

An assessment of the biodiversity - ecosystem function relationship in southern African woodlands

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1 Introduction

A large number of studies have shown relationships between biodiversity and ecosystem functionality (Liang et al., 2016; Cardinale et al., 2009). The strength and direction of these observed Biodiversity - Ecosystem Function Relationships (BEFRs) varies depending on the ecosystem being studied, the ecosystem function(s) of interest (Hector and Bagchi, 2007), and the inclusion of environmental covariates in statistical models (Vilà et al., 2005), but there appears to be a generalisable positive correlation between biodiversity and ecosystem functionality (Liang et al., 2016). Over the past decade, many observational studies of the BEFR have been conducted, mostly in tropical and temperate forests, and grasslands (Chen et al., 2011). These studies support early findings from small scale experimental studies which began in earnest during the 1990s as concern grew over the global loss of biodiversity (Tilman and Downing, 1994; Tilman et al., 2014).

Ecosystem functions can be defined in broad terms as rate processes and properties of ecosystems which describe the degree of biotic activity within an ecosystem (Jax, 2005). This includes basic processes of primary production such as gross primary productivity and atmospheric nitrogen fixation, but can be extended to indirect aggregate measures of function such as resistance of productivity to disturbance, and even to static properties which themselves influence process, such as trophic complexity and total vegetative biomass. The frequently reported and intuitive relationship between biodiversity and ecosystem functionality invokes three main mechanisms which drive the relationship (Tilman et al., 2014): 1) niche complementarity, whereby communities with greater diversity fill a greater breadth of realised niche space and avoid competition due to differences in their traits, 2) selection effects, whereby communities with greater diversity are more likely to include a species which contributes highly to the measured ecosystem function, and 3) facilitation effects, whereby communities with greater diversity are more likely to include combinations of species which together increase the others functional contribution.

The relationship between biodiversity and ecosystem functionality has informed conservation policy which aims to maintain and strengthen biodiversity under the assumption that this strengthens ecosystem resilience to stressors such as anthropogenic climate change and ecosystem service use by humans (Thompson et al., 2011). It is probable however, that the strength and relative importance of the BEFR varies with ecosystem studied and environmental conditions, with many ecosystems remaining understudied in the BEFR literature.

The representation of dry tropical ecosystems in the BEFR literature is poor compared to other ecosystems. Clarke et al. (2017) conducted a meta-analysis of 182 published BEFR studies, finding that only 13% were conducted in the tropics generally, with 42% of those being conducted in the wet tropical forests of Costa Rica, which are unique in terms of species assemblage, being an oceanic island with many endemic species. In wet tropical forests, much of the observed effect of biodiversity on ecosystem function can be attributed to niche complementarity (Wright et al., 2017; Poorter et al., 2015; van der Sande et al., 2017). In the dry tropics however, low precipitation, high seasonality of rainfall and temperature, and high levels of disturbance from fire and herbivory mean that niche complementarity and competition may not play as great a role in the relationship between biodiversity and ecosystem function. Instead, stress tolerance and the functional contribution of abundant species (selection effects) may be the predominant forces which influence ecosystem functionality (Lasky et al., 2014; Tobner et al., 2016). Similarly, more diverse species assemblages may lead to facilitation effects between certain species combinations in environments which are more hostile to growth (Ratcliffe et al., 2017).

Savannas and sparse woodlands are the dominant vegetation type across the southern African region, spanning >4 million km^2 (Ryan et al., 2016) (Figure 1a). The carbon stored in these woodlands is comparable to that found in the wet forests of the Congo basin and is of global importance to the carbon cycle (Houghton et al., 2009; Mayaux et al., 2008). Climatic conditions and biogeography vary across southern African woodlands, resulting in a diverse range of woodland tree species assemblages, which retain the common features of an open tree canopy and an understorey generally dominated by C4 grass species (Frost, 1996). Southern African woodlands are highly diverse, thought to harbour ~ 8500 plant species of which there are >300 tree species (Frost, 1996). Southern African woodlands have been identified by previous studies as a priority for conservation efforts (Brooks et al., 2006; Mittermeier et al., 2003). Many conservation projects in the region currently aim to conserve biodiversity and woody biomass stocks simultaneously under the directive of the United Nations REDD+ programme or the similar Forest Carbon Partnership Facility (FCPF) (Hinsley et al., 2015). Despite these efforts however, human actions are driving rapid changes in biodiversity. Southern African woodlands are also relied heavily upon for their ecosystem service provision, which is itself affected by ecosystem functionality (Schulze and Mooney, 1994). This has raised interest in how biodiversity influences ecosystem function in these ecosystems. Resource extraction by humans is directly influencing biodiversity via selective tree-felling for timber, charcoal making, non-timber forest products and through land use change to agriculture (Aleman et al., 2016; Ryan et al., 2016). Climate change is also indirectly affecting the biodiversity of southern African woodlands, altering

temperature and precipitation, and affecting climate seasonality which heavily influences the degree of seasonal drought and thus woodland structure (Scholes et al., 2004; Eldridge et al., 2012). While rapid biodiversity change is being observed in southern African woodlands, research into the relationship between biodiversity and ecosystem functionality remains scarce.

A small number of academic studies in southern African woodlands have shown that above ground woody carbon/biomass stocks correlate positively with tree species richness (McNicol et al., 2018; Shirima et al., 2015; Mutowo and Murwira, 2012), but the scientific impact of all these studies has been hampered by a restricted climatic and biogeographical range of study sites. Due to the highly variable environmental conditions within which southern African woodlands occur (Frost, 1996), and given the potential importance of environment and biogeography in defining the strength and form of a relationship between biodiversity and woody biomass, it is important to sample across these gradients to be able to infer a truly regional scale relationship between diversity and biomass. Studies conducted over small environmental gradients often find that at local scales, diversity shows a strong effect on ecosystem function, but at large scales diversity effects pale in significance compared to abiotic factors such as climate (Pasari et al., 2013). Small scale studies may therefore be over-estimating the effect of diversity.

In forests, climatic variation is known to affect both woody biomass (Michaletz et al., 2014, 2018) and species diversity independently (Spasojevic et al., 2014), so it is important to account for climatic factors and understand how they interact with biomass and diversity to effectively model and correctly attribute the effects of biodiversity on biomass in analyses at large spatial scales in woodlands. Sankaran et al. (2005) used data from 854 African woodland field sites to show that mean annual precipitation (MAP) sets the upper limit for woody cover in savannas, which is positively correlated with biomass (Chisholm et al., 2013; Prado-Junior et al., 2016), while other factors such as herbivory, fire regime and soil properties also contribute at a local scale to reduce woody cover below this maximum. However, this study did not consider the role of species diversity in the complex interactions between environment and tree cover.

Environmental conditions may indirectly moderate the observed relationship between biodiversity and ecosystem function (Vilà et al., 2005). Condit et al. (2013) found that dry season intensity was the main determinant of tree species distribution and abundance evenness in Panamanian tropical forests. Dry season length may therefore indirectly affect ecosystem level productivity in forest ecosystems and thus woody biomass through selection effects, promoting the growth of a certain highly productive species only under specific environmental conditions. In European forests (Ratcliffe et al., 2017) found a general trend towards stronger positive relationships between tree species richness and various ecosystem functions in more arid environments. They suggested variation in the balance between competitive and facilitative effects along the aridity gradient as the driver of this moderation relationship, an example of the Stress Gradient Hypothesis (Dohn et al., 2013). Water availability imposes a physiological limit on growth rate, which interacts with mortality due to stochastic processes such as fire and herbivory to limit maximum potential woody biomass (Sankaran et al., 2005).

While southern African woodlands are species rich in the herbaceous understorey (Murphy et al., 2016), the tree layer is relatively species poor. Solbrig et al. (1996) however, writes that they remain structurally diverse in the tree layer, with trees occupying distinct layers of the canopy at different growth stages and between species. This structural diversity may be one mechanism through which diversity influences woody biomass. (Kunz et al., 2019) found that crown complementarity and crown plasticity both increased with species richness in a seasonally dry subtropical forest. They also found that trees growing in species rich neighbourhoods exhibited enhanced biomass production. Tree species in southern African woodlands may differ in their occupation of different canopy layers allowing a more full canopy with a greater total foliage density, enhancing productivity and allowing greater standing woody biomass in a smaller area as a form of niche complementarity. This theory however, which has been supported by experiments and observational studies in temperate and wet tropical ecosystems (Hardiman et al., 2011; Stark et al., 2012), may not occur in savannas, which are structured by disturbance rather than competition. Instead, disturbance history may override

the effects of species diversity on structural diversity nullifying the effects of species diversity on structural diversity.

(Zhang et al., 2012) conducted a meta-analysis of 54 studies of the BEFR in forest ecosystems and found that greater tree species abundance evenness led to an increased effect size of species richness on productivity. They suggest this effect is caused by an increased heterogeneity of functional traits over space at small scales, leading to a more evenly filled environmental niche volume, which allows a greater number of species to maximise their productivity. We expect that this effect will occur in southern African woodlands as well. A decrease in species abundance evenness may be found as a result of increased disturbance, which prevents dominance by one species, or due to more extreme environmental conditions which preclude the growth of many species, instead favouring those few species who can effectively tolerate sub-optimal conditions. It is important ~~to~~ therefore to consider not only species richness as a measure of diversity, but also the evenness of the distribution of trees among species when assessing the impact of diversity on ecosystem function. ~~ecosystem function.~~

High levels of disturbance in the form of seasonal fires and herbivory from large mammals may moderate the strength of observable BEFRs in southern African woodlands (Staver et al., 2009; Bond and Keeley, 2005). Fire disturbance in forests has been linked to abundance dependent mortality among smaller stems (Roques et al., 2001). A dominant and competitive species which tends to produce many small fast-growing stems is more likely to experience mortality during a fire as the higher surface area to volume ratio of the trunk increases the likelihood of combustion. These fast-growing bushy species exist in southern African woodlands and often proliferate when fire is deliberately excluded, leading to the growth of thicket vegetation which precludes the growth of high biomass tree species (Higgins et al., 2007). Additionally, the dominance of a single species is likely to decrease overall functioning by decreasing niche space utilisation (Cardinale et al., 2002).

In this study, we made the first known regional estimation of the Biodiversity-Ecosystem Function Relationship in southern African woodlands, using inventory plots which straddle multi-dimensional environmental and biogeographical gradients (Figure 1a, 1b). We investigated the relationship between aboveground woody biomass and tree biodiversity. We compared the relative effects of tree species biodiversity with that of environmental factors known to affect ecosystem productivity and biomass accumulation: precipitation, temperature and soil fertility. We also investigated the potential moderation effects of environmental covariates on the relationship between tree species diversity and biomass. We incorporated vegetation type clusters based on major tree species compositional units into our analyses to understand how species composition as well as species biodiversity affected ecosystem functionality. Initially, we made four hypotheses:

1.1 Hypotheses

1. Plots with a higher tree species diversity will maintain higher above-ground woody biomass stocks.
2. More arid plots and plots with less fertile soil will show a stronger positive effect of tree species diversity on above-ground woody biomass.
3. There will be a positive effect of precipitation and soil fertility on AGB. Part of this positive effect will exist as a mediating effect through tree species diversity.
4. Tree structural diversity of the woodland will interact with tree species diversity to provide an indirect path of influence between species diversity and woody biomass stocks.
5. The observed effect of diversity on AGB will increase in strength as stem density increases owing to an increased importance of niche complementarity as competition increases.

To address these hypotheses we used Structural Equation Models with path analysis as a preferred method to simultaneously account for environmental factors and biotic factors, which may interact in their effect on ecosystem structure and therefore biomass.

2 Methods

2.1 Study location

The study used 1766 woodland monitoring plots from the larger SEOSAW network (SEOSAW, 2019) located across 10 countries within southern Africa in so-called Miombo woodlands (Figure 1a, Hopkins and White 1987). The study region spans a precipitation gradient from $\sim 460 \text{ mm y}^{-1}$ in southern Mozambique and southern Zimbabwe and $\sim 1700 \text{ mm y}^{-1}$ in northern Zambia, Malawi and northern Mozambique. The study sites straddle the climate space of the region as a whole (Figure 1b). The 2D convex hull of Mean Annual Precipitation (MAP) and Mean Annual Temperature (MAT) of the study sites covers 94.4% of the pixel-wise climate space of the Miombo woodland area as defined by Hopkins and White (1987), using WorldClim estimates of temperature and precipitation between the year 1970 and 2000 with a pixel size of 30 arc seconds (0.86 km^2 at the equator) (Fick and Hijmans, 2017).

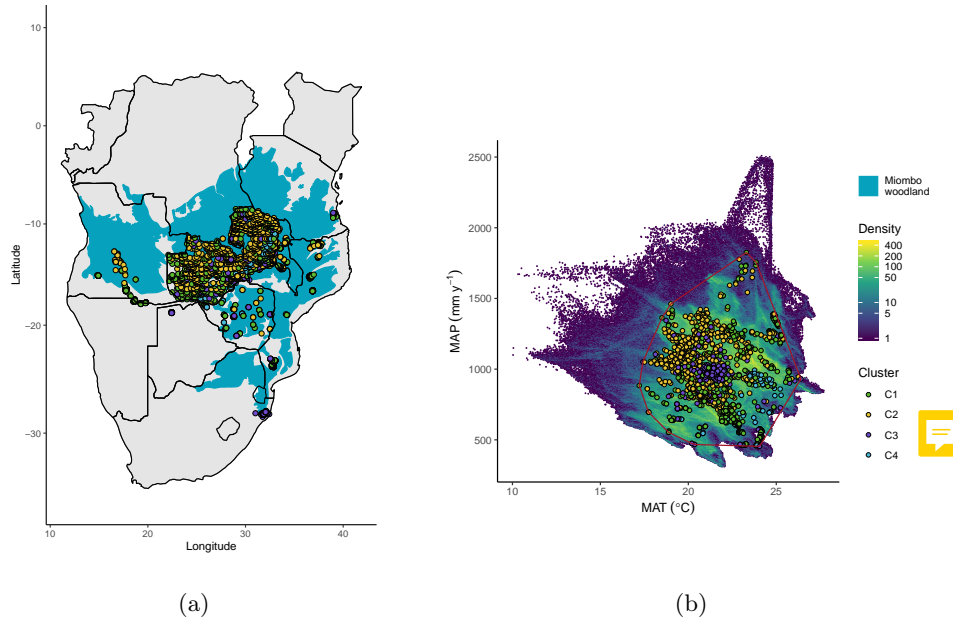


Figure 1: The locations of the 1766 plots used in this study, by geographic location (a) with respect to the distribution of Miombo woodland vegetation according to Hopkins and White (1987), and in climate space (b), showing the plot locations as points compared to the climate space of the whole region as estimated using the WorldClim dataset over the Miombo woodland vegetation extent (Fick and Hijmans, 2017). Note that the density colour scale is log-transformed.

