

CS 294 - Partition Functions

Joon Kim



Introduction

Def) Partition Function: polynomial $Z(\lambda) = \sum_{k=0}^D a_k \lambda^k$ where $a_k \geq 0$.

Goal: Compute $Z(\lambda)$ for various λ .

↪ Why? Intrinsic interest in combinatorics, coeff. $\{a_k\}$ are interesting, connection to Gibbs distribution ($\pi_\lambda(\sigma) = \frac{1}{Z(\lambda)} \lambda^{E(\sigma)}$ "energy" of config σ), derivatives of Z carry additional info, notion of phase transitions by singularities in $\log Z$

Course Scope: Complexity/Exact Computation, Approximation Algorithms, Approximating $Z(\lambda) \Leftrightarrow$ Sampling from π_λ , MCMC & Mixing Times, Correlation Decay (Spatial Properties), Geometry of Polynomials & Zeros, Connections with Phase Transitions

Counting Problems

Decision Problems → Counting Problems (how many sat. assignments?)

ex) For bipartite graph G , \exists perfect matching? If so, how many?

Let $Z(x) := \sum_k a_k x^k$ be the generating function for integer seq $\{a_k\}$.

ex) $a_k := \#$ of assignments that satisfy exactly k clauses of Ψ . Then, $a_m = \#$ of satisfying assigns ($m = \#$ clauses).

ex) $a_k := \#$ of matching of k edges. Then, $a_n = \#$ of perfect matchings ($n = \frac{|V|}{2}$).

Spin Systems

Def) Spin System: graph G , set of spins $\{1, \dots, q\}$. Configurations are assigns $\sigma: V \rightarrow [q]$. Hamiltonian $H(\sigma) := \sum_{\substack{(u,v) \in E \\ \text{symmetric}}} f(\sigma(u), \sigma(v)) + \sum_{v \in V} g(\sigma(v))$.

$\hookrightarrow \Pr[\text{system is in } \sigma] = \frac{1}{Z(\beta)} \exp(-\beta H(\sigma))$ where $\beta > 0$ is a parameter.
(inverse temp.)

$Z(\beta) := \sum_{\sigma} \exp(-\beta H(\sigma))$. Let $\lambda = \exp(-\beta)$. Then, we can rewrite to

$Z(\beta) = \sum_k a_k \lambda^k$ where $a_k = \#$ of configs where $H(\sigma) = k$.

Examples of Spin Systems:

1) Ferromagnetic Ising Model: spin $\{\uparrow, \downarrow\} / \{+, -\}$.

+	-	+	+	+	+	-
-	-	+	-	-	+	
+	+	-	+	-	+	

$$f(+,+) = f(-,-) = \emptyset, f(+,-) = f(-,+) = 1.$$

$g(+)= -h, g(-)=0$ where $h \geq 0$ is external.

$$H(\sigma) = \# \text{Disagreements}(\sigma) - h \# \text{Pluses}(\sigma).$$

$$Z_G(\beta, h) = \sum_{\sigma} \exp(-\beta H(\sigma)) \rightarrow Z_G(\lambda, \mu) = \sum_{\sigma} \lambda^{\# \text{Dis}(\sigma)} \mu^{\# \text{Plus}(\sigma)} \text{ where}$$

$$\lambda \leftarrow \exp(-\beta), \mu \leftarrow \exp(\beta h) \Rightarrow \sum_{j,k} a_{jk} \lambda^j \mu^k \text{ where } a_{jk} := \# \text{ of configgs}$$

with j disagreements and k pluses, enumerating all σ s!

→ If we set $\mu=1$ ($h=0$, no ext. field), $Z_G(\lambda) = \sum_j b_j \lambda^j$ where $b_j :=$
 $\# \text{ of cuts in } G \text{ with } j \text{ edges} \Rightarrow \text{Cut generating polynomial!}$

2) Anti-Ferromagnetic Ising: $f(+,+) = f(-,-) = 1, f(+,-) = f(-,+) = \emptyset$.

$$\hookrightarrow H(\sigma) = \# \text{Agreements}(\sigma) - h \# \text{Pluses}(\sigma) = |E| - \# \text{Dis}(\sigma) - h \# \text{Pluses}(\sigma)$$

$$\rightarrow Z_G(\lambda, \mu) = \lambda^{-|E|} \sum_{j,k} a_{jk} \lambda^j \mu^k \text{ where } \lambda \leftarrow \underline{\exp(\beta)}, \mu \leftarrow \exp(\beta h).$$

3) Potts Model: Ising, but an arbitrary number $q \geq 2$ of spins.

$$f(i,j) = \begin{cases} 1 & \text{if } i \neq j \\ 0 & \text{if } i=j \end{cases} \text{ for Ferro.} \rightarrow Z(\lambda) = \sum_{S_1, \dots, S_q} \lambda^{|E(S_1, \dots, S_q)|}, \lambda \in (0,1] \text{ ferro.} \\ \lambda \geq 1 \text{ antif.}$$

4) Independent Sets/Hard-Core Model: $\sigma = \{\text{ind. sets in } G\}$.

$[q] = \{0,1\}$ where 1 is occupied, 0 unoccupied.

$f(i,i) = \infty$, $f(i,j) = 0$ otherwise. $g(k) = \begin{cases} l & \text{if } k=1 \\ 0 & \text{if } k=0. \end{cases}$

$\rightarrow Z(\lambda) = \sum_I \lambda^{|I|}$ where $\lambda = \exp(-\beta l) \rightarrow Z(\lambda) = \sum_k a_k \lambda^k$ where $a_k := \# \text{ of ind. sets in } G \text{ with exactly } k \text{ vertices}, \lambda \in (0, \infty)$

* Spin Glass Model has (random) mixture of ferro. and antif. interactions.

Markov Random Fields

a.k.a. undirected graphical models. For a graph $G(V,E)$,
RVs $\{X_v | v \in V\}$ (assume discrete) have the markov property s.t.
if a set S separates A and B , $X_A \perp\!\!\!\perp X_B$ given X_S .

Obs) Spin systems are MRFs.

Thm) Hammersly-Clifford: Let μ be an MRF on $G(V,E)$ with discrete
r.v.s $\{X_v\}$ s.t. μ assigns nonzero prob. to every config. $\{X_v | v \in V\}$.

Then, μ is equivalent to a spin system on G with Hamiltonian
 $H(x) := \sum_C \psi_c(x_c)$ where C ranges only over cliques in G .

Proof [Bessy]: WLOG, assume \emptyset is a valid value for all r.v.s and that
 $\pi(\vec{0}) = \pi(00..0) \neq 0$. Define $g(x) := \log \frac{\pi(x)}{\pi(\vec{0})}$. We can uniquely write

$g(x) = \sum_{u \in V} \psi_u(x_u)$ for real valued function ψ_u that only depend on x_u and

satisfy $\Psi_u(X_u) = \emptyset$ unless $X_u > \vec{0}$. E.g., s.p.s all r.v.s are \emptyset -1 valued.

$\rightarrow \Psi_{\{u,v\}}(X_{uv}) = \log \frac{\pi(\vec{I}_{uv})}{\pi(\vec{0})} \forall v$ where \vec{I}_v assigns $X_v = 1, X_u = \emptyset \forall u \neq v$.

Now, we want $g(\vec{I}_{uv}) = \Psi_{uv}(\vec{I}_{uv}) + \Psi_u(\vec{I}_u) + \Psi_v(\vec{I}_v)$ by definition.

Thus, $\Psi_{uv}(X_{\{u,v\}}) = g(\vec{I}_{uv}) - \Psi_u(X_u) - \Psi_v(X_v)$, and build up iteratively.

$\rightarrow \pi(x) = \pi(\vec{0}) \exp\left(\sum_{u \in V} \Psi_u(X_u)\right)$. We shall show that $\Psi_u = \emptyset \forall u$ that is not a clique.

Consider u, v s.t. $(u, v) \notin E$. Let $u, v \in U$. For any valuation x , set

$x' := x$ except $x'_u = \emptyset$. Then $g(x) - g(x') = \log \frac{\pi(x)}{\pi(x')} = \log \frac{\pi(X_u | X_{V \setminus \{u\}})}{\pi(\emptyset | X_{V \setminus \{u\}})}$

is independent of X_v by the Markov property. Also consider that

$g(x) - g(x') = \sum_{u \in u} \Psi_u(X_u)$, which should also be indep. of X_v .

Set $X_w = \emptyset \forall w \notin \{u, v\}$ and set $X_u \neq \emptyset$. Then, $\Psi_u(X_u) + \Psi_{uv}(X_{uv})$ is indep. of X_v . $\Rightarrow \Psi_{uv} = \emptyset$.

#P-Completeness

[Cook]

For NP-Completeness, we have 3-SAT as the canonical problem.

Def) #P: a function $f: \Sigma^* \rightarrow \mathbb{N}$ is in # if \exists a poly. time NDTM M s.t. # of accepting comps. of M on input x is exactly $f(x)$.

Def) #P-Complete: $f \in \#P$ and all $g \in \#P$ reduces to it in poly. time.

A "parsimonious" reduction preserves the # of solutions.

Fact) \exists many natural decision problems $\in P$ whose counting version is $\in \#P$ -Complete.

ex) DNF-SAT: $(x_1 \wedge x_2 \wedge \bar{x}_3) \vee (\dots)$ \rightarrow ORs of ANDs

Let φ be any boolean formula in 3-CNF. Then we can reduce

$$\#\text{SAT}(\varphi) = 2^n - \#\text{DNFSAT}(\neg \varphi) \Rightarrow \text{DNF counting is } \#P\text{-C.},$$

ex) Permanent: Let A be an $n \times n$ matrix. $\text{per}(A) = \sum_{\sigma \in S_n} \prod_{i=1}^n a_{i\sigma(i)}$.

Thm) [Valiant] Computing $\text{per}(A)$ for a 0-1 matrix A is $\#P$ -Complete.

Fact) $\text{per}(A) = \#$ of perfect matchings in G_A , the bipartite graph of A .

Proof of Thm: Observe that $\text{per}(A) = \#$ of weighted cycle covers of G'_A ,

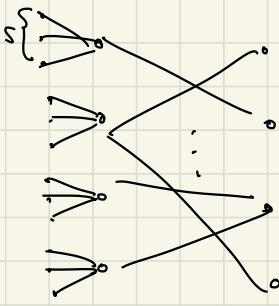
the directed graph of A . The goal is to take an arbitrary 3-CNF φ and construct a directed graph G_φ s.t. # of cycle covers of $G_\varphi = \chi(n,m) \cdot \#\text{SAT}(\varphi)$ where $n := \#$ of variables in φ , $m := \#$ of clauses in φ , and

χ is some poly. time pre-computable function. [Gadgets in slides/notes]

With appropriate gadget construction, $\alpha(m,n) := 4^{3^m}$. Lastly, make all integer edge weights into $\{\pm 1\}$ by splitting, and since $\text{per}(A) \in [N!, N!]$ where $N := \#$ of vertices in G_φ , we can actually compute $\text{per}(A) \bmod 2^{N^2+1}$ as $N! < 2^{N^2+1}$. Then, we can replace -1 by 2^{N^2} , and make them a series of N^2 edges of weight 2, and split.,

Thm) #MATCH(G) is #P-Complete. [For simplicity, assume G is bipartite]

Proof: Consider $Z_G(\lambda) = \sum_k M_k \lambda^k$ where $M_k := \#$ of k -edge matchings.



Consider an augmented graph G_s . Then, #MATCH(G_s)
 $= \sum_k M_k (s+1)^{n-k}$. If we have, say, $2n$ vertices,
we can evaluate this for $(n+1)$ distinct s and
compute all M_k by Lagrange Interpolation!

Thm') Evaluating $Z_G(\lambda) \in \#P$ -Complete $\forall \lambda > 0$.

Proof: weight each original edge by λ . Then we get $Z_{G_s}(\lambda) = ((s+\lambda))^n Z_G\left(\frac{\lambda}{s+\lambda}\right)$. This is the same hardness as above.,

Exact Counting

1) # Spanning Trees (Matrix - Tree Theorem)

2) FKT Algorithm for planar perfect matchings

"Holographic Algorithms" by Valiant

Laplacian
↓

Thm) (Kirchoff's) Matrix - Tree Theorem: $\# ST(G_i) = (1,1)$ minor of $L(G_i)$

* $L(G_i) = \text{diag}(d_i) - A$, $(1,1)$ minor is determinant with first row and column deleted. (all (i,i) minors of $L(G_i)$ are equal [Exercise])

ex) G_i :

$$\begin{bmatrix} 1 & -1 & 0 & 0 & 0 \\ -1 & 3 & -1 & 0 & 0 \\ 0 & -1 & 3 & -1 & -1 \\ 0 & -1 & -1 & 3 & -1 \\ 0 & 0 & -1 & -1 & 2 \end{bmatrix}$$

$ST(G) = 8 = \det(\boxed{\quad})$.

Proof: Let $B :=$ edge/vertex incidence matrix, e.g., for the above G ,

	$\{1,2\}$	$\{2,3\}$	$\{2,4\}$	$\{3,4\}$	$\{3,5\}$	$\{4,5\}$
1	+1					
2	-1	+1	+1			
3		-1		-	-	-
4			-1	-	-	-
5				-	-	-

$$L = BB^T$$
 [Exercise]. Then,

$$\det(L) = \sum_s \det(B_s) \det(B_s^T)$$
 where

B_s are square submatrices of B .

$$\rightarrow \det(L) = \sum_s \det(B_s)^2$$
. Similarly, $\det(L_{(1,1)}) = \sum_s \det(B'_s)^2$ where

$B' := B$ with first row removed. To prove the theorem, we need

to show that $\det(B'_s)^2 = \begin{cases} 1 & \text{if } S \text{ is a spanning tree} \\ 0 & \text{if } S \text{ is not a spanning tree.} \end{cases}$ Observe that

choosing some nonzero diagonal. If it is acyclic, it corresponds to a directed spanning tree rooted at vertex 1 and has $\det(B'_S) = \pm 1$.

On the other hand, if the subset S contains a cycle, we can construct a linear combination of the cycling edge columns s.t. they sum to 0, hence $\det(B'_S) = 0$. „ \rightarrow just greedily set all row sums to 0

(Corollary) We can uniformly sample spanning trees by fixing a random edge and finding the decision tree prob. of including it in the ST.

Alg) FKT Planar Perfect Matchings: $G(V, E)$ where $|V|$ is even $\xrightarrow{|V|=2k}$ # perf. match
 (Inspired by dimer coverings of monomer-dimer system of lattice)

Let $\pi := \{\{i_1, j_1\}, \dots, \{i_k, j_k\}\}$, a perfect matching. Let $A :=$ skew-symmetric matrix of G , $a_{ij} := \begin{cases} 1 & \text{if } \{i, j\} \in E \text{ and } i < j \\ -1 & \text{if } \{i, j\} \in E \text{ and } i > j \\ 0 & \text{if } \{i, j\} \notin E \end{cases}$. The Pfaffian of A , $Pf(A)$

$:= \sum_{\pi} \text{sgn}(\pi) \prod_{i=1}^k a_{i_2, j_2}$ where $\text{sgn}(\pi) :=$ sign of $\begin{bmatrix} 1 & 2 & \dots & (2k-1) & 2k \\ i_1 & j_1 & \dots & i_k & j_k \end{bmatrix}$. Then

$Pf(A) = \sum_M s(M)$ where M are perfect matchings and $s(M) := \text{sgn}(\pi_M) \prod_{i=1}^k a_{i_2, j_2}$.

Lemma: For any $2k \times 2k$ skew-symmetric matrix, $Pf(A)^2 = \det(A)$ [Optional Ex].

$\hookrightarrow Pf(A)$ can be calculated in polytime.

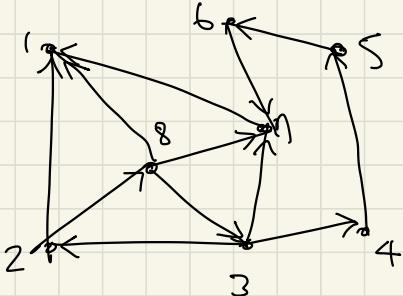
Lemma 2: A Pfaffian orientation of G is an orientation of its edges s.t. $s(M)$ has the same sign for all perfect matchings M . Any planar graph G

has a Pf orientation. * := # of edges of each orientation is odd.

Proof: Sps. \exists orientation in which every even cycle C in G for which $G \setminus V(C)$ contains a perfect matching is oddly oriented. Then, this is a Pf orientation. Let M_1, M_2 be any 2 pf. matchings. It is s.t.p. $s(M_1) \cdot s(M_2) = +1$. Take any cycle C in $M_1 \sqcup M_2$. By the condition, C must be oddly oriented. Analyzing, $s(M_1)s(M_2) = +1$. Now, sps. a planar drawing of G has an orientation in which the # of clockwise edges around (not necessarily even) face (except possibly the outer face) is odd. Then, this orientation is Pfaffian. This is a reduction to the previous more general statement. [In lecture notes, omitted here]

↳ Every planar graph has a Pf. orientation & can be found in polytime.
(Consider a ST and its dual ST of faces. Greedily orienting suffices.)

ex) Graph from Note 3:



	1	2	3	4	5	6	7	8
1	O	H						
2	H	O	H					
3		H	O	H				
4			H	O	H			
5				H	O	H		
6					H	O	H	
7						H	O	H
8							H	O

→ $\det(A) = 16$,
perfect
Matchings
 $= 4 = \sqrt{16}$

Approximations

$f: \Sigma^* \rightarrow \mathbb{R}^+$. What does it mean to approximate this function?

Def) Fully Polynomial (Randomized) Approximation Scheme (FPTAS/FPRAS):

algorithm A s.t. on inputs (x, ε) where $x \in \Sigma^*$, $\varepsilon \in (0, 1]$, outputs a (random) values $A(x, \varepsilon)$ s.t. $\Pr[(1+\varepsilon)^{-1}f(x) \leq A(x, \varepsilon) \leq (1+\varepsilon)f(x)] \geq 3/4$ and runs in time $\text{poly}(|x|, \varepsilon^{-1})$.

↪ we can boost $3/4$ with repeated trials and median finding [CS'74].

⇒ it suffices to run $t = O(\log \delta^{-1})$ trials to boost $3/4 \rightarrow 1 - \delta$!

ex) Let $f(G)$ be # of proper colorings of G with g colors. Sps. ∃ an algo. that approximates $f(G)$ within constant C . Consider $G^{(m)}$, the m -copied version of G . Then $f(G^{(m)}) = (f(G))^m$. If we set $C^m = 1 + \varepsilon$, $m = \frac{\log C}{\log(1+\varepsilon)} = O(\varepsilon^{-1})$. ⇒ arbitrary precision with polynomial blowup

Approximate Counting $\xleftarrow{\text{equiv.}}$ Random Sampling $\xrightarrow{\text{polytime alg. that outputs config. } \sigma \text{ from (close to) Gibbs distrib.}}$

For $Z(\lambda)$, we allow distribution $\hat{\pi}$ s.t. $\|\pi - \hat{\pi}\|_{TV} \leq \delta$.

Sampling → Counting Methods:

(1) ex) g -colorings, uniform. Define $\Omega := \text{set of proper } g\text{-coloring of } G$.

$|\Omega| = \frac{|\Omega_0|}{|\Omega_1|} \times \frac{|\Omega_1|}{|\Omega_2|} \times \dots \times \frac{|\Omega_{m-1}|}{|\Omega_m|} \times |\Omega_m|^l = \prod_{i=1}^m \frac{1}{z_i} \times |\Omega_m|^l$ where $\Omega_i :=$ set of proper q_i -colorings of G with ordered vertices $1, \dots, i$ pinned to some values and $Z_i := \frac{|\Omega_i|}{|\Omega_{i-1}|} \leq 1$. Then, each Z_i can be estimated by sampling u.a.r. from Ω_{i-1} and observing proportion \hat{Z}_i of colorings in which vertex i has coloring $c_i \leftarrow$ maximal proportion, to reduce variance.

Thm) Unbiased Estimator: Let X_1, \dots, X_n be iid nonneg. r.v.s with $\mu = E[X_i]$, $\sigma^2 = \text{Var}(X_i)$. Let $y = \frac{1}{t} \sum_{i=1}^t X_i$. Then $\Pr[y - \mu \geq \varepsilon\mu] \leq \frac{1}{4}$ provided $t \geq \frac{4\sigma^2}{\varepsilon^2 \mu^2}$. Proof: Chebyshov.

→ sample size of Z_i : $O(\frac{1}{\varepsilon^2} \times \frac{1}{z_i}) \rightarrow O(\frac{1}{z_i} \times q)$ since $\frac{1}{z_i} \leq q$ w.h.p.

* the sampling blackbox algo. should be able to sample with prior pinnings!

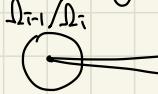
↳ this is not a big problem as long as the problem is "self-reducible".

finally, we actually need $\varepsilon \mapsto \frac{\varepsilon}{n}$ bound for each trial for $(1 + \frac{\varepsilon}{n})^n \approx 1 + \varepsilon$, so each Z_i actually needs $O(\frac{n^2}{\varepsilon^2} q)$ trials, and since we have n Z_i s, we need $O(\frac{n^3}{\varepsilon^2} q)$ samples (or by boosting, $O(\frac{n^2 \log n}{\varepsilon^2})$). q^n

(2) q -coloring again, hard constraints. $|\Omega| = \frac{|\Omega_m|}{|\Omega_{m-1}|} \times \frac{|\Omega_{m-1}|}{|\Omega_{m-2}|} \dots \times \frac{|\Omega_1|}{|\Omega_0|} \times |\Omega_0| = \prod_{i=1}^m \frac{1}{y_i} \times q^n$ where $\Omega_i :=$ set of q -colorings with only i edges present.

