Diagrammatic Monte Carlo

Exercises

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Task 1: Monte Carlo Basics Estimating PI

Estimating PI - Task

Two different methods to estimate π :

- · Area Method (as seen in the picture)
- MC Integration Method $\pi/4 = \int_0^1 \sqrt{1 x^2} dx = \int_0^1 \frac{\sqrt{1 x^2}}{p(x)} p(x) dx =$ $\langle \frac{\sqrt{1 x^2}}{p(x)} \rangle_p \approx \frac{1}{N} \sum_{i=1}^N \frac{\sqrt{1 x_i^2}}{p(x_i)} \text{ for } p(x) = 1$ and $x_i \sim U(0, 1)$

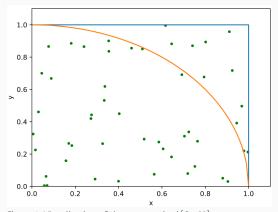


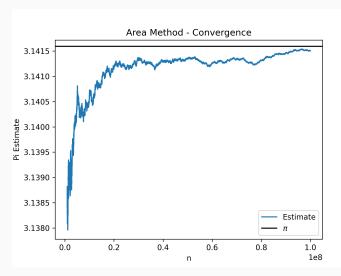
Figure 1: Visualization of the area method.[Ces22]

Estimating PI: Area Method - Implementation

```
def area_method(n, plot=False):
         n in = 0
2
         estimates_n = np.array([], dtype=int)
         estimates area = np.array([])
         for i in range(n):
             x = np.random.random()
             y = np.random.random()
             if x**2 + y**2 <= 1:
                 n in += 1
12
             if plot:
13
                 if i % (int(n / 100000)) == 0 and i > 1000000:
14
                      estimates n = np.append(estimates n, i + 1)
1.5
                      estimates area = np.append(estimates area, 4 * n in / (i + 1))
16
         if plot:
17
18
19
         return 4 * n in / n
20
```

Estimating PI: Area Method - Convergence

area_method(int(1e8), plot=True)



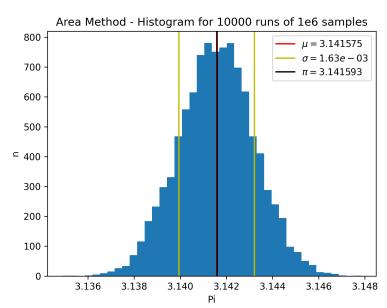
Estimating PI: Area Method - We can do better

```
def area_method_fast(n):
    x = np.random.random(n)
    y = np.random.random(n)
    n_in = np.sum(x**2 + y**2 <= 1)
    return 4 * n_in / n</pre>
```

```
pis = np.zeros(10000)
for i in range(10000):
    pis[i] = area_method_fast(int(1e6))

std = np.std(pis)
mean = np.mean(pis)
# Do the plotting here
# ...
print("Area Method:")
print(f"{mean = }")
print(f"{std = }")
```

Estimating PI: Area Method - Histogram

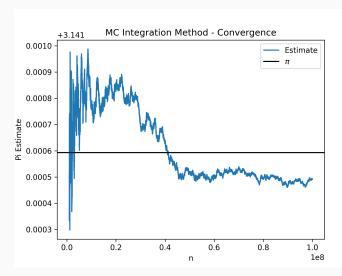


Estimating PI: MC Integration Method - Implementation

```
def p(x):
        return 1
3
    def f(x):
        return np.sqrt(1 - x**2)
5
6
    def mc_integrate(n, plot=False):
        sum = 0
         estimates n = np.array([], dtype=int)
         estimates area = np.array([], dtype=np.float32)
10
        for i in range(n):
             x = np.random.random()
12
                 sum += f(x) / p(x)
13
                 if plot:
14
                     if i % (int(n / 100000)) == 0 and i > 1000000:
1.5
                         estimates n = np.append(estimates n, i+1)
16
                         estimates area = np.append(estimates area, 4 * sum / (i+1))
        if plot:
18
19
         return sum / n * 4
20
```

