

# A Novel Approach to Modeling Variable Annuity Policyholder Surrender Using Joint Modeling Techniques

Pstat 296-Graduate Research in Actuarial Science

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# Project Background

- Industry sponsor/data provider: [REDACTED]
- Event of interest: Policyholder **surrender** of Variable Annuity contract
  - **Surrender:** Policyholder withdrawals all money out of account and terminates the contract.
- Fit an Averaged Logistic Regression Model
- Fit a **Joint Model**
  - Fit a linear mixed effects model to a longitudinal, endogenous covariate
  - Fit a Cox Proportional Hazards Survival Model

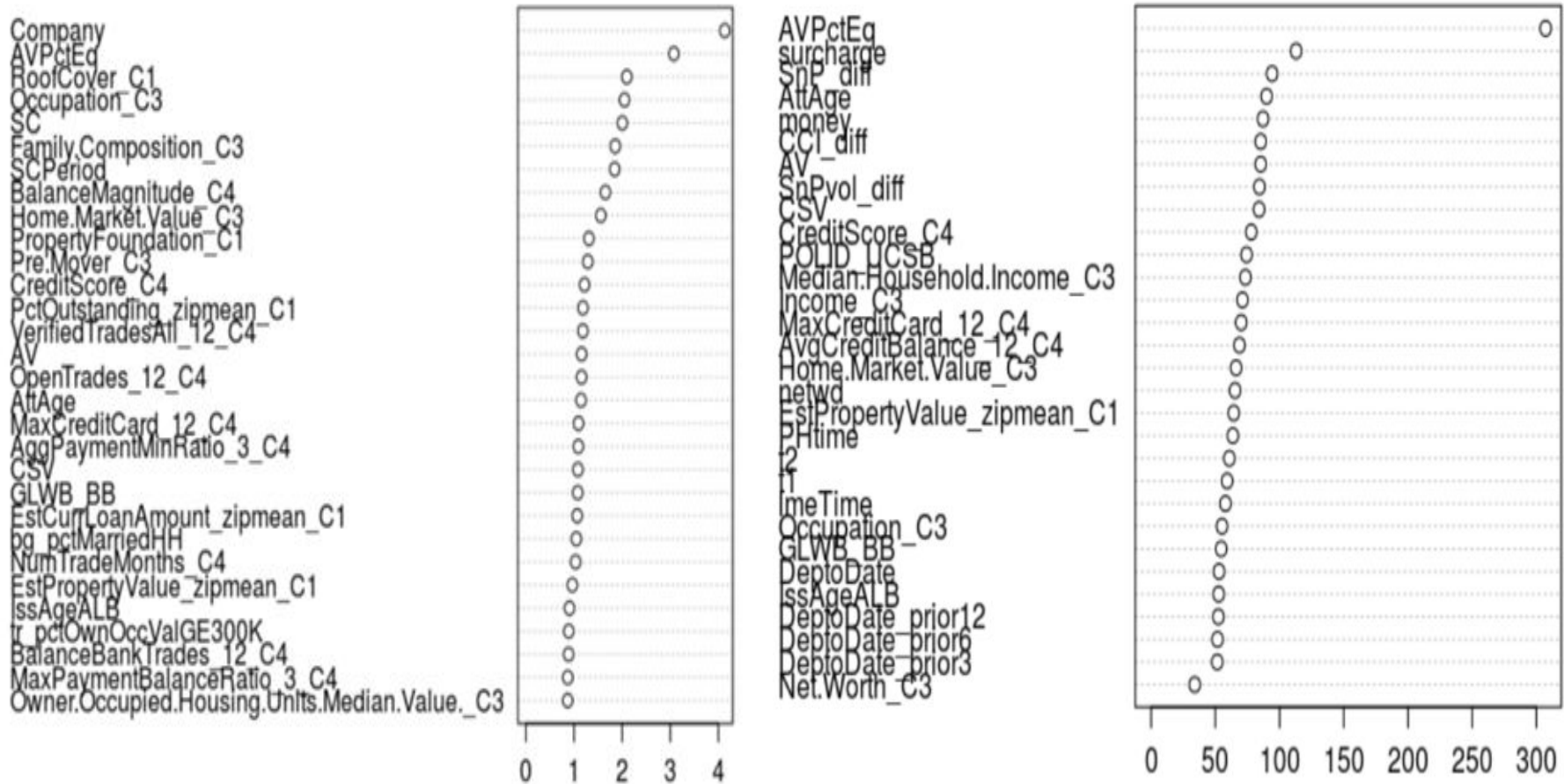
# Data Description

- Data obtained through the good graces of [REDACTED] consist of:
  - VA contract data from 2 companies
  - U.S Bureau census data
  - Consumer data (credit card, property, etc.)
- 4,732,698 rows consisting of quarterly observations on 518,120 policyholders
- 11 valuation quarters (June 2012 - December 2014)

# Data Preprocessing

- Remove extra observation after the policyholder surrendered
  - Left with 4,512,705 rows (95% of the original size)
- Add in new data from market
  - Historical closing prices for S&P 500 - we normalize the differences
  - $\text{S\&P500 at Valuation Date} - \text{S\&P 500 at Issue Date}$
- Created new variables for Joint Model
- Variable selection through random Forest and Variable Importance Plot.
- Missing values filled with Median or Mode

