

Problem 1: Random Number Transformation

1) What is the probability density function of $Z = \max(X, Y)$?

$$\begin{aligned}
 \text{Ans: } P(Z \leq a) &= P(X \leq a \text{ and } Y \leq a) \\
 &= P(X \leq a) P(Y \leq a) \\
 &= F_X(a) F_Y(a) \\
 &= \left(\int_0^a e^{-x} dx \right) \left(\int_0^a e^{-y} dy \right) \\
 &= (1 - e^{-a})^2 = 1 - 2e^{-a} + e^{-2a}
 \end{aligned}$$

 \therefore The pdf of Z is

$$f_Z(a) = \begin{cases} -2e^{-a} + 2e^{-2a} & \text{for } a > 0 \\ 0 & \text{elsewhere} \end{cases}$$

2) What is the probability density function of $W = \min(X, Y)$?

$$\begin{aligned}
 \text{Ans: } P(W > a) &= P(X > a \text{ and } Y > a) = P(X > a) P(Y > a) \\
 &= 1 - P(W \leq a)
 \end{aligned}$$

$$\begin{aligned}
 F_W(a) &= 1 - P(W > a) = 1 - P(X > a) P(Y > a) \\
 &= 1 - (1 - P(X \leq a)) (1 - P(Y \leq a)) \\
 &= 1 - (1 - F_X(a)) (1 - F_Y(a)) \\
 &= 1 - (1 - (1 - e^{-a})) (1 - (1 - e^{-a})) \\
 &= 1 - (e^{-a})^2 \\
 &= 1 - e^{-2a}
 \end{aligned}$$

 \therefore The pdf of W is

$$f_W(a) = \begin{cases} 2e^{-2a} & \text{for } a > 0 \\ 0 & \text{elsewhere} \end{cases}$$

Problem 2: Statistical Distances

1) What are the four conditions that f must satisfy for $f(\cdot)$ to be considered a metric?

Ans: $f: X \times X \rightarrow \mathbb{R}^+$ (\mathbb{R}^+ : the set of non-negative real numbers) such that for all x, y, z in X , f satisfies the following condition:

1. $f(x, y) \geq 0$ (non-negativity)
2. $f(x, y) = 0$ if and only if $x = y$ (identity of indiscernibles)
3. $f(x, y) = f(y, x)$ (symmetry)
4. $f(x, z) \leq f(x, y) + f(y, z)$ (subadditivity/triangle inequality)

2) Prove that the squared Euclidean distance is a Bregman divergence.

$$\text{Ans: } d(x, y) = \|x - y\|^2 = \langle x - y, x - y \rangle$$

$$\begin{aligned}
 &= \|x\|^2 - 2xy + \|y\|^2 \\
 &= \|x\|^2 - \|y\|^2 + 2\|y\|^2 - 2xy \\
 &= \|x\|^2 - \|y\|^2 - 2y \cdot (x - y) \\
 &= \|x\|^2 - \underbrace{(\|y\|^2 + 2y \cdot (x - y))}_{\text{Tangent of } f \text{ at } y}
 \end{aligned}$$

$$\Rightarrow F(x) - (F(y) + \langle \nabla F(y), x - y \rangle) \rightarrow \text{The Bregman distance definition}$$

 \therefore squared Euclidean distance is a Bregman divergence.

3) Prove that the entropy of the output of this neural network will always be equal or less than that of the input.