Estimating PI: MC Integration Method - Convergence

mc_integrate(int(1e8), plot=True)



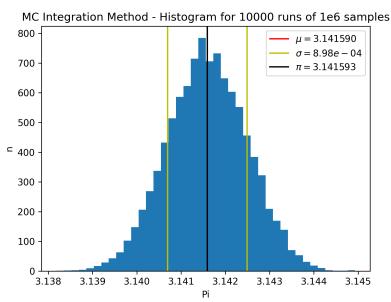
Estimating PI: MC Integration Method - We can do better

```
def mc_integrate_fast(n):
    x = np.random.random(n)
    return np.sum(f(x)) / n * 4
```

```
pis = np.zeros(10000)
for i in range(10000):
    pis[i] = mc_integrate_fast(int(1e6))

std = np.std(pis)
mean = np.mean(pis)
# Do the plotting here
# ...
print("MC Integration Method:")
print(f"{mean = }")
print(f"{std = }")
```

Estimating PI: Area Method - Histogram



Task 2: CDF Inversion

CDF Inversion - Task

Let $Q(\tau, \alpha) = \exp(-\alpha \tau)$. and $F(\tau) = \frac{1 - \exp(-\alpha \tau)}{\alpha}$ the CDF. Sample from Q by inverting:

$$r = \frac{\int_{\tau_{\min}}^{\tau} Q(\tau') d\tau'}{\int_{\tau_{\min}}^{\tau_{\max}} Q(\tau') d\tau'} = \frac{F(\tau) - F(\tau_{\min})}{F(\tau_{\max}) - F(\tau_{\min})}$$

with $r \sim U(0,1)$. Setting $\tau_{\min} = 0$ and $\tau_{\max} = 5$ and solving:

$$\tau = -\frac{\log(1 - \alpha \cdot r \cdot F(\tau_{\text{max}}))}{\alpha}$$

We can now

- Sample from Q and plot the histogram
- Calculate $I_1 = \int_0^5 \tau Q(\tau, \alpha) d\tau = \langle \tau \rangle \cdot \text{Norm} = \langle \tau \rangle \cdot F(\tau_{\text{max}})$
- Calculate $I_2 = \int_0^5 \tau^2 Q(\tau, \alpha) d\tau = \langle \tau^2 \rangle \cdot F(\tau_{\text{max}})$
- Calculate $\sigma(I_1)$ and $\sigma(I_2)$: $\sigma(\overline{\tau}) = \frac{\sigma_{\text{Samples}}}{\sqrt{n}} [\text{GKW16, p.47}]$

CDF Inversion - Implementation

```
alpha = 1
    x max = 5
    n = int(1e8)
3
    def F(x):
         return (1 - np.exp(-alpha*x))/alpha
 7
    F \max = F(x \max)
 8
9
    def F inv(r):
10
         return -np.log(1 - alpha*r*F_max)/alpha
11
12
13
     r = np.random.random(n)
14
     x = F inv(r)
15
16
17
     print("I1 exact: 1-6*exp(-5) = 0.9595723")
18
     print(f"I1 mean: { np.mean(x)*F_max:.6f}")
19
     print(f"I1 std: {np.std(x)*F max/np.sqrt(n):.3e}")
20
     print("I2 exact: 2-37*exp(-5) = 1.750696")
21
    print(f"I2 mean: {np.mean(x**2)*F_max:.6f}")
22
    print(f"I2 std: {np.std(x**2)*F_max/np.sqrt(n):.3e}")
23
```

CDF Inversion - Histogram

Console Output:

3

5

```
I1 exact: 0.9595723

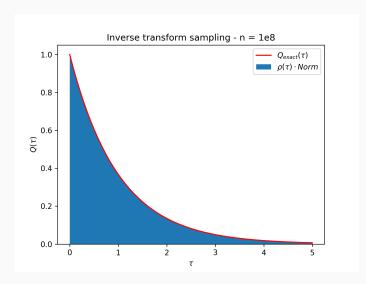
I1 mean: 0.959501

I1 std: 9.045e-05

I2 exact: 1.750696

I2 mean: 1.750502

I2 std: 3.205e-04
```



Task 3: Towards DMC - MCMC

Markov Chain Monte Carlo - Task

Basic principle behind DMC is Markov Chain Monte Carlo (MCMC). We again sample from a distribution *f*, this time using the Metropolis-Hastings algorithm:

- Propose a random configuration x' from a proposal distribution p
- Accept the configuration with probability $A(x, x') = \min \left(1, \frac{f(x')p(x)}{f(x)p(x')}\right)$
- p can be optimized to increase acceptance rate (similarly to CDF inversion, see also [FSW08, Chapter 12.3.1.1])
- This Markov Chain converges to desired distribution f, guarantees ergodicity
- Samples now correlated \rightarrow Blocking Analysis to estimate σ

Markov Chain Monte Carlo - Implementation

```
alpha = 1
     n equ = int(1e4)
     n sam = int(1e6)
     n acc = 0
     n rej = 0
     rng = np.random.default rng()
 7
     #function to sample from
     def f(x):
          return np.exp(-alpha*x)
10
11
12
     uni = rng.uniform
13
14
15
     def acc ratio(x, x new):
16
          return f(x \text{ new})/f(x)
17
18
     #initial sample
19
     curr = uni(0,5)
20
21
     #array to store samples
22
     samples = np.zeros(n_sam)
23
```

Markov Chain Monte Carlo - Implementation

```
for i in range(n_equ):
         prop = uni(0,5)
3
         if acc ratio(curr, prop) > np.random.random():
             curr = prop
             n acc += 1
             n rej += 1
9
10
     for i in range(n sam):
11
         prop = uni(0,5)
12
         if acc_ratio(curr, prop) > np.random.random():
13
             curr = prop
14
             n acc += 1
1.5
16
             n_rej += 1
         samples[i] = curr
18
19
20
21
```

Markov Chain Monte Carlo - Blocking Analysis

```
def blocking(self, samples, min blocks: int = 32) -> tuple[float, float]:
         means = np.copy(samples) # Copy the samples to avoid changing the original array
2
         mean = np.mean(means).astype(float) # Calculate the mean of the samples
3
         n = np.log2(len(means) // min blocks).astype(int) # Calculate the iteration steps
         block sizes = np.logspace(0, n, n + 1, base=2) # Minumum block size is 1, max is 2**n
         vars = np.zeros(n + 1, dtype=float) # Initialize the array for the variances
         vars[0] = 1 / (len(means) - 1) * (np.mean(means**2) - mean**2) # Naive variance
         Rx = np.zeros(n + 1) # Initialize Rx array
         Rx[0] = 1
         for i in range(1, n + 1): # Perform blocking
10
             if means.size % 2 != 0: # Make sure the number of blocks is even in order to divide by 2
11
                 means = means[:-1]
12
             means = 0.5 * (means[::2] + means[1::2]) # Double the block size
13
             n blocks = means.size
14
             varXBlock = n_blocks / (n_blocks - 1) * (np.mean(means**2) - mean**2)
15
             vars[i] = varXBlock / n blocks # Calculate the variance
16
             Rx[i] = block sizes[i] * varXBlock / vars[0] / samples.size
17
         plateau_index = np.argmin(np.abs(Rx[1:] - Rx[:-1])[3:]) + 4 # Find plateau in Rx
18
19
         return mean, vars[plateau index]
20
```

For the maths behind the blocking analysis, see [GKW16, Chapter 3.4].

Markov Chain Monte Carlo - Results

The results are very similar to the CDF inversion method, as we sample from the same distribution in a different way.

We will take a closer look at the blocking analysis after the next task.