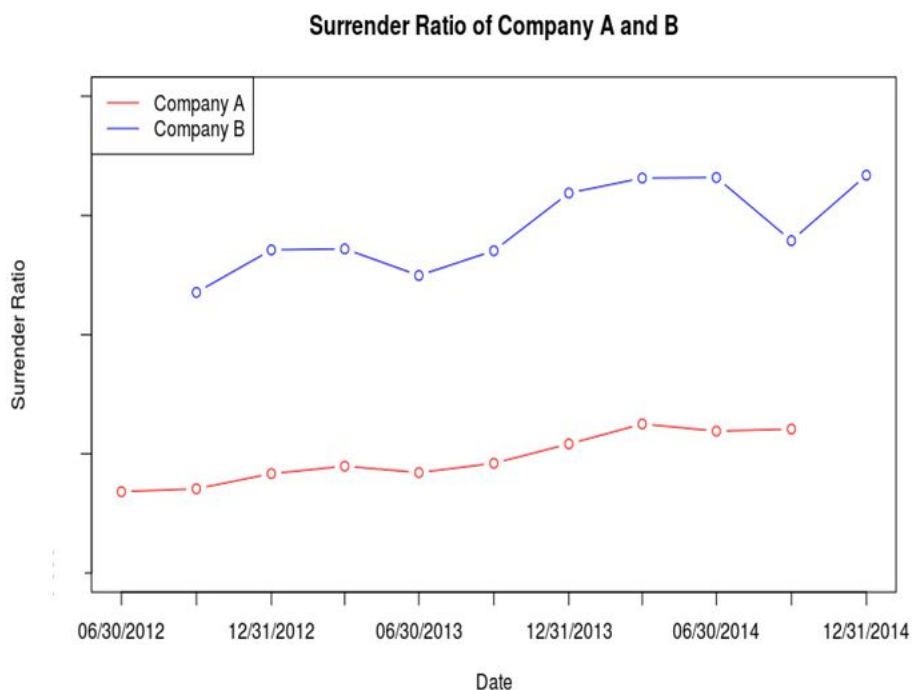
# Variable Importance Plot



# Important Terminology/Covariates

Variables	
AV	Account Value: dollar amount in the variable annuity account at observation time
GLWB_BB	Guaranteed Lifetime Withdrawal Benefit (Benefit Base), the amount of guaranteed return on the variable annuity product
Moneyiness	$GLWB\_BB/AV$ Moneyiness $> 1 \rightarrow$ “In the money” (less incentive to surrender) Moneyiness $< 1 \rightarrow$ “Out of the money” (higher incentive to surrender) $0 < \text{Moneyiness} < 2$
Phtime	The number of quarters a Policyholder has been in the contract at observation time
SCI	Surrender Charge Indicator In – The policyholder is in Surrender Charge Period Shock – The policyholder is at the end of Surrender Charge Period Out – The policyholder is out of the Surrender Charge Period

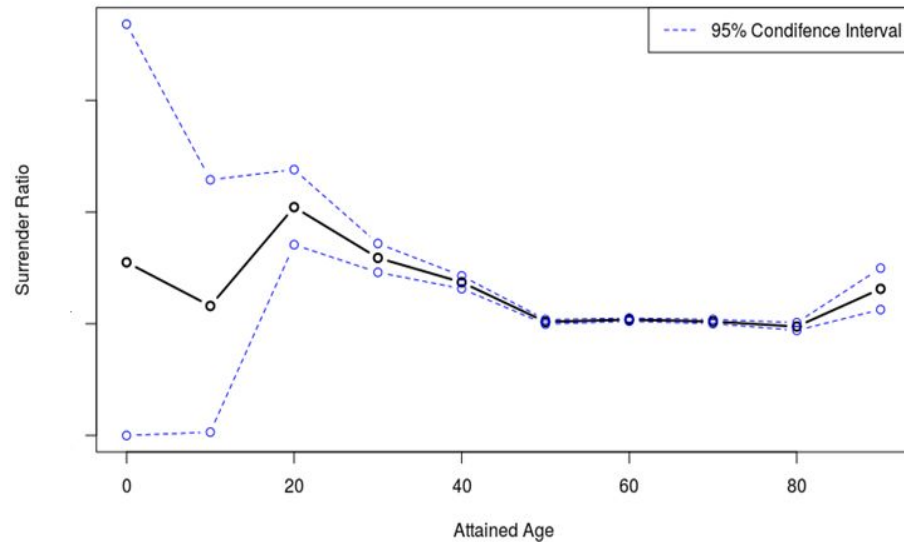
# Observation: Company



- Glm Model: Separate Glm for each company.
  - Different distributions in two companies.
  - Different observation time periods.
- Joint Model: Only modeling people from Company B.
  - Higher Surrender frequency.
  - Provides data on how much of each policyholder's account is in equity

