

1 Title: Structural diversity and tree density drives variation in the biodiversity-
2 ecosystem function relationship of woodlands and savannas

3 Summary

- 4 • Positive biodiversity-ecosystem function relationships (BEFRs) have been widely documented,
5 but it is unclear if BEFRs should be expected in disturbance-driven systems. Disturbance
6 may limit competition and niche differentiation, which are frequently posited to underlie BE-
7 FRs. We explore the relationship between tree species diversity and biomass, one measure of
8 ecosystem function, across southern African woodlands and savannas, an ecological system
9 rife with disturbance from fire, herbivores and humans.
- 10 • We used >1000 vegetation plots distributed across 10 southern African countries, and struc-
11 tural equation modelling, to determine the relationship between tree species diversity and
12 aboveground woody biomass, accounting for interactive effects of resource availability, dis-
13 turbance by fire, stem density and vegetation type.
- 14 • We found positive effects of tree species diversity on aboveground biomass, operating via
15 increased structural diversity. The observed BEFR was highly dependent on stem density,
16 with a minimum threshold of c. 180 stems ha⁻¹. We found that resource availability mainly
17 affects biomass indirectly, via increasing species diversity.
- 18 • The study underlines the close association between tree diversity, ecosystem structure, en-
19 vironment and function in highly disturbed savannas and woodlands. We suggest that tree
20 diversity is an under-appreciated determinant of savanna and woodland structure and func-
21 tion.

22 **Keywords:** biodiversity, biomass, ecosystem function, forest structure, miombo, savanna, struc-
23 tural equation modelling, woodland.

24 1 Introduction

25 Understanding the relationship between biodiversity and ecosystem function has become a cent-
26 ral endeavour in ecological science, as we seek to predict the consequences of global biodiversity
27 change (Naeem, Duffy & Zavaleta, 2012). Over the past two decades, study of the Biodiversity-
28 Ecosystem Function Relationship (BEFR) has grown from small-scale experimental studies mostly
29 in temperate grasslands (Cardinale et al., 2009; Tilman & Downing, 1994; Tilman, Isbell & Cowles,
30 2014), to observational studies in natural ecosystems (van der Plas, 2019). While positive BEFRs
31 which align with theory have been frequently reported, as research has expanded a complex pic-
32 ture has emerged whereby the strength and direction of the BEFR varies depending on the eco-
33 system studied (Liang et al., 2016), the ecosystem function(s) of interest (Hector & Bagchi, 2007),
34 and the inclusion of environmental covariates in statistical models (Vilà et al., 2005). The goal
35 now should be to study the BEFR in different environmental and ecological contexts, in order to

36 develop an ecosystem-agnostic understanding of the complex interactions between biodiversity,
37 abiotic environment, and ecosystem function.

38 Ecosystem functions are defined in broad terms as rate processes and aggregate properties of eco-
39 systems that describe the nature of biotic activity within those ecosystems (Jax, 2005). Woody
40 productivity and biomass storage are two of the most commonly studied ecosystem functions in
41 forested ecosystems (Brockerhoff et al., 2017), due to their importance in the global carbon cycle
42 (Pan et al., 2011) and their potential leverage as a tool to mitigate the effects of anthropogenic
43 climate change while maintaining biodiversity (Pichancourt, Firn, Chadès & Martin, 2013). The-
44 ory predicts that biodiversity will have a positive effect on ecosystem function via three principle
45 mechanisms: 1) niche complementarity, whereby niche partitioning and ecosystem resource use ef-
46 ficiency is increased in diverse communities, minimising negative density dependent effects such
47 as intraspecific competition and abundance of natural enemies (Barry et al., 2019); 2) selection
48 effects, whereby diverse communities are more likely to include a high-yield species; and 3) facilit-
49 ation effects, whereby diverse communities are more likely to contain species combinations which
50 enhance each others' functional contribution (Wright, Wardle, Callaway & Gaxiola, 2017).

51 There has been extended debate on whether positive BEFRs should be expected in all forest sys-
52 tems (Liang et al., 2016). In temperate and wet tropical forests, where the majority of BEFR
53 studies in natural forest ecosystems have been conducted (van der Plas, 2019), the mechanism of
54 niche complementarity, which contributes the majority of the observed biodiversity effect (Poorter
55 et al., 2015; van der Sande et al., 2017; Wright et al., 2017), hinges on the condition that intraspe-
56 cific competition between trees is the limiting factor to ecosystem functioning, but this may not
57 hold true in all systems.

58 Disturbance-driven mesic savannas and open canopy woodlands cover >20% of the global land
59 surface (Pennington, Lehmann & Rowland, 2018; Solbrig, Medina & Silva, 1996), and represent
60 the dominant vegetation type in Africa, spanning >4 million km² (Hopkins & White, 1987; Rat-
61 nam et al., 2011; Ryan et al., 2016) (Figure 1). Taken together, the above- and below-ground car-
62 bon stored in African mesic savannas is comparable to that found in the wet forests of the Congo
63 basin (Houghton, Hall & Goetz, 2009; Mayaux, Eva, Brink, Achard & Belward, 2008; Spawn, Sul-
64 livan, Lark & Gibbs, 2020). Despite their global importance however, African mesic savannas are
65 severely under-represented in BEFR studies (Clarke, York, Rasheed & Northfield, 2017; Liang et
66 al., 2016). In mesic savannas disturbance by fire (Lehmann et al., 2014) and herbivory (Levick,
67 Asner, Kennedy-Bowdoin & Knapp, 2009; Sankaran, Ratnam & Hanan, 2008) reduces woody stem
68 density and biomass, diminishing competitive interactions between individuals, allowing competit-
69 ors to co-exist where they would normally be excluded (Grime, 1979; Keddy, 1990). Instead, stress
70 tolerance and the functional contribution of particular species (selection effects) may be the pre-
71 dominant biotic forces influencing ecosystem function in these ecosystems (Lasky et al., 2014; Tob-
72 ner et al., 2016). It is possible that a threshold woody stem density exists below which the lack
73 of competition precludes the detection of a biodiversity effect on ecosystem function. Addition-
74 ally, facilitation effects may play a larger role in driving the BEFR in stressful environments like
75 mesic savannas. Ratcliffe et al. (2017) found stronger positive relationships between tree species
76 richness and various ecosystem functions with increasing aridity. They demonstrate that com-
77 petition diminishes where environmental stress limits species' abundances. This potential mis-

match in the contribution of different mechanisms to the BEFR among resource-, disturbance-, and competition-limited ecosystems requires further investigation if we are to derive a generalisable BEFR.

A number of studies in miombo woodlands, the dominant mesic savanna type in southern Africa (Campbell, 1996), found that above-ground woody carbon/biomass stocks correlate positively with tree species richness (McNicol, Ryan, Dexter, Ball & Williams, 2018; Mutowo & Murwira, 2012; Shirima, Pfeifer, Platts, Totland & Moe, 2015). These studies however, lacked the spatial extent required to account for the interacting effects of variation in abiotic environment, disturbance regime, and biogeography which are expected to affect the BEFR. Studies of the BEFR often find that within a small spatial extent (<50 km), biodiversity shows a strong effect on ecosystem function, but at broader extents (>1000 s km) biodiversity effects pale in significance compared to abiotic factors such as climate (Gonzalez et al., 2020). In West Africa, Mensah, Salako and Seifert (2020) found that woodlands and forests showed a positive effect of tree species richness on above-ground carbon, while open savannas did not, implying that stem density and the presence of dominant large trees may affect the strength of the observed BEFR. In the Brazilian Cerrado savanna, Loiola, Scherer-Lorenzen and Batalha (2015) found that disturbance by fire reduced tree productivity via its effect on functional trait values. In contrast however, Carvalho, Batalha, Silva, Cianciaruso and Petchey (2014) also working in the Brazilian Cerrado, found that while disturbance by fire did reduce soil fertility, which is often closely related to productivity in savannas, functional trait diversity was unaffected by fire frequency. Other studies focussing on the herbaceous diversity in disturbance-prone savannas and grasslands in North America and Europe have shown that disturbance by fire, mowing and herbivory reduces herbaceous productivity (Grace et al., 2007), and allows for weak competitors to co-exist (Mason, de Bello, Doležal & Lepš, 2011). While these studies together offer some glimpse into how the BEFR may operate in disturbance-prone systems, none provide a full and comprehensive assessment of the interacting effects of diversity, abiotic environment, disturbance and ecosystem function. Additionally, due to differences in community assembly, evolutionary history and contemporary drivers between the neotropics and the African tropics, inferences from one continent cannot necessarily be applied to the other (Dexter et al., 2015).