Task 4&5: Diagrammatic Monte Carlo

Diagrammatic Monte Carlo - Task

We are interested in the following function:

$$Q(\tau,\alpha,V) = \exp(-\alpha\tau) + \sum_{\beta} \int_0^{\tau} d\tau_2 \int_0^{\tau_2} d\tau_1 e^{-\alpha\tau_1} V e^{-\beta(\tau_2-\tau_1)} V e^{-\alpha(\tau-\tau_2)}$$

which can be represented as zero and second order Feynman Diagrams of the following form:

$$\frac{1}{0} \frac{1}{\alpha} \frac{1}{\tau_j} = D_0^{\xi_0}(\tau_j, \alpha)$$

$$\frac{1}{0} \frac{1}{\alpha} \frac{1}{\tau_1} \frac{1}{\tau_2} \frac{1}{\alpha} \frac{1}{\tau_2} \frac{1}{\tau_2} \frac{1}{\sigma} \frac{1}{\tau_2} \frac{1}$$

How can we sample from this function?

Diagrammatic Monte Carlo - Task

We use the Metropolis-Hastings algorithm again, using the diagrams as weights. For ergodicity, we need to implement the following updates:

- \cdot Change of au
- · Increasing the order of the diagram
- · Decreasing the order of the diagram
- (Change of α)

Mind the context factors of the updates when changing the order of the diagram^[FSW08, ch. 12.3.1.2]:

$$A_{0\to 2} = \min \left\{ 1, \frac{p_{\text{rem}}}{p_{\text{add}}} \frac{D_2^{\xi_2} \left(\tau, \alpha, V; \tau_1, \tau_2, \beta\right) d\tau_1 d\tau_2}{D_0^{\xi_0} \left(\tau_0, \alpha_0\right) p_1 \left(\tau_1\right) p_2 \left(\tau_2\right) p_3(\beta) d\tau_1 d\tau_2} \right\}$$

We implement a Python class to sample, analyse and plot the results of the DMC algorithm. We will first take a look at some of the methods of this class, then have a look at the whole class **DMC.py** and finally look at the results.

Diagrammatic Monte Carlo Implementation - Initialization

```
def init (...) -> None:
       if tau max <= tau_min:</pre>
           raise ValueError("tau_max must be greater than tau_min")
       self.rng = np.random.default_rng() # set up random number generator
       self.uni = lambda a, b: self.rng.random() * (b - a) + a
       self.n_equ, n_sam, tau_min; tau_max, V, use_change_alpha = ... # Set up config parameters
       self.tau, alpha, beta, tau1, tau2 = ... # Set up state variables
       self.n tau acc, n tau rej, n alpha acc, n alpha rej... = 0 # Counters
10
       self.samples = np.zeros((n sam, 2) if use change alpha else n sam) # Set up sample array
1.1
12
       self.updaters = [self.change tau]
13
       if use change alpha:
           self.updaters.append(self.change alpha)
1.5
16
17
```

Diagrammatic Monte Carlo Implementation - Weight Calculation

```
def weight(self, tau=None, alpha=None, beta=None, tau1=None, tau2=None) -> float:
         if tau is None:
             tau = self.tau
         if beta == 0:
             return np.exp(-alpha * tau)
         return (
             np.exp(-alpha * tau1)
             * self.V
13
             * np.exp(-beta * (tau2 - tau1))
14
             * self.V
1.5
             * np.exp(-alpha * (tau - tau2))
16
17
```

As we will see, this method is used in every update to calculate the acceptance ratio.