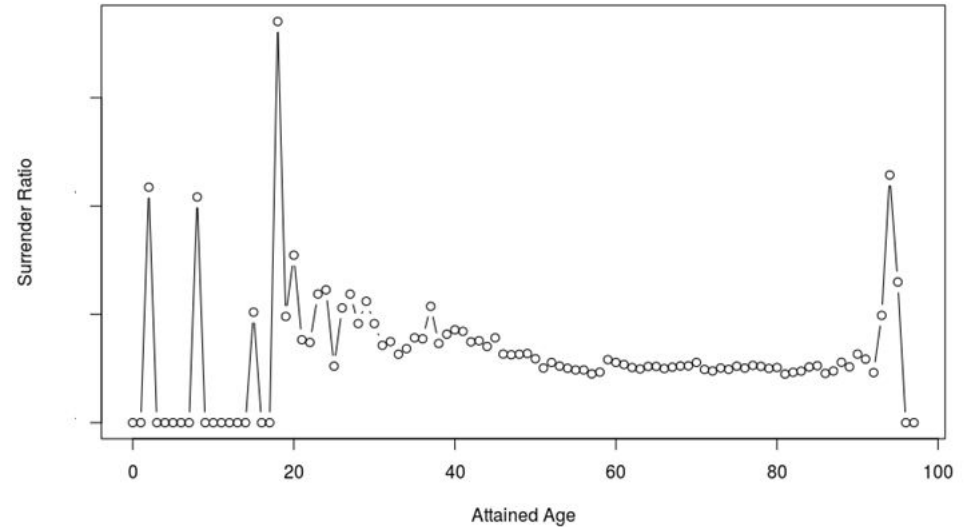
# Observation: Age

Estimated Surrender Ratio vs. Age



Time interval = 10 years

Estimated Surrender Ratio vs. Age



Time interval = 1 year

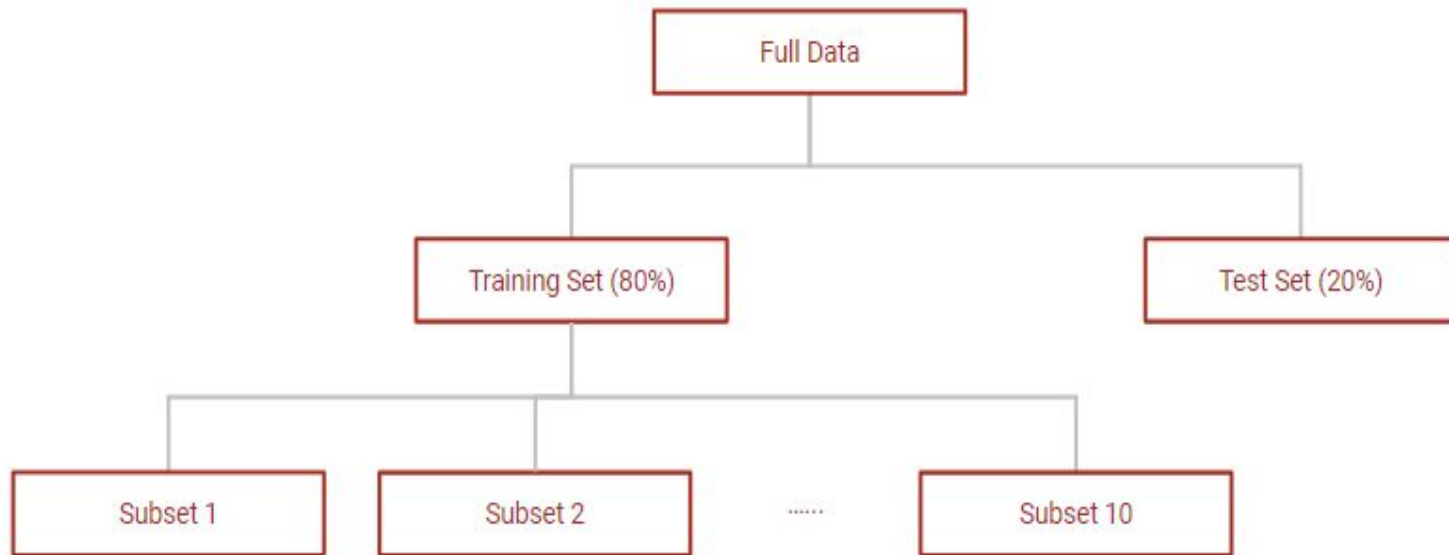


## Observation: Age Cont.

- Distributions that occur before the IRA owner reaches the age of 59½ are subject to a 10% early-distribution penalty, in addition to any income tax owed.
- Required Minimum Distributions (RMDs) generally are minimum amounts that a retirement plan account owner must withdraw annually starting with the year that he or she reaches 70½ years of age or, if later, the year in which he or she retires.

	Z Value
Age 59 - IRA	5.39915
Age 59 - No IRA	2.92530
Age 70 - IRA	-3.70759
Age 70 - No IRA	-1.16711

