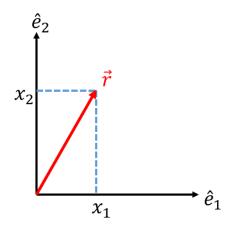
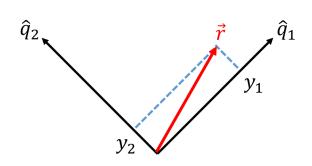
Ellipse and Gaussian Distribution

Industrial AI Lab.

Coordinates with Basis

• Basis $\{\hat{e}_1 \ \hat{e}_2\}$ or basis $\{\hat{q}_1 \ \hat{q}_2\}$

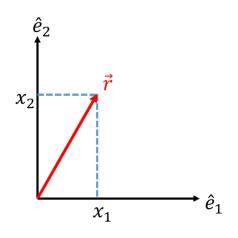


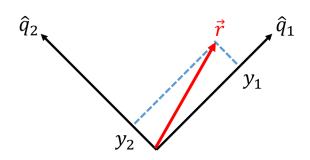


$$\begin{array}{l} \overrightarrow{r}_I = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \text{: coordinate of } \overrightarrow{r} \text{ in basis } \{ \hat{e}_1 \ \hat{e}_2 \} \ (=I) \\ \overrightarrow{r}_Q = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix} \text{: coordinate of } \overrightarrow{r} \text{ in basis } \{ \hat{q}_1 \ \hat{q}_2 \} \ (=Q) \end{array}$$

$$\overrightarrow{r}_Q = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}$$
: coordinate of \overrightarrow{r} in basis $\{\hat{q}_1 \ \hat{q}_2\}$ (= Q

Coordinate Transformation





$$egin{aligned} \overrightarrow{r} &= x_1 \hat{e}_1 + x_2 \hat{e}_2 = y_1 \hat{q}_1 + y_2 \hat{q}_2 \ &= \left[egin{aligned} \hat{e}_1 & \hat{e}_2 \end{array}
ight] egin{bmatrix} x_1 \ x_2 \end{bmatrix} = \left[egin{aligned} \hat{q}_1 & \hat{q}_2 \end{array}
ight] egin{bmatrix} y_1 \ y_2 \end{bmatrix} \end{aligned}$$

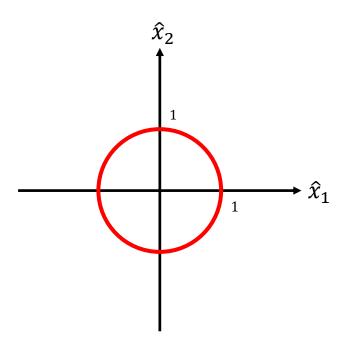
$$egin{aligned} \Longrightarrow egin{aligned} egin{aligned} & \Longrightarrow egin{aligned} egin{aligned} x_1 \ x_2 \end{bmatrix} & = Q egin{bmatrix} y_1 \ y_2 \end{bmatrix} = Q^{-1} egin{bmatrix} x_1 \ x_2 \end{bmatrix} = Q^T egin{bmatrix} x_1 \ x_2 \end{bmatrix} \end{aligned}$$

Coordinate Transformation

• Coordinate change to basis of $\{\hat{q}_1 \; \hat{q}_2\}$

$$egin{bmatrix} egin{bmatrix} x_1 \ x_2 \end{bmatrix} & Q^T & egin{bmatrix} y_1 \ y_2 \end{bmatrix} \ ext{coordinate in } I & ext{coordinate in } Q \end{pmatrix}$$