We can estimate $y_i = \frac{|\Omega_i|}{|\Omega_{i-1}|}$ by random sampling from Ω_{i-1} and observing proportion of them that is still consistent with i -th edge ($\sigma(u) \neq \sigma(v)$).

Claim: Assuming $\Omega_i \geq \Delta + 1$ ($\Delta := \max \deg. \text{ of } G_i$), $\frac{|\Omega_i|}{|\Omega_{i-1}|} \geq \frac{1}{2}$.

Proof:  $|\Omega_i| \geq |\Omega_{i-1}| - |\Omega_{i-1}| \Rightarrow \frac{|\Omega_i|}{|\Omega_{i-1}|} \geq \frac{1}{2}$. //

↪ the rest of analysis is equivalent to (1).

(3) $Z(\beta) = \sum_{\sigma} \exp(-\beta H(\sigma))$. WLOG, assume that $H: \sigma \rightarrow \mathbb{N}^+$ and that $|H(\sigma)| \leq n^c$ for constant c . Goal is estimating $Z(\beta)$.

Let $Z(\beta) = \frac{Z(\beta_t)}{Z(\beta_{t-1})} \times \frac{Z(\beta_{t-1})}{Z(\beta_{t-2})} \times \dots \times \frac{Z(\beta_1)}{Z(\beta_0)} \times Z(\beta_0)$ where $0 = \beta_0 < \beta_1 < \dots < \beta_t = \beta$.

Choose $\beta_i = \beta_{i-1} + \frac{1}{n^c}$. $\rightarrow t \sim n^{2c}$ since WLOG, $\beta \leq n^2$ (later justified).

$\frac{Z(\beta_i)}{Z(\beta_{i-1})} = \frac{1}{Z(\beta_{i-1})} \sum_{\sigma} \exp(-\beta_i H(\sigma)) = \frac{1}{Z(\beta_{i-1})} \sum_{\sigma} \exp(-\beta_{i-1} H(\sigma)) \exp((\beta_{i-1} - \beta_i) H(\sigma))$

$= \mathbb{E}_{\beta_{i-1}} [\exp((\beta_{i-1} - \beta_i) H(\sigma))]$, and this is just an observable.

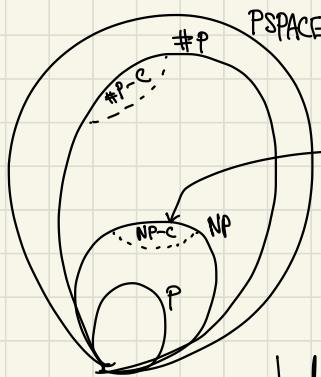
$|\beta_{i-1} - \beta_i| = \frac{1}{n^c}$, $|H(\sigma)| \leq n^c$, so the r.v. $\exp(\sim)$ lies in const. range $[e^{-1}, e]$.

$\rightarrow \beta \leq n^2$ is a reasonable assumption since $\beta \geq n^2$ is a trivial ground state.

Counting → Sampling: Consider the decision tree  where each layer picks a color for each vertex. Every leaf corresponds to a proper coloring. As long as the counting error is $\leq 1 + \frac{1}{n}$, error is $(1 + \frac{1}{n})^n \leq e$, a constant. Let's say we sampled σ . $P_\sigma := \Pr[\text{reach leaf } \sigma]$ is known. We output σ w.p. $[2eN P_\sigma]^{-1}$. Otherwise, reject & try again. $N := \text{approx (upto factor 2) \# of total colorings}$.

Claim: $1 \leq 2e^{\hat{N}} p_e \leq 4e^2$. Proof: use bounds above.

Thm) $\forall f \in \#P$, \exists a polytime randomized algorithm with an NP-oracle



that satisfies the FPRAS requirements.
 $\rightarrow \leq \varepsilon^{0.13^m}$

Algo: $\forall x \in \Sigma^*$, let $\Omega(x)$ be the set of accepting computations of TM M on x .

Repeatedly intersect $\Omega(x)$ with indep. uniform half-spaces in $\{0,1\}^m$. Define $S_t := \{z \in \Omega \mid z \cdot y_i = 0$

$(\bmod 2)$ for $1 \leq i \leq t\}$. Let $t^* := \min_t \{S_t = \emptyset\}$. Output 2^{t^*} as an estimate.

Analysis: Assume $|\Omega| = N$ lies in $2^{k-1} < N \leq 2^k$. Observe that

$|S_t|$ is a sum of N 0-1 r.v.s, each with prob 2^{-t} , pairwise indept.

$$\rightarrow E[|S_t|] = N2^{-t}, \text{Var}[|S_t|] = N2^{-t}(1-2^{-t}) \leq N2^{-t} = E[|S_t|].$$

Claim: (i) if $t \geq k+3$, $\Pr[|S_t| > 0] \leq \frac{1}{8}$. $\rightarrow E[|S_t|] = N2^{-t} \leq 2^k 2^{-(k+3)}$

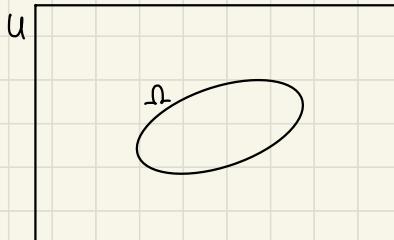
$= \frac{1}{8}$. By Markov's Inequality, $\Pr[|S_t| > 0] \leq \frac{1}{8}$.

(ii) if $t \leq k-4$, $\Pr[|S_t| = 0] \leq \frac{1}{8}$. $\rightarrow E[|S_t|] = N2^{-t} \geq 2^{(k-1)} 2^{-(k-4)} = 8$.

By Chebyshov's Inequality, $\Pr[|S_t| = 0] \leq \frac{\text{Var}[|S_t|]}{E[|S_t|]^2} \leq \frac{1}{E[|S_t|]} \leq \frac{1}{8}$.

$\Rightarrow 2^{t^*}$ lies within a factor of 16 of N w.o.p. $\geq 3/4$. Replicate r TMs and take the r -th root to find a suitable $(1+\varepsilon)$ factor.

Basic Monte Carlo



$$U := \{0,1\}^m, \Omega := \{z \in \{0,1\}^m \mid z \text{ is a "solution"}\}$$

Repeat t times: pick $z \in U$ u.a.r.

$$x_i = \mathbb{1}\{z \in \Omega\}. \text{Output } |U| \cdot \frac{\sum x_i}{t}.$$

Fact: # samples to get a $(1+\varepsilon)$ -approx. w.p. $\geq 3/4 \rightarrow t \geq \frac{4}{\varepsilon^2} \cdot \frac{1}{p}$ where $p = \frac{|\Omega|}{|U|}$.
 ↳ can be exp. large!

#DNF: ψ is OR of ANDs. DNF $\in P$, but #DNF $\in \#P$ -Complete.

Let $S_i :=$ set of assignments that satisfy the i -th term. Then, $|\Omega| = \prod_{i=1}^m |S_i|$.

$|S_i| = 2^{n-r_i}$ where $r_i :=$ # of variables in i -th term. We use a new universe

	a_1	a_2	-	.	-	-	-	a_n
s_1	\otimes				$\otimes \otimes$			\otimes
s_2		\otimes				\otimes		\times
:			\times					
s_n			\otimes			\times		\otimes
	\times	\times						

$U = \{(a, i) \mid a \text{ satisfies the } i\text{-th term}\}$.

→ Pick $(a, i) \in U$ u.a.r. Set $x_i :=$

$\mathbb{1}\{s_i \text{ is the lowest numbered term satisfied by } a\}$.

by $a\}$. $\Omega := \{(a, i) \mid s_i \text{ is the lowest term satisfied by } i\}$. Repeat t times,
 and output $|U| \cdot \frac{\sum x_i}{t}$.

Analysis: (i) $\frac{|\Omega|}{|U|} = \frac{\#\text{ of } \otimes \text{ in table}}{\#\text{ of } \times \text{ in table}} \geq 1/m$ since $\exists \otimes$ for every column.

→ Sample size is $t = O(m/\varepsilon^2)$.

(ii) Implementation is efficient: picking $(a, i) \in U$ is poly time, checking
 $(a, i) \in \Omega$ is also easy. → $O(nm)$ per sample.

→ total runtime $O(M^2/\epsilon^2)$, can improve to $(\frac{M}{\epsilon^2})$ [Karp/Luby].,,

* this extends to probabilistic DNF, $\Pr[\varphi \text{ is sat.}]$ when each $\Pr[X_i=T]=p_i$.

Network Reliability: Connected $G(V, E)$, fail pr. p_e for each edge. What is

$P_{\text{fail}} := \Pr[G \text{ becomes disconnected under random edge failures}]$?

→ $P_{\text{fail}} := \Pr[U S_i]$ where $S_i := i\text{-th cut fails}$. However, # of cuts is large.

* Karger's MinCut Algorithm: while $|V| \geq 2$, pick an edge $(u, v) \in E$ u.a.r., merge $u \& v$. Output the remaining cut. If C is a min-cut, $\Pr[C \text{ is output}]$

$\geq \frac{1}{\binom{n}{2}}$. [Analysis omitted, can be found in CS74 notes] $\Pr[C \text{ is output}] \geq \gamma \frac{1}{\binom{n}{2}}$

↪ Corollary: # of min-cuts $\leq \binom{n}{2}$. Also, # of α -min-cuts $\leq \binom{n}{2\alpha}$.

Ingredients for Algo: 1) probabilistic DNF Karp-Luby → FPRAS

2) Karger's Min-Cut → # of α -min-cuts $\leq \binom{n}{2\alpha} \leq n^{2\alpha} \rightarrow$ can be enumerated in poly time w.h.p. due to coupon collecting ($n^{2\alpha} \log(n^{2\alpha})$)

Algo) [Assume min-cut size is C , and $p_e = p \forall e$] ($G(V, E), p$) $\rightarrow P_{\text{fail}}$

- If $p^C \geq \frac{1}{n^4}$, then use basic Monte Carlo. (runtime $O(\epsilon^2 \cdot n^4)$)

- Else, set $\alpha := 2 + \frac{1}{2} \log_n(3/\epsilon)$.

↪ "absorbed" into $(1+\epsilon)P_{\text{fail}}$ error!

Claim*: $\Pr[\text{some cut size } \geq NC \text{ fails}] \leq \epsilon p^C \leq \epsilon P_{\text{fail}}$.

- Left with only α -min-cuts. Then, $\Pr[\text{some } \alpha\text{-min-cut fails}] = \Pr\left[\bigvee_{i=1}^t (X_{e_{i1}} \wedge X_{e_{i2}} \wedge \dots \wedge X_{e_{ir}})\right]$ where $X_{e_{ij}}$ is 1 w.p. p . and $t \leq \# \text{ of } \alpha\text{-min-cuts} \leq n^{2\alpha} = \frac{2n^4}{\epsilon}$.

↪ Use Karp-Luby to compute $(1+\epsilon)$ approx. in $\text{poly}(n, 1/\epsilon)$ time.

Proof of Claim^{*}: Enumerate the cuts of size $\geq \alpha c$ in increasing order of size

$$\alpha c \leq C_1 \leq C_2 \dots$$

(i) Consider the first $n^{2\alpha}$ of these cuts. \forall such cuts, $\Pr[\text{cut fails}] \leq p^{\alpha c}$.

$\rightarrow \Pr[\text{any of such cuts fails}] \leq n^{2\alpha} \cdot p^{\alpha c}$. By assumption, $p^c = n^{-(4+\delta)}$.

$$\rightarrow \leq n^{2\alpha} \cdot n^{-(4+\delta)\alpha} = n^{-(2+\delta)\alpha} \quad . \quad \begin{matrix} \uparrow \\ C_i = \text{size of } i\text{-th cut} \end{matrix}$$

(ii) If $\beta > 0$, # of β -min-cuts $\leq n^{2\beta}$. Hence, $C_{n^\beta} \geq \beta c \Rightarrow C_k \geq \frac{c}{2} \log n$

$$\rightarrow p^{c_k} \leq p^{\frac{c}{2} \log n} = n^{-(4+\delta)(\log n)/2} = k^{-(2+\delta/2)} \quad .$$

$\rightarrow \Pr[\text{any of remaining cuts fail}] \leq \sum_{k>n^{2\alpha}} k^{-(2+\delta/2)} \leq \int_{n^{2\alpha}}^{\infty} x^{-(2+\delta/2)} dx \leq n^{-(2+\delta)\alpha}$.

$\Rightarrow \Pr[\text{any of cuts size } \geq \alpha c \text{ fails}] \leq 2n^{-(2+\delta)\alpha}$. Substituting $\alpha = 2 + \frac{1}{2} \log(\frac{1}{\epsilon})$,

$$\leq \epsilon n^{(4+\delta)} = \epsilon p^c \leq \epsilon P_{\text{fail}} . . .$$

* Calculating $P_{\text{not fail}} = 1 - P_{\text{fail}}$ is actually not reducible to P_{fail} !

[Guo, Jerrum] came up with a different algorithm for this.

Markov Chain Monte Carlo

Input: implicit description of a very large set Ω , weight function $w: \Omega \rightarrow \mathbb{R}^+$.

Goal: sample $\sigma \in \Omega$ from distribution $\pi(\sigma) := \frac{w(\sigma)}{\sum_{\sigma'} w(\sigma')}$.

→ Equivalently, compute $\sum_{\sigma' \in \Omega} w(\sigma')$, the partition function.

Def) Markov Chain: sequence of r.v.s s.t. $\Pr[X_{t+1}=y | X_0, \dots, X_t] = \Pr[X_{t+1}=y | X_t]$.

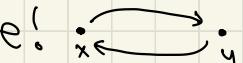
Then we can let $P_{ij} := \Pr[X_{t+1}=j | X_t=i]$. P is the transition matrix.

→ P is stochastic, i.e., $P_{xy} \geq 0$, $\sum_y P_{xy} = 1 \forall x$. [Denote $P_{xy} \equiv P(x,y)$]

Notation: $P_x^{(t)} :=$ distribution of X_t given that $X_0=x$.

→ $P_x^{(t+1)} = P_x^{(t)} P$, i.e., $P_x^{(t+1)}(u) = \sum_z P_x^{(t)}(z) P(z,u)$.

Def) Reversibility: ∃ distribution π s.t. $\pi(x)P(x,y) = \pi(y)P(y,x) \forall x,y$.

↪ think of the "flow" of $x \rightarrow y$ and $y \rightarrow x$ being same! 

↪ then we can represent P as an undirected graph, $Q(x,y) = \pi(x)P(x,y)$
(and finite)

Thm) Fundamental Theorem of MC: If P is irreducible and aperiodic,
then it has a unique stationary distribution $\pi > 0$; this is the
unique left eigenvector of P with eigenvalue 1. Also, $P^{(\ell)}(x,y) \xrightarrow{\ell \rightarrow \infty} \pi(y) \forall x,y$.

Def) Irreducibility: $\forall x, y, \exists t$ s.t. $P^{(t)}(x, y) > 0$.

Def) Aperiodic: $\forall x, y, \gcd\{t \mid P^{(t)}(x, y) > 0\} = 1$.

Fact: If $\overset{\leftarrow}{P}$ is reversible w.r.t. π , then $\pi > 0$ is the unique stationary dist.

Proof: $(\pi P)(y) = \sum_x \pi(x) P(x, y) = \sum_x \pi(y) P(y, x) = \pi(y) \sum_x P(y, x) \stackrel{1.}{=}$

Fact: If P is reversible, then P is similar to a symmetric matrix. Hence, all eigenpairs will be real and P is diagonalizable.

Proof: Let $D := \text{diag}(\sqrt{\pi(x)})$. Then $DPD^{-1}(x, y) = \frac{\sqrt{\pi(x)}}{\sqrt{\pi(y)}} P(x, y)$
 $= \frac{\sqrt{\pi(y)}}{\sqrt{\pi(x)}} P(y, x) = DPD^{-1}(y, x)$.

Thm) [Perron-Frobenius]: Any irreducible, aperiodic stochastic matrix P has an eigenvalue $\lambda_1 = 1$ with unique left eigenvector $e_1 > 0$ and all other eigenvalues λ_i satisfy $|\lambda_i| < 1$.

↳ This implies the Fund. Thm. of MC since we can represent $p^{(0)} = \sum \alpha_i e_i$,

$P^{(t)} = p^{(0)} P^t = \sum \alpha_i \lambda_i^t e_i \xrightarrow{t \rightarrow \infty} \alpha_1 e_1 = e_1$ ($\alpha_1 = 1$ b/c conservation of "mass").

* λ_2 dictates how fast $P^{(t)}$ converges to e_1 , and the quantity $(1 - \lambda_2)$ is called the spectral gap.

ex1) Random walk on undirected graph $G(V, E)$, $\Omega = V$.

at state $u \in V$, pick v u.a.r. from $\text{neighbor}(u)$ and transition to v [Exercise!]

- irreducible iff G is connected, aperiodic iff G is not bipartite.

- reversible w.r.t. $\pi(x) = \frac{\deg(x)}{2|E|}$.

ex2) Card Shuffling, $\Omega = \text{set of permutations of } n \text{ cards}$.

$\pi(\sigma) = 1$ (uniform dist.)

→ $i=j$ is allowed

(i) Random Transposition: pick two positions i, j . switch them.

- irreducible & aperiodic. reversible w.r.t. π as actions are symmetric.

(ii) Top-in-at-Random: Take top card, move to a uniform random position.

- irreducible & aperiodic. reversible? → no, it's not symmetric. However,

we can observe that P is doubly stochastic. → uniform π must be stationary.

(iii) Gilbert-Shannon-Reeds Riffle: split deck into $L \& R \sim Bi(n, \frac{1}{2})$.

drop cards from L/R w.p. $\frac{|L|}{|L|+|R|} / \frac{|R|}{|L|+|R|}$ (equiv. to uniform interleaving).

- irreducible? think of L being only the top card. reduces to (ii). ✓

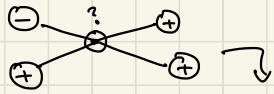
- aperiodic? think of L being empty. ✓

- not reversible, but doubly stochastic (consider the "inverse riffle").

ex3) Glauber Dynamics, spin system on $G(V, E)$, spins = $\{1, \dots, q\}$.

Hamiltonian $H(\sigma)$, Gibbs $\pi(\sigma) \propto \exp(-\beta H(\sigma))$. $\Omega = \{\sigma: V \rightarrow [q]\}$.

In config σ , pick vertex $v \in V$ u.a.r., resample $\sigma(v)$ from correct conditional dist. $\pi(\cdot | \sigma_{V \setminus v})$.



Consider Ising, $\pi(\sigma) \propto \lambda^{\# \text{dis}(\sigma)}$, $\lambda = \exp(-2\beta)$. $\rightarrow \sigma(v) \propto \begin{cases} \lambda^3 & (-) \\ \lambda & (+) \end{cases}$

- irreducible? trivial since we can set any v to anything. (no hard constraints)

\hookrightarrow (in general, this is not true. consider q -coloring on q -clique.)

- aperiodic? trivial.

- reversible? $\pi(\sigma)P(\sigma, \tau) = \pi(\tau)P(\tau, \sigma) \rightarrow \frac{\pi(\sigma)}{\pi(\tau)} = \frac{P(\cdot, \sigma)}{P(\cdot, \tau)}$. \checkmark

ex 4) Metropolis Process. Ω , $w: \Omega \rightarrow \mathbb{R}^+$, sample $\pi(\sigma) \propto w(\sigma)$.

construct an undirected connected graph on Ω ("neighborhood structure")

define a proposal dist. $K(x, \cdot)$ on $N(x)$, neighbors of x s.t. $K(x, y) > 0 \forall y \in N(x)$.

at state x , pick $y \in N(x)$ w.p. $K(x, y)$.

go to y w.p. $\min\left\{1, \frac{w(y)K(y, x)}{w(x)K(x, y)}\right\}$. (trivially irred. & aper.)

- reversible w.r.t. π ? take neighbors x, y , $w(y)K(y, x) \leq w(x)K(x, y)$.

$$w(x)P(x, y) = \cancel{w(x) \cdot K(x, y)} \cdot \frac{w(y)K(y, x)}{\cancel{w(x)K(x, y)}} = w(y)K(y, x) = w(y)P(y, x).$$

ex 5) Swendsen-Wang Algorithm. q -state Potts model, $[q]$.

$\pi(\sigma) \propto \lambda^{\# \text{dis}(\sigma)}$, $\lambda = \exp(-2\beta)$. in config σ , delete all disagreeing (bichromatic)

edges. in the remaining subgraph, retain each edge w.p $p = (1 - \lambda)$, otherwise delete it. assign uniformly random spins to each remaining connected components. go to the resulting configuration.

[Exercise] Check that SW dynamic is reversible w.r.t. π .

↳ consider all intermediate configurations after the $p = (1 - \lambda)$ thinning.