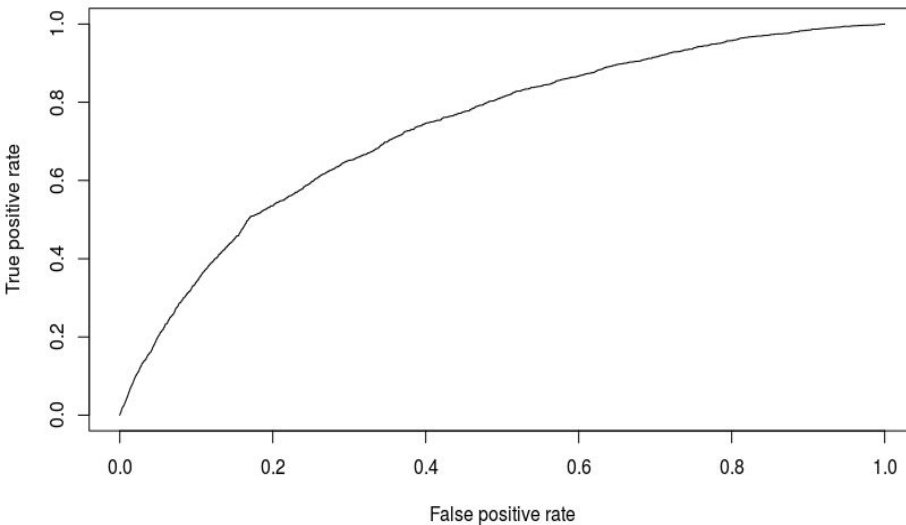
# Averaged Logistic Regression Model Steps



Use the model with averaged coefficients in glm's from 10 subsets to predict the response in the test set

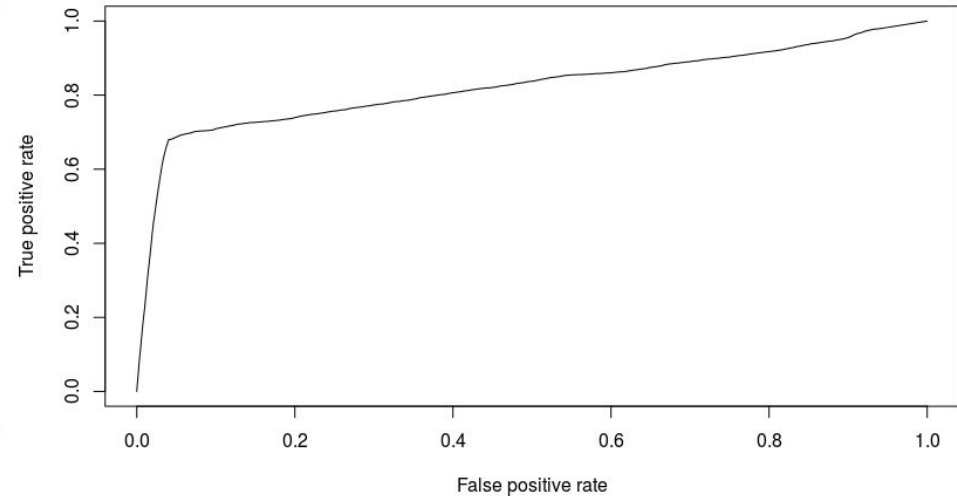
# Model Performance on the Test Set

ROC curve Admissions for Company A



**AUC = 0.735**

ROC curve Admissions for Company B



**AUC = 0.8221**

# Averaged Logistic Regression Model Results

Variable Name	Relationship to Surrender
Maximum Credit Card Owned	Positive
Credit Score	Negative
Surrender Charge Indicator - Out	Positive
Snp 500 Normalized Difference	Negative
Moneyness	Negative

# A Crash Course on Joint Modeling

## Cox Proportional Hazard Model

- The most popular approach to modeling a survival process (such as time-to-surrender) is to use a Cox Proportional Hazards model:

$$\lambda(t|X_i) = \lambda_0(t) \exp(\beta_1 X_{i1} + \cdots + \beta_p X_{ip}) = \lambda_0(t) \exp(X_i \cdot \beta).$$

- The crux of the matter is that the model relies on the Proportional Hazards assumption, that the effects of any covariates included in the model are constant over time (time-independent), ie:

$$\frac{\lambda(t|X_i)}{\lambda(t|X_j)} = c \in \mathbb{R}, \forall t$$

# Why Not a Simple Cox Model?

- The Cox PH model can be adjusted to accommodate time-varying effects, but we run into another problem that invalidates this approach
- The Cox approach assumes that all covariates are **exogenous** (external), meaning that they are deterministic and tell the “full story”
  - Examples: age, sex, occupation, number of beers drunk
- However, in most cases, variables of interest can be classified as **endogenous** (internal), meaning that they are prone to measurement error
  - Examples: T cell count, blood alcohol levels, moneyness
- One can model an endogenous marker's evolution over time using a linear mixed effects model, which fits a model with random and fixed effects

# A Crash Course on Joint Modeling Linear Mixed Effect Model

- The mixed model approach remedies our problem of measurement error by postulating that the observed longitudinal outcome  $y_i(t)$  equals the “true” level plus a random term

$$\begin{aligned} y_i(t) &= m_i(t) + \epsilon_i(t) \\ m_i(t) &= x_i^T \cdot \beta_i + z_i^T \cdot b_i \\ b_i &\sim \mathcal{N}(0, D), \epsilon_i \sim \mathcal{N}(0, \sigma^2) \end{aligned}$$

- $\mathbf{X}_i$  = fixed effects vector
- $\mathbf{Z}_i$  = random effects vector
- $\mathbf{B}_i$  = fixed effects coefficients
- $\mathbf{b}_i$  = random effects coefficients
- $\mathbf{y}_i(\mathbf{t})$  is the observed level of the longitudinal outcome at time  $t$  for subject  $i$
- $\mathbf{m}_i(\mathbf{t})$  is the fitted value of the longitudinal outcome at time  $t$  for subject  $i$

