# CSL003P1M: Probability and Statistics Lecture 35 (Joint Distribution Of A Function Of Continuous Random Variables)

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## Joint Density Function of Continuous Random Variables

Suppose that the joint density function of the n random variables  $X_1, X_2, \ldots, X_n$  is given and we want to compute the joint density function of  $Y_1, Y_2, \ldots, Y_n$ , where

$$Y_1 = g_1(X_1, \ldots, X_n), Y_2 = g_2(X_1, \ldots, X_n), \ldots, Y_n = g_n(X_1, \ldots, X_n)$$

#### Assumptions:

- **1** The functions  $g_i$  have continuous partial derivatives.
- ② The equations  $y_1 = g_1(x_1, ..., x_n)$ ,  $y_2 = g_2(x_1, ..., x_n)$ , ...,  $y_n = g_n(x_1, ..., x_n)$  have a unique solution, say,  $x_1 = h_1(y_1, ..., y_n)$ ,  $x_2 = h_2(y_1, ..., y_n)$ , ...,  $x_n = h_n(y_1, ..., y_n)$ .
- **3**  $J(x_1, x_2, \ldots, x_n) \neq 0$ .



### Joint Density Function of Continuous Random Variables

$$J(x_{1},...,x_{n}) = \begin{vmatrix} \frac{\partial g_{1}}{\partial x_{1}} & \frac{\partial g_{1}}{\partial x_{2}} & \dots & \frac{\partial g_{1}}{\partial x_{n}} \\ \frac{\partial g_{2}}{\partial x_{1}} & \frac{\partial g_{2}}{\partial x_{2}} & \dots & \frac{\partial g_{2}}{\partial x_{n}} \\ \vdots & \vdots & \vdots & \vdots \\ \frac{\partial g_{n}}{\partial x_{1}} & \frac{\partial g_{n}}{\partial x_{2}} & \dots & \frac{\partial g_{n}}{\partial x_{n}} \end{vmatrix} \neq 0$$

Under these assumptions, the joint density function of the random variables  $Y_i$  is given by

$$f_{Y_1,...,Y_n}(y_1,...,y_n) = f_{X_1,...,X_n}(x_1,...,x_n)|J(x_1,...,x_n)|^{-1}$$

where  $x_i = h_i(y_1, ..., y_n), i = 1, 2, ..., n$ .



Let  $X_1, X_2$  and  $X_3$  be independent standard normal random variables. If  $Y_1 = X_1 + X_2 + X_3$ ,  $Y_2 = X_1 - X_2$ , and  $Y_3 = X_1 - X_3$ , compute the joint density function of  $Y_1, Y_2, Y_3$ .

Solution: Let  $y_1 = g_1(x_1, x_2, x_3) = x_1 + x_2 + x_3$ ,  $y_2 = g_2(x_1, x_2, x_3) = x_1 - x_2$ ,  $y_3 = g_3(x_1, x_2, x_3) = x_1 - x_3$ , the Jacobian of these transformations is given by

$$J = \begin{vmatrix} 1 & 1 & 1 \\ 1 & -1 & 0 \\ 1 & 0 & -1 \end{vmatrix} = 3$$

Moreover,

$$x_1 = \frac{y_1 + y_2 + y_3}{3}, \ x_2 = \frac{y_1 - 2y_2 + y_3}{3}, \ x_3 = \frac{y_1 + y_2 - 2y_3}{3}$$

So,

$$f_{Y_1,Y_2,Y_3}(y_1,y_2,y_3) = \frac{1}{3} f_{X_1,X_2,X_3} \left( \frac{y_1 + y_2 + y_3}{3}, \frac{y_1 - 2y_2 + y_3}{3}, \frac{y_1 + y_2 - 2y_3}{3} \right)$$

Since,

$$f_{X_1,X_2,X_3}(x_1,x_2,x_3) = \frac{1}{(2\pi)^{3/2}} e^{-\sum_{i=1}^3 x_i^2/2}$$

hence,

$$f_{Y_1,Y_2,Y_3}(y_1,y_2,y_3) = \frac{1}{3(2\pi)^{3/2}}e^{-Q(y_1,y_2,y_3)/2}$$

where

$$Q(y_1, y_2, y_3) = \left(\frac{y_1 + y_2 + y_3}{3}\right)^2 + \left(\frac{y_1 - 2y_2 + y_3}{3}\right)^2 + \left(\frac{y_1 + y_2 - 2y_3}{3}\right)^2$$
  