As well as variation in disturbance regime, southern African savannas occur over a wide range of precipitation, temperature, and soil conditions (Campbell, 1996). Environmental heterogeneity has been shown to affect both woody biomass and tree species diversity independently, across a number of different biomes (Michaletz, Cheng, Kerkhoff & Enquist, 2014; Michaletz, Kerkhoff & Enquist, 2018; Spasojevic, Grace, Harrison & Damschen, 2014). It is important therefore to account for environmental heterogeneity and understand how it influences both biomass and biodiversity to effectively model and correctly attribute the effects of biodiversity on woody biomass. Sankaran et al. (2005) and Lehmann et al. (2014) both report that total precipitation sets the upper limit for woody biomass in African savannas. Lehmann et al. (2014) also report complex indirect relationships between climate, disturbance by fire and woody biomass, demonstrating the need for directional multi-faceted modelling techniques to properly account for the effects of climate on ecosystem function in this system.

In this study, we make the first known estimation of the Biodiversity-Ecosystem Function Rela-

120 tionship (BEFR) across a sub-continental area of disturbance-driven mesic woody savannas in
 121 southern Africa. We aim to understand the synergistic effects of environmental and biotic drivers
 122 of variation in the BEFR in this system, with a view to generating a general model of the BEFR
 123 in disturbance-prone forested ecosystems, which is currently lacking in the BEFR literature. We
 124 posit three hypotheses: (1) water availability and soil fertility will indirectly positively affect woody
 125 biomass via an increase in tree species diversity, (2) the effect of tree species diversity on woody
 126 biomass will increase with stem density (number of stems ha^{-1}), as competitive interactions be-
 127 come more influential as stem proximity increases. In addition, we expect that an increase in dis-
 128 turbance by fire will decrease stem density and therefore competition, weakening the effect of tree
 129 species diversity on woody biomass. Finally, we expect that (3) tree species diversity will increase
 130 tree structural diversity (i.e. physiognomic diversity), providing an indirect path by which tree di-
 131 versity increases woody biomass.

132 2 Materials and Methods

133 2.1 Study location

134 The study used 1235 woodland monitoring plots from a larger pool of 5395 plots in the larger
 135 SEOSAW database (SEOSAW, 2020), located across 10 countries within southern Africa in the
 136 miombo ecoregion (Figure 1, Hopkins & White, 1987). The study area spans the core climate
 137 space of the region, with a precipitation gradient from c. 460 mm y^{-1} in southern Mozambique
 138 and southern Zimbabwe to c. 1700 mm y^{-1} in northern Zambia, Malawi and northern Mozam-
 139 bique. A 2D convex hull of Mean Annual Precipitation (MAP) and Mean Annual Temperature
 140 (MAT) of the study sites covers 96.5% of the pixel-wise climate space of the miombo woodland
 141 ecoregion (Hopkins & White, 1987), using WorldClim estimates of Mean Annual Temperature
 142 (MAT, BIO1) and Mean Annual Precipitation (MAP, BIO12) between 1970 and 2000 with a pixel
 143 size of 30 arc seconds (926 m at equator) (Fick & Hijmans, 2017).

144 Plots were chosen from the SEOSAW database based on the quality and completeness of data
 145 collection, and plot configuration. Plot vegetation was identified under the broad term of ‘sa-
 146 vanna’, which includes ‘woodland’, ‘savanna woodland’, and ‘tree savanna’, variously defined in
 147 other areas of the scientific literature and here referred to collectively as southern African wood-
 148 lands (Hill & Hanan, 2011; Ratnam et al., 2011). Plots with evidence of farming, human resource
 149 extraction or experimental treatments such as prescribed burning or herbivore exclusion were ex-
 150 cluded from the initial pool. Only plots >0.1 hectares were used in analyses, as area-based bio-
 151 mass estimation from small plots is highly influenced by rare large trees (Stegen et al., 2011), lead-
 152 ing to inaccurate biomass estimates. Only plots with a stem density >50 trees ha^{-1} (>10 cm stem
 153 diameter) were used, to ensure all plots represented woodland rather than ‘grassy savanna’, which
 154 is considered here a separate biome with very different species composition (Parr, Lehmann, Bond,
 155 Hoffmann & Andersen, 2014). 3760 plots within the SEOSAW database were arranged in clusters
 156 of four 20x50 m plots, with 20 m between plots. Plots within each spatial cluster were combined
 157 and treated as a single plot in analyses, resulting in 940 aggregate plots which were then subject
 158 to the plot filtering process described above.

2.2 Data collection

We considered only trees and shrubs in our calculations of Above-Ground woody Biomass (AGB), including woody species such as palms and cycads, which are functionally tree-like. Woody lianas are scarce in our study plots and were not measured. Only stems >10 cm DBH (Diameter at Breast Height, 1.3 m) were included in analyses. Many plots in the dataset did not include data on stems <10 cm DBH. For those plots which contained stem measurements <10 cm DBH, small stems only accounted for a median of 2.1% of the plot level AGB.

All stems >10 cm DBH were measured within each plot resulting in a total of 66,758 stems with measurements. A tree may be comprised of multiple stems and so tree-level richness estimates, rather than stem-level estimates, were used to prevent bias from species which readily coppice. For each tree, we recorded species, DBH and tree height to the top of the highest branch material. Height was measured through a variety of means including laser rangefinders, manual clinometers and measuring sticks. When DBH could not be measured at 1.3 m due to trunk abnormalities, it was measured at the closest regular portion of the trunk to 1.3 m. The height of this measurement was recorded and used to estimate the DBH_e at 1.3 m using a cubic polynomial regression, with parameters estimated using a test dataset from Ryan C., (unpublished), see Godlee et al. (2020). AGB for each plot (t ha^{-1}) was calculated using Equation 1, taken from Chave et al. (2014):

$$AGB = 0.0673 \times (\rho D^2 H)^{0.976} \quad (1)$$

where ρ is the species mean wood density (g cm^{-3}), D is the DBH_e (cm) at 1.3 m, and H is the tree height (m). Wood density estimates were taken from the global wood density database for each species where possible (Chave et al., 2009; Zanne et al., 2009). Wood density for species without species level estimates was estimated from the means of their respective genera. For stems where tree height was unknown, the plots' climatic parameters, estimated from plot location, were used to estimate tree height, according to Chave et al. (2014).

Climatic data were taken from the WorldClim database, using the BioClim variables (Fick & Hijmans, 2017). In addition to MAT and MAP, temperature stress was calculated as the mean diurnal temperature range (BIO2) and precipitation seasonality was calculated as the mean of the coefficient of variation of monthly mean precipitation (BIO15). Soil fertility data were extracted from the ISRIC gridded soil information data product at 250 m resolution, taking the grid cell value for each plot centre (Hengl et al., 2017). We extracted Cation Exchange Capacity (CEC) (cmolc kg^{-1}), soil organic carbon stocks (kg m^{-2}) percentage soil sand content (0.05-2 mm) by weight and soil nitrogen content (g kg^{-1}). These data are a modelled product derived from various remotely sensed and directly measured data sources. The degree of fire disturbance was calculated using the MODIS monthly burned area product at 500 m resolution (MCD64A1, Giglio, Justice, Boschetti and Roy 2015), counting the total number of times the plot pixel was classified as burning, between 2001 and 2018. We initially aimed to include disturbance by herbivory in our model, including total herbivore biomass from the Hempson, Archibald and Bond (2017) modelled herbivory product, but this inclusion prevented models from converging due to its collinearity with other observed variables, notably MAP and disturbance by fire.

2.3 Data analysis

2.3.1 Species diversity and structural diversity metrics

Estimated tree species richness was calculated for each plot using ‘ChaoRichness()’ from the ‘iN-EXT’ package in R (Hsieh, Ma & Chao, 2016). This procedure uses Hill numbers of the order $q = 0$ (Jost, 2006) to extrapolate a species rarefaction curve to its predicted asymptote and uses this value as its estimated species richness value. Extrapolated species richness accounts for variation in plot size present in the dataset (0.1-10 ha) and therefore sampling effort among plots. Larger plots will tend to encompass more individuals, and therefore more species (Dengler, 2009). To measure tree species evenness, the Shannon Equitability index ($E_{H'}$) (Smith & Wilson, 1996) was calculated as the ratio of the estimated Shannon diversity index to the natural log of estimated species richness. Abundance evenness allows for greater niche complementarity at small scales due to potentially increased heterogeneity of functional traits. We quantified tree structural diversity for each plot by calculating the Coefficient of Variation of DBH (DBH CoV) and tree height (Height CoV).