# Arriving at the Joint Model

- The mixed model approach remedies our problem of measurement error by postulating that the observed longitudinal outcome  $y_i(t)$  equals the “true” level  $m_i(t)$  plus a random term representing measurement error
- We then use our estimates for the “true” value of  $m_i(t)$  and impute it into the Cox model, yielding a joint model:

$$\lambda_i(t) = \lambda_0(t) \exp(x_i^T \cdot \beta + \alpha m_i(t))$$



# Important Terminology/Covariates

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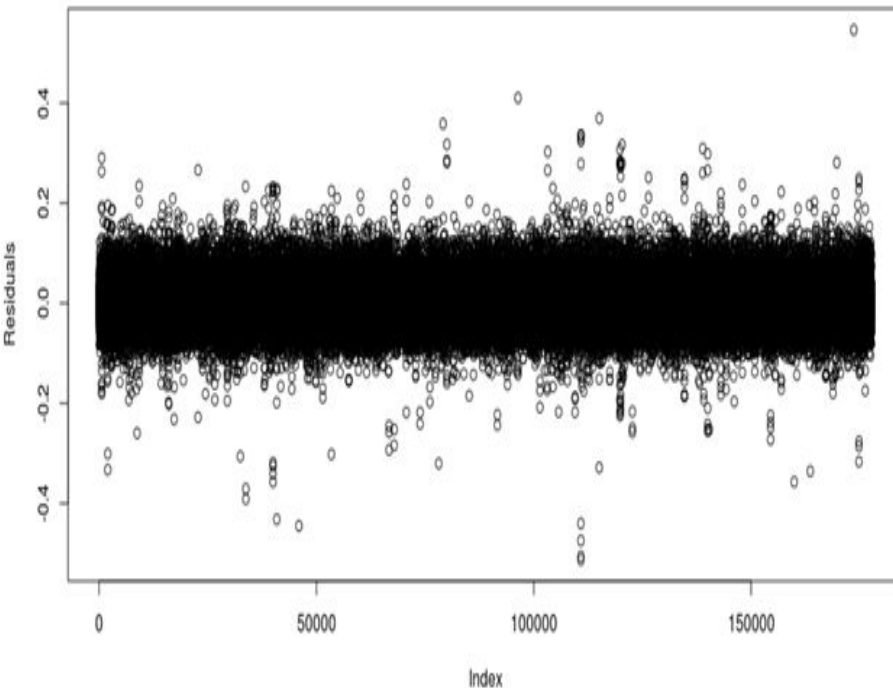
# Our Linear Mixed Effects Model

## Moneyiness as response variable

Variable Name	Effect	P-value
Normalized S&P Difference	Positive	0
Policyholder Time	Positive	0
Number of Withdrawals in past 3 months	Positive	0
Average Percentage of Equity	Positive	0
Interaction between S&P_Diff and Policyholder Time	Positive	0