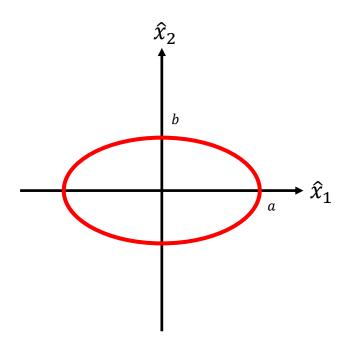
• Unit circle



$$x_1^2 + x_2^2 = 1 \implies$$

$$\left[egin{array}{cc} x_1 & x_2 \end{array}
ight] \left[egin{array}{cc} 1 & 0 \ 0 & 1 \end{array}
ight] \left[egin{array}{cc} x_1 \ x_2 \end{array}
ight] = 1$$

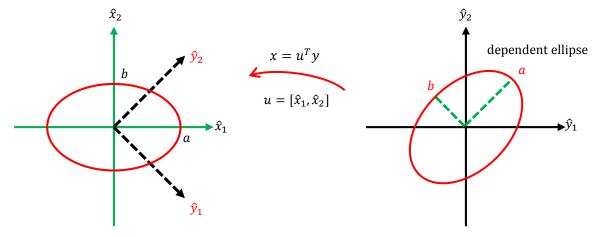
Independent ellipse



$$egin{aligned} rac{x_1^2}{a^2} + rac{x_2^2}{b^2} &= 1 \implies [\ x_1 \ \ x_2] egin{bmatrix} rac{1}{a^2} & 0 \ 0 & rac{1}{b^2} \end{bmatrix} egin{bmatrix} x_1 \ x_2 \end{bmatrix} &= 1 \ & \Rightarrow [\ x_1 \ \ x_2] \ \Sigma_x^{-1} egin{bmatrix} x_1 \ x_2 \end{bmatrix} &= 1 \end{aligned}$$

where
$$\Sigma_x^{-1}=egin{bmatrix} rac{1}{a^2} & 0 \ 0 & rac{1}{b^2} \end{bmatrix},~\Sigma_x=egin{bmatrix} a^2 & 0 \ 0 & b^2 \end{bmatrix}$$

- Dependent ellipse (Rotated ellipse)
 - Coordinate changes



$$egin{bmatrix} x_1 \ x_2 \end{bmatrix} = u^T egin{bmatrix} y_1 \ y_2 \end{bmatrix}, \hspace{5mm} x = u^T y \ ux = y \end{bmatrix}$$

• Now we know in basis $\{\hat{x}_1, \hat{x}_2\} = I$

$$x^T \Sigma_x^{-1} x = 1 \quad ext{and} \quad \Sigma_x = egin{bmatrix} a^2 & 0 \ 0 & b^2 \end{bmatrix}$$

• Then, we can find Σ_y such that

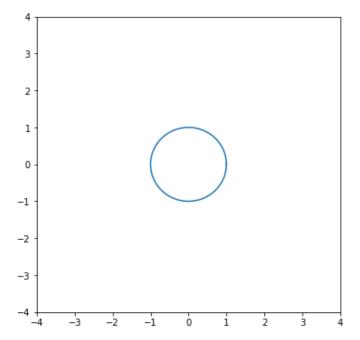
$$y^T \Sigma_y^{-1} y = 1 \quad \text{and} \quad \Sigma_y = ?$$
 $\implies x^T \Sigma_x^{-1} x = y^T u \Sigma_x^{-1} u^T y = 1 \quad (\Sigma_y^{-1} : \text{similar matrix to } \Sigma_x^{-1})$

$$\Sigma_y^{-1} = u \Sigma_x^{-1} u^T \quad \text{or} \quad \Sigma_y = u \Sigma_x u^T$$

```
import numpy as np
import matplotlib.pyplot as plt
%matplotlib inline
```

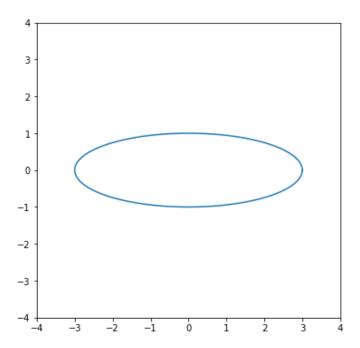
```
theta = np.arange(0,2*np.pi,0.01)
x1 = np.cos(theta)
x2 = np.sin(theta)

plt.figure(figsize=(6,6))
plt.plot(x1, x2)
plt.xlim([-4,4])
plt.ylim([-4,4])
plt.show()
```