2.3.2 Vegetation clusters

Plots were assigned to vegetation type groups based on tree species composition. Groups were defined in a manner adapted from Fayolle et al. (2018) in an Africa-wide analysis of floristic units using plot data in savannas and woodlands with tree species diversity and relative abundance data. Group identification was conducted using unconstrained correspondence analysis, followed by hierarchical clustering based on dominant ordination axes. Plot data used in this study occurred in four compositional vegetation types. See Table 1 for a description of each vegetation cluster and Figure 1 for the spatial distribution of plots from each of these clusters. Cluster names were assigned post-hoc based on the dominant and indicator species in each cluster.

2.3.3 Structural Equation Modelling

We used Structural Equation Modelling (SEM) to investigate the determinants of AGB. All SEMs were constructed and analysed in the ‘lavaan’ package (Rosseel, 2012) in R version 3.6.0 (R Core Team, 2019). SEM was used because of its suitability for modelling complex causal interactions in ecological systems (Lee, 2007). A key aspect to our decision to use SEM is that they can explicitly model and partition variance attributed to indirect effects, which is challenging in standard multiple regressions. Using SEMs also allowed us to describe latent variables such as ‘water availability’, ‘soil fertility’, and ‘disturbance’ which have been suggested to act upon biodiversity and biomass/productivity in previous studies despite these factors not having directly observable measures in our dataset. SEM is also necessary to properly account for potential feedback mechanisms between aspects of environment and tree species diversity, which could otherwise increase the chances of Type I error and wrongly attribute inference due to the covariance of explanatory variables when using conventional regression analyses (Nachtigall, Kroehne, Funke & Steyer, 2003).

234 We specified a conceptual model with factors expected to affect AGB: water availability, soil fertil-
 235 ity, disturbance, tree species diversity, tree structural diversity and stem density (Figure 2).
 236 Observed variables were transformed to achieve normality where necessary and standardised to Z-
 237 scores prior to analysis (Fig. S1, Fig. S2). Standardisation allows path regression coefficients to be
 238 easily compared between paths in the same model to assess their relative effect size, and elimin-
 239 ates confusion in model interpretation arising from the observed variables being on different scales
 240 (Beaujean, 2014). Standardisation also controls for variables with variation across different orders
 241 of magnitude, which could otherwise prevent adequate model estimation from the covariance mat-
 242 rix in ‘lavaan’. To ensure that observed variables within a latent variable had consistent directions
 243 of influence, some observed variables had their sign reversed. For example, overall water availab-
 244 ility is expected to decrease as soil sand content increases, therefore sand content was reversed for
 245 use in the water availability latent variable. Precipitation seasonality, and temperature stress were
 246 also reversed in this way to account for the direction of their effect on water availability.
 247 The factor loadings of the observed variable assumed to contribute most to each latent variable
 248 were set to one, as per convention, with other observed variables being allowed to vary (Beaujean,
 249 2014). We tested the robustness of our assumptions with a chi-squared test of all possible com-
 250 binations of observed variable factor loadings set to one, while ensuring no factor loadings were in
 251 excess of one. We found no significant difference between model specifications ($p > 0.05$). Full In-
 252 formation Maximum Likelihood (FIML) was used in each model to estimate the values of missing
 253 data in each latent variable (Cham, Reshetnyak, Rosenfeld & Breitbart, 2017).
 254 We assessed the role of tree species diversity and tree structural diversity in determining AGB via
 255 a simple mediation model which allowed species diversity to influence AGB both directly and in-
 256 directly via structural diversity. Structural diversity can also directly influence AGB in this model,
 257 separate to the effect of species diversity. To account for variation in stem density, which may
 258 covary with species diversity, we included it as an observed variable in our model. To explore vari-
 259 ation in the model among woodland vegetation types, we fit the model both at the regional scale
 260 and for each vegetation type separately. We compared unstandardised path coefficients among the
 261 models for different vegetation types to understand the effect that vegetation type has on the re-
 262 lationship between tree species diversity, structural diversity, stem density and AGB. Path coeffi-
 263 cients show the effect of a given path with other paths held constant. Models were estimated using
 264 the ‘MLM’ estimator, because it is robust to multivariate non-normality (Shapiro, 1983). Model
 265 fit was evaluated using the robust Comparative Fit Index (CFI), the robust Tucker Lewis Index
 266 (TLI), the Root Mean Squared Error of Approximation (RMSEA) and the R^2 coefficient of de-
 267 termination for AGB. We critically assessed model fit in each case, taking into consideration the
 268 recommendations of Hu and Bentler (1999) who define threshold values of acceptability for these
 269 model fit indices: CFI > 0.85 , TLI > 0.85 , RMSEA < 0.15 , alongside our judgement of the model
 270 estimates.
 271 To explore the hypothesis that biodiversity effects on ecosystem function increase in strength as
 272 stem density increases, we repeatedly sub-sampled the available plot dataset to create 50 data
 273 subsets with similar stem density. For each data subset we separately fitted a model including tree
 274 species and structural diversity latent variables to predict AGB. As we controlled for stem density
 275 via the dataset sub-sampling process, the effect of stem density on AGB was not included in the

276 model. We examined how the unstandardised path coefficients for each path in the SEM varied
277 according to the median stem density of the data subsets.

278 Preliminary models that included herbivore biomass (Hempson et al., 2017) did not converge, pos-
279 sibly due to the spatially coarse nature of the available data, we therefore did not include herb-
280 ivory in our final model. We incorporated environmental covariates into our model to understand
281 the relative effects of water availability, soil fertility and disturbance on AGB both directly and in-
282 directly via species diversity and stem density. We compared standardised path coefficients between
283 paths in the model to understand the relative contribution of each path to explain variance in
284 AGB. Vegetation type specific models could not be reliably fitted for this more complex model
285 specification with environmental covariates, due to sample size issues and because some vegetation
286 types were narrow in their climate space, leading to a lack of environmental variation, particularly
287 in the water availability latent variable.

288 **3 Results**

289 Pairwise correlations between all observed variables used in the Structural Equation Models (SEMs)
290 showed that all tree species diversity (extrapolated tree species richness, Shannon equitability in-
291 dex) and structural diversity (coefficients of variation of DBH and height) variables had moderate
292 positive correlations with AGB (Figure 3, Fig. S3). Stem density had the strongest correlation
293 with AGB of all variables considered ($r = 0.59$, $p < 0.01$). Environmental variables had weaker
294 correlations with AGB than diversity variables, with all environmental variables having signifi-
295 cant correlations with AGB, except fire frequency. The direction of these correlations was used as
296 a test of our assumptions for the direction of influence of latent variables later used in the SEMs.
297 MAP had positive correlations with all tree species diversity and structural diversity variables.
298 Tree species diversity variables had clear positive correlations with stem density (species richness:
299 $r = 0.24$, $p < 0.01$; Shannon equitability: $r = 0.58$, $p < 0.01$), but structural diversity variables
300 showed weak correlations with stem density (DBH CoV: $r = 0.11$, $p < 0.01$, Height CoV: $r = 0.01$,
301 $p = 0.86$).

302 **3.1 Structural and species diversity models**

303 In an SEM describing the effect of tree species diversity on AGB via the mediating effects of tree
304 structural diversity and stem density (Figure 4), species diversity showed no direct effect on AGB
305 ($\beta = 0.01 \pm 0.053$, $p = 0.88$), but did have an indirect positive effect via structural diversity ($\beta =$
306 0.18 ± 0.039 , $p < 0.01$) (Figure 4). Model fit was good with high factor loadings for all observed
307 variables. All other path coefficients were significant ($p < 0.01$) (Table 2). The R^2 of AGB was
308 0.49. The strongest direct effect on AGB was from stem density ($\beta = 0.5 \pm 0.033$, $p < 0.01$).