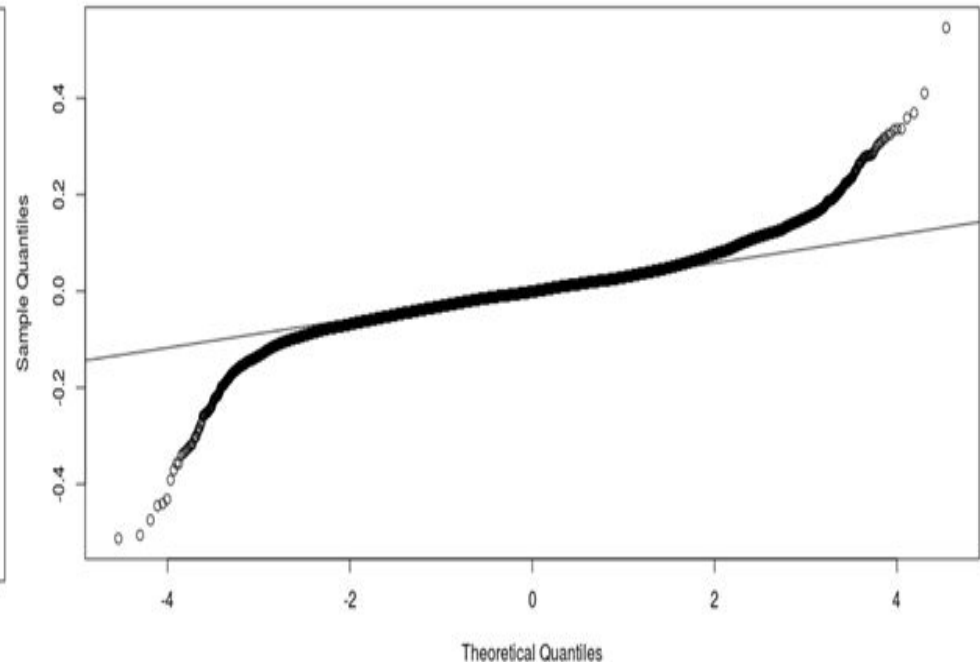
# LME Diagnostic

LME Model Residual Plots



Residual Plot

Normal Q-Q Plot



The Normal Q-Q Plot Showed Heavy Tailed Distribution of Residuals

Normal Q-Q Plot

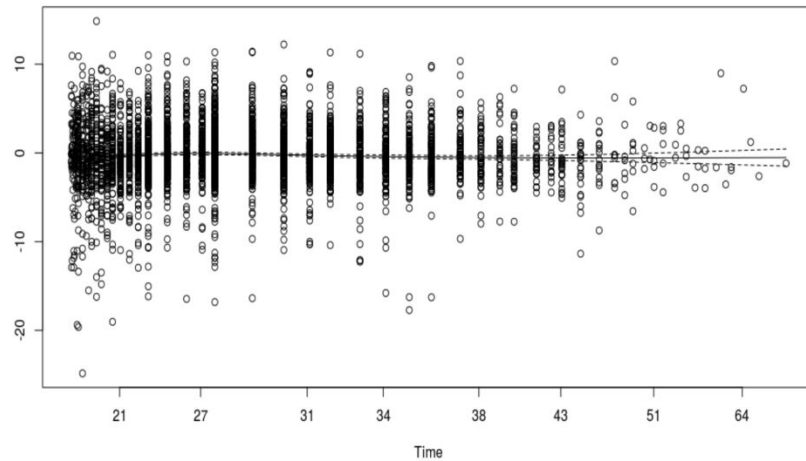
**The Q-Q plot demonstrates heavy-tailed distribution of residuals due to left skewed distribution and capping Moneyness**

# Our Cox PH Model

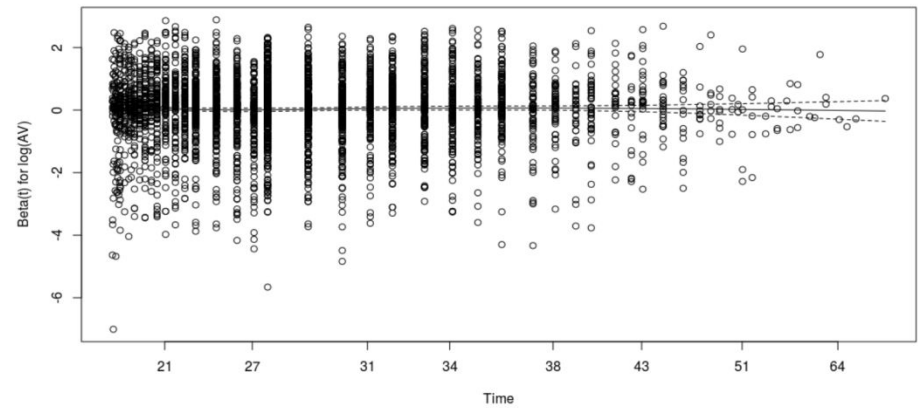
Variable	Effect	P-value
S&P Normalized Difference	Positive	0
Log(Account Value)	Positive	0.0.12
Average Percentage in Equity	Positive	0.35
Surrender Charge Indicator-Out	Negative	0
Surrender Charge Indicator - Shock	Negative	0
Credit Score	Negative	0
Interaction: Log(Account Value) and Average Percentage in Equity	Negative	0

# Cox PH Model Diagnostic

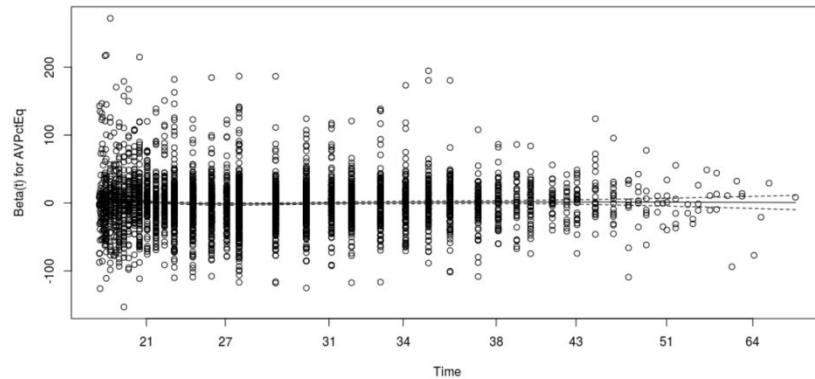
Scaled Schoenfeld Residuals-LME Model



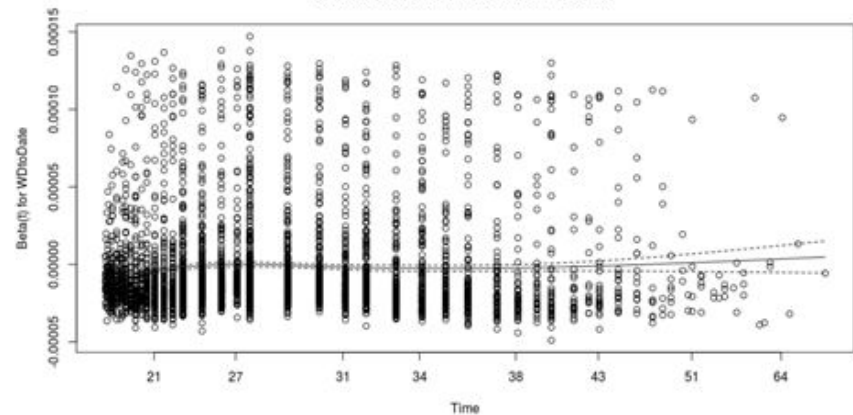
Scaled Schoenfeld Residuals-AccountValue