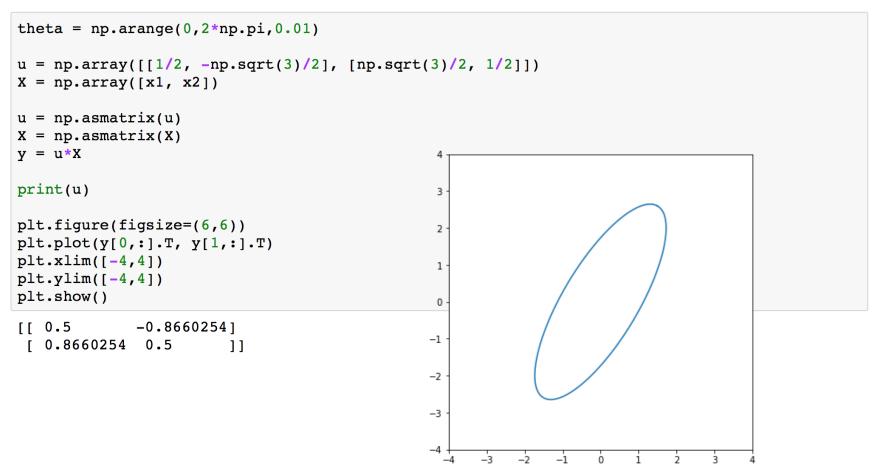


```
x1 = 3*np.cos(theta);
x2 = np.sin(theta);

plt.figure(figsize=(6,6))
plt.plot(x1, x2)
plt.xlim([-4,4])
plt.ylim([-4,4])
plt.show()
```



$$u=[\hat{x}_1\;\hat{x}_2]=\left[egin{array}{cc} rac{1}{2} & -rac{\sqrt{3}}{2} \ rac{\sqrt{3}}{2} & rac{1}{2} \end{array}
ight]$$



Question (Reverse Problem)

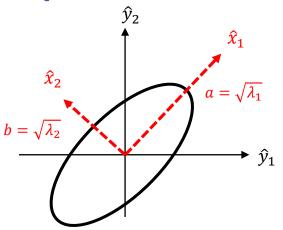
- Given Σ_y^{-1} (or Σ_y), how to find a (major axis) and b (minor axis) or
- How to find the proper matrix u
- Eigenvectors of Σ

$$A = S\Lambda S^T$$
 where $S = [v_1 \ v_2]$ eigenvector of A , and $\Lambda = \begin{bmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{bmatrix}$

here,
$$\Sigma_y = u \Sigma_x u^T = u \Lambda u^T$$

where
$$u = \begin{bmatrix} \hat{x}_1 & \hat{x}_2 \end{bmatrix}$$
 eigenvector of Σ_y , and $\Lambda = \begin{bmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{bmatrix} = \begin{bmatrix} a^2 & 0 \\ 0 & b^2 \end{bmatrix}$

Question (Reverse Problem)



$$\text{eigen-analysis} \left\{ \begin{array}{l} \Sigma_y \hat{x}_1 = \lambda_1 \hat{x}_1 \\ \Sigma_y \hat{x}_2 = \lambda_2 \hat{x}_2 \end{array} \right. \implies \left. \Sigma_y \underbrace{\left[\left. \hat{x}_1 \quad \hat{x}_2 \right]}_{u} = \underbrace{\left[\left. \hat{x}_1 \quad \hat{x}_2 \right]}_{u} \underbrace{\left[\left. \hat{x}_1 \quad \hat{x}_2 \right]}_{u} \underbrace{\left[\left. \hat{x}_1 \quad \hat{x}_2 \right]}_{u} \right] \right. \right\}$$

