# **Probabilistic Machine Learning**

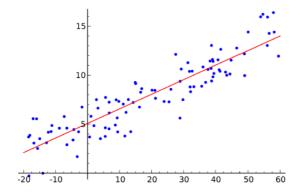
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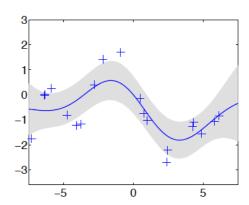
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# 1. Probabilistic Linear Regression

$$P(X \mid \theta) = \text{Probability [data | pattern]}$$





· Inference idea

data = underlying pattern + independent noise

· each response generated by a linear model plus some Gaussian noise

$$y = \omega^T x + arepsilon, \quad arepsilon \sim \mathcal{N}\left(0, \sigma^2
ight)$$

each response y then becomes a draw from the following Gaussian:

$$y \sim \left(\omega^T x, \sigma^2
ight)$$

· Probability of each response variable

$$P(y \mid x, \omega) = \mathcal{N}\left(\omega^T x, \sigma^2
ight) = rac{1}{\sqrt{2\pi\sigma^2}} \mathrm{exp}igg(-rac{1}{2\sigma^2}ig(y - \omega^T xig)^2igg)$$

• Given observed data  $D=\{(x_1,y_1),(x_2,y_2),\cdots,(x_m,y_m)\}$ , we want to estimate the weight vector  $\omega$ 

#### 1.1. Maximum Likelihood Solution

· Log-likelihood:

$$egin{aligned} \ell(\omega) &= \log L(\omega) = \log P(D \mid \omega) \ &= \log P(Y \mid X, \omega) \ &= \log \prod_{n=1}^m P\left(y_n \mid x_n, \omega
ight) \ &= \sum_{n=1}^m \log P\left(y_n \mid x_n, \omega
ight) \ &= \sum_{n=1}^m \log rac{1}{\sqrt{2\pi\sigma^2}} \mathrm{exp}\left(-rac{\left(y_n - \omega^T x_n
ight)^2}{2\sigma^2}
ight) \ &= \sum_{n=1}^m \left\{-rac{1}{2} \mathrm{log}ig(2\pi\sigma^2ig) - rac{\left(y_n - \omega^T x_n
ight)^2}{2\sigma^2}
ight\} \end{aligned}$$

· Maximum Likelihood Solution:

$$egin{aligned} \hat{\omega}_{MLE} &= rg \max_{\omega} \log P(D \mid \omega) \ &= rg \max_{\omega} \ -rac{1}{2\sigma^2} \sum_{n=1}^m ig(y_n - \omega^T x_nig)^2 \ &= rg \min_{\omega} rac{1}{2\sigma^2} \sum_{n=1}^m ig(y_n - \omega^T x_nig)^2 \ &= rg \min_{\omega} \sum_{n=1}^m ig(y_n - \omega^T x_nig)^2 \end{aligned}$$

It is equivalent to the least-squares objective for linear regression

#### 1.2. Maximum-a-Posteriori Solution

• Let's assume a Gaussian prior distribution over the weight vector  $\omega$ 

$$P(\omega) \sim \mathcal{N}\left(\omega \mid 0, \lambda^{-1}I
ight) = rac{1}{(2\pi)^{D/2}} \mathrm{exp}igg(-rac{\lambda}{2}\omega^T\omegaigg)$$

· Log posterior probability:

$$\log P(\omega \mid D) = \log rac{P(\omega)P(D \mid \omega)}{P(D)} = \log P(\omega) + \log P(D \mid \omega) - \underbrace{\log P(D)}_{ ext{constant}}$$

