# A Joint Model for Longitudinal and Event History Data in Life Course Research: Occupational Trajectories and Time to Retirement

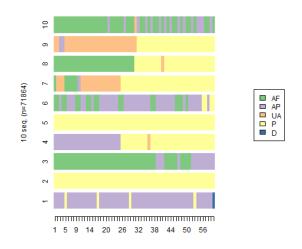
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# Labour market registry (2008 - 2017)

id	month	sex	birthdate	status	hours	start	end
1	January	М	08-03-1950	sick leave	148	01-01-2011	31-01-2011
2	January	F	21-09-1950	employed	10	20-01-2011	31-01-2011
2	Febuary	F	21-09-1950	employed	50	01-02-2011	28-02-2011
3	January	M	02-11-1950	employed	10	01-01-2011	15-01-2011
3	January	M	02-11-1950	employed	10	01-01-2011	15-01-2011
3	January	M	02-11-1950	unemployed	74	16-01-2011	31-01-2011
3	Febuary	M	02-11-1950	unemployed	37	01-02-2011	07-02-2011
3	Febuary	M	02-11-1950	pension	109	08-02-2011	28-02-2011
4	March	F	07-04-1950	pension	148	01-03-2011	31-03-2011
5	January	M	08-12-1950	employed	148	01-01-2011	31-01-2011

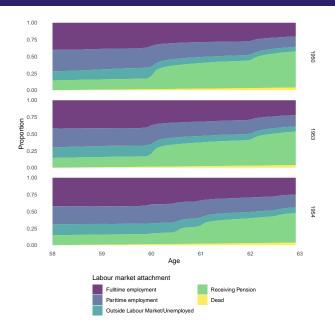
# Occupational Trajectories for 10 individuals



# Early Retirement Pension

Date of Birth	ERP age
< 1954	60
$\geq$ 01-01-1954	60.5
$\geq$ 01-07-1954	61
$\geq$ 01-01-1955	61.5
$\geq 01$ -07-1955	62
$\geq$ 01-01-1956	62.5
$\geq 01$ -07-1956	63
$\geq$ 01-01-1959	63.5
$\geq 01$ -07-1959	64
> 1963	computed in relation to life expectancy

# Occupational Trajectories for 3 Cohorts



## Joint Models

Individuals  $i=1,\ldots,N$ , Timepoints  $t=1,\ldots,n_i$ 

Longitudinal submodel:

$$y_{it} = \boldsymbol{x}_{it}^{(1)t} \boldsymbol{\beta}^{(1)} + \boldsymbol{z}_{it}^t \boldsymbol{b}_i + \epsilon_{ij}$$

Survival submodel:

$$\lambda_i(t) = \lambda_0(t) \exp(\boldsymbol{x}_i^{(2)t} \boldsymbol{\beta}^{(2)} + w_i(t))$$

## Joint Models

## Parameterization $w_i(t)$ :

- current value:  $w_i(t) = \alpha z_{it}^t b_i$ , implemented in stan\_jm/JM/JMbayes
- ullet correlated random-effects:  $w_i(t)=lpha heta_i$ , where

$$\begin{bmatrix} \boldsymbol{b}_i \\ \boldsymbol{\theta}_i \end{bmatrix} \sim N \left( \begin{bmatrix} \boldsymbol{0} \\ \boldsymbol{0} \end{bmatrix}, \begin{bmatrix} \Sigma_{bb} \Sigma_{b\theta} \\ \Sigma_{b\theta} \sigma_{\theta}^2 \end{bmatrix} \right)$$

See Hickey et.al.  $(2018)^1$  for more parameterizations.

<sup>&</sup>lt;sup>1</sup>Hickey, G.L., Philipson, P., Jorgensen, A. & Kolamunnage-Dona, R. (2018). A comparison of joint models for longitudinal and competing risks data, with application to a epilepsy drug randomized controlled trial. Journal of the Royal Statistical Society A, 181(4), p. 1105-1123.