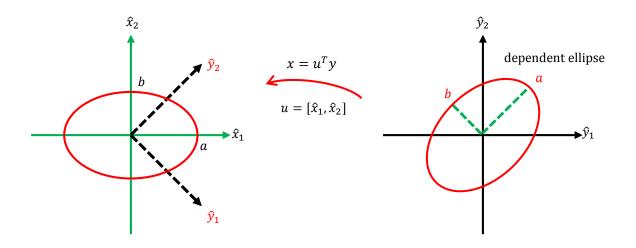
$$egin{aligned} \Sigma_y u &= u \Lambda \ \Sigma_y &= u \Lambda u^T = u \Sigma_x u^T \end{aligned}$$

Therefore

$$egin{aligned} x = u^T y & a = \sqrt{\lambda_1} \ egin{bmatrix} x_1 \ x_2 \end{bmatrix} = u^T egin{bmatrix} y_1 \ y_2 \end{bmatrix} & a = \sqrt{\lambda_1} \ b = \sqrt{\lambda_2} \ & ext{major axis} = \hat{x}_1 \ & ext{minor axis} = \hat{x}_2 \end{aligned}$$

Question (Reverse Problem)

Summary

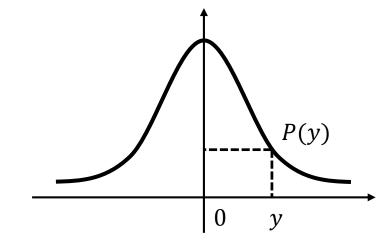


- Independent ellipse in $\{\hat{x}_1, \hat{x}_2\}$
- Dependent ellipse in $\{\hat{y}_1, \hat{y}_2\}$
- Decouple
 - diagonalize
 - eigen-analysis

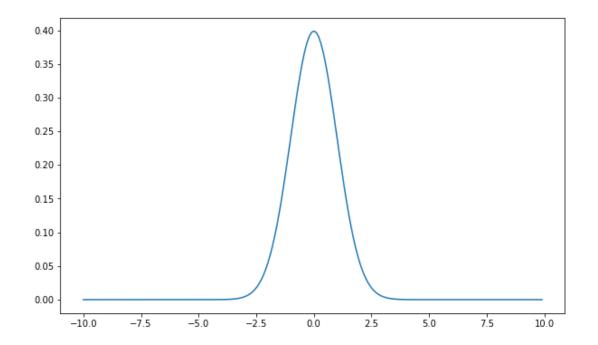
$$egin{aligned} x &= u^T y \ u &= [\ \hat{x}_1 \quad \hat{x}_2 \] \end{aligned}$$