Maximum-a-Posteriori Solution:

$$\begin{split} \hat{\omega}_{MAP} \\ &= \arg\max_{\omega} \log P(\omega \mid D) \\ &= \arg\max_{\omega} \left\{ \log P(\omega) + \log P(D \mid \omega) \right\} \\ &= \arg\max_{\omega} \left\{ -\frac{D}{2} \log(2\pi) - \frac{\lambda}{2} \omega^{T} \omega + \sum_{n=1}^{m} \left\{ -\frac{1}{2} \log(2\pi\sigma^{2}) - \frac{\left(y_{n} - \omega^{T} x_{n}\right)^{2}}{2\sigma^{2}} \right\} \right\} \\ &= \arg\min_{\omega} \frac{1}{2\sigma^{2}} \sum_{n=1}^{m} \left(y_{n} - \omega^{T} x_{n}\right)^{2} + \frac{\lambda}{2} \omega^{T} \omega \\ &\text{(ignoring constants and changing max to min)} \end{split}$$

- For  $\sigma=1$  (or some constant) for each input, it's equivalent to the regularized least-squares objective
- BIG Lesson: MAP  $= l_2$  norm regularization

# 1.3. Summary: MLE vs MAP

· MLE solution:

$$\hat{\omega}_{MLE} = rg \min_{\omega} rac{1}{2\sigma^2} \sum_{n=1}^{m} ig(y_n - \omega^T x_nig)^2$$

· MAP solution:

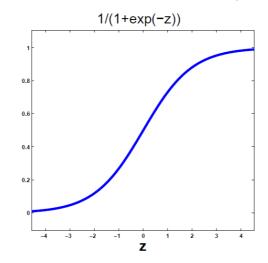
$$\hat{\omega}_{MLE} = rg \min_{\omega} rac{1}{2\sigma^2} \sum_{n=1}^m ig(y_n - \omega^T x_nig)^2 + rac{\lambda}{2} \omega^T \omega^T$$

- · Take-Home messages:
  - MLE estimation of a parameter leads to unregularized solutions
  - MAP estimation of a parameter leads to regularized solutions
  - The prior distribution acts as a regularizer in MAP estimation
- · Note: for MAP, different prior distributions lead to different regularizers
  - ullet Gaussian prior on  $\omega$  regularizes the  $l_2$  norm of  $\omega$
  - Laplace prior  $\exp(-C\|\omega\|_1)$  on  $\omega$  regularizes the  $l_1$  norm of  $\omega$

### 2. Probabilistic Linear Classification

- Often we do not just care about predicting the label y for an example
- Rather, we want to predict the label probabilities  $P(y \mid x, \omega)$ 
  - ullet E.g.,  $P(y=+1\mid x,\omega)$ : the probability that the label is +1
  - In a sense, it is our confidence in the predicted label
- Probabilistic classification models allow us do that
- Consider the following function in a compact expression (y=-1/+1):

$$P(y \mid x, \omega) = \sigma\left(y\omega^T x
ight) = rac{1}{1 + \exp(-y\omega^T x)}$$



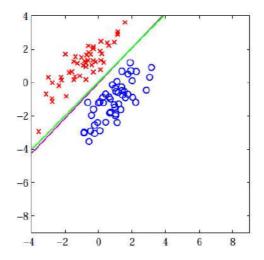
•  $\sigma$  is the logistic function which maps all real number into (0,1)

### 2.1. Logistic Regression

- · What does the decision boundary look like for logistic regression?
- ullet At the decision boundary labels -1/+1 becomes equiprobable

$$egin{aligned} P(y = +1 \mid x, \omega) &= P(y = -1 \mid x, \omega) \ rac{1}{1 + \exp(-\omega^T x)} &= rac{1}{1 + \exp(\omega^T x)} \ \exp\left(-\omega^T x
ight) &= \exp\left(\omega^T x
ight) \ \omega^T x &= 0 \end{aligned}$$

ullet The decision boundary is therefore linear  $\Longrightarrow$  logistic regression is a linear classifier



