

---

# *FISH 621 Laboratory #5: Bayesian Mark-Recapture*

---

Curry Cunningham 2022

## Instructions

---

The purpose of this lab is to:

- Explore the Jolly-Seber mark-recapture model for open populations.
- Develop familiarity with implementing simple Bayesian analyses within the Stan platform.

If you have a question during the lab, please un-mute yourself and ask, or type it into the chat box. There is a high likelihood that someone else has the same question. It is more fun if we all learn together in our distance-learning world.

I have posted the lecture slides to the **Canvas site**, so you can reference this material as you work through the lab.

This and all other labs will be graded based on your attendance and participation.

## Lab Contents

---

- |   |   |
|---|---|
| • <b>621_Lab 5_Bayes.pdf</b>                  | (this file)   |
| • <b>621_Lab 5_Bayes.R</b>                    | R script with exercises.  |
| • <b>Jolly Seber.csv</b>                      | Example of Jolly-Seber analysis of Large-mouth bass at Par Pond from Hightower and Gilbert (1984) |
| • <b>Bristol Bay Spawner-Recruit Data.csv</b> | Spawner-recruit data for Bristol Bay Alaska sockeye salmon  |

## Exercise 1: Jolly-Seber

---

We will explore implementation of the Jolly-Seber model for a dataset of large-mouth bass at Par Pond from Hightower and Gilbert (1984). The sampling periods  $i$  in  $1:s = 6$  are subsequent weeks. The available data include:

- The number of bass captured during each sampling period:  $n_i$
- The total number of sampled individuals returned to the population during each sampling event, with tags:  $R_i$
- The matrix describing the number of individuals from each sample release that are captured during subsequent sampling events:  $m_{hi}$ . Recall in this type of experiment we mark individuals in

such a way that upon recapture we can tell in which sampling period they were originally marked (e.g. using differently colored tags during each sampling event).

- The rows  $h$  reference the release period (release cohorts)
- The columns  $i$  are the recovery periods
- The elements of the  $m_{hi}$  matrix are the number of fish from release period  $h$  that are recaptured during period  $i$

Please open the spreadsheet called **Jolly Seber.xlsx**. Green cells are data, and differently colored cells represent different derived parameters of the Jolly-Seber model. Please use the following steps to generate estimates:

1. Calculate  $m_i$  the total number of marked individuals observed during each sampling event (Cells B11:G11) for periods  $i = 1:6$ , as:  $m_i = \sum_{h=1}^{i-1} m_{hi}$ .
2. Calculate  $r_h$ , the number of  $R_i$  releases that are later recaptured (Cells I5:I9), for release groups  $h = 1:5$ , as:  $r_h = \sum_{i=h+1}^S m_{hi}$ .
3. Calculate the  $c_{hi}$  matrix (Cells B15:G19) where:
  - a. The first row of  $c_{hi}$  is equal to the first row of  $m_{hi}$ , or  $c_{1i} = m_{1i}$
  - b. Remaining rows for each recapture period  $i$  are the sum of captures of all earlier release groups  $h$ , or  $c_{hi} = c_{h-1,i} + m_{hi}$
4. Calculate the number of individuals for each release group  $h$  that before period  $i$ , that were not captured in period  $i$ , and are captured after period  $i$ , or  $z_{h+1}$ , based on the  $c_{hi}$  matrix.
  - a. Where  $z_{h+1}$  is calculated by summing each row of the  $c_{hi}$  matrix, as:  $z_{h+1} = \sum_{i=h+2}^S c_{hi}$
5. Next, we will calculate summary statistics in **Cells B24:O29**
6. Begin by copying summary values calculated from your  $m_{hi}$  and  $c_{hi}$  matrices into the appropriate columns of the **Summary Statistics** table.
  - a. Copy the total marks observed in each sampling period  $m_i$  from Cells B11:G11, into Cells D24:D29.
  - b. Copy the total number of releases that are later recaptured  $r_h$  from Cells I5:I9, into Cells E24:E28.
  - c. Copy  $z_{h+1}$  from Cells I15:I18 into Cells F25:F28.
7. Calculate the mark fraction in each sampling event  $\rho_i = m_i/n_i$ , in Cells G24:G29.
8. Calculate the **unbiased** estimate for the total number of marks  $M_i^* = \frac{(R_i+1)}{(r_i+1)}(z_i) + m_i$  for  $i = 2:5$ , in Cells H25:H28.
9. Calculate the **unbiased** estimate for the total number of unmarked individuals in the population  $U_i^* = \frac{M_i^*(n_i+2)}{m_i+1} - M_i^*$  for  $i = 2:5$ , in Cells I25:I28.
10. Calculate  $M_i^* - m_i$  for  $i = 2:5$ , in Cells J25:J28.
11. Calculate  $M_i^* - m_i + R_i$  for  $i = 2:5$ , in Cells K25:K28.
12. Calculate  $(1/r_i) - (1/R_i)$  for  $i = 1:5$ , in Cells L24:L28.
13. Calculate the **potentially biased** estimate of the total number of marked individuals in the population  $\hat{M}_i = \frac{R_i z_i}{r_i} + m_i$  for  $i = 2:5$ , in Cells M25:M28.
14. Calculate  $\hat{M}_i - m_i$  for  $i = 2:5$ , in Cells N25:N28.
15. Finally, calculate  $\hat{M}_i - m_i + R_i$  for  $i = 2:5$ , in cells O25:O28.
16. Now we have all of the pieces to calculate estimates for survival. Under the Population Estimates section please calculate

