

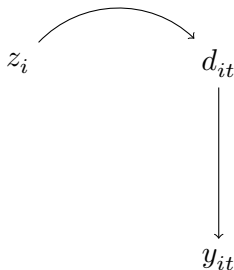
The Problem With Historical Instrumental Variables

Alternative title: Identification of Bi-Directional Two-Variable System With Time-Invariant Instrument

John T.H. Wong

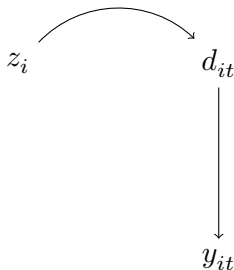
Illustrating the issue with a DAG

- ▶ Many historical IV papers use the following strategy.
 - ▶ e.g., AJR use settlers mortality (z_i) to instrument for constraints on the government's executive (d_{it}), and then estimate the latter's effect on output growth (y_{it}).



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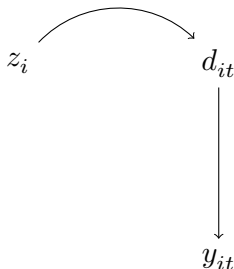
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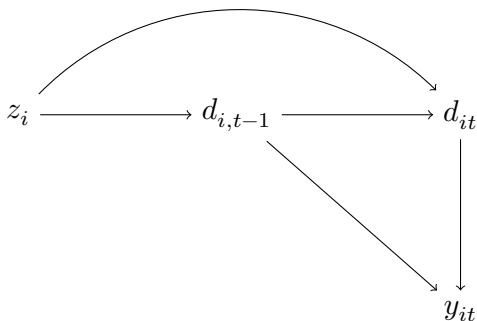
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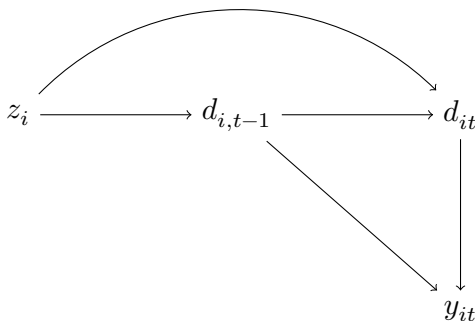
- ▶ z_i must affect y_{it} only through d_{it} .
- ▶ But note that z_i is time-invariant, whereas d_{it} is time-variant.

Illustrating the issue with a DAG



- If z_i affects d_{it} , then it must also affect $d_{i,t-1}$.

Illustrating the issue with a DAG



- ▶ If z_i affects d_{it} , then it must also affect $d_{i,t-1}$.
- ▶ We can add $d_{i,t-2}$, $d_{i,t-3}$, and so forth.

Theoretical model

Second-stage equation

$$y_{it} = \beta_0 d_{it} + \beta_1 d_{i,t-1} + \epsilon_{y,it}.$$

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First-stage equation

$$d_{it} = \delta z_i + \alpha_0 d_{i,t-1} + \epsilon_{d,it}.$$

What happens when we omit d_{t-1} ?

- ▶ Obtain the particular solution of the first-stage equation:

$$d_{it} = (\delta \sum_{j=0}^{\infty} a_1^j) z_i + \underbrace{\sum_{j=0}^{\infty} a_1^j \epsilon_{i,t-j}}_{\text{Not iid!}}.$$

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|-----------|-------|--------------------|
| Intercept | 0.00 | 0.00 (0.01) |
| d_t | 0.30 | -0.10*** (0.02) |
| Ld_t | -0.40 | |
| Num. obs. | | 49950 |

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 1: Two-Stage Least Squares Results With Omitted Treatment Lag

Monte Carlo Results

- These results are consistently biased across samples. (Each unit has 100 observations.)

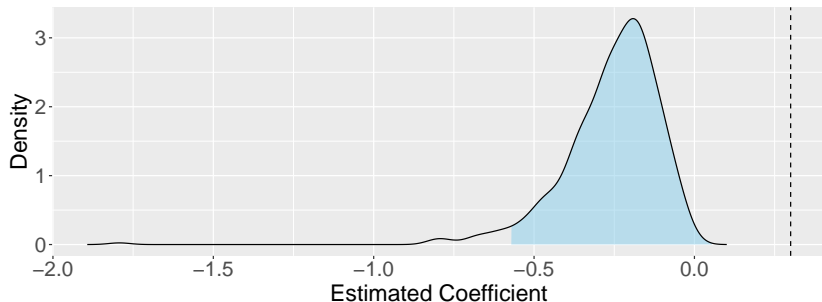


Figure 1: Monte Carlo Results, Omitted Treatment Lag (500 iterations; ± 2 SD shaded)

Solution

- Including Ld_{it} in both stages of the equation leads to a consistent estimator on all variables.

| | True | TOLS | TOLS With Lag |
|-----------|-------|--------------------|--------------------|
| Intercept | 0.00 | 0.00 (0.01) | 0.00 (0.00) |
| d_t | 0.30 | -0.10*** (0.02) | 0.30*** (0.03) |
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Table 2: Two-stage least squares results with treatment lag