Plots were chosen from a larger pool of 5395 plots based on the quality and completeness of data collection, and plot setup. Plot vegetation was identified under the broad term of “savanna”, which includes “woodland”, “savanna woodland”, and “tree savanna”, variously defined in other areas of the scientific literature (Ratnam et al., 2011; Hill and Hanan, 2011). Plots with evidence of farming, human resource extraction or experimental treatments such as prescribed burning or herbivore exclusion were excluded from the initial pool. Only plots >0.1 hectares were used in analysis, as area based biomass estimation from small plots is highly influenced by rare large trees (Stegen et al., 2011), leading to inaccurate estimates. Only plots with a stem density $>10 \text{ stems ha}^{-1}$ were used, to ensure all plots were within woodland rather than “grassy savanna”, which are considered a separate biome with very different species composition (Parr et al., 2014) (Figure 12).

Many plots provided by the Zambian Forestry Commission were arranged in clusters of up to four $20 \times 50 \text{ m}$ plots, 20 metres apart. Plots within a cluster were aggregated before the plot dataset filtering described above and treated as a single plot in analyses.

After the initial plot data cleaning described above, we conducted an outlier removal procedure of plots with rare tree species composition. We used the `outlier()` function from the `dave` R package (Wildi, 2017), which uses a nearest neighbor criterion for each plot in species abundance ordination space and a threshold value for the minimum nearest neighbour distance to identify outliers. We set the threshold value to remove the top 5% of plots with the largest nearest neighbour distances in multidimensional species composition space (Otto, 2013), removing 97 plots.

2.2 Data collection

We considered only trees and shrubs in our calculations of AGB, including woody species such as palms and cycads which are functionally tree-like but excluding lianas, which fill a different ecological niche (Selaya and Anten, 2008). Only stems >5 cm DBH (Diameter at Breast Height, 1.3 m) were included in analyses. Most plots in the dataset did not include data on stems <5 cm DBH, with these small stems comprising a very small proportion of the total AGB in a plot.

All stems >5 cm DBH were measured within each plot resulting in a total of 160,076 stems with measurements. A tree may be comprised of multiple stems, but for this analysis each stem is treated as an individual. For each stem we measured species, DBH and tree height to the top of the highest branch material. Height was measured through a variety of means including laser rangefinders, manual clinometers and measuring sticks. When DBH could not be measured at 1.3 m due to trunk abnormalities, it was measured at the closest regular portion of the trunk to 1.3 m. The height of this measurement was recorded and used to estimate the DBH_e at 1.3 m using a cubic polynomial regression, with parameters estimated using a test dataset from (Ryan, unpublished).

Woody AGB for each plot was calculated using Equation 1, taken from Chave et al. (2014). Wood density estimates were taken from the global wood density database for each species where possible (Chave et al., 2009; Zanne et al., 2009). Wood density for species without species level estimates was estimated from the mean of their respective genus.

$$AGB = 0.0673 \times (\rho D^2 H)^{0.976} \quad (1)$$

Where ρ is the species level mean wood density, D is the DBH at 1.3 m, and H is the tree height.

Climatic data was collected from the ECMWF ERA5 dataset, generated using Copernicus Climate Change Service Information (Copernicus Climate Change Service, 2017). Values of Mean Annual Temperature (MAT) and Mean Annual Precipitation (MAP) were calculated from daily data between 2000 and 2018, then averaged across years to provide a single mean annual estimate per plot.

Temperature and precipitation seasonality were both calculated as the coefficient of variation of daily MAT and MAP, respectively, across the 18 years of available data. Soil fertility data was extracted from the ISRIC gridded soil information data product at 250 m resolution, taking the grid cell value for each plot (Hengl et al., 2017). We extracted Cation Exchange Capacity (CEC), percentage soil organic carbon by volume (Org. C %), and percentage soil sand content by volume (Sand %). These data are a modelled product compiled from various remote sensed and directly measured data sources.

2.3 Data analysis

Estimated tree species richness was calculated for each plot using `ChaoRichness()` from the `iNEXT` package in R (Hsieh et al., 2016). This procedure uses Hill numbers to extrapolate a species rarefaction curve to its predicted asymptote and uses this value as its estimated species richness value. Extrapolated species richness was preferred over raw Hill numbers as they are more interpretable, representing actual species numbers. Extrapolated species richness accounts for variation in plot size (0.1-10 ha) and therefore sampling effort. Larger plots will tend to encompass more individuals, and therefore more species (Dengler, 2009).

To measure tree species abundance evenness, the Shannon Equitability index ($E_{H'}$) (Smith and Wilson, 1996) (Equation 2) was calculated:

$$E_{H'} = \frac{H'_e}{\ln S} \quad (2)$$

Where H'_e is an estimation of the Shannon diversity index of tree species by extrapolation of the observed Shannon diversity index (H') to its asymptote via Hill numbers using the `ChaoShannon()` function from the `iNEXT` package in R (Hsieh et al., 2016), and S is the extrapolated tree species richness in the plot. We calculated tree structural diversity for each plot by calculating the coefficient of variation of DBH and tree height.

2.3.1 Vegetation clusters

Plots were assigned to vegetation type groups based on tree species composition. Groups were identified in Fayolle et al. (2018) using an Africa wide analysis of floristic units using plot data in savannas and woodlands with tree species diversity and relative abundance data. Groups were identified using unconstrained correspondence analysis and ordination. Plot data used in this study occurred in five vegetation type groups. See Table 1 for a description of each vegetation cluster and Figure 2 for the spatial distribution of plots from each of these clusters .

Table 1: Description of the biogeographical clusters (C1-C5) to which each plot in the study was assigned. Indicator species were generated using Dufrene-Legendre indicator species analysis (Dufrêne and Legendre, 1997) implemented with `indval()` from the `labdsv` R package (Roberts, 2019) and represent species which define the given cluster. Dominant species were identified by choosing the species with the largest AGB contribution within each cluster. Numeric values of species richness, stems ha⁻¹ and AGB are medians and interquartile ranges

| Cluster | Dominant species | Indicator species | N plots | Species rich. | Stems ha ⁻¹ | AGB (t ha ⁻¹) |
|---------|---------------------------------|--|---------|---------------|------------------------|---------------------------|
| 1 | <i>Julbernadia</i> spp. | <i>Diplorhynchus condylocarpon</i> | 734 | 6(5) | 276(208) | 35.4(31.41) |
| | <i>Brachystegia spiciformis</i> | <i>Burkea africana</i> | | | | |
| | <i>Baikaea plurijuga</i> | <i>Pseudolachnostylis maprouneifolia</i> | | | | |
| 2 | <i>Julbernadia</i> spp. | <i>Julbernardia paniculata</i> | 728 | 8(5) | 272(195) | 47.6(44.01) |
| | <i>Brachystegia</i> spp. | <i>Isoberlinia angolensis</i> | | | | |
| | <i>Isoberlinia angolensis</i> | <i>Brachystegia longifolia</i> | | | | |
| 3 | <i>Spirostachys africana</i> | <i>Baikaea plurijuga</i> | 299 | 9(9) | 422(450.9) | 45(45.92) |
| | <i>Senegalia</i> spp. | <i>Senegalia ataxacantha</i> | | | | |
| | <i>Euclea racemosa</i> | <i>Combretum collinum</i> | | | | |
| 4 | <i>Colophospermum mopane</i> | <i>Colophospermum mopane</i> | 102 | 4(4) | 322(228.8) | 43.8(36.49) |
| | | <i>Combretum</i> spp. | | | | |

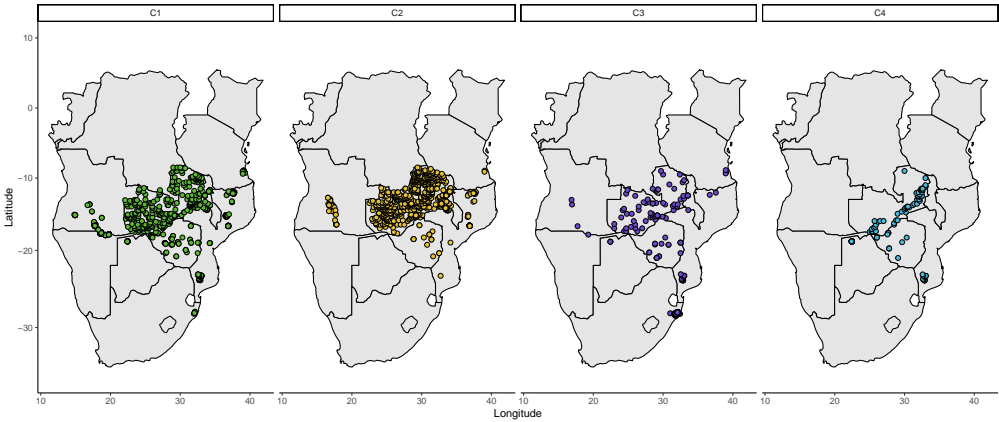


Figure 2: The spatial distribution of plots according to vegetation cluster within southern Africa.

2.3.2 Structural Equation Modelling

Structural Equation Models (SEM) investigated the determinants of AGB. All SEMs were constructed and analysed in the `lavaan` package (Rosseel, 2012) in R version 3.6.0 (R Core Team, 2019). SEM was used because of its suitability for modelling complex causal interactions in ecological systems (Lee, 2007). A key aspect to our decision to use SEMs is that they can explicitly model and partition variance to indirect effects, which is impossible in multiple regression. Using SEMs also allowed us to describe theoretical latent constructs which have been suggested to act upon diversity and biomass/productivity in previous studies despite these factors not having single observable values in our dataset. For example, moisture availability is expected to affect AGB (Saito et al., 2014; Campbell, 1996), but moisture availability itself is determined by the interaction of multiple observable variables over the time scales relevant to tree lifetime growth: precipitation, the seasonality of that precipitation and temperature which affects the rate of evapotranspiration. Independently of total precipitation, precipitation seasonality determines whether water arrives uniformly over a given time period or as a few high volume floods, the latter leading to much water being lost before it can be used for plant growth. Structural equation modelling is also necessary to properly account for potential feedback mechanisms between aspects of climate and woody species diversity, which could otherwise increase the chances of Type I error and wrongly attributing inference due to covariance of explanatory variables, if testing for significant effects of biodiversity on AGB using conventional regression analyses (Nachtigall et al., 2003).

Prior to analysis, we specified a conceptual model with factors expected to affect AGB: moisture availability, soil fertility, tree species diversity, tree structural diversity and stem density (Figure 3).

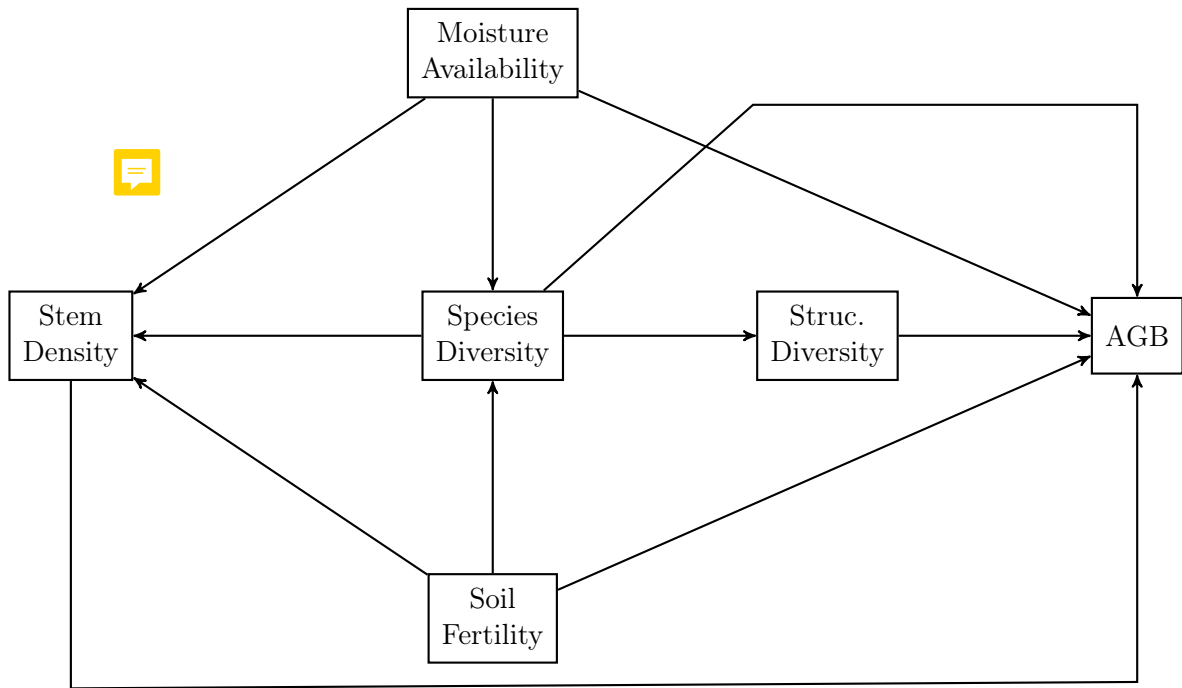


Figure 3: Conceptual Directed Acyclic Graph (DAG) showing the theoretical relationships between environmental factors, tree species diversity, tree structural diversity, tree stem density, and AGB. Hypothesised paths of causation are depicted as arrows from predictor to response.

Observed variables were standardised to Z-scores prior to analysis. Standardization put each latent variable on the same scale, with a mean of zero and a standard deviation of one. Standardization allows path regression coefficients to be easily compared between paths in the same model to assess their relative effect strength, and eliminates confusion in model interpretation arising from the observed variables being on different scales (Beaujean, 2014). Standardization also controls for variables with different orders of magnitude which could otherwise prevent adequate model estimation

from the covariance matrix in `lavaan`. To ensure that observed variables within a latent variable had consistent directions of influence, some observed variables were reversed by multiplying by -1. For example, soil fertility is expected to decrease as soil sand content increases, so soil percentage sand content was reversed for model fitting. Precipitation seasonality, temperature seasonality, and mean temperature were also reversed in this way to account for the direction of their effect on moisture availability.