$$P_{Y}\left(Y=y
ight)=rac{1}{\sqrt{2\pi}}\mathrm{exp}igg(-rac{1}{2}y^{2}igg)$$

$$\frac{1}{2}y^2 = \mathrm{const} \implies \mathrm{prob.\ contour}$$



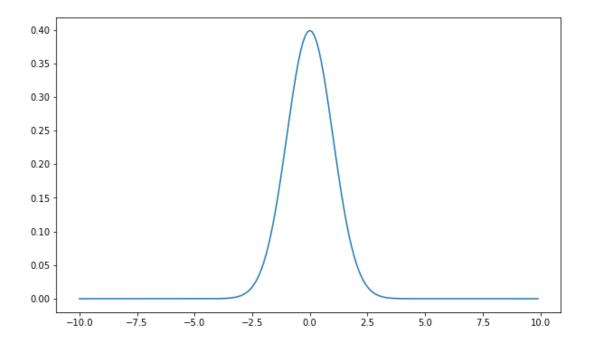
```
y = np.arange(-10,10,0.1)
ProbG = 1/np.sqrt(2*np.pi)*np.exp(-1/2*y**2)
plt.figure(figsize=(10,6))
plt.plot(y, ProbG)
plt.show()
```



```
from scipy.stats import norm

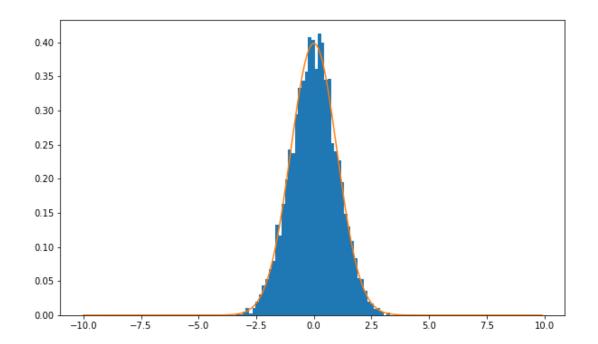
ProbG2 = norm.pdf(y)

plt.figure(figsize=(10,6))
plt.plot(y, ProbG2)
plt.show()
```



```
x = np.random.randn(5000,1)

plt.figure(figsize=(10,6))
plt.hist(x, bins=51, normed=True)
plt.plot(y, ProbG2, label='G2')
plt.show()
```



• Gaussian or normal distribution, 1D (mean μ , variance σ^2)

$$N(x;\,\mu,\sigma) = rac{1}{\sqrt{2\pi}\sigma} \mathrm{exp} \Biggl(-rac{1}{2} rac{(x-\mu)^2}{\sigma^2} \Biggr)$$

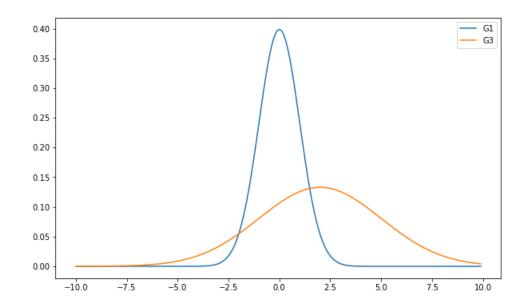
$$x \sim N\left(\mu, \sigma^2\right)$$
 $\implies P_Y\left(y\right) = P_X\left(x\right), \quad y = \frac{x - \mu}{\sigma}$
 $P_X\left(X = x\right) \sim \exp\left(-\frac{1}{2}\left(\frac{x - \mu}{\sigma}\right)^2\right)$
 $= \exp\left(-\frac{1}{2}\frac{\left(x - \mu\right)^2}{\sigma^2}\right)$

```
mu = 2
sigma = 3

x = np.arange(-10, 10, 0.1)

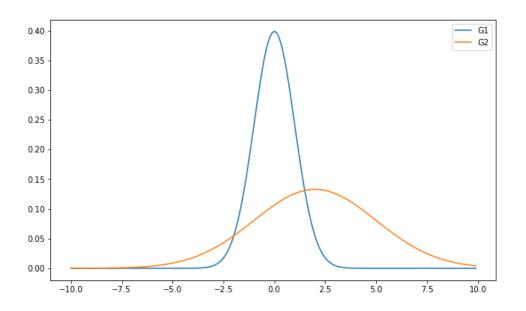
ProbG3 = 1/(np.sqrt(2*np.pi)*sigma) * np.exp(-1/2*(x-mu)**2/(sigma**2))

plt.figure(figsize=(10,6))
plt.plot(y,ProbG, label='G1')
plt.plot(x,ProbG3, label='G3')
plt.legend()
plt.show()
```