- a. For survival during the first sampling period use the approximation:  $\phi_{i=1}^* = M_{i+1}^*/R_i$  or  $\phi_2^* = M_2^*/R_1$  in Cell F33.
  - b. For subsequent sampling periods  $i = 2: (s - 2) = 2: 4$ , calculate survival as  $\phi_i^* = \frac{M_{i+1}^*}{\tilde{M}_i - m_i + R_i}$
17. Next, we can calculate our total population size estimates  $N_i^*$
- a. For periods  $i = 2: 5$  calculate  $M_i^* + U_i^* = \frac{M_i^*(n_i+2)}{m_i+1}$ , in Cells B34:B37.
  - b. Set cells in the  $N_i^*$  column (B34:B37) for periods  $i = 2: 5$ , equal to the  $M_i^* + U_i^*$  column (Cells C34:C37).
  - c. To estimate total abundance at the start ( $N_{i=1}^*$  in Cell C33) of period  $i = 1$ , we will use  $N_{i=2}^*$  and the estimated survival rate  $\phi_i^*$ , as:  $N_{i=1}^* = \frac{N_{i=2}^*}{\phi_{i=1}^*}$
  - d. Congratulations! You have estimated abundance at the start of each sampling period!**
18. Now that we have point estimates for abundance, we will focus on uncertainty in our estimates.
- a. Remember an estimate isn't all that useful if we don't have a sense of its precision.
19. First calculate the uncaptured number of individuals in the population at each time point  $N_i^* - n_i$  in Cells D33:D37.
20. We will first calculate the standard error in our estimate of  $N_i^*$  in Cells E33:E37, as
- a.  $SE(N_i^*) = \sqrt{N_i^*(N_i^* - n_i) \left( \left( \frac{M_i^* - m_i + R_i}{M_i^*} \right) \left( \frac{1}{r_i} - \frac{1}{R_i} \right) + \left( \frac{1 - \rho_i}{m_i} \right) \right)}$  For periods  $i = 2: 5$  in Cells E34:E37.
21. Calculate the coefficient of variation for the abundance estimates  $N_i^*$  as:  $CV(N_i^*) = SE(N_i^*)/N_i^*$ .
22. Next, we will calculate our estimate of recruitment  $B_i^*$  for periods  $i = 2: 4$  in Cells I35:I37, using the formula  $B_i^* = N_{i+1}^* - \phi_i^*(N_i^* - n_i + R_i)$ .