Coupling

Def) Total Variance Distance: For two prob. dist. μ, ν on Ω ,

$$\|\mu - \nu\|_{\text{TV}} := \frac{1}{2} \sum_{\sigma \in \Omega} |\mu(\sigma) - \nu(\sigma)| = \max_{A \subseteq \Omega} |\mu(A) - \nu(A)|.$$

Def) Coupling: For any two prob. dist. μ, ν on Ω , ω on $(\Omega \times \Omega)$ s.t.

$$\mu(x) = \sum_y \omega(x, y) \quad \forall x, \quad \nu(y) = \sum_x \omega(x, y) \quad \forall y.$$

Thm) Coupling Lemma: Let X, Y be RVs with dist. μ, ν on Ω . Then,

$$(i) \Pr[X \neq Y] \geq \|\mu - \nu\|_{\text{TV}}, \quad (ii) \exists \text{ coupling of } (\mu, \nu) \text{ s.t. } \Pr[X \neq Y] = \|\mu - \nu\|_{\text{TV}}.$$

Proof: Consider the coupling $\Pr[X=Y=z=\sigma] = \min\{\mu(\sigma), \nu(\sigma)\}$.

Setup of Fundamental Thm of MC: P is an irreducible & aperiodic MC, and assume P has a stationary distribution π . Let $\Delta_x(t) = \|P_x^{(t)} - \pi\|_{\text{TV}}$,

$$\Delta(t) = \max_x \Delta_x^{(t)}, D_{xy}(t) = \|P_x^{(t)} - P_y^{(t)}\|, D(t) = \max_{x,y} D_{xy}^{(t)}. \text{ Then,}$$

$\Delta(t) \leq D(t) \leq 2\Delta(t)$ by lower bound and TI.

Claim 1: $\Delta_x(t)$ is non-increasing with t .

Proof: Let $X_0 = x, Y_0 \sim \pi$. Fix t and couple X_t, Y_t s.t. $\Pr[X_t \neq Y_t] = \|P_x^{(t)} - \pi\|_\pi$.

Couple (X_{t+1}, Y_{t+1}) as follows: if $X_t = Y_t, X_{t+1} = Y_{t+1}$. Else, evolve independently.

$$\Delta_x^{(t+1)} = \|P_x^{(t+1)} - \pi\| \leq \Pr[X_{t+1} \neq Y_{t+1}] \leq \Pr[X_t \neq Y_t] = \Delta_x^{(t)}.$$

Claim 2: $\forall s, t \in \mathbb{N}, D(s, t) \leq D(s)D(t)$.

Proof: Let $X_0 = x, Y_0 = y$. Couple $P_x^{(t)}, P_y^{(t)}$ s.t. $\Pr[X_t \neq Y_t] = \|P_x^{(t)} - P_y^{(t)}\|_\pi$.

Then, if $X_t = Y_t$, set $X_{t+i} = Y_{t+i}$ for $i = 1, \dots, s$. Else, let $x' = X_t, y' = Y_t$,

where $x' \neq y'$. Use coupling lemma to couple distributions X_{t+s}, Y_{t+s}

$$\text{conditioned on } X_t = x', Y_t = y' \text{ s.t. } \Pr[X_{t+s} \neq Y_{t+s} \mid X_t = x', Y_t = y'] = \|P_{x'}^{(s)} - P_{y'}^{(s)}\|_\pi$$

$$= D_{x'y'}(s) \leq D(s). \text{ Then, } D_{xy}(s+t) = \|P_x^{(s+t)} - P_y^{(s+t)}\| \leq \Pr[X_{t+s} \neq Y_{t+s}]$$

$$\leq D(s)D_{xy}(t) \leq D(s)D(t).$$

Claim 3: $D(t) < 1$ for some finite t .

Proof: Irreducibility & Aperiodicity $\Rightarrow P^t(x, y) > 0 \quad \forall x, y \in \Omega$ for some finite t .

Proof of Fund. Thm. of MC: $\|P_x^{(t)} - \pi\|_\pi \xrightarrow{t \rightarrow \infty} 0$ suffices. Let t be s.t.

$$D(t) = 1 - \delta \text{ for some } \delta > 0. \quad \forall k > 0, \Delta(kt) \leq D(kt) \leq D(t)^k \leq (1-\delta)^k.$$

$\Rightarrow \Delta(t) \rightarrow 0$ as $t \rightarrow \infty$, so $\Delta_x(t) \rightarrow 0$. //

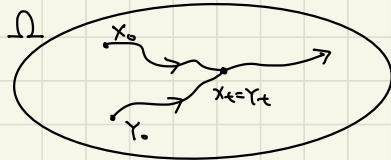
Let $T_x(\varepsilon) := \min\{t \mid \Delta_x(t) = \varepsilon\}$, $T(\varepsilon) := \max_x T_x(\varepsilon)$, $T_{\text{mix}} := T(\frac{1}{2e})$.

Claim: $\Delta(t) \leq \exp(-\lfloor \frac{t}{T_{\text{mix}}} \rfloor)$.

Proof: $\Delta(kT_{\text{mix}}) \leq D(kT_{\text{mix}}) \leq D(T_{\text{mix}})^k \leq (2\Delta(T_{\text{mix}}))^k \leq e^{-k}$. //

Corollary: $T(\varepsilon) \leq T_{\text{mix}} \cdot \lceil \ln(\varepsilon^{-1}) \rceil$.

Coupling for Mixing Times



Def) Coupling for MC: a pair process (X_t, Y_t)

s.t. (i) if $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$, (ii) $\forall t, a, b,$

$$\Pr[X_{t+1} = b \mid X_t = a] = \Pr[Y_{t+1} = b \mid Y_t = a] = P(a, b).$$

Def) $T_{xy} := \min\{t \mid X_t = Y_t, X_0 = x, Y_0 = y\}$ under a given coupling.

Claim: $\Delta(t) \leq \max_{x,y} \{\Pr[T_{xy} > t]\}$.

Proof: $\Delta(t) \leq D(t) = \max_{x,y} \|P_x^{(+)}/P_y^{(+)} - 1\|_\infty \leq \max_{x,y} \Pr[X_t \neq Y_t \mid X_0 = x, Y_0 = y]$

$$= \max_{x,y} \Pr[T_{xy} > t]. //$$

Corollary: for any coupling, $T_{\text{mix}} \leq \max_{x,y} \Pr[T_{xy} > \frac{1}{2e}] \leq \max_{x,y} 2e E[T_{xy}]$.

ex1) Random Walk on cube $\{0,1\}^n$. $|\Omega| = 2^n$. Make the graph lazy.
(self-loop w.p. $1/2$)

If moving, pick $i \in [n]$ and flip the i -th bit. Equivalently, pick $i \in [n]$ and $b \in \{0, 1\}$ u.a.r., and set $x_i \leftarrow b$.

Coupling: $X_t \& Y_t$ both use the same $i \& b \rightarrow d(X_t, Y_t)$ is nonincreasing.

$E[\text{time until } d(X_t, Y_t) = 0] \leq \text{coupon-collecting } n \text{ coupons} = n \ln n + \dots$

$$\Rightarrow T_{\text{mix}} \leq n \ln n + O(n), T_{\text{mix}} = \frac{1}{2} n \ln n.$$

ex2) Top-in-at-Random Shuffle: Symmetric, so inverse path is isomorphic to our real path. Analyzing the inverse is easier.

Coupling: (X_t, Y_t) , choose a random card (not position!). Then, the top k cards that are ordered cannot be disturbed.

$$\Rightarrow \text{again coupon-collecting} \rightarrow T_{\text{mix}} = n \ln n + \dots \quad //$$

ex3) Random Transposition Shuffle: Pick position $i \& \text{ card } c$. Exchange c with card at position i .

$$d_t =$$

Coupling: (X_t, Y_t) chooses same $i \& c$. $d(X_t, Y_t) := \# \text{ of disagreements}$.

(i) if c is in same position in $X_t \& Y_t$, $d_{t+1} = d_t$.

(ii) else, if card at position i is same, $d_{t+1} = d_t$. else, $d_{t+1} \leq d_t - 1$.

$\Rightarrow d_t$ is non-increasing. $\Pr[d_{t+1} \leq d_t - 1] = \left(\frac{d_t}{n}\right)^2 \rightarrow T_{\text{mix}} = O\left(\sum_{t=1}^n \left(\frac{1}{d_t}\right)^2\right)$

$$(* \text{ tight } T_{\text{mix}} \text{ is } \frac{1}{2} n \ln n + \dots) = O(n^2).$$

Graph Coloring

Connected Undirected $G(V, E)$, # of colors q , assume max degree Δ .

↪ Sample from uniform distribution over proper q -colorings of G .

Conjecture: Standard Glauber dynamics on q -colorings has mixing time $\text{poly}(n)$ (actually $O(n \log n)$) provided that $q \geq \Delta + 2$.

Glauber Dynamics: pick $v \in V, c \in Q$ u.a.r. recolor if possible.

Claim: G.D. is connected (irreducible) if $q \geq \Delta + 2$.

Proof Sketch: we can always greedily recolor every vertex.

Claim: When $q \geq 4\Delta + 1$, mixing time for G.D. is $O(n \log n)$.

Proof: Coupling. Both X_t & Y_t choose the same v and c .

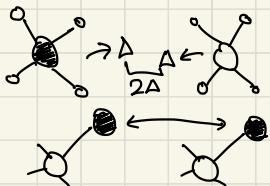
Distance $d(X_t, Y_t) :=$ # of disagreeing vertices.

"Good" move ($d(X_{t+1}, Y_{t+1}) = d(X_t, Y_t) - 1$) when v is a disagreeing vertex & color c is not in $N(v)$ for both X_t & Y_t .

"Bad" move ($d(X_{t+1}, Y_{t+1}) = d(X_t, Y_t) + 1$) when v is an agreeing vertex & color c is in $N(v)$ of either X_t or Y_t but not both.

Other moves are neutral (distance stays the same).

At some state, # good moves $\geq d_t(q - 2\Delta)$.



bad moves $\leq \Delta d_t \cdot 2$. $\Rightarrow q \geq 4\Delta$ ensures that

of good moves $\geq d_t(q - 2\Delta) \geq d_t(2\Delta) \geq \# \text{ bad moves} !$

$$\Rightarrow E[d_{t+1} | X_t, Y_t] \leq d_t - \frac{d_t(q - 2\Delta)}{q n} + \frac{2d_t \Delta}{2} = d_t \left(1 - \frac{q - 4\Delta}{qn}\right).$$

$$\rightarrow E[d_t | X_0, Y_0] \leq \left(1 - \frac{q - 4\Delta}{qn}\right)^t \stackrel{\text{upper bound on } d_0}{=} \left(1 - \frac{1}{\alpha n}\right)^t n \text{ where } \alpha := \frac{q}{q - 4\Delta}.$$

$$Pr[d_t > 0 | X_0, Y_0] = Pr[d_t \geq 1 | X_0, Y_0] \leq E[d_t | X_0, Y_0] \leq \left(1 - \frac{1}{\alpha n}\right)^t n$$

$$\leq \varepsilon \text{ if we take } t = C_\varepsilon \left(\frac{q}{q - 4\Delta}\right) n \ln n. \Rightarrow T_{\text{mix}} = \mathcal{O}(n \ln n). //$$

Path Coupling: a pre-metric on Ω is a connected graph with the edge weights s.t. every edge is a shortest path. Say x, y are adjacent if \exists edge (x, y) .

Thm) [Bubley, Dyer]: Sps. \exists a coupling $(X, Y) \rightarrow (X', Y')$ only defined on pairs (X, Y) that are adjacent in pre-metric, and s.t.

* $E[d(X', Y') | X, Y] \leq (1 - \alpha) d(X, Y)$ for some $\alpha \in [0, 1]$ where d is the metric defined by the pre-metric. Then, this coupling can be extended to a coupling on all pairs (X, Y) also satisfying the above condition *.

Proof: For arbitrary (X, Y) , let $X = Z_0 \rightarrow \dots \rightarrow Z_k = Y$ be a

shortest path in the pre-metric. $E[d(X', Y') | X, Y] \leq E\left[\sum_{i=0}^{k-1} d(Z'_i, Z'_{i+1}) | X, Y\right] = \sum_{i=0}^{k-1} E[d(Z'_i, Z'_{i+1}) | Z_i, Z_{i+1}] \leq (1-\alpha) \sum_{i=0}^{k-1} d(Z_i, Z_{i+1})$

 $= (1-\alpha) d(X, Y).$

For colorings, we now extend Ω to allow invalid colorings. However, there are no transitions into invalid colorings.

Claim: we can get down to $q \geq 2\Delta + 1$ with $O(n \log n)$ mixing time.

Proof: we are now considering X, Y that differ in exactly only v_0 .

Consider the coupling st. if $v \in N(v_0)$, then couple $\text{color}(v_0^X), \text{color}(v_0^Y)$.

"Good" moves when $v = v_0, c \notin \text{Color}(N(v_0)) \rightarrow \# \text{good moves} \geq 1 \times (q - \Delta)$

"Bad" moves when $v \in N(v_0), c = (\text{color}(v_0^Y), \text{color}(v_0^X)) \rightarrow \# \text{bad moves} \leq \Delta$.

$\Rightarrow T_{\text{mix}} = O(n \log n)$ provided that $q \geq 2\Delta + 1$ by path coupling. //

Thm [Vigoda]: $q \geq \frac{11}{6}\Delta + 1$ is sufficient.

Thm [Hayes, Vigoda, et Al]: $q \geq 1.76\Delta + 1$ is suff. if $\Delta = \Omega(\log n)$ and G is triangle-free.

Proof Sketch: # available colors $\geq q - \Delta$ is too pessimistic. Let $A(X, v) := \# \text{available colors at } v \text{ in coloring } X$. Suff. to have $A(X, v) \geq \Delta$.

Consider a uniformly random coloring & assume $N(v)$ are independent.

$$\rightarrow E[A(X_v)] = q \left(1 - \frac{1}{q}\right)^{\Delta} \propto q e^{-\Delta/q} > \Delta \text{ provided that } q \geq \alpha \Delta \text{ where } \alpha \text{ is unique solution to } x = e^x \rightarrow x \approx 1.76.$$

" $q \geq \Delta + 3$ is suff. provided that girth is a large constant dependent on Δ ".

For $q < \Delta$, it is NP-Hard to even approximate # of q -colorings.

For $q \leq \frac{\Delta}{2}$, it is "decide if \exists a q -coloring.

Dobrushin Conditions

Def) Influence Matrix: for a spin system on G with Gibbs dist. π ,

$R := (R_{ij})$, an $n \times n$ matrix, s.t. $R_{ij} := \max_{\sigma_i, \tau \in S_j} \{ \|\pi_h(\sigma, \cdot) - \pi_h(\tau, \cdot)\|_{TV} \}$

and where $S_j := \{(\sigma, \tau) \mid \sigma, \tau \text{ agree everywhere except } j\}$

Def) Spectral (L_2 Operator) Norm: $\|A\| := \sup_{v \neq 0} \frac{\|Av\|_2}{\|v\|_2}$

Thm) [Hayes]: If $\|R\| \leq (1-\delta)$ for some $\delta > 0$, then the mixing time of Glauber dynamics is $\leq \frac{C}{\delta} n \ln n$ for some constant C .

Proof: $(X_t), (Y_t)$. Coupling - at every step, X_t & Y_t choose the same vertex i & use a maximal coupling for the update, i.e.,

$$\Pr[X_{t+1}(i) \neq Y_{t+1}(i) \mid X_t, Y_t] = \|\pi_i(X_t, \cdot) - \pi_i(Y_t, \cdot)\|.$$

\rightarrow Write $p_t(i) := \Pr[X_t(i) \neq Y_t(i) \mid X_0, Y_0]$. coupling argument!

Then $p_{t+1}(i) \leq (1 - \frac{1}{n})p_t(i) + \frac{1}{n} \sum_{j \in N(i)} p_{ij} p_t(j)$. [Exercise]

Let $A := \frac{n-1}{n} I + \frac{1}{n} R$. Then $p_{t+1} \leq A p_t \Rightarrow p_t \leq A^t p_0$. union bound

$\|A\| \leq \frac{n-1}{n} \|I\| + \frac{1}{n} \|R\| \leq (1 - \frac{\delta}{n})$. Now, $\Pr[X_t \neq Y_t] \leq \|p_t\|_1$

$\leq \sqrt{n} \|p_t\|_2 \leq \sqrt{n} \|A^t p_0\|_2 \leq \sqrt{n} \|A^t\| \|p_0\|_2 \leq \sqrt{n} (1 - \frac{\delta}{n})^t \cdot \sqrt{n}$

$= n (1 - \frac{\delta}{n})^t \Rightarrow T_{\text{mix}} = O(\frac{1}{\delta} \ln n)$. //

ex1) Hard-Core Model: G , spins := $\{0, 1\}$. Ω = ind. sets of G .

$\pi(\sigma) \propto \lambda^{|\sigma|}$. If some $j \in N(i)$ is occupied, $\Pr[i \text{ occ.}] = 0$. Else,

$\Pr[i \text{ occ.}] = \frac{\lambda}{1+\lambda}$. $\rightarrow p_{ij}$, worst case is when all $j \in N(i) - i$ are unocc.,

so $p_{ij} = \frac{\lambda}{1+\lambda}$. $\rightarrow R = \frac{\lambda}{1+\lambda} A_{G_i, \text{matrix of } G}^{\leftarrow \text{adjacency}}$ $\rightarrow \|R\| = \frac{\lambda}{1+\lambda} \|A_G\| = \frac{\lambda}{1+\lambda} \cdot \lambda_o(G)$. \downarrow max eigenv.

We know that $\lambda_o(G) \leq \Delta \rightarrow \|R\| \leq \frac{\lambda}{1+\lambda} \Delta \rightarrow$ Setting $\lambda \leq \frac{1-\delta}{\Delta}$,

we get $\|R\| \leq (1 - \delta)$. //

ex2) Ferromag. Ising Model: Spins $\{+, -\}$, $\pi(\sigma) \propto \lambda^{\# \text{Dis}(\sigma)}$.

$\Pr[\sigma(i) = +] = \frac{\lambda^d}{\lambda^{d+} + \lambda^{d-}}$ where $d^\pm := \# \text{ of neighbors with sign } \pm$.

Worst case influence is when $\deg(i)$ is odd & $|d^+ - d^-| = 1$.

→ This gives a bound on $P_{ij} \leq \frac{\lambda^{-\frac{1}{2}} - \lambda^{\frac{1}{2}}}{\lambda^{-\frac{1}{2}} + \lambda^{\frac{1}{2}}} = \tanh(\beta/2)$ where $\lambda = \exp(-\beta)$.

So $R \leq \tanh(\beta/2) A_0 \Rightarrow \|R\| \leq \tanh(\beta/2) \cdot \lambda_0(G) \leq \tanh(\beta/2) \cdot \Delta$

⇒ Rapid mixing provided $\tanh(\beta/2) \leq \frac{1-\delta}{\Delta}$.

ex3) Linear Extensions: Partial order \preceq on n elements. Sample uniformly/
count # of linear extensions of \preceq to a total order. (#P-Complete)

Dynamics: pick a position $p \in [n-1]$. Exchange $(p, p+1)$ if legal.

[Exercise: Check irreducibility (diameter $\binom{n}{2}$)
(w.p. $\frac{1}{2}$, do nothing)]

Pre-metric: σ, τ are adjacent iff they differ at exactly $i \neq j$, $i < j$.

$$d(\sigma, \tau) := j - i.$$

Path Coupling: Case 1) $j \neq i+1$. both σ, τ pick the same p & attempt.

Case 2) $j = i+1$. both σ, τ pick the same p . if $p \neq i$, make the same move.

if $p = i$, couple doing nothing in σ & making a move in τ , vice versa.

Analysis: Case I) $p \notin \{i-1, i, j-1, j\}$. moves in σ, τ are identical.
 \hookrightarrow no change.

Case II) $p = i-1$ or $p = j$. Symmetrical, so WLOG assume $p = i-1$.

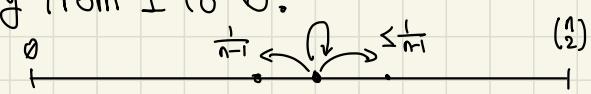
Worst outcome is d increasing by 1 w.p. $\leq 2 \cdot \frac{1}{2(n-1)}$.

Case III) $p = i$ or $p = j-1$. if $j \neq i+1$, then d decreases by 1 since

every such attempted move must be legal. $\rightarrow \Pr[\text{decrease}] = 2 \cdot \frac{1}{2(n-1)}$.

if $j=i+1$, then d decreases surely from 1 to 0.

→ we are doing a random walk



⇒ $O(n^5 \log n) \approx O(n^5)$ for high prob. to hit 0.

* Tight bound is $O(n^3 \log n)$ [Wilson] (For any partial ordering!)

↳ If we use p sampling w.p. $\propto p(n-p)$, we get this tight bound.

Functional Analysis of MCMC

Motivation: view the MC as an operator on its state space.

Let P be the transition matrix of an ergodic finite MC, state space Ω , and its stationary distribution π .

Def) Expectation & Variance: A real-valued function $\varphi: \Omega \rightarrow \mathbb{R}$,

$$E_\pi[\varphi] := \sum_{x \in \Omega} \pi(x) \varphi(x), \quad \text{Var}_\pi[\varphi] := \sum_{x \in \Omega} \pi(x) \cdot (\varphi(x) - E_\pi[\varphi])^2$$

The transition matrix P is an operator on φ , i.e. $P\varphi(x) := \sum_{y \in \Omega} P(x,y) \varphi(y)$.

↳ $P\varphi(x)$ is the "one-step average" of φ starting at x .