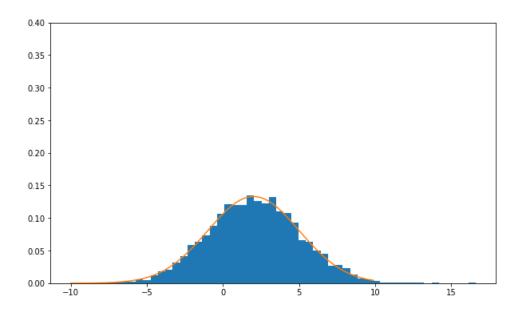
```
ProbG2 = norm.pdf(x, 2, 3)

plt.figure(figsize=(10,6))
plt.plot(y,ProbG, label='G1')
plt.plot(x,ProbG2, label='G2')
plt.legend()
plt.show()
```



```
x = mu + sigma*np.random.randn(5000,1)

plt.figure(figsize=(10,6))
plt.hist(x, bins=51, normed=True)
plt.plot(y,ProbG2, label='G2')
plt.ylim([0,0.4])
plt.show()
```



Multivariate Gaussian Models

Similar to a univariate case, but in a matrix form

$$egin{aligned} Nig(x;\,\mu,\Sigmaig) &= rac{1}{(2\pi)^{rac{n}{2}}|\Sigma|^{rac{1}{2}}} \mathrm{exp}igg(-rac{1}{2}(x-\mu)^T\Sigma^{-1}\,(x-\mu)igg) \ &= \mathrm{length}\,\,n\,\,\mathrm{column}\,\,\mathrm{vector} \ \Sigma &= n imes n\,\,\mathrm{matrix}\,\,(\mathrm{covariance}\,\,\mathrm{matrix}) \ |\Sigma| &= \mathrm{matrix}\,\,\mathrm{determinant} \end{aligned}$$

- Multivariate Gaussian models and ellipse
 - Ellipse shows constant Δ^2 value...

$$\Delta^2 = (x - \mu)^T \Sigma^{-1} (x - \mu)$$

$$P(X_1 = x_1, X_2 = x_2) = P_{X_1}(x_1) P_{X_2}(x_2)$$

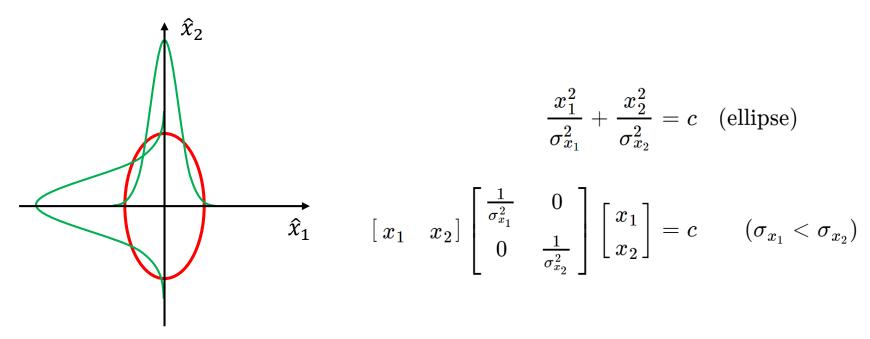
$$\sim \exp\left(-\frac{1}{2} \frac{\left(x_1 - \mu_{x_1}\right)^2}{\sigma_{x_1}^2}\right) \cdot \exp\left(-\frac{1}{2} \frac{\left(x_2 - \mu_{x_2}\right)^2}{\sigma_{x_2}^2}\right)$$

$$\sim \exp\left(-\frac{1}{2} \left(\frac{x_1^2}{\sigma_{x_1}^2} + \frac{x_2^2}{\sigma_{x_2}^2}\right)\right)$$

In a matrix form

$$P(x_1) \cdot P(x_2) = \frac{1}{Z_1 Z_2} \exp\left(-\frac{1}{2}(x - \mu)^T \Sigma^{-1}(x - \mu)\right)$$

$$\left(x = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}, \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \quad \Sigma = \begin{bmatrix} \sigma_{x_1}^2 & 0 \\ 0 & \sigma_{x_2}^2 \end{bmatrix}\right)$$



