# stochastics and probability

Lecture 2

Dr. Johannes Pahlke

### random variable

#### example:

 $\Omega$ 

 $\mathbb{B}$ 

#### random variables

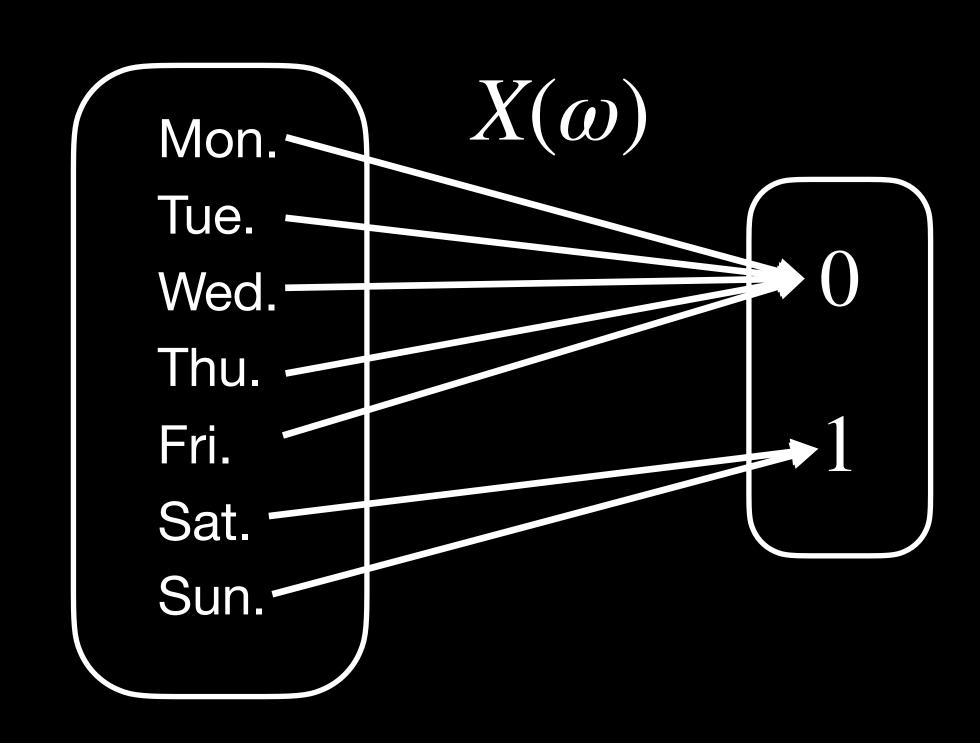
$$X:\Omega o J$$

$$X(\omega) = x$$

$$\omega \in \Omega, \quad x \in J$$

event:  $A \subseteq \Omega$ 

$$P(X > x_i) := P(\{\omega \in \Omega : X(\omega) > x_i\})$$



range of X:  $Range(X) := \{x \in \mathbb{R} : x = X(\omega), \omega \in \Omega\}$ 

## probability mass function

$$p_X(x) = P(X = x)$$
  $p: J \to [0,1]$ 

$$\sum_{x \in J} p_X(x) = 1$$

## exercise: probability mass function

$$X: \Omega \to J$$
 
$$P(\Omega) = 1$$
 
$$P(X > x_i) := P(\{\omega \in \Omega : X(\omega) > x_i\}) \qquad P(A \cup B) = P(A) + P(B) \quad \text{if } P(A \cap B) = 0$$
 
$$p_X(x) = P(X = x)$$

proof: 
$$\sum_{x \in J} p_X(x) = \sum_{x \in J} P(X = x) = \sum_{x \in J} P(\{\omega \in \Omega : X(\omega) = x\})$$
$$= P\left(\bigcup_{x \in J} \{\omega \in \Omega : X(\omega) = x\}\right) = P(\Omega) = 1$$