Similarly, $P^t \varphi(x) := \sum_{y \in \Omega} P^t(x,y) \varphi(y)$ is the " t -step average".

Obs) π is stationary, so $E_\pi[P^t \varphi] = E_\pi[\varphi] \ \forall t$, and as $t \rightarrow \infty$,

$P^t \varphi$ converges to $E_\pi[\varphi]$. $\Rightarrow \text{Var}_\pi[P^t \varphi] \rightarrow 0$ as $t \rightarrow \infty$!

The convergence of MC can be linked to convergence of $\text{Var}_{\pi}[\mathbb{P}^t \varphi]$!

Def) Dirichlet Form: for functions $\varphi, \psi: \Omega \rightarrow \mathbb{R}$, $\mathcal{E}_p(\varphi, \psi) := \langle \varphi, L \psi \rangle_{\pi}$

where Laplacian $L := I - P$ and inner product $\langle f, g \rangle_{\pi} := \sum_x \pi(x) f(x) g(x)$
 (also $E_{\pi}[fg]$)

The symmetric Dirichlet form $\mathcal{E}_p(\varphi, \varphi) := \langle \varphi, L \varphi \rangle_{\pi}$

$$\begin{aligned} &= \sum_x \pi(x) \varphi(x) [(I - P)\varphi(x)] = \sum_x \pi(x) \varphi(x) \sum_y (I(x, y) - P(x, y)) \varphi(y) \\ &= \sum_{xy} \pi(x) \varphi(x) (I(x, y) - P(x, y)) \varphi(y) = \sum_x \pi(x) \varphi(x)^2 - \sum_{xy} \pi(x) \varphi(x) P(x, y) \varphi(y) \\ &= \frac{1}{2} \sum_{xy} \pi(x) P(x, y) (\varphi(x)^2 + \varphi(y)^2) - \sum_{xy} \pi(x) \varphi(x) P(x, y) \varphi(y) \\ &= \underbrace{\frac{1}{2} \sum_{xy} \pi(x) P(x, y)}_{\geq \frac{1}{2} \text{ self-loop } \forall x} (\varphi(x) - \varphi(y))^2 \end{aligned}$$

Comparing to the variance $\text{Var}_{\pi}[\varphi] := \sum_x \pi(x) (\varphi(x) - E_{\pi}(x))^2$

$$\begin{aligned} &= \sum_x \pi(x) \varphi(x)^2 - E_{\pi}[\varphi]^2 \quad (\text{use variance formula } \text{Var}(x) = E[x^2] - E[x]^2) \\ &= \frac{1}{2} \sum_{xy} \pi(x) \pi(y) (\varphi(x)^2 + \varphi(y)^2) - \sum_x \pi(x) \varphi(x) \sum_y \pi(y) \varphi(y) \\ &= \underbrace{\frac{1}{2} \sum_{xy} \pi(x) \pi(y)}_{\geq \frac{1}{2} \text{ self-loop } \forall x} (\varphi(x) - \varphi(y))^2. \end{aligned}$$

Obs) $\mathcal{E}_p(\varphi, \varphi)$ and $\text{Var}_{\pi}[\varphi]$ are similar, except that $\mathcal{E}_p(\varphi, \varphi)$ is "local".

Def) Poincare Constant: $\alpha := \inf_{\varphi \text{ non-const}} \frac{\mathcal{E}_p(\varphi, \varphi)}{\text{Var}_{\pi}[\varphi]}$.

Thm) For any $\overset{\geq \frac{1}{2} \text{ self-loop } \forall x}{\text{lazy}}$, ergodic P and any initial state $x \in \Omega$,

$$T_x(\varepsilon) \leq \frac{1}{\alpha} (2 \ln(\varepsilon)^{-1} + \ln(4 \pi(x)^{-1})).$$

Proof: The proof of the theorem relies on the following lemma:

~~Lemma: $\forall \varphi: \Omega \rightarrow \mathbb{R}, \text{Var}_{\pi}[P\varphi] \leq \text{Var}[\varphi] - \mathcal{E}_p(\varphi, \varphi)$. [proof deferred]~~

↳ This immediately gives a $(1-\alpha)$ contraction at every step.

Corollary: \forall non-constant $\varphi: \Omega \rightarrow \mathbb{R}, \text{Var}_{\pi}[P^t \varphi] \leq (1-\alpha)^t \text{Var}_{\pi}[\varphi]$.

Relating $\text{Var}_{\pi}[P^t \varphi]$ to variation distance, we get by Cauchy-Schwarz

$$\begin{aligned} \|P_x^{(t)} - \pi\|_{TV}^2 &= \left(\frac{1}{2} \sum_{\sigma} |P_x^{(t)}(\sigma) - \pi(\sigma)| \right)^2 = \frac{1}{4} \left(\sum_{\sigma} \pi(\sigma) \left| \frac{P_x^{(t)}(\sigma)}{\pi(\sigma)} - 1 \right| \right)^2 \\ &\leq \frac{1}{4} \left(\sum_{\sigma} \pi(\sigma) \left(\frac{P_x^{(t)}(\sigma)}{\pi(\sigma)} - 1 \right)^2 \right) \left(\sum_{\sigma} \pi(\sigma) (1)^2 \right)^{-1} \\ &= \frac{1}{4} \left(\sum_{\sigma} \pi(\sigma) \left(\frac{P_x^{(t)}(\sigma)}{\pi(\sigma)} - E_{\pi}\left[\frac{P_x^{(t)}}{\pi}\right] \right)^2 \right) \left[E_{\pi}\left[\frac{P_x^{(t)}}{\pi}\right] = \sum_{\sigma} \pi(\sigma) \frac{P_x^{(t)}(\sigma)}{\pi(\sigma)} = 1 \right] \\ &= \frac{1}{4} \text{Var}\left[\frac{P_x^{(t)}}{\pi}\right]. \end{aligned}$$

Now define $\varphi := \frac{P_x^{(t)}}{\pi}$. Note that $P^t \varphi \neq \frac{P_x^{(t)}}{\pi}$.

Def) Time Reversal: $P^*(x,y) := \frac{\pi(y)}{\pi(x)} P(y,x)$. Some properties:

(i) P^* is ergodic and has the same π .

$$(P^*)^*(x) = \sum_y \pi(y) P^*(y,x) = \sum_y \pi(y) \frac{\pi(x)}{\pi(y)} P(x,y) = \sum_y P(x,y) \pi(x) = \pi(x).$$

$$(ii) (P^*)^* = P. \text{ Let } Q := P^*. Q^*(x,y) = \frac{\pi(x)}{\pi(y)} P^*(y,x) = \frac{\pi(x)}{\pi(y)} \frac{\pi(y)}{\pi(x)} P(x,y).$$

$$(iii) P^* = P \text{ iff } P \text{ is reversible. } P(x,y) = \frac{P(y)}{P(x)} P(y,x) = P^*(x,y) \quad \forall x,y.$$

$$(iv) \mathcal{E}_{P^*}(\varphi, \varphi) = \mathcal{E}_p(\varphi, \varphi) \quad \forall \varphi. \quad \mathcal{E}_{P^*}(\varphi, \varphi) = \frac{1}{2} \sum_{xy} \pi(x) P^*(x,y) (\varphi(x) - \varphi(y))^2$$

$$= \frac{1}{2} \sum_{xy} \pi(x) \frac{\pi(y)}{\pi(x)} P(y,x) (\varphi(x) - \varphi(y))^2 = \frac{1}{2} \sum_{yx} \pi(y) P(y,x) (\varphi(y) - \varphi(x))^2$$

$$= \frac{1}{2} \sum_{xy} \pi(x) P(x,y) (\varphi(x) - \varphi(y))^2.$$

\Rightarrow The Poincaré constant for P and P^* are the same!

$$\text{Obs: } P^{*t} \varphi = P^{*t} \frac{P_x^{(0)}}{\pi} = \frac{P_x^{(0)} P^t}{\pi} = \frac{P_x^{(t)}}{\pi}$$

$$\begin{aligned} \text{Proof: } & \left(P^{*t} \frac{P_x^{(0)}}{\pi} \right) (y) = \sum_{z \in \Omega} P^* (y, z) \frac{P_x^{(0)}(z)}{\pi(z)} = \sum_z \frac{\pi(z)}{\pi(y)} P(z, y) \frac{P_x^{(0)}(z)}{\pi(z)} \\ & = \sum_z P(z, y) \frac{P_x^{(0)}(z)}{\pi(y)} = \frac{P_x^{(0)}(y)}{\pi(y)}. \text{ Induct on } t \text{ to get } P^{*t} \varphi = \frac{P_x^{(t)}}{\pi}. \end{aligned}$$

$$\Rightarrow 4 \|P_x^{(t)} - \pi\|_{\text{TV}}^2 \leq \text{Var}_{\pi} \left[\frac{P_x^{(t)}}{\pi} \right] = \text{Var}_{\pi} [P^{*t} \varphi] \leq (1-\alpha)^t \text{Var} [\varphi].$$

$$\text{Since } \text{Var}_{\pi} [\varphi] = \frac{1}{\pi(x)} - 1 \leq \frac{1}{\pi(x)}, \text{ setting } t = \frac{1}{\alpha} (2 \ln \varepsilon^{-1} + \ln(4\pi(x))^{-1})$$

$$\text{ensures } (1-\alpha)^t \text{Var}_{\pi} [\varphi] \leq 4\varepsilon^2 \Rightarrow \|P_x^{(t)} - \pi\|_{\text{TV}} \leq \varepsilon.$$

$$\begin{aligned} \text{Proof of Lemma*}: & P \text{ is lazy, so } P := \frac{1}{2}(I + \hat{P}) \text{ where } \hat{P} \text{ is} \\ & \text{stochastic \& has same } \pi. \text{ Then, } [P\varphi](x) = \sum_y P(x, y) \varphi(y) \\ & = \frac{1}{2} \varphi(x) + \frac{1}{2} \sum_y \hat{P}(x, y) \varphi(y) = \frac{1}{2} \sum_y \hat{P}(x, y) (\varphi(x) + \varphi(y)). \end{aligned}$$

WLOG, assuming that $E_{\pi} [\varphi] = 0$ (shifting $\varphi(x)$ doesn't affect $\text{Var}_{\pi} [\varphi]$),

$$\begin{aligned} \text{Var}_{\pi} [\varphi] &= \sum_x \pi(x) ([P\varphi](x))^2 = \frac{1}{4} \sum_x \pi(x) \left(\sum_y \hat{P}(x, y) (\varphi(x) + \varphi(y)) \right)^2 \\ &\leq \frac{1}{4} \sum_x \pi(x) \hat{P}(x, y) (\varphi(x) + \varphi(y))^2 \text{ (due to Cauchy-Schwarz).} \end{aligned}$$

$$\begin{aligned} \text{Also, } \text{Var}_{\pi} [\varphi] &= \frac{1}{2} \sum_x \pi(x) \varphi(x)^2 + \frac{1}{2} \sum_y \pi(y) \varphi(y)^2 \\ &= \frac{1}{2} \sum_{xy} \pi(x) \varphi(x)^2 \hat{P}(x, y) + \frac{1}{2} \sum_{xy} \pi(x) \hat{P}(x, y) \varphi(y)^2 \\ &= \frac{1}{2} \sum_{xy} \pi(x) \hat{P}(x, y) (\varphi(x)^2 + \varphi(y)^2). \end{aligned}$$

$$\text{Taking differences, } \text{Var}_{\pi} [\varphi] - \text{Var}_{\pi} [P\varphi] \geq \frac{1}{4} \sum_{xy} \pi(x) \hat{P}(x, y) (\varphi(x) - \varphi(y))^2.$$

Observe that entries in \hat{P} are twice as large as P except for diagonal entries. However, the diagonals disappear in this sum, so the RHS = $\frac{1}{2} \sum_{xy} \pi(x) \hat{P}(x,y) (\varphi(x) - \varphi(y))^2 = \mathbb{E}_P(\varphi, \varphi)$.
 $\Rightarrow \text{Var}_{\pi}[\varphi] - \text{Var}_{\pi}[P\varphi] \geq \mathbb{E}_P(\varphi, \varphi).$

If P is reversible ($\pi(x)P(x,y) = \pi(y)P(y,x)$), then an alternative proof is given by spectral graph theory.

\rightarrow Recall that if P is reversible, $S := DPD^{-1}$ where $D := \text{diag}(\sqrt{\pi(x)})$ is non-negative & symmetric \Rightarrow ev of P is same as S & all are real.

By Perron-Frobenius, spectrum of P is $| = \lambda_1 > \lambda_2 \geq \dots \geq \lambda_n > -1$.

Then, the initial distribution $p^{(0)} = \sum_i \alpha_i \overset{\text{evecs}}{\downarrow} e_i$ where $e_i = \pi$, $\alpha_i = 1$.

$\rightarrow p^{(0)} P^t = \sum_i \alpha_i \lambda_i^t e_i$, and since $|\lambda_i| < 1 \quad \forall i > 1$, $p^{(0)} P^t \xrightarrow{t \rightarrow \infty} \pi$,

and the slowest decay is given by $\max_{i>1} |1 - \lambda_i|$. If P is lazy,

$\lambda_i \geq 0 \quad \forall i$, so the rate of convergence is given by spectral gap ($1 - \lambda_2$).

Claim: Ergodic, reversible, lazy NC P and $\forall x \in \Omega$, [Exercise!]

$\Delta_x(t) \leq \frac{\lambda_2^t}{2\sqrt{\pi(x)}}$, hence $T_x(\varepsilon) \leq \frac{1}{1-\lambda_2} (\ln \varepsilon^{-1} + \frac{1}{2} \ln (4\pi(x))^{-1})$.

Claim: for an ergodic & reversible P , $1 - \lambda_2 = \inf_{\varphi \text{ non-const}} \frac{\mathbb{E}_P(\varphi, \varphi)}{\text{Var}_{\pi}[\varphi]}$

Proof: $L := I - P$ has evals $\mu_i = (1 - \lambda_i)$, so $0 = \mu_1 < \mu_2 \leq \dots \leq \mu_n < 2$

with the same evecs. So the spectral gap of P is just μ_2 .

Suppose P is symmetric, then μ_1 minimizes the Raleigh Quotient,

$$\mu_1 = \inf_{\psi \neq 0} \frac{\langle \psi, L \psi \rangle}{\langle \psi, \psi \rangle} \text{ where } \langle \psi, \psi \rangle := \sum_x \psi(x) \psi(x).$$

This is given by the principal evec $\vec{1}$. Now, $\mu_2 = \inf_{\psi \perp \vec{1}} \frac{\langle \psi, L \psi \rangle}{\langle \psi, \psi \rangle} = \inf_{\psi \perp \vec{1}} \frac{\langle \psi, L \psi \rangle}{\langle \psi, \psi \rangle - \langle \psi, \vec{1} \rangle^2}$.

Extending to reversible P , apply this to $(I - DPD^{-1})$ to obtain

$$\mu_2 = \inf_{\psi \perp \vec{1}} \frac{\langle \psi, L \psi \rangle_\pi}{\langle \psi, \psi \rangle_\pi} = \inf_{\psi \perp \vec{1}} \frac{\langle \psi, L \psi \rangle}{\langle \psi, \psi \rangle - \langle \psi, \vec{1} \rangle^2} \quad [\text{Exercise!}] \text{ where } \langle \psi, \psi \rangle_\pi$$

is defined the same as before. Then, this is just $\frac{\text{Var}_\pi[\psi]}{\text{Var}_\pi[\psi] - \langle \psi, \vec{1} \rangle^2} \cdot //$

Multicommodity Flow

To estimate the Poincaré constant α , we incorporate flow into the MC.

Here are some basic definitions for an ergodic MC P with sd. π :

Capacity/Ergodic Flow: $C(e) := \pi(z)P(z, z')$ \forall directed $e = (z, z')$.

\hookrightarrow represents the flow of probability mass on e when stationary

Demand: $D(x, y) := \pi(x)\pi(y) \quad \forall (x, y) \in (\Omega \times \Omega)$.

Flow: $f: P \rightarrow \mathbb{R}^+ \cup \{0\}$ where $P := \bigcup_{x, y} P_{xy}$ and P_{xy} is the set of all simple paths from x to y s.t. $\sum_{p \in P_{xy}} f(p) = D(x, y) \quad \forall (x, y) \in (\Omega \times \Omega)$.

Cost of flow: $c(f) := \max_e f(e)/C(e)$ where $f(e) := \sum_{p \ni e} f(p)$.

Length of flow: $l(f) := \max_{p: f(p) > 0} |p|$, length of longest flow-carrying path.

Thm) + ergodic P, + flow f for P, $\propto \geq \frac{1}{p(f)l(f)}$.

Proof: start from symmetrized variance equation and using $D(x,y)$,

$$2\text{Var}_{\pi}[\varphi] = \sum_{xy} \pi(x)\pi(y)(\varphi(x) - \varphi(y))^2 = \sum_{xy} \sum_{p \in P_{xy}} f(p)(\varphi(x) - \varphi(y))^2.$$

Use the following telescoping sum over paths: $\varphi(x) - \varphi(y) = \sum_{u \in \text{dep}} (\varphi(u) - \varphi(v))$.

$$\Rightarrow 2\text{Var}_{\pi}[\varphi] = \sum_{xy} \sum_{p \in P_{xy}} f(p) \left(\sum_{u \in \text{dep}} (\varphi(u) - \varphi(v))^2 \right)^2 \quad [\text{by Cauchy-Schwarz}]$$

$$\leq \sum_{xy} \sum_{p \in P_{xy}} f(p) \cdot |p| \cdot \sum_{u \in \text{dep}} (\varphi(u) - \varphi(v))^2 \quad [\text{switch order of summations}]$$

$$= \sum_{e \in \text{edges}} (\varphi(u) - \varphi(v))^2 \sum_{p \ni e} f(p) \cdot |p| \leq l(f) \sum_{e \in \text{edges}} (\varphi(u) - \varphi(v))^2 \sum_{p \ni e} f(p).$$

[Recall that $\sum_{p \ni e} f(p) = f(e)$ and $p(f) = \max_e f(e) / C(e)$]

$$\leq l(f) p(f) \sum_{e \in \text{edges}} (\varphi(u) - \varphi(v))^2 C(e) \quad [\text{substitute } C(e) = \pi(u)P(u,v)]$$

$$= l(f) p(f) \sum_{e \in \text{edges}} (\varphi(u) - \varphi(v))^2 \pi(u)P(u,v) = 2l(f)p(f)\mathcal{E}_p(\varphi, \varphi).$$

$$\Rightarrow \text{Var}_{\pi}[\varphi] \leq l(f)p(f)\mathcal{E}_p(\varphi, \varphi) \Rightarrow \propto = \inf_{\varphi \text{ non-const}} \frac{\mathcal{E}_p(\varphi, \varphi)}{\text{Var}_{\pi}[\varphi]} \geq \frac{1}{l(f)p(f)}. //$$

Corollary: +ergodic, lazy MC P and +flow f, the mixing time ($\min_{x \in E} \pi(x)$)

$$\tau_{\text{mix}}(\varepsilon) \leq l(f)p(f) \left(2 \ln \varepsilon^{-1} + \ln(4\pi(x))^{-1} \right) \Rightarrow \tau_{\text{mix}} = O(l(f)p(f) \ln(\frac{1}{\varepsilon}))$$

$\hookrightarrow \tau_{\text{mix}}$ is essentially bounded by $p(f)$ since $l(f) \downarrow$ $\ln(\pi_{\min}) \downarrow$ are usually small.
 \downarrow diameter of MC \downarrow π uniform, then $O(n)$

ex1) Hypercube: lazy random walk on $\{0,1\}^n$. Let $N = 2^n$ be #vertices.

$$\pi(x) = \frac{1}{N}, C(u,v) = \pi(u)P(u,v) = \frac{1}{N} \cdot \frac{1}{2^n} = \frac{1}{2N} \quad \forall u,v \in E, D(x,y) = \frac{1}{N^2}.$$

Intuitively, flow maximizing $\frac{1}{pc(f)l(f)}$ spreads across all shortest paths of (x,y) .

Using the symmetry of the hypercube, $f(e) = f(e')$ $\forall e, e' \in E$, so

$$f(e) = \frac{\sum_{e \in E} f(e)}{|E|} = \frac{\frac{1}{N^2} \sum_{x,y \in V} \text{dist}(x,y)}{N_n} = \frac{n/2}{N_n} = \frac{1}{2N}$$

where we use the fact that the average distance $\frac{1}{N^2} \sum_{x,y} \text{dist}(x,y)$ is $n/2$ for hypercubes.

$$\Rightarrow pc(f) = \max_{e \in E} f(e) / C(e) = \frac{1}{2N} \cdot 2N_n = n. \text{ All flows take shortest paths, so } l(f) = n. \text{ Thus, } T_{\text{mix}} = O(pc(f)l(f) \ln \pi_{\text{min}}^{-1}) = \underline{O(n^3)}.$$

↪ This is not tight, since we know that $T_{\text{mix}} \sim \frac{1}{2} n \ln n$. The slack is induced by the heavy-duty approximations used in this method.

* for hypercubes, α is actually $\sim \frac{1}{n}$, so it is still not tight ($O(n^2)$).

ex2) Random Walk on a Line: MC for lazy walk on line $\{1, 2, \dots, N\}$,

Self-loop w.p. $1/2$ except endpoints w.p. $3/4$. π is uniform $\rightarrow \pi(x) = \frac{1}{N}$.

$$C(e) = \pi(x)P(x,y) = \frac{1}{4N} \text{ } \forall \text{ non self-loop edges, } D(x,y) = \frac{1}{N^2} \text{ } \forall x,y.$$

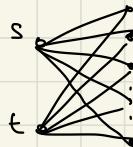
(max in middle)

There is a unique path for each pair, $f(i, i+1) = i(N-i) \frac{1}{N^2} \leq \frac{1}{4}$ is unique.