309 **3.2 Variation among vegetation types**

310 When the tree species and structural diversity model (Figure 4) was refitted separately using data
311 from each of the four vegetation types, we found that the effect sizes of each latent variable re-

312 maintained largely similar, though model fit varied. The direct effect of tree species diversity on AGB
 313 was positive and marginally significant in ex-Acacia ($\beta = 0.16 \pm 0.121$, $p = 0.18$) but negligible in
 314 Mopane ($\beta = 0.24 \pm 0.099$, $p < 0.05$), sparse miombo / *Baikiaea* ($\beta = 0.23 \pm 0.045$, $p < 0.01$) and
 315 Core miombo ($\beta = 0.23 \pm 0.041$, $p < 0.01$) (Figure 5). Relationships between structural diversity
 316 and AGB remained generally similar, with the same sign and overlap between the 95% confidence
 317 intervals of path coefficients. The R^2 of AGB was highest in ex-Acacia shrubland ($R^2 = 0.83$) and
 318 lowest in sparse miombo / *Baikiaea* ($R^2 = 0.46$). The total effect of species diversity on AGB re-
 319 mained strongly positive and there was a positive direct effect of species diversity on structural
 320 diversity, across all vegetation types. All models had adequate goodness-of-fit (Table 2), though
 321 confidence intervals around the unstandardised path coefficients were wide particularly for Mopane
 322 and ex-Acacia. χ^2 statistics were high for some vegetation types, but this appears to be highly
 323 correlated with sample size for each vegetation type (Hooper, Coughlan & Mullen, 2008).

324 **3.3 Moderation of Diversity-AGB relationship by stem density**

325 In the sub-sampling of the plot dataset by stem density, we found an increasing positive effect of
 326 tree species diversity on AGB as stem density increased (Figure 6e). There appears to be a min-
 327 imum stem density threshold at c. 180 trees ha^{-1} below which there appears to be a reasonably
 328 constant baseline effect of tree diversity on biomass (Figure 6b). The effect of structural diversity
 329 on AGB appears to remain constant with increasing stem density (Figure 6d). The indirect effect
 330 of tree species diversity on AGB via structural diversity increases as stem density increases (Fig-
 331 ure 6c).

332 **3.4 Environmental covariates and tree diversity**

333 A model incorporating the latent variables of water availability, soil fertility and disturbance by
 334 fire showed that the total effect of tree species diversity on biomass was similar to that of water
 335 availability, soil fertility and disturbance (Figure 7, Fig. S4). The direct effects of water availabil-
 336 ity, soil fertility and disturbance on AGB were negligible (water: $\beta = 0.1 \pm 0.13$, $p = 0.43$, soil: $\beta =$
 337 0.1 ± 0.155 , $p = 0.51$, disturbance: $\beta = -0.04 \pm 0.043$, $p = 0.32$), with nearly all of their observed ef-
 338 fects on AGB coming from the indirect paths via stem density (water: $\beta = 0.14 \pm 0.091$, $p = 0.12$,
 339 soil: $\beta = -0.22 \pm 0.109$, $p < 0.05$, disturbance: $\beta = -0.12 \pm 0.03$, $p < 0.01$) and species diversity (wa-
 340 ter: $\beta = 0.62 \pm 0.172$, $p < 0.01$, soil: $\beta = -0.24 \pm 0.209$, $p = 0.26$, disturbance: $\beta = 0.19 \pm 0.058$, p
 341 < 0.01). MAP and soil sand content had the greatest contributions to the latent variable of wa-
 342 ter availability. Model fit was acceptable: CFI = 0.925, TLI = 0.900, and RMSEA = 0.153, R^2 of
 343 AGB = 0.34.

344 Similar to the model that only considered tree species and structural diversity (Figure 4), the dir-
 345 ect effect of species diversity on structural diversity was positive, while structural diversity itself
 346 had a positive effect on AGB, leading to a strong positive indirect effect of species diversity on
 347 AGB via structural diversity ($\beta = 0.19 \pm 0.026$, $p < 0.01$) when environmental covariates were ac-
 348 counted for. Again, the direct effect of species diversity on AGB was negligible ($\beta = -0.05 \pm 0.041$,
 349 $p = 0.27$). The total effect of species diversity on AGB was positive ($\beta = 0.34 \pm 0.044$, $p < 0.01$).
 350 Compared to the simple model with no environmental covariates, the total explanatory power of

tree species diversity and structural diversity in this model decreased, but the predictive power of the model as a whole increased.

4 Discussion

We assessed the importance of a) tree species diversity, b) tree structural diversity, c) resource availability, d) disturbance by fire, e) stem density and their interactions on above-ground woody biomass (AGB) across southern African woodlands, using a network of 1235 woodland plots in conjunction with Structural Equation Modelling (SEM). We found support for a general positive relationship between tree species diversity and AGB, operating indirectly via structural diversity (H_1). Tree species diversity, structural diversity and stem density accounted for 49% of the variation in AGB across the region, while models for specific vegetation types showed even greater explanatory power in some cases (Table 2). We found that the effect of tree species diversity on AGB increased with stem density (H_2), with an apparent threshold of 180 stems ha^{-1} below which the effect of species diversity on AGB remained at a low baseline level. The strongest direct effect on AGB was that of stem density. When the effects of water availability, soil fertility and disturbance by fire were controlled for, the total explanatory power of tree species diversity and structural diversity decreased, but the predictive power of the model increased, suggesting that it is important to control for environmental covariates to understand the true effect of tree species diversity on AGB in regional scale assessments in southern African woodlands.

4.1 Inter-related effects of tree species and structural diversity on AGB

We found a consistent positive effect of tree species diversity on AGB. Within southern African woodlands we therefore find support for the hypothesis that higher tree species richness and evenness leads to higher above-ground woody biomass. This finding is in agreement with many other studies across different ecosystems and biomes, supporting the idea that there is a generalisable positive association between biodiversity and ecosystem function (Cardinale et al., 2009; Liang et al., 2016). Our study provides a novel dissection of the mechanisms underlying this relationship, particularly in the context of southern African woodlands, a disturbance-driven and poorly studied ecological system.

Much of the total variation in AGB was driven by variation in stem density. It is possible that within southern African woodlands a higher species diversity allows for a higher stem density through niche separation, which reduces competition between species occupying varying niche space, leading to an increase in total AGB per unit area. The opposite causation is also plausible however, with increased stem density causing higher species richness through an increased probability of encountering new species. We attempted to correct for the correlation between species richness and stem density using extrapolated species richness, which extrapolates a rarefaction curve to its predicted asymptote, thus estimating the total landscape-level species richness which is unaffected by plot size and stem density. We suggest therefore that an increase in tree species diversity through species richness and evenness produces an assemblage of species which can utilise more available light and moisture, resulting in greater plot-level AGB. This is supported by the

moderately strong indirect positive effect of tree species diversity on AGB via structural diversity, and the positive effect of water availability on AGB via stem density in the model which included environmental covariates.

We found evidence that tree species diversity led to an increase in AGB indirectly via tree structural diversity, and we therefore find support for our second hypothesis H_2 . A higher tree species diversity allows for a greater structural diversity of trees, i.e. greater variation in DBH and height. This may act as a mechanism for niche complementarity, with a canopy of diversely sized trees able to take advantage of a greater proportion of the available light. Although we did not measure them here, we would also expect that tree species diversity allows for a greater range of tree functional forms (Pretzsch, 2014), i.e. wider variation in canopy shape and overall growth form; broad flat crowns vs. narrow deep crowns, for example. In forests, where the tree canopy is effectively closed, as the stand matures a more diverse canopy emerges via competition and tree mortality events which open canopy gaps (Muscolo, Bagnato, Sidari & Mercurio, 2014). Indeed, our finding that the strength of the effect of tree diversity on AGB increases with stem density supports this mechanism. In frequently disturbed woodlands such as those studied here however, a woodland canopy similar to that of a forest is frequently not reached. Instead, a simple open canopy is maintained that can be made more complex and productive via an increase in species diversity. Previous studies have found that southern African woodlands with higher species diversity tend to experience less frequent disturbance by fire and tend to form a more closed canopy and a more sparse understorey (Chidumayo, 2013; Mutowo & Murwira, 2012). In our study however, we found a positive effect of disturbance on species diversity, perhaps suggesting that disturbance prevents domination of woodlands by a single dominant species (Chidumayo, 2013).