## probability density function (PDF)

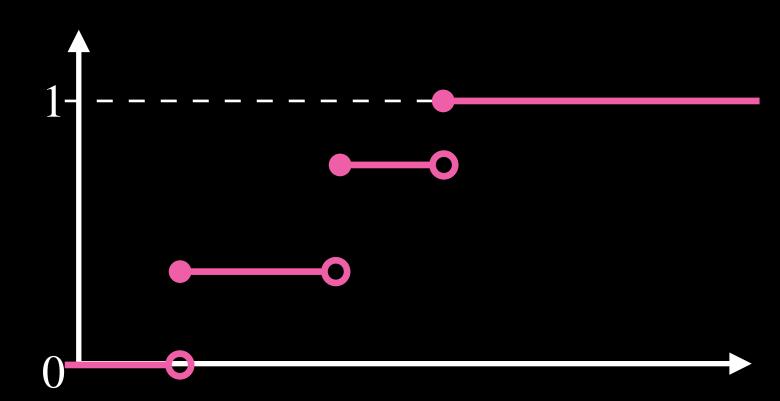
$$P(X = x) = 0 f: \mathbb{R} \to [0, \infty]$$

$$P(a \le X \le b) = \int_{a}^{b} f(x) \ dx$$

## cumulative distribution function (CDF)

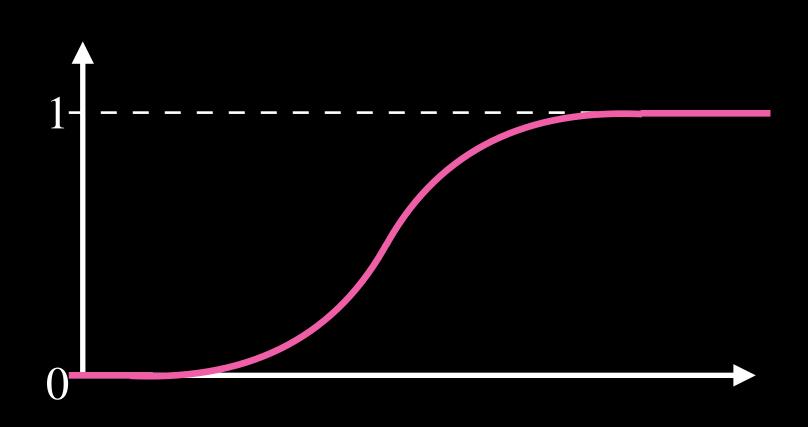
#### discrete

$$F_X(x) = P(X \le x) = \sum_{t \le x} P(X = t) = \sum_{t \le x} p_X(t)$$



#### continuous

$$F_X(x) = P(X \le x) = \int_{-\infty}^{x} f_X(t) dt$$



## exercice

## $F_X(x) = P(X \le x) = \int_{-\infty}^{x} f(t) dt$

How to calculate the PDF  $f_X(x)$  given the CDF  $F_X(x)$ ?

$$f_X(x) = \frac{d}{dx} F_X(x)$$

 $Range(X) := \{x \in \mathbb{R} : x = X(\omega), \omega \in \Omega\}$ 

What is the range of X?

$$X:\Omega o \mathbb{R}$$

$$X: \Omega \to \mathbb{R}$$
  $\Omega := [0cm, 100cm]$ 

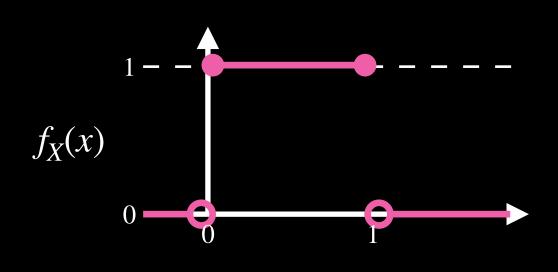
$$X(\alpha \text{cm}) = \frac{\alpha}{100}$$

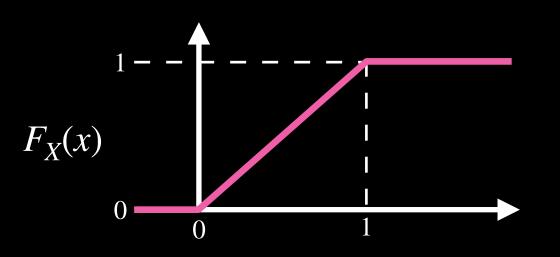
$$Range(X) = [0,1]$$

What is  $F_X(x)$  of the uniform distribution  $\mathcal{U}(0,1)$ ?

$$f_X(x) = \begin{cases} 1 & \text{for } 0 \le x \le 1 \\ 0 & \text{else} \end{cases}$$

$$f_X(x) = \begin{cases} 1 & \text{for } 0 \le x \le 1 \\ 0 & \text{else} \end{cases} \qquad F_X(x) = \begin{cases} 0 & \text{for } x < 0 \\ x & \text{for } 0 \le x \le 1 \\ 1 & \text{else} \end{cases}$$