$$\text{Ans: } H(X, Y) = H(X|Y) + H(Y) = H(Y|X) + H(X) \dots \textcircled{1}$$

$$\text{We know that } H(Y|X) = H(F_{X,Y}|X) = 0 \dots \textcircled{2}$$

$$\begin{aligned}
 \text{use } \textcircled{1}, H(X, F_{X,Y}) &= H(X) + H(F_{X,Y}|X) = H(F_{X,Y}) + H(X|F_{X,Y}) \\
 &= H(X) + 0 = H(F_{X,Y}) + H(X|F_{X,Y})
 \end{aligned}$$

$$\Rightarrow H(X) = H(F_{X,Y}) + H(X|F_{X,Y}) \dots \textcircled{3}$$

$$\text{and we know } H(X|F_{X,Y}) \text{ always } \geq 0, \therefore H(X) = -\sum_{i=1}^n P(X_i) \log(P_{X_i})$$

$$P_{X_i} \text{ always } \leq 1, \text{ then } \log(P_{X_i}) < 0, \therefore H(X) \text{ always } \geq 0 \dots \textcircled{4}$$

$$\therefore \text{ with } \textcircled{3} \text{ and } \textcircled{4}, \text{ we know } H(X|F_{X,Y}) \geq 0, \therefore H(X) - H(F_{X,Y}) \geq 0$$

 \therefore the entropy of the output of this NN will always be equal or less than that of the input.

Problem 4: Goodness of Estimation

1) What is the unbiased estimator with the lowest variance that you can construct from a linear combination of θ_1 and θ_2 , and what's its variance?

$$\text{Ans: } \text{Var}(\theta_1) = 1, \text{Var}(\theta_2) = 2, \text{Cov}(\theta_1, \theta_2) = \frac{1}{4}$$

$$\text{set linear combination of } \theta_1 \text{ and } \theta_2 \text{ is } c_1 \hat{\theta}_1 + c_2 \hat{\theta}_2$$

$$\text{we set } c_1 + c_2 = 1, \text{ then } E(c_1 \hat{\theta}_1 + c_2 \hat{\theta}_2) = c_1 \theta_1 + c_2 \theta_2 = \theta, \text{ Using } c_2 = 1 - c_1,$$

$$\begin{aligned}
 \text{we have: } \text{Var}(c_1 \hat{\theta}_1 + c_2 \hat{\theta}_2) &= c_1^2 \text{Var}(\hat{\theta}_1) + c_2^2 \text{Var}(\hat{\theta}_2) + 2c_1 c_2 \text{Cov}(\hat{\theta}_1, \hat{\theta}_2) \\
 &= c_1^2 + 2c_2^2 + \frac{1}{2} c_1 c_2 \\
 &= c_1^2 + 2(1 - c_1)^2 + \frac{1}{2} c_1 (1 - c_1) \\
 &= \frac{5}{2} c_1^2 - \frac{7}{2} c_1 + 2 \\
 &= \frac{5}{2} (c_1 - \frac{7}{10})^2 + \frac{31}{10}
 \end{aligned}$$

$$\text{minimizing the expression, we have } c_1 = \frac{7}{10}, c_2 = \frac{3}{10}$$

$$\text{Var}(c_1 \hat{\theta}_1 + c_2 \hat{\theta}_2) = \frac{5}{2} \times \frac{49}{100} - \frac{7}{2} \times \frac{7}{10} + 2 = \frac{31}{40}$$

The minimum variance estimator of this unbiased class is

$$\frac{7}{10} \hat{\theta}_1 + \frac{3}{10} \hat{\theta}_2$$

(2) Also answer the same question for θ_3 and θ_4 .

Ans: $\text{Var}(\theta_3) = 1, \text{Var}(\theta_4) = 2, \text{Cov}(\theta_3, \theta_4) = \frac{3}{4}$

Set linear combination of θ_3 and θ_4 is $C_3\hat{\theta}_3 + C_4\hat{\theta}_4$
 We set $C_3 + C_4 = 1$, then $E(C_3\hat{\theta}_3 + C_4\hat{\theta}_4) = C_3\theta + C_4\theta = \theta$
 Using $C_4 = 1 - C_3$, we have

$$\begin{aligned}\text{Var}(C_3\hat{\theta}_3 + C_4\hat{\theta}_4) &= C_3^2 \text{Var}(\hat{\theta}_3) + C_4^2 \text{Var}(\hat{\theta}_4) + 2C_3C_4 \text{Cov}(\hat{\theta}_3, \hat{\theta}_4) \\ &= C_3^2 + 2C_4^2 + \frac{3}{2}C_3C_4 \\ &= C_3^2 + 2(1-C_3)^2 + \frac{3}{2}C_3(1-C_3) \\ &= \frac{3}{2}C_3^2 - \frac{5}{2}C_3 + 2 \\ &= \frac{3}{2}(C_3 - \frac{5}{6})^2 + \frac{23}{24}\end{aligned}$$

minimizing the expression, we have $C_3 = \frac{5}{6}, C_4 = \frac{1}{6}$.