Summary in a matrix form

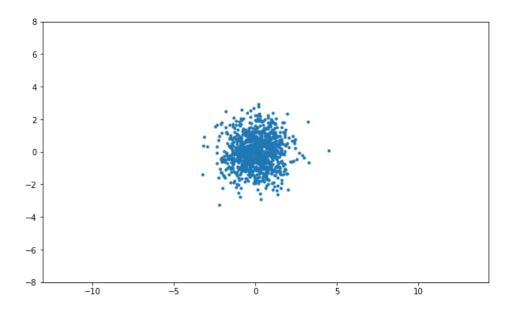
$$egin{aligned} N\left(0,\Sigma_{x}
ight) \sim \expigg(-rac{1}{2}x^{T}\Sigma_{x}^{-1}xigg) \ N\left(\mu_{x},\Sigma_{x}
ight) \sim \expigg(-rac{1}{2}(x-\mu_{x})^{T}\Sigma_{x}^{-1}\left(x-\mu_{x}
ight)igg) \end{aligned}$$

```
mu = np.array([0, 0])
sigma = np.eye(2)

m = 1000
x = np.random.multivariate_normal(mu, sigma, m)
print(x.shape)

plt.figure(figsize=(10,6))
plt.plot(x[:,0], x[:,1], '.')
plt.axis('equal')
plt.ylim([-8,8])
plt.show()
```

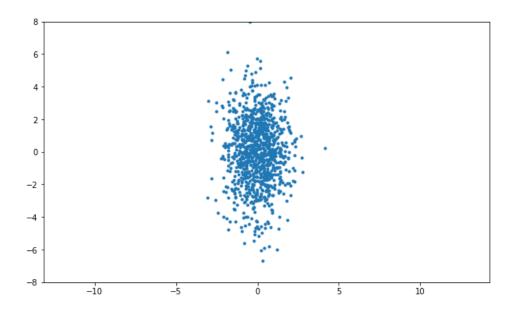
(1000, 2)

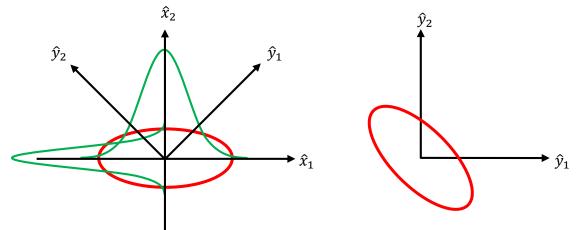


```
mu = np.array([0, 0])
sigma = np.array([[1, 0], [0, 4]])

m = 1000
x = np.random.multivariate_normal(mu, sigma, m)

plt.figure(figsize=(10,6))
plt.plot(x[:,0], x[:,1], '.')
plt.axis('equal')
plt.ylim([-8,8])
plt.show()
```





• Compute $P_Y(y)$ from $P_X(x)$

$$P_X(x) = P_Y(y) \;\; ext{where} \;\; x = \left[egin{array}{c} x_1 \ x_2 \end{array}
ight], y = \left[egin{array}{c} y_1 \ y_2 \end{array}
ight]$$

• Relationship between y and x

$$oldsymbol{x} = \left[egin{array}{cc} \hat{x}_1 & \hat{x}_2
ight]^T y = u^T y$$

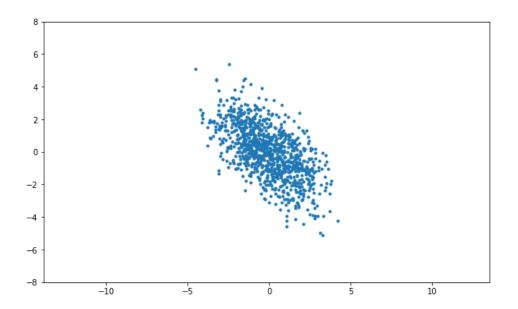
$$egin{aligned} x^T \Sigma_x^{-1} x &= y^T u \Sigma_x^{-1} u^T y = y^T \Sigma_y^{-1} y \ dots & \Sigma_y^{-1} &= u \Sigma_x^{-1} u^T \ & o \Sigma_y &= u \Sigma_x u^T \end{aligned}$$