=  $\frac{1}{3}y_1^2 + \frac{2}{3}y_2^2 - \frac{2}{3}y_2y_3$ 

Let  $X_1, X_2, \ldots, X_n$  be independent and identically distributed exponential random variables with parameter  $\lambda$ . Let

$$Y_i = X_1 + X_2 + \cdots + X_i$$
  $i = 1, \ldots, n$ 

- Find the joint density function of  $Y_1, \ldots, Y_n$ .
- **②** Use the result of part (1) to find the density of  $Y_n$ .

Solution: (1)

$$f_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \prod_{i=1}^n \lambda e^{-\lambda x_i} \quad 0 < x_i < \infty, i = 1, \dots, n$$



The Jacobian is

$$J = \begin{vmatrix} 1 & 0 & 0 & 0 & \cdots & 0 \\ 1 & 1 & 0 & 0 & \cdots & 0 \\ 1 & 1 & 1 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & 1 & 1 & 1 & \cdots & 1 \end{vmatrix} = 1$$

Moreover, we obtain

$$X_1 = Y_1, X_2 = Y_2 - Y_1, \dots, X_i = Y_i - Y_{i-1}, \dots, X_n = Y_n - Y_{n-1}$$

Now, the joint density function of  $Y_i$  is given by

$$f_{Y_1,...,Y_n}(y_1,...,y_n) = f_{X_1,...,X_n}(x_1,...,x_n)|J(x_1,...,x_n)|^{-1}$$



$$f_{Y_{1},...,Y_{n}}(y_{1},...,y_{n}) = f_{X_{1},...,X_{n}}(x_{1},...,x_{n})|J(x_{1},...,x_{n})|^{-1}$$

$$= f_{X_{1},...,X_{n}}(y_{1},y_{2}-y_{1},...,y_{n}-y_{n-1})$$

$$= \lambda^{n} \exp \left\{-\lambda \left[y_{1} + \sum_{i=2}^{n} (y_{i}-y_{i-1})\right]\right\}$$

$$= \lambda^{n} e^{-\lambda y_{n}} \quad 0 < y_{1}, 0 < y_{i} - y_{i-1}, i = 2,...,n$$

$$= \lambda^{n} e^{-\lambda y_{n}} \quad 0 < y_{1} < y_{2} < \cdots < y_{n}$$

Solution (2): Try!!

# Miscellaneous

# Jointly Continuous

### Joint Probability Density Function

We say that X and Y are jointly continuous if there exists a function f(x, y), defined for all real x and y, having the property that, for every set C of pairs of real numbers (that is, C is a set in the two-dimensional plane),

$$P\{(X,Y)\in C\}=\int\int_{(x,y)\in C}f(x,y)dxdy$$

The function f(x, y) is called the **joint probability density** function of X and Y.

If A and B are any sets of real numbers, then, by defining  $C = \{(x, y) : x \in A, y \in B\}$ , we obtain

$$P\{X \in A, Y \in B\} = \int_{B} \int_{A} f(x, y) dx dy$$

### Joint Distribution Function

#### Joint Distribution Function

$$F(a,b) = P\{X \in (-\infty, a], Y \in (-\infty, b]\}$$
$$= \int_{-\infty}^{b} \int_{-\infty}^{b} f(x, y) dx dy$$

## Joint Density Function

#### Joint Density Function

$$f(a,b) = \frac{\partial^2}{\partial a \partial b} F(a,b)$$

where the partial derivatives are defined.

Alternatively,

$$P\{a < X < a + da, b < Y < b + db\} = \int_{b}^{d+db} \int_{a}^{a+da} f(x, y) dx dy$$
$$\approx f(a, b) da db$$

when da and db are small and f(x, y) is continuous at a, b.



