

S722 HW6

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To save on typing, I will denote $\frac{\partial^k}{\partial x^k} f(x) = \partial_x^k f(x)$.

Part 1

1.33

$$\begin{aligned} P(M|CB) &= \frac{P(CB|M)P(M)}{P(CB|M)P(M)+P(CB|F)P(F)} \\ &= \frac{(.05)(.5)}{(.05)(.5)+(.0025)(.5)} \\ &\approx 0.952 \end{aligned}$$

2.11

a

From example 2.1.7, $f_Y(y) = \frac{1}{2\sqrt{y}}(f_X(\sqrt{y}) + f_X(-\sqrt{y})) = (2\pi y)^{-1/2}e^{-y/2}$.

Then $E[Y] = \int_0^\infty \sqrt{\frac{y}{2\pi}} e^{-y/2} dy$.

Let $u = \sqrt{y}$ and $dv = e^{-y/2} dy$. Then $du = \frac{1}{2}y^{-1/2}$ and $v = -2e^{-y/2}$.

Then the integral becomes $(2\pi)^{-1/2} \int_0^\infty y^{-1/2} e^{-y/2} dy$.

Let $u = \sqrt{y} \implies du = \frac{1}{2}y^{-1/2} dy$. Then the integral becomes $(2\pi)^{-1/2} \int_0^\infty 2e^{-u^2/2} du = (2\pi)^{-1/2} (2)(\pi/2)^{1/2} = 1$.

b

$$\begin{aligned} Y = |X| &\implies X = \pm Y \implies |X'| = 1 \\ \implies f_Y(y) &= f_X(y) + f_X(-y) = \sqrt{2/\pi} e^{-y^2/2}. \end{aligned}$$

$$\begin{aligned} E[Y] &= \int_0^\infty y(2\pi)^{1/2} e^{-y^2/2} dy \\ u = y^2/2 &\implies du = y dy \text{ so the above becomes} \\ &= (2/\pi)^{1/2} \int_0^\infty e^{-u} du = (2/\pi)^{1/2}. \end{aligned}$$

$$\begin{aligned} E[Y^2] &= (2/\pi)^{1/2} \int_0^\infty y^2 e^{-y^2/2} dy \\ u = y \text{ and } dv &= y e^{-y^2/2} \implies du = dy \text{ and } v = -e^{-y^2/2} \text{ so the above becomes} \\ &= \sqrt{2/\pi} \int_0^\infty e^{-y^2/2} dy = 1. \end{aligned}$$

$$\text{So } \text{Var}(Y) = E[Y^2] - (E[Y])^2 = 1 - \frac{2}{\pi}.$$

2.33

a

$$M_X(t) = E[e^{tX}] = \sum_x e^{tx} \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum \frac{(e^t \lambda)^x}{x!} = e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t - 1)}$$

$$E[X] = M'_X(t=0) = e^{\lambda(e^t - 1)} \lambda e^t|_{t=0} = e^{\lambda(1-1)} \lambda e^0 = \lambda$$

$$E[X^2] = M_X''(t=0) = \lambda e^t e^{\lambda(e^t-1)} \lambda e^t + \lambda e^t e^{\lambda(e^t-1)}|_{t=0} = \lambda^2 + \lambda$$

$$Var(X) = E[X^2] - (E[X])^2 = \lambda$$

$$\phi_X(t) = M_X(it) = e^{\lambda(e^{it}-1)}$$

$$E[X] = -i\phi_X'(0) = -ie^{\lambda(e^{it}-1)} \lambda i e^{it}|_{t=0} = (-1)i^2 \lambda = \lambda$$

$$\begin{aligned} E[X^2] &= -\phi_X''(0) = -i\lambda(e^{\lambda(e^{it}-1)} i e^{it} \lambda + e^{\lambda(e^{it}-1)} i e^{it})|_{t=0} \\ &= -i\lambda(i\lambda + i) = \lambda^2 + \lambda \end{aligned}$$

$$Var(X) = E[X^2] - (E[X])^2 = \lambda$$

c

$$\begin{aligned} M_X(t) &= E[e^{tX}] = \int e^{tx} (2\pi\sigma^2)^{-1/2} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \\ &= (2\pi\sigma^2)^{-1/2} \int e^{-\frac{1}{2\sigma^2}(x^2-2\mu x+\mu^2-2\sigma^2 tx)} dx \\ &= (2\pi\sigma^2)^{-1/2} e^{-\frac{\mu^2}{2\sigma^2}} \int e^{-\frac{1}{2\sigma^2}(x^2-2(\mu+\sigma^2 t)x+(\mu+\sigma^2 t)^2-(\mu+\sigma^2 t)^2)} dx \\ &= (2\pi\sigma^2)^{-1/2} e^{-\frac{\mu^2}{2\sigma^2}} e^{\frac{(\mu+\sigma^2 t)^2}{2\sigma^2}} \int e^{-\frac{1}{2\sigma^2}(x-(\mu+\sigma^2 t))^2} dx \\ &= (2\pi\sigma^2)^{-1/2} (2\pi\sigma^2)^{1/2} e^{-\frac{\mu^2}{2\sigma^2} + \frac{\mu^2}{2\sigma^2} + \frac{\sigma^4 t^2}{2\sigma^2} + \frac{2\mu\sigma^2 t}{2\sigma^2}} \\ &= e^{\frac{\sigma^2 t^2}{2} + \mu t} \end{aligned}$$