Then the cost is $pc(f) = \max_e \frac{f(e)}{C(e)} = \frac{1/4}{1/(4N)} = N$, and length is $l(f) = N$.

Thus, $\alpha \geq \frac{1}{pc(f)l(f)} \geq \frac{1}{N^2} \Rightarrow T_{\text{mix}} = O(N^2 \log N)$, tight to $\log N$ factor.

ex3) Random Walk on $K_{2,N}$: lazy random walk w.p. $1/2$, small side is st.


 Stationary distribution is $\pi(s) = \pi(t) = \frac{1}{4}$, $\pi(x) = \frac{1}{2N} \forall x \neq s, t$.
 $C(e) = \frac{1}{8N} \forall e$, $D(s,t) = \frac{1}{16}$, $D(s,x) = D(t,x) = \frac{1}{8N}$, $D(x,y) = \frac{1}{(AN)^2}$.

If we send flow through one shortest path, then $s \rightarrow t \& t \rightarrow s$ has $\frac{1}{16}$ units, so $f(f) \geq \frac{1/16}{\frac{1}{8N}} = \frac{N}{2}$, and $l(f) = 2 \Rightarrow T_{\text{mix}} = O(N)$, which is a bad bound since true mixing time is $\Theta(1)$ due to trivial coupling.

If we distribute flow across all shortest paths, then

$$\max_e f(e) \leq \frac{1}{8N} + \frac{1}{16N} + \frac{1}{4N^2} \cdot \frac{N-1}{2} \leq C \times \frac{1}{N} \text{ for constant } C.$$

$$\Rightarrow f(f) = \frac{C/N}{1/8N} = 8C, \text{ so mixing time from } s \text{ or } t \text{ is } O(1), \text{ tight.}$$

Flow Encoding: how to count paths & calculate flows in general MCs

Let $|\Omega| = N$ be the size of ergodic, lazy MC, where N is exponential in n , the natural measure of problem size. Assume π is uniform ($\pi(x) = \frac{1}{N}$),

$P(x,y) \geq \frac{1}{\text{poly}(n)}$ & non-zero transitions, i.e. degree is not too large. Then,

$C(u_N) = \pi(u) \cdot P(u,N) \geq \frac{1}{N \text{poly}(n)}$. To get a polytime T_{mix} , we need f st.

$l(f) \leq \text{poly}(n)$ and $\frac{f(e)}{C(e)} \leq \text{poly}(n) \forall e$. Hence, we need $f(e) \leq \frac{\text{poly}(n)}{N}$.

On the other hand, # edges $\leq N \times \text{poly}(n)$, and total flow is $\sum_{xy} \frac{1}{N^2} \approx 1$.

Thus, some edge must carry at least $\frac{1}{N \times \text{poly}(n)}$ flow, $f(e) \geq \frac{1}{N \times \text{poly}(n)}$.

\Rightarrow any good flow is optimal upto a polynomial factor.

Suppose flow $x \rightarrow y$ goes along a single path γ_{xy} . Let $\text{paths}(e)$ be the set of paths through edge e under f . Then $f(e) = |\text{paths}(e)| \times \frac{1}{N^2}$.

Since we want $f(e) \leq \text{poly}(n)/N$, we need $|\text{paths}(e)| \leq \text{poly}(n) \cdot |\Omega|$.

Multiple paths from x to y is the same for averaged values, and π being non-uniform is same except everything is weighted accordingly.

\Rightarrow we need to compare $|\text{paths}(e)|$ with $N = |\Omega| \rightarrow$ injective mapping!

Def) Encoding: for a flow f , a set of functions $\eta_e: \text{paths}(e) \rightarrow \Omega$ s.t.

(i) η_e is injective, (ii) $\forall \beta \leq \text{poly}(n), \pi(x)\pi(y) \leq \beta \pi(z)\pi(\eta_e(x,y)) \forall x,y \in \text{paths}(e)$

\hookrightarrow (i) says η_e is an injection, and (ii) says η_e is in addition weight-preserving upto poly factor β . If π is uniform, $\beta=1$ satisfies trivially.

Claim: If \exists encoding for f as above, then $f(f) \leq \beta \max_{P(z,z') > 0} \frac{1}{P(z,z')}$.

Proof: Let $e=(z,z')$ be an arbitrary edge. Then $f(e) = \sum_{(x,y) \in \text{paths}(e)} \pi(x)\pi(y)$

$\leq \beta \sum_{(x,y) \in \text{paths}(e)} \pi(z)\pi(\eta_e(x,y)) \leq \beta \pi(z)$, using properties (i) and (ii). Finally,

$C(e) = \pi(z)P(z,z')$, so $\frac{f(e)}{C(e)} \leq \beta \cdot \frac{1}{P(z,z')} //$

ex) Hypercube revisited: use encodings, not symmetry of hypercube.

$C(e) = \frac{1}{2^n N}, D(x,y) = \frac{1}{N^2}$. Consider f that sends $x \rightarrow y$ flow through a "left-right bit fixing path" γ_{xy} . Then, $l(f) = n$ trivially.

Consider an arbitrary edge $e = (z, z')$ where they differ in bit i .

Consider any pair $(x, y) \in \text{paths}(e)$. Observe that y agrees with z in the first $(i-1)$ bits and x agrees with z in the last $(n-i)$ bits.

We define $\eta_e(x, y) := x_1 x_2 \dots x_i y_{i+1} y_{i+2} \dots y_n$ st. $\eta_e(x, y)$ is the string that agrees with x on first i bits and with y on the rest. (missing information!)

→ We can recover a unique pair (x, y) from $\eta_e(x, y)$ and e , so η_e is an injection. Since π is uniform, it is trivially weight-preserving.

→ η_e is a valid encoding! So $p(f) \leq \max_{z, z'} \frac{1}{P(z, z')} \leq 2n$.

Up to a constant, we recover the same $T_{\text{mix}} = O(n^3)$ bound.

Matchings

$G(V, E)$, \mathcal{M} := set of matchings in G . $\pi_\lambda(M) := \frac{\lambda^{|M|}}{Z(\lambda)}$, $Z(\lambda) := \sum_k m_k \lambda^k$

where $m_k := \# \text{ matchings of size } k$. Assume $\lambda \geq 1$.

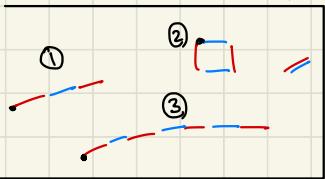
MC: in a matching M , w.p. $1/2$ do nothing. Else, pick $e = (u, v)$ u.a.r.

If $u \& v$ are both unmatched, match them ($M \mapsto M + e$). If $e \in M$,

$M \mapsto M - e$ w.p. $1/\lambda$. If exactly one of u/v is matched, go to $M + e'$ where e' is the edge touching u/v . Else, do nothing.

How to get from $x \rightarrow y$, $x, y \in M$? The union $x+y$ is alternating paths.

Order the alt. paths with some ordering. "Unwind" each path in order.



If we consider some transition in $x \rightarrow y$, we can infer (x, y) from z and some "fill-in" $\eta_e(x, y)$.

This is injective, and $\pi(x)\pi(y) \leq \lambda^2 \pi(z)\pi(\eta_e(x, y))$. Also,

$$\min P(z, z') = \frac{1}{2|E|\lambda} \rightarrow T_{\text{mix}} = O(\lambda^2 \cdot |E|\lambda \cdot |V| \cdot \log(\pi(x_0))).$$

But we can always find a maximal matching X_0 , and $\pi(x_0) \geq \frac{1}{\# \text{matchings}} \geq \frac{1}{2|E|}$.

$$\Rightarrow T_{\text{mix}} = O(\lambda^3 |V| |E|^2).$$

Now, we have a black-box $\xrightarrow{G, \lambda} A$ and A runs in $\text{poly}(G, \lambda)$.

1. Approximating $Z(\lambda)$: $Z(\lambda) = \frac{Z(\lambda_r)}{Z(\lambda_{r-1})} \cdots \frac{Z(\lambda_1)}{Z(\lambda_0)} \cdot Z(\lambda_0)$ where $\lambda_0 = 0$

$$< \lambda_1 < \cdots < \lambda_r = \lambda \text{ and } \lambda_i = \lambda_{i-1} \left(1 + \frac{1}{n}\right). \frac{Z(\lambda_i)}{Z(\lambda_{i-1})} = E_{\lambda_{i-1}} \left[\left(\frac{\lambda_i}{\lambda_{i-1}}\right)^{|M|} \right]$$

$$\leq \left(1 + \frac{1}{n}\right)^n \leq e. \rightarrow \text{Low variance, good approximation.}$$

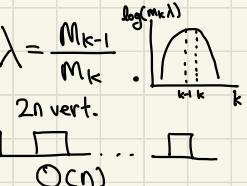
2. Approximating coeffs. $\{m_k\}$: Use the fact that $\{m_k\}$ is log-concave,

i.e. $m_{k+1} m_{k-1} \leq m_k^2 \forall k$ [Exercise, use injection?]. Then $\log(m_k \lambda^k)$ is concave,

so we can adjust its peak. $M_k = \frac{m_k}{m_{k-1}} \times \cdots \times \frac{m_1}{m_0} \times M_0$, each term can be approx-

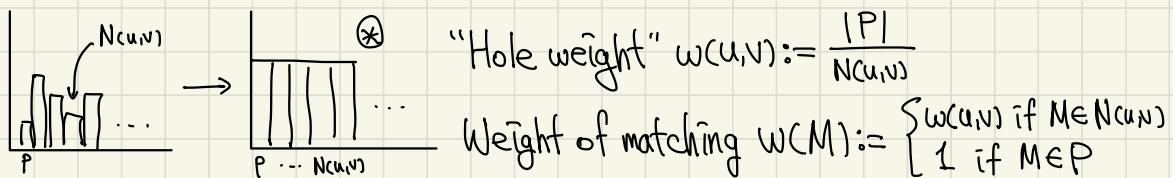
imated by finding λ s.t. distribution peaks at $M_{k-1} \& M_k \rightarrow \lambda = \frac{M_{k-1}}{M_k}$.

*Problem in graphs where $\frac{m_{n-1}}{m_n} \gg \text{poly}(n)$. e.g.,



where $M_n = 1$ but $M_{n-1} = \exp(c_n)$ due to each square being indep!

Idea: consider a modified MC with $\Omega = M_{n-1} \cup M_n$ and uniform π .



Using Metropolis, we can get the distribution \otimes as the new stationary.

$\rightarrow T_{\text{mix}} = O(\text{poly}(n))$, π puts weight at $\Omega(1/n^2)$ on P .

... except for the fact that we don't know $\frac{|P|}{Nc(u,v)}$ values!

Introduce edge weights $\{\lambda_e\}_{e \in [n] \times [n]}$. Initialize $\lambda_e \leftarrow 1 \forall e$.

Reduce λ_e for $e \in E$ down to $\ll \frac{1}{n!}$.

Each iteration, we reduce $\lambda_e \rightarrow \lambda_e/2$ for one $e \notin E$. $\rightarrow \# \text{iters} \text{ is poly}(n)$.

Then $\lambda(M) = \prod_{e \in M} \lambda_e$ will be negligible $\forall M \in \Omega$ ($\exists e \in M$ s.t. $e \notin E$).

Let $\lambda(P) = \sum_{M \in P} \lambda(M)$, $\lambda(Nc(u,v)) = \sum_{M \in Nc(u,v)} \lambda(M)$.

Hole weights $w_\lambda(u,v) := \frac{\lambda(P)}{\lambda(Nc(u,v))}$. $w(M) = \begin{cases} w_\lambda(u,v) & \text{if } M \in Nc(u,v) \\ 1 & \text{if } M \in P \end{cases}$.

For current values $\{\lambda_e\}$, now cut $\lambda_e \rightarrow \lambda_e/2$.

Claim: By observing the output of the MC with new edge weights $\{\lambda_e\}$,

we can learn the required hole weights $\frac{\lambda(P)}{\lambda(Nc(u,v))}$ w/ arbitrary accuracy.

Proof Sketch: Current MC has π s.t. satisfying $\pi(Nc(u,v)) = \frac{\lambda(Nc(u,v)) w(Nc(u,v))}{Z}$ ↳ can calculate using flows

$$\pi(P) = \frac{\lambda(P)}{\sum} \rightarrow w'(u,v) = \frac{\pi(N(u,v))}{\lambda(N(u,v))} \times \frac{\lambda(P)}{\pi(P)}.$$

Recall that $w_\lambda(u,v) = \frac{\lambda(P)}{\lambda(N(u,v))} = w'(u,v) \times \frac{\pi(P)}{\pi(N(u,v))}$. Induction complete. //

Ferromagnetic Ising Model



$$\sigma = \{\pm 1\}^V, w(\sigma) = \exp\left(\beta \sum_{(u,v) \in E} \sigma(u)\sigma(v) + \beta h \sum_{u \in V} \sigma(u)\right).$$

$$\pi(\sigma) = \frac{w(\sigma)}{Z(\beta, h)} \text{ where } Z(\beta, h) = \sum_{\sigma} w(\sigma).$$

Glauber dynamics: pick $v \in V$ u.a.r., reassign $\sigma(v)$ according to conditional distribution given its neighbors.

Thm) For a $\sqrt{n} \times \sqrt{n}$ box in \mathbb{Z}^2 , when $h=0$, T_{mix} of Gl. dyn. is:

$O(n \log n)$ if $\beta < \beta_c$, $\exp(O(\sqrt{n}))$ if $\beta > \beta_c$, $O(\text{poly}(n))$ if $\beta = \beta_c$.

↪ how do we deal with the supercritical ($\beta > \beta_c$) case?

Idea: Transform Ising spin configs $\Omega \rightarrow$ "subgraphs-world" configs Ω' where $Z(\beta, h) = C_{\beta, h} Z'(\eta, \xi)$ where $\eta = \tanh(\beta)$, $\xi = \tanh(\beta h)$.

Subgraphs world: $\Omega' = \{A \mid A \subseteq E\}$. $w'(A) = \eta^{|A|} \xi^{\lfloor \# \text{odd deg.}(A) \rfloor}$

$$Z'(\beta, h) = \sum_A w'(A), Z(\beta, h) = \sum_{\sigma} w(\sigma).$$

$$w'(A) = \eta^6 \xi^6 \rightarrow \dots$$

Claim: $Z(\beta, h) = C_{\beta, h} Z'(\eta, \xi)$ where $C_{\beta, h} = (\cosh \beta)^{|E|} \cdot (2 \cosh(\beta h))^{|V|}$.

Thm) "Glauber dynamics" of subgraphs word has $T_{\text{mix}} = O(\text{poly}(n, \xi^{-1}))$,

(forall graphs, $\# \beta$).

\rightarrow if $|h| > \varepsilon > 0$, then $T_{\text{mix}} = O(\text{poly}(n)) \wedge \beta \cdot (Z(\beta, h) \approx Z(\beta, 0) \text{ for } \beta \ll h)$.

Proof of Claim: Use $\exp(x) = \cosh(x)(1 + \tanh(x))$.

$$\begin{aligned} Z(\beta, h) &= (\cosh \beta)^{|E|} (\cosh \beta h)^{|V|} \sum_{\sigma \in \{0,1\}^V} \prod_u ((1 + \tanh(\beta \sigma(u) \sigma(v))) \prod_v (1 + \tanh(\beta h \sigma(u))) \\ &= 2^{-|V|} C_{\beta, h} \sum_{\sigma} \prod_u ((1 + \sigma(u) \sigma(v) \tanh(\beta)) \prod_v ((1 + \sigma(u) \tanh(\beta h))) \\ &= 2^{-|V|} C_{\beta, h} \sum_{\sigma} \left(\sum_{A \subseteq E} \prod_{\substack{u \in V \\ e \in A}} \sigma(u) \cdot \sigma(v) \right) \left(\sum_{u \leq v} \sum_{u \in u} \sigma(u) \right) \\ &= 2^{-|V|} C_{\beta, h} \sum_{A \subseteq E} \sum_{u \in V} \sum_{\sigma} W(A, u, \sigma) \text{ where } W(A, u, \sigma) := \prod_{e \in A} \sigma(u) \sigma(v) \prod_{v \in u} \sigma(u). \end{aligned}$$

Observe that $\sum_{\sigma} W(A, u, \sigma) = \begin{cases} 0 & \text{if } u \neq \text{odd}(A) \\ 2^{|V|} & \text{if } u = \text{odd}(A) \end{cases} //$

Entropy & Log-Sobolev Constant

Motivation: Stronger T_{mix} bounds than Poincare constant! (log vs log-log)

CTMC: P irreducible, heat kernel $H_t(x, y) = \Pr[X_t=y | X_0=x]$, computed as

$$H_t(x, y) := \sum_{k=0}^{\infty} \frac{e^{tL}}{k!} P^k(x, y) = \exp(-tL) \text{ where } L = I - P \text{ is Laplacian.}$$

* $\exp(A) = I + A + \frac{1}{2!} A^2 + \dots$ is a Taylor expansion on matrices.

Lemma: Let P be irreducible (not necessarily aperiodic) MC with st.dist. π .
not needed in cont. time!

$\tilde{P} = \frac{1}{2}(I + P)$ is the lazy version of P . \forall initial state x ,

(i) for suff. large k , if $\|\tilde{P}_x^{(k)} - \pi\|_{\text{TV}} \leq \varepsilon$, then $\|h_x^{(k)} - \pi\|_{\text{TV}} \leq 2\varepsilon$.

(ii) for suff. large t , if $\|h_x^{(t)} - \pi\|_{\text{TV}} \leq \varepsilon$, then $\|\tilde{P}_x^{(t+1)} - \pi\|_{\text{TV}} \leq 2\varepsilon$.

Proof Sketch: $\|\tilde{P}_x^{(k)} - \pi\|_{TV} \leq \varepsilon \rightarrow \tilde{H}$ after $2k$ steps, $\Pr[\text{fewer than } k \text{ transitions happened}] = \Pr[P_0(2k) < k] \rightarrow 0$ as $k \rightarrow \infty$. $\tilde{H}_{2t} = \tilde{H}$.

$\|h_x^{(t)} - \pi\|_{TV} \leq \varepsilon \rightarrow$ same holds after extra t P transitions. # of transitions in P is $P_0(t) + t$. \rightarrow $4t$ transitions in $\tilde{P} \rightarrow$ # trans. in P is $\text{Bin}(4t, \frac{1}{2})$.
 $\|(P_0(t) + t) - \text{Bin}(4t, \frac{1}{2})\|_{TV} \rightarrow 0$ as $t \rightarrow \infty$.

Previously, $\text{Var}_\pi[P_\varphi] - \text{Var}_\pi[\varphi] \leq -\mathcal{E}_p(\varphi, \varphi)$. (difference form)

Thm) For an irreducible P & heat kernel H_t , $\forall \varphi: \Omega \rightarrow \mathbb{R}$,

$$\frac{d}{dt} \text{Var}_\pi[H_t \varphi] = -2 \mathcal{E}_p(H_t \varphi, H_t \varphi). \quad (\text{derivative form})$$

Proof: Let $\varphi_t := H_t \varphi$. Since $H_t = \exp(-tL)$, $\frac{d}{dt} \varphi_t = -L \varphi_t$.

$$\begin{aligned} \frac{d}{dt} \text{Var}_\pi[\varphi_t] &= \frac{d}{dt} \left[E_\pi[\varphi_t]^2 - \overbrace{E_\pi[\varphi_t^2]}^0 \right] = \frac{d}{dt} \left[\sum_x \pi(x) \cdot \varphi_t(x)^2 \right] \\ &= 2 \sum_x \pi(x) \varphi_t(x) \frac{d}{dt} \varphi_t(x) = -2 \sum_x \pi(x) \varphi_t(x) [L \varphi_t(x)] = -2 \mathcal{E}_p(\varphi_t, \varphi_t), \end{aligned}$$

Corollary: $\forall \varphi$, $\frac{d}{dt} \text{Var}_\pi[H_t \varphi] \leq -2\alpha \text{Var}_\pi[H_t \varphi]$ where $\alpha := \inf_{\text{non-const}} \frac{\mathcal{E}_p(\varphi_t, \varphi_t)}{\text{Var}_\pi[\varphi_t]}$.

Corollary: $T_x(\varepsilon) \leq \frac{1}{2\alpha} (2 \ln \varepsilon^{-1} + \ln(4\pi(x))^{-1}) = O\left(\frac{1}{\alpha} \ln \pi_{\min}^{-1}\right)$.

Def) Entropy: $\text{Ent}_\pi[\varphi] := \sum_x \pi(x) \varphi(x) \log \varphi(x) - E_\pi[\varphi] \log E_\pi[\varphi]$
 $= \sum_x \pi(x) \varphi(x) \log \frac{\varphi(x)}{E_\pi[\varphi]}$.

Set $\varphi = \frac{\mu}{\pi}$ where μ is a prob. dist. over Ω . $\rightarrow E_\pi[\varphi] = 1$, and

$\text{Ent}_\pi[\varphi] = \sum_x \mu(x) \log \frac{\mu(x)}{\pi(x)} = D_{KL}(\mu \parallel \pi)$, the KL-Divergence.