The factor loadings of the observed variable assumed to contribute most to each latent variable were set to 1 as per convention, with other observed variables being allowed to vary (Beaujean, 2014). While it is recommended by some to set exact factor loadings in the SEM from the regression coefficients of multiple regressions (Rosseel, 2012), because some latent variables were regressed against both structural diversity and AGB, exact factor loadings from simple multiple regressions could not be used. We therefore allowed factor loadings to be estimated by the SEM itself. We tested the robustness of our assumptions with a chi-squared test of all possible combinations of observed variable factor loadings set to 1, finding no significant difference between model specifications. Full Information Max-Likelihood (FIML) was used in each model to estimate the values of missing data in each latent variable (Cham et al., 2017).

First, we assessed the interacting factors of structural diversity and species diversity in determining AGB. We constructed a simple mediation model which allowed species diversity to influence AGB both directly and indirectly via structural diversity. To explore variation in the model among woodland vegetation types, we fit the model both at the regional scale and for each vegetation cluster separately. We compared unstandardised path coefficients among these vegetation cluster scale models to understand the effect that vegetation type has on the relationship between tree species diversity, structural diversity, stem density and AGB. Path coefficients show the effect of a path with other paths of inference held constant. Models were estimated using the “MLM” estimator, because it is robust to multivariate non-normality (Shapiro, 1983). Model fit was evaluated using the robust Comparative Fit Index (CFI), the robust Tucker Lewis Index (TLI), the Root Mean Squared Error (RMSEA) and the R^2 coefficient of determination for AGB. We critically assess model fit in each case, taking into consideration the recommendations of Hu and Bentler (1999) which define threshold values of acceptability for these model fit indices: CFI > 0.9, TLI > 0.9, RMSEA < 0.1, alongside our judgement of the model estimates.

To explore the hypothesis that complementarity effects increase in strength as stem density increases, we repeatedly sub-sampled the available plot dataset to create 900 datasets of similar size with varying median stem density. We used each of these datasets to fit the model including only tree species and structural diversity latent variables to predict AGB. We then examined how the unstandardised path coefficients for each path in the SEM varied according to the median stem density of subsampled dataset.

Second, we incorporated environmental covariates into our model to understand the relative effects of moisture availability and soil fertility on AGB both directly and indirectly via species diversity and stem density. We compared standardised path coefficients between paths in the model to understand the relative contribution of each path to explain variance in AGB.

We fitted separate moderation models to investigate whether there was an interaction effect whereby the strength of the relationship between species diversity and AGB was influenced by moisture availability and soil fertility. The `lavaan` R package does not natively support moderation of latent variables in its model specification. Instead we manually calculated interaction variables for both soil fertility and moisture availability from the product of predicted values of these latent variables in a Confirmatory Factor Analysis (CFA). Interaction variables were the products: tree species diversity \times moisture availability, and tree species diversity \times soil fertility. These interaction terms were included as explanatory terms in multiple regressions alongside the latent variable of species diversity to predict AGB. Regression coefficients and model fit were analysed to determine the presence and form of interaction effects.

3 Results

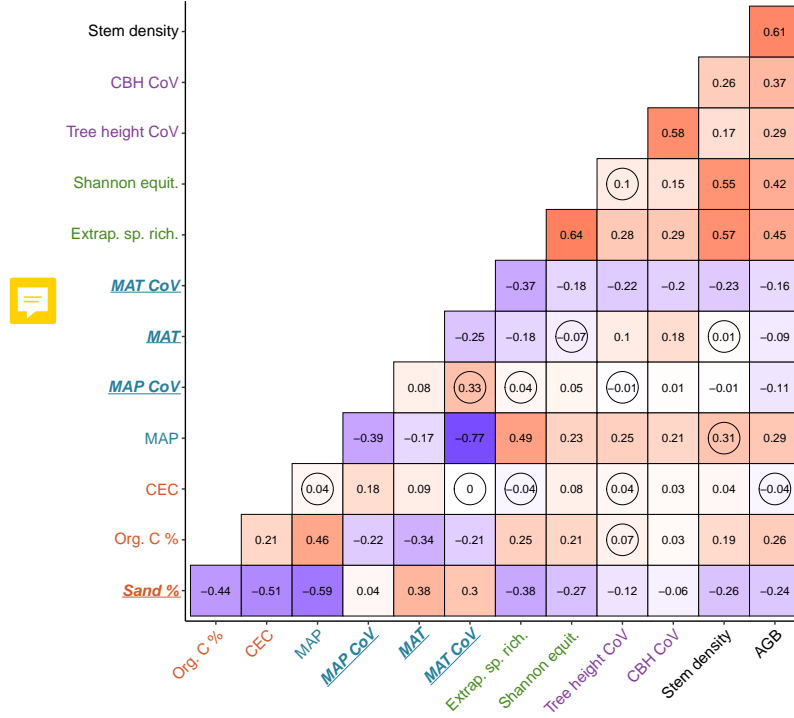


Figure 4: Correlogram of standardised observed variables used in the SEMs, with Pearson correlation coefficients (ρ) coloured according to sign (+ve red, -ve blue) and shaded by strength of correlation. Variables in bold and underlined on the axis labels were later reversed for SEMs to maintain positive correlations for all observed variables within each latent variable. Correlation coefficients marked by a circle indicate that the 95% confidence interval of this correlation overlapped zero. Colours of variable names group them into latent variables used in the SEMs: red = Soil fertility, blue = Moisture availability, green = tree species diversity, purple = tree structural diversity. See Table 3 for a full assessment of correlation fit statistics.

Pairwise correlations between all observed variables used in the Structural Equation Models (SEMs) showed that all tree species diversity and structural diversity variables had moderate positive correlations with AGB. Stem density had the strongest correlation with AGB of all variables ($\rho = 0.56$, $p < 0.01$). Environmental variables had weaker correlations with AGB than diversity variables, with all environmental variables having significant correlations with AGB, except CEC.

The direction of these correlations was used as a test of our assumptions of the direction of influence of latent variables later used in the SEMs. As expected, there was a positive correlation between MAP and AGB ($\rho = 0.21$, $p < 0.01$), and a negative correlation between the seasonality of precipitation and AGB ($\rho = -0.1$, $p < 0.01$). MAT and the seasonality of temperature negatively correlated weakly with AGB (MAT: $\rho = -0.08$, $p < 0.01$; MAT seasonality: $\rho = -0.14$, $p < 0.01$). As expected, there was a negative correlation between soil sand content and AGB ($\rho = -0.23$, $p < 0.01$), and a positive correlation between soil organic carbon and AGB ($\rho = 0.25$, $p < 0.01$).

MAP had positive correlations with tree species richness ($\rho = 0.43$, $p < 0.01$), abundance evenness ($\rho = 0.1$, $p < 0.01$), tree height diversity ($\rho = 0.25$, $p < 0.01$) and a positive but not significant correlation with tree stem density. MAT had weak correlations with tree species and structural diversity variables. Tree species diversity variables had clear positive correlations with stem density (Species richness: $\rho = 0.51$, $p < 0.01$; Shannon equitability: $\rho = 0.61$, $p < 0.01$).

3.1 Structural and species diversity models

In an SEM describing the effect of tree species diversity on AGB via the mediating effects of stand structural diversity and stem density (Figure 5), species diversity had a positive effect on AGB, both directly and indirectly via stem density and structural diversity (Figure 5, Figure 6). Tree species diversity had a strong positive effect on stem density. Model fit was good with high factor loadings for all observed variables, all path coefficients were significant ($p < 0.01$) (Table 2). The R-squared of AGB was 0.46. The strongest effect of tree species diversity on AGB was via stem density ($\beta = 0.32 \pm 0.033$, $p < 0.01$)

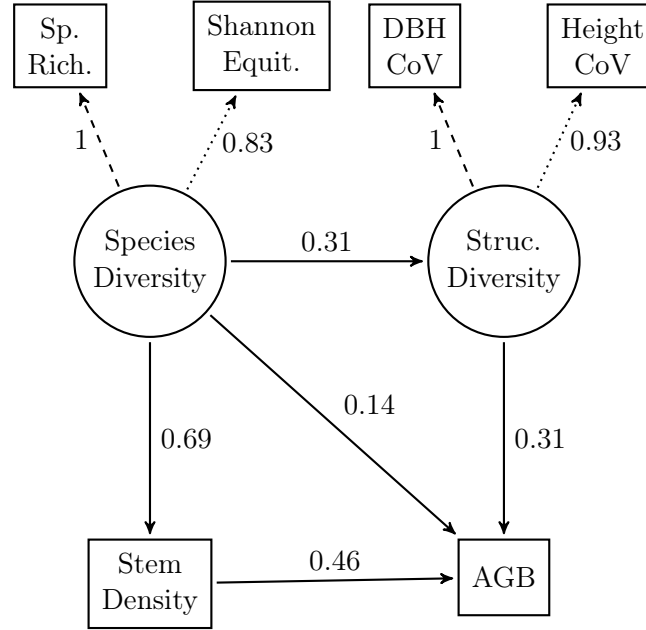


Figure 5: Path diagram with regression coefficients for the tree diversity SEM, including plots from all five vegetation clusters. Latent variables are shown as circles while observed variables are shown as rectangles. Path coefficients are solid arrows pointing from predictor to response. The observed variables which inform the unmeasured latent variables are connected by dotted arrows, observed variables with loading set to one are connected by dashed arrows. Measurement errors of exogenous variables are omitted for clarity.

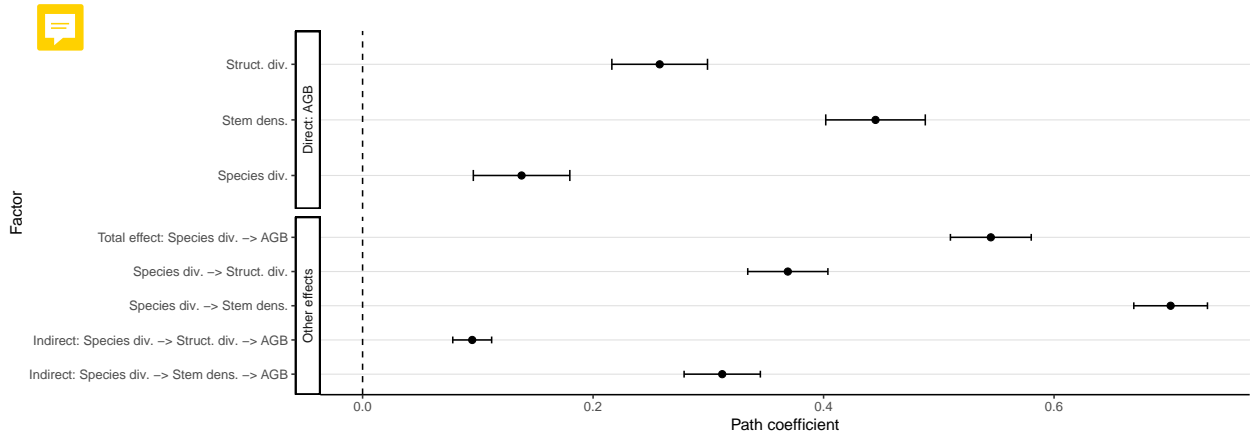


Figure 6: Standardised path coefficients for the effects of tree diversity on AGB, mediated by the effect of stand structural diversity. Due to all observed variables being standardised and centred, path coefficients are expressed in terms of standard deviations on the latent variable response scale ± 1 standard error. Path coefficients where the standard error does not overlap zero are considered significant effects.

3.2 Variation among vegetation clusters in structural and species diversity effects

When the tree species and structural diversity model (Figure 5) was refitted separately using data from each of the 4 vegetation clusters the strengths of unstandardised path coefficients varied but relationships between tree diversity and AGB remained generally similar with the same sign and significant overlap between the 95% confidence intervals of path coefficients. All models exhibited adequate goodness-of-fit (Table 2), though wide confidence intervals around the unstandardised path coefficients, particularly for Clusters 3 and 4 indicate an issue with model prediction due to low sample size for these clusters. χ^2 statistics were high for some vegetation clusters, but this appears to be highly correlated with sample size in our data ().

Cluster 2, which contained vegetation dominated by *Julbernardia paniculata* showed no significant effect of tree species diversity on AGB ($\beta = -0.01 \pm 0.057$, $p = 0.8$), in contrast to the other vegetation clusters, and had a smaller positive effect of structural diversity on AGB ($\beta = 0.08 \pm 0.045$, $p = 0.09$) (Figure 7).

The strongest total effect of tree species diversity on AGB was in Cluster 3 ($\beta = 0.25 \pm 0.163$, $p = 0.13$), which was species rich but highly variable in species diversity compared to other vegetation clusters (Table 1). The greatest effect of structural diversity on AGB was found in Clusters 3 and 4. The R-squared of AGB was highest in cluster 3 ($R^2 = 0.69$) and lowest in cluster 2 ($R^2 = 0.37$).

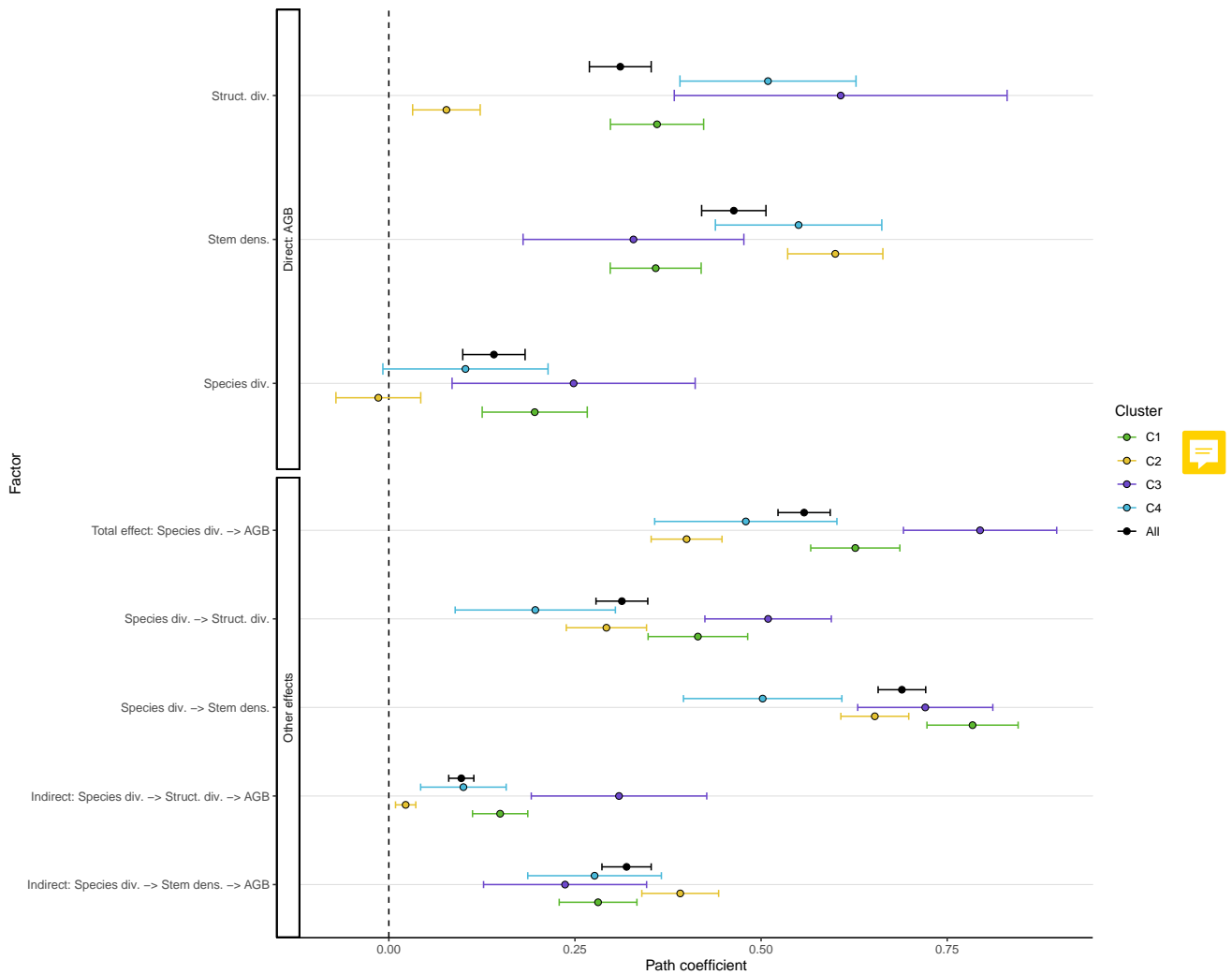


Figure 7: Unstandardised path coefficients for the effects of tree diversity on AGB, mediated by the effect of stand structural diversity. Path coefficients are ± 1 standard error. Path coefficients where the standard error does not overlap zero are considered to be significant effects.