Diagrammatic Monte Carlo Implementation - Change-au&lpha

```
def change tau(self) -> bool:
         prop_tau = self.uni(self.tau2, self.tau_max)
2
         M = self.weight(tau=prop tau) / self.weight()
         if M > self.rng.random():
             self.tau = prop tau
             self.n_tau_acc += 1
             return True
         else:
             self.n tau rej += 1
             return False
10
11
     def change alpha(self) -> bool:
12
         prop alpha = [0.5, 1][int(self.rng.random() * 2)]
13
         M = self.weight(alpha=prop_alpha) / self.weight()
14
         if M > self.rng.random():
15
             self.alpha = prop alpha
16
             self.n alpha acc += 1
17
             return True
18
19
             self.n alpha rej += 1
20
             return False
21
```

Diagrammatic Monte Carlo Implementation - Change-Order

```
def add beta(self) -> bool:
         if self.heta != 0:
             self.n beta rej += 1
3
             return False
         prop tau2 = self.uni(self.tau min, self.tau)
         prop tau1 = self.uni(self.tau min, prop tau2)
         prop_beta = [0.25, 0.75][int(self.rng.random() * 2)]
         M = self.weight(beta=prop beta, tau1=prop tau1, tau2=prop tau2) / (
             self.weight() / (self.tau - self.tau min) / (prop tau2 - self.tau min) * 0.5
10
         if M > self.rng.random():
11
             self.beta = prop beta
12
             self.tau1 = prop tau1
13
             self.tau2 = prop tau2
14
             self.n_beta_acc += 1
1.5
             return True
         else:
17
             self.n beta rei += 1
18
             return False
19
```

The remove- β update works the same way. The Metropolis ratio is the inverse of the add- β ratio.

Diagrammatic Monte Carlo Implementation - Updating, Sampling

```
def update(self) -> bool:
         updater = self.updaters[int(self.rng.random() * len(self.updaters))]
2
         return updater()
4
    def sample(self) -> None:
         for i in trange(self.n sam):
             self.update()
             if self.use_change_alpha:
                 self.samples[i, 0] = self.tau
                 self.samples[i, 1] = self.alpha
                 if self.beta == 0:
                      self.n_zero_order[self.alpha] = self.n_zero_order.get(self.alpha, 0) + 1
12
13
                 self.samples[i] = self.tau
14
                 if self.beta == 0:
15
                     self.n zero order += 1
16
17
```

Equilibration works the same way, but without sampling.

Diagrammatic Monte Carlo Implementation - Normalization

The normalization constant could still be calculated analytically, but normally this is not possible.

Solution: Use normalization constant C_0 of zero order diagram and number of zero order diagrams N_0 to estimate $C^{[RFT20, ch. 3.3]}$:

$$C \approx C_0 \frac{N}{N_0}$$

This is implemented as an alternative to the analytical normalization. Furthermore, helper methods for plotting and analysis are implemented.

Let's take a look at the whole class DMC.py, before we come back to the results.

Switching to VS Code

Diagrammatic Monte Carlo - Results

```
from DMC import DMC
     dmc = DMC(
         n = qu = int(1e4),
         n sam=int(5e6),
4
         tau min=0.
         tau \max=5.
         alpha 0=1,
         tau 0=1,
         use change alpha=True,
         use change beta=False.
10
         use_analytical=True,
12
     dmc.equilibrate()
     dmc.sample()
14
     dmc.print I1(0.5)
15
     dmc.print I1(1)
16
     dmc.print I2(0.5)
17
     dmc.print I2(1)
18
     dmc.plot hist(1)
19
    dmc.plot_hist(0.5)
20
```

```
Warming up: 10000 steps Warmup complete: \tau acc. ratio of 0.54, \alpha acc. ratio of 0.86 Sampling: 5000000 steps 100% | 5000000 steps | 5000000/5000000 [00:25<00:00, 194393.91it/s] Sampling complete: \tau acc. ratio of 0.54, \alpha acc. ratio of 0.85 I1(\alpha=0.5) mean: 2.85430, var: 9.53e-06 I1(\alpha=1) mean: 0.96035, var: 2.40e-06 I2(\alpha=0.5) mean: 7.31758, var: 1.63e-04 I2(\alpha=1) mean: 1.75439, var: 2.34e-05
```

Plots are shown on the next slide.

