Probability

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Lecture 1: Probability Space Example. If we have a die with outcomes 1, 2,, 6. 1. $\mathbb{P}(2) = \frac{1}{6}$ 2. $\mathbb{P}(\text{multiple of } 3) = \frac{2}{6} = \frac{1}{3}$ 3. $\mathbb{P}(\text{pair or a multiple of } 3) = \frac{4}{6} = \frac{2}{3}$				20 Jan. 11:00

1 Formal Setup

We try to define a probability space rigorously in this section.

Definition 1.1 (Probability Space). We have the following,

- 1. Sample space Ω , a set of outcomes.
- 2. \mathcal{F} , a collection of subsets of Ω (called events).
- 3. \mathcal{F} is a σ -algebra if
 - (a) **F1**: $\Omega \in \mathcal{F}$
 - (b) **F2**: if $A \in \mathcal{F}$ then $A^c \in \mathcal{F}$
 - (c) **F3**: For all countable collections $\{A_n\}$ in \mathcal{F} , $\cup_n A_n \in \mathcal{F}$.

Given σ -algebra \mathcal{F} on Ω , function $\mathbb{P}: \mathcal{F} \to [0,1]$ is a probability measure if

- 1. **P1**: The probability function is nonnegative.
- 2. **P2**: $\mathbb{P}(\Omega) = 1$
- 3. **P3**: For all countable collection $\{A_n\}$ of disjoint events in \mathcal{F} , we have $\mathbb{P}(\cup_n A_n) = \sum_{n=1}^{\infty} \mathbb{P}(A_n)$.

Then $(\Omega, \mathcal{F}, \mathbb{P})$ is a probability space.

Problem. Why $\mathbb{P}: \mathcal{F} \to [0,1]$, not $\mathbb{P}: \Omega \to [0,1]$?

We will justify the definition in the following examples.

Example. When Ω is finite or countable,

- 1. In general: $\mathcal{F} = \mathcal{P}(\Omega)$.
- 2. $\mathbb{P}(2)$ is shorthand for $\mathbb{P}(\{2\})$.
- 3. \mathbb{P} is determined by $\mathbb{P}(\{w\}), \forall w \in \Omega$.

Remark. When Ω is uncountable, a probability space behaves differently, as shown in the following example.

Example. If $\Omega = [0, 1]$, and we want to choose a real number, all equally likely.

If $\mathbb{P}\{0\} = \alpha > 0$, then $\mathbb{P}(\{0,1,\frac{1}{2},\ldots,\frac{1}{n}\} = n\alpha)$. This cannot happen if n large, because we would have $\mathbb{P} > 1$. So $\mathbb{P}(\{0\}) = 0$ or undefined.

Example. When Ω is infinitely countable (e.g., $\Omega = \mathbb{N}$ or $\Omega = \mathbb{Q} \cap [0,1]$), however, it is not possible to choose uniformly. Suppose it is possible, there are two possibilities

• If $\mathbb{P}(\{\omega\}) = \alpha \quad \forall \omega \in \Omega$,

then
$$\mathbb{P}(\Omega) = \sum_{\omega \in \Omega} \mathbb{P}(\{\omega\}) = \infty$$
. \nleq

• If $\mathbb{P}(\{\omega\}) = 0 \quad \forall \omega \in \Omega$,

then
$$\mathbb{P}(\Omega) = \sum_{\omega \in \Omega} \mathbb{P}(\{\omega\}) = 0.$$
 \nleq

So it is not possible to have one such uniform probability space. But that's fine as there exists many other interesting probability measures on a infinite countably set.

Property. From the axioms, we want to prove the following properties of a probability space.

1. $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$.

Proof.
$$A, A^c$$
 disjoint. $A \cup A^c = \Omega$. So $\mathbb{P}(A) + \mathbb{P}(A^c) = \mathbb{P}(\Omega) = 1$

- 2. $\mathbb{P}(\varnothing) = 0$
- 3. If $A \subseteq B$, then $\mathbb{P}(A) \leq \mathbb{P}(B)$.
- 4. $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) \mathbb{P}(A \cap B)$

1.1 Examples of Probability Spaces

Example. Here we list some concrete examples of probability spaces.

1. Ω finite, $\Omega = \{w_1, \dots, w_n\}$, $\mathcal{F} = \text{all subsets under uniform choice.}$

$$\mathbb{P}: \mathcal{F} \to [0,1], \mathbb{P}(A) = \frac{|A|}{|\Omega|}$$
. In particular: $\mathbb{P}(\{w\}) = \frac{1}{|\Omega|} \forall w \in \Omega$.