Table 2: Model fit statistics for SEMs investigating the effects of tree diversity and stem density on AGB (Figure 5).

| Cluster | n | χ^2 | DoF | CFI | TLI | LogLik | RMSEA | R^2 AGB |
|---------|------|----------|-----|-------|-------|-----------|-------|-----------|
| C1 | 525 | 38.850 | 6 | 0.970 | 0.924 | -3842.200 | 0.100 | 0.460 |
| C2 | 629 | 49.100 | 6 | 0.966 | 0.915 | -3961.500 | 0.110 | 0.370 |
| C3 | 92 | 4.570 | 6 | 0.968 | 0.935 | -609.900 | 0.090 | 0.690 |
| C4 | 84 | 16.760 | 6 | 0.940 | 0.851 | -606.400 | 0.150 | 0.600 |
| All | 1326 | 74.710 | 6 | 0.976 | 0.939 | -9305 | 0.090 | 0.460 |

3.2.1 Moderation of Diversity-AGB relationship by stem density

We repeatedly sub-sampled the plot dataset to build 80 datasets of varying mean stem density in order to test how the relationship between species diversity, structural diversity and biomass varied with stem density. Each dataset consisted of approximately 361 plots with overlap of plot identity between subsampled datasets. The same SEM specification was used as above (Figure 5) except that paths via stem density were removed. Figure 8 shows a saturating effect of tree species diversity on AGB as stem density increases. There appears to be a minimum stem density threshold at ~ 295 stems ha^{-1} below which there appears to be no effect of tree diversity on biomass. The effect of structural diversity on AGB appears to decrease with increasing stem density, though the relationship is less clear.

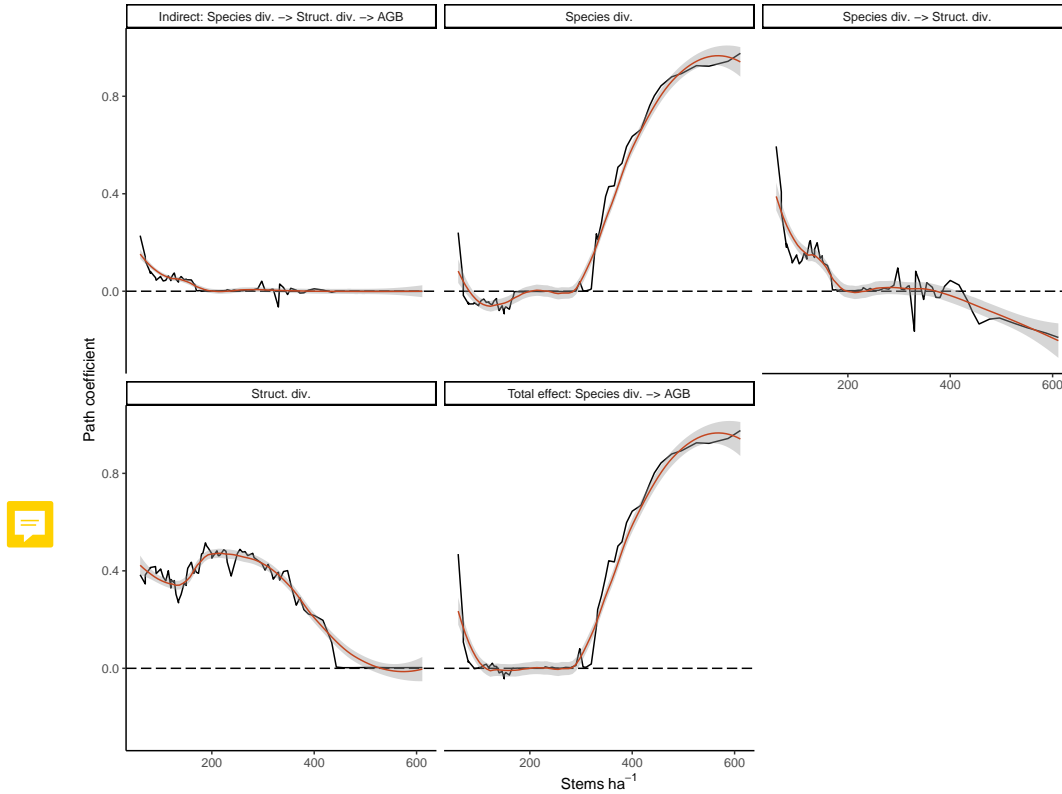


Figure 8: Line plots showing the variation in path coefficients in the SEM, using datasets with different mean stem density. Smoothed lines are loess curves with standard error shaded bars.

3.3 Mediation of environmental covariates via diversity

A model incorporating the latent variables of moisture availability and soil fertility showed that the effect of diversity on biomass was greater than that of both moisture availability and soil fertility (Figure 9). Surprisingly, the direct effects of moisture availability and soil fertility on biomass were negligible, with nearly all of their observed effect on AGB coming from the indirect path via species diversity (moisture: $\beta = 0.02 \pm 0.008$, $p = 0.06$, soil: $\beta = 0.02 \pm 0.008$, $p = 0.05$) (Figure 10). MAP and temperature seasonality had the greatest contributions to the latent variable of moisture availability. Model fit was acceptable: CFI = 0.924, TLI = 0.905, and RMSEA = 0.166. The R^2 coefficient of determination on AGB was 0.5.

Moisture availability and soil fertility also had negligible direct effects on stem density, but both had a positive total effect on stem density via species diversity, where species diversity itself had a strong positive effect on stem density ($\beta = 0.62 \pm 0.034$, $p < 0.01$). As species diversity increases, stem density also increases.

Similar to the model which only considered tree species and structural diversity, the direct effect of species diversity on structural diversity was positive, while structural diversity itself had a positive effect on AGB ($\beta = 0.33 \pm 0.042$, $p < 0.01$).

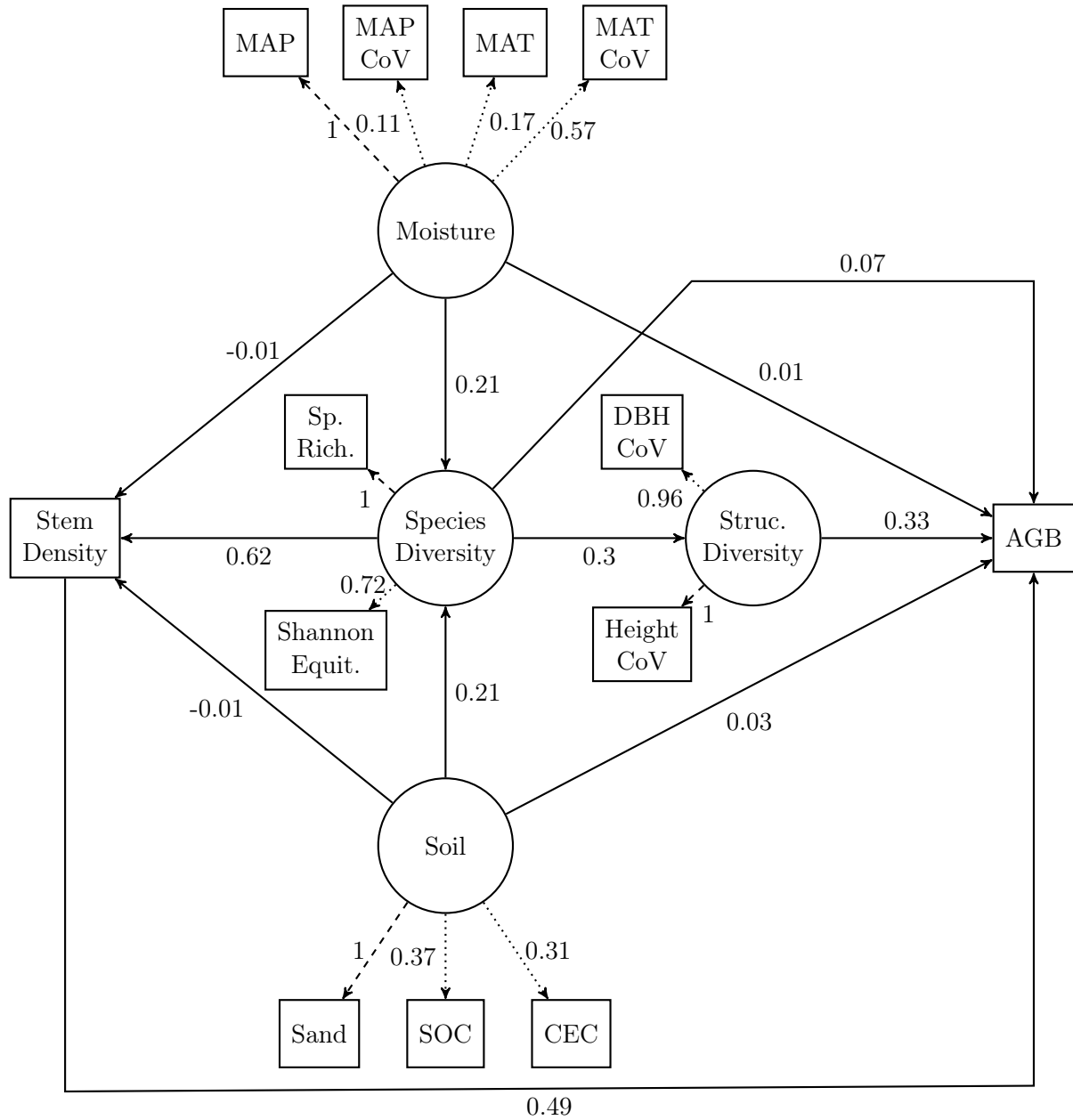


Figure 9: Path diagram with regression coefficients for the SEM incorporating environmental co-variates and tree species and structural diversity across all five vegetation type clusters. Latent variables are shown as circles while observed variables are shown as rectangles. Path coefficients are solid arrows pointing from predictor to response. The observed variables which inform the unmeasured latent variables are connected by dotted arrows, observed variables with loading set to one are connected by dashed arrows. Measurement errors of exogenous variables are omitted for clarity.

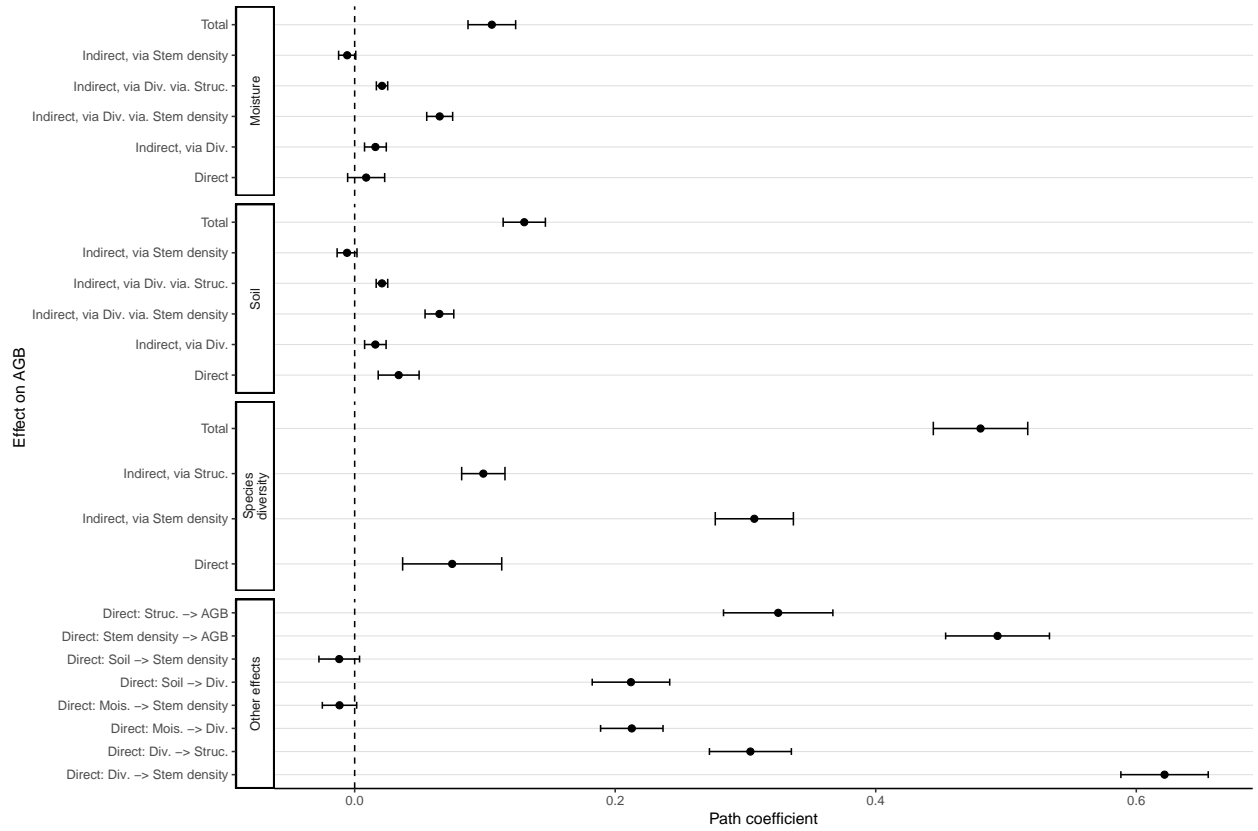


Figure 10: Standardised path coefficients for the interactive effects of abiotic environment and tree diversity on AGB across all plots. Path coefficients are ± 1 standard error. Path coefficients where the standard error does not overlap zero are considered significant effects.