### Marginal Density Function

Let X and Y be jointly continuous and furthermore they are individually continuous. Their probability density function can be obtained as follows:

$$P\{X \in A\} = P\{X \in A, Y \in (-\infty, \infty)\}$$
$$= \int_{A} \int_{-\infty}^{\infty} f(x, y) dy dx$$
$$= \int_{A} f_{X}(x) dx$$

where

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

is thus the probability density function of X. Similarly, the probability density function of Y is given by

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$



The joint density function of X and Y is given by

$$f(x,y) = \begin{cases} 2e^{-x}e^{-2y} & 0 < x < \infty, 0 < y < \infty \\ 0 & \text{otherwise} \end{cases}$$

Compute (a)  $P\{X > 1, Y < 1\}$ , (b)  $P\{X < Y\}$  and (c)  $P\{X < a\}$ .

Solution: (a)

$$P\{X > 1, Y < 1\} = \int_0^1 \int_1^\infty 2e^{-x} e^{-2y} dx dy$$
  
= 
$$\int_0^1 2e^{-2y} \left( -e^{-x} |_1^\infty \right) dy$$
  
= 
$$e^{-1} \int_0^1 2e^{-2y} dy = e^{-1} (1 - e^{-2})$$

(b) 
$$P\{X < Y\} = \int \int_{(x,y):x < y} 2e^{-x}e^{-2y}dxdy$$

$$= \int_{0}^{\infty} \int_{0}^{y} 2e^{-x}e^{-2y}dxdy$$

$$= \int_{0}^{\infty} 2e^{-2y}(1 - e^{-y})dy$$

$$= \int_{0}^{\infty} 2e^{-2y}dy - \int_{0}^{\infty} 2e^{-3y}dy$$

$$= 1 - \frac{2}{3} = \frac{1}{3}$$

(c) 
$$P\{X < a\} = \int_{0}^{a} \int_{0}^{\infty} 2e^{-2y} e^{-x} dy dx \\ = \int_{0}^{a} e^{-x} dx = 1 - e^{-a}$$

## Independent Random Variables

#### Independent Random Variables

The random variables X and Y are said to be independent if, for any two sets of real numbers A and B,

$$P{X \in A, Y \in B} = P{X \in A}P{Y \in B}$$

In other words, X and Y are independent if, for all A and B, the events  $E_A = \{X \in A\}$  and  $F_B = \{Y \in B\}$  are independent.

It can be shown that the above equation holds if and only if

$$P\{X \le a, Y \le b\} = P\{X \le a\}P\{Y \le b\}$$



### Independent Random Variables

In terms of the joint distribution function of X and Y, X and Y are independent if

$$F(a,b) = F_X(a)F_Y(b)$$
 for all  $a,b$ 

In the jointly continuous case, the condition of independence is equivalent to

$$f(x,y) = f_X(x)f_Y(y)$$
 for all  $x, y$ 



### Conditional Distributions

If X and Y have a joint probability density function f(x, y), then the conditional probability density function of X given that Y = y is defined, for all values of y such that  $f_Y(y) > 0$ , by

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$$

### Conditional Distributions

• If X and Y are jointly continuous, then, for any set A,

$$P\{X \in A|Y = y\} = \int_A f_{X|Y}(x|y)dx$$

② In particular, by letting  $A=(-\infty,a]$ , we can define the conditional cumulative distribution function of X given that Y=y by

$$F_{X|Y}(a|y) = P\{X \le a|Y = y\} = \int_{-\infty}^{a} f_{X|Y}(x|y)dx$$



The joint density of X and Y is given by

$$f(x,y) = \begin{cases} \frac{12}{5}x(2-x-y) & 0 < x < 1, \ 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$

Compute the conditional density of X given that Y = y, where 0 < y < 1.

Solution: For 0 < x < 1, 0 < y < 1, we have

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$$
  
=  $\frac{f(x,y)}{\int_{-\infty}^{\infty} f(x,y)dx}$ 

$$\frac{f(x,y)}{\int_{-\infty}^{\infty} f(x,y) dx} = \frac{x(2-x-y)}{\int_{0}^{1} x(2-x-y) dx}$$
$$= \frac{x(2-x-y)}{\frac{2}{3} - \frac{y}{2}}$$
$$= \frac{6x(2-x-y)}{4-3y}$$

# Thank You