Monte Carlo results

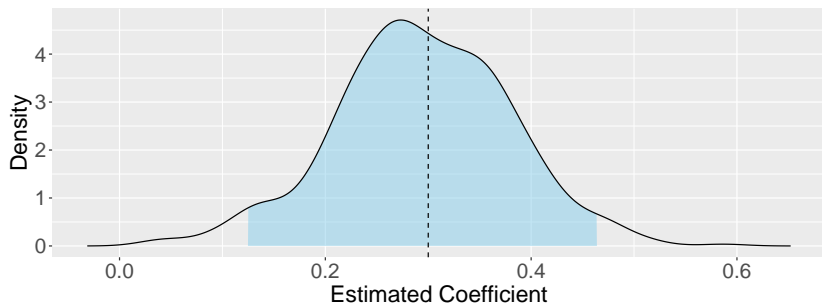
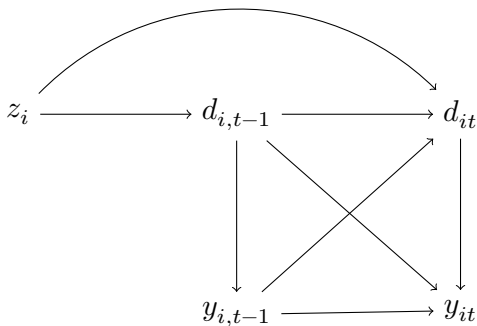


Figure 2: Monte Carlo results, with treatment lag (500 iterations; 50 units; 100 observations per unit; ± 2 SD shaded)

Generalize to bi-directional Granger causation

► What if y_{t-1} feeds into y_{it} and d_{it} ?



For example, Solow model

$$(1) \Delta k = k_t - k_{t-1} = sy_{t-1} - \delta k_{t-1}$$

$$\implies k_t = sy_{t-1} + (1 - \delta)k_{t-1}$$

$$(2) y_t = f(k_t)$$

What we are doing, in VAR terms

$$\begin{aligned}y_{it} &= \beta d_{it} + \alpha_{11}y_{i,t-1} + \alpha_{12}d_{i,t-1} + \epsilon_{y,it} \\d_{it} &= \alpha_{11}y_{i,t-1} + \alpha_{12}d_{i,t-1} + \delta z_i + \epsilon_{it}.\end{aligned}$$

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$$\text{or } \begin{bmatrix} 1 & -\beta \\ \textcolor{red}{0} & 1 \end{bmatrix} \begin{bmatrix} y_{it} \\ d_{it} \end{bmatrix} = \begin{bmatrix} \alpha_{11} & \alpha_{12} \\ \alpha_{21} & \alpha_{22} \end{bmatrix} \begin{bmatrix} y_{i,t-1} \\ d_{i,t-1} \end{bmatrix} + \begin{bmatrix} \textcolor{red}{0} \\ \delta \end{bmatrix} z_i + \begin{bmatrix} \epsilon_{y,it} \\ \epsilon_{d,it} \end{bmatrix}.$$

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What most IV papers are doing, in VAR terms

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Results

| | 1st-Stage (d_t) | | | 2nd-Stage (y_t) | | |
|-----------|---------------------|-------------------|-------------------|---------------------|--------------------|--------------------|
| | True | (a) | (b) | True | (c) | (d) |
| Intercept | 0.00 | 0.02 (0.01) | 0.00 (0.00) | 0.00 | 0.01 (0.01) | 0.01 (0.00) |
| z_t | 0.20 | 0.29*** (0.01) | 0.20*** (0.00) | | | |
| Ld_t | 0.50 | | 0.50*** (0.00) | -0.40 | | -0.40*** (0.01) |
| Ly_t | 0.70 | | 0.70*** (0.00) | 0.60 | | 0.60*** (0.02) |
| d_t | | | | 0.30 | -0.24*** (0.03) | 0.30*** (0.03) |
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Table 3: Two-stage least squares results, comparison

Monte Carlo results

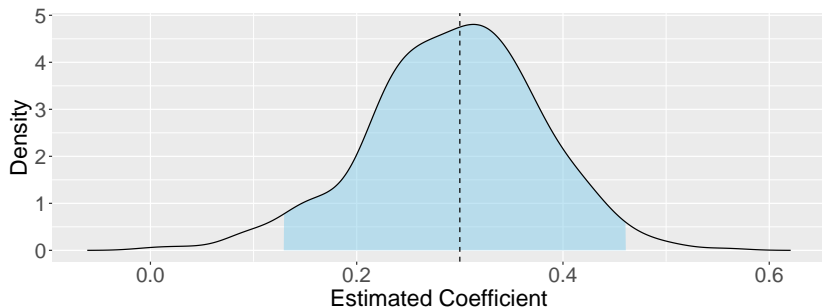


Figure 3: Monte Carlo results, with treatment and outcome lag (500 iterations; 50 units; 100 observations per unit; ± 2 SD shaded)

Note that

1. We did not need a second set of instruments.
2. Our instruments needn't be time-variant.
3. Once d_{it} is instrumented for, the coefficient on d_{t-1} is unbiased.

Results by sample size

(Work in progress...)

Discussion: is anyone talking about this?

- ▶ Most of the papers are in epidemiology (see Labrecque and Swanson 2018 in particular).
 - ▶ This is because they use genetic variants to predict disease's effect (e.g., smoking) on health outcome (e.g., life expectancy), i.e., Mendelian randomization.
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- ▶ There is one development econ paper which talks about this (Casey and Klemp 2021), but their solution is questionable.
 - ▶ They propose to estimate d_{it} and $d_{i,t-Q}$, and instrument for the latter with z_i , and use the resulting parameter to adjust the second-stage equation parameter.

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 - ▶ They propose to estimate d_{it} and $d_{i,t-Q}$, and instrument for the latter with z_i , and use the resulting parameter to adjust the second-stage equation parameter.
 - ▶ The lag length is arbitrary. And even if it works, their method can't recover only a “long-run” parameter, not the instantaneous parameter that is of policy interest.
- ▶ There are actually a lot of time series tools that can help analyze and solve the problem. But the Anderson-Rubin causal inference people don't talk to the time series people or something?

Discussion: how important is this result?

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2. The causal literature has been misspecifying a lot of IV papers.
 - ▶ Part of the causal revolution-revolution: TWFE DID estimator (Callaway & Sant'Anna 2020), issues with geographic IVs (Mellon 2022).