# Normal Mixture Model, Gibbs/Metropolis-Hastings Sampler

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 $z_i \sim Bern(\frac{a}{a+b})$  where  $a = (p * f(x_i|\mu_1, \sigma_1))$  and  $b = ((1-p) * f(x_i|\mu_2, \sigma_2))$ . 1000  $z_i's$  are drawn each iteration.

```
p \sim beta(\sum z, n - \sum z)
```

 $\sigma_1 \sim inverse - gamma(\frac{\kappa_{n_1}}{2}, \frac{\kappa_{n_1}}{2}\sigma_{n_1}^2)$  where  $\kappa_{n_1} = \kappa_{0_1} + n_1$ ,  $\kappa_{0_1}$  is the prior sample size for population 1, and  $n_1$  is the sample size of the data for population 1

 $\sigma_2 \sim inverse - gamma(\frac{\kappa_{n_2}}{2}, \frac{\kappa_{n_2}}{2}\sigma_{n_2}^2)$  where  $\kappa_{n_2} = \kappa_{0_2} + n_2$ ,  $\kappa_{0_2}$  is the prior sample size for population 2, and  $n_2$  is the sample size of the data for population 2

 $\theta_1 \sim normal(\mu_{n_1}, \frac{\sigma_1^2}{\kappa_{n_1}})$  where  $\mu_{n_1} = \frac{\kappa_{0_1}}{\kappa_{n_1}} \mu_{0_1} + \frac{n_1}{\kappa_{n_1}} \bar{y_1}$ ,  $\mu_{0_1}$  is the prior mean for population 1, and  $\bar{y_1}$  is the mean of the data for population 1

 $\theta_2 \sim normal(\mu_{n_2}, \frac{\sigma_2^2}{\kappa_{n_2}})$  where  $\mu_{n_2} = \frac{\kappa_{0_2}}{\kappa_{n_2}} \mu_{0_2} + \frac{n_2}{\kappa_{n_2}} \bar{y_2}$ ,  $\mu_{0_2}$  is the prior mean for population 1, and  $\bar{y_2}$  is the mean of the data for population 1

If the Metropolis-Hastings algorihtm is used to calculate p:

We are attempting to sample from  $P(p|\sum z)$ .

The prior for p is:  $P(p) = \frac{\sin(5*\pi*p)^2}{c}$  where c is some normalizing constant.

As the  $z_i's$  are distributed the same as above,  $P(p|\sum z) = {n\choose \sum z}*p^{\sum z}*(1-p)^{n-\sum z}$ 

The proposal distribution for  $p^*$  is the beta distribution

We then calculate  $\log(r)$  as it is more stable than r.  $\log(r) = \log(\frac{P(p^*|\sum z)}{P(p^{(s)}|\sum z)} * \frac{J_s(p^{(s)}|p^*)}{J_s(p^*|p^{(s)})}) = \log(\frac{P(\sum z|p^*)P(p^*)}{P(\sum z|p^{(s)})P(p^{(s)})} * \frac{J_s(p^{(s)})}{J_s(p^*)}) = \sum(\log(P(z|p^*)) - \log(P(z|p^{(s)}))) + (\log(P(p^*)) - \log(P(p^{(s)}))) + (\log(J_s(p^{(s)})) - \log(J_s(p^*))).$  Note that both the c's and the  $\binom{n}{z}$  are cancelled out.

 $J_s(p)$  randomly samples from a beta distribution with parameters  $b = (\frac{p(1-p)}{v} - 1)(1-p)$ ,  $a = b\frac{p}{1-p}$ , and v is a constant (set at .0005)