Vegetation cluster level models could not be reliably fitted for this more complex model specification with environmental covariates, due to sample size issues and because some vegetation clusters were narrow in their climate space, leading to a lack of variance particularly in moisture availability.

3.4 Moderation effects of moisture availability and soil

To address the hypothesis (H_2) that more arid plots and plots with less fertile soil will show a stronger positive effect of tree species richness on above ground woody biomass, we fit separate linear multiple regressions with the latent variables of moisture availability and soil fertility as moderators on the relationship between species diversity and AGB.

Both soil fertility and moisture availability had small positive interaction effects on the strength of the relationship between tree species diversity and biomass (Figure 11). The regression coefficient for the interaction effect of tree species diversity and moisture availability was significant ($\beta = 0.06$, $t(1762) = 2.73$, $p < 0.01$), while the interaction term of tree species diversity and soil fertility was not ($\beta = 0.04$, $t(1762) = 1.72$, $p = 0.09$). As moisture availability increased, the effect of tree species diversity on AGB became stronger. See Table 4 and Table 5 for full regression tables.

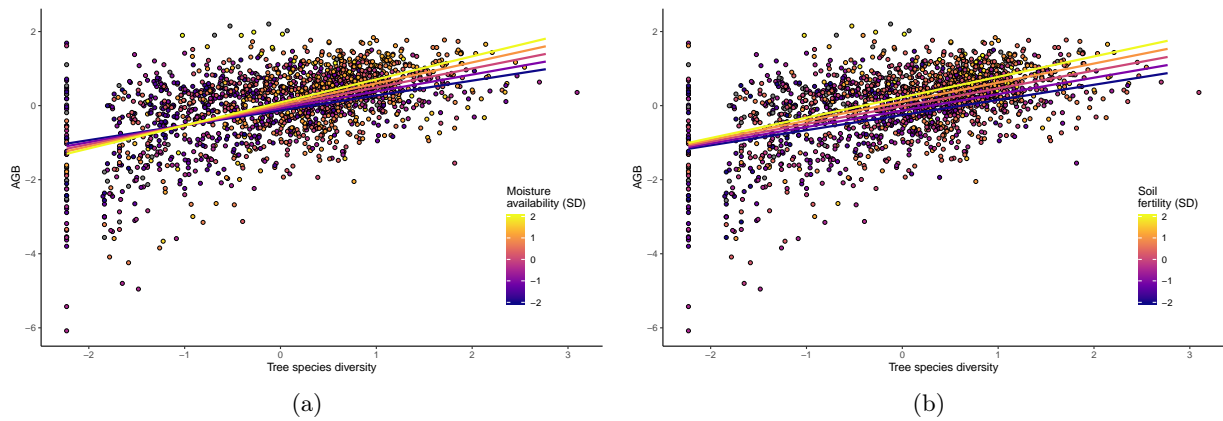


Figure 11: Scatter plots showing variation in the relationship between tree species diversity and AGB with moisture availability (a) and soil fertility (b). All variables are standardised and centred to a mean of zero and a standard deviation of 1. Lines of best fit are drawn for standard deviations of moisture availability and soil fertility between -2 and +2. Moisture availability, tree species diversity and soil fertility are latent variables comprised of the same observed variable loadings.

4 Discussion

In this study, we assessed the importance of [a] tree species richness, [b] tree structural diversity, [c] moisture availability and soil fertility, [d] stem density and their interactions on aboveground woody biomass (AGB) across southern African woodlands, using a plot network of 1766. Using latent variables and Structural Equation Modelling (SEM), we found support for a general relationship between tree species diversity and AGB (H_1), and an indirect relationship between tree species diversity and AGB via structural diversity (H_4) and via stem density (H_5). Tree diversity and stem density accounted for 46% of the variation in AGB across the region, while models for certain vegetation types showed even greater explanatory power (Table 2). The strongest effect on AGB was that of stem density, which itself was highly dependent on species diversity. Interestingly, when tree species diversity, structural diversity and stem density were controlled for, we found little evidence of a direct effect of resource availability, in the form of moisture availability and soil fertility, on AGB (H_3). In separate multiple regressions however, we did find weak indirect moderation effects of both soil fertility and moisture availability on the strength of the relationship between species diversity and AGB (H_2). We found that vegetation composition in the form of discrete vegetation type clusters affected the relationship between species diversity and AGB, with true Miombo woodland representing an outlier in our models (vegetation type C2).

Why is Cluster 2 different?

C3 had the strongest effect of species div on AGB

4.1 Effect of tree species diversity on AGB

The latent variable of tree species diversity had a strong positive effect on woody above-ground biomass (AGB) in the environment+diversity SEM model (Figure 9). Within the savanna woodlands of southern Africa we therefore find support for our hypothesis (H_1), that higher tree species richness causes higher woody AGB. This finding is in agreement with many other studies across different ecosystems and biomes, showing that there is a generalisable positive association between species richness and ecosystem functionality (Liang et al., 2016; Cardinale et al., 2009).

Most of the total species diversity effect was through the indirect effect via stem density. This suggests that within southern African woodlands a higher species diversity allows for a greater density of tree stems and this leads to an increase in total AGB. While reverse causation is plausible,

with increased stem density causing higher species richness through an increased probability of encountering new species, we accounted for this variation via the extrapolation of species richness according to stem density (sampling effort), meaning that stem density should have little causal effect on species diversity in our models. This suggests that the observed direction of causation in our model is real. We cannot decompose the relative effects of tree species richness and abundance evenness in our model, but previous studies have shown that both richness and evenness have similar importance in their effects on ecosystem function (Valéry et al., 2009; Zhang et al., 2012). We suggest that an increase in tree species diversity through species richness and evenness produces an assemblage of species which can occupy a greater proportion of the total woodland canopy volume with leaf area, utilising more of the available light resulting in greater total AGB at the plot level. This is supported by the moderately strong positive effect of tree species diversity on AGB via structural diversity.

While we did not explicitly measure net primary production (NPP) in this study, other studies have shown a strong positive correlation between woody AGB and NPP in woodland and forest ecosystems (Chisholm et al., 2013; Prado-Junior et al., 2016). This suggests that as has been found in many other woodland/forest ecosystems, woody biomass and woody productivity in southern African woodlands can be maximised by increasing species diversity.

4.2 Structural diversity as a mechanism for the BEFR

We found evidence that species diversity led to an increase in AGB via tree structural diversity, we can therefore accept our hypothesis (H_4). A higher tree species diversity allows for a greater diversity of tree functional forms within a plot and this may act as a mechanism of niche complementarity, with a highly diverse canopy being able to take advantage of a greater proportion of the available light. Variation in structural diversity may be a joint result of disturbance history and tree species diversity, with highly disturbed plots generally having a less structurally diverse canopy (LaRue et al., 2019). In forests, where the tree canopy is effectively closed, as the stand matures a more diverse canopy emerges via competition and tree mortality events which open canopy gaps (Muscolo et al., 2014). ~~In highly disturbed woodland plots, we~~ suggest that disturbance history may be an important unmeasured abiotic factor which may account for a high proportion of the variation in AGB and in tree species diversity. In highly disturbed forests, maturity in the sense of classical forest succession is never reached, and a simple open canopy is maintained. Unfortunately we did not have access to a dataset of high enough spatial and temporal resolution to adequately quantify the effects of disturbance history.

4.3 Effects of moisture availability and soil fertility

Surprisingly, moisture availability and soil fertility had only small effects on AGB compared to that of tree species diversity. We expected that higher moisture availability and soil fertility would lead to higher AGB under the assumption that higher resource availability would allow for greater resource partitioning between individual trees () and a greater stem density per unit area ().

Previous studies in tropical forests have shown that moisture availability increases AGB both directly and indirectly via increasing tree species diversity and via increasing stand structural diversity (Ali et al., 2019a,b; Poorter et al., 2017). In this study, while we observed the indirect effects on species diversity, we saw no evidence for a direct effect on AGB. Compared to moist tropical forests, climatic water availability is more of a limiting factor to tree growth in southern African woodlands, which are frequently droughted. It is possible that the range of observed moisture availability in this study ($\sim 460\text{--}1700\text{ mm y}^{-1}$) may not have been able to capture variation in AGB. We deliberately excluded plots with very low stem density as they are not considered woodlands, but grassy savannas. It may be that by excluding the bottom end of this stem density continuum we prevented a relationship being observed between moisture and AGB/stem density. Additionally, due to the high levels of adaptation of tree species to drought conditions in southern Africa, at the large scale we conducted our experiment turnover in species composition may have prevented a direct relationship being

observed between resource availability and AGB.

In southern African woodlands moisture availability is closely linked with the intensity of disturbance from seasonal fires. The growth of C4 grasses in wetter woodlands leads to more intense seasonal fires which limit tree growth (Charles-Dominique et al., 2018), and may also limit species diversity (Linder, 2014). It is possible therefore that the effect of moisture availability is confounded in its effect on AGB with the unmeasured variable of intensity of disturbance by fire. The direct effect of moisture availability on stem density may also be confounded in this way.

In a separate set of linear multiple regressions, we found a weak positive interaction effect of moisture availability on the relationship between tree species diversity and AGB. As moisture availability increased, the relationship between tree species diversity and AGB became stronger. This is in contrast to Ratcliffe et al. (2017) who found that in European forests, a decrease in water availability due to drought led to a stronger effect of tree species diversity on AGB. They attributed this to an increase in selection effects which allowed dominance of stress tolerant species which were more likely to occur in high diversity assemblages. **Why were our results different?**

4.4 Vegetation type cluster specific responses

I have no idea what to write here just yet



4.5 Stem density effects

We found a non linear saturating effect of stem density on the relationship between tree species diversity and AGB (Figure 8). The effect of tree species diversity on AGB peaked at ~700 stems ha⁻¹. At low stem densities competition between trees may not occur, meaning that the niche complementarity provided by an increase in tree species richness might not make any difference to plot level AGB.

It is possible that at very high stem density there are other factors which become more important than tree species diversity in affecting AGB. For example, very high stem density plots in our dataset often had tree species commonly found in thicket vegetation, which is highly structured through disturbance by fire. There may also be an effect whereby the high stem density plots have many small stems, meaning that AGB cannot be high. An abundance of small stems may prevent the growth of large trees which hold the majority of the AGB in a plot. The peak in species diversity effect may occur at a given stem density due to a balance point between tree species diversity effects and the offsetting effects of stem crowding which preclude high biomass large trees.

Maybe at very low stem density, no niche complementarity

Decreasing effect of struct. on AGB with stem density



4.6 Conclusion

References

- Aleman, J. C., Blarquez, O. and Staver, C. A. (2016), 'Land-use change outweighs projected effects of changing rainfall on tree cover in sub-saharan africa', *Global Change Biology* **22**, 3013–3025.
- Ali, A., Lin, S., He, J., Kong, F., Yu, J. and Jiang, H. (2019a), 'Big-sized trees overrule remaining trees' attributes and species richness as determinants of aboveground biomass in tropical forests', *Global Change Biology* **25**, 2810–2824.
- Ali, A., Lin, S., He, J., Kong, F., Yu, J. and Jiang, H. (2019b), 'Climate and soils determine aboveground biomass indirectly via species diversity and stand structural complexity in tropical forests', *Forest Ecology and Management* **432**, 823–831.

- Beaujean, A. A. (2014), *Latent variable modeling using R*, Routledge, New York NY, USA.
- Bond, W. J. and Keeley, J. E. (2005), ‘Fire as a global ‘herbivore’: the ecology and evolution of flammable ecosystems’, *Trends in Ecology and Evolution* **20**(7), 387–394.
- Brooks, T. M., Mittermeier, R. A., da Fonseca, G. A. B., Gerlach, J., Hoffmann, M., Lamoreux, J. F., Mittermeier, C. G., Pilgrim, J. D. and Rodrigues, S. L. (2006), ‘Global biodiversity conservation priorities’, *Science* **313**, 58–61.
- Campbell, B. M., ed. (1996), *The miombo in transition: woodlands and welfare in Africa*, Centre for International Forestry Research, Bogor, Indonesia.
- Cardinale, B. J., Palmer, M. A. and Collins, S. L. (2002), ‘Species diversity enhances ecosystem functioning through interspecific facilitation’, *Nature* **415**, 426–429.
- Cardinale, B. J., Srivastava, D. S., Duffy, J. E., Wright, J. P., Downing, A. L., Sankaran, M., Jouseau, C., Cadotte, M. W., Carroll, I. T., Weis, J. J., Hector, A. and Loreau, M. (2009), ‘Effects of biodiversity on the functioning of ecosystems: a summary of 164 experimental manipulations of species richness’, *Ecology* **90**(3), 854–854.
- Cham, H., Reshetnyak, E., Rosenfeld, B. and Breitbart, W. (2017), ‘Full information maximum likelihood estimation for latent variable interactions with incomplete indicators’, *Multivariate Behavioural Research* **52**(1), 12–30.
- Charles-Dominique, T., Midgley, G. F., Tomlinson, K. W. and Bond, W. J. (2018), ‘Steal the light: shade vs fire adapted vegetation in forest-savanna mosaics’, *New Phytologist* **218**, 1419–1429.
- Chave, J., Coomes, D., Jansen, S., Lewis, S. L., Swenson, N. G. and Zanne, A. E. (2009), ‘Towards a worldwide wood economics spectrum’, *Ecology Letters* **12**, 351–366.
- Chave, J., Réjou-Méchain, M., Búrquez, A., Chidumayo, E., Colgan, M. S., Delitti, W. B. C., Duque, A., Eid, T., Fearnside, P. M., Goodman, R. C., Henry, M., Martínez-Yrízar, A., Mugasha, W. A., Muller-Landau, H. C., Mencuccini, M., Nelson, B. W., Ngomanda, A., Nogueira, E. M., Ortiz-Malavassi, E., Péliissier, R., Ploton, P., Ryan, C. M., Saldarriaga, J. G. and Vieilledent, G. (2014), ‘Improved allometric models to estimate the aboveground biomass of tropical trees’, *Global Change Biology* **20**(10), 3177–3190.
- Chen, I., Hill, J. K., Ohlemüller, R., Roy, D. B. and Thomas, C. D. (2011), ‘Rapid range of species associated with high levels of climate warming’, *Science* **333**, 1024–1026.
- Chisholm, R. A., Muller-Landau, H. C., Rahman, K. A., Bebbler, D. P., Bin, Y., Bohlman, S. A., Bourg, N. A., Brinks, J., Bunyavejchewin, S., Butt, N., Cao, H., Cao, M., Cárdenas, D., Chang, L., Chiang, J., Chuyong, G., Condit, R., Dattaraja, H. S., Davies, S., Duque, A., Fletcher, C., Gunatilleke, N., Gunatilleke, S., Hao, Z., Harrison, R. D., Howe, R., Hsieh, C., Hubbell, S. P., Itoh, A., Kenfack, D., Kiratiprayoon, S., Larson, A. J., Lian, J., Lin, D., Liu, H., Lutz, J. A., Ma, K., Malhi, Y., McMahon, S., McShea, W., Meegaskumbura, M., Razman, S. M., Morecroft, M. D., Nytch, C. J., Oliveira, A., Parker, G. G., Pulla, S., Punchi-Manage, R., Romero-Saltos, H., Sang, W., Schurman, J., Su, S., Sukumar, R., Sun, I., Suresh, H. S., Tan, S., Thomas, D., Thomas, S., Thompson, J., Valencia, R., Wolf, A., Yap, S., Ye, W., Yuan, Z. and Zimmermann, J. K. (2013), ‘Scale-dependent relationships between tree species richness and ecosystem function in forests’, *Journal of Ecology* **101**, 1214–1224.
- Clarke, D. A., York, P. H., Rasheed, M. A. and Northfield, T. D. (2017), ‘Does biodiversity - ecosystem function literature neglect tropical ecosystems’, *Trends in Ecology & Evolution* **32**(5), 320–323.
- Condit, R., Engelbrecht, B. M. J., Pino, D., Pérez, R. and Turner, B. L. (2013), ‘Species distributions in response to individual soil nutrients and seasonal drought across a community of tropical trees’, *Proceedings of the National Academy of Sciences* **110**(13), 5064–5068.

