

## ON THE DISTRIBUTION OF THE PRODUCT OF INDEPENDENT BETA RANDOM VARIABLES

Jen TANG

*Marquette University, Milwaukee, WI 53233, USA*

A.K. GUPTA

*Bowling Green State University, Bowling Green, OH 43403, USA*

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*Abstract:* In this paper, exact distribution of the product of independent beta random variables has been derived and its structural form is given together with recurrence relations for the coefficients of this representation. These recurrence relations yield a direct computational algorithm for computing the percentage points of many test criteria in multivariate statistical analysis.

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### 1. Introduction

Wilks (1932) has considered the following type-*B* integral equation:

$$\int_0^1 w^h g_p(w) dw = \prod_{j=1}^p \frac{\Gamma[c(j)] \Gamma[b(j) + h]}{\Gamma[b(j)] \Gamma[c(j) + h]} \quad (h = 0, 1, 2, \dots), \quad (1.1)$$

where  $g_p(w)$  is unknown and independent of  $h$ ,  $b(j)$  and  $c(j)$  are real and positive such that  $c(j) > b(j)$ , for  $j = 1, 2, \dots, p$ , and  $\Gamma[\cdot]$  is the gamma function. Evidently,  $g_p(w)$  is the density function of  $W_p \cong \prod_{j=1}^p t_j$  say, where  $\mathcal{L}(t_j) = \text{Beta}[b(j), c(j) - b(j)]$ ,  $j = 1, \dots, p$ , are independent. Here we define the Beta[ $m, n$ ] density as  $B^{-1}[m, n] u^{m-1}(1-u)^{n-1}$ ,  $0 < u < 1$ .

The distribution of  $W_p$ , known as Wilks' type-*B* distribution is very useful in multivariate statistical hypothesis testing. In fact, many likelihood ratio test criteria of multivariate hypotheses can be expressed as such products; for example, (i) general linear hypothesis (see Anderson, 1958, p. 187), (ii) equality of several mean vectors with same covariance matrix (see Anderson, 1958, p. 211); that is, the MANOVA hypothesis, (iii) independence between every pair of sets of variables (Anderson, 1958, p. 237), (iv) the sphericity hypothesis (Anderson, 1958, p. 259), (v) equality of covariance matrices (Anderson, 1958, p. 247).

Wilks (1932) has expressed  $g_p(w)$  in the form of a  $(p-1)$ -fold multiple integral, which he was able to evaluate for some special cases only. Using the Fourier inversion theorem, Wald and Brookner (1941) have obtained the structural form for  $g_p(w)$  as a mixture of beta functions. Recently, Gupta (1977), Walster and Trotter (1980) and Nandi (1980) have obtained some formulas as  $(p-1)$ -fold series for computing the

coefficients of the mixture representation. Schatzoff (1966), Pillai and Gupta (1969), and Gupta (1971) have applied convolution technique to obtain expressions for the MANOVA model for small  $p$ , which Khatri and Bhargava (1981) have modified to obtain  $g_p(w)$  in the form of a  $p$ -fold multiple series. Consul (1968) has used the inverse Mellin transform and some special properties of hypergeometric functions to derive the exact distributions of  $W_p$  for a limited number of cases in various problems. Using contour integration, Mathai (1971) has obtained the pdf of  $W_p$  in series form. Mathai (1973) has written a review of different methods of obtaining exact distributions of multivariate test criteria and pointed out shortcomings of these methods.

In this paper, we obtain the exact distribution of  $W_p$  by deriving both the structural form for  $g_p(w)$  and *simple recurrence relations* for the coefficients of this representation. Recurrence relations yield a direct computational algorithm for computing the percentage points for many test criteria, hence our results are very useful from theoretical as well as practical points of view.

## 2. Distribution of product of independent beta random variables

We state the main result of this paper as follows.

**Theorem 2.1.** *The exact density function of  $W_p$  is given by*

$$g_p(w) = K_p w^{b(p)-1} (1-w)^{f(p)-1} \sum_{r=0}^{\infty} \sigma_r^{(p)} (1-w)^r, \quad 0 < w < 1, \quad (2.1)$$

where  $K_p, f(p)$  are defined by

$$K_m = \prod_{j=1}^m \{ \Gamma[c(j)] / \Gamma[b(j)] \}, \quad f(m) = \sum_{j=1}^m [c(j) - b(j)],$$

for  $1 \leq m \leq p$ , and  $\sigma_r^{(p)}$  can be obtained from the following recurrence relation:

$$\sigma_r^{(k)} = \frac{\Gamma[f(k-1)+r]}{\Gamma[f(k)+r]} \sum_{s=0}^r [(c(k) - b(k-1))_s / s!] \sigma_{r-s}^{(k-1)} \quad \text{for } r = 0, 1, 2, \dots, \quad k = 2, 3, \dots, p, \quad (2.2)$$

with initial values

$$\sigma_0^{(1)} = \Gamma^{-1}[f(1)], \quad \sigma_r^{(1)} = 0 \quad \text{for } r = 1, 2, \dots$$