- Σ_x : covariance matrix of x
- Σ_{v} : covariance matrix of y
- If u is an eigenvector matrix of Σ_{v} , then Σ_{x} is a diagonal matrix

```
mu = np.array([0, 0])
sigma = 1./2.*np.array([[5, -3], [-3, 5]])

m = 1000
x = np.random.multivariate_normal(mu, sigma, m)

plt.figure(figsize=(10,6))
plt.plot(x[:,0], x[:,1], '.')
plt.axis('equal')
plt.ylim([-8,8])
plt.show()
```



Remark

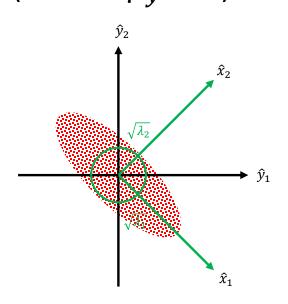
$$x \sim N(\mu_x, \Sigma_x)$$
 and $y = Ax + b$ affine transformation

$$\implies y \sim N(\mu_y, \Sigma_y) = N(A\mu_x + b, A\Sigma_x A^T)$$

$$\implies$$
 y is also Gaussian with $\mu_y = Ax + b$, $\Sigma_y = A\Sigma_x A^T$

Decouple using Covariance Matrix

• Given data, how to find Σ_y and major (or minor) axis (assume $\mu_v=0$)



$$\Sigma_y = egin{bmatrix} ext{var}(y_1) & ext{cov}(y_1,y_2) \ ext{cov}(y_2,y_1) & ext{var}(y_2) \end{bmatrix}$$

$$egin{align} ext{eigen-analysis} & \Sigma_x^{-1} = egin{bmatrix} rac{1}{\sqrt{\lambda_1}^2} & 0 \ 0 & rac{1}{\sqrt{\lambda_2}^2} \end{bmatrix} \ \Sigma_y \hat{x}_1 = \lambda_1 \hat{x}_1 \ \Sigma_y \hat{x}_2 = \lambda_2 \hat{x}_2 & \Sigma_x = egin{bmatrix} \sqrt{\lambda_1}^2 & 0 \ 0 & \sqrt{\lambda_2}^2 \end{bmatrix} \end{split}$$

$$egin{aligned} \Sigma_y \left[egin{array}{ccc} \hat{x}_1 & \hat{x}_2
ight] & = \left[\hat{x}_1 & \hat{x}_2
ight] \left[egin{array}{ccc} \lambda_1 & 0 \ 0 & \lambda_2 \end{array}
ight] & y = ux \implies u^T y = x \ & = \left[\hat{x}_1 & \hat{x}_2
ight] \Sigma_x & \left[egin{array}{ccc} \hat{x}_1 & \hat{x}_2 \end{array}
ight] & = u \end{aligned}$$

$$\Sigma_y = u \Sigma_x u^T$$

Decouple using Covariance Matrix

```
x.shape
S = np.cov(x.T)
print ("S = \n", S)
S =
[[ 2.59216411 -1.54924881]
 [-1.54924881 2.54567035]]
D, U = np.linalq.eiq(S)
idx = np.argsort(-D)
D = D[idx]
U = U[:,idx]
print ("U = \n", U)
print ("D = \n", D)
U =
[[ 0.7123916  0.70178217]
[-0.70178217 \quad 0.7123916]
D =
 [ 4.11834045 1.01949402]
```

Decouple using Covariance Matrix

```
xp = np.arange(-10, 10)
plt.figure(figsize=(10,6))
plt.plot(x[:,0],x[:,1],'.')
plt.plot(xp, U[1,0]/U[0,0]*xp, label='u1')
plt.plot(xp, U[1,1]/U[0,1]*xp, label='u2')
plt.axis('equal')
plt.ylim([-8, 8])
plt.legend()
plt.show()
```

