

## CS 475 Machine Learning: Homework 5

## Graphical Models

Due: Friday Nov 18, 2016, 11:59pm

100 Points Total

Version 1.0

**1 Analytical (40 points)**

**1. (12 points)** Consider the Bayesian Network given in Figure 1(a). Are the sets **A** and **B** d-separated given set **C** for each of the following definitions of **A**, **B** and **C**? Justify each answer.

- $\mathbf{A} = \{x_1\}, \mathbf{B} = \{x_9\}, \mathbf{C} = \{x_5, x_{14}\}$
- $\mathbf{A} = \{x_{11}\}, \mathbf{B} = \{x_{13}\}, \mathbf{C} = \{x_1, x_{15}\}$
- $\mathbf{A} = \{x_4\}, \mathbf{B} = \{x_5\}, \mathbf{C} = \{x_{10}, x_{16}\}$
- $\mathbf{A} = \{x_3, x_4\}, \mathbf{B} = \{x_{13}, x_9\}, \mathbf{C} = \{x_{10}, x_{15}, x_{16}\}$

Now consider a Markov Random Field in Figure 1(b), which has the same structure as the previous Bayesian network. Re-answer each of the above questions with justifications for your answers.

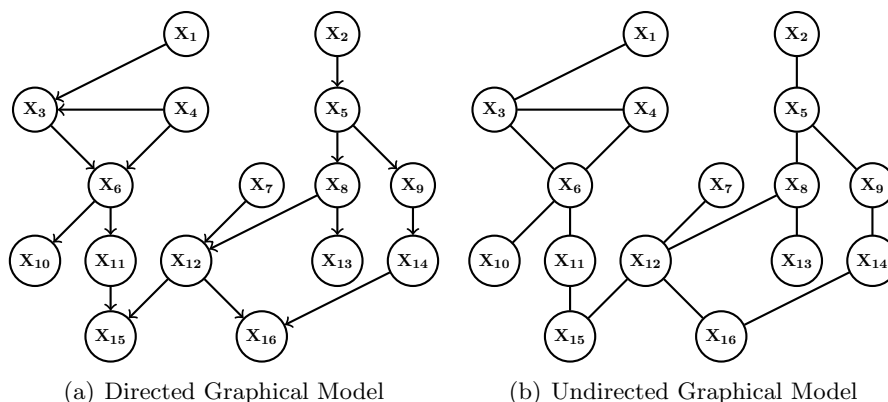


Figure 1: Two graphs are the same. However since (a) is directed and (b) is undirected the two graphs have different a conditional independence interpretation.

Directed:

- Sets **A** and **B** are d-separated because  $x_5$  is observed which blocks off  $x_9$  from the top because it is a tail to tail path, and  $x_{14}$  is observed which blocks off  $x_9$  from the bottom because it is a head to tail connection.
- Sets **A** and **B** are not d-separated because  $x_{15}$  is observed and it is a head to head connection, so it remains unblocked, so the path between  $x_{11}$  and  $x_{13}$  is unblocked.

- (c) Sets A and B are d-separated because  $x_{15}$  is unobserved, so it blocks the path between  $x_4$  and  $x_5$ .
- (d) Sets A and B are not d-separated because  $x_{15}$  is observed, and  $x_{16}$  is observed so the head to head connection at  $x_{16}$  is unblocked and the head to head connection at  $x_{15}$  is also unblocked.

Undirected

- (a) It is d-separated because all paths from  $x_1$  to  $x_9$  must pass through set C.
- (b) It is d-separated because all paths from set A to set B pass through  $x_{15}$
- (c) Not d-separated because not all paths from set A to set B pass through set C
- (d) It is d-separated because all paths from set A to set B have to pass through  $x_{15}$

**2. (15 points)** Let  $X = (X_1, \dots, X_d^T)$  be a random vector, which follows a  $d$ -variate Gaussian distribution  $N(0, \Sigma)$ . Define the precision matrix as  $\Omega = \Sigma^{-1}$ . Please prove that  $\Omega_{ij} = 0$  implies the conditional independence between variables  $X_i$  and  $X_j$  given the remaining variables, i.e., a sparsity pattern  $\Omega$  is equivalent to an adjacency matrix of a Gaussian Markov Random Field.