**Proof.** For  $2 \leq k \leq p$ ,  $p$  fixed, write

$$W_{k,p} = \prod_{j=1}^k t_j = \left( \prod_{j=1}^{k-1} t_j \right) t_k = W_{k-1,p} t_k$$

say, where  $W_{p,p} = W_p$  and we assume that the density of  $W_{k-1,p}$  can be written as

$$g_{k-1,p}(w) = K_{k-1} w^{b(k-1)-1} (1-w)^{f(k-1)-1} \sum_{r=0}^{\infty} \sigma_r^{(k-1)} (1-w)^r.$$

Using the usual transformation technique for products and Euler's formula for the hypergeometric function  ${}_2F_1$  (see Luke, 1969, p. 57), we can obtain the density of  $W_{k,p}$  as follows:

$$g_{k,p}(w) = K_k w^{b(k)-1} (1-w)^{f(k)-1} \sum_{r=0}^{\infty} \sigma_r^{(k-1)} \Gamma[f(k-1)+r] \\ \times \Gamma^{-1}[f(k)+r] (1-w)^r {}_2F_1[f(k-1)+r, c(k)-b(k-1); f(k)+r; 1-w]. \quad (2.3)$$

Since  ${}_2F_1$  is a power series in  $1-w$ , we can rewrite (2.3) by combining powers of  $1-w$  as

$$g_{k,p}(w) = K_k w^{b(k)-1} (1-w)^{f(k)-1} \sum_{r=0}^{\infty} \sigma_r^{(k)} (1-w)^r, \quad (2.4)$$

where  $g_{p,p} = g_p$  and  $\sigma_r^{(k)}$ ,  $r = 0, 1, 2, \dots$ , clearly satisfy (2.2) for  $2 \leq k \leq p$ .

Note that both (2.1) and (2.3) with  $k = p$  are representations for the density of  $W_p$ . It is easily seen from (2.2) that  $\sigma_r^{(p)}$ ,  $r = 0, 1, \dots, m$ , can be computed from  $\sigma_r^{(k)}$ ,  $r = 0, 1, 2, \dots, m$ ,  $k = 1, 2, \dots, p-1$ . Since  $b(j)$  and  $c(j)$  may depend on  $p$ ,  $g_{k,p}(w)$  given in (2.3) may not be the density of  $W_k$  unless  $b(j)$  and  $c(j)$  are independent of  $p$ . (2.2) provides formulas for computing distributions of  $W_k$ ,  $k = 1, 2, 3, \dots$ , successively and its pdf is given in (2.4).

The following corollary, which gives the exact density of the product of two independent beta variables, is an immediate consequence of Theorem 2.1.

**Corollary 2.1.** *The exact density of  $W_2$  is given by*

$$g_2(w) = K_2 \Gamma^{-1}[f(2)] w^{b(2)-1} (1-w)^{f(2)-1} {}_2F_1[f(1), c(2)-b(1); f(2); 1-w] \quad \text{for } 0 < w < 1. \quad (2.5)$$

Note that many of Consul's results can easily be obtained from this corollary and simple transformations with the help of duplication formula for gamma functions (see Srivastava and Khatri, 1979, Chapter 7).

In many applications, it is possible to evaluate  $\sigma_r^{(p)}$  using the fact that  $\int_0^1 g_p(w) dw = 1$ . From (2.1) we obtain

$$K_p^{-1} = \sum_{r=0}^{\infty} \sigma_r^{(p)} \Gamma[b(p)] \Gamma[f(p)+r] / \Gamma[b(p)+f(p)+r]. \quad (2.6)$$

Both sides of (2.6) can, in general, be expressed as power series in  $n^{-1} = (\text{sample size})^{-1}$ . Upon equating powers of  $n^{-1}$ , we can obtain different recurrence relations for  $\sigma_r^{(p)}$ ,  $r = 0, 1, 2, \dots$  (see Tang, 1983).

### 3. Concluding remarks

The cumulative distribution function of  $W_p$ , obtained by integrating (2.1) term by term, is a mixture of incomplete beta functions which can also be evaluated recursively. The exact null pdf's of test criteria mentioned in Section 1 can be obtained by simple substitution of the parameters. Hence our results are useful for obtaining exact distributions and exact percentage points of many test criteria in multivariate hypothesis testing.

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