# Joint Model: Longitudinal Submodel

Hierarchical Multinomial:

## Probability state

 $k \in \{1, \dots, K\} = \{ \text{Fulltime}, \text{Parttime}, \text{Outside the labour market} \} :$ 

$$\pi_{itk} = P(Y_{it} = k \mid \boldsymbol{x}_{it}, b_{ih}) = \frac{\frac{1}{1 + \sum_{h=1}^{K-1} \exp(\boldsymbol{x}_{it}\boldsymbol{\beta}_h + b_{ih})}} \text{ if } k = K}{\frac{\exp(\boldsymbol{x}_{it}\boldsymbol{\beta}_h + b_{ih})}{1 + \sum_{h=1}^{K-1} \exp(\boldsymbol{x}_{it}\boldsymbol{\beta}_h + b_{ih})}} \text{ if } k = 1, \dots, K-1,$$

$$x_{it}\beta_h = \beta_{0h} + \beta_{1h} * \mathsf{Sex}_i + \beta_{2h} * \mathsf{Education}_i$$
  
 $b_{ih} = \mathsf{random}$  intercept

## Joint Model: Survival Submodel

Hierarchical Poisson log-linear model (proportional hazards model with piece-wise constant baseline hazard):

$$\log \mu_{it} = \log t_{it} + \boldsymbol{x}_{it}\boldsymbol{\eta} + \alpha_t + u_i,$$

