

MACHINE LEARNING

LINEAR REGRESSION

AGENDA

O1 Supervised Learning Training data, hypotheses, an example.

O2 LMS Algorithm
Linear model, cost function, gradient descent.

O3 Normal Equations

Matrix linear model, deriving the normal equations

O4 Probabilistic Interpretation
Assumptions for errors, likelihood function

O5 Base Function

Derivation, functions, maximum likelihood





SUPERVISED LEARNING DEFINITIONS AND NOMENCLATURE



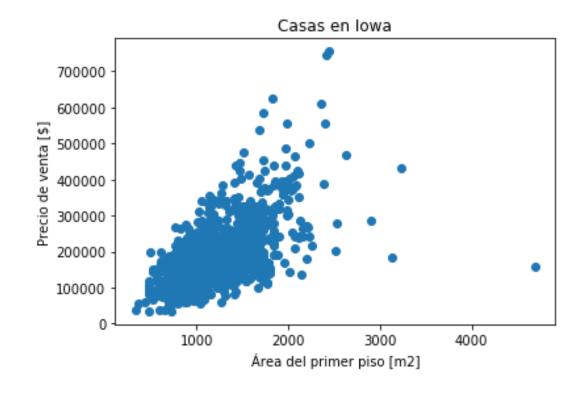
We start with a data set that represents first-floor living space and sales prices for 1,460 homes in Ames, Iowa.

1^{st} floor square feet	Sale Price
856	208500
1262	181500
920	223500
961	140000
1145	250000
:	:

SUPERVISED LEARNING DEFINITIONS AND NOMENCLATURE



We graph the 1,460 elements using Python:



SUPERVISED LEARNING DEFINITIONS AND NOMENCLATURE



The question that naturally arises would be:

HOW DO WE PREDICT THE PRICES FOR OTHER HOMES IN AMES, IOWA?

SUPERVISED LEARNING DEFINITIONS AND NOMENCLATURE



To answer the question, we must first define the nomenclature that will be used:

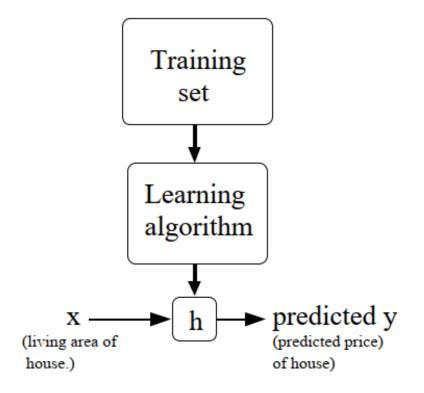
Variable / Symbol	Description	Example
$x^{(i)}$	Input variable or characteristics	The house's first floor area
$y^{(i)}$	Output, target, or response variable	The house's sale price
$(\boldsymbol{x^{(i)}}, \boldsymbol{y^{(i)}})$	Training example	(area, price)
m	Number of training examples	1460 houses with their first floor área and their price.
$\{(x^{(i)}, y^{(i)}); i = 1,, m\}$	Training data set	NA
χ	Input value space	NA
γ	Output value space	

SUPERVISED LEARNING DEFINITIONS AND NOMENCLATURE



The main goal of any supervised algorithm is to learn a function $h: \chi \to \Upsilon$, such that h(x) can predict the corresponding value of y.

Where h is defined as the **hypothesis**.



SUPERVISED LEARNING DEFINITIONS AND NOMENCLATURE



There are **2 types** of **supervised learning problems**:

Type of problem	Description
Regression problem	y takes continuous values.
Classification problem	y takes discrete values



L M S A L G O R I T H M L I N E A R M O D E L

Now let's look at the same model but with another variable: the living space on the second floor.

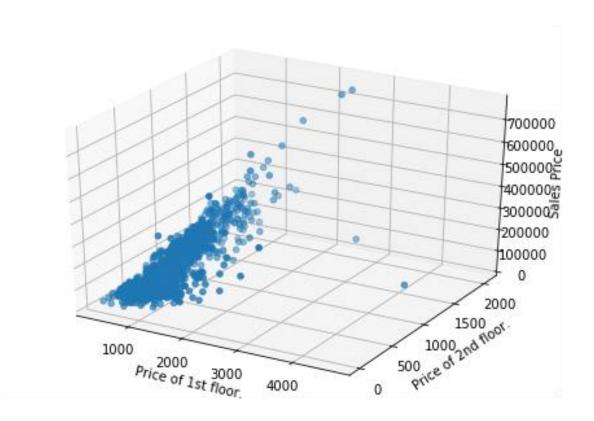
1 st floor square feet	2^{nd} floor square feet	Sale Price
856	854	208500
1262	0	181500
920	866	223500
961	756	140000
1145	1053	250000
:	:	:

In this case each $x^{(i)} \in \mathbb{R}^2$, so it is defined as a vector $x^{(i)} = [x^{(i)}_1 \ x^{(i)}_2]$.

L M S A L G O R I T H M L I N E A R M O D E L



We graph the 1,460 elements again using Python:



It is **hypothesized** that the data is distributed in a **linear fashion**, therefore:

$$h_w(x) = w_0 + w_1 x_1 + w_2 x_2$$

In this case the variables w_i are defined as the parameters or weights that parameterize the space of all linear functions that map from χ to Υ .

The expression is **simplified** into **vector form**, where n represents the **number** of **input** variables (omitting x_0),:

$$h(x) = \sum_{i=1}^{n} w_i x_i = w^T x$$



Then, the question would be:

WHAT IS THE BEST COMBINATION OF PARAMETERS w THAT RESULTS IN THE BEST HYPOTHESIS h?

L M S A L G O R I T H M C O S T F U N C T I O N

To **answer** the **question**, we need to **measure** the **error** produced by the **estimates** produced by h.