2. If we are choosing without replacement n indistinguishable marbles that are labelled $\{1, \ldots, n\}$. Pick $k \leq n$ marbles uniformly at random.

Here we have
$$\Omega = \{A \subseteq \{1, \dots, n\}, |A| = k, |\Omega| = \binom{n}{k}\}$$
.

3. If we have a well-shuffled deck of cards, and we uniformly chose permutation of 52 cards.

$$\Omega = \{\text{all permutations of 52 cards}\}. |\Omega| = 52!.$$

Then we have

$$\mathbb{P}(\text{first three cards have the same suit}) = \frac{52 \cdot 12 \cdot 11 \cdot 49!}{52!} = \frac{22}{425}$$

Lecture 2: Finite Probability Space

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Example (Coincidental Birthday). There we have n people, what is the probability that at least two share a birthday? To be precise, we first make the following assumptions,

- No leap years; (365 days in a year)
- All birthdays are equally likely.

We have the probability space

$$\Omega = \{1, \dots, 365\}^n$$

$$\mathcal{F} = \mathcal{P}(\Omega)$$

 $A = \{ \text{at least 2 people share birthday} \}$

 $A^c = \{ \text{all } n \text{ birthdays are different} \}.$

So we have the probability

$$\mathbb{P}(A^c) = \frac{365 \times 364 \times \ldots \times (365 - n - 1)}{365^n},$$

$$\mathbb{P}(A) = 1 - \frac{365 \times 364 \times \ldots \times (365 - n - 1)}{365^n}.$$

Remark.

• We note several special n values,

n = 22 : $\mathbb{P}(A) \approx 0.479$ n = 23 : $\mathbb{P}(A) \approx 0.507$ $n \ge 366$: $\mathbb{P}(A) = 1$

- The probability of birthday is not equal in real life though. It is more likely to be born about 9 months after christmas.
- Sometimes it would be easier to calculate the probability of the complement of an event.

1.2 Combinatorial Analysis

If Ω is a finite set such that $|\Omega| = n$,

Problem. How many ways to partition Ω into k disjoint subsets $\Omega_1, \ldots \Omega_k$ with $|\Omega_i| = n_i \ (\sum_{i=1}^k n_i = n)$?

The total number of ways M is

$$M = \binom{n}{n_i} \binom{n-n_1}{n_2} \binom{n-n_1-n_2}{n_3} \cdots \binom{n-n_1-n_2\cdots-n_{k-1}}{n_k}$$

$$= \binom{n}{n_i} \binom{n-n_1}{n_2} \binom{n-n_1-n_2}{n_3} \cdots \binom{n_k}{n_k}$$

$$= \frac{n!}{n!(n-n_1)!} \times \frac{(n-n_1)!}{n_2!(n-n_1-n_2)!} \times \cdots \times \frac{(n-n_1-n_2-\cdots-n_{k-1})!}{x_k!0!}$$

$$= \frac{n!}{n_1!n_2!\cdots n_k!}$$

$$= \binom{n}{n_1, n_2, \dots, n_k}$$

which is called the *multinomial coefficient*, and denoted by the last term in the equations.

Remark. The ordering of the subsets do matter in this setting.

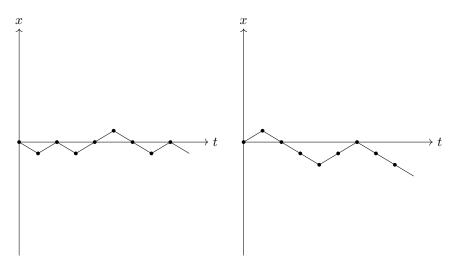


Figure 1: Random Walks

1.3 Random Walks

We have the following uniform probability space

$$\Omega = \{ (x_0, x_1, \dots, x_n) \mid x_0 = 0, |x_k - x_{k-1}| = 1, k = 1, \dots, n \},$$

$$|\Omega| = 2^n.$$

Problem. What's $\mathbb{P}(x_n = 0)$ and $\mathbb{P}(x_n = n)$?

We have $\mathbb{P}(x_n = n) = \frac{1}{2^n}$.

When n is odd, $\mathbb{P}(x_n = 0) = 0$ because after every step the value changes parity. To find the probability when n is even, we need to choose $\frac{n}{2}$ ks for which $x_k = x_{k-1} + 1$, and the rest $x_k = x_{k-1} - 1$. So

$$\mathbb{P}(x_n = 0) = 2^{-n} \binom{n}{n/2}$$
$$= \frac{n!}{2^n \left[\left(\frac{n}{2} \right) \right]!}.$$

Problem. What happens when n is large?

We next present Stirling's Formula, and we adopt the following notation for the time being.