## uniform random number generation

```
data-driven { not statistical random not reproducable } 
algorithmic { statistical random reproducable } 
}
```

random numbers  $x_1, \ldots, x_n \in [0,1]$ 

$$\hat{F}_U(x) := \frac{\left| \left\{ x_i \in \{x_1, \dots, x_n\} : x_i \le x \right\} \right|}{n}$$

require 
$$\lim_{n \to \infty} \left| \hat{F}_U(x) - F_U(x) \right| = \lim_{n \to \infty} \left| \hat{F}_U(x) - x \right| = 0$$

## pseudo-random number generation

Lehmer generator /a Linear congruential generator (LCG)

$$z_{i+1} = az_i \mod m$$
  $i = 1,2,3,...$   $x_i = \frac{z_i}{m} \in [0,1)$ 

Need to choose:

$$m \in \mathbb{N}_{>0}$$
  $0 < a < m$   
seed:  $z_1 \in \mathbb{N}_{>0}$ 

$$m = 2^{31} - 1$$
 $a = 48271$ 

cycle length:

$$x_1, x_2, \dots, x_T$$
  $x_1 = x_T$ 

## pseudo-random number generation

generators with lager cycle length T

Mersenne Twister (1998)

Park-Miller (1988)

XOR-Shift (2003)

XoroShiro (2018)

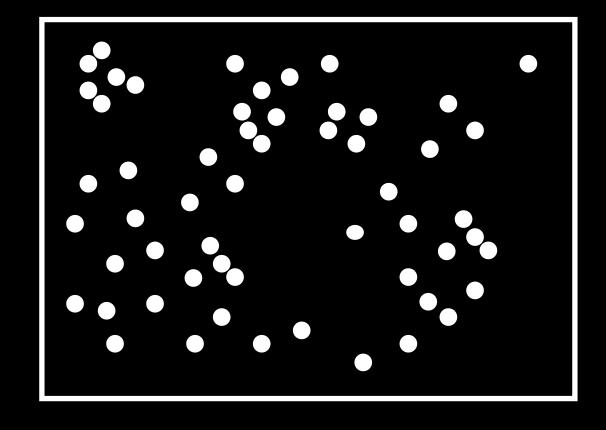
parallel random number generators

$$z_1(\text{proc id})$$

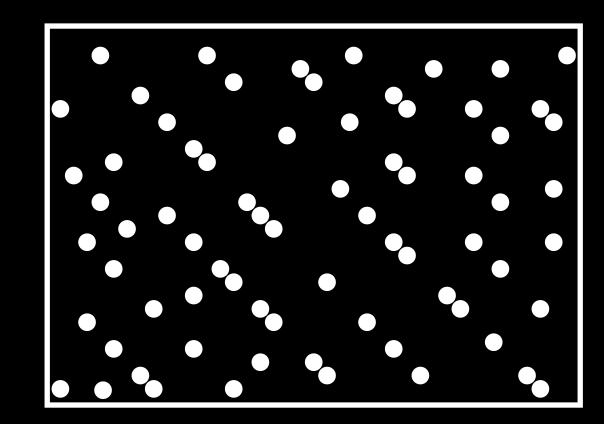
$$T_{parallel} = \frac{T}{\text{#proc}}$$

## quasi-RNG (low discrepancy sequence)

#### pseudo-RNG



#### quasi-RNG



Additive recurrence sequence:

$$x_{i+1} = (x_i + \alpha) \mod 1$$

$$\alpha$$
 irrational  $\alpha = \sqrt{2} - 1$   $\alpha = \frac{\sqrt{5} - 1}{2}$ 