We know that we can **measure** the **error** of a single **estimate** i as a **difference between** the estimate $h_w(x^{(i)})$ and the response variable $y^{(i)}$.

$$e = h_w(x^{(i)}) - y^{(i)}$$

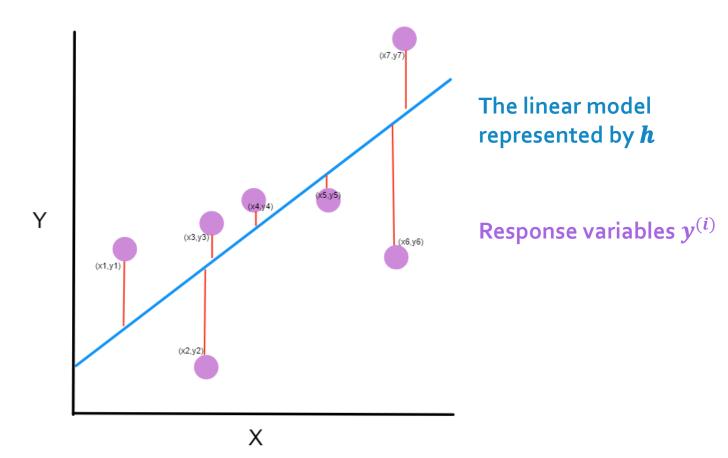
To ensure positive amounts of error, the answer is squared:

$$e^2 = (h_w(x^{(i)}) - y^{(i)})^2$$

L M S A L G O R I T H M
C O S T F U N C T I O N



The **error** that we are **calculating graphically** is **interpreted** with an **example**:



So far, we have calculated the **squared error** of a **single training data**. The **mean square error** *MSE* is calculated **for all training data**.

$$MSE = \frac{1}{2m} \sum_{i=1}^{m} (h_w(x^{(i)}) - y^{(i)})^2$$

$$MSE = J(w)$$

Where J(w) is the **cost function**.



Therefore, the **best combination** of weights w is the one that **minimizes** the **cost function** J(w), which **measures** our **mean square error**.

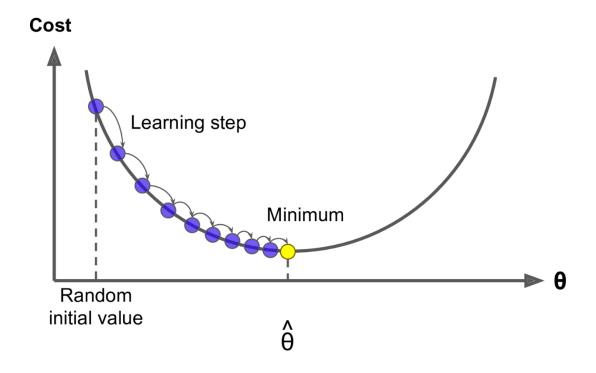
To find the best combination, let's design a search algorithm that starts with a random initial value of w, and updates the values of w until it converges to a minimum value of J(w). The constant α is defined as the learning rate.

$$w_j \coloneqq w_j - \alpha \frac{\partial}{\partial w_j} J(w)$$

NOTE: the above equation **only updates** the **value** of a **single weight** w_j of the n weights that **parameterize** the **linear model**. In reality, **all weights** w_j are **updated simultaneously**.



In the gradient descent optimization algorithm, we update the weights in the direction with the greatest decrement of J(w).





The derivative of the cost function J(w) with respect to a specific weight w_i is calculated.

Derive the result:

$$\frac{\partial}{\partial w_i}J(w) = \frac{1}{m}(h(x) - y)x_j$$



Therefore, the **gradient descent weights update** for a **single training data** would look like this:

$$w_j \coloneqq w_j + \frac{\alpha}{m} (y^{(i)} - h(x^{(i)})) x_j$$

This **equation** is called the **LMS update rule** ("Least Mean Squares") or also the Widrow-Hoff learning rule.

This **method** is also called **batch gradient descent**, where **all training data** is **analyzed simultaneously**.



For m training data, and defining $x^{(i)}$ and w as vectors, the learning rule would look like this:

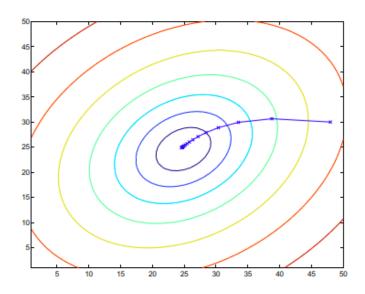
$$w \coloneqq w + \frac{\alpha}{m} \sum_{i=1}^{m} (y^{(i)} - h(x^{(i)})) x^{(i)}$$

This **equation** is called the **LMS update rule** ("Least Mean Squares") or also the Widrow-Hoff learning rule.



In general, the **gradient descent method** can suffer from **several local minima**. In this case, for **linear regression models** there is only a **single global minimum**.

Consequently, the **algorithm will always converge** assuming that the **learning rate is not too high.**



When the **method** only **observes one training data** at a time, and **updates** the **weights with a single data**, it is called **stochastic or incremental gradient descent**.

$$w \coloneqq w + \alpha(y^{(i)} - h(x^{(i)}))x^{(i)}$$

This variant of the algorithm is used when it is very expensive to evaluate the update for very large data sets (when m is very large), but it has the minimal divergence problem.

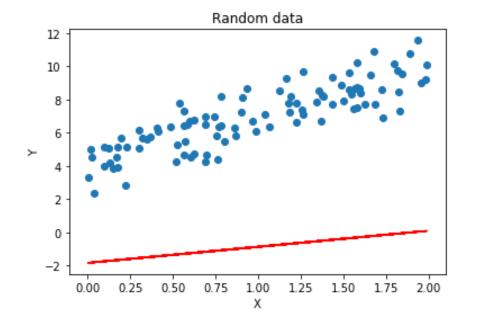


Example:

Random training data was generated with a certain degree of error *e*:

$$y = 3x + 4 + e$$

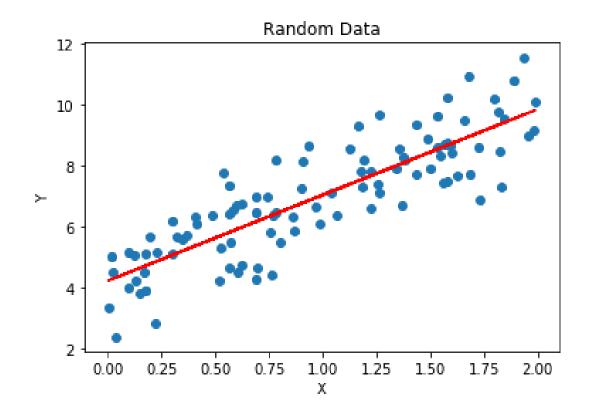
and both weights were randomly initialized: w_0 and w_1





Example:

After 1000 iterations the following model was obtained:



$$w_0 = 4.2174$$

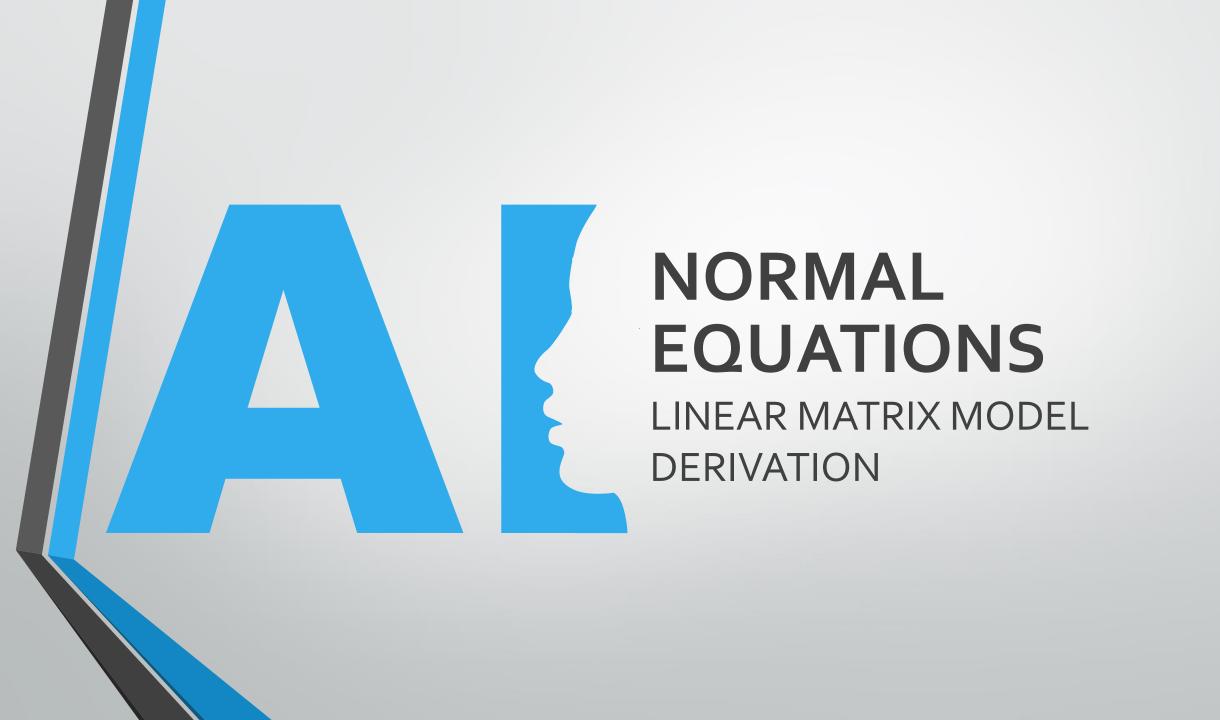
 $w_1 = 2.816$



Example:

Error based on 100 iterations.





NORMAL EQUATIONS MATRIX NOTATION



Now we want to calculate the minimum of the cost function analytically. For this, matrix notation is introduced, which allows the derivatives of the cost function to be expressed in an elegant way without getting into so much verbiage.

The m training data $\{(x^{(i)}, y^{(i)}); i = 1, ..., m\}$ is represented as a matrix $X \in \mathbb{R}^{m \times n}$, the set of **outputs** as a **vector** $\vec{y} \in \mathbb{R}^m$ and the n **weights** w_i as a vector $\vec{w} \in \mathbb{R}^n$.

$$x^{(i)} = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix}$$

$$x^{(i)} = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix} \qquad X = \begin{bmatrix} -(x^{(1)})^T - \\ -(x^{(2)})^T - \\ \vdots \\ -(x^{(m)})^T - \end{bmatrix} \qquad \vec{y} = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(m)} \end{bmatrix} \qquad \vec{w} = \begin{bmatrix} w_1 \\ w_2 \\ \vdots \\ w_n \end{bmatrix}$$

$$\vec{y} = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(m)} \end{bmatrix}$$

$$\overrightarrow{w} = \begin{bmatrix} w_1 \\ w_2 \\ \vdots \\ w_n \end{bmatrix}$$

NORMAL EQUATIONS MATRIX NOTATION



The absolute error function is represented in matrix notation where $h_w(x^{(i)}) = (x^{(i)})^T \vec{w}$:

$$X\overrightarrow{w} - \overrightarrow{y} = \begin{bmatrix} -(x^{(1)})^T \overrightarrow{w} - \\ -(x^{(2)})^T \overrightarrow{w} - \\ \vdots \\ -(x^{(m)})^T \overrightarrow{w} - \end{bmatrix} - \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(m)} \end{bmatrix}$$

$$X\overrightarrow{w} - \overrightarrow{y} = \begin{bmatrix} h_w(x^{(1)}) - y^{(1)} \\ h_w(x^{(2)}) - y^{(2)} \\ \vdots \\ h_w(x^{(m)}) - y^{(m)} \end{bmatrix}$$

NORMAL EQUATIONS MATRIX NOTATION



The **mean square error term** (cost function):

$$J(w) = \frac{1}{2m} \sum_{i=1}^{m} (h_w(x^{(i)}) - y^{(i)})^2$$

can be **represented** in **vector** form using the **property** $z^Tz = \sum_i z_i^2$

$$J(w) = \frac{1}{2m} (X\overrightarrow{w} - \overrightarrow{y})^T (X\overrightarrow{w} - \overrightarrow{y})$$

NORMAL EQUATIONS DERIVATION OF EQUATIONS



We have to **obtain** the **derivatives** of J(w) with respect to the vector w.

$$\nabla_{w}J(w) = \nabla_{w}\frac{1}{2m}(X\overrightarrow{w} - \overrightarrow{y})^{T}(X\overrightarrow{w} - \overrightarrow{y})$$

The gradient calculates all the derivatives with respect to all the weights w_j simultaneously.

Also, because we **represent** the **training dataset** in **matrix form**, we can **calculate** the **gradient** for **all training data**.

