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### Goals for the talk

- $\rightarrow$  Introduce the idea of the conformal prediction (CP) algorithm
- → Describe why this idea is so important
- ightarrow Explain how the CP framework lacks versatility
- $\rightarrow$  Propose a resolution relying on:
  - → Generalized fiducial (GF) inference
  - → Imprecise probability calculus

CP is a relatively general-purpose approach to uncertainty quantification for prediction problems in machine intelligence

### Desirable properties:

- $\rightarrow$  Finite sample control of type 1 error rates for predictions
- ightarrow Can be built on top of virtually any machine learning algorithm
- ightarrow Requires only weak assumptions on data generating mechanisms

```
Input: Nonconformity measure \Psi : \mathcal{T} \mathbb{R}^n \cap \times \mathbb{R} \to \mathbb{R}, measurable;
             Exchangeable examples y_1, \ldots, y_n, and an arbitrary y;
             Significance level \alpha \in (0,1)
   Output: Logical value;
               1 indicates that y_1, \ldots, y_n, y are exchangeable;
               0 else
1 Denote y_{n+1} := y;
2 for i \in \{1, ..., n+1\} do
3 Compute t_i(y_i) = \Psi(y_{-i}^{n+1}, y_i);
4 end
5 Set p_{n+1} := \frac{1}{n+1} \sum_{i=1}^{n+1} 1\{t_i(y_i) \ge t_{n+1}(y_{n+1})\};
6 return 1\{p_{n+1} > \alpha\};
```

### Finite sample control of type 1 errors:

Let  $\{\Gamma_n^{\alpha}: \alpha \in (0,1)\}$  be a family of CP sets for  $Y_{n+1}$  constructed from observed data  $Y_1, \ldots, Y_n$ 

The set  $\Gamma_n^{\alpha}$  is comprised of the values  $y_{n+1}$  such that  $p_{n+1} > \alpha$ 

#### **Theorem**

If  $Y_1, \ldots, Y_n, Y_{n+1} \sim P$  are exchangeable, then the CP sets are valid in the sense that for all  $(\alpha, n, P)$ ,

$$P(\Gamma_n^{\alpha} \ni Y_{n+1}) \ge 1 - \alpha$$

Remark: It suffices to assume only that  $t_1(Y_1), \ldots, t_n(Y_n), t_{n+1}(Y_{n+1})$  are exchangeable



# Why does validity matter?

Validity matters because accountability and reliability in uncertainty quantification matters — in the same way that:

- ightarrow Financial reporting standards exist to facilitate security valuation of insurance companies
- ightarrow Building codes and standards exist to ensure the integrity of engineering and construction practices
- ★ There is no generally accepted standard of accountability of stated uncertainties in all of data science

# Why does validity matter?

At the American Society of Clinical Oncology conference in Chicago last June:

A new liquid biopsy can help identify the need for adjuvant therapy in stage II colon cancer

- → thereby avoiding post-operative chemotherapy,
- $\rightarrow$  which for bowel cancer can cause peripheral neuropathy

# Why does validity matter?

Suppose the results of this biopsy is 95% confidence reported . . .

How is this confidence defined?

- $\rightarrow$  Is it defined as the reported error on a test set?
- $\rightarrow$  Is it a Bayesian posterior probability?
- $\rightarrow$  Is it some sort of averaging over a collection of predictions?
- ★ All are widely accepted notions of *confidence*
- ★ Varying (if any) guarantees for how the algorithm might perform on future data

The CP algorithm provides valid, general purpose uncertainty quantification, but lacks versatility:

 $\rightarrow$  Does not prescribe how to quantify the degree to which a data set provides evidence in support of (or against) an arbitrary event from a general class of events.

e.g., within the Bayesian paradigm, the degree to which a data set provides evidence in support of (or against) an event is quantified by the posterior probability of the event, for any *measurable* event.

### Approach:

- → Construct CP sets from the GF statistical framework
  - ightarrow Motivated by a rank-based data generating association
- → Apply imprecise probability tools
  - ightarrow e.g., belief/plausibility functions or lower/upper probabilities
- → Approximate imprecise GF distribution by a precise distribution

### Generalized fiducial inference

Assume a data generating model for some Y:

$$Y = G(U, \theta),$$

#### where

- $\rightarrow$  G is a deterministic function
- ightarrow heta is an unknown population parameter(s) of interest
- ightarrow U has a completely known and fully specified distribution

### Generalized fiducial inference

## Definition (Hannig et al., 2016)

Given an observed data set  $y_1, \ldots, y_n$  generated independently from  $Y = G(U, \theta)$ , a GF distribution on a parameter space  $\Theta$  is defined as the weak limit,

$$\lim_{\epsilon \to 0} \left\{ \underset{\vartheta \in \Theta}{\operatorname{argmin}} \sum_{i=1}^{n} \|y_i - G(U_i, \vartheta)\|^2 \mid \min_{\vartheta \in \Theta} \sum_{i=1}^{n} \|y_i - G(U_i, \vartheta)\|^2 \le \epsilon \right\}$$