- Copernicus Climate Change Service (2017), *ERA5: Fifth generation of ECMWF atmospheric re-analyses of the global climate*, Copernicus Climate Change Service Climate Data Store (CDS). Accessed: 2019-09-03.
- Dengler, J. (2009), ‘Which function describes the species-area relationship best? a review and empirical evaluation’, *Journal of Biogeography* **36**, 728–744.
- Dohn, J., Dembélé, F., Karembé, M., Moustakas, A., Amévor, K. A. and Hanan, N. P. (2013), ‘Tree effects on grass growth in savannas: competition, facilitation and the stress-gradient hypothesis’, *Journal of Ecology* **101**, 202–209.
- Dufrêne, M. and Legendre, P. (1997), ‘Species assemblage and indicator species: the need for a flexible asymmetrical approach’, *Ecological Monographs* **67**(3), 345–366.
- Eldridge, D. J., Maestre, F. T., Maltez-Mouro, S. and Bowker, M. A. (2012), ‘A global database of shrub encroachment effects on ecosystem structure and functioning’, *Ecology* **93**(11), 2499.
- Fayolle, A., Swaine, M. D., Aleman, J., Azihou, A. F., Bauman, D., te Beest, M., Chidumayo, E. N., Cromsigt, J. P. G. M., Dessard, H., Finkch, M., Gonçalves, F. M. P., Gillet, J., Gorel, A., Hick, A., Holdo, R., Kirunda, B., Mahy, G., McNicol, I., Ryan, C. M., Revermann, R., Plumptre, A., Pritchard, R., Nieto-Quintano, P., Schmitt, C. B., Seghier, J., Swemmer, A., Talila, H. and Woollen, E. (2018), ‘A sharp floristic discontinuity revealed by the biogeographic regionalization of african savannas’, *Journal of Biogeography* **46**(2), 1–12.
- Fick, S. E. and Hijmans, R. J. (2017), ‘Worldclim 2: new 1-km spatial resolution climate surfaces for global land areas’, *International Journal of Climatology* **37**(12), 4302–4315.
- Frost, P. (1996), The ecology of miombo woodlands, in B. Campbell, ed., ‘The miombo in transition: woodlands and welfare in Africa’, Center for International Forestry Research, Bogor, Indonesia, pp. 11–55.
- Hardiman, B. S., Bohrer, G., Gough, C. M., Vogel, C. S. and Curtis, P. S. (2011), ‘The role of canopy structural complexity in wood net primary production of a maturing northern deciduous forest’, *Ecology* **92**(9), 1818–1827.
- Hector, A. and Bagchi, R. (2007), ‘Biodiversity and ecosystem multifunctionality’, *Nature* **448**, 188–190.
- Hengl, T., Mendes de Jesus, J., Heuvelink, G. B. M., Gonzalez, M. R., Kilibarda, M., Blagotić, A., Shangquan, W., Wright, M. N., Geng, X., Bauer-Marschallinger, B., Guevara, M. A., Vargas, R., MacMillan, R. A., Batjes, N. H., Leenaars, J. G. B., Ribeiro, E., Wheeler, I., Mantel, S. and Kempen, B. (2017), ‘Soilgrids250m: Global gridded soil information based on machine learning’, *PLoS ONE* **12**(2), 1–40.
- Higgins, S. I., Bond, W. J., February, E. C., Bronn, A., Euston-Brown, D. I. W., Enslin, B., Govender, N., Rademan, L., O’Regan, S., Potgieter, A. L. F., Scheiter, S., Sowry, R., Trollope, L. and Trollope, W. S. W. (2007), ‘Effects of four decades of fire manipulation on woody vegetation structure in savanna’, *Ecology* **88**(5), 1119–1125.
- Hill, M. J. and Hanan, N. P., eds (2011), *Ecosystem function in savannas*, CRC Press, Boca Raton FL, USA.
- Hinsley, A., Entwistle, A. and Pio, D. V. (2015), ‘Does the long-term success of redd+ also depend on biodiversity’, *Oryx* **49**(2), 216–221.
- Hopkins, B. and White, F. (1987), ‘Vegetation map of africa. the vegetation of africa: A descriptive memoir to accompany the unesco/aetfat/unso vegetation map of africa.’, *The Journal of Ecology* **75**(4).

- Houghton, R. A., Hall, F. and Goetz, S. J. (2009), ‘Importance of biomass in the global carbon cycle’, *Journal of Geophysical Research* **114**, 1–13.
- Hsieh, T. C., Ma, K. H. and Chao, A. (2016), ‘inext: an r package for rarefaction and extrapolation of species diversity (hill numbers)’, *Methods in Ecology and Evolution* **7**, 1451–1456.
- Hu, L. and Bentler, P. M. (1999), ‘Cutoff criteria for fit indexes in covariance structure analysis: conventional criteria versus new alternatives’, *Structural Equation Modeling* **6**(1), 1–55.
- Jax, K. (2005), ‘Function and “functioning” in ecology: what does it mean?’, *Oikos* **111**(3), 641–648.
- Kunz, M., Fichtner, A., Hardtle, W., Raunonen, P., Bruelheide, H. and von Oheimb, G. (2019), ‘Neighbour species richness and local structural variability modulate aboveground allocation patterns and crown morphology of individual trees’, *Ecology Letters* pp. 1–11.
- LaRue, E. A., Hardiman, B. S., Elliott, J. M. and Fei, S. (2019), ‘Structural diversity as a predictor of ecosystem function’, *Environmental Research Letters* **14**, 114011.
- Lasky, J. R., Uriarte, M., Boukili, V. K., Erickson, D. L., Kress, W. J. and Chazdon, R. L. (2014), ‘The relationship between tree biodiversity and biomass dynamic changes with tropical forest succession’, *Ecology Letters* **17**, 1158–1167.
- Lee, S. Y. (2007), *Structural Equation Modeling: A Bayesian Approach*, Wiley, New York, USA.
- Liang, J., Crowther, T. W., Picard, N., Wiser, S., Zhou, M., Alberti, G., Schulze, E.-D., McGuire, A. D., Bozzato, F., Pretzsch, H., De-Miguel, S., Paquette, A., Herault, B., Scherer-Lorenzen, M., Barrett, C. B., Glick, H. B., Hengeveld, G. M., Nabuurs, G.-J., Pfautsch, S., Viana, H., Vibrans, A. C., Ammer, C., Schall, P., Verbyla, D., Tchebakova, N., Fischer, M., Watson, J. V., Chen, H. Y. H., Lei, X., Schelhaas, M.-J., Lu, H., Gianelle, D., Parfenova, E. I., Salas, C., Lee, E., Lee, B., Kim, H. S., Bruelheide, H., Coomes, D. A., Piotto, D., Sunderland, T., Schmid, B., Gourlet-Fleury, S., Sonke, B., Tavani, R., Zhu, J., Brandl, S., Vayreda, J., Kitahara, F., Searle, E. B., Neldner, V. J., Ngugi, M. R., Baraloto, C., Frizzera, L., Ba azy, R., Oleksyn, J., Zawadi a Nied wiecki, T., Bouriaud, O., Bussotti, F., Finer, L., Jaroszewicz, B., Jucker, T., Valladares, F., Jagodzinski, A. M., Peri, P. L., Gonmadje, C., Marthy, W., O’Brien, T., Martin, E. H., Marshall, A. R., Rovero, F., Bitariho, R., Niklaus, P. A., Alvarez-Loayza, P., Chamuya, N., Valencia, R., Mortier, F., Wortel, V., Engone-Obiang, N. L., Ferreira, L. V., Odeke, D. E., Vasquez, R. M., Lewis, S. L. and Reich, P. B. (2016), ‘Positive biodiversity-productivity relationship predominant in global forests’, *Science* **354**(6309), aaf8957–aaf8957.
- Linder, H. P. (2014), ‘The evolution of african plant diversity’, *Frontiers in Ecology and Evolution* **2**(38), 1–14.
- Mayaux, P., Eva, H., Brink, A., Achard, F. and Belward, A. (2008), Remote sensing of land-cover and land-use dynamics, in ‘Earth Observation of Global Change: The Role of Satellite Remote Sensing in Monitoring the Global Environment’, Springer-Verlag, Berlin, Germany, pp. 85–108.
- McNicol, I. M., Ryan, C. M., Dexter, K. G., Ball, S. M. J. and Williams, M. (2018), ‘Aboveground carbon storage and its links to stand structure, tree diversity and floristic composition in south-eastern tanzania’, *Ecosystems* **21**, 740–754.
- Michaletz, S. T., Cheng, D., Kerkhoff, A. J. and Enquist, B. J. (2014), ‘Convergence of terrestrial plant production across global climate gradients’, *Nature* **512**, 39–43.
- Michaletz, S. T., Kerkhoff, A. J. and Enquist, B. J. (2018), ‘Drivers of terrestrial plant production across broad geographical gradients’, *Global Ecology and Biogeography* **27**, 166–174.
- Mittermeier, R. A., Mittermeier, C. G., Brooks, T. M., Pilgrim, J. D., Konstant, W. R., da Fonseca, G. A. B. and Kormos, C. (2003), ‘Wilderness and biodiversity conservation’, *Proceedings of the National Academy of Sciences* **100**(18), 10309–10313.

- Murphy, B. P., Andersen, A. N. and Parr, C. L. (2016), ‘The underestimated biodiversity of tropical grassy biomes’, *Philosophical Transactions of the Royal Society B: Biological Sciences* **371**, 20150319.
- Muscolo, A., Bagnato, S., Sidari, M. and Mercurio, R. (2014), ‘A review of the roles of forest canopy gaps’, *Journal of Forestry Research* **25**(4), 725–736.
- Mutowo, G. and Murwira, A. (2012), ‘Relationship between remotely sensed variables and tree species diversity in savanna woodlands of southern africa’, *International Journal of Remote Sensing* **33**(20), 6378–6402.
- Nachtigall, C., Kroehne, U., Funke, F. and Steyer, R. (2003), ‘(why) should we use sem? pros and cons of structural equation modeling’, *Methods of Psychological Research* **8**(2), 1–22.
- Otto, W. (2013), *Data analysis in vegetation ecology*, Wiley-Blackwell, Chichester, UK.
- Parr, C. L., Lehmann, C. E. R., Bond, W. J., Hoffmann, W. A. and Andersen, A. N. (2014), ‘Tropical grassy biomes: misunderstood, neglected, and under threat’, *Trends in Ecology and Evolution* **29**(4), 205–213.
- Pasari, J. R., Levi, T., Zavaleta, E. S. and Tilman, D. (2013), ‘Several scales of biodiversity affect ecosystem multifunctionality’, *Proceedings of the National Academy of Sciences* **110**(25), 10219–10222.
- Poorter, L., van de Sande, M. T., Thompson, J., Arets, E. J. M. M., Alarcón, A., Álvarez-Sánchez, J., Ascarrunz, N., Balvanera, P., Barajas-Guzmán, G., Boit, A., Bongers, F., Carvalho, F. A., Casanoves, F., Cornejo-Tenorio, G., Costa, F. R. C., de Castilho, C. V., Duivenvoorden, J. F., Dutrieux, L. P., Enquist, B. J., Fernández-Méndez, F., Finegan, B., Gormley, L. H. L., Healey, J. R., Hoosbeek, M. R., Ibarra-Manríquez, G., Junqueira, A. B., Levis, C., Licona, J. C., Lisboa, L. S., Magnusson, W. E., Martínez-Ramos, M., Martínez-Yrizar, A., Martorano, L. G., Maskell, L. C., Mazzei, L., Meave, J. A., Mora, F., Muñoz, R., Nytch, C., Pansonato, M. P., Parr, T. W., Paz, H., Pérez-García, E. A., Rentería, L. Y., Rodríguez-Velázquez, J., Rozendaal, D. M. A., Ruschel, A. R., Sakschewski, B., Salgado-Negret, B., Schiatti, J., Simões, M., Sinclair, F. L., Souza, P. F., Souza, F. C., Stropp, J., ter Steege, H., Swenson, N. G., Thonicke, K., Toledo, M., Uriarte, M., van der Hout, P., Walker, P., Zamora, N. and Peña-Claros, M. (2015), ‘Diversity enhances carbon storage in tropical forests’, *Global Ecology and Biogeography* **24**, 1314–1328.
- Poorter, L., van der Sande, M. T., Arets, E. J. M. M., Ascarrunz, N., Enquist, B. J., Finegan, B., Licona, J. C., Martínez-Ramos, M., Mazzei, L., Meave, J. A., Muñoz, R., Nytch, C. J., de Oliveira, A. A., Pérez-García, E. A., Prado-Junior, J., Rodríguez-Velázquez, J., Ruschel, A. R., Salgado-Negret, B., Schiavini, I., Swenson, N. G., Tenorio, E. A., Thompson, J., Toledo, M., Uriarte, M., van der Hout, P., Zimmerman, J. K. and Peña0-Claros, M. (2017), ‘Biodiversity and climate determine the functioning of neotropical forests’, *Global Ecology and Biogeography* **26**, 1423–1434.
- Prado-Junior, J. A., Schiavini, I., Vale, V. S., Arantes, C. S., van der Sande, M. T., Lohbeck, M. and Poorter, L. (2016), ‘Conservative species drive biomass productivity in tropical dry forests’, *Journal of Ecology* **104**, 817–827.
- R Core Team (2019), *R: A Language and Environment for Statistical Computing*, R Foundation for Statistical Computing, Vienna, Austria.
- Ratcliffe, S., Wirth, C., Jucker, T., van der Plas, F., Scherer-Lorenzen, M., Verheyen, K., Allan, E., Benavides, R., Bruelheide, H., Ohse, B., Paquette, A., Ampoorter, E., Bastias, C. C., Bauhus, J., Bonal, D., Bouriaud, O., Bussotti, F., Carnol, M., Castagneyrol, B., Češko, E., Dawud, S. M., de Wandeler, H., Domisch, T., Finér, L., Fischer, M., Fotelli, M., Gessler, A., Granier, A., Grossiord, C., Guyot, V., Haase, J., Hättenschwiler, S., Jactel, H., Jaroszewicz, B., Joly, F., Kambach, S., Kolb, S., Koricheva, J., Liebersgesell, M., Milligan, H., Müller, S., Muys, B., Nguyen, D., Nock, C.,

