# Question 2

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$$f_Y(y) = \frac{\lambda^y e^{-\lambda}}{y!}, \ \lambda \ge 0, \ y \in \{0, 1, 2, ...\}$$
$$f_Y^*(y^*) = I(Y_i > 0)$$

# 0.1 Likelihood of $Y_i^*$

$$\begin{split} Y_i^* \sim bernoulli(p) \\ p &= P(Y=1) \\ f_Y^*(y^*) &= P(Y_i = y|p) = p^y (1-p)^{1-y} I_{(0,1)}^{(p)}, y \in \{0,1\} \\ L(p|Y_1^*,Y_2^*,...Y_n^*) &= \prod_{i=1}^n f_{Y_i^*}(y_i^*|p) \\ L(p|Y_1^*,Y_2^*,...Y_n^*) &= \prod_{i=1}^n \left[ p_i^{y*} (1-p)^{1-y_i^*} \right] \\ L(p|Y_1^*,Y_2^*,...Y_n^*) &= p^{\sum_{i=1}^n y_i^*} (1-p)^{\sum_{i=1}^n 1-y_i^*} \\ L(p|Y_1^*,Y_2^*,...Y_n^*) &= p^{\sum_{i=1}^n y_i^*} (1-p)^{n-\sum_{i=1}^n y_i^*} \\ L(p|Y_i^*) &= ln(L(p|Y_1^*,Y_2^*,...Y_n^*)) \\ l(p|Y_i^*) &= \sum_{i=1}^n y_i^* ln(p) + (n-\sum_{i=1}^n y_i^*)(ln(1-p)) \end{split}$$

#### 0.2 Maximum Likelihood Estimation

$$U(p) = \frac{\delta l(p|Y_i^*)}{\delta p}$$

$$U(p) = \frac{\delta}{\delta p} \left[ \sum_{i=1}^n y_i^* ln(p) + (n - \sum_{i=1}^n y_i^*) (ln(1-p)) \right]$$

$$U(p) = \frac{\sum_{i=1}^n y_i^*}{p} - \frac{n - \sum_{i=1}^n y_i^*}{1-p}$$

Set U(p) = 0 and solve for p:

$$0 = \frac{\sum_{i=1}^{n} y_{i}^{*}}{p} - \frac{n - \sum_{i=1}^{n} y_{i}^{*}}{1 - p}$$

$$\frac{\sum_{i=1}^{n} y_{i}^{*}}{p} = \frac{n - \sum_{i=1}^{n} y_{i}^{*}}{1 - p}$$

$$\frac{1 - p}{p} = \frac{n - \sum_{i=1}^{n} y_{i}^{*}}{\sum_{i=1}^{n} y_{i}^{*}}$$

$$\frac{1}{p} - \frac{p}{p} = \frac{n}{\sum_{i=1}^{n} y_{i}^{*}} - \frac{\sum_{i=1}^{n} y_{i}^{*}}{\sum_{i=1}^{n} y_{i}^{*}}$$

$$\frac{1}{p} = \frac{n}{\sum_{i=1}^{n} y_{i}^{*}}$$

$$p = \frac{\sum_{i=1}^{n} y_{i}^{*}}{n}$$

From a poisson distribution:

$$P(Y = 1) = 1 - P(Y = 0) = 1 - f_y(0)$$

$$f_Y(0) = \frac{\lambda^0 e^{-\lambda}}{0!}$$

$$1 - f_Y(0) = 1 - e^{-\lambda}$$

$$p = 1 - e^{-\lambda}$$

Hence:

$$1 - e^{-\lambda} = \frac{\sum_{i=1}^{n} y_i^*}{n}$$
$$1 - \frac{\sum_{i=1}^{n} y_i^*}{n} = e^{-\lambda}$$
$$\ln\left(1 - \frac{\sum_{i=1}^{n} y_i^*}{n}\right) = -\lambda$$
$$\hat{\lambda}_{MLE} = -\ln\left(1 - \frac{\sum_{i=1}^{n} y_i^*}{n}\right)$$

### **0.3** Information about $\lambda$

$$\begin{split} I(p) &= Var[U(\lambda)] = E[U(\lambda)^2] = -E[U'(\lambda)] \\ I(p) &= -E\left[\frac{\delta}{\delta p} \frac{\sum_{i=1}^n y_i^*}{p} - \frac{n - \sum_{i=1}^n y_i^*}{1 - p}\right] \\ I(p) &= E\left[\frac{\sum_{i=1}^n y_i^*}{p^2} + \frac{n - \sum_{i=1}^n y_i^*}{(1 - p)^2}\right] \\ I(p) &= \frac{np}{p^2} + \frac{n - np}{(1 - p)^2} \\ I(p) &= \frac{n}{p} + \frac{n(1 - p)}{(1 - p)^2} \\ I(p) &= \frac{n}{p} + \frac{n}{(1 - p)} \\ I(p) &= \frac{n}{p} + \frac{n}{(1 - p)} \\ I(\lambda) &= \frac{n}{(1 - e^{-\lambda})(1 - (1 - e^{-\lambda}))} \\ I(\lambda) &= \frac{n}{(1 - e^{-\lambda})(-e^{-\lambda})} \end{split}$$

## 0.4 Using the invariance property

Due to the invariance properties of MLE's, which implies that we know the mean of any function of  $\lambda$ . So  $\tau(\hat{\lambda}_{MLE})$  is the MLE of  $\tau(\lambda)$ .

$$P(Y = 1) = 1 - P(Y = 0) = 1 - f_y(0)$$
$$f_Y(0) = \frac{\lambda^0 e^{-\lambda}}{0!}$$
$$1 - f_Y(0) = 1 - e^{-\lambda}$$
$$p = 1 - e^{-\lambda}$$

Hence:

$$1 - e^{-\lambda} = \frac{\sum_{i=1}^{n} y_i^*}{n}$$

$$1 - \frac{\sum_{i=1}^{n} y_i^*}{n} = e^{-\lambda}$$

$$ln\left(1 - \frac{\sum_{i=1}^{n} y_i^*}{n}\right) = -\lambda$$

$$\hat{\lambda}_{MLE} = -ln\left(1 - \frac{\sum_{i=1}^{n} y_i^*}{n}\right)$$

$$\hat{\lambda}_{MLE} = -ln(1 - \bar{y})$$

Where  $\bar{y}$  is the mean of  $Y_i^* \sim bernoulli(p)$