**My Answer:**

Using the hint from piazza, we have to prove that:

$$P(X_i, X_j | X_{-i, -j}) = P(X_i | X_{-i, -j}) P(X_j | X_{-i, -j})$$

And using the rules of conditional probability we can rearrange these probabilities such that the equation becomes:

$$\frac{P(X_i, X_j, X_{-i, -j})}{P(X_{-i, -j})} = \frac{P(X_i, X_{-i, -j})}{P(X_{-i, -j})} \frac{P(X_j, X_{-i, -j})}{P(X_{-i, -j})}$$

$$P(X_i, X_j, X_{-i, -j}) = P(X_i, X_{-i, -j}) \frac{P(X_j, X_{-i, -j})}{P(X_{-i, -j})}$$

Now looking at the the multivariate distribution, we can see that the above equation is only true if:

$$\exp(X^T \Omega X) = \frac{\exp(X_i^T \Omega X_i) \exp(X_j^T \Omega X_j)}{\exp(X_{-i, -j}^T \Omega X_{-i, -j})}$$

which in terms of sums is:

$$X^T \Omega X = X_i^T \Omega X_i + X_j^T \Omega X_j - X_{-i, -j}^T \Omega X_{-i, -j}$$

Looking at the left hand side, we have:

$$X^T \Omega X = \sum_{p, q | p, q \neq j} X_p^T \Omega_{pq} X_q + \sum_{p, q | p, q \neq i} X_p^T \Omega_{pq} X_q + \sum_{p, q | p, q \in \{i, j\}} X_p^T \Omega_{pq} X_q - \sum_{p, q | p, q \notin \{i, j\}} X_p^T \Omega_{pq} X_q$$

The third summation becomes 0 because  $\Omega_{ij}$  we know to be 0. So the LHS becomes:

$$\sum_{p, q | p, q \neq j} X_p^T \Omega_{pq} X_q + \sum_{p, q | p, q \neq i} X_p^T \Omega_{pq} X_q - \sum_{p, q | p, q \notin \{i, j\}} X_p^T \Omega_{pq} X_q$$

Now if we exponentiate this, it becomes:

$$\frac{\exp(X_i^T \Omega X_i) \exp(X_j^T \Omega X_j)}{\exp(X_{-i, -j}^T \Omega X_{-i, -j})}$$

which means the LHS equals RHS and our proof is complete.

**3. (13 points)** The probability density function of most Markov Random Fields cannot be factorized as the product of a few conditional probabilities. This question explores some MRFs which can be factorized in this way.

Consider the graph structure in Figure 2. From this graph, we know that  $X_2$  and  $X_3$

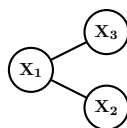


Figure 2: The Original Undirected Graph

are conditionally independent given  $X_1$ . We can draw the corresponding directed graph as Figure 3. This suggests the following factorization of the joint probability:

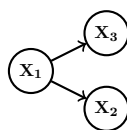


Figure 3: The Converted Directed Graph

$$P(X_1, X_2, X_3) = P(X_3|X_1)P(X_2|X_1)P(X_1)$$

Now consider the following graphical model in Figure 4.

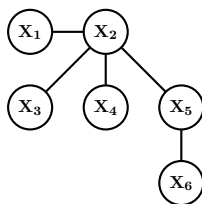


Figure 4: An Undirected Graph

As before, we can read the conditional independence relations from the graph.

- Following the example above, write a factorization of the joint distribution:  $P(X_1, X_2, X_3, X_4, X_5, X_6)$ .
- Is this factorization unique, meaning, could you have written other factorizations that correspond this model?
- If the factorization is unique, explain why it is unique. If it is not unique, provide an alternate factorization.

**My Answer:**

- $P(X_1, X_2, X_3, X_4, X_5, X_6) = P(X_1|X_2)P(X_3|X_2)P(X_4|X_2)P(X_6|X_5, X_2)$
- No its not unique.
- $P(X_1, X_2, X_3, X_4, X_5, X_6) = P(X_2|X_1)P(X_3|X_2, X_1)P(X_4|X_2, X_1)P(X_6|X_5, X_2, X_1)$