$$\text{Var}(C_3\hat{\theta}_3 + C_4\hat{\theta}_4) = \frac{3}{2} \times \frac{25}{36} - \frac{5}{2} \times \frac{5}{6} + 2 = \frac{23}{24}$$

The minimum variance estimator of this unbiased class is

$$\frac{5}{6}\hat{\theta}_3 + \frac{1}{6}\hat{\theta}_4$$

(3) Ans: When $\text{Cov}(\theta_1, \theta_2) = \frac{1}{4}$ can produce an estimator with lower variance, so when the lower Covariance can produce an estimator with lower variance.

Problem 3: Point Estimation

(1) The moment method

$$\text{for } f(x) = \begin{cases} \frac{1}{\theta}, & \text{for } 0 \leq x \leq \theta \\ 0, & \text{elsewhere} \end{cases} \Rightarrow F(x) = \begin{cases} 0, & \text{for } x < 0 \\ \frac{x}{\theta}, & \text{for } 0 \leq x \leq \theta \\ 1, & \text{for } x > \theta \end{cases}$$

in this case, we have $\mu = E(X) = \frac{\theta}{2}$.

$$\therefore \bar{X} = \frac{\theta}{2} \therefore \hat{\theta} = 2\bar{X}$$

(2) The MAP method

$$\hat{\theta}_{\text{MAP}}(x) = \arg \max_{\theta} f(x|\theta)g(\theta).$$

$\therefore g(\theta)$ is the prior distribution and is a uniform prior.

$$\therefore \hat{\theta}_{\text{MAP}}(x) = \arg \max_{\theta} f(x|\theta) = \hat{\theta}_{\text{MLE}}(x).$$

For a uniform distribution, the likelihood function can be written as:

$$L(\theta) = \prod_{i=1}^n f(x_i; \theta) = \prod_{i=1}^n \frac{1}{\theta} = \frac{1}{\theta^n}$$

$$\ln L(\theta) = \ln \prod_{i=1}^n f(x_i; \theta) = \ln \prod_{i=1}^n \frac{1}{\theta} = \ln(\theta^{-n}) = -n \ln \theta$$

$$\frac{\partial \ln L(\theta)}{\partial \theta} = \frac{-n}{\theta}$$

notice that the derivative with respect to θ is monotonically decreasing.

Thus, the $\hat{\theta}_{\text{MLE}}$ would be the smallest θ possible, which is:

$$\hat{\theta}_{\text{MLE}} = \max(X_1, X_2, \dots, X_n)$$

$$\therefore \hat{\theta}_{\text{MAP}} = \max(X_1, X_2, \dots, X_n)$$

(3) The Bayesian method using squared error loss function.

Population $f(x; \theta) = \frac{1}{\theta}$, for $0 \leq x < \theta$

Prior distribution $h(\theta) = 1$, for $0 < \theta < 1$

Joint distribution $u(x, \theta) = h(\theta)f(x; \theta) = 1(\frac{1}{\theta}) = \frac{1}{\theta}$, for $0 < x < \theta < 1$

Marginal distribution of x

$$g(x) = \int_{\theta=x}^1 u(x, \theta) d\theta = \int_x^1 \frac{1}{\theta} d\theta = -\ln x, \text{ for } 0 < x < 1$$