$$E[X] = M_X'(0) = e^{\sigma^2 t^2/2 + \mu t} (\sigma^2 t + \mu)|_{t=0} = \mu$$

$$E[X^2] = M_X''(0) = e^{\sigma^2 t^2/2 + \mu t} (\sigma^2 t + \mu)^2 + e^{\sigma^2 t^2/2 + \mu t} \sigma^2|_{t=0} = \mu^2 + \sigma^2$$

$$Var(X) = E[X^2] - (E[X])^2 = \sigma^2$$

$$\phi_X(t) = M_X(it) = e^{-\sigma^2 t^2/2 + i\mu t}$$

$$E[X] = -i\phi_X'(0) = -ie^{-\sigma^2 t^2/2 + i\mu t} (-\sigma^2 t + i\mu)|_{t=0} = -i(i\mu) = \mu$$

$$E[X^2] = -\phi_X''(0) = -e^{-\sigma^2 t^2/2 + i\mu t} ((-\sigma^2 t + i\mu)^2 - \sigma^2)|_{t=0} = \mu^2 + \sigma^2$$

$$Var(X) = \mu^2 + \sigma^2 - \mu^2 = \sigma^2$$

2.38

a

$$\begin{aligned} M_X(t) &= E[e^{tX}] = \sum_{x=0}^{\infty} e^{tx} \binom{r+x-1}{x} p^r (1-p)^x \\ &= p^r \sum_x \binom{r+x-1}{x} (e^t (1-p))^x \\ &= \frac{p^r}{(1-e^t(1-p))^r} \end{aligned}$$

b

$$M_Y(t) = M_X(2pt) = \left(\frac{p}{1-e^{2pt}(1-p)} \right)^r$$

$$\begin{aligned} \lim_{p \rightarrow 0} \frac{p}{1-e^{2pt}(1-p)} &= \lim_{p \rightarrow 0} \frac{1}{e^{2pt} - 2te^{2pt}(1-p)} = \frac{1}{1-2t} \\ \implies \lim_{p \rightarrow 0} \left(\frac{p}{1-e^{2pt}(1-p)} \right)^r &= \left(\frac{1}{1-2t} \right)^r \end{aligned}$$

3.24

a

$$\begin{aligned} Y = X^{1/\gamma} &\implies X = Y^\gamma \implies X' = \gamma Y^{\gamma-1} \\ &\implies f_Y(y) = f_X(y^\gamma) \gamma y^{\gamma-1} = \frac{\gamma}{\beta} e^{-y^\gamma/\beta} y^{\gamma-1} \end{aligned}$$

b

$$\begin{aligned} Y = (2X/\beta)^{1/2} &\implies X = \beta Y^2/2 \implies X' = \beta Y \\ &\implies f_Y(y) = y e^{-y^2/2} \end{aligned}$$

c

$$\begin{aligned} Y = 1/X &\implies X = 1/Y \implies |X'| = Y^{-2} \\ &\implies f_Y(y) = \frac{1}{\Gamma(a)b^a} y^{-(a+1)} e^{-\frac{1}{by}} \end{aligned}$$

d

$$\begin{aligned} Y = (X/\beta)^{1/2} &\implies X = \beta Y^2 \implies X' = \beta Y/2 \\ &\implies f_Y(y) = \frac{2}{\Gamma(3/2)} y^2 e^{y^2} \end{aligned}$$

e

$$\begin{aligned} Y = \alpha - \gamma \log X &\implies X = \exp(\frac{\alpha-Y}{\gamma}) \implies |X'| = \frac{1}{\gamma} \exp(\frac{\alpha-Y}{\gamma}) \\ &\implies f_Y(y) = \frac{1}{\gamma} \exp(-e^{\frac{\alpha-y}{\gamma}} + \frac{\alpha-y}{\gamma}) \end{aligned}$$

3.48

$$\begin{aligned} f(x+1) &= \binom{n}{x+1} p^{x+1} (1-p)^{n-x-1} \\ &= \binom{n}{x} \frac{n-x}{x+1} \frac{p}{1-p} p^x (1-p)^{n-x} \\ &= \frac{n-x}{x+1} \frac{p}{1-p} f(x) \end{aligned}$$

