ightharpoonup OR_{TT} = 1$  implies no association with disease.
- ▶  $OR_{TT} > 1$  or  $OR_{TT} < 1$  implies association with the disease.

## Genotypic Association Tests: Logistic regression

- Generally used to estimate odds ratios and get confidence intervals for genotypes.
- Let  $\pi_i$  be the probability that individual i is has the disease and let  $G_i$  be the genotype at the SNP:

$$\log \underbrace{\left(\frac{\pi_i}{1-\pi_i}\middle|G_i\right)}_{\text{odds of disease}} = \beta_0 + \beta_{CT}I\{G_i = CT\} + \beta_{TT}I\{G_i = TT\}$$

where  $I\{G_i = CT\}$  is 1 if  $G_i = CT$  and 0 otherwise, and similarly for  $I\{G_i = TT\}$ .

## Genotypic Association Tests: Logistic regression

► The coefficient estimates for  $\hat{\beta}_{CT}$  and  $\hat{\beta}_{TT}$  can be used to calculate odds ratios:

$$\begin{split} \textit{OR}_{\textit{CT}} &= \frac{\text{odds of disease with the CT genotype}}{\text{odds of disease with the CC genotype}} \\ &= \frac{e \times p(\hat{\beta}_0 + \hat{\beta}_{\textit{CT}})}{e \times p(\hat{\beta}_0)} = e \times p(\hat{\beta}_{\textit{CT}}) \end{split}$$

Similarly,  $OR_{TT} = exp(\hat{\beta}_{TT})$ 

▶ 95% CI for  $OR_{CT}$  is

$$exp(\hat{eta}_{CT}\pm 1.96 imes s.e.(\hat{eta}_{CT}))$$

#### **Odds Ratios for LHON Example**

Leber Hereditary Optic Neuropathy (LHON) disease and genotypes for marker rs6767450:

	CC	СТ	TT
Cases	6	8	75
Controls	10	66	163

- We will use the R to obtain odds ratios and confidence intervals for this data set
- We will contrast this test with an allelic test (Pearson's  $\chi^2$  test).

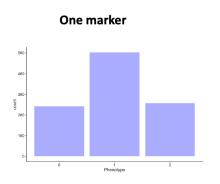
#### Introduction to Quantitative Trait Mapping

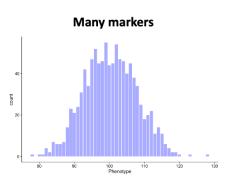
- Quantitative trait loci (QTL) mapping involves identifying genetic loci that influence the phenotypic variation of a quantitative trait.
- QTL mapping is commonly conducted with GWAS, often involving directly genotyped variants along with imputation through reference panels to result in millions of genetic variants

#### Introduction to Quantitative Trait Mapping

- Some quantitative traits can be largely influenced by a single gene as well as by environmental factors (e.g. monogenic or Mendelian traits)
- Influences on a quantitative trait can also be due to a number of genes → polygenicity
- Many quantitative traits of interest are complex where phenotypic variation is due to a combination of both multiple genes and environmental factors
- Examples: Blood pressure, cholesterol levels, IQ, height, weight, etc.

#### Mendelian & Quantitative Genetics





#### **Quantitative Genetic Model**

► The classical quantitative genetics model introduced by Ronald Fisher (1918) is

$$Y = G + E$$

- where Y is the phenotypic value, G is the genetic value, and E is the environmental deviation.
- ► *G* is the combination of all genetic loci that influence the phenotypic value and *E* consists of all non-genetic factors that influence the phenotype (mean set to 0)

#### **Components of Genetic Variance**

Consider a single locus. Fisher modeled the genotypic value G with a linear regression model (least squares) where the genotypic value can be partitioned into an additive component (A) and deviations from additivity as a result of dominance (D), where

$$G = A + D,$$
 $\underbrace{Var(G)}_{\sigma_G^2} = \underbrace{Var(A)}_{\sigma_A^2} + \underbrace{Var(D)}_{\sigma_D^2}$ 

- $\sigma_A^2$  is the **additive genetic variance**. It is the genetic variance associated with the average additive effects of alleles
- $\sigma_D^2$  is the **dominance genetic variance**. It is the genetic variance associated with the dominance effects.

#### Heritability

Remember

$$Y = G + E$$

$$= A + D + E,$$

$$\underbrace{Var(Y)}_{\sigma_Y^2} = \underbrace{Var(A)}_{\sigma_A^2} + \underbrace{Var(D)}_{\sigma_D^2} + \underbrace{Var(E)}_{\sigma_E^2}$$

Narrow-sense heritability (or simply heritability) is

$$h^2 = \frac{\sigma_A^2}{\sigma_Y^2}$$

- ▶  $h^2$  is the proportion of the total phenotypic variance due to additive effects.
- ► It can also be viewed as the extent to which phenotype is determined by the alleles transmitted from the parents.