We found a non-linear positive effect of stem density on the relationship between tree species diversity and AGB (Figure 6). At low stem densities, competition between mature trees may not occur, meaning that the niche complementarity effect provided by an increase in tree species richness may not be present, accounting for the small effect of tree species diversity on AGB below c. 180 trees ha^{-1} . At very high stem density, there is also an increase in the effect of species diversity on structural diversity. This could be because at high stem density, the adaptation of different species to growth form become important. At low stem density, individual trees tend to spread out rather than growing tall, but at high stem density, only certain species are able to exist in the understorey, while others are able to grow tall above the woodland canopy, leading to greater variation in tree height over the plot.

Within the latent variable of tree species diversity we found similarly strong factor loadings for both species richness and abundance evenness. This demonstrates that species richness and abundance evenness measure different and largely uncorrelated aspects of diversity. In terms of the theory of niche complementarity, both species richness and abundance evenness are important to produce the positive diversity effect on ecosystem function. Holding either species richness or abundance evenness constant while increasing the other will have the identical effect of reducing the likelihood that a neighbour is a conspecific, thus reducing the occurrence of negative density dependent competition effects and therefore increasing ecosystem function.

4.2 Effects of water availability, soil fertility and disturbance

Water availability had a positive total effect on AGB, comparable in size to the total effect of tree species diversity on AGB, while soil fertility had a negative total effect. We expected that higher water availability and soil fertility would lead to higher AGB under the assumption that higher resource availability would allow for a greater stem density per unit area, greater productivity per unit area and additionally greater tree species diversity due to niche partitioning (Kraaij & Ward, 2006; Shirima et al., 2015). Previous studies in tropical forests have shown that water availability increases AGB both directly and indirectly via increasing tree species diversity and via increasing stand structural diversity (Ali et al., 2019a, 2019b; Poorter et al., 2017). In this study, we observed indirect positive effects of water availability on AGB via species diversity and a positive but only marginally significant direct effect on AGB. Compared to moist tropical forests, water availability is more of a limiting factor to tree growth in southern African woodlands, which experience frequent drought.

Disturbance by fire had a negative total effect on AGB, with most of this negative effect coming from the indirect pathway via stem density. This is expected as increased fire frequency is a key mechanism by which mesic savannas maintain an open canopy, rather than shift to a closed canopy forest (Staver, Archibald & Levin, 2011). There were small positive effects of disturbance by fire on tree species diversity, a finding which is supported by studies elsewhere (Durigan et al., 2020; Staver, Bond, Stock, van Rensburg & Waldram, 2009). It is suggested that in mesic savannas where the tree-species pool is largely adapted to fire, increased fire may actually increase tree species diversity by allowing weak competitors to co-exist.

A negative total effect of soil fertility on AGB is in contrast to other studies in the region and against general ecological theory, which predicts a positive effect of soil nutrients on biomass. The negative total effect of soil fertility on AGB was driven mostly by an indirect negative effect via stem density. The direct effect on AGB however, remained positive and marginally significant, as expected. Model estimates of the effect of soil on AGB were poorly constrained compared with other latent variables. This wide standard error on the model predictions is possibly due to the coarseness and nature of the soil data we used. SoilGrids provides modelled data at 250 m resolution, while soil structure and nutrient content varies at much finer scales in southern African woodlands (Bucini & Hanan, 2007; Muledi et al., 2017), notably by termite mounds which offer small areas of highly fertile soil with peculiar tree species diversity (Joseph, Seymour, Cumming, Cumming & Mahlangu, 2014). Due to the plots used in this study often being situated non-randomly in the landscape, coupled with the coarseness of the SoilGrids data, it is not surprising that this model path is poorly constrained. Soil data is time-consuming to collect and difficult to compare across studies when different protocols are used, though this study prompts the need for further effort in this regard, which may reveal interesting findings about the complex interactions between soil, disturbance and tree diversity in southern African woodlands. Lehmann et al. (2014) found similarly weak and poorly constrained relationships for soil in a Structural Equation Model including precipitation, temperature, soil, and fire to predict tree basal area.

4.3 Vegetation type responses

All four vegetation types produced similar results in the simple SEM, with a positive total effect of species diversity on AGB, the majority being indirectly via structural diversity. This demonstrates the robustness of our results, showing they are generalisable across vegetation types in southern Africa. It also demonstrates that similar ecosystem processes are occurring in these vegetation types, despite variation in species composition, overall species richness and mean biomass. Core miombo and sparse miombo / *Baikiaea* woodland vegetation exhibited a small negative direct effect of tree species diversity on AGB, while the total effect, incorporating the indirect effect via structural diversity, remained positive in these vegetation types. Compared to ex-Acacia and Mopane woodlands, miombo woodlands have higher median tree species richness. Ex-Acacia and Mopane woodlands are dominated by fewer tree species, notably *Senegalia* spp. in ex-Acacia woodlands and *Colophospermum mopane* in Mopane woodlands which often produce large canopy dominating trees. We postulate that the slight negative effect of tree species richness on AGB in miombo woodlands may be due to an increase in interspecific competition through canopy crowding, but that this effect is not present in ex-Acacia and Mopane woodlands, where the top level of the woodland canopy is dominated often by a single species.

Higher functional redundancy among tree species in miombo woodlands may lead to smaller trees with lower AGB in the most diverse plots, more resembling thicket vegetation and suppressing the few species which tend to create high biomass, such as *Julbernardia* and *Brachystegia* spp.. In the species-poor Mopane and ex-Acacia woodlands however, the addition of extra species may fill a greater proportional niche space, thus increasing total AGB more.

Despite Mopane woodland having very low species diversity generally, with often monospecific stands (Timberlake, Chidumayo & Sawadogo, 2010), a positive effect of tree species diversity on AGB was observed. In previous studies across ecosystem types it has been found often that the effect on ecosystem function of adding species is stronger in low diversity assemblages (Cardinale et al., 2006; Srivastava & Vellend, 2005). This has been attributed to an increase in functional redundancy as species diversity increases. In other words, with more species, it is more likely that the addition of a new species will occupy the same ecological niche space as an existing species, meaning niche complementarity will not occur and competition will not lead to niche partitioning, making little difference to overall ecosystem functioning. Mopane woodlands also have a negligible effect of species diversity on structural diversity. This may be due to the species which tend to co-exist with *C. mopane*, many of which are small shrub-like trees and which do not grow into large canopy trees (Timberlake et al., 2010). Larger canopy trees tend to have greater variation in physical structure (Seidel et al., 2019).

Ex-Acacia woodlands showed the strongest total effect of species diversity on AGB and was the only vegetation type to show a significant positive direct effect of species diversity on AGB. Ex-Acacia woodlands also had relatively low median species richness compared to miombo, but the addition of new species appears to make a larger difference to the AGB of these plots than in Mopane woodlands. We suggest that this is due mostly to the particular identity of species found in ex-Acacia woodlands and their contribution to ecosystem functioning. Unlike Mopane woodlands, ex-Acacia woodlands contain a wider variety of species which can grow to large canopy trees, al-

beit at low densities, especially in transition zones with miombo woodlands. Additionally, many more species species in ex-Acacia woodlands are found in the Mimosoideae and Papilionoideae sub-families, of which most are nitrogen-fixing (Tedersoo et al., 2018). Nitrogen availability is often a limiting factor in productivity, making nitrogen-fixing species strong competitors. It is possible that in ex-Acacia dominated woodlands, the presence of a large number of nitrogen-fixing tree species reduces functional redundancy, meaning that the effect of adding species on ecosystem function saturates at a higher species richness.

5 Conclusions

In this study we found that even in highly disturbed southern African woodlands, there exists a generalisable positive association between tree species diversity and ecosystem function, quantified as above-ground woody biomass (AGB). Our findings contribute to our understanding of a universal biodiversity-ecosystem function relationship, one which is moderated in a predictable manner by environmental covariates and their interaction with biodiversity and ecosystem structure. We found that the multiple vegetation types which comprise southern African woodlands exhibit similarities in the relationship between species diversity and woody biomass, suggesting that similar processes operate across the region to determine ecosystem function. We advocate for explicit inclusion of environmental covariates in regional scale models of biodiversity and ecosystem function. We assert that this is necessary to develop our understanding of the biodiversity-ecosystem function relationship in real-world ecosystems, to progress from experimental mesocosms. We found that much of the effect of species diversity on biomass exists as an indirect effect by increasing the structural diversity of trees, exemplifying a key vector through which tree species diversity determines ecosystem function in savannas, woodlands and forests, i.e. all wooded ecosystems. The presence of a stem density threshold above which the effect of tree species diversity on AGB increases clearly implies the presence of niche complementarity effects in southern African woodlands, an aspect which has often been overlooked in previous studies despite its intuitive logic as a determinant of niche complementarity effects in wooded ecosystems. Our study shows that biodiversity change through extensive human actions in this region will have the greatest negative impact on ecosystem function in areas of high stem density, and low species diversity, which are those areas predominantly targeted for tree felling. This raises concerns about the robustness of these ecosystems to further resource extraction and biodiversity loss.