3.49

a

$$\begin{aligned} E[g(X)(X - \alpha\beta)] &= \int_0^\infty g(x)(x - \alpha\beta) \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} dx \\ u = g(x) \text{ and } dv &= (x - \alpha\beta) x^{\alpha-1} e^{-x/\beta} e^{-x/\beta} dx \implies du = g'(x) dx \text{ and } v = -\beta x^\alpha e^{-x/\beta} \text{ so the above becomes} \\ &\frac{1}{\Gamma(\alpha)\beta^\alpha} \beta \int g'(x) x^\alpha e^{-x/\beta} dx \\ &= \beta \int x g'(x) f(x) dx = \beta E[X g'(X)] \end{aligned}$$

4.4

a

$$1 = \int_0^1 \int_0^2 C(x+2y) dx dy = 4C \implies C = 1/4$$

b

$$f_X(x) = \int_0^1 \frac{1}{4}(x+2y) dy = \frac{x+1}{4}$$

c

$$F(x, y) = \int_0^x \int_0^y \frac{s+2t}{4} ds dt = \frac{x^2 y + 2y^2 x}{8}$$

d

$$\begin{aligned} Z = 9/(X+1)^2 &\implies X = 3Z^{-1/2} - 1 \implies |X'| = \frac{3}{2} Z^{-3/2} \\ \implies f_Z(z) &= \frac{1}{4} (3z^{-1/2} - 1 + 1) \frac{3}{2} z^{-3/2} = \frac{9}{8} z^{-2} \end{aligned}$$

4.40

a

In the case where a , b , and c are natural numbers ...

$$\begin{aligned} \int \int x^{a-1} y^{b-1} (1-x-y)^{c-1} dx dy &= \Gamma(a) \int y^{b-1} (1-y)^{c+a-1} dy = \frac{\Gamma(a)\Gamma(b)\Gamma(c)}{\Gamma(a+b+c)} \\ \implies C &= \frac{\Gamma(a+b+c)}{\Gamma(a)\Gamma(b)\Gamma(c)} \end{aligned}$$

b

By symmetry, we only have to show this for one, say, X .

$$\begin{aligned} \text{Reusing some of the math from part (a), we get } f_X(x) &= \int_0^{1-x} f_{X,Y}(x, y) dy = \frac{\Gamma(a+b+c)}{\Gamma(a)\Gamma(b)\Gamma(c)} \Gamma(b) x^{a-1} (1-x)^{c+b-1} \\ \implies X &\sim \text{Beta}(a, b+c) \end{aligned}$$

c

$$\begin{aligned} f_{Y|X}(y) &\propto f_{X,Y}(x, y) \propto y^{b-1} (1-x-y)^{c-1} \\ &= y^{b-1} \sum_{i=0}^{c-1} (-y)^i (1-x)^{c-1-i} \\ &\propto y^{b-1} \sum_i (-y)^i = y^{b-1} (1-y)^{c-1} \\ \implies Y | X &\sim \text{Beta}(b, c) \end{aligned}$$

d

$$\begin{aligned} E[X, Y] &= \int \int \frac{\Gamma(a+b+c)}{\Gamma(a)\Gamma(b)\Gamma(c)} x^a y^b (1-x-y)^{c-1} dx \\ &= \frac{\Gamma(a+b+c)}{\Gamma(a)\Gamma(b)\Gamma(c)} \frac{\Gamma(a+1)\Gamma(b+1)\Gamma(c)}{\Gamma(a+b+c+2)} \\ &= \frac{ab}{(a+b+c+1)(a+b+c)} \end{aligned}$$

Part 2

1

$$E[e^{itX}] \leq E[(e^{itX})^2] = E[\sin^2 tX + \cos^2 tX] = E[1] = 1$$

2

$$\phi_{aX+b}(t) = E[e^{it(aX+b)}] = E[e^{itb}e^{itaX}] = e^{itb}E[e^{i(at)X}] = e^{itb}\phi_X(at)$$

3

$$\phi_{\sum_j X_j}(t) = E[e^{it\sum_j X_j}] = E[\prod_j e^{itX_j}] = \prod_j E[e^{itX_j}] = \prod_j \phi_{X_j}(t)$$

4

Let $\delta > 0$.

$$\begin{aligned} |\phi(t+\delta) - \phi(t)|^2 &= |E[e^{i(t+\delta)X} - e^{itX}]|^2 = |E[e^{itX}(e^{i\delta X} - 1)]|^2 \leq E[|e^{itX}(e^{i\delta X} - 1)|^2] = E[(e^{i\delta X} - 1)^2] \\ &\leq E[\delta^2 X^2] = \delta^2 E[X^2] \end{aligned}$$

So letting $\epsilon = \delta\sqrt{E[X^2]}$, we get $|\phi(t+\delta) - \phi(t)| \leq \epsilon$, assuming that the second moment exists.