NORMAL EQUATIONS DERIVATION OF EQUATIONS



Therefore, we obtain:

$$\nabla_{w}J(w) = \nabla_{w}\frac{1}{2m}(X\overrightarrow{w} - \overrightarrow{y})^{T}(X\overrightarrow{w} - \overrightarrow{y})$$

DERIVE THE EQUALITY

$$\nabla_{w}J(w) = \frac{1}{m} (X^{T}X\overrightarrow{w} - X^{T}\overrightarrow{y})$$

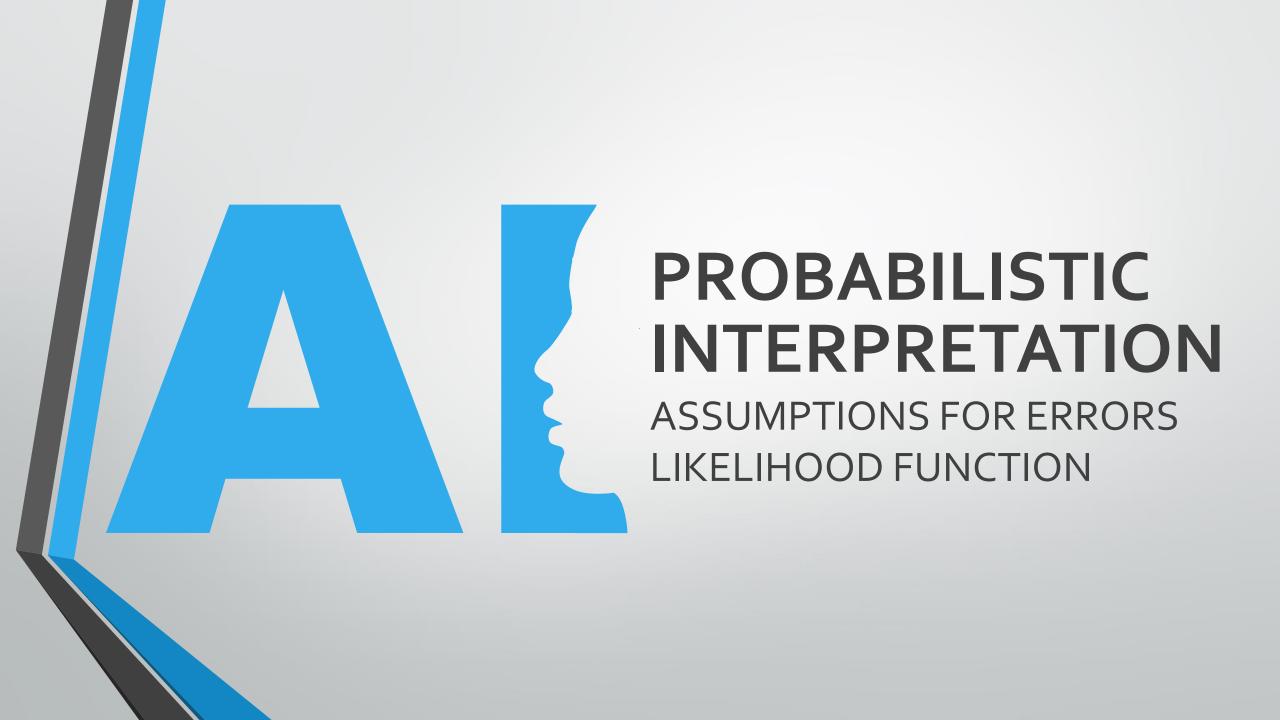
NORMAL EQUATIONS DERIVATION OF EQUATIONS



Equating the **derivative to zero** to find the **minimum**, we obtain the **vector of weights** that gives this **minimum**:

$$\nabla_{w}J(w) = \frac{1}{m} (X^{T}X\overrightarrow{w} - X^{T}\overrightarrow{y}) = 0$$

$$\overrightarrow{w} = \left(X^T X\right)^{-1} X^T \overrightarrow{y}$$



PROBABILISTIC INTERPRETATION DEVELOPMENT MOTIVATION



WHY DO WE MINIMIZE THE MEAN QUADRATIC ERROR FUNCTION AND NOT ANOTHER FUNCTION?



The output variables $y^{(i)}$ can be defined as the hypothesis h_w plus an estimation error $\varepsilon^{(i)}$ that captures the effects that we did not consider in our hypothesis or contemplates the random errors.

$$y^{(i)} = h_w(x^{(i)}) + \varepsilon^{(i)}$$



The following assumptions are made for the errors $\varepsilon^{(i)}$ (random sampling): IID

Independent Errors - The probability of one error does not affect the probability of other errors.

Identically distributed errors: errors are sampled from the same probability distribution.

Distributed by a Gaussian probability distribution with mean $\mu=0$ and variance σ^2 .

 $\varepsilon^{(i)} \sim N(0, \sigma^2)$



Therefore, the **probability density function** of $\varepsilon^{(i)}$ is given by:

$$p(arepsilon^{(i)}) = rac{1}{\sqrt{2\pi}\sigma}e^{\left(-rac{\left(arepsilon^{(i)}
ight)^2}{2\sigma^2}
ight)}$$



WHY PROPOSE A MODEL THAT IS NORMALLY DISTRIBUTED?



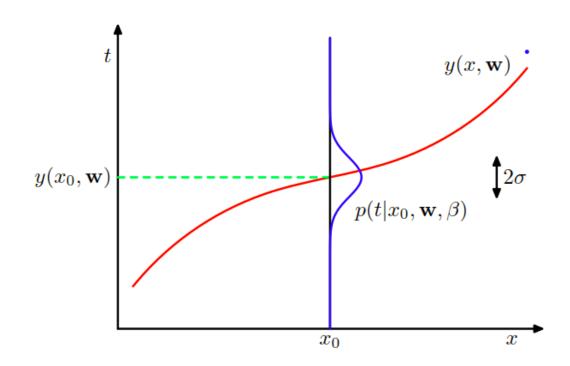
Furthermore, the **outputs** $y^{(i)}$ and the **inputs** $x^{(i)}$ are set as **random variables**.

Consequently, the following question is posed:

Given my input data set X and my weight vector \overrightarrow{w} , what is the probability distribution that models my outputs \overrightarrow{y} ?



Following the assumptions from the normal distribution of errors, it naturally occurs that the distribution of the outputs $y^{(i)}$ follows a normal form with mean $h_w(x^{(i)})$ and variance σ^2 .