### Heritability

▶ The **broad-sense heritability** is defined to be

$$H^2 = \frac{\sigma_G^2}{\sigma_Y^2} = \frac{\sigma_A^2 + \sigma_D^2}{\sigma_Y^2}$$

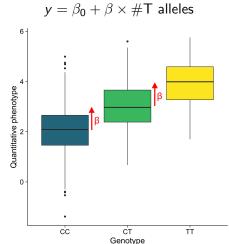
- $ightharpoonup H^2$  is the proportion of the total phenotypic variance that is due to all genetic effects (additive and dominance)
- Heritability can vary over time and with the study population as it depends also on environmental effects

### **QTL** Mapping

- For traits that are heritable, i.e., traits with a non-negligible genetic component that contributes to phenotypic variability, identifying (or mapping) QTLs that influence the trait is often of interest.
- Linear regression models are commonly used for QTL mapping
  - ► They will often include a single genetic marker (e.g., a SNP) as predictor in the model, in addition to other relevant covariates (e.g. age, sex), with the quantitative phenotype as the response

### Linear regression with SNPs

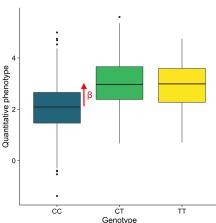
Many analyses fit the 'additive model'



#### Linear regression, with SNPs

An alternative is the 'dominant model';

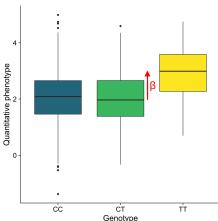
$$y = \beta_0 + \beta \times I\{G \neq CC\}$$



#### Linear regression, with SNPs

or the 'recessive model';

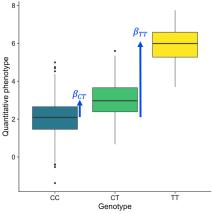
$$y = \beta_0 + \beta \times I\{G == TT\}$$



#### Linear regression, with SNPs

Finally, the 'two degrees of freedom model';

$$y = \beta_0 + \beta_{CT} \times I\{G == CT\} + \beta_{TT} \times I\{G == TT\}$$



#### **Additive Genetic Model**

- ► Most GWAS perform single SNP association testing with linear regression assuming an additive model.
- ▶ The coefficient of determination  $(r^2)$  of an additive linear regression model gives an estimate of the proportion of phenotypic variation that is explained by the SNP (or SNPs) in the model, e.g., the "SNP heritability"

#### **Additive Genetic Model**

Consider the following additive model for association testing with a quantitative trait and a SNP with alleles C and T:

$$Y = \beta_0 + \beta_1 G + \epsilon$$

where G is the number of copies of the allele T.

▶ How would you interpret  $\epsilon$  in this model?

#### **Association Testing with Additive Model**

$$Y = \beta_0 + \beta_1 G + \epsilon$$

▶ Two test statistics for  $H_0$ :  $\beta_1 = 0$  versus  $H_a$ :  $\beta_1 \neq 0$ 

$$\mathcal{T} = rac{\hat{eta}_1}{\sqrt{\mathit{var}(\hat{eta}_1)}} \sim \mathbf{t}_{N-2} pprox \mathit{N}(0,1)$$
 for large  $\mathit{N}$ 

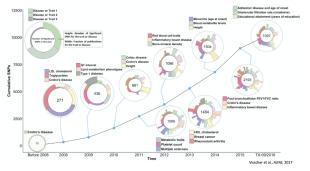
$$T^2 = rac{\hat{eta}_1^2}{var(\hat{eta}_1)} \sim \mathbf{F}_{1,N-2} pprox \chi_1^2$$
 for large  $N$ 

where

$$var(\hat{eta}_1) = \frac{\sigma_{\epsilon}^2}{S_{CC}}$$

and  $S_{GG}$  is the corrected sum of squares for the  $G_i$ 's

#### **Missing Heritability**



- Gap between SNP-based and pedigree-based heritability estimates
- Causal variants not well tagged? Rare variants involved?

#### Time for practical exercises!

- Head to your breakout room
- Introduce yourselves:
  - Name
  - Academic background
  - Research interests
  - ► 1 hobby
- Work on Session 1 exercise (link on class website) https://joellembatchou.github.io/SISG2022\_ Association\_Mapping/Session01\_practical.html

#### References

- Phasukkijwatana, N., Kunhapan, B., Stankovich, J., Chuenkongkaew, W. L., Thomson, R., Thornton, T., Bahlo, M., Mushiroda, T., Nakamura, Y., Mahasirimongkol, S., et al. (2010). Genome-wide linkage scan and association study of PARL to the expression of LHON families in Thailand. Human genetics, 128(1), 39–49.
- Visscher, P.M. et al. (2017) 10 Years of GWAS Discovery: Biology, Function, and Translation. The American Journal of Human Genetics 101, 5-22.