more modern: Sobol - sequence

### transform random variables

$$Y = g(X)$$
  $g: J \to J'$   $Domain(g) := J$   $g: \mathbb{R} \to \mathbb{R}$ 

 $Range(X) \subseteq Domain(g)$ 

inverse 
$$g^{-1}(y) := \{x \in J : g(x) = y\}$$

### transform discrete random variables

given 
$$p_X$$
,  $Y = g(X)$ 

 $g^{-1}(y) := \{x \in J : g(x) = y\}$ 

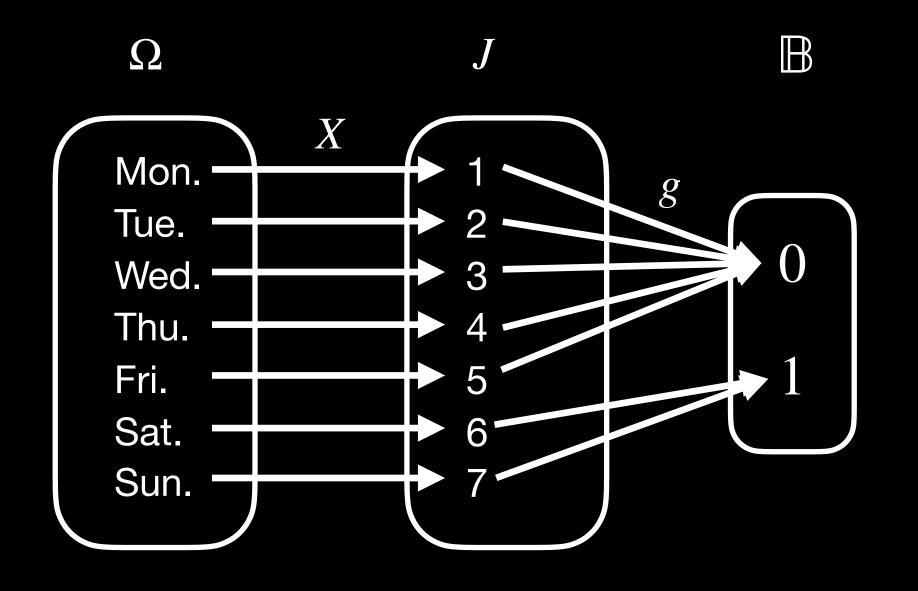
 $P(A \cup B) = P(A) + P(B) \quad \text{if} \quad P(A \cap B) = 0$ 

What is  $p_Y$ ?

$$p_Y(y) = P(Y = y) = P(g(X) = y) = P(X \in g^{-1}(y))$$

$$= P\left(\bigcup_{x \in g^{-1}(y)} \{X = x\}\right) = \sum_{x \in g^{-1}(y)} p_X(x)$$

### exercise: transform discrete random variables



Mon. 
$$Y$$
Tue.  $0$ 
Thu. Fri. Sat.  $g(X)$ 

$$g(x) = \begin{cases} 0 & \text{for } x \in \{1, ..., 5\} \\ 1 & \text{for } x \in \{6, 7\} \end{cases}$$

$$p_{Y}(y) = \sum_{x \in g^{-1}(y)} p_{X}(x)$$
$$g^{-1}(y) := \{x \in J : g(x) = y\}$$

$$p_X(x) = \frac{1}{7} \qquad x \in \{1, ..., 7\}$$

derive: 
$$p_Y(1) = \sum_{x \in g^{-1}(1)} p_X(x) = \sum_{x \in \{6,7\}} p_X(x) = p_X(6) + p_X(7) = \frac{2}{7}$$

### transform continuous random variables

given 
$$f_X$$
,  $Y = g(X)$ 

with g increasing in Range(X)  $\rightarrow g^{-1}$  is a function \* not the full story

What is  $f_Y$ ?

$$F_Y(y) = P(Y \le y) = P(g(X) \le y) = P(X \le g^{-1}(y))$$
  
=  $F_X(g^{-1}(y))$ 

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(g^{-1}(y)) = f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

### transform continuous random variables

$$f_X$$

given 
$$f_X$$
,  $Y = g(X)$ 

with g decreasing in Range(X)  $\rightarrow g^{-1}$  is a function \* not the full story

What is  $f_Y$ ?

$$F_Y(y) = P(Y \le y) = P(g(X) \le y) = P(X \ge g^{-1}(y))$$
$$= 1 - P(X \le g^{-1}(y)) = 1 - F_X(g^{-1}(y))$$

$$P(\overline{A}) = 1 - P(A)$$

$$f_Y(y) = \frac{d}{dy} F_Y(y) = -\frac{d}{dy} F_X(g^{-1}(y)) = -f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

### exercise: transform continuous random variables

$$X \sim \mathcal{U}(0,1)$$

$$f_X(x) = \begin{cases} 1 & \text{for } 0 \le x \le 1 \\ 0 & \text{else} \end{cases}$$

$$Y = g(X) = 1 - X$$

#### derive:

$$g^{-1}(y) = 1 - y$$

$$f_Y(y) = -f_X(1-y)\frac{d}{dy}(1-y) = -f_X(1-y)(-1)$$

$$= \begin{cases} 1 & \text{for } 0 \le 1-y \le 1 \\ 0 & \text{else} \end{cases} = \begin{cases} 1 & \text{for } 0 \le y \le 1 \\ 0 & \text{else} \end{cases}$$