Formally it would look like this:

$$p(y^{(i)}/x^{(i)};w) = \frac{1}{\sqrt{2\pi}\sigma}e^{\left(-\frac{\left(y^{(i)}-h_w(x^{(i)})\right)^2}{2\sigma^2}\right)}$$

$$p(y^{(i)}/x^{(i)};w) = \frac{1}{\sqrt{2\pi}\sigma}e^{\left(-\frac{\left(y^{(i)}-w^Tx^{(i)}\right)^2}{2\sigma^2}\right)}$$



We define the **conditional probability** for **all training data** (assuming the **samples** were **collected independently**):

$$p(\vec{y}/X; w) = \prod_{i=1}^{m} p(y^{(i)}/x^{(i)}; w) = \prod_{i=1}^{m} \frac{1}{\sqrt{2\pi}\sigma} e^{\left(-\frac{(y^{(i)}-w^{T}x^{(i)})^{2}}{2\sigma^{2}}\right)}$$

INTERPRETACIÓN PROBABILÍSTICA LIKELIHOOD FUNCTION



This equation is called the **likelihood function**, because it describes the **probability** that our **beliefs** about the **reality** of the **observations** (data) are **true**. In other words, that the **hypothesis** we proposed is **correct**.

$$p(\overrightarrow{y}/X; \overrightarrow{w}) = \prod_{i=1}^{m} p(y^{(i)}/x^{(i)}; \overrightarrow{w}) = \prod_{i=1}^{m} \frac{1}{\sqrt{2\pi}\sigma} e^{\left(-\frac{\left(y^{(i)}-w^{T}x^{(i)}\right)^{2}}{2\sigma^{2}}\right)}$$

$$p(\mathbf{w}|\mathcal{D}) = \underbrace{\frac{p(\mathcal{D}|\mathbf{w})p(\mathbf{w})}{p(\mathcal{D})}}$$

INTERPRETACIÓN PROBABILÍSTICA LIKELIHOOD FUNCTION



Therefore, we want to **maximize** the **probability** that **our beliefs** about how the **data** is **distributed** (in a **normal** way).

That is, we want to **maximize** the **likelihood function**. From a **frequentist** perspective (the value of \vec{w} is **not random**), we want to **find** the **value** of \vec{w} that **maximizes** the **probability** that the **observations** made of **reality** are **true**.

$$\arg\max_{\overrightarrow{w}} L(\overrightarrow{w}; X, \overrightarrow{y}) = \arg\max_{\overrightarrow{w}} \prod_{i=1}^{m} \frac{1}{\sqrt{2\pi}\sigma} e^{\left(-\frac{\left(y^{(i)} - w^T x^{(i)}\right)^2}{2\sigma^2}\right)}$$

INTERPRETACIÓN PROBABILÍSTICA LOGARITHMIC LIKELIHOOD



Maximizing a summation is much easier than doing so with a multiplication

$$\arg\max_{\overrightarrow{w}} log(L(\overrightarrow{w})) =_{\arg\max_{\overrightarrow{w}}} log(\prod_{i=1}^{m} p(y^{(i)}/x^{(i)}; \overrightarrow{w}))$$

$$\underset{\overrightarrow{w}}{\operatorname{arg\,max}} \ \log(L(\overrightarrow{w})) =_{\underset{\overrightarrow{w}}{\operatorname{arg\,max}}} \sum_{i=1}^{m} \log(p(y^{(i)}/x^{(i)}; \overrightarrow{w}))$$

INTERPRETACIÓN PROBABILÍSTICA

LOGARITHMICLOSS



The maximization problem is converted into a minimization one by scaling the function with a minus sign.

This function $-log(L(\overrightarrow{w}))$ is called **logarithmic loss.**

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) = \underset{\overrightarrow{w}}{\operatorname{arg\,min}} - \sum_{i=1}^{m} \log(p(y^{(i)}/x^{(i)}; \overrightarrow{w}))$$

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) = \underset{\overrightarrow{w}}{\operatorname{arg\,min}} - \sum_{i=1}^{m} \log(\frac{1}{\sqrt{2\pi}\sigma} e^{\left(-\frac{\left(y^{(i)} - w^{T}x^{(i)}\right)^{2}}{2\sigma^{2}}\right)})$$

INTERPRETACIÓN PROBABILÍSTICA (

LOGARITHMIC LOSS



We develop the equation:

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) = \underset{\overrightarrow{w}}{\operatorname{arg\,min}} - \sum_{i=1}^{m} log\left(\frac{1}{\sqrt{2\pi}\sigma}\right) + log(e^{\left(-\frac{\left(y^{(i)} - w^{T}x^{(i)}\right)^{2}}{2\sigma^{2}}\right)}\right)$$

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) =_{\underset{\overrightarrow{w}}{\operatorname{arg\,min}}} - m \log\left(\frac{1}{\sqrt{2\pi}\sigma}\right) - \sum_{i=1}^{m} -\frac{\left(y^{(i)} - w^{T}x^{(i)}\right)^{2}}{2\sigma^{2}}$$

INTERPRETACIÓN PROBABILÍSTICA





We develop the equation:

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) =_{\underset{\overrightarrow{w}}{\operatorname{arg\,min}}} c + \sum_{i=1}^{m} \frac{\left(y^{(i)} - w^{T}x^{(i)}\right)^{2}}{2\sigma^{2}}$$

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) =_{\underset{\overrightarrow{w}}{\operatorname{arg\,min}}} \frac{1}{2\sigma^2} \sum_{i=1}^{m} (y^{(i)} - w^T x^{(i)})^2$$

INTERPRETACIÓN PROBABILÍSTICA MEAN SQUARED ERROR



Minimizing the logarithmic loss is the same as minimizing the mean square error $MSE = J(\overrightarrow{w})$:

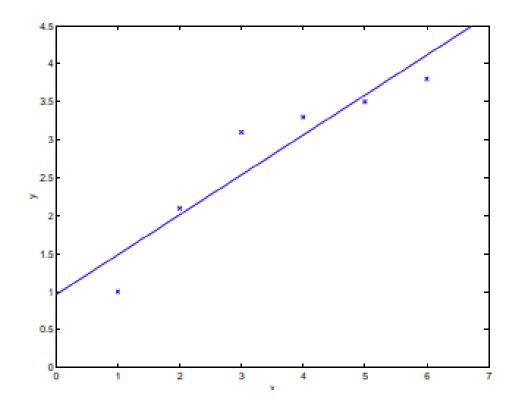
$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} \, MSE = \underset{\overrightarrow{w}}{\operatorname{arg\,min}} \frac{1}{2m} \sum_{i=1}^{m} (y^{(i)} - w^{T} x^{(i)})^{2}$$