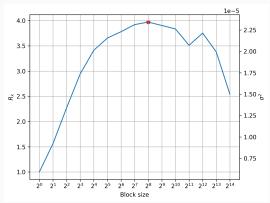
3

4

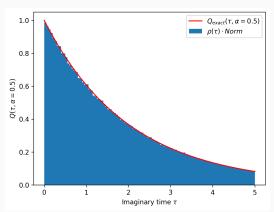
6

9

Diagrammatic Monte Carlo - Results



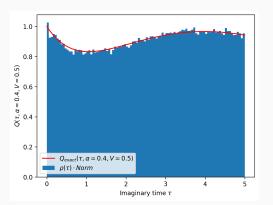
One of the blocking analysis plots



Histogram for $\alpha = 0.5$

Diagrammatic Monte Carlo - Results

```
DMC import DMC
dmc = DMC(
    n = qu = int(1e4),
    n sam=int(3e6),
    tau min=0,
    tau \max=5.
    alpha 0=0.4.
    tau 0=1,
    V=0.5.
    use change alpha=False,
    use change_beta=True,
    use analytical=True.
dmc.equilibrate()
dmc.sample()
dmc.plot_hist()
```



3

4

5

12

13

14

16

17

2

3

4

5

Task 6: DMC - Green Function
Estimator

Diagrammatic Monte Carlo - Greens Function Estimator

An exact MC estimator for the Greens function can be derived [MPSS00, ch. III.C]:

$$\langle q_{\tau_0} \rangle_{MC} = C * Q(\tau_0, \alpha, V)$$

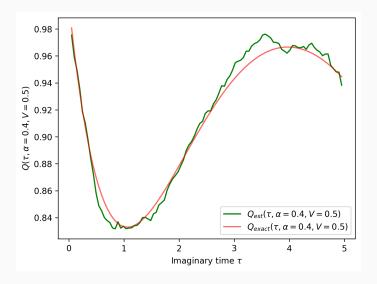
$$q_{\tau_0}(\nu, \tau) = \begin{cases} q(\nu) \mathcal{D}_{\nu}(\tau_0) / \mathcal{D}_{\nu}(\tau) = |\Gamma_0|^{-1} e^{-\alpha(\tau_0 - \tau)}, & \text{if } \tau \in \Gamma_0 \text{ and } \mathcal{D}_{\nu}(\tau) \neq 0 \\ 0, & \text{otherwise.} \end{cases}$$

Where i choose $\Gamma_0 = [\tau_0 - a, \tau_0 + a]$ with $a = \min(0.2, \tau_0, \tau_{\max} - \tau_0)$ This eliminates the discretization error of the histogram method.

```
def green_estimator(self, tau0: float, alpha: float | None = None) -> float:
    if alpha is None:
        alpha = self.alpha
    tau = self.get_samples(alpha)
    diff = np.subtract.outer(tau0, tau)
    tau0_stacked = np.repeat(tau0, tau.size).reshape(diff.shape)
    mins = np.minimum(np.minimum(tau0_stacked, self.tau_max - tau0_stacked), 0.2)
    res = np.exp(-alpha * (diff)) / (2 * mins)
    res[np.abs(diff) > mins] = 0
    meaned = np.mean(res, axis=1) *self.norm(alpha, analytical=self.use_analytical)
    return meaned
```

Diagrammatic Monte Carlo - Greens Function Estimator - Results

Plot of the Greens function estimator for $n = 3 \cdot 10^6$ samples:



References i

[Ces22] Cesare Franchini.

Basic diagrammatic monte carlo - exercise sheet.

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 Diagrammatic quantum monte carlo study of the fröhlich polaron.

 Physical Review B, 62(10):6317–6336, September 2000.

References ii

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 Diagrammatic monte carlo study of the holstein polaron.
 Master's thesis, Alma Mater Studiorum - University of Bologna, 2020.
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