#### For discrete-valued data:

- ightarrow The limit  $\epsilon 
  ightarrow 0$  reduces to setting  $\epsilon = 0$
- $\rightarrow$  Leads to an imprecise probability distribution over  $\Theta$

## Generalized fiducial inference

e.g., for  $Y_1, \ldots, Y_n \stackrel{\text{iid}}{\sim} \text{Bernoulli}(\theta)$ ,

$$Y_i = \underbrace{1\{U_i < \theta\}}_{=G(U_i,\theta)},$$

where  $U_1, \ldots, U_n \stackrel{\text{iid}}{\sim} \operatorname{uniform}(0,1)$ 

Leads to the (imprecise) GF distribution for  $\theta$ :

$$\big\{(U^\star_{(\sum_1^n y_i)}, U^\star_{(1+\sum_1^n y_i)}] \ : \ \text{where} \ U^\star_1, \dots, U^\star_n \overset{\text{iid}}{\sim} \text{uniform}(0,1)\big\} \subseteq \Theta,$$

 $\rightarrow U_{(k)}^{\star}$  denotes the k-th order statistic

Suppose  $Y_1, \ldots, Y_{n+1}$  are exchangeable and continuous

A model-free data generating association for  $Y_{n+1}$ :

$$\mathsf{rank}(t_{n+1}(Y_{n+1})) = V \sim \mathsf{uniform}\{1, \dots, n+1\},\$$

where

- $\rightarrow t_i(Y_i) := \Psi(Y_{-i}^{n+1}, Y_i)$  is a nonconformity score
- $\rightarrow$  rank $(t_{n+1}(Y_{n+1}))$  denotes position in ascending order

Using the rank-based data association,

The (imprecise) GF distribution of the to-be-predicted value  $y_{n+1}$  is a distribution over the random sets:

$$A_n(V^\star) := \underset{y}{\operatorname{argmin}} \left\{ |\operatorname{rank}(t_{n+1}(y)) - V^\star| \right\}$$

$$= \left\{ y : \operatorname{rank}(t_{n+1}(y)) = V^\star \right\}$$

where  $V^{\star} \sim \mathsf{uniform}\{1,\ldots,n+1\}$ 

Illustration of the imprecise GF distribution of  $y_{n+1}$ 

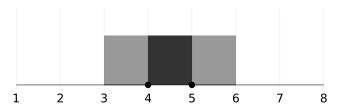


Figure: Hypothetical observed univariate data with  $y_1 = 4$ ,  $y_2 = 5$ , and n = 2. With nonconformity measure  $\Psi(y_{-i}^{n+1}, y_i) := |\text{mean}(y_{-i}^{n+1}) - y_i|$ 

- $\rightarrow A_n(1) = \text{black region}$
- $\rightarrow A_n(2) = \text{grey region}$
- $\rightarrow A_n(3) = \text{white region}$

With respect to the discrete uniform measure  $\mu$ ,

$$\mu\big(A_n(V^\star)\ni y_{n+1}\big)=\mu\bigg(V^\star=\mathsf{rank}(t_{n+1}(y_{n+1}))\bigg)=\frac{1}{n+1}$$

i.e.,  $A_n(1), \ldots, A_n(n+1)$  are all equally likely to contain  $y_{n+1}$ 

How to construct a prediction set with at least  $1-\alpha$  level confidence?

 $\rightarrow$  Accumulate k of the prediction sets such that

$$\frac{k}{n+1} \ge 1 - \alpha$$

Accordingly, for  $k \in \{1, \dots, n+1\}$ ,

$$\Omega_n(k) := \bigcup_{1 \le v \le k} A_n(v) = \left\{ y : \operatorname{rank}(t_{n+1}(y)) \le k \right\}$$

is a  $\frac{k}{n+1}$  level prediction set

$$\to f_n(y) := \mu(\Omega_n(V^\star) \ni y)$$
 is a conformal transducer

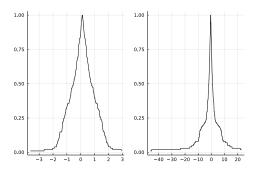


Figure:  $f_n(y) = \mu(\Omega_n(V^*) \ni y)$ ; the left and right plots are based on simulated samples of n = 100 realizations from the standard Gaussian and standard Cauchy distribution, respectively

 $\Upsilon_n^{\alpha} := \{ y : f_n(y) > \alpha \}$  is a CP set, i.e., valid in the sense that

$$P(\Upsilon_n^{\alpha} \not\ni Y_{n+1}) = P(f_n(Y_{n+1}) \le \alpha) \le \alpha$$