Scaled Schoenfeld Residuals-AvPctEq



Scaled Schoenfeld Residuals-WDtoDate



# At Long Last... Our Joint Model

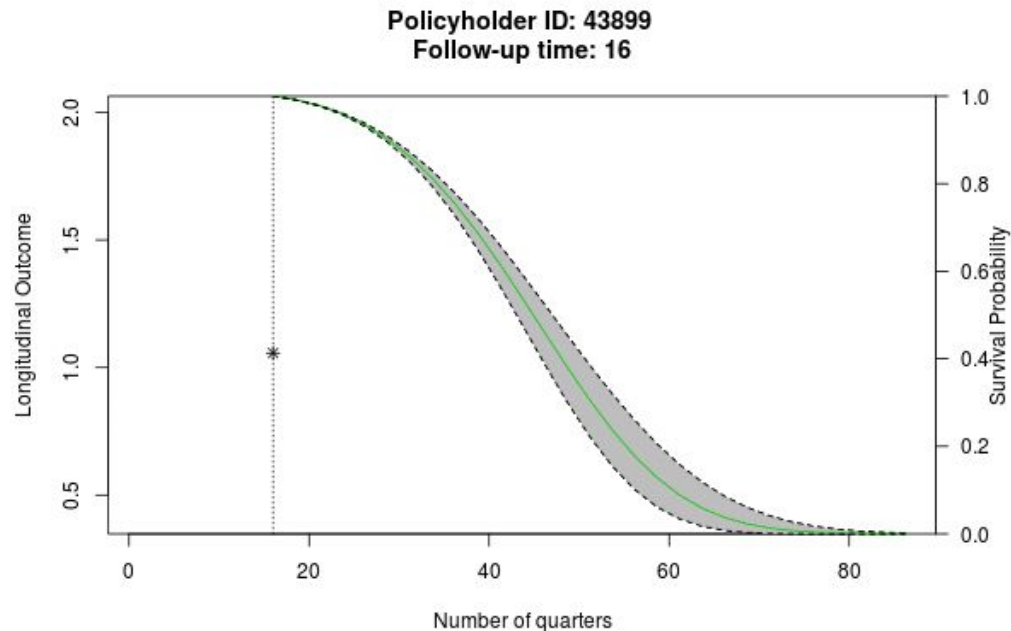
**Negative Association**

**P-value = 0**

Variable Name-Longitudinal	Effect	P	Variable	Effect	P
Normalized S&P Difference	+	0	S&P Normalized Difference	+	0
Policyholder Time	+	0	Log(Account Value)	-	0.3697
Number of Withdrawals in past 3 months	+	0	Average Percentage in Equity	-	0
Average Percentage of Equity	+	0	Surrender Charge Indicator-Out	-	0
Interaction between S&P_Diff and Policyholder Time	+	0	Surrender Charge Indicator - Shock	-	0
			Credit Score	-	0
			Interaction: Log(Account Value) and Average Percentage in Equity	+	0.3667

# Model Prediction

- The R package “JM” has a really cool function “runDynPred()” which allows us to visualize dynamic survival predictions and how the longitudinal response’s fluctuation affects them over time