Table 1: Description of the biogeographical clusters to which each plot in the study was assigned. Indicator species were generated using Dufrene-Legendre indicator species analysis (Dufrene & Legendre, 1997) implemented with `indval()` from the `labdsv` R package (Roberts, 2019) and represent species which define the given cluster. Dominant species were identified by choosing the species with the largest mean plot level proportional AGB within each cluster. Numeric values of species richness, stems ha⁻¹ and AGB represent medians and interquartile ranges (75th percentile - 25th percentile).

Cluster	Dominant species	Indicator species	n	Species Richness	Stem density (stems ha ⁻¹)	AGB (t ha ⁻¹)
Core miombo	<i>Brachystegia spiciformis</i>	<i>Parinari curatellifolia</i>	523	20(16.9)	204(142.5)	44.2(36.11)
	<i>Julbernardia paniculata</i>	<i>Uapaca kirkiana</i>				
	<i>Brachystegia boehmii</i>	<i>Brachystegia spiciformis</i>				
ex-Acacia	<i>Spirostachys africana</i>	<i>Euclea racemosa</i>	188	12(10.3)	181(166.5)	54.5(61.33)
	<i>Senegalia burkei</i>	<i>Vachellia nilotica</i>				
	<i>Senegalia nigrescens</i>	<i>Spirostachys africana</i>				
Mopane	<i>Colophospermum mopane</i>	<i>Colophospermum mopane</i>	58	10(10.2)	186(125.6)	42.7(32.83)
	<i>Androstachys johnsonii</i>	<i>Psuedolachnostylis maprouneifolia</i>				
	<i>Kirkia acuminata</i>	<i>Lannea discolor</i>				
Sparse miombo / Baikiaea	<i>Baikiaea plurijuga</i>	<i>Burkea africana</i>	466	12(13.7)	178(129.5)	36.9(26.98)
	<i>Burkea africana</i>	<i>Baikiaea plurijuga</i>				
	<i>Pterocarpus angolensis</i>	<i>Pterocarpus angolensis</i>				

Table 2: Model fit statistics for SEMs investigating the effects of tree diversity and stem density on AGB (Figure 4). n = number of plots in cluster, χ^2 = Chi-squared fit statistic, DoF = model degrees of freedom, CFI = Comparative Fit Index, TLI = Tucker-Lewis Index, RMSEA = Root Mean Square Error of Approximation, R^2 AGB = R-squared of AGB (Above-Ground Biomass).

Cluster	n	χ^2	DoF	CFI	TLI	RMSEA	R^2 AGB
Core miombo	523	78.670	6	0.904	0.759	0.140	0.490
ex-Acacia	188	9.570	6	0.952	0.879	0.130	0.830
Mopane	58	19.880	6	0.834	0.584	0.240	0.510
Sparse miombo / Baikiaea	466	43.870	6	0.914	0.784	0.130	0.580
All	1235	91.380	6	0.937	0.843	0.120	0.490

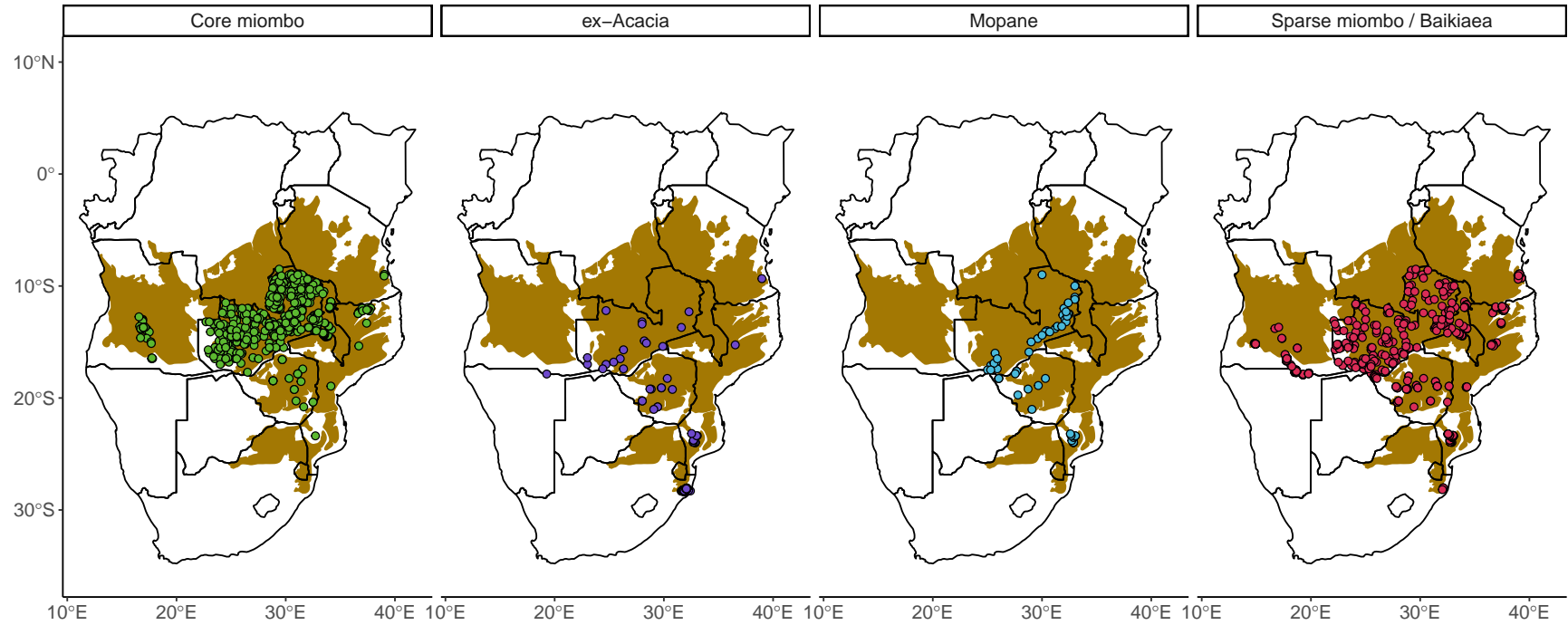


Figure 1: The locations of the 1235 plots used in this study, with respect to the distribution of miombo woodland vegetation according to Hopkins and White (1987). Each panel shows plots categorized by their vegetation type as defined by the vegetation types in Table 1.

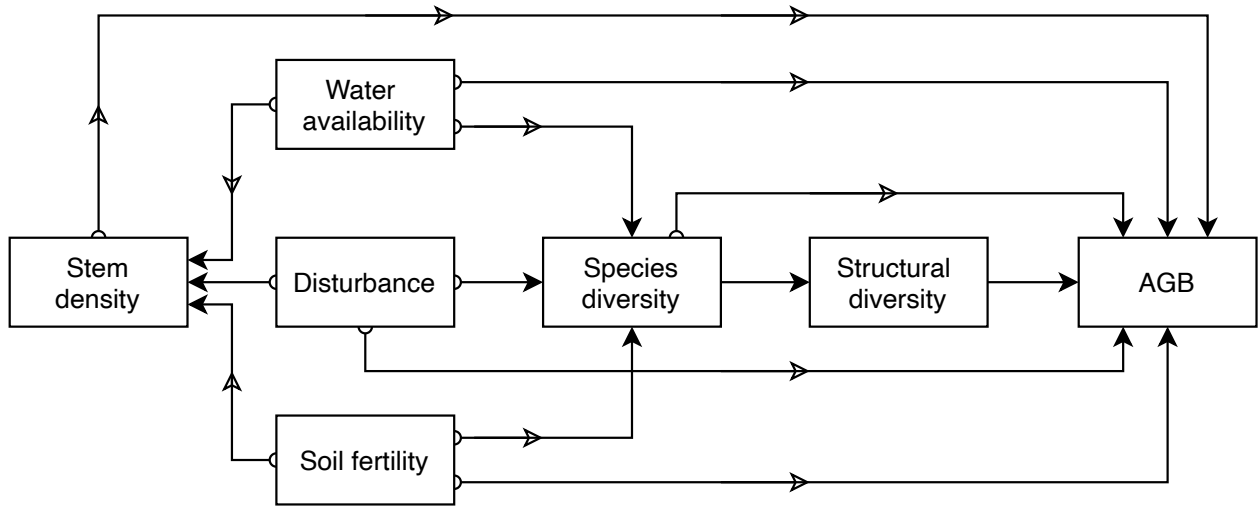


Figure 2: Conceptual Directed Acyclic Graph (DAG) showing the theoretical relationships between environmental factors, tree species diversity, tree structural diversity, stem density, and AGB. Hypothesised paths of causation are depicted as arrows from predictor to response.