### Exercise 3: Bayesian Linear Regression

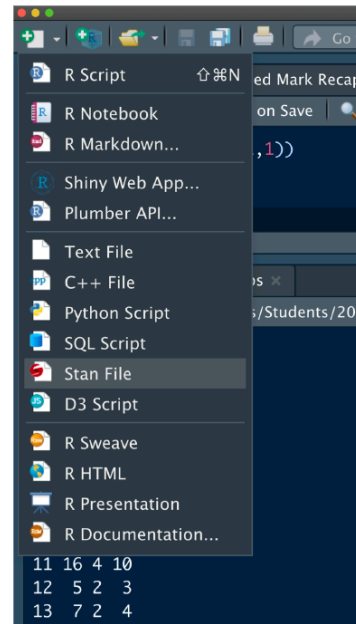
---

To familiarize ourselves with how we simulate data with R, and define and fit a Bayesian model with Stan we will start with a simple linear regression:  $y_i = \alpha + \beta x_i + e_i$ , where our observation errors are normally-distributed:  $e_i \sim \text{Normal}(0, \sigma^2)$  with standard deviation  $\sigma$ .

Please follow through the R script as we simulate our data, and then open the **lin\_reg.stan** script to see how we encode the Bayesian regression model using our Stan syntax.

# The .stan Script

- data
  - Define data inputs
    - Including dimensionality
- parameters
  - Define “free” or estimated (true) parameters
    - Names, dimensions, ect.
- transformed parameters
  - Define derived parameters
    - Quantities that depend on your estimated (true) parameters
  - Do calculations
  - *This is where the meat of your code is likely to exist!*
- model
  - Define priors for estimated parameters
  - Define likelihoods for the data
    - Probability of the data, given the model
- generated quantities
  - Calculations you want to do based on your
    - Estimated or derived parameters and data



## Stan Data Types

- Primitive types
  - real
    - Continuous values: 1.4, 0.9, -99.1, 100, ect.
  - int
    - Integer values: 1,2,3, ect.
- Vector and matrix types
  - Matrix-based types
    - vector, matrix, and row\_vector
  - Examples
    - `vector[3] myVect` – vector of length 3, named “myVect”
    - `matrix[3,3] myMat` – matrix with 3 rows, 3 columns named “myMat”

*I don't use these often*

- Array types
  - Any data type can be made into an array type
  - `array[10] real x;`
    - One-dimensional array of size 10 containing real values
  - `array[6,7] matrix[3,3] m;`
    - Declares “m” to be a two-dimensional array of size 6 x 7
    - Containing values that are *each* 3 x 3 matrices
- Alternative declarations
  - `real x[10];`
  - `matrix[3,3] m[6,7];`
- A vector of vectors
  - `vector[N] pred[S];`
    - A vector of length S where each element is a vector of length N
    - Accessed like a matrix or 2d-array: `pred[s,n]`

#### Exercise 4: Poisson Regression

---

R script

## Poisson Regression in Stan

- Estimate counts of peregrine falcons over  $n$  years
  - Linear predictor is a cubic polynomial
  - Random part of the response (statistical distribution)
    - $C_i \sim \text{Poisson}(\lambda_i)$
  - Link function of random and systematic part (log link)
    - $\log(\lambda_i) = \eta_i$
  - The systematic part of the response (linear predictor of  $\eta_i$ )
    - $\eta_i = \alpha + \beta_1 \text{year}_i + \beta_2 \text{year}_i^2 + \beta_3 \text{year}_i^3$

# Binomial GLM for Bounded Counts or Proportions

- While the Poisson distribution is a standard model for unbounded count data
  - Frequently we have counts that are bounded by an upper limit
- Example: when modelling number of fish in a population
  - The number counted cannot exceed the total population size
    - $n \sim \text{Binomial}(N, p)$
- Special case: Binary
  - Outcome of independent survival events
    - $\text{Survived}_t \sim \text{Binomial}(N = 1, p)$
- Example: successful bird breeding pairs
  - Data:
    - Successful breeding pairs ( $C_i$ ) out of some number of monitored pairs ( $N_i$ )
  - Goal:
    - Model the probability of successful breeding ( $p_i$ ) as a function of time
- Random part of the response (statistical distribution)
  - $C_i \sim \text{Binomial}(N_i, p_i)$
- Link of random and systematic part (logit link function)
  - $\text{logit}(p_i) = \log\left(\frac{p_i}{1-p_i}\right) = \eta_i$
- Systematic part of response (linear predictor  $\eta_i$ )
  - $\eta_i = \alpha + \beta_1 X_i + \beta_2 X_i^2$ 
    - Polynomial