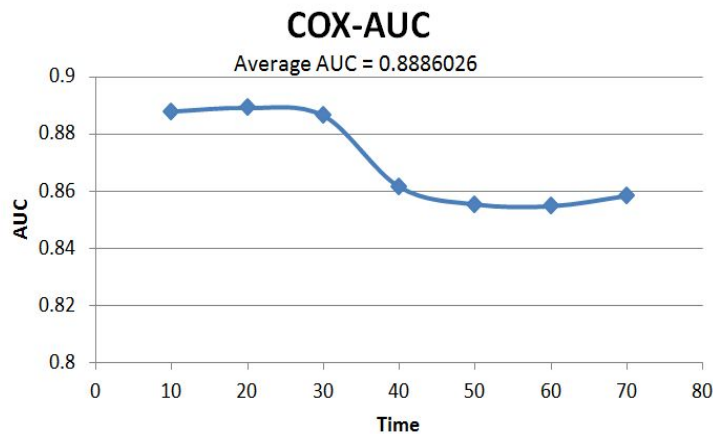


# Model “Comparison” by AUC

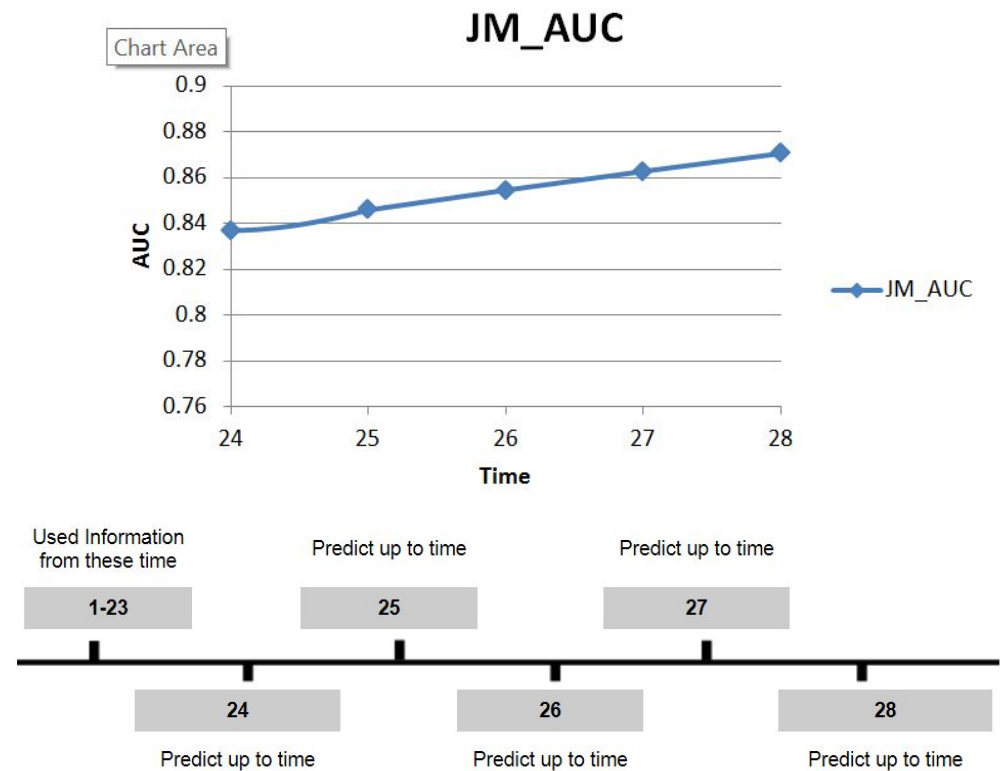
- GLM Model:

Com	AUC
Company A	0.735
Company B	0.8221

- Cox Model



- Joint Model:





# Challenges and Insights

- Joint models are very complex from a computational perspective
- Data need to be well-organized before a joint model can be fitted without issue
- We couldn't get some of the diagnostic functions in library(JM) to work if our model had categorical variables
- Models are imperfect - there are other endogenous covariates related to surrender that can be modeled through fixed and random effects
- library(JM) supports multivariate approaches... not for the squeamish. There's also a Bayesian approach.

# Challenges and Insights

- Joint model
- Difficult to fit

```
> data$money <- (data$GLWB_BB - data$AV) / data$GLWB_BB
```

```
> coxph(Surv(data$t1,data$t2,data$Surr) ~ data$money)
```

```
Error in fitter(X, Y, strats, offset, init, control, weights = weights, :
```

```
routine failed due to numeric overflow.This should never happen. Please contact the author.
```

can be

```
Error in `contrasts<-'(`*tmp`, value = contr.funs[1 + isOF[nn]]) :
```

```
contrasts can be applied only to factors with 2 or more levels
```

```
Error in lme.formula((money) ~ SnP_diff_rnlminb problem, convergence error code message = false convergence (8)
```

Warning message:

```
In jointModel(mod33, mod22, timeVar = "PHtime") :
infinite or missing values in Hessian at convergence.
```

```
Estimation: Monte Carlo ( samples)
```

```
Error in if (x$diffType == "absolute") { : argument is of length zero
```

```
96, 97, 101, 102, 105, 112, 114, 122, 123, 124, 125, 126, 127, 128, 129, 130, 131, 132, 133, 134, 135, 136, 137, 138, 139, 140, 141, 142, 143, 144, 145, 146, 147, 148, 149, 150, 151, 152, 153, 154, 155, 156, 157, 158, 159, 160, 161, 162, 163, 164, 165, 166, 167, 168, 169, 170, 171, 172, 173, 174, 175, 176, 177, 178, 179, 180, 181, 182, 183, 184, 185, 186, 187, 188, 189, 190, 191, 192, 193, 194, 195, 196, 197, 198, 199, 200, 201, 202, 203, 204, 205, 206, 207, 208, 209, 210, 211, 212, 213, 214, 215, 216, 217, 218, 219, 220, 221, 222, 223, 224, 225, 226, 227, 228, 229, 230, 231, 232, 233, 234, 235, 236, 237, 238, 239, 240, 241, 242, 243, 244, 245, 246, 247, 248, 249, 250, 251, 252, 253, 254, 255, 256, 257, 258, 259, 260, 261, 262, 263, 264, 265, 266, 267, 268, 269, 270, 271, 272, 273, 274, 275, 276, 277, 278, 279, 280, 281, 282, 283, 284, 285, 286, 287, 288, 289, 290, 291, 292).
```

# Conclusion

- Joint Modeling is the correct approach for the phenomenon at hand.
- Moneyness has a direct inverse relation with surrenders, although many explanatory variables exist to predict surrenders.
- A multivariate approach is needed to arrive at an industry-quality model.
- This problem deserves more attention than we were able to give it

# Closing Remarks

- In particular, we think that the process of applying JM to this context and building a model from the ground up (rigorously and correctly) deserves to be investigated in a Master's thesis or PhD dissertation
- We maintain that the joint modeling approach is a viable one for the phenomena of interest in this study, and urge you to consider it in your approaches for studying time-to-event and longitudinal data

# Acknowledgments

## **Major props to:**

- Dimitris Rizopolous (author of R package “JM”)
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- Jose Pinhiero (author of R package “nlme”)
- Mr. and Mrs. Duncan
- Prof. Jiyoung Myung
- Shannon Nicponski
- Nhan Huynh

Thank You For Listening!

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Questions/Observations?