Goal: bound rate of decay of $D(p_x^{(\cdot)} \parallel \pi) = \text{Ent}_\pi\left[\frac{p_x^{(\cdot)}}{\pi}\right]$.

Def) (Modified) Log-Sobolev Constant: $\rho := \inf_{\varphi \geq 0} \frac{\mathbb{E}_P(\varphi, \log \varphi)}{\text{Ent}_\pi[\varphi]}.$

Thm) $T_x(\varepsilon) \leq \frac{1}{\rho} (2 \ln \varepsilon^{-1} + \ln \ln \pi_x^{-1}) \Rightarrow T_{\min} = O\left(\frac{1}{\rho} \log \log \pi_{\min}^{-1}\right).$

Thm*) $\forall \varphi: \Omega \rightarrow \mathbb{R}^+$, $\frac{d}{dt} \text{Ent}_\pi\left[\frac{\varphi_t}{\pi}\right] = -\mathcal{E}_P(\varphi_t, \log \varphi_t)$ where $\varphi_t := \frac{\varphi H_t}{\pi}$.

Proof*: $\varphi_t = \frac{\varphi H_t}{\pi} = H_t^*(\frac{\varphi}{\pi})$ where H_t^* is the heat kernel of reversalization of P ,

$$p^*(x, y) = \frac{\pi(y) P(x, y)}{\pi(x)} \quad [\text{Exercise}]. \rightarrow \frac{d}{dt} \varphi_t = \frac{d}{dt} H_t^*\left(\frac{\varphi}{\pi}\right) = -L^* \varphi_t.$$

$$\begin{aligned} \frac{d}{dt} \text{Ent}_\pi[\varphi_t] &= \frac{d}{dt} \left[E_\pi[\varphi_t \log \varphi_t] - E_\pi[\varphi_t] \log E_\pi[\varphi_t] \right] \quad (\sum_x \pi(x) \frac{d}{dt} \varphi_t(x) = 0 \\ &= \frac{d}{dt} \sum_x (\pi(x) \varphi_t(x) \log \varphi_t(x)) = \sum_x \pi(x) (1 + \log \varphi_t(x)) \frac{d}{dt} \varphi_t(x) \\ &= - \sum_x \pi(x) \log \varphi_t(x) [L^* \varphi_t](x). \quad L^* \text{ is adjoint on } \langle \cdot, \cdot \rangle_\pi \langle \eta, L^* \psi \rangle_\pi = \langle \psi, L \eta \rangle_\pi. \\ &\rightarrow - \sum_x \pi(x) \varphi_t(x) [L \log \varphi_t](x) = -\mathcal{E}_P(\varphi_t, \log \varphi_t). \end{aligned}$$

Corollary: $\forall \varphi: \Omega \rightarrow \mathbb{R}^+$, $\text{Ent}_\pi[\varphi] \neq 0$, $\text{Ent}_\pi\left[\frac{\varphi H_t}{\pi}\right] \leq \exp(-pt) \text{Ent}_\pi\left[\frac{\varphi}{\pi}\right]$.

Proof^o: Set $\varphi = p_x^{(0)}$ (initial dist. concentrated at x).

$$\begin{aligned} \text{Recall } \text{Ent}_\pi\left[\frac{p_x^{(\cdot)}}{\pi}\right] &= D_{KL}(p_x^{(\cdot)} \parallel \pi). \rightarrow D_{KL}(p_x^{(\cdot)} \parallel \pi) \leq \exp(-pt) D_{KL}(p_x^{(0)} \parallel \pi) \\ &= \exp(-pt) \log \pi(x)^{-1}. \quad \text{ Pinsker's Inequality: } 2 \|p - \pi\|_{TV}^2 \leq D_{KL}(p \parallel \pi). \\ &\rightarrow \|p_x^{(\cdot)} - \pi\|_{TV} \leq \varepsilon \quad \forall t \geq \frac{1}{p} (\ln \frac{1}{2\varepsilon^2} + \ln \ln \pi(x)^{-1}) \Rightarrow T_{\min} = O\left(\frac{1}{p} \ln \ln \pi_{\min}^{-1}\right). \end{aligned}$$

Fact [Exercise]: if P is reversible, $\sum_p (\varphi, \psi) = \frac{1}{2} \sum_{xy} \pi(x) P(x,y) (\varphi(x) - \varphi(y)) (\psi(x) - \psi(y))$

ex) Random Walk on Hypercube: $\Omega = \{0,1\}^n$. Decompose $\Omega = \Omega_1 \times \dots \times \Omega_n^{\{0,1\}}$.

Base Case: $P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$. Let φ be an arbitrary non-negative function

$$\text{s.t. } \alpha := \varphi(0), \beta := \varphi(1). \quad \sum_p (\varphi, \log \varphi) = \frac{1}{2} \sum_{xy} \pi(x) P(x,y) (\varphi(x) - \varphi(y)) \log \frac{\varphi(x)}{\varphi(y)}$$

$$= \frac{1}{2} (a-b) \log \frac{a}{b}. \quad \underline{\text{Ent}_n[\varphi]} = E_\pi[\varphi \log \varphi] - E_\pi[\varphi] \log E_\pi[\varphi]$$

$$= \frac{a \log a + b \log b}{2} - \left(\frac{a+b}{2} \right) \log \left(\frac{a+b}{2} \right) \Rightarrow \text{Find } \min_C \max_{a,b \geq 0} \frac{a \log a + b \log b - (a+b) \log \left(\frac{a+b}{2} \right)}{(a-b) \log \left(\frac{a}{b} \right)} \leq C.$$

With some Hessian argument [omitted, in notes], $C = 1/4 \Rightarrow \underline{P \geq 4}$.

If we have some operator P_i on Ω_i , $P = \sum_i q_i (I \otimes I \otimes \dots \otimes P_i \otimes \dots \otimes I)$ where $\sum_i q_i = 1$.

Claim*: Let $(\Omega_1, P_1), (\Omega_2, P_2)$ be two reversible MCs on disjoint state spaces $\Omega_1 \& \Omega_2$, stat. dist. π_1, π_2 . Then the combined MC $q_1(P_1 \otimes I) + q_2(I \otimes P_2)$ is reversible w.r.t. $(\pi_1 \otimes \pi_2)$ [ex] and $P \geq \min\{P_1 q_1, P_2 q_2\}$ where P_1, P_2 are log-Sobolev const. for P_1, P_2 .

Corollary: for $\{0,1\}^n$, $n=1 \rightarrow P \geq 4$. $n > 1 \rightarrow P \geq 4/n$ by induction.

Proof: $\Omega = \Omega_1 \times \{0,1\}^{n-1}$. $P = \frac{1}{n} (P_1 \otimes I) + (1 - \frac{1}{n}) (I \otimes P_{n-1})$. By IH,

$$P_1 \geq 4, P_{n-1} \geq \frac{4}{n-1} \rightarrow P \geq \min\left\{\frac{4}{n}, \frac{n-1}{n} \cdot \frac{4}{n-1}\right\} = \frac{4}{n} \Rightarrow T_{\text{mix}} = O(n \ln n) !!!$$

(tight for hypercubes)

To prove the Claim*, we first introduce two lemmas.

Lemma 1: $\mathbb{E}_p(\varphi, \psi) = q_1 \sum_{x_2} \pi_2(x_2) \mathbb{E}_{p_1}(\varphi(\cdot, x_2), \psi(\cdot, x_2)) + q_2 \sum_{x_1} \pi_1(x_1) \mathbb{E}_{p_2}(\varphi(x_1, \cdot), \psi(x_1, \cdot))$.

Lemma 2: $\text{Ent}_\pi[\varphi] \leq \sum_{x_1} \pi(x_1) \text{Ent}_{\pi_2}[\varphi(x_1, \cdot)] + \sum_{x_2} \pi_2(x_2) \text{Ent}_{\pi_1}[\varphi(\cdot, x_2)]$.

Proof of Claim*: $\mathbb{E}_p(\varphi, \log \varphi) = q_1 \sum_{x_2} \pi_2(x_2) \mathbb{E}_{p_1}(\sim) + q_2 \sum_{x_1} \pi_1(x_1) \mathbb{E}_{p_2}(\sim)$
 $\geq q_1 p_1 \sum_{x_2} \pi_2(x_2) \text{Ent}_{\pi_1}[\varphi(\cdot, x_2)] + q_2 p_2 \sum_{x_1} \pi_1(x_1) \text{Ent}_{\pi_2}[\varphi(x_1, \cdot)]$
 $\geq \min\{q_1 p_1, q_2 p_2\} \cdot \left[\sum_{x_2} \pi_2(x_2) \text{Ent}_{\pi_1}[\varphi(\cdot, x_2)] + \sum_{x_1} \pi_1(x_1) \text{Ent}_{\pi_2}[\varphi(x_1, \cdot)] \right]$
 $\geq \min\{q_1 p_1, q_2 p_2\} \cdot \text{Ent}_\pi[\varphi].$

$x := (x_1, x_2), y := (y_1, y_2)$

Proof of Lemma 1: $\mathbb{E}_p(\varphi, \psi) = \frac{1}{2} \sum_{xy} \pi(x) P(x, y) (\varphi(x) - \varphi(y)) (\psi(x) - \psi(y))$
 $= \frac{q_1}{2} \sum_{x_1, y_1, y_2} \pi_1(x_1) \pi_2(x_2) P_1(x_1, y_1) (\varphi(x_1, x_2) - \varphi(y_1, x_2)) (\psi(x_1, x_2) - \psi(y_1, x_2))$
 $+ \frac{q_2}{2} \sum_{x_1, y_1, y_2} \pi_1(x_1) \pi_2(x_2) P_2(x_2, y_2) (\varphi(x_1, x_2) - \varphi(x_1, y_2)) (\psi(x_1, x_2) - \psi(x_1, y_2))$
 $= q_1 \sum_{x_2} \pi_2(x_2) \mathbb{E}_{p_1}(\varphi(\cdot, x_2), \psi(\cdot, x_2)) + q_2 \sum_{x_1} \pi_1(x_1) \mathbb{E}_{p_2}(\varphi(x_1, \cdot), \psi(x_1, \cdot)).$

Proof of Lemma 2: $\text{Ent}_\pi[\varphi] = \sum_x \pi(x) \varphi(x) \log \left(\frac{\varphi(x)}{\mathbb{E}_\pi[\varphi]} \right)$
 $= \sum_{x_1, x_2} \pi_1(x_1) \pi_2(x_2) \varphi(x_1, x_2) \log \left(\frac{\varphi(x_1, x_2)}{\sum_{x_1, x_2} \pi_1(x_1) \pi_2(x_2) \varphi(x_1, x_2)} \right)$
 $= \underbrace{\sum_{x_1, x_2} \pi_1(x_1) \pi_2(x_2) \varphi(x_1, x_2)}_{\sum_{x_1} \pi_1(x_1) \text{Ent}_{\pi_2}[\varphi(x_1, \cdot)]} \log \left(\frac{\varphi(x_1, x_2)}{\sum_{x_2} \pi_2(x_2) \varphi(x_1, x_2)} \right) + \underbrace{\sum_{x_1, x_2} \pi_1(x_1) \pi_2(x_2) \varphi(x_1, x_2)}_{\sum_{x_2} \pi_2(x_2) \text{Ent}_{\pi_1}[\varphi(\cdot, x_2)]} \log \left(\frac{\sum_{x_2} \pi_2(x_2) \varphi(x_1, x_2)}{\sum_{x_1, x_2} \pi_1(x_1) \pi_2(x_2) \varphi(x_1, x_2)} \right)$
 $\hookrightarrow \sum_{x_1} \pi_1(x_1) \text{Ent}_{\pi_2}[\varphi(x_1, \cdot)] / \text{use log-Sobolev inequality: } \sum a_i \log \frac{\sum a_i}{\sum b_i} \leq \sum a_i \log \left(\frac{a_i}{b_i} \right) \forall a_i \geq 0, b_i > 0$
 $\rightarrow \leq \sum_{x_1} \pi_1(x_1) \sum_{x_2} \pi_2(x_2) \varphi(x_1, x_2) \log \left(\frac{\pi_2(x_2) \varphi(x_1, x_2)}{\pi_1(x_1) \sum_{x_2} \pi_2(x_2) \varphi(x_1, x_2)} \right) = \sum_{x_2} \pi_2(x_2) \sum_{x_1} \pi_1(x_1) \varphi(x_1, x_2) \log(\sim)$

$$= \sum_{x_2} \pi_2(x_2) \text{Ent}_{\pi_1} [\varphi(\cdot, x_2)] .$$

Matroid Bases

Def) Matroid: $M(E, \mathcal{I})$, $\mathcal{I} \subseteq 2^E$ where (i) $I \in \mathcal{I}$ and $J \in \mathcal{I} \Rightarrow J \in \mathcal{I}$.
 (ii) $I, J \in \mathcal{I}, |I| > |J| \Rightarrow \exists e \in I \setminus J \text{ s.t. } J + e \in \mathcal{I}$. \rightarrow augmentation property

Def) Basis: maximal indep. set. of a matroid M .

Fact: all bases have same cardinality \rightarrow call it "rank" of M .

ex) Graphical Matroids: Connected $G(V, E)$, $E = E$. \mathcal{I} = sets of forests.

\rightarrow Bases are spanning trees.

ex) Uniform Matroids: E , a finite set. \mathcal{I} = subsets of E of size $\leq k$.

\rightarrow Bases are k -subsets of E .

ex) Linear/Representable Matroids: $n \times m$ matrix over \mathbb{F} , $m \geq n$. $E = \text{columns}$.

\mathcal{I} = linearly independent subsets of E . \rightarrow bases are maximally indep. columns.

* We assume a membership oracle; given $J \subseteq E$, tells if $J \in \mathcal{I}$.

Algorithm: use a suitable MCMC to sample bases to \mathcal{B} , set of bases.

Natural algorithm: Bases Exchange MCMC. With current basis B ,

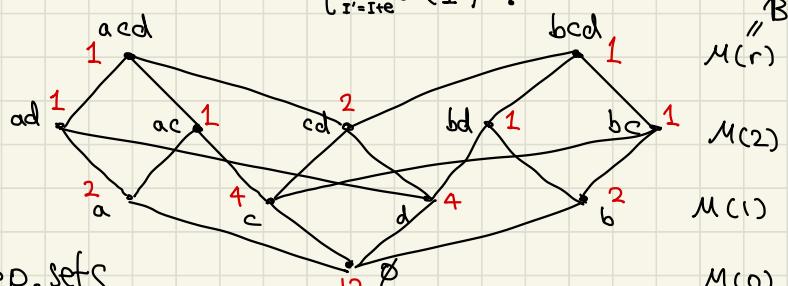
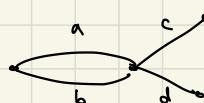
Pick $e \in B$ u.a.r and pick f u.a.r. from $F = \{f \in E \mid B \setminus \{e\} \cup \{f\} = I\}$.

Irreducibility? $\forall B \rightarrow B'$, $\{e_1, \dots, e_{2k}\} = B \oplus B'$. By augmentation property, if we throw out e_i from B , $\exists e_j$ s.t. $B \setminus e_i + e_j \in I$. induct.,

\downarrow optimal!
Thm) T_{mix} of Basis Exchange MC is $O(n \log n)$ where $n = |E|$ (can be improved to $r \log r$)

Def) Weight Function: $w: I \rightarrow \mathbb{Z}^+$. $w(I) = \begin{cases} 1 & \text{if } I \in I \\ \sum_{I' \in M_k} w(I') & \text{otherwise} \end{cases}$.

ex)



$$\text{Obs: } w(\emptyset) = |\mathcal{B}| \times (r!)$$

$M(k) := k\text{-th level indep. sets.}$

π_k on M_k : $\pi_k(I) := \frac{w(I)}{\sum_{I \in M_k} w(I)}$ where $Z_k = \sum_{I \in M_k} w(I)$.

Def) Contraction: $\forall I \in J$, $M_I = (E \setminus I, \overbrace{I}^{J \subseteq E \setminus I \mid J \cup I = J}) \rightarrow \{J \subseteq E \setminus I \mid J \cup I = J\}$.

Write $\pi_{I,k}$ for conditional distribution on M_I , $\pi_{I,k}(J) := \begin{cases} \frac{k! w(J)}{w(I)} & \text{if } J \supseteq I \\ 0 & \text{o.w.} \end{cases}$

A Down-Up Walk P_k^v on $M(k)$: remove $e \in I$ uar, let $S = \{J \in M(k) \mid J = I - e + e'\text{ for some } e'\}$. Move to $J \in S$ w.p. $\propto w(J)$.

Modified Thm) $\forall 2 \leq k \leq r$, M -L-S const. of P_k^v satisfies $p_k \geq \frac{1}{|\mathcal{B}| \leq \binom{n}{k} \leq n^r}$

$$\Rightarrow T_{\text{mix}} = O\left(\frac{1}{p_r} \times \log \log \overbrace{\pi_{\min}^{-1}}^{M(k-1)}\right) = O(r(\log r + \log \log n)) = O(r \log n)!$$

Let $A_{k-1}(I, J) := \mathbb{1}_{\{J \text{ s.t. } J = I + e\}} \overbrace{\pi_{M(k-1)}}^{M(k-1)}(\dots \overset{i}{\dots})$ Also let

$P_{k-1}^{\downarrow} := \frac{1}{k} A_{k-1}^T$, $P_{k-1}^{\uparrow} := A_{k-1}$ with (I, J) replaced by $\frac{w(J)}{w(I)}$. Then $P_k^V = P_k^{\downarrow} P_{k-1}^{\uparrow}$.

Claim: [Exercise] $\forall I, J, \pi_k(I) P_k^{\downarrow}(I, J) = \pi_{k-1}(J) P_{k-1}^{\uparrow}(J, I)$.

$\forall f^{(k)}: M_k \rightarrow \mathbb{R}^{>0}$ and any $i < k$, define $f^{(i)} := P_i^{\uparrow} P_{i+1}^{\uparrow} \dots P_{k-1}^{\uparrow} f^{(k)}$.

Fact: Always assume wlog functions have mean 1, i.e. $E_{\pi_k}[f^{(k)}] = 1$.

Then this implies that $E_{\pi_i}[f^{(i)}] = 1 \forall i < k$ as well.

Main Lemma: $\forall k \geq 2, \forall f^{(k)}: M_k \rightarrow \mathbb{R}^{>0}$ with $E_{\pi_k}[f^{(k)}] = 1$, we have

$$\text{Ent}_{\pi_{k-1}}[f^{(k-1)}] \leq ((-\frac{1}{k}) \text{Ent}_{\pi_k}[f^{(k)}]). \quad (\text{Entropy Contraction})$$

Lemma \rightarrow Thm: s.t.o.p. $P_k = \inf_{f^{(k)} \sim} \frac{\mathcal{E}_{P_k^V}(f^{(k)}, \log f^{(k)})}{\text{Ent}_{\pi_k}[f^{(k)}]}$.

Claim: $\text{Ent}_{\pi_{k-1}}[f^{(k-1)}] \geq \text{Ent}_{\pi_k}[f^{(k)}] - \mathcal{E}_{P_k^V}(f^{(k)}, \log f^{(k)})$. If Claim holds,

$$\mathcal{E}_{P_k^V}(f^{(k)}, \log f^{(k)}) \geq \text{Ent}_{\pi_k}[f^{(k)}] - \text{Ent}_{\pi_{k-1}}[f^{(k-1)}]$$

$\geq \frac{1}{k} \text{Ent}_{\pi_k}[f^{(k)}]$ by the Main Lemma $\rightarrow P_k \geq \frac{1}{k}$.