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\begin{array}{l} \mu_{it} = \mathsf{hazard} \\ \alpha_t = \log \lambda_t = \mu_\lambda + N(\log \lambda_{t-1}, \sigma_\lambda) \text{ (random walk baseline hazard)} \\ \boldsymbol{x}_{it} \boldsymbol{\eta} = \eta_1 * \mathsf{Sex}_i + \eta_2 * \mathsf{Education}_i \\ u_i = \mathsf{frailty} \\ t_{it} = \mathsf{offset} \end{array}
```

## Joint Model: Link

Random intercept  $b_{ih}$  and frailty  $u_i$  assumed to follow a multivariate normal distribution with the following variance-covariance matrix:

$$\boldsymbol{\Sigma} = \begin{bmatrix} \Sigma_{\boldsymbol{b}} & \Sigma'_{\boldsymbol{b}u} \\ \Sigma_{\boldsymbol{b}u} & \sigma_u^2 \end{bmatrix},$$

$$\Sigma_{\boldsymbol{b}} = \begin{bmatrix} \sigma_{\boldsymbol{b}_1}^2 & \sigma_{\boldsymbol{b}_1, \boldsymbol{b}_2} \\ \sigma_{\boldsymbol{b}_2, \boldsymbol{b}_1} & \sigma_{\boldsymbol{b}_2}^2 \end{bmatrix}, \ \Sigma_{\boldsymbol{b}\boldsymbol{u}} = (\sigma_{\boldsymbol{b}_1, \boldsymbol{u}}, \sigma_{\boldsymbol{b}_2, \boldsymbol{u}})^t$$

## Joint Model: Estimation

Bayesian Model in Stan on a subset of the data (  $\pm 10\%$  stratified sample)

#### Priors:

- regression coefficients  $(\beta, \eta)$ :
  - $N(\mu_0 = 0, \sigma_0 = 10000)$
- baseline hazards:
  - $\mu_{\lambda} \sim N(0,1)$
  - $\sigma_{\lambda} \sim N(0,1)$
  - $\log(\lambda_0) \sim N(0,1)$
- random effects:
  - $(\boldsymbol{b}_i, u_i) \sim MVN(\boldsymbol{0}, \boldsymbol{\Sigma})$
- ullet variance-covariance random effects  $(\Sigma)$ :
  - LKJ(2) prior for correlations (LKJ(1) = U(-1,1))
  - $\exp(\lambda = 0.5)$  prior for standard deviation

## Random Effects/Frailties

Instead of looking at the variance-covariance matrix we look at a matrix with correlations on the off-diagnonals and standard deviations on the diagonals:

$$R = \begin{bmatrix} \sigma_{b_1} & r_{(b_1,b_2)} & r_{(b_1,u)} \\ r_{(b_1,b_2)} & \sigma_{b_2} & r_{(b_2,u)} \\ r_{(b_1,u)} & r_{(b_2,u)} & \sigma_u \end{bmatrix}$$
(1)

# Random Effects/Frailties

	19	950 cohort	19	953 cohort	1954 cohort		
	mean	CI	mean	CI	mean	CI	
$r_{(\boldsymbol{b}_1,\boldsymbol{b}_2)}$	0.61	(0.53, 0.69)	0.64	(0.56, 0.71)	0.64	(0.56, 0.72)	
$r_{(\boldsymbol{b}_1, \boldsymbol{u})}$	-0.79	(-0.95, -0.58)	-0.82	(-0.96, -0.63)	-0.77	(-0.95, -0.53)	
$r_{(oldsymbol{b}_2,oldsymbol{u})}$	-0.35	(-0.59, -0.11)	-0.53	(-0.73, -0.31)	-0.49	(-0.76, -0.20)	
$\sigma_{\boldsymbol{b}_1}$	5.67	(5.16, 6.25)	5.97	(5.41, 6.59)	6.31	(5.73, 6.98)	
$\sigma_{m{b}_2}$	3.90	(3.55, 4.28)	4.90	(4.45, 5.42)	5.21	(4.70, 5.79)	
$\sigma_{m{u}}$	0.50	(0.32, 0.70)	0.57	(0.38, 0.79)	0.44	(0.25, 0.65)	

## Fixed Effects: Survival Submodel

#### Hazard Ratios:

	1950 cohort		1953 cohort		1954 cohort	
	mean	CI	mean	CI	mean	CI
sex	1.47	(1.14, 1.87)	1.51	(1.17, 1.91)	1.92	(1.45, 2.49)
education	0.78	(0.59, 1.01)	0.78	(0.59, 1.01)	0.64	(0.46, 0.85)

# Fixed Effects: Longitudinal Submodel

#### Coefficients and odds ratios:

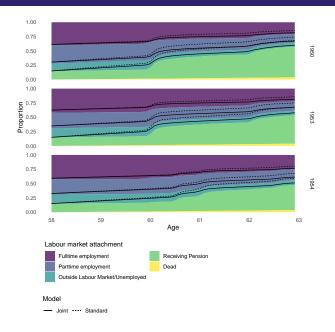
			1950 co	hort	1954 cohort			
		mean		CI	mean		CI	
		β	$\exp(\beta)$		β	$\exp(\beta)$		
	intercept	2.65	14.15	(1.91, 3.37)	3.00	20.09	(2.12, 3.90)	
fulltime employment	sex	-2.23	0.11	(-3.28, -1.23)	-2.06	0.13	(-3.25, -0.88	
	education	1.81	6.11	(0.69, 2.97)	2.29	9.87	(1.05, 3.52)	
	intercept	1.51	4.53	(1.00, 2.05)	1.19	3.29	(0.43, 1.97)	
parttime employment	sex	0.65	1.92	(-0.07, 1.34)	0.31	1.36	(-0.68, 1.30	
	education	0.53	1.70	(-0.28, 1.39)	1.71	5.53	(0.60, 2.77)	

# Fixed Effects: Longitudinal Submodel

#### Probabilities:

		1950 cohort		1954 cohort	
		mean	CI	mean	CI
mean	CI				
	male + low	0.71	(0.59, 0.82)	0.82	(0.71, 0.90)
fulltime	male + high	0.90	(0.81, 0.96)	0.90	(0.81, 0.96)
Tulltime	female + low	0.14	(0.21, 0.57)	0.32	(0.18, 0.48)
	female + high	0.37	(0.10, 0.39)	0.49	(0.29, 0.70)
	male + low	0.23	(0.14, 0.34)	0.14	(0.08, 0.22)
parttime	male + high	0.09	(0.04, 0.17)	0.09	(0.03, 0.18)
parttime	female + low	0.77	(0.67, 0.85)	0.55	(0.40, 0.70)
	female + high	0.59	(0.40, 0.74)	0.49	(0.29, 0.69)
	male + low	0.05	(0.03, 0.09)	0.04	(0.02, 0.09)
outside	male + high	0.01	(0.00, 0.03)	0.01	(0.00, 0.01)
outside	female + low	0.09	(0.05, 0.15)	0.13	(0.06, 0.23)
	female + high	0.04	(0.02, 0.08)	0.02	(0.01, 0.05)

# Comparison to standard survival model



#### Extensions

- Increase computational speed
  - Replace multinomial setup with multiple logistic models
  - Parallell MCMC computation
  - Different parameterization (random effects + covariances)
- Include death in a competing risks setup
- Inclusion of health related variables (e.g. disease diagnosis, medicine usage) as additional longitudinal or survival outcomes