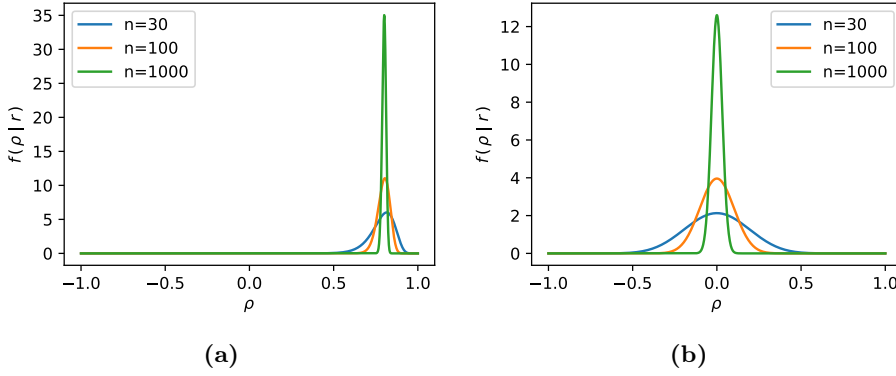


## 1.1 Confidence interval for absolute correlation in bivariate Gaussian

From [?], given a bivariate Gaussian, the confidence distribution of  $\rho$  given the empirical correlation  $r$  based on  $n$  observations is given by

$$f(\rho | r, \nu) = \frac{\nu(\nu-1)\Gamma(\nu-1)}{\sqrt{2\pi}\Gamma(\nu+\frac{1}{2})} \frac{(1-r^2)^{\frac{\nu-1}{2}} (1-\rho^2)^{\frac{\nu-2}{2}}}{(1-r\rho)^{\frac{2\nu-1}{2}}} F\left(\frac{3}{2}, -\frac{1}{2}, \nu+\frac{1}{2}, \frac{1+r\rho}{2}\right)$$

where  $F(a, b, c, z)$  is the Gaussian hypergeometric function and  $\nu = n - 1$ . That is, given a sample correlation  $r$ , what is the confidence in  $\rho$  in terms of a distribution. In the following figure, a sample correlation  $r = 0.8$  and  $r = 0$  has been used with varying number of observations (degrees of freedom) in figures Figure 1.1a and Figure 1.1b respectively. A key property is that  $f$  is *even symmetric* in  $\rho, r$ . That is  $f(\rho | r) = f(-\rho | -r)$ . Thus, a confidence interval for  $\rho$  given  $r$  is the negative of the confidence interval given  $-r$ . In particular, if we only observe  $|r|$ , we can calculate a confidence interval for  $\rho$  up to the sign of the bounds of the interval. Furthermore, as we want a CI for  $|\rho|$ , it does not matter if  $r$  is negative or positive. Hence, without loss of generality, we assume that  $r \geq 0$ . At this point, to construct a confidence interval for  $|\rho|$  we



**Figure 1.1:**  $f(\rho | r, \nu)$  shown for  $r = 0.8$  and  $r = 0$  in (a) and (b) with  $n \in \{30, 100, 1000\}$ . As one would expect, the power i.e. the width of the peak decreases with increasing  $n$  and for correlations closer to 0, the width is the largest.

list the following desired properties. Firstly, it should be an exact confidence interval, meaning that for a given significance level  $\alpha$ , the CI includes the true value exactly  $1 - \alpha$  fraction of the times. Secondly, if for a given  $r$ , it can not be rejected that  $\rho$  is 0, 0 should also be contained in the interval. Finally, if we reject that  $\rho = 0$ , we shall have  $\alpha/2$  probability mass above and below the bounds of the interval. The above is enough to uniquely define a confidence interval in all cases. Before continuing with how this CI is calculated, we mention that as  $r$  is an unbiased estimator of  $\rho$ , we would preferably want  $|r| \in CI_{1-\alpha}(|\rho|)$  (where  $CI_{1-\alpha}(|\rho|)$  denotes the  $1 - \alpha$  confidence interval for  $|\rho|$ ). However, although this will in almost every scenario be the case, we can not be sure of this from the above properties and in fact examples with large  $\alpha$  can be constructed such that  $|r|$  lies just outside the constructed CI.

First, to conform with the second desired property, if it can not be rejected that  $\rho = 0$  on a significance level  $\alpha$ , we will initially compute a CI for  $\rho$  (not  $|\rho|$ ) based on  $r$  (wlog chosen to be non-negative). This CI will just be a symmetric CI in the sense that  $\alpha/2$  of the probability mass will lie below the lower bound of the CI and above the upper bound of the CI respectively. If 0 is contained in this CI, we can not reject that  $\rho = 0$  and vice versa on an  $\alpha$  significance level. Thus, if 0 is contained in this initial CI for  $\rho$ , we will start the CI for  $|\rho|$  at 0 and determine an upper bound  $b$  such that  $\alpha$  probability mass is above this  $b$ . Otherwise, we shall find  $a$  and  $b$  such that  $\alpha/2$  probability mass is below  $a$  and above  $b$  respectively. Choosing  $a$  and  $b$  this way also conforms with the third property. Finally, to ensure that the CI contains exactly  $1 - \alpha$ , we define  $\hat{f}$  as