- Pollastrini, M., Purschke, O., Radoglou, K., Raulund-Rasmussen, K., Roger, F., Ruiz-Benito, P., Seidl, R., Selvi, F., Seiferling, I., Stenlid, J., Valladares, F., Vesterdal, L. and Baeten, L. (2017), 'Biodiversity and ecosystem functioning relations in european forests depend on environmental context', *Ecology Letters* **20**, 1414–1426.
- Ratnam, J., Bond, W. J., Fensham, R. J., Hoffmann, W. A., Archibald, S., Lehmann, C. E. R., Anderson, M. T., Higgins, S. I. and Sankaran, M. (2011), 'When is a 'forest' a savanna, and why does it matter?', *Global Ecology and Biogeography* **20**, 653–660.
- Roberts, D. W. (2019), *labdsv: Ordination and Multivariate Analysis for Ecology*. R package version 2.0-1.
- Roques, K. G., O'Connor, T. G. and Watkinson, A. R. (2001), 'Dynamics of shrub encroachment in an african savanna: relative influences of fire, herbivory, rainfall and density dependence', *Journal of Applied Ecology* **38**, 268–280.
- Rosseel, Y. (2012), 'lavaan: An R package for structural equation modeling', *Journal of Statistical Software* **48**(2), 1–36.
- Ryan, C. M., Pritchard, R., McNicol, I., Owen, M., Fisher, J. A. and Lehmann, C. (2016), 'Ecosystem services from southern african woodlands and their future under global change', *Philosophical Transactions of the Royal Society B: Biological Sciences* **371**(1703), 1–16.
- Saito, M., Luyssaert, S., Poulter, B., Williams, M., Ciais, P., Bellassen, V., Ryan, C. M., Yue, C., Cadule, P. and Peylin, P. (2014), 'Fire regimes and variability in aboveground woody biomass in miombo woodland', *Journal of Geophysical Research: Biogeosciences* **119**, 1014–1029.
- Sankaran, M., Hanan, N. P., Scholes, R. J., Ratnam, J., Augustine, D. J., Cade, B. S., Gignoux, J., Higgins, S. I., Le Roux, X., Ludwig, F., Ardo, J., Banyikwa, F., Bronn, A., Bucini, G., Caylor, K. K., Coughenour, M. B., Diouf, A., Ekaya, W., Feral, C. J., February, E. C., Frost, P. G. H., Hiernaux, P., Hrabar, H., Metzger, K. L., Prins, H. H. T., Ringrose, S., Sea, W., Tews, J., Worden, J. and Zambatis, N. (2005), 'Determinants of woody cover in african savannas', *Nature* **438**(8), 846–849.
- Scholes, R. J., Frost, P. G. H. and Tian, Y. (2004), 'Canopy structure in savannas along a moisture gradient on kalahari sands', *Global Change Biology* **10**, 292–302.
- Schulze, E. D. and Mooney, H. A. (1994), *Ecosystem function of biodiversity: A summary*, Springer Berlin Heidelberg, Berlin, Germany.
- Selaya, N. G. and Anten, P. R. (2008), 'Differences in biomass allocation, light interception and mechanical stability between lianas and trees in early secondary tropical forest', *Functional Ecology* **22**, 30–39.
- SEOSAW (2019), 'Seosaw: A socio-ecological observatory for southern african woodlands'. [Online; 15-January-2019].
- Shapiro, A. (1983), 'Asymptotic distribution theory in the analysis of covariance structures (a unified approach)', *South African Statistical Journal* **17**, 33–81.
- Shirima, D. D., Pfeifer, M., Platts, P. J., Totland, Ø. and Moe, S. R. (2015), 'Interactions between canopy structure and herbaceous biomass along environmental gradients in moist forest and dry miombo woodland of tanzania', *PLoS ONE* **10**(11), 1–15.
- Smith, B. and Wilson, J. B. (1996), 'A consumer's guide to evenness indices', *Oikos* **76**(1), 70–82.
- Solbrig, O. T., Medina, E. and Silva, J. F. (1996), *Biodiversity and Savanna Ecosystem Processes*, Springer-Verlag, Berlin, Germany.

- Spasojevic, M. J., Grace, J. B., Harrison, S. and Damschen, E. I. (2014), 'Functional diversity supports the physiological tolerance hypothesis for plant species richness along climatic gradients', *Journal of Ecology* **102**, 447–455.
- Stark, S. C., Leitold, V., Wu, J. L., Hunter, M. O., de Castilho, C. V., Costa, F. R. C., McMahon, S. M., Parker, G. G., Shimabukuro, M. T., Lefsky, M. A., Keller, M., Alves, L. F., Schiatti, J., Shimabukuro, Y. E., Brandão, D. O., Woodcock, T. K., Higuchi, N., de Camargo, P. B., de Oliveira, R. C. and Saleska, S. R. (2012), 'Amazon forest carbon dynamics predicted by profiles of canopy leaf area and light environment', *Ecology Letters* **15**, 1406–1414.
- Staver, A. C., Bond, W. J., Stock, W. D., van Rensburg, S. J. and Waldram, M. S. (2009), 'Browsing and fire interact to suppress tree density in an african savanna', *Ecological Applications* **19**(7), 1909–1919.
- Stegen, J. C., Swenson, N. G., Enquist, B. J., White, E. P., Phillips, O. L., Jørgensen, P. M., Weiser, M. D., Mendoza, A. M. and Vargas, P. N. (2011), 'Variation in above-ground forest biomass across broad climatic gradients', *Global Ecology and Biogeography* **20**, 744–754.
- Thompson, I. D., Okabe, K., Tylianakis, J. M., Kumar, P., Brockerhoff, E. G., Schellhorn, N. A., Parrotta, J. A. and Nasi, R. (2011), 'Forest biodiversity and the delivery of ecosystem goods and services: translating science into policy', *BioScience* **61**(12), 972–981.
- Tilman, D. and Downing, J. A. (1994), 'Biodiversity and stability in grasslands', *Nature* **367**, 363–365.
- Tilman, D., Isbell, F. and Cowles, J. M. (2014), 'Biodiversity and ecosystem functioning', *Annual Review of Ecology, Evolution, and Systematics* **45**, 471–493.
- Tobner, C. M., Paquette, A., Gravel, D., Reich, P. B., Williams, L. J. and Messier, C. (2016), 'Functional identity is the main driver of diversity effects in young tree communities', *Ecology Letters* **19**, 638–647.
- Valéry, L., Fritz, H., Lefeuvre, J. and Simberloff, D. (2009), 'Ecosystem-level consequences of invasions by native species as a way to investigate relationships between evenness and ecosystem function', *Biological Invasions* **11**, 609–617.
- van der Sande, M. T., Poorter, L., Kooistra, L., Balvanera, P., Thonicke, K., Thompson, J., Arets, E. J. M. M., Alaniz, N. G., Jones, L., Mora, F., Mwampamba, T. H., Parr, T. and Peña-Claros, M. (2017), 'Biodiversity in species, traits, and structure determines carbon stocks and uptake in tropical forests', *Biotropica* **49**(5), 593–603.
- Vilà, M., Inchausti, P., Vayreda, J., Barrantes, O., Gracia, C., Ibáñez, J. J. and Mata, T. (2005), 'Confounding factors in the observational productivity-diversity relationship in forests', *Ecological Studies* **176**, 65–86.
- Wildi, O. (2017), *dave: Functions for "Data Analysis in Vegetation Ecology"*. R package version 2.0.
- Wright, A. J., Wardle, W. D. A., Callaway, W. R. and Gaxiola, A. (2017), 'The overlooked role of facilitation in biodiversity experiments', *Trends in Ecology and Evolution* **32**(5), 383–390.
- Zanne, A. E., Lopez-Gonzalez, G., Coomes, D. A., Ilic, J., Jansen, S., Lewis, S. L., Miller, R. B., Swenson, N. G., Wiemann, M. C. and Chave, J. (2009), *Global wood density database*. <http://hdl.handle.net/10255/dryad.235>.
- Zhang, Y., Chen, H. Y. H. and Reich, P. B. (2012), 'Forest productivity increases with evenness, species richness and trait variation: a global meta-analysis', *Journal of Ecology* **100**, 742–749.

Appendix 1 - Data cleaning process

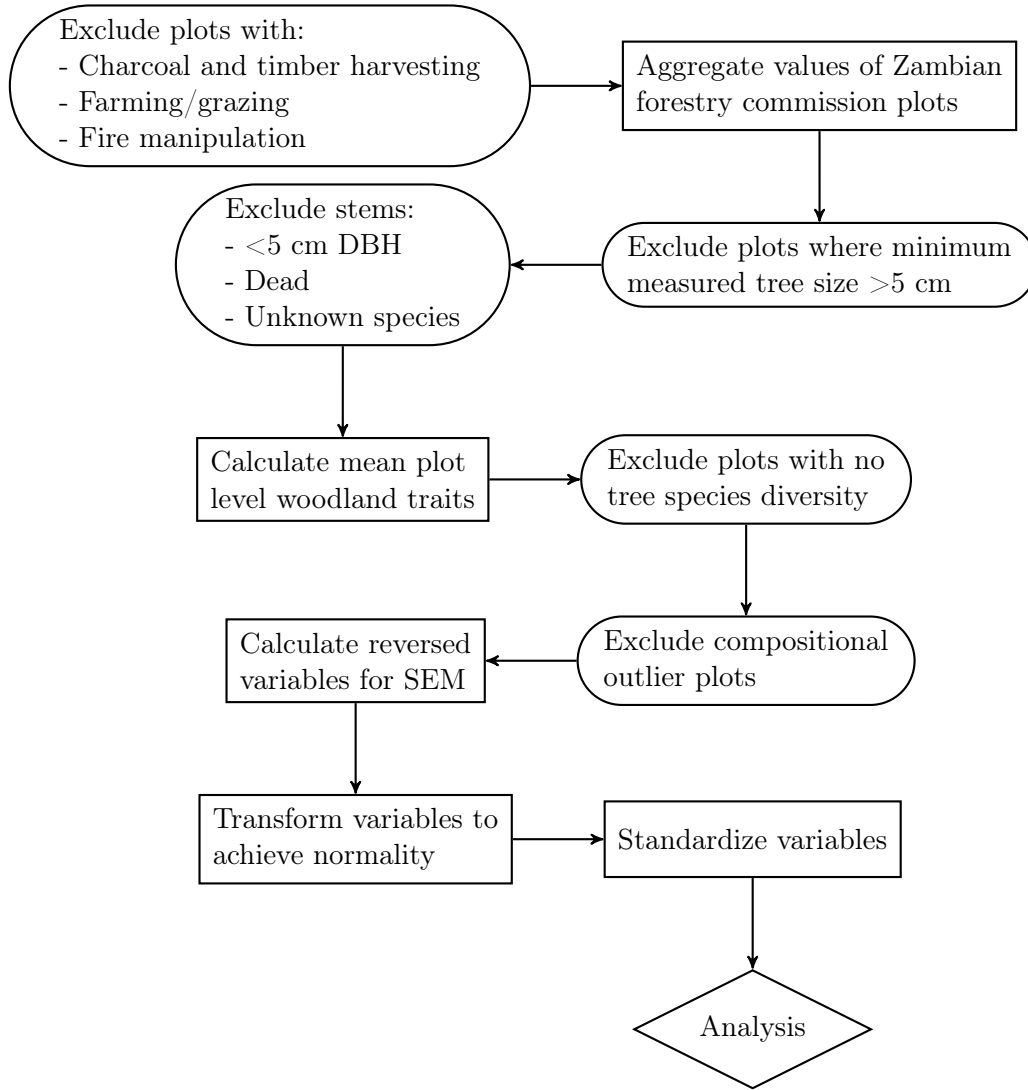


Figure 12: Flow diagram of the data filtering and cleaning process prior to analysis. Rounded boxes indicate filtering events while regular boxes indicate calculation events.