The conditional density of θ given x :

$$k(\theta/x) = \frac{u(x, \theta)}{g(x)} = \frac{1}{\theta} \times \frac{1}{-\ln x} = -\frac{1}{\theta \ln x}, \text{ for } 0 < x < \theta < 1$$

$$\hat{\theta} = E[\theta/x] = \int_x^1 \theta k(\theta/x) d\theta$$

$$= \int_x^1 \theta \frac{-1}{\theta \ln x} d\theta = -\frac{1}{\ln x} \int_x^1 d\theta = \frac{-(1-x)}{\ln x} = \frac{x-1}{\ln x}$$

$$\therefore \hat{\theta} = \frac{x-1}{\ln x}$$

Problem 4: Interval Estimation

(1) Show that $Q = X(1) - \theta$ is a pivotal quantity.

$$F(x) = \int_0^x e^{-(x-\theta)} dx = 1 - e^{-(x-\theta)}, \text{ for } 0 < x < \infty$$

for $f(x; \theta)$, the probability density of the r -th order statistic X_r is given by

$$f_{X_r}(x) = \frac{n!}{(r-1)!(n-r)!} [F(x)]^{r-1} [1-F(x)]^{n-r} f(x)$$

$$\text{when } r=1, f_{X_1}(x) = \frac{n!}{(1-1)!(n-1)!} [F(x)]^{1-1} [1-F(x)]^{n-1} f(x)$$

$$= n(e^{-(x-\theta)})^{n-1} \cdot e^{-(x-\theta)}$$

$$= n e^{-n(x-\theta)}, \text{ for } x > \theta$$

$$\text{for } Q = X(1) - \theta, f_Q(q) = n e^{-nq}, \text{ for } q > 0$$

$$F_Q(q) = 1 - e^{-nq}, \text{ for } q > 0$$

the distribution of Q is not depend on θ , $\therefore Q = X(1) - \theta$ is a pivotal quantity.

(2) Use this pivotal quantity find a $100(1-\alpha)\%$ confidence interval for θ .

we know $X(1) = \min(X_1, X_2, \dots, X_n)$ and $Q = X(1) - \theta, F_Q(q) = 1 - e^{-nq}$, then $p(Q > q) = e^{-nq}$

$$\text{for } p(\mu_1 \leq Q \leq \mu_2) = 1 - \alpha \therefore \frac{\alpha}{2} = P(Q \leq \mu_1) = 1 - e^{-n\mu_1}$$

$$\Rightarrow e^{-n\mu_1} = 1 - \frac{\alpha}{2}$$

$$\Rightarrow -n\mu_1 = \ln(1 - \frac{\alpha}{2})$$

$$\Rightarrow \mu_1 = -\frac{\ln(1 - \frac{\alpha}{2})}{n}$$

$$\frac{\alpha}{2} = P(Q > \mu_2) = e^{-n\mu_2}$$

$$\Rightarrow e^{-n\mu_2} = \frac{\alpha}{2}$$

$$\Rightarrow -n\mu_2 = \ln(\frac{\alpha}{2})$$

$$\Rightarrow \mu_2 = -\frac{\ln(\frac{\alpha}{2})}{n}$$

$$\therefore -\frac{\ln(1 - \frac{\alpha}{2})}{n} \leq Q \leq -\frac{\ln(\frac{\alpha}{2})}{n}$$

$$\Rightarrow -\frac{\ln(1 - \frac{\alpha}{2})}{n} \leq X(1) - \theta \leq -\frac{\ln(\frac{\alpha}{2})}{n}$$

$$\Rightarrow \frac{\ln(\frac{\alpha}{2})}{n} + X(1) \leq \theta \leq \frac{\ln(1 - \frac{\alpha}{2})}{n} + X(1)$$

Thus, this pivotal quantity find a $100(1-\alpha)\%$ confidence interval for θ is

$$\left[\frac{\ln(\frac{\alpha}{2})}{n} + X(1), \frac{\ln(1 - \frac{\alpha}{2})}{n} + X(1) \right]$$