Proof of Claim: $\text{Ent}_{\pi_{k-1}}[f^{(k-1)}] = \sum_{I \in M(k-1)} \pi_{k-1}(I) f^{(k-1)}(I) \log(f^{(k-1)}(I))$

$$= \sum_{I \in M(k-1)} \pi_{k-1}(I) [P_{k-1}^{\uparrow} f^{(k)}](I) \log([P_{k-1}^{\uparrow} f^{(k)}](I))$$

$$= \sum_{I \in M(k-1), J \in M(k)} \pi_{k-1}(I) P_{k-1}^{\uparrow}(I, J) f^{(k)}(J) \log(\sum_{L \in M(k)} P_{k-1}^{\uparrow}(I, L) f^{(k)}(L))$$

$$= \sum_{I, J \in M(k-1), M(k)} \pi_k(J) f^{(k)}(J) P_k^{\downarrow}(J, I) \log(\quad) [\text{by "detailed balance"}]$$

$$\geq \sum_{\substack{I \in M(k) \\ J, L \in M(k)}} \pi_k(J) f^{(k)}(J) P_k^{\downarrow}(J, I) P_{k-1}^{\uparrow}(I, L) \log(f^{(k)}(L)) [\text{by Jensen's Inequality}]$$

$$= \sum_{S, L \in M(k)} \pi_k(S) f^{(k)}(S) P_k^V(S, L) \log(f^{(k)}(L))$$

$$= \text{Ent}_{\pi_k} [f^{(k)}] - \sum_{e \in E} (\pi_e f^{(k)}, \log f^{(k)}) . //$$

Proof of Main Lemma: Induction step for $k > 2$.

$$\text{Assume } \text{Ent}_{\pi_{k-1}} [f^{(k-1)}] \leq ((1 - \frac{1}{k}) \text{Ent}_{\pi_k} [f^{(k)}]).$$

$$\text{To prove: } \text{Ent}_{\pi_k} [f^{(k)}] \leq ((1 - \frac{1}{k+1}) \text{Ent}_{\pi_{k+1}} [f^{(k+1)}]).$$

$$\text{Decompose } \pi_k = \sum_e \pi_e (e) \pi_{e,k-1}.$$

$$\text{By Standard Procedure [Ex]} \quad \text{Ent}_{\pi_k} [f^{(k)}] = \text{Ent}_{\pi_1} [f^{(1)}] + \sum_{e \in E} \pi_e (e) \text{Ent}_{\pi_{e,k-1}} [f^{(k)}].$$

$$\text{By IH applied to all of } f^{(k)}, f^{(k-1)}, \dots, f^{(2)}, \text{Ent}_{\pi_k} [f^{(k)}] \geq k \text{Ent}_{\pi_1} [f^{(1)}].$$

$$\rightarrow \sum_{e \in E} \pi_e (e) \text{Ent}_{\pi_{e,k-1}} [f^{(k)}] \geq (k-1) \text{Ent}_{\pi_1} [f^{(1)}].$$

$$\begin{aligned} \text{Ent}_{\pi_{k+1}} [f^{(k+1)}] &= \text{Ent}_{\pi_1} [f^{(1)}] + \sum_e \pi_e (e) \text{Ent}_{\pi_{e,k}} [f^{(k+1)}] \\ &\geq \text{Ent}_{\pi_1} [f^{(1)}] + \sum_e \pi_e (e) \frac{k}{k-1} \text{Ent}_{\pi_{e,k-1}} [f^{(k)}] \quad [\text{By IH on Me}] \\ &= \text{Ent}_{\pi_1} [f^{(1)}] + \frac{1}{k(k-1)} \sum_e (\pi_e (e) \text{Ent}_{\pi_{e,k-1}} [f^{(k)}]) + \frac{k+1}{k} \sum_e (//) \\ &\geq \text{Ent}_{\pi_1} [f^{(1)}] + \frac{k-1}{k(k-1)} \text{Ent}_{\pi_1} [f^{(1)}] + \frac{k+1}{k} \sum_e \pi_e (e) \text{Ent}_{\pi_{e,k-1}} [f^{(k)}] \\ &= \frac{k+1}{k} (\text{Ent}_{\pi_1} [f^{(1)}] + \sum_e \pi_e (e) \text{Ent}_{\pi_{e,k-1}} [f^{(k)}]) = \frac{k+1}{k} \text{Ent}_{\pi_k} [f^{(k)}]. // \end{aligned}$$

*Base Case of $k=2$ is deferred to Lecture Notes. It uses a completely orthogonal set of analyses from the induction. This content may be covered later if time permits.

Correlation Decay

\sqrt{n}	-	-
+	-	-
-	+	+
-	+	-

\emptyset field, λ interaction. For some λ_c , $\lambda < \lambda_c \Rightarrow$ no long-range correlations as $n \rightarrow \infty$ & GD has $T_{\text{mix}} = O(n \log n)$.
 OTOTH, $\lambda > \lambda_c \Rightarrow$ long-range correlations exist as $n \rightarrow \infty$ and GD has $T_{\text{mix}} = \exp(\Omega(\sqrt{n}))$.

Hard-Core Model: $G(V, E)$, $\mathcal{I} = \text{indep. sets in } G$, $\pi(I) = \frac{\lambda^{|I|}}{Z(\lambda)}$, $Z(\lambda) = \sum_k a_k \lambda^k$

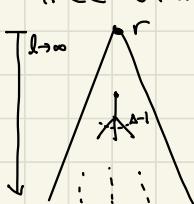
.	.
.	.

λ small

.	.	.	.
.	.	.	.
.	.	.	.

λ large

Tree view: infinite Δ -regular tree $T(\Delta)$



Fact: $\forall \Delta, \exists \lambda_c(\Delta) = \frac{(\Delta-1)^{\Delta-1}}{(\Delta-2)^\Delta}$ s.t. "phase transition" occurs at $\lambda = \lambda_c(\Delta)$.

Thm) For graphs of max degree $\Delta \geq 3$:

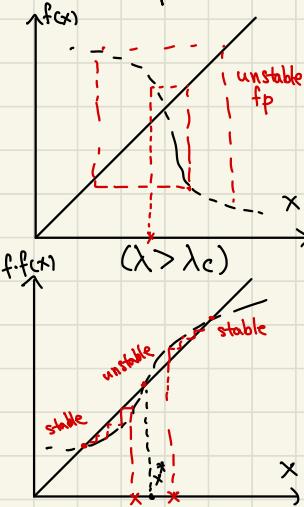
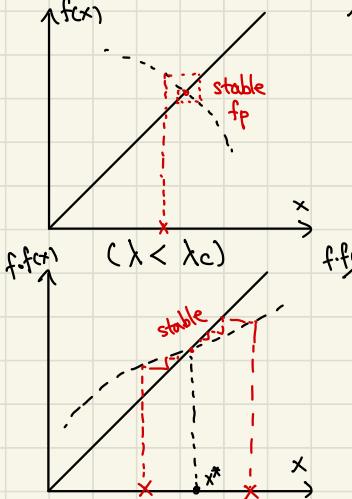
- (i) [Weitz] If $\lambda < \lambda_c(\Delta)$ then \exists FPTAS for $Z(\lambda)$
- (ii) [Sly] If $\lambda > \lambda_c(\Delta)$ then \nexists FPRAS for $Z(\lambda)$ (unless RP=NP)

Def) Weak Spatial Mixing (WSM): Let S be any subset of vertices in $T(\Delta)$ and T be some arbitrary configuration of S . Let \Pr^T be the probability that the root r is occupied. WSM holds if $|\Pr^T - \Pr^{\bar{T}}| \leq \exp(-c \cdot \text{dist}(r, S))$ for some $c > 0$ $\forall T, \bar{T}$ configs of S .

Proof of Fact: Let $p_v := \Pr_{T_v}[\sigma(v) = \emptyset]$, prob. that root v is unoccupied.

$$p_v = \frac{\sum_{T_v} [\sigma(v) = \emptyset]}{\sum_{T_v} [\sigma(v) = \emptyset] + \sum_{T_v} [\sigma(v) = 1]} = \frac{\prod_{i=1}^d Z_{T_{v,i}}}{\prod_{i=1}^d Z_{T_{v,i}} + \lambda \cdot \prod_{i=1}^d Z_{T_{v,i}} [\sigma(v) = 1]} = \frac{1}{1 + \lambda \cdot \prod_{i=1}^d R_{v,i}}.$$

Translation invariance $\rightarrow p_v = p \forall v \Rightarrow p = \frac{1}{1 + \lambda p^d}$, p is f.p. of $f(x) = \frac{1}{1 + \lambda x^d}$.



stability is $f'(x^*) < 1$.

$$(f \circ f)'(x) = f'(f(x))f'(x).$$

$$\text{At } x^*, (f \circ f)'(x) = f'(x^*)^2.$$

$$\text{So } (f \circ f)'(x^*) > 1 \Leftrightarrow f'(x^*) > -1.$$

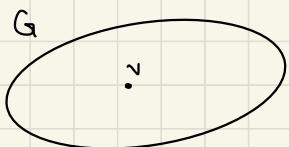
$$\text{At } x=x^*, x^* = f(x^*) = \frac{1}{1 + \lambda x^{*d}}. \quad (\text{a})$$

$$f'(x^*) = -1 = -\lambda \frac{x^{*(d-1)}}{(1 + \lambda x^{*d})^2}. \quad (\text{b})$$

$$(\text{a}) \rightarrow (\text{b}) \text{ shows } d(x^* - 1) = 1 \Rightarrow x^* = \frac{d-1}{d}. \text{ plug into (a)} \rightarrow \lambda = \frac{d^d}{(d-1)(d+1)}. //$$

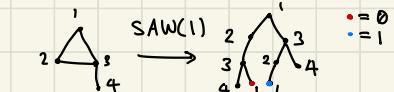
Sketch of (ii) of Theorem: 1. SAW construction 2. Spatial mixing for truncation

Consider graph G with max degree Δ . Compute $\Pr[v \text{ occupied}]?$ $\lambda < \lambda_c(\Delta)$.



- Construct a Self-Avoiding Walk Tree $T_{SAW}(G, v)$.
 \nearrow can't go back immediately
- Use corr. decay on $T_{SAW}(G, v)$ to truncate the tree at

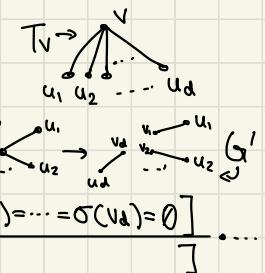
level $O(\log n)$ and do explicit recurrence.



Claim 1: $\Pr_G[\sigma(v) = \emptyset] = \Pr_{T_{SAW}(G, v)}[\sigma(v) = \emptyset]$.

Proof: Let $R_v := \frac{\Pr[\sigma(v) = \emptyset]}{\Pr[\sigma(v) = 1]}$. s.t.p. that $\underline{R}_v^G = \overline{R}_v^{T_v}$.

$$R_V^{TV} = \frac{\sum_{T_U} [\sigma(v) = 1]}{\sum_{T_U} [\sigma(v) = 0]} = \frac{\lambda \cdot \prod_{i=1}^d Z_{T_{U_i}} [\sigma(u_i) = 0]}{\prod_{i=1}^d (Z_{T_{U_i}} [\sigma(u_i) = 0] + Z_{T_{U_i}} [\sigma(u_i) = 1])} = \lambda \prod_{i=1}^d \frac{1}{1 + R_{U_i}^{T_U}}.$$



$$R_V^G = \frac{Pr_G[\sigma(v) = 1]}{Pr_G[\sigma(v) = 0]} = \lambda \frac{\prod_{i=1}^{d-1} Pr_{G_i}[\sigma(v_1) = \dots = \sigma(v_d) = 1]}{Pr_{G'_i}[\sigma(v_1) = \dots = \sigma(v_d) = 0]} = \lambda \frac{\prod_{i=1}^{d-1} Pr_{G_i}[\sigma(v_1) = 1 | \sigma(v_2) = \dots = \sigma(v_d) = 0] \cdot Pr[\sigma(v_2) = 1 | \sigma(v_1) = 1, \sigma(v_3) = \dots = \sigma(v_d) = 0]}{Pr_{G'_i}[\sigma(v_1) = 0 | \sigma(v_2) = \dots = \sigma(v_d) = 0] \cdot Pr[\sigma(v_2) = 0] \dots}.$$

(Using $\frac{Pr[AB]}{Pr[CD]} = \frac{Pr[A|D]}{Pr[C|D]} \cdot \frac{Pr[B|A]}{Pr[D|A]}$)

$$= \lambda \prod_{i=1}^{d-1} R_{v_i}^{G_i, T_i} \text{ where } T_i(v_j) = \begin{cases} 1 & j < i \\ 0 & \text{otherwise} \end{cases}.$$

$$\rightarrow R_{v_i}^{G_i, T_i} = \frac{\lambda \sum_{G_i - v_i}^{T_i} [\sigma(u_i) = 0]}{\sum_{G_i - v_i}^{T_i} [\sigma(u_i) = 0] + \sum_{G_i - v_i}^{T_i} [\sigma(u_i) = 1]} = \frac{\lambda}{1 + R_{u_i}^{G_i - v_i, T_i}}.$$

$$\rightarrow R_V^G = \lambda \prod_{i=1}^{d-1} \frac{\lambda}{1 + R_{u_i}^{G_i - v_i, T_i}} = \lambda \prod_{i=1}^d \frac{1}{1 + R_{u_i}^{G_i - v_i, T_i}}.$$

$R_{u_i}^{G_i - v_i, T_i} = R_{u_i}^{T_{U_i}}$ by induction on subtree. //

Def) Strong Spatial Mixing: $|Pr^T - Pr^{T'}| \leq \exp(-c \cdot \text{dist}(r, \hat{S}))$ where $T = T'$ for $S \setminus \hat{S}$.

Thm 2) For hard-core model on T_u , WSM \Rightarrow SSM.

Proof: Fix boundary conditions accordingly. Let $P_v = Pr[\sigma(v) = 0]$.

$$\text{As before, } P_v = \frac{1}{1 + \lambda \prod_{i=1}^d P_{u_i}} = f(P_{u_1}, \dots, P_{u_n}).$$

Need to show that $|P_v - P'_v| \leq (1-\gamma) \max_i |P_{v_i} - P'_{v_i}|$. This is false. But,

Def) Message/Potential: cont. diff. function $\varphi: [\frac{1}{1+\lambda}, 1] \rightarrow \mathbb{R}$ with positive derivative. \rightarrow same holds for φ' .

$$\rightarrow \text{Use } \varphi(x) = \frac{1}{S} \log\left(\frac{x}{S-x}\right) \text{ where } S := \frac{d+1}{d}.$$

Work with $\varphi(P_v)$ instead of P_v . $P_v \rightarrow P_i$. $m = \varphi(P_v)$, $m_i = \varphi(P_i)$.

$$\rightarrow M = \varphi(f(\varphi^{-1}(m_1), \varphi^{-1}(m_2), \dots, \varphi^{-1}(m_n)) = F(m_1, \dots, m_n).$$

Key Fact: $\forall \lambda < \lambda_c(\Delta)$, $\exists \gamma > 0$ s.t. $\forall \vec{m}, \vec{m}' \in (\varphi[\frac{1}{1+\lambda}, 1])^d$,

$$|F(\vec{m}) - F(\vec{m}')| \leq (-\gamma) \|\vec{m} - \vec{m}'\|_\infty \rightarrow \text{Implies original goal.}$$

Sketch of (ii) of Theorem: We will actually show the following:

[Sly] A poly time algorithm for sampling ind. sets from Gibbs distribution $\pi(I) = \frac{1}{Z(\lambda)} \lambda^{|I|}$ would allow us to solve MAXCUT in polytime w.h.p. (for $\lambda > \lambda_c(\Delta)$). 

Consider the bipartite graph G with independent, random Δ -deg. perfect matchings. An (α, β) -ind. set has αn vertices in L, βn in R.

$$\text{Let } Z_G^{(\alpha, \beta)}(\lambda) := \sum_{\alpha, \beta \text{ IS } I} \lambda^{(\alpha + \beta)n}.$$

Claim: $E[Z_G^{(\alpha, \beta)}(\lambda)] = \exp\{\Phi_\lambda(\alpha, \beta) \cdot n \cdot (1 + o(1))\}$ where

$$\Phi_\lambda(\alpha, \beta) := (\alpha + \beta) \ln \lambda + H(\alpha) + H(\beta) + \Delta \left[(1 - \beta) H\left(\frac{\alpha}{1-\beta}\right) - H(\alpha) \right]$$

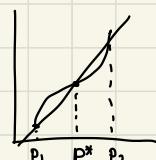
and $H(x) = -x \ln x - (1-x) \ln(1-x)$ (two-point entropy).

$$\text{Proof: } E[Z_G^{(\alpha, \beta)}(\lambda)] = \lambda^{(\alpha + \beta)n} \sum_{|I_L|=\alpha n, |I_R|=\beta n} \Pr[I_L \cup I_R \text{ is ind. set}]$$

$$= \lambda^{(\alpha + \beta)n} \binom{n}{\alpha n} \binom{n}{\beta n} \Pr[I_L \cup I_R \text{ is ind. set}] = \lambda^{(\alpha + \beta)n} \binom{n}{\alpha n} \binom{n}{\beta n} \left[\frac{\binom{n-\beta n}{\alpha n}}{\binom{n}{\alpha n}} \right]^\Delta.$$

Using $\ln \binom{n}{\alpha n} \sim n H(c) + o(1)$, we get the result. //

Fact: $\Phi_\lambda(\alpha, \beta)$ satisfies: (i) max on line $\alpha = \beta$ is at $\alpha = \beta = p^*$.



(iii) if $\lambda < \lambda_c(\Delta)$ then $x = \beta = p^*$ is a global maximum.

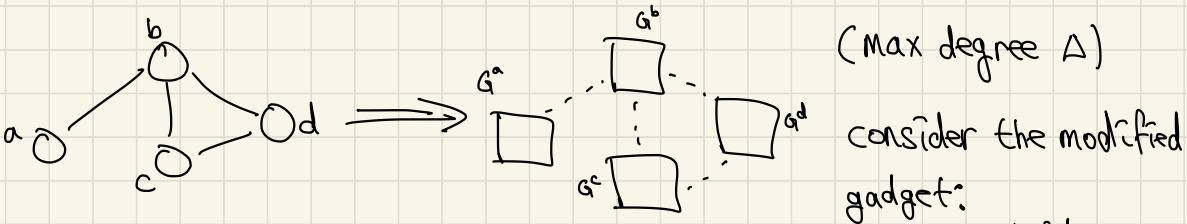
(iv) if $\lambda > \lambda_c(\Delta)$ then (p^*, p^*) is a saddle point and (p_1, p_2) and (p_2, p_1) are two global maxima. (ind. sets are likely to be lop-sided)

Fact*: w.h.p. over the random bipartite graph G of degree Δ ,

$$Z_{G_i}^{\alpha\beta}(\lambda) > \frac{1}{n} E[Z_{G_i}^{\alpha\beta}(\lambda)] \quad \forall \alpha, \beta \text{ not close to } 0, 1.$$

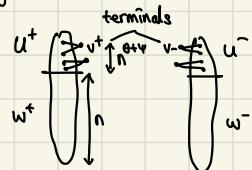
Thm) Glauber Dynamics for ind. sets with $\lambda > \lambda_c(\Delta)$ has mixing time $\exp(n)$ w.h.p. over G . (Intuition: conductance b/w $p_1 \leftrightarrow p_2$ is small)

Thm) Sly's Reduction: H (input for MaxCut) $\rightarrow H^G$ (input for sampling ind. sets)



we have n^θ trees with n^ψ vertices each in terminals.

also, $0 < \theta + \psi < 1$.



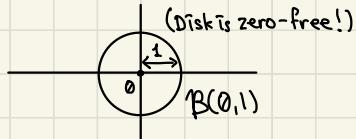
Fact: For any Δ and $\lambda > \lambda_c(\Delta)$, \exists constants θ, ψ with $\theta + \psi < 1$ s.t.

(i) $\Pr_{\sigma}[S(\sigma) = +] \geq \frac{1}{2}$ and $\Pr_{\sigma}[S(\sigma) = -] \geq \frac{1}{2}$, (S is the \pm phase)

(ii) Conditioned on phase S , marginal of Gibbs dist. on terminals is very close to a product dist. with parameters p_1, p_2 . i.e.,

$$\max_{\sigma_V} \left\{ \frac{\pi(\sigma_V | S=+)}{Q_V^+(\sigma_V)} - 1 \right\} \leq n^{-2\theta}. \quad [\text{deferred to notes due to time}]$$

Lee-Yang Theorem



$Z(\lambda) = \sum_k a_k \lambda^k$ where $a_k \geq 0$. What if $\lambda \in \mathbb{C}$?

Ferromagnetic Ising Model $Z_G(\lambda, \mu) = \sum_{S \in V} \lambda^{|E(S, \bar{S})|} \cdot \mu^{|S|}$ where $\lambda \in [0, 1]$, $\mu \geq 0$.

$\hookrightarrow Z(\mu) = \sum a_k \mu^k$ where $a_k = \sum_{|S|=k} \lambda^{|E(S, \bar{S})|}$. Let $f(\mu) := \log Z(\mu)$.

Claim: $\forall \varepsilon \leq \frac{1}{4}, \mu$, if $|\tilde{f} - f(\mu)| \leq \varepsilon$, then $|\exp(\tilde{f}) - Z(\mu)| \leq 4\varepsilon Z(\mu)$.

Proof: $|\exp(\tilde{f}) - Z(\mu)| = |\exp(\tilde{f}) - \exp(f(\mu))| = \exp(f(\mu)) |\exp(\tilde{f} - f(\mu)) - 1| \leq \exp(f(\mu)) \cdot 4\varepsilon \leq 4\varepsilon Z(\mu)$ (Geometric fact!)

Taylor expansion of $f(\mu) = \log Z(\mu)$ around $\mu=0$: $f(\mu) = \sum_{i=0}^{\infty} \frac{f^{(i)}(0) \mu^i}{i!}$.

Truncate it after m terms, $f_m(\mu) = \sum_{i=0}^m \frac{f^{(i)}(0) \mu^i}{i!}$. $|f_m(\mu) - f(\mu)| \leq \varepsilon$?

Write $Z(\mu) = \prod_{i=1}^n (1 - \frac{\mu}{r_i})$ where $\{r_i\}$ are complex roots of Z s.t. $|r_i| \geq 1$.

$\hookrightarrow f(\mu) = \sum_{i=1}^n \log(1 - \frac{\mu}{r_i}) = - \sum_{i=1}^n \sum_{j=0}^{\infty} \frac{1}{j} \left(\frac{\mu}{r_i}\right)^j$.

$|f_m(\mu) - f(\mu)| \leq n \cdot \sum_{j=m+1}^{\infty} \frac{(\mu/\epsilon)^j}{j} \leq n \frac{(\mu/\epsilon)^{m+1}}{m+1} \sum_{j=0}^{\infty} (\mu/\epsilon)^j = n \frac{(\mu/\epsilon)^{m+1}}{(m+1)(1 - \mu/\epsilon)}$.

To achieve $|f_m(\mu) - f(\mu)| \leq \frac{\varepsilon}{4}$, $m \geq \frac{1}{\log(\mu/\epsilon)} \left(\log \frac{4n}{\varepsilon} + \log \frac{1}{1 - \mu/\epsilon} \right) = O(\log(\mu/\epsilon))$.