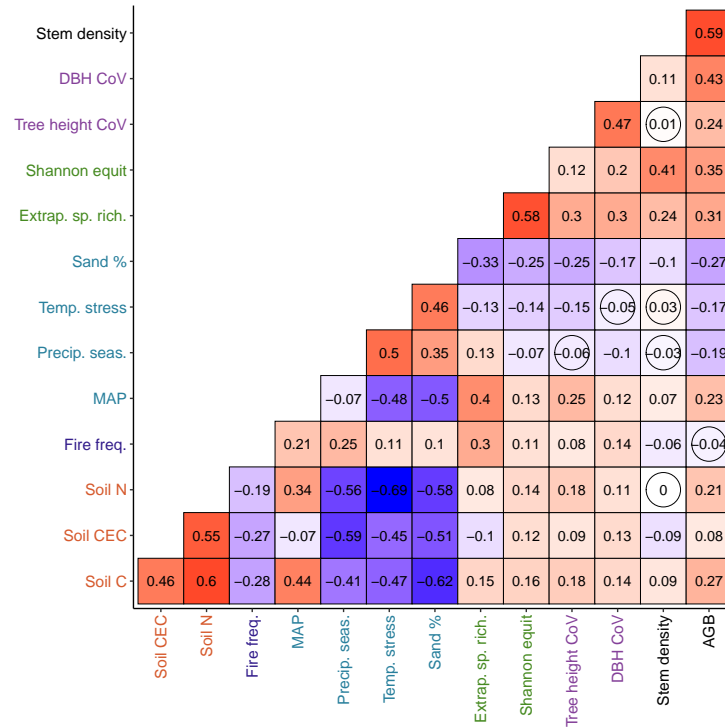


Figure 3: Correlation matrix of standardised observed variables used in the SEMs, with Pearson correlation coefficients (r) coloured according to sign (+ve red, -ve blue) and shaded by strength of correlation. Correlation coefficients marked by a circle indicate that the 95% confidence interval of r overlapped zero. Colours of variable names group them into latent variables used in the SEMs: red = soil fertility, blue = disturbance, turquoise = water availability, green = tree species diversity, purple = tree structural diversity. See Table S1 for a full assessment of correlation fit statistics.

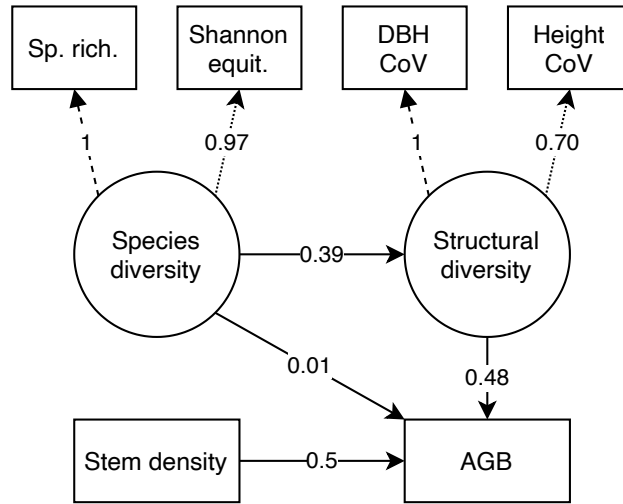


Figure 4: Path diagram with regression coefficients for the tree diversity SEM, including plots from all vegetation clusters. Latent variables are shown as circles while observed variables are shown as rectangles. Standardised path coefficients are shown as solid arrows pointing from predictor to response with the effect size of the path coefficient expressed in terms of standard deviations on the latent variable response scale. The observed variables that inform the latent variables are connected by dotted arrows, and observed variables with loadings set to one are connected by dashed arrows. Measurement errors of exogenous variables are omitted for clarity.

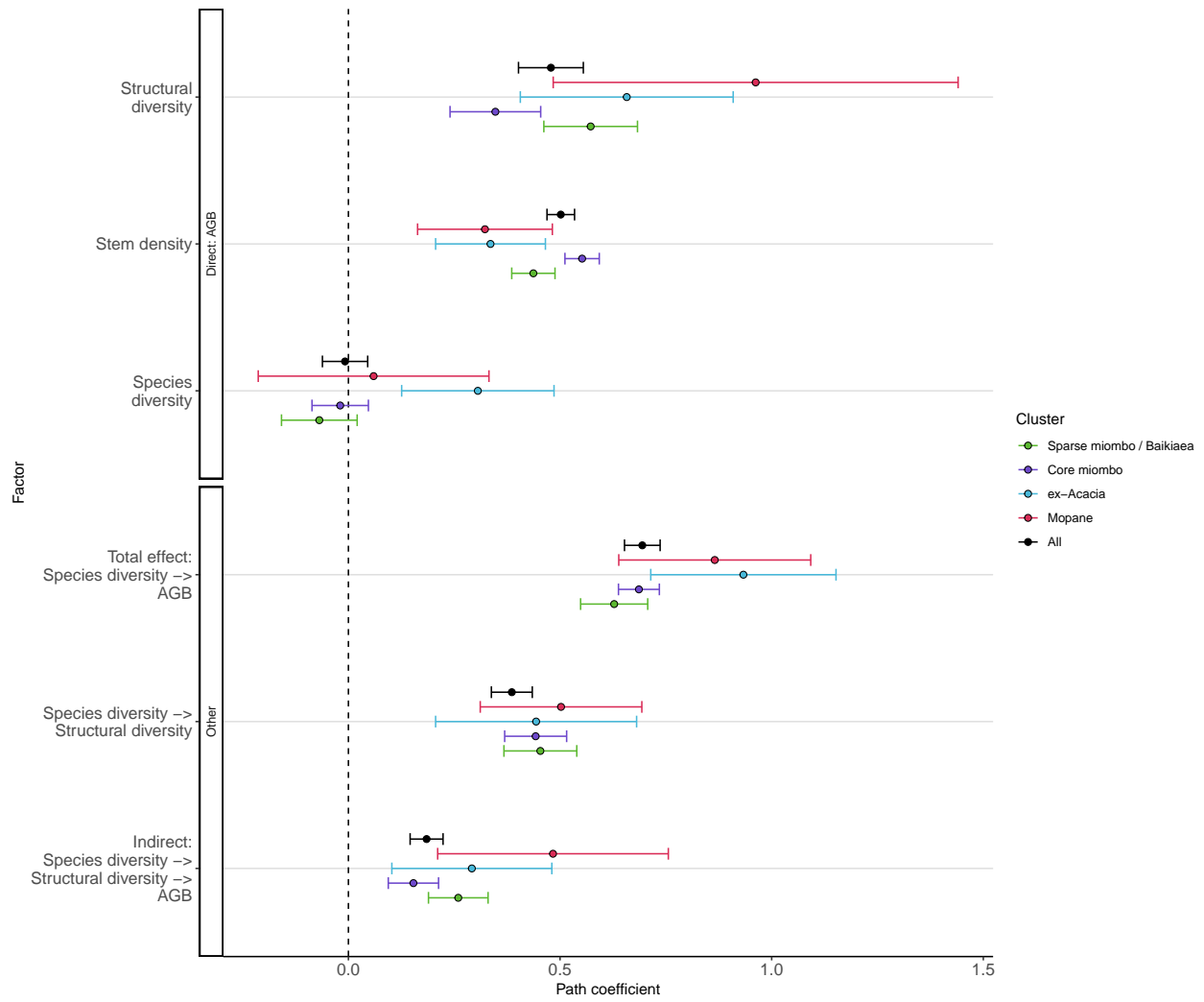


Figure 5: Unstandardised path coefficients for the effects of tree diversity on AGB, mediated by the effect of stand structural diversity. Path coefficients are ± 1 standard error. Path coefficients where the interval (standard error) does not overlap zero are considered to be significant effects.