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) =_{\underset{\overrightarrow{w}}{\operatorname{arg\,min}}} \frac{1}{2\sigma^2} \sum_{i=1}^{m} (y^{(i)} - w^T x^{(i)})^2$$



B A S I S F U N C T I O N S L I N E A R I T Y P R O B L E M

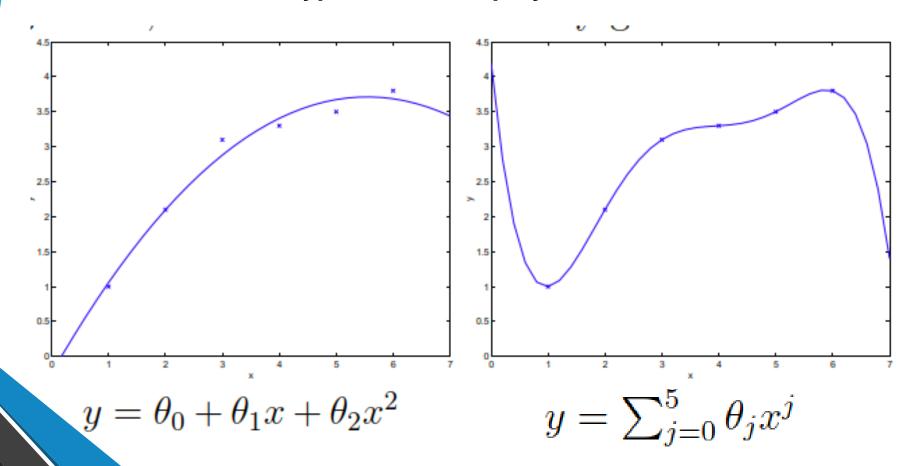
The **real world** has **non-linear** behaviors, so the **linear regression models** seen so far have their **limitations**.



BASIS FUNCTIONS LINEARITY PROBLEM



We can **establish** our **hypothesis** as a **polynomial model**.



BASIS FUNCTIONS



AN INTRODUCTION TO BASIS FUNCTIONS

If we take this kind of reasoning further, we can **build** models as **linear** combinations of nonlinear functions $\phi_j(x)$, called **base** functions. Where $\phi: \mathbb{R}^n \to \mathbb{R}^k$

$$h_w(x^{(i)}) = \sum_{j=1}^k w_j \phi_j(x^{(i)})$$

$$h_w(x^{(i)}) = w^T \phi(x^{(i)})$$

NOTA $w \in \mathbb{R}^k$

$$x^{(i)} \in \mathbb{R}^n$$

Thus we can **build** a **non-linear model**, which is still **parametrized** by **linear weights** *w*.

B A S I S F U N C T I O N S

NS

AN INTRODUCTION TO BASIS FUNCTIONS

In a more detailed fashion, we have that:

$$\boldsymbol{\phi}(x^{(i)}) = \begin{bmatrix} \boldsymbol{\phi}_1(x^{(i)}) \\ \boldsymbol{\phi}_j(x^{(i)}) \\ \vdots \\ \boldsymbol{\phi}_k(x^{(i)}) \end{bmatrix}$$

In the case of classical linear regression we have for a single training data $\phi(x) \in \mathbb{R}^k$ in this case k = n:

$$\phi_j(x^{(i)}) = x^{(i)}$$
 $x^{(i)} = \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix}$ $w = \begin{bmatrix} w_1 \\ \vdots \\ w_k \end{bmatrix}$

In the case of classical linear regression we have for a single training data $\phi(x) \in \mathbb{R}^k$ in this case k = n:

$$\phi(x) = \begin{bmatrix} \phi_0(x^{(1)}) \\ \phi_1(x^{(i)}) \end{bmatrix} = \begin{bmatrix} 1 \\ x^{(i)} \end{bmatrix} = \begin{bmatrix} 1 \\ x_1 \\ \vdots \\ x_n \end{bmatrix}$$

$$h_w(x^{(i)}) = w^T \phi(x^{(i)}) = w_0 + w_1(x_1) + \dots + w_k(x_n)$$



In the case of classical linear regression, we have for m training data:

$$\phi(x) = \begin{bmatrix} \phi_0(x^{(1)}) & \phi_1(x^{(1)}) \\ \vdots & \vdots \\ \phi_0(x^{(m)}) & \phi_1(x^{(m)}) \end{bmatrix} = \begin{bmatrix} 1 & x^{(1)}^T \\ \vdots & \vdots \\ 1 & x^{(m)}^T \end{bmatrix}$$

$$w = [w_0 \quad \dots \quad w_k]$$



In the case of classical linear regression, we have for m training data:

$$h_{w}(x) = w^{T} \phi(x) = \begin{bmatrix} w_{0} + \dots + w_{k} x_{n}^{(1)} \\ \vdots \\ w_{0} + \dots + w_{k} x_{n}^{(m)} \end{bmatrix}$$

$$h_w(x) = \begin{bmatrix} h_w(x^{(1)}) \\ \vdots \\ h_w(x^{(m)}) \end{bmatrix}$$

B A S I S F U N C T I O N S D E S I G N M A T R I X

Therefore, the **design matrix** for any $\phi(x)$ would be given by:

$$\phi(x) = \begin{bmatrix} \phi_0(x^{(1)}) & \cdots & \phi_{k-1}(x^{(1)}) \\ \vdots & \ddots & \vdots \\ \phi_0(x^{(m)}) & \cdots & \phi_{k-1}(x^{(m)}) \end{bmatrix}$$

$$x^{(i)} \in \mathbb{R}^n$$

Where each row of the matrix $\phi(x)$ is given by $\phi_i = \phi(x^{(i)})^T$

BASIS FUNCTIONS LINEAR BASIS FUNCTIONS



Classic linear functions:

$$\phi_j(x^{(i)}) = x^{(i)}; x^{(i)} \in \mathbb{R}^{n+1}$$

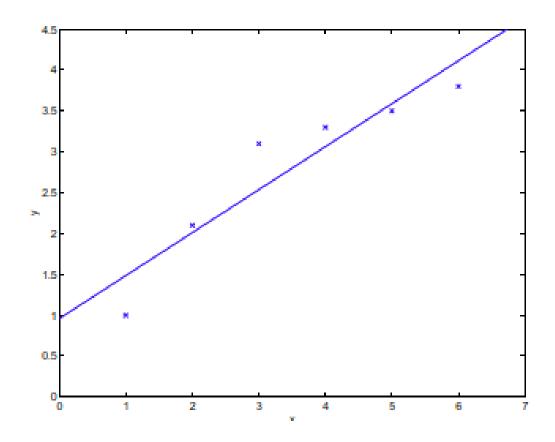
$$\phi(x) = \begin{bmatrix} \mathbf{1} & x^{(1)}^T \\ \vdots & \vdots \\ \mathbf{1} & x^{(m)}^T \end{bmatrix} = \begin{bmatrix} \mathbf{1} & \cdots & x_n^{(1)} \\ \vdots & \ddots & \vdots \\ \mathbf{1} & \cdots & x_n^{(m)} \end{bmatrix}$$

$$w = \begin{bmatrix} w_0 \\ \vdots \\ w_k \end{bmatrix}; k = n$$

$$h_{w}(x) = w^{T} \phi(x) = \begin{bmatrix} w_{0} + w_{1} x_{1}^{(1)} + \dots + w_{k} x_{n}^{(1)} \\ \vdots \\ w_{0} + w_{1} x_{1}^{(m)} + \dots + w_{k} x_{n}^{(m)} \end{bmatrix}$$

BASIS FUNCTIONS LINEAR BASIS FUNCTIONS

Classic linear functions:



BASIS FUNCTIONS



POLYNOMIAL BASIS FUNCTIONS

Polynomial basis functions:

$$\phi_j(x) = (x^{(i)})^j; x^{(i)} \in \mathbb{R}^n$$

$$\phi_{j}(x) = (x^{(i)})^{j}; x^{(i)} \in \mathbb{R}^{n}$$

$$\phi(x) = \begin{bmatrix} 1 & x^{(1)^{T}} & \cdots & (x^{(1)^{T}})^{(k-1)} \\ \vdots & \vdots & & \vdots \\ 1 & x^{(m)^{T}} & \cdots & (x^{(m)^{T}})^{(k-1)} \end{bmatrix}$$

$$w = \begin{bmatrix} w_{0} \\ \vdots \\ w_{k*n+1} \end{bmatrix}$$

$$\begin{bmatrix} w_{0} + w_{1}x_{1}^{(1)} + \cdots + w_{k*n}(x_{n}^{(1)})^{k-1} \end{bmatrix}$$

$$w = \begin{bmatrix} w_0 \\ \vdots \\ w_{k*n+1} \end{bmatrix}$$

$$h_{w}(x) = w^{T} \phi(x) = \begin{bmatrix} w_{0} + w_{1}x_{1}^{(1)} + \dots + w_{k*n}(x_{n}^{(1)})^{k-1} \\ \vdots \\ w_{0} + w_{1}x_{1}^{(m)} + \dots + w_{k*n}(x_{n}^{(m)})^{k-1} \end{bmatrix}$$

B A S I S F U N C T I O N S

POLYNOMIAL BASIS FUNCTIONS

Polynomial basis functions: concrete example polynomial degree 2 and two training data m=2

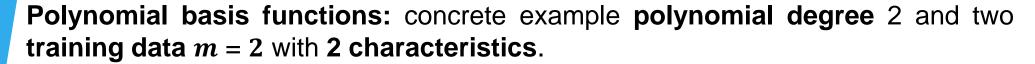
$$x^{(i)} = \begin{bmatrix} 1 \\ x_1 \\ x_2 \end{bmatrix}$$

we have example polynomial degree 2 and
$$w=\begin{bmatrix} w_0\\w_1\\w_2\\w_3\\w_4 \end{bmatrix}$$

$$\phi(x) = \begin{bmatrix} 1 & x_1^{(1)} & x_2^{(1)} & (x_1^{(1)})^2 & (x_2^{(1)})^2 \\ 1 & x_1^{(2)} & x_2^{(2)} & (x_1^{(2)})^2 & (x_2^{(2)})^2 \end{bmatrix}$$

BASIS FUNCTIONS

POLYNOMIAL BASIS FUNCTIONS



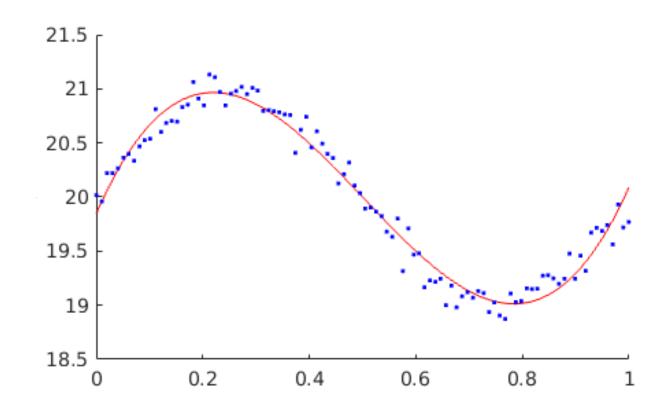
$$h_w(x) = w^T \phi(x)$$

$$h_{w}(x) = \begin{bmatrix} 1 & x_{1}^{(1)} & x_{2}^{(1)} & (x_{1}^{(1)})^{2} & (x_{2}^{(1)})^{2} \\ 1 & x_{1}^{(2)} & x_{2}^{(2)} & (x_{1}^{(2)})^{2} & (x_{2}^{(2)})^{2} \end{bmatrix} \begin{bmatrix} w_{0} \\ w_{1} \\ w_{2} \\ w_{3} \\ w_{4} \end{bmatrix}$$

$$h_w(x) = \begin{bmatrix} w_0 + w_1 x_1^{(1)} + w_2 x_2^{(1)} + w_3 (x_1^{(1)})^2 + w_4 (x_2^{(1)})^2 \\ w_0 + w_1 x_1^{(2)} + w_2 x_2^{(2)} + w_3 (x_1^{(2)})^2 + w_4 (x_2^{(2)})^2 \end{bmatrix}$$

BASIS FUNCTIONS POLYNOMIAL BASIS FUNCTIONS

Polynomial basis functions:



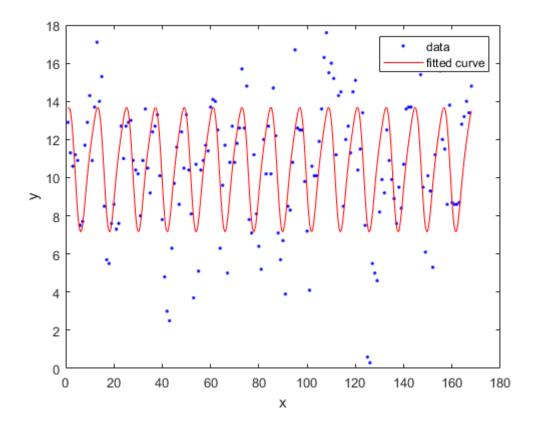
BASIS FUNCTIONS FOURIER SERIES



Fourier series

$$\phi_0(x)=1$$

$$\phi_j(x) = \cos(\boldsymbol{\varpi}_j x^{(i)} + \psi_j); j > 0$$

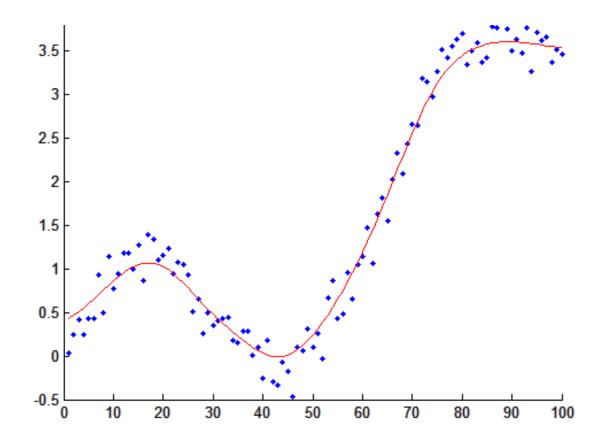


BASIS FUNCTIONS
RADIAL BASIS



Radial basis function:

$$\phi_j(x) = e^{-\frac{1}{2l}\sum_{r=1}^n (x_r^{(i)} - \mu_{j,r})^2}$$



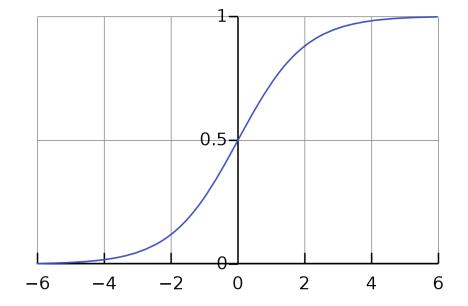
BASIS FUNCTIONS

SIGNMOIN DALS

Sigmoid function:

$$\phi_j(x^{(i)}) = \sigma\left(\frac{x^{(i)} - \mu_j}{s}\right)$$

$$\sigma(a) = \frac{1}{1 + e^{-a}}$$





The **same** estimation **model** is proposed **plus** an **error**:

$$y^{(i)} = h_w(x^{(i)}) + \varepsilon^{(i)}$$

Where:

$$h_w(x^{(i)}) = w^T \phi(x)$$

Assuming again that the **error** is **normally distributed** with **mean 0** and **variance** β^{-1} . It is defined for m training data:

$$p(\overrightarrow{y}/X, w, \beta) = \prod_{i=1}^{m} p(y^{(i)}/x^{(i)}; w) = \prod_{i=1}^{m} \sqrt{\frac{\beta}{2\pi}} e^{\left(-\frac{\beta}{2}\left(y^{(i)}-w^{T}\phi(x^{(i)})\right)^{2}\right)}$$

Since the **likelihood function** is **parameterized** in w and β the logarithmic likelihood is written like this:

$$\log p(\vec{y}/w, \beta) = \log \prod_{i=1}^{m} \sqrt{\frac{\beta}{2\pi}} e^{\left(-\frac{\beta}{2}\left(y^{(i)} - w^{T}\phi(x^{(i)})\right)^{2}\right)}$$

$$\log p(\vec{y}/w, \beta) = \sum_{i=1}^{m} \log \sqrt{\frac{\beta}{2\pi}} e^{\left(-\frac{\beta}{2}\left(y^{(i)} - w^{T}\phi(x^{(i)})\right)^{2}\right)}$$

Developing:

$$\log p(\vec{y}/w, \beta) = \sum_{i=1}^{m} \log \sqrt{\frac{\beta}{2\pi}} + \log e^{\left(-\frac{\beta}{2}\left(y^{(i)} - w^{T}\phi(x^{(i)})\right)^{2}\right)}$$

$$\log p(\vec{y}/w, \beta) = m \log \sqrt{\frac{\beta}{2\pi}} - \sum_{i=1}^{m} \frac{\beta}{2} \left(y^{(i)} - w^{T} \phi(x^{(i)}) \right)^{2}$$



Developing:

$$\log p(\vec{y}/w, \beta) = \frac{m}{2} \log \beta - \frac{m}{2} \log 2\pi - \sum_{i=1}^{m} \frac{\beta}{2} \left(y^{(i)} - w^T \phi(x^{(i)}) \right)^2$$

$$\log p(\vec{y}/w, \beta) = \frac{m}{2} \log \beta - \frac{m}{2} \log 2\pi - \beta E_D(w)$$

$$E_D(w) = \frac{1}{2} \left(y^{(i)} - w^T \phi(x^{(i)}) \right)^2$$

Developing:

$$-\log p(\vec{y}/w,\beta) = -\frac{m}{2}\log\beta + \frac{m}{2}\log2\pi + \beta E_D(w)$$

$$\underset{\overrightarrow{w}}{\operatorname{arg\,min}} - log(L(\overrightarrow{w})) =_{\underset{\overrightarrow{w}}{\operatorname{arg\,min}}} \frac{\beta}{2} \sum_{i=1}^{m} \left(y^{(i)} - w^{T} \phi(x^{(i)}) \right)^{2}$$

BASIS FUNCTIONS NORMAL EQUATIONS

Calculating the **gradient**, **setting** it equal to **zero** and solving for the **vector** *w*:

$$\boldsymbol{w} = \left(\boldsymbol{\phi}^T \boldsymbol{\phi}\right)^{-1} \boldsymbol{\phi}^T \vec{\boldsymbol{y}}$$