Algorithm: compute $f_m(\mu)$ for $m = C \cdot \log(\mu/\epsilon)$. Output $\exp(f(\mu))$.

Computing $f_m(\mu)$: $f'(\mu) = \frac{1}{Z(\mu)} \frac{dZ(\mu)}{d\mu} \rightarrow \frac{dZ(\mu)}{d\mu} = Z(\mu) f'(\mu)$.

$$Z^{(m)}(0) = \sum_{j=0}^{m-1} \binom{m-1}{j} Z^{(j)}(0) f^{(m-j)}(0) \quad [\text{exercise; by induction}]$$

↪ Non-singular triangular system \Rightarrow can compute all $f^{(i)}(0)$ w.r.t. $\{Z^{(j)}(0)\}_{j=0}^m$

Finding $Z^{(j)}(0)$ is exhaustive enumeration of $\binom{n}{j} = O(n^{\log(n)})$ subsets in G

\Rightarrow quasi-polynomial algorithm (if Δ is bounded, polynomial)

$$\text{Reminder: } Z_G(\lambda, \mu) = \sum_{S \subseteq V} \lambda^{|E(S, \bar{S})|} \cdot \mu^{|S|} = \sum_{|S|=k} a_k \mu^k.$$

Thm) Lee-Yang: For any finite, connected graph G and fixed $\lambda < 1$, all the zeros of the Ising model partition function on \mathbb{C} lie on $B(0, 1)$.

Corollary: \exists FPTAS for $Z_G(\lambda, \mu) \forall G, \forall \lambda, \mu \neq 1$. (b/c $Z(\mu) = \mu^n Z(\mu')$).

Proof: [Asano] write $Z_G(\mu_1, \dots, \mu_n) := \sum_{S \subseteq V} \lambda^{|E(S, \bar{S})|} \cdot \prod_{i \in S} \mu_i$. Say that Z_G is "Lee-Yang" if $\forall i, |\mu_i| > 1 \Rightarrow Z_G(\mu_1, \dots, \mu_n) \neq 0$. S.t.p. Z_G is L-Y.

Structural Induction on G : (i) $\xrightarrow{G \cong 2} Z(\mu_1, \mu_2) = \mu_1 \mu_2 + \lambda(\mu_1 + \mu_2) + 1$.

Suppose $Z(\mu_1, \mu_2) = 0 \rightarrow |\mu_2| = \left| \frac{1 + \lambda \mu_1}{\lambda + \mu_1} \right| \rightarrow \text{if } |\mu_1| > 1, |\mu_2| \leq 1$.

(ii) For vertex-disjoint G, H , if $Z_G \& Z_H$ are both L-Y, $Z_{G \oplus H}$ is also L-Y

since roots of $Z_{G \oplus H}$ are union of roots of $Z_G \& Z_H$. $\xrightarrow{G \Delta \rightarrow H \setminus \setminus}$

(iii) Merging: s.p.s. by induction that Z_G is L-Y. Fix all μ_3, \dots, μ_n s.t. $|\mu_i| > 1$

$$\xrightarrow{G \xrightarrow{\mu_1} G'} \xrightarrow{G' \xrightarrow{\mu_2} G} \quad \forall 3 \leq i \leq n. \quad Z_G(\mu_1, \mu_2) = A\mu_1 \mu_2 + B\mu_1 + C\mu_2 + D.$$

We know that Z_G is L-Y by induction. Thus the above cannot be 0 if $|\mu_1|, |\mu_2| > 1$. Hence roots of $A\mu^2 + (B+C)\mu + D$ have magnitude ≤ 1 . Hence magnitude of product of these roots is $|D/A| \leq 1$. (Technicality: show $A \neq 0$).

$$Z_G(\mu) = A\mu + D \rightarrow \text{root of this is } \mu = -D/A, \text{ so } |\mu| < 1 \dots$$

* this does not work with $\mu \xrightarrow{\lim} 1$ (cannot approximate zero-field)

// End of official lectures, below are Student Presentations

Langevin Algorithm for Log-Concave [Altschuler & Talwar, COLT 23']

Continuous energy function $f: \mathbb{R}^d \rightarrow \mathbb{R}$, sample $x \in \mathbb{R}^d$ w.p. $\propto e^{-f(x)}$ $\rightarrow \pi$

Assumptions on f : f is convex and M -smooth. $\rightarrow \lambda_{\max}(\nabla^2 f) \leq M$
 $\forall x, y, \|\nabla f(x) - \nabla f(y)\| \leq M \|x - y\|$

Motivation: Sampling from / Approximating volume of convex body

Langevin Diffusion: Start from $x_0 \in \mathbb{R}^d$. Consider the following diff eq,
 $dx_t = -\nabla f(x_t) \cdot dt + \sqrt{2} dB_t$. $x_t \rightarrow \pi$ as $t \rightarrow \infty$.

Langevin Algorithm (Discrete): step size $\eta > 0$, convex subset $K \subseteq \mathbb{R}^d$.

$X_{t+1} = \underbrace{\Pi_K(X_t + \eta \cdot \nabla f(X_t) + Z_t)}_{\text{projection to } K} \text{ where } Z_t \sim N(0, 2\eta \cdot I_d)$

Question: How fast is T_{mix} for Langevin Algorithm?

Main Thm) Let $K \subseteq \mathbb{R}^d$ with diameter D . Let $f: \mathbb{R}^d \rightarrow \mathbb{R}$ be a convex & M -smooth potential. Assume

$$\begin{array}{ccc} LA_\eta & \xrightarrow{\substack{t \rightarrow \infty \\ \downarrow \eta \rightarrow 0}} & \Pi_\eta \\ \downarrow \eta \rightarrow 0 & & \downarrow \eta \rightarrow 0 \\ LD & \xrightarrow{\substack{t \rightarrow \infty}} & \Pi \end{array}$$

step size $\eta < \frac{2}{M}$. Consider the of above LA_η . $\underline{T_{\text{mix}}(K)}(\frac{1}{4}) \leq O(\frac{D^2}{\eta})$.

Remarks: 1) $O(\frac{D^2}{\eta})$ is also a lower bound! Consider a 1-D line.

2) Stochastic update $f = \frac{1}{n} \sum_i f_i$ also works WLOG.

3)

LA

GD

Task: sample x w.p. $\propto e^{-f(x)}$ find $\min_x \{f(x)\}$

Iteration: $x_t - \eta \cdot \nabla f(x_t) + N(0, 2\eta I_d)$ $x_t - \eta \cdot \nabla f(x_t)$

Error bound: $e^{-\frac{D^2}{\eta t}}$ $\frac{\eta t}{D^2}$

Def) α -Contractivity: $\phi: \mathbb{R}^d \rightarrow \mathbb{R}$ s.t. $\forall x, y \in \mathbb{R}^d$, $\|\phi(x) - \phi(y)\| \leq \alpha \cdot \|x - y\|$

Obs) GD step and projection step are 1-contractive! (uses smoothness)

Consider $x_0, x'_0 \in K$. Run $x_0 \rightarrow \dots \rightarrow x_t, x'_0 \rightarrow \dots \rightarrow x'_t$ with LA_η .

Claim: $TV \leq \sqrt{K \cdot KL}$ for this coupling.

1) Data Processing Inequality: $x \sim \mu, y \sim \nu$. Then $D_{KL}(\phi(x) \parallel \phi(y)) \leq D_{KL}(x \parallel y)$.

2) DPI $\Rightarrow D_{KL}(x_t \parallel y_t) \leq D_{KL}(x_t \parallel z_t) + D_{KL}(z_t \parallel y_t)$.

3) Shifted KL $D_{KL}^{(z)}(X||y) := \inf_{\substack{(x, x') \sim \pi}} D_{KL}(x' || y)$ where $W_\infty(x, x') \leq z$. $\overbrace{\|x - x'\|_2}$

↪ monotonically decreases with z .

Claims: 1') for 1-contractive ϕ , $D_{KL}^{(z)}(\phi(x) || \phi(y)) \leq D_{KL}^{(z)}(x || y)$.

2') $\forall a \geq 0$, $D_{KL}^{(z)}(X+z || y+z) \leq D_{KL}^{(z+a)}(X || y) + \frac{a^2}{2\sigma^2}$ where $z \sim N(0, \sigma^2 I)$.

Proof: 1') $(x, x') \sim (\mu, \mu')$. $\|x - x'\| \leq z \Rightarrow \|\phi(x) - \phi(x')\| \leq z$.

$$\Rightarrow D_{KL}^{(z)}(\phi(x) || \phi(y)) \leq D_{KL}(\phi(x') || \phi(y)) \leq D_{KL}(x' || y) = D_{KL}^{(z)}(x || y). //$$

$$2') D_{KL}(X+z || y+z) = D_{KL}(\overbrace{X-w}^{x'} + w + z || y+z) \leq D_{KL}(x', w+z || y, z)$$

$$= D_{KL}(x' || y) + \mathbb{E}_{x', w, y} [D_{KL}(w+z || z) | x', x, y] = D_{KL}(x' || y) + \frac{\mathbb{E}[\|w\|_2^2]}{2\sigma^2} \xrightarrow{a^2} //$$

Consider the coupling (y_t, y'_t) with ϕ . $\begin{cases} y_{t+1} = \phi_t(y_t) + z_t \\ y'_{t+1} = \phi_t(y'_t) + z_t \end{cases}$

$$\text{Prop: } KL(y_T || y'_T) \leq \frac{\|y_0 - y'_0\|_2^2}{2\sigma^2 T} \leq \frac{D^2}{2\sigma^2 T}.$$

MCMC for Computing Volumes [Cousins, Vempala '18']

Problem: Compute volume of convex body $K \in \mathbb{R}^n$ given a membership oracle

↳ requires $(1-\epsilon) \text{Vol}(K) \leq V \leq (1+\epsilon) \text{Vol}(K)$ w.p. $1-p$.

Complexity measured by # of oracle calls

$$\text{DFK: } \tilde{O}(\text{poly}(n, 1/\epsilon)) = \tilde{O}\left(\frac{n^3}{\epsilon^2} + n^4\right)$$

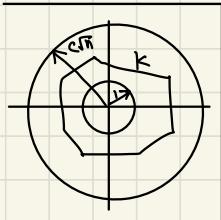
$$E_{x \sim K}[\|x\|^2] = O(n)$$



Assume that $B_n \subseteq K \subseteq C\sqrt{n}B_n$, that is, K is "well-rounded".

↑ takes $\tilde{O}(n^2)$

Ideally, K should be isotropic, $E_{x \sim K}[xx^T] = I_n$. Then, the problem is easy due to concentration bound decaying exponentially out of $C\sqrt{n}B_n$.



Define $F(\sigma^2) := \int_{\mathbb{R}^n} f(\sigma^2, x) dx$ where

$$f(\sigma^2, x) = \mathbf{1}\{x \in K\} \cdot \exp\left(-\frac{\|x\|^2}{2\sigma^2}\right). \text{ Then } F(\infty) = \int_K dx = \text{Vol}(K).$$

Strategy: incrementally approximate $\frac{F(\sigma_{i+1}^2)}{F(\sigma_i^2)}$.

Algorithm Sketch: consider some σ^2 schedule of $O(\frac{1}{n}) \rightarrow O(n) \rightarrow \infty$.

Sample $x_j \sim \mu_i$ where $\mu_i(x) := \frac{f(\sigma_i^2, x)}{F(\sigma_i^2)}$. estimate $\frac{F(\sigma_{i+1}^2)}{F(\sigma_i^2)}$ with

Monte Carlo: $\frac{1}{k} \sum_{j=1}^k \frac{f(\sigma_{i+1}^2, x_j)}{f(\sigma_i^2, x_j)}$. Then, $V = F(\infty) \approx \prod \frac{F(\sigma_{i+1}^2)}{F(\sigma_i^2)} \cdot F(\sigma_0^2)$.

To sample $x_j \sim \mu_i$ efficiently, we use a Speedy Walk that has conductance $\phi = \Omega\left(\frac{\delta}{\sigma\sqrt{n}}\right)$ where $\delta \leq \frac{\sigma}{8\sqrt{n}}$. [Proof omitted]

Gaussian Cooling: σ^2 scheduled in two phases, $O(\frac{1}{n}) \rightarrow 1$, $1 \rightarrow O(n)$.

phase 1 $\rightarrow \sigma_{i+1}^2 \propto \sigma_i^2(1 + \frac{1}{n})$, phase 2 $\rightarrow \sigma_{i+1}^2 \propto \sigma_i^2(1 + \frac{\sigma_i^2}{n})$

Analysis: Let $Y_j := \frac{f(\sigma_{i+1}^2, x_j)}{f(\sigma_i^2, x_j)}$, $x_j \sim M_i$, $\hat{Y} = \frac{1}{k} \sum_{j=1}^k Y_j$. $E[Y] = \frac{F(\sigma_{i+1}^2)}{F(\sigma_i^2)}$. (unbiased)

Y is concentrated around $E[Y]$ with Chebyshev, and $O(\frac{1}{\epsilon^2})$ samples suffice.

Each sampling takes mixing $T \sim \max\{1, \sigma^2\} \cdot n^2$. First phase requires

$\tilde{O}(n)$ updates $\times \tilde{O}(\frac{1}{\epsilon^2})$ samples $\times \tilde{O}(n^2)$ sampling time $= \tilde{O}(\frac{n^3}{\epsilon^2})$.

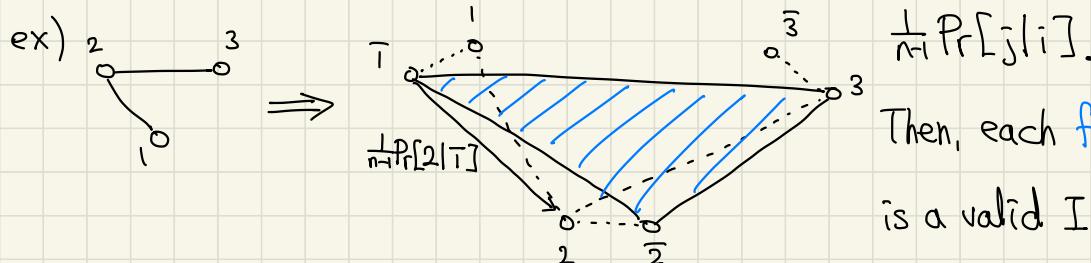
Second phase has $\tilde{O}(\frac{1}{\sigma^2})$ updates $\times \tilde{O}(\frac{1}{\epsilon^2})$ samples $\times \tilde{O}(\sigma^2 \cdot n^2)$ sample time $= \tilde{O}(\frac{n^3}{\epsilon^2})$. \Rightarrow Total runtime $\tilde{O}(n^4 + \frac{n^3}{\epsilon^2})$ (isotropic transformation & SW)

Spectral Independence [ALG]

Problem: $\mu: 2^{[n]} \rightarrow [0, 1]$, hard-core model, Glauber dynamics.

pick i u.a.r., move to $S \cup \{i\}$ w.p. $\frac{\mu(S \cup \{i\})}{\mu(S \cup \{i\}) + \mu(S \setminus \{i\})}$.

Simplicial Complex: X^n , make two nodes i, \bar{i} for each i , representing inclusion/exclusion of i in the IS. Then draw valid edges w.p.



Generally, conditioning on m states has transitions $\frac{1}{n-m-1} \Pr[j | i_1, \dots, i_m]$.

We view transitions as down-up $P_d^v(\sigma, \sigma')$.

\downarrow (pure) d -dimensional SC

Local Expander: X^M is (x_0, \dots, x_{d-2}) LE if conditioned on i items,

$\lambda_2(P_{n-i}) \leq \alpha_i$. By [Alev & Lau 20'], $1 - \lambda_2(P_d^v) \leq \frac{1}{d} \prod_{k=0}^{d-2} (1 - \alpha_k)$.

$\rightarrow (n \times n)$

Influence Matrix: $\Psi_n(i, j) := \Pr[j | i] - \Pr[j | \bar{i}]$.

η -Spectral Independence: $\lambda_{\max}(\Psi_n) \leq \eta$.

$(\eta_0, \dots, \eta_{n-2})$ -SI: $\lambda_{\max}(\Psi_{n-1, S}) \leq \eta_j$, where $|S| = j$.

\hookrightarrow This allows $1 - \lambda_2(P_d^v) \leq \frac{1}{n} \prod_{k=0}^{n-2} \left(1 - \frac{\eta_k}{n-k-1}\right)$.

Quantum Metastability [BCV 25']

Background: consider spin system on $G(V, E)$, $|V| = n$.

Classical

configurations

$v \in V$ to $\{\pm 1\}$

state

$\sigma: v \rightarrow \{\pm 1\}$

Hamiltonian

$H(\sigma)$

Gibbs?

distribution $\pi(\sigma) \propto e^{-\beta H(\sigma)}$

Markov Generator

heat kernel $H_t(x|y) = e^{tL^* - P_I}$

Quantum

$v \in V$ to \mathbb{C}^2 $\text{tr}(p)=1$, PSD

\downarrow
quantum state $p \in \mathbb{C}^{2^n \times 2^n}$

H is an operator.

state $p \propto e^{-\beta H}$

Lindbladian L

Quantum Gibbs sampling? design \mathcal{L} s.t. $\mathcal{L}[\rho] = \emptyset$? $\simeq \emptyset$? (meta stable)

↳ Are meta stable states useful? Do they satisfy properties of p ?

Def) ε -metastability: $\|\mathcal{L}[\sigma]\|_1 \leq \varepsilon$.

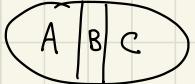
Prop: Fix $T > 0$, any initial state σ_0 . Then the time-averaged state

$\bar{\sigma}_T := \frac{1}{T} \int_0^T e^{t\mathcal{L}} [\sigma_0] dt$ is ε -metastable with $\varepsilon < \frac{2}{T}$.

Proof: equation of motion is $\frac{d}{dt} \sigma_t = \mathcal{L}[\sigma_t]$. $\rightarrow \sigma_T - \sigma_0 = \int_0^T \mathcal{L}[\sigma_t] dt$

$$= \mathcal{L}\left[\int_0^T \sigma_t dt\right] = T \cdot \mathcal{L}[\bar{\sigma}_T]. \rightarrow T \cdot \|\mathcal{L}[\bar{\sigma}_T]\|_1 \leq \|\sigma_T\| + \|\sigma_0\| \leq 2.$$

$$\Rightarrow I(A;C|B) = 0.$$

Classical (Global) Markov Property:  $\sigma \sim \pi$. $\Rightarrow \sigma_A \perp \sigma_C | \sigma_B$

Area Law: $S(x) := \sum_x p(x) \log(\frac{1}{p(x)})$, $I(x;y) := S(x) + S(y) - S(x,y) \leq S(x)$

$I(x;y,z) := I(x;y) - I(x,z|y)$. $\Rightarrow I(A;A_c)_\pi \leq |\partial A|$ (boundary of A)

Quantum Area Law: $S(p) := \text{Tr}[p \log p]$. $I(A;C|B)_p = 0$? not in local.

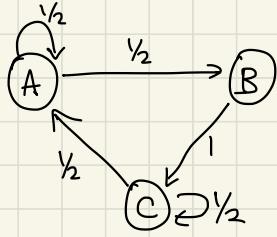
Thm 1) $\forall \varepsilon$ -metastable state σ , $\forall A \subseteq [n]$, $I(A;\bar{A})_\sigma \leq 2\beta \|2H\|_\infty + n^2 \varepsilon^\lambda e^{\mu|A|}$.

Local Approx. Markov Property: $I(A;C|B)_p$ exponentially decays, if or C is "small". something like $I(A;C|B) \lesssim e^{-\min(A,C)} \cdot e^{-\text{dist}(A,C)}$.

Thm 2) $\forall \varepsilon$ -MS σ , $\forall A \subseteq [n]$, $\exists R$ s.t. \forall noise N_A , $\|\sigma - R[N_A[\sigma]]\|_1 \leq n e^{\mu|A|} \varepsilon^\lambda$.

Bottom Line w/o proof: Metastability $\Rightarrow \dots \Rightarrow$ Locally Markov \Rightarrow Area Law

Coupling From the Past [Propp & Wilson '96']



Naïve MC: for $i \in [N]$, sample random seed r_0, \dots, r_t .

start from $M=2$ initial states, evolve with r_i .

first time they meet \rightarrow coupled, return that state.

Observe: if we collapse at state B at time A, then at time $(t-1)$, we must have collapsed at state A. \Rightarrow Naïve forward MC never outputs state B!

Let $f: \Omega \times [0,1] \rightarrow \Omega$ be a "global" coupling, i.e. $\forall x \in \Omega, r \sim U([0,1])$,

$\Pr[r|f(x,r)=y] = P(x,y)$. Fix a time t , then $f_t(x) := f(x, r_t)$.

From t_1 to t_2 : $F_{t_2}^{t_2}(x) := (f_{t_2} \circ \dots \circ f_{t_1})(x)$ (composition of f_t)

Prop: $\Pr[F_{t_1}^{t_2}(x)=y] = P^{t_2-t_1}(x,y)$. Proof: Induction.

Let T be the first time when $\forall x \in \Omega, |F_T^T(x)| = 1$.

\hookrightarrow Is $F_T^T(x) = \pi(x)$? No, but motivates coupling from past, $F_M^0(x)$!

Thm) F_M^0 has the same distribution as π .