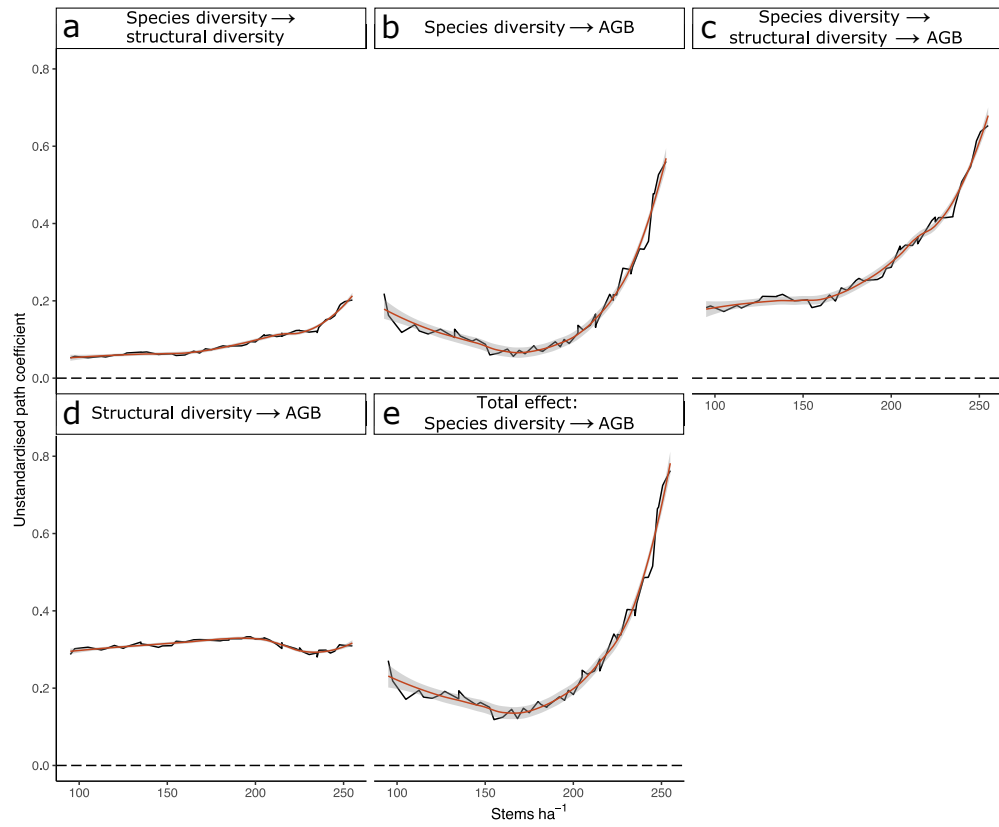


Figure 6: Line plots showing the variation in SEM path coefficients across datasets with different mean stem density. Smoothed lines are loess curves with ± 1 standard error shaded bars. AGB = Above-Ground woody Biomass, arrows in plot titles indicate causal paths in SEM models.

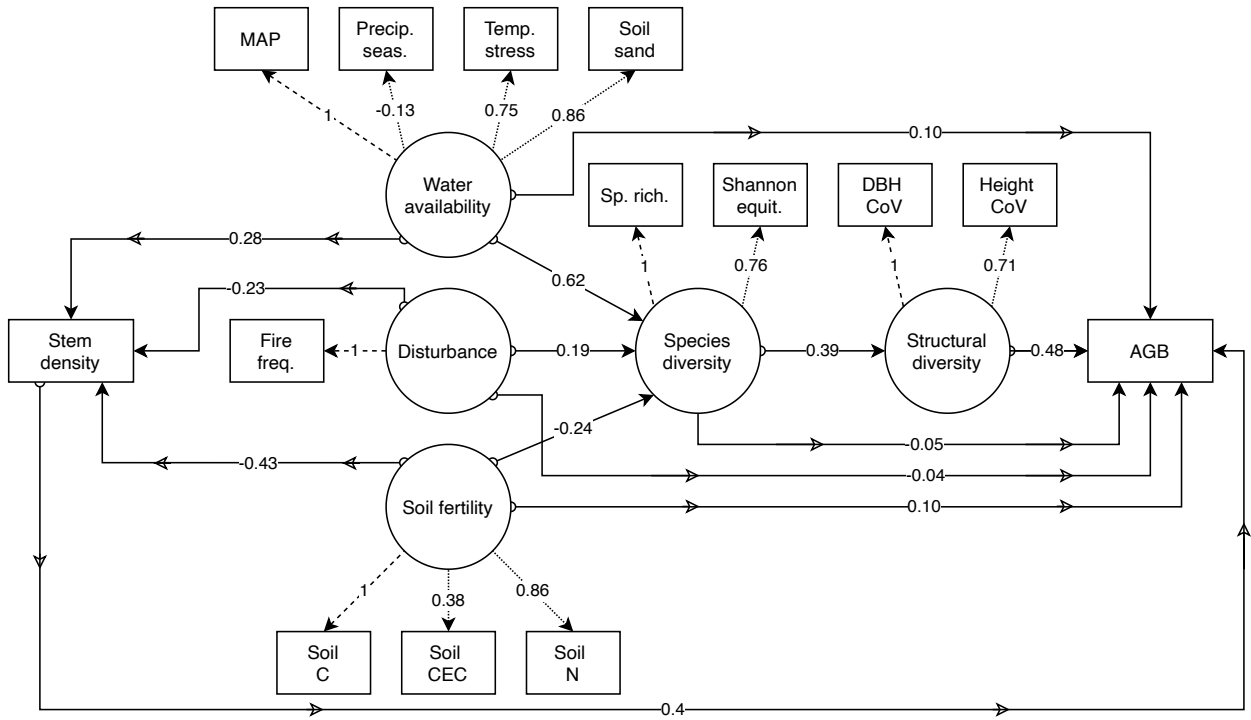


Figure 7: Path diagram with regression coefficients for the SEM incorporating environmental covariates and tree species and structural diversity across all five vegetation types. Latent variables are shown as circles while observed variables are shown as rectangles. Standardised path coefficients are shown as solid arrows pointing from predictor to response, with the effect size of the path coefficient expressed in terms of standard deviations on the latent variable response scale. Observed variables that inform the latent variables are connected by dotted arrows, observed variables with loading set to one are connected by dashed arrows. Measurement errors of exogenous variables are omitted for clarity.

8 Acknowledgements

This work is funded by a NERC E3 Doctoral Training Partnership PhD studentship at the University of Edinburgh (John L. Godlee, Grant No. NE/L002558/1). The data for this study was contributed by a number of independently funded projects and was assembled and prepared by SEOSAW (A Socio-Ecological Observatory for Southern African Woodlands, <https://seosaw.github.io>), an activity of the Miombo Network and a NERC-funded project (Grant No. NE/P008755/1). Revisions of the SEOSAW dataset were funded by SavannaChange, a GCRF/University of Edinburgh funded project. We thank all data providers and the field assistance they received when collecting plot data. JMBC was supported by the Natural Environment Research Council (Agreement PR140015 between NERC and the National Centre for Earth Observation).

9 Author contribution

JG and KD conceived the study. JG conducted data analysis, data management for further versions of the SEOSAW dataset, and wrote the manuscript. CR conceived the SEOSAW database and conducted data management for earlier versions of the SEOSAW dataset. JG, CR, DB, JMBC, MF, RH, EM, SS, HT, HT, MB, MW, and KD contributed to manuscript revisions. JG, CR, SB, VC, JPGMC, DD, MF, FG, SM, IM, AM, RR, NR, AS, SS, JT, JW, MB, and MW contributed to experimental design, field data collection, data preparation and data management of parts of the dataset used in this study.

10 Data Availability

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Supporting Information

Additional supporting information may be found in the online version of this article.

Fig. S1 Histograms of raw untransformed observed variables used in final analyses.

Fig. S2 Histograms of observed variables transformed to achieve a normal frequency distribution.

Fig. S3 Bivariate scatter plots for each observed variable used in SEMs, based on hypothesised paths of causality.

Fig. S4 Unstandardised path coefficients for full SEM model.

Table S1 Table of correlation fit statistics for each pairwise Pearson correlation test of observed variables used in Structural Equation Models.

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