Notation. If (a_n) , b_n are two sequences, we say $a_n \sim b_n$ as $n \to \infty$ if $\frac{a_n}{b_n} \to 1$ as $n \to \infty$.

Theorem 1.1 (Stirling's Formula).

$$n! \sim \sqrt{2\pi} n^{n+\frac{1}{2}} e^{-n}$$
 as $n \to \infty$.

We also have the weaker version

$$\log(n!) \sim n \log n$$
.

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Proof. We have

$$\log(n!) = \log 2 + \log 3 + \ldots + \log n.$$

So

$$\underbrace{\int_{1}^{n} \log x dx} \leq \log(n!) \leq int_{1}^{n+1} \log x dx$$

$$\underbrace{n \log n - n + 1}_{n \log n} \leq \log(n!) \leq \underbrace{(n+1) \log(n+1) - n}_{n \log n}.$$

 $\log(n!)$ is sandwiched between the lower and upper integrals, so $\log(n!)$ must be approximately $n \log n$ as well. In this calculation, these facts helped

- 1. $\log x$ is increasing, so it's easier to bounded by the integrals.
- 2. $\log x$ has a nice integral. So the integrals have closed forms.

(Ordered) Compositions

Definition 1.2. A composition of m with k parts is sequence $(m1, \ldots, m_k)$ of non-negative integers with $\sum_{i=1}^k m_i = m$.

We use stars and bars. There are m stars and k-1 bars, and

$$\#\text{Compositions} = \binom{m+k-1}{m}.$$

1.4 Properties of Probability Measures

Recall Definition (1.1). We prove the following properties.

Property.

1. Countable sub-additivity

Let $(A_n)_{n>1}$ sequence of events in \mathcal{F} . Then

$$\mathbb{P}(\cup_{n\geq 1} A_n) \leq \sum_{n\geq 1} \mathbb{P}(A_n).$$

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Proof. We rewrite $\cup_{n\geq 1}$ as a disjoint union.

Define
$$B_1 = A_1$$
 and $B_n = A_n \setminus (A_1 \cup \ldots \cup A_{n-1})$.

So

- $\bullet \cup_{n>1} B_n = \cup_{n>1} A_n,$
- $(B_n)_{n>1}$ disjoint (by construction),
- $B_n \subseteq A_n \implies \mathbb{P}(B_n) \le \mathbb{P}(A_n)$.

And we have

$$\mathbb{P}(\cup_{n\geq 1}A_n) = \mathbb{P}(\cup_{n\geq 1}B_n) = \sum_{n\geq 1}\mathbb{P}(B_n) = \sum_{n\geq 1}\mathbb{P}(A_n).$$

2. Continuity $(A_n)_{n\geq 1}$ increasing sequence of events in \mathcal{F} that is $A_n\subseteq A_{n+1}$ for all n.

In fact,
$$\lim_{n\to\infty} \mathbb{P}(A_n) = \mathbb{P}(\cup_{n\geq 1} A_n)$$
.

Proof. We reuse the B_n s, and we have

- $\bullet \ \sqcup_{k=1}^n B_k = A_n,$
- $\bullet \ \cup_{n>1} B_n = \cup_{n>1} A_n.$

So we have

$$\mathbb{P}(A_n) = \sum_{k=1}^n \mathbb{P}(B_k) \to \sum_{k>1} \mathbb{P}(B_k) = \mathbb{P}(\cup_{n\geq 1} B_n) = \mathbb{P}(\cup_{n\geq 1} A_n).$$

3. Inclusion-Exclusion Principle

Background:
$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$$
.

Similarly, for $A, B, C \in \mathcal{F}$,

$$\mathbb{P}(A \cup B \cup C) = \mathbb{P}(A) + \mathbb{P}(B) + \mathbb{P}(C) - \mathbb{P}(A \cap B) - \mathbb{P}(B \cap C) - \mathbb{P}(C \cap A) + \mathbb{P}(A \cap B \cap C).$$

The full Inclusion-Exclusion principle statement is the following. Let $A_1, \ldots, A_n \in \mathcal{F}$, then

$$\mathbb{P}(\cup_{i=1}^{n} A_{i}) = \sum_{i=1}^{n} \mathbb{P}(A_{i}) - \sum_{1 \leq i_{1} < i_{2} \leq n} \mathbb{P}(A_{i_{1}} \cap A_{i_{2}}) + \dots + (-1)^{n+1} \mathbb{P}(A_{1} \cap \dots \cap A_{n})$$

$$= \sum_{I \subseteq \{1,\dots,n\}} (-1)^{|I|+1} \mathbb{P}(\cap_{i \in I} A_{i}).$$

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