Appendix 2 - Table of correlation fit statistics

Table 3: Table of correlation fit statistics for each pairwise Pearson correlation test of observed variables used in Structural Equation Models.

| X | Y | ρ | lower 95% CI | upper 95% CI | n | Prob. |
|--------|-----------------|--------|--------------|--------------|------|---------|
| Sand % | Org. C % | -0.500 | -0.540 | -0.470 | 1766 | p <0.01 |
| Sand % | CEC | -0.560 | -0.600 | -0.530 | 1766 | p <0.01 |
| Sand % | MAP | -0.490 | -0.520 | -0.450 | 1766 | p <0.01 |
| Sand % | MAP CoV | 0.330 | 0.280 | 0.370 | 1766 | p <0.01 |
| Sand % | MAT | 0.290 | 0.250 | 0.330 | 1766 | p <0.01 |
| Sand % | MAT CoV | 0.220 | 0.180 | 0.270 | 1766 | p <0.01 |
| Sand % | Sp. rich. | -0.360 | -0.400 | -0.310 | 1766 | p <0.01 |
| Sand % | Shannon equit. | -0.300 | -0.340 | -0.260 | 1766 | p <0.01 |
| Sand % | Tree height CoV | -0.120 | -0.170 | -0.060 | 1326 | p <0.01 |
| Sand % | DBH CoV | -0.190 | -0.230 | -0.140 | 1766 | p <0.01 |

| | | | | | | |
|----------|-----------------|--------|--------|--------|------|----------|
| Sand % | Stems ha | -0.270 | -0.310 | -0.230 | 1766 | p < 0.01 |
| Sand % | AGB | -0.230 | -0.280 | -0.190 | 1766 | p < 0.01 |
| Org. C % | CEC | 0.300 | 0.250 | 0.340 | 1766 | p < 0.01 |
| Org. C % | MAP | 0.440 | 0.400 | 0.480 | 1766 | p < 0.01 |
| Org. C % | MAP CoV | -0.310 | -0.350 | -0.270 | 1766 | p < 0.01 |
| Org. C % | MAT | -0.310 | -0.350 | -0.270 | 1766 | p < 0.01 |
| Org. C % | MAT CoV | -0.140 | -0.190 | -0.100 | 1766 | p < 0.01 |
| Org. C % | Sp. rich. | 0.260 | 0.210 | 0.300 | 1766 | p < 0.01 |
| Org. C % | Shannon equit. | 0.220 | 0.180 | 0.270 | 1766 | p < 0.01 |
| Org. C % | Tree height CoV | 0.070 | 0.020 | 0.130 | 1326 | p < 0.01 |
| Org. C % | DBH CoV | 0.140 | 0.090 | 0.190 | 1766 | p < 0.01 |
| Org. C % | Stems ha | 0.170 | 0.120 | 0.210 | 1766 | p < 0.01 |
| Org. C % | AGB | 0.250 | 0.200 | 0.290 | 1766 | p < 0.01 |
| CEC | MAP | -0.010 | -0.060 | 0.040 | 1766 | p = 0.66 |
| CEC | MAP CoV | -0.510 | -0.540 | -0.480 | 1766 | p < 0.01 |
| CEC | MAT | 0.190 | 0.150 | 0.240 | 1766 | p < 0.01 |
| CEC | MAT CoV | 0.030 | -0.020 | 0.070 | 1766 | p = 0.26 |
| CEC | Sp. rich. | -0.020 | -0.060 | 0.030 | 1766 | p = 0.46 |
| CEC | Shannon equit. | 0.220 | 0.170 | 0.260 | 1766 | p < 0.01 |
| CEC | Tree height CoV | 0.040 | -0.010 | 0.090 | 1326 | p = 0.15 |
| CEC | DBH CoV | 0.240 | 0.190 | 0.280 | 1766 | p < 0.01 |
| CEC | Stems ha | 0.200 | 0.150 | 0.240 | 1766 | p < 0.01 |
| CEC | AGB | 0.030 | -0.020 | 0.080 | 1766 | p = 0.23 |
| MAP | MAP CoV | -0.120 | -0.160 | -0.070 | 1766 | p < 0.01 |
| MAP | MAT | -0.120 | -0.170 | -0.070 | 1766 | p < 0.01 |
| MAP | MAT CoV | -0.690 | -0.720 | -0.670 | 1766 | p < 0.01 |
| MAP | Sp. rich. | 0.430 | 0.390 | 0.470 | 1766 | p < 0.01 |
| MAP | Shannon equit. | 0.100 | 0.060 | 0.150 | 1766 | p < 0.01 |
| MAP | Tree height CoV | 0.250 | 0.200 | 0.300 | 1326 | p < 0.01 |
| MAP | DBH CoV | 0.150 | 0.100 | 0.190 | 1766 | p < 0.01 |
| MAP | Stems ha | 0.070 | 0.030 | 0.120 | 1766 | p < 0.01 |
| MAP | AGB | 0.210 | 0.170 | 0.260 | 1766 | p < 0.01 |
| MAP CoV | MAT | -0.170 | -0.210 | -0.120 | 1766 | p < 0.01 |
| MAP CoV | MAT CoV | 0.070 | 0.020 | 0.120 | 1766 | p < 0.01 |
| MAP CoV | Sp. rich. | -0.010 | -0.060 | 0.040 | 1766 | p = 0.71 |
| MAP CoV | Shannon equit. | -0.240 | -0.280 | -0.200 | 1766 | p < 0.01 |
| MAP CoV | Tree height CoV | -0.010 | -0.060 | 0.050 | 1326 | p = 0.77 |
| MAP CoV | DBH CoV | -0.340 | -0.380 | -0.300 | 1766 | p < 0.01 |
| MAP CoV | Stems ha | -0.260 | -0.300 | -0.210 | 1766 | p < 0.01 |
| MAP CoV | AGB | -0.100 | -0.140 | -0.050 | 1766 | p < 0.01 |
| MAT | MAT CoV | -0.310 | -0.350 | -0.270 | 1766 | p < 0.01 |
| MAT | Sp. rich. | -0.180 | -0.220 | -0.130 | 1766 | p < 0.01 |
| MAT | Shannon equit. | 0 | -0.050 | 0.040 | 1766 | p = 0.9 |
| MAT | Tree height CoV | 0.100 | 0.050 | 0.160 | 1326 | p < 0.01 |
| MAT | DBH CoV | 0.140 | 0.100 | 0.190 | 1766 | p < 0.01 |
| MAT | Stems ha | 0.050 | 0.010 | 0.100 | 1766 | p < 0.05 |
| MAT | AGB | -0.080 | -0.130 | -0.030 | 1766 | p < 0.01 |
| MAT CoV | Sp. rich. | -0.340 | -0.380 | -0.300 | 1766 | p < 0.01 |
| MAT CoV | Shannon equit. | -0.130 | -0.180 | -0.090 | 1766 | p < 0.01 |
| MAT CoV | Tree height CoV | -0.220 | -0.280 | -0.170 | 1326 | p < 0.01 |
| MAT CoV | DBH CoV | -0.140 | -0.190 | -0.100 | 1766 | p < 0.01 |
| MAT CoV | Stems ha | -0.150 | -0.200 | -0.110 | 1766 | p < 0.01 |
| MAT CoV | AGB | -0.140 | -0.190 | -0.100 | 1766 | p < 0.01 |

| | | | | | | |
|-----------------|-----------------|-------|-------|-------|------|------------|
| Sp. rich. | Shannon equit. | 0.600 | 0.570 | 0.630 | 1766 | $p < 0.01$ |
| Sp. rich. | Tree height CoV | 0.280 | 0.230 | 0.330 | 1326 | $p < 0.01$ |
| Sp. rich. | DBH CoV | 0.320 | 0.270 | 0.360 | 1766 | $p < 0.01$ |
| Sp. rich. | Stems ha | 0.510 | 0.480 | 0.540 | 1766 | $p < 0.01$ |
| Sp. rich. | AGB | 0.490 | 0.460 | 0.530 | 1766 | $p < 0.01$ |
| Shannon equit. | Tree height CoV | 0.100 | 0.040 | 0.150 | 1326 | $p < 0.01$ |
| Shannon equit. | DBH CoV | 0.250 | 0.200 | 0.290 | 1766 | $p < 0.01$ |
| Shannon equit. | Stems ha | 0.610 | 0.580 | 0.640 | 1766 | $p < 0.01$ |
| Shannon equit. | AGB | 0.430 | 0.390 | 0.470 | 1766 | $p < 0.01$ |
| Tree height CoV | DBH CoV | 0.580 | 0.550 | 0.620 | 1326 | $p < 0.01$ |
| Tree height CoV | Stems ha | 0.170 | 0.110 | 0.220 | 1326 | $p < 0.01$ |
| Tree height CoV | AGB | 0.290 | 0.240 | 0.340 | 1326 | $p < 0.01$ |
| DBH CoV | Stems ha | 0.330 | 0.280 | 0.370 | 1766 | $p < 0.01$ |
| DBH CoV | AGB | 0.450 | 0.410 | 0.480 | 1766 | $p < 0.01$ |
| Stems ha | AGB | 0.560 | 0.530 | 0.590 | 1766 | $p < 0.01$ |

Appendix 3 - Frequency distribution of observed variables

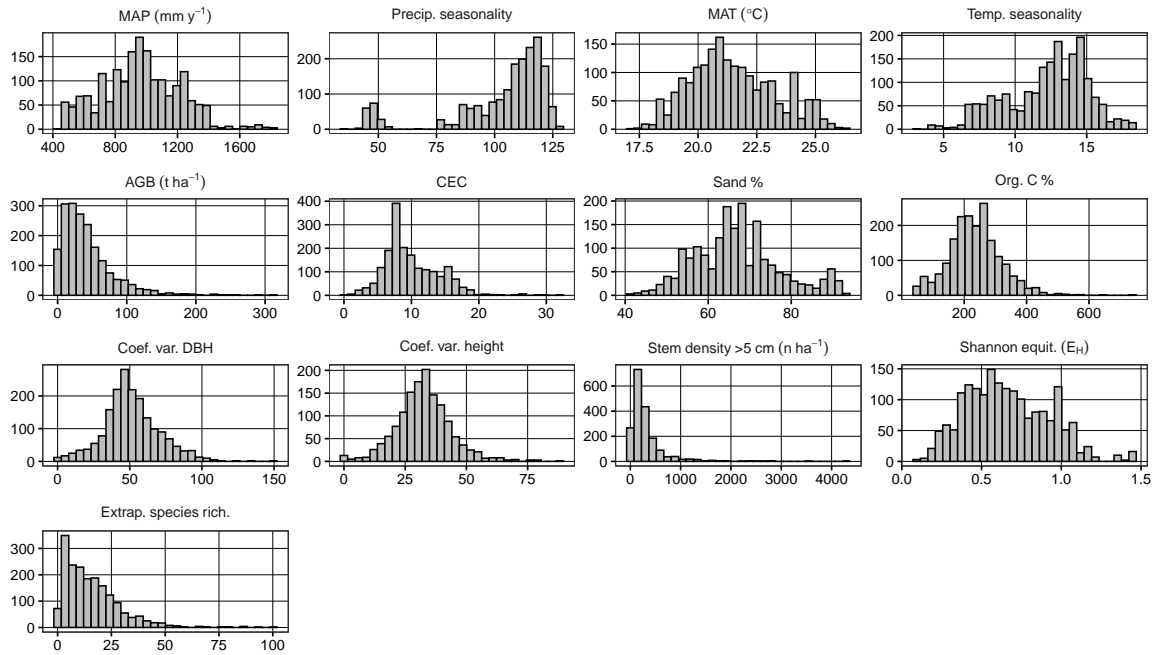


Figure 13: Histograms of raw untransformed observed variables used in final analyses.

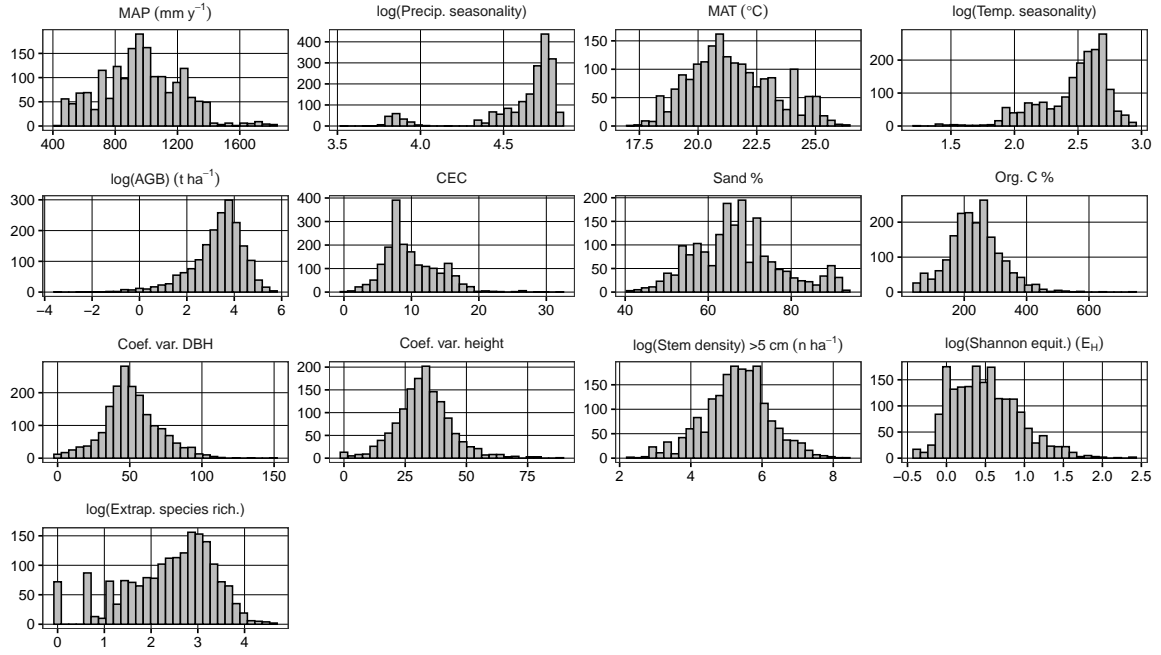


Figure 14: Histograms of observed variables transformed to achieve a normal frequency distribution.

Appendix 4 - Regression tables for moderation by environmental variables

Table 4: Regression fit for a linear multiple regression including the latent variables of moisture availability, tree species diversity and their interaction term on AGB.

| | bchave_std |
|---------------------------------|-------------------|
| Tree sp. div. | 0.512*** (0.022) |
| Moisture avail. | 0.054** (0.022) |
| Tree sp. div. * Moisture avail. | 0.055*** (0.020) |
| Constant | -0.021 (0.022) |
| Observations | 1,766 |
| R ² | 0.277 |
| Residual Std. Error | 0.851 (df = 1762) |
| F Statistic | 224.932*** |

*p<0.1; **p<0.05; ***p<0.01

Table 5: Regression fit for a linear multiple refression including the latent variables of soil fertility, tree species diversity and their interaction term on AGB.

| | bhave_std |
|----------------------------|-------------------|
| Tree sp. div. | 0.480*** (0.022) |
| Soil fert. | 0.120*** (0.022) |
| Tree sp. div. * Soil fert. | 0.036* (0.021) |
| Constant | −0.014 (0.022) |
| Observations | 1,766 |
| R ² | 0.284 |
| Residual Std. Error | 0.847 (df = 1762) |
| F Statistic | 233.151*** |

*p<0.1; **p<0.05; ***p<0.01