· note: it is possible to kernelize and make it nonlinear

#### 2.2. Maximum Likelihood Solution

- Goal: want to estimate  $\omega$  from the data  $D = \{(x_1, y_1), \cdots, (x_m, y_m)\}$
- · Log-likelihood:

$$egin{aligned} \ell(\omega) &= \log L(\omega) = \log P(D \mid \omega) \ &= \log P(Y \mid X, \omega) \ &= \log \prod_{n=1}^m P(y_n \mid x_n, \omega) \ &= \sum_{n=1}^m \log P(y_n \mid x_n, \omega) \ &= \sum_{n=1}^m \log rac{1}{1 + \exp(-y_n \omega^T x_n)} \ &= \sum_{n=1}^m - \log igl[ 1 + \exp(-y_n \omega^T x_n) igr] \end{aligned}$$

· Maximum Likelihood Solution:

$$\hat{\omega}_{MLE} = rg \max_{\omega} \log L(\omega) = rg \min_{\omega} \sum_{n=1}^m \log igl[ 1 + \expigl( -y_n \omega^T x_n igr) igr]$$

- No closed-form solution exists but we can do gradient descent on  $\omega$ 

$$egin{aligned} 
abla_{\omega} \log L(\omega) &= \sum_{n=1}^m -rac{1}{1+\exp(-y_n\omega^Tx_n)} \expigl(-y_n\omega^Tx_nigr)(-y_nx_nigr) \ &= \sum_{n=1}^m rac{1}{1+\exp(y_n\omega^Tx_n)} y_nx_n \end{aligned}$$

#### 2.3. Maximum-a-Posteriori Solution

- Let's assume a Gaussian prior distribution over the weight vector  $\omega$ 

$$P(\omega) = \mathcal{N}\left(\omega \mid 0, \lambda^{-1}I
ight) = rac{1}{(2\pi)^{D/2}} \mathrm{exp}igg(-rac{\lambda}{2}\omega^T\omegaigg)$$

Maximum-a-Posteriori Solution:

$$\begin{split} \hat{\omega}_{MAP} \\ &= \arg\max_{\omega} \log P(\omega \mid D) \\ &= \arg\max_{\omega} \{\log P(\omega) + \log P(D \mid \omega) - \underbrace{\log P(D)}_{\text{constant}} \} \\ &= \arg\max_{\omega} \{\log P(\omega) + \log P(D \mid \omega) \} \\ &= \arg\max_{\omega} \left\{ -\frac{D}{2} \log(2\pi) - \frac{\lambda}{2} \omega^T \omega + \sum_{n=1}^m -\log \left[1 + \exp\left(-y_n \omega^T x_n\right)\right] \right\} \\ &= \arg\min_{\omega} \sum_{n=1}^m \log \left[1 + \exp\left(-y_n \omega^T x_n\right)\right] + \frac{\lambda}{2} \omega^T \omega \\ &\text{(ignoring constants and changing max to min)} \end{split}$$

- BIG Lesson: MAP  $= l_2$  norm regularization
- No closed-form solution exists but we can do gradient descent on  $\omega$
- See "A comparison of numerical optimizers for logistic regression (http://research.microsoft.com/enus/um/people/minka/papers/logreg/minka-logreg.pdf)" by Tom Minka on optimization techniques (gradient descent and others) for logistic regression (both MLE and MAP)

## 2.4. Summary: MLE vs MAP

· MLE solution:

$$\hat{\omega}_{MLE} = rg\min_{\omega} \sum_{n=1}^{m} \logigl[1 + \expigl(-y\omega^T x_nigr)igr]$$

· MAP solution:

$$\hat{\omega}_{MAP} = rg \min_{\omega} \sum_{n=1}^{m} \log igl[ 1 + \expigl( -y \omega^T x_n igr) igr] + rac{\lambda}{2} \omega^T \omega^T$$

- · Take-home messages (we already saw these before)
  - MLE estimation of a parameter leads to unregularized solutions
  - MAP estimation of a parameter leads to regularized solutions
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- · Note: For MAP, different prior distributions lead to different regularizers
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# 3. Probabilistic Clustering

· will not cover in this course

### 4. Probabilistic Dimension Reduction

· will not cover in this course

In [1]:

%%javascript

\$.getScript('https://kmahelona.github.io/ipython\_notebook\_goodies/ipython\_notebook\_toc.
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