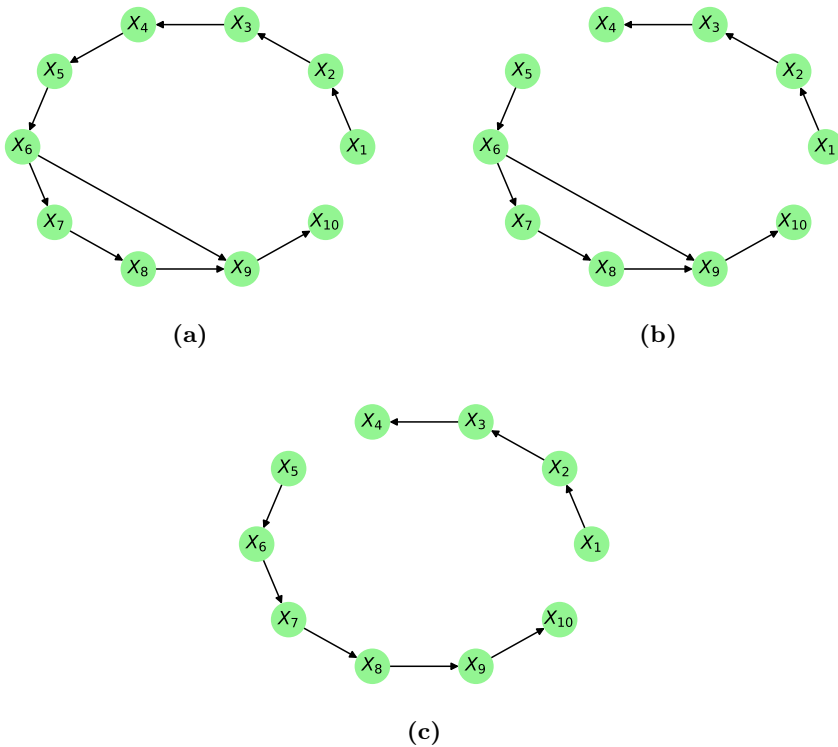
the reflected  $f$  in  $\rho$  such that

$$\tilde{f}(\rho_a | r_a, \nu) = f(\rho_a | r_a, \nu) + f(-\rho_a | r_a, \nu), \quad \rho_a, r \in [0, 1]$$

where  $\rho_a$  and  $r_a$  is the absolute correlation and empirical correlation respectively. With this  $\tilde{f}$ , the density at  $\rho_a$  is both the density for the negative and positive correlation ensuring that the  $\tilde{f}$  has probability mass 1. Thus, if  $a = 0$  (i.e. the CI must contain 0), we find  $b$  as the  $1 - \alpha$  percentile of  $\tilde{f}$  and if  $a \neq 0$ , we take  $a$  as the  $\alpha/2$  percentile and  $b$  as the  $1 - \alpha/2$  percentile of  $\tilde{f}$ .

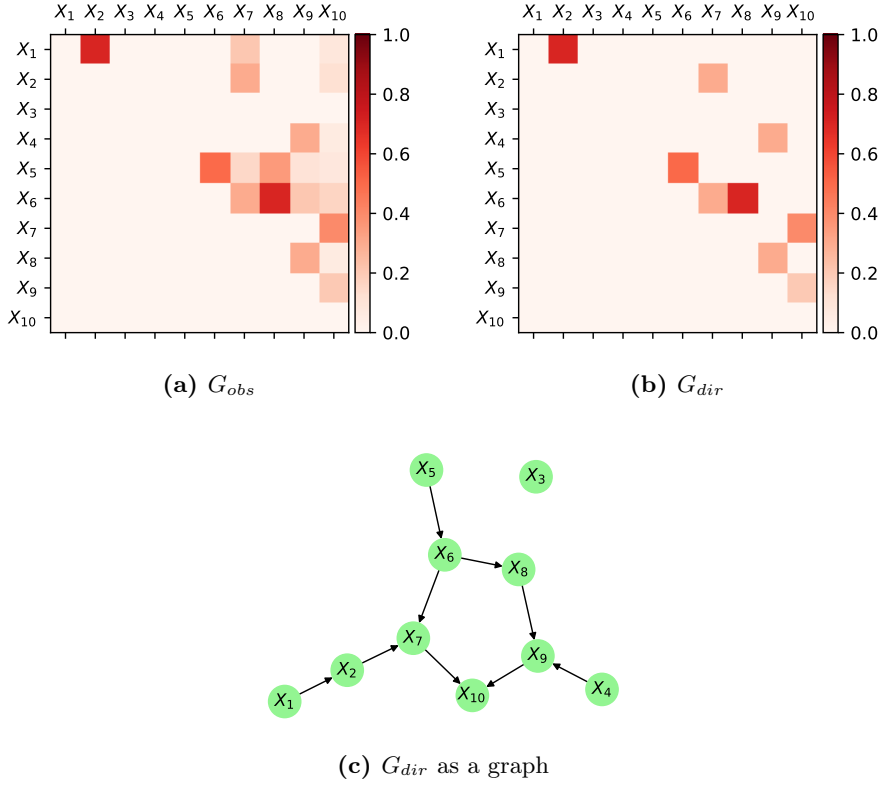
As an example, suppose  $r_a = 0.06$  with 1000 observations. Then a 95% CI for  $|\rho|$  is  $[0, 0.11164]$  whereas if on had observed  $r_a = 0.07$  the CI would be  $[0.01071, 0.1314]$ . These CI could then be used to test the absolute correlation of a bivariate Gaussian i.e. for  $r_a = 0.07$  based on 1000 observations would be rejected as stemming from a Gaussian with absolute correlation 0.01 on a 5% significance level.

## 1.2 Gaussian chain deconvolution



**Figure 1.2:** Triangular, mutual information, cutoff  $2 \cdot 10^{-10}$ ,  $2.1 \cdot 10^{-2}$  and  $4.51 \cdot 10^{-2}$ .

### 1.3 Gaussian network deconvolution



**Figure 1.3:** For the linear network defined in ??, using a triangular  $G_{obs}$  (a) with the true topological structure we are able to perfectly rediscover the causal structure as seen in (b) and (c).