#### To summarize:

- $\rightarrow$  The sets  $A_n(1), \dots, A_n(n+1)$  are the atoms of the random set GF predictive distribution
- $\rightarrow$  Each set has  $\frac{1}{n+1}$  GF probability
- ightarrow These sets can be arranged to construct any CP set
- $\rightarrow$  Further, for any assertion B not necessarily a CP set:

(lower probability) 
$$\underline{\Pi}_n(B) := \mu\{A_n(V^*) \subseteq B \mid A_n(V^*) \neq \emptyset\}$$
 (upper probability) 
$$\overline{\Pi}_n(B) := 1 - \underline{\Pi}_n(B^c)$$

What if the lower and upper probabilities are difficult to compute?

- ightarrow Construct a precise approximation for model-free GF inference
- ightarrow Uniform sampling over  $A_n(1),\ldots,A_n(n+1)$  seems to be the sensible thing to do
- $\rightarrow$  Leads to the precise model-free GF (MFGF) distribution with density:

$$\pi_{y}(y_{n+1}) = \sum_{v^{\star}=1}^{n+1} \pi_{y,v}(y_{n+1}, v^{\star})$$

$$= \sum_{v^{\star}=1}^{n+1} \frac{1}{\mu \{A_{n}(v^{\star})\}} \cdot \frac{1}{n+1} \cdot 1\{y_{n+1} \in A_{n}(v^{\star})\},$$

### **Algorithm 1:** Computing the MFGF predictive distribution.

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Input: Prediction regions A_n(1), \ldots, A_n(n+1) and a desired sample size N.
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**Output:** A sample from the MFGF distribution of  $Y_{n+1}$ .

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1 Initialize an N-dimensional vector \widetilde{y};

2 for j \in \{1, ..., N\} do

3 | Sample v^* \sim \text{uniform}\{1, ..., n+1\};

4 | Sample y_{n+1} \sim \text{uniform}(A_n(v^*));

5 | Set \widetilde{y}_j := y_{n+1};
```

6 end

7 return  $\widetilde{y}$ ;

### **Theorem**

If  $Y_1, \ldots, Y_n, Y_{n+1} \stackrel{iid}{\sim} P$  is a collection of continuous random variables, then for any  $\epsilon > 0$ ,  $\alpha \in (0,1)$ , and  $v \in \{1,\ldots,n\}$ ,

$$P\Big(n^{\alpha}\big|\pi_{y}\{A_{n}(v)\}-P\{A_{n}(v)\}\big|>\epsilon\Big)\leq e^{-n^{1-\alpha}\epsilon}.$$

 $\rightarrow$  MFGF distribution converges to the true distribution of  $Y_{n+1}$ .



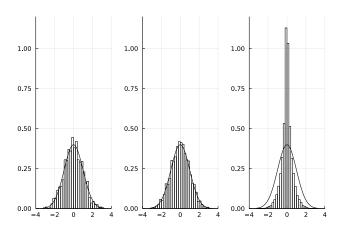


Figure: The middle and right panels display histograms of samples of size 100,000 drawn from the MFGF distribution and CP-induced distribution, respectively, based on n=1,000 data points drawn from a Gaussian distribution

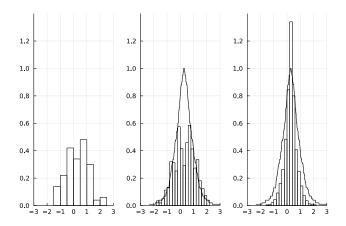


Figure: The middle and right panels display histograms of samples of size 10,000 drawn from the MFGF distribution and CP-induced distribution, respectively, based on n=100 data points drawn from the standard Gaussian distribution

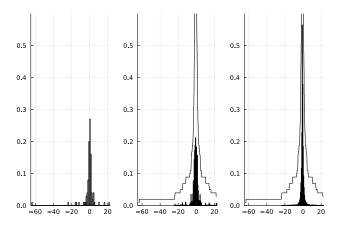


Figure: The middle and right panels display histograms of samples of size 10,000 drawn from the MFGF distribution and CP-induced distribution, respectively, based on n=100 data points drawn from the standard Cauchy distribution

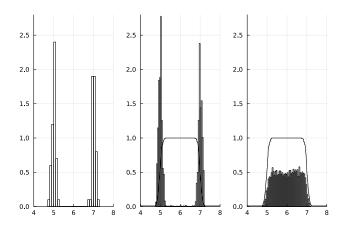


Figure: The middle and right panels display histograms of samples of size 10,000 drawn from the MFGF distribution and CP-induced distribution, respectively, based on n=100 data points drawn from a mixture of two Gaussian distributions

### Link to preprint:

Coming soon

My personal academic website:

https://jonathanpw.github.io

The end