#### increasing g

$$f_Y(y) = f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

#### decreasing g

$$f_Y(y) = -f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

### inversion transform

$$Y = g(X) := F_X(X) \qquad F_X : \mathbb{R} \to [0,1]$$

$$F_Y(y) = P(Y \le y)$$

$$= P(F_X(X) \le y)$$

$$= P(X \le F_X^{-1}(y))$$

$$= F_X(F_X^{-1}(y))$$

$$= y \qquad \longrightarrow Y \sim \mathcal{U}(0,1)$$

$$U \sim \mathcal{U}(0,1)$$

$$X = F_X^{-1}(U) \longrightarrow X \text{ has the CDF } F_X(x)$$

### non-uniform distributions

exponential distribution  $Exp(\lambda)$ 

$$y = F_X(x) = 1 - e^{-\lambda x} \qquad x \ge 0$$

$$e^{-\lambda x} = 1 - y$$

$$-\lambda x = \ln(1 - y)$$

$$x = -\frac{1}{\lambda} \ln(1 - y)$$

$$Y \sim \mathcal{U}(0, 1) \qquad \longrightarrow \qquad U := 1 - Y \sim \mathcal{U}(0, 1)$$

$$\longrightarrow \qquad X \sim Exp(\lambda)$$

### non-uniform distributions

standard Gaussian/ normal distribution  $\mathcal{N}(0,1)$ 

Box-Muller Transform

$$X_1, X_2 \sim \mathcal{N}(0,1)$$

$$R^2 = X_1^2 + X_2^2 \sim Exp\left(\frac{1}{2}\right)$$

$$\Theta \sim \mathcal{U}(0,2\pi)$$

exercise: write  $X_1, X_2$  using  $U_1, U_2 \sim \mathcal{U}(0,1)$ 

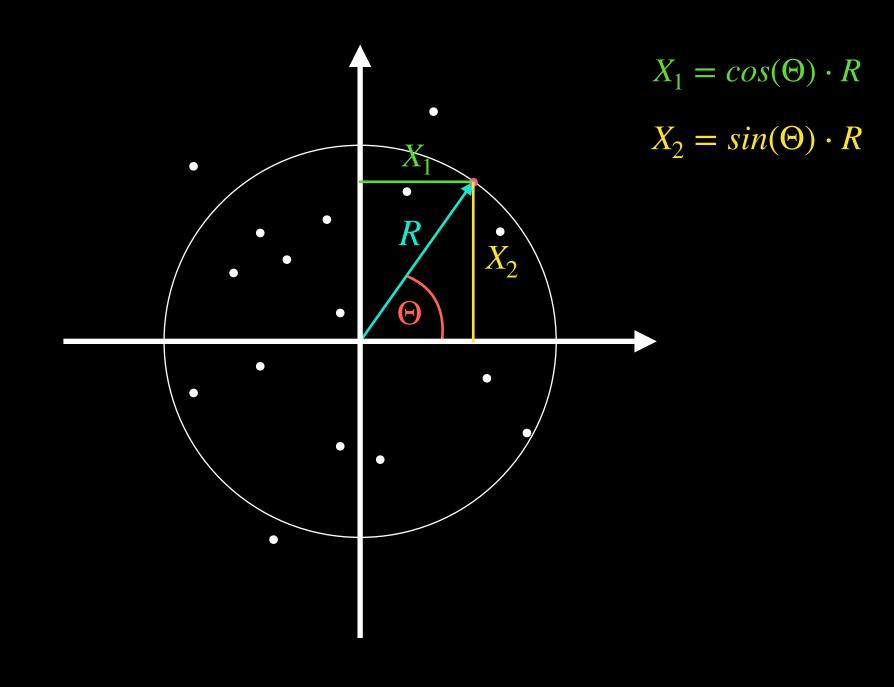
$$R^2 = -2\ln(U_1)$$

$$\Theta = 2\pi U_2$$

$$X_1 = \cos(2\pi U_2) \cdot \sqrt{-2\ln(U_1)}$$

$$X_2 = \sin(2\pi U_2) \cdot \sqrt{-2\ln(U_1)}$$

$$X = -\frac{1}{\lambda} \ln(U) \qquad U \sim \mathcal{U}(0,1)$$
$$X \sim Exp(\lambda)$$



## end