We accept  $p^*$  as a new value with probability r. If rejected p is instead accepted.

```
#function of Gibbs/Metropolis-Hastings algorithm
gibbs.MH <- function(data, trials, use.MH){
   library(car)

#Dats is the data used for the algorithm
   #trials is number of trials to be done
#Use.MH determines whether or not to use a metropolis-hastings algorithm for p
#A 0 is no, anything else is yes

n <- length(data)
p<-0.5
p.star<-0.5</pre>
```

```
r.log < -1
#creates priors
prior.pop1 <- rnorm(pop1.prior.sampleSize, pop1.prior.trueMean, pop1.prior.trueSD)</pre>
prior.pop2 <- rnorm(pop2.prior.sampleSize, pop2.prior.trueMean, pop2.prior.trueSD)</pre>
#creates matrix where gibbs will be done
simlist <- matrix(rep(0,5*trials), ncol = 5)</pre>
colnames(simlist) <- c("p", "SD of Pop1", "SD of Pop2", "Mean of Pop1", "Mean of Pop2")</pre>
#saves value for later usage
pop1.prior.mean <- mean(prior.pop1)</pre>
pop1.prior.variance <- var(prior.pop1)</pre>
pop2.prior.mean <- mean(prior.pop2)</pre>
pop2.prior.variance <- var(prior.pop2)</pre>
#initialize first entry in gibbs so it will run correctly
#p sigma1 sigma2 theta1 theta2
simlist[1,1] <- .5
simlist[1,2] <- sqrt(pop1.prior.variance)</pre>
simlist[1,3] <- sqrt(pop2.prior.variance)</pre>
simlist[1,4] <- pop1.prior.mean</pre>
simlist[1,5] <- pop2.prior.mean</pre>
#The actual gibbs sampler
\#starts at 2 since we initialized values for when i=1
for(i in 2:trials){
  #max size will be z
  z <- numeric(n)</pre>
  a <- numeric(n)
  b <- numeric(n)
  j = 1
  #calculates a_j and b_j for each data_j
  #then calculates z_i's
  for (j in 1:n){
    a[j] = simlist[i-1,1]*dnorm(data[j],simlist[i-1,4],simlist[i-1,2])
    b[j] = (1-simlist[i-1,1])*dnorm(data[j],simlist[i-1,5],simlist[i-1,3])
    z[j] = rbinom(1,1,(a[j]/(a[j]+b[j])))
  #calculates p using beta distribution for gibbs
  if (use.MH == 0){
    simlist[i,1] \leftarrow rbeta(1,sum(z), n-sum(z))
  #calculates p ising metropolis hastings with beta as the proposal distribution
  } else {
    #metropolis hastings algorithm to calculate p
    #constant, larger values allow larger skips, but also can produce NaN
    variance.constant = .0005
    c = ((p.star*(1-p.star)/variance.constant) - 1)*(1 - p.star)
```

```
d = c*(p.star/(1-p.star))
  p.star = rbeta(1,d,c)
  #The algorithm ran into problems with really large and small values of p.star
  #Played around with these values, using 0.01 and 0.99 produced a weird shaped gamma for the next
  #doesnt impact calculations done later
  #One caveat is this method can't be used on a population where p is extrmeley small
  if (p.star < 0.03){
   p.star <-0.03
  if (p.star > 0.97){
   p.star <- 0.97
 hast.log <- log(dbeta(p.star,d,c)) - log(dbeta(p,d,c))</pre>
 r.log <- metro.log + hast.log</pre>
 if (log(runif(1)) < r.log){
   if (p.star != 0 & p.star != 1){
     p <- p.star
 simlist[i,1] <- p</pre>
pop1 <- numeric(n)</pre>
pop2 <- numeric(n)
#dividing data into pop1 and pop2 as best as we can at the moment
for (j in 1:n){
 pop1[j] = z[j]*data[j]
 pop2[j] = ((1-z[j])*-1)*data[j]
 pop2[j] = -1*pop2[j]
#gets rid of zeros in list
pop1 <- pop1[pop1!=0]</pre>
pop2 <- pop2[pop2!=0]</pre>
#calculates length, used later
n1 <- length(pop1)</pre>
n2 <- length(pop2)
#calculates SD for population 1
#first calculated precision with gamma, then variance, finally SD
pop1.Vn <- pop1.prior.sampleSize + n1</pre>
pop1.sigmaN <- ((pop1.prior.sampleSize*pop1.prior.variance) + ((n1 - 1)*var(pop1)) + (((pop1.prior.</pre>
prec1 <- rgamma(1,pop1.Vn/2,(pop1.Vn * pop1.sigmaN)/2)</pre>
variance1 = 1/prec1
simlist[i,2] <- sqrt(variance1)</pre>
#similar to above except uses population 2 parameters
```

```
pop2.Vn <- pop2.prior.sampleSize + n2
pop2.sigmaN <- ((pop2.prior.sampleSize*pop2.prior.variance) + ((n2 - 1)*var(pop2)) + (((pop2.prior.prec2 <- rgamma(1,pop2.Vn/2,(pop2.Vn * pop2.sigmaN)/2)
variance2 = 1/prec2
simlist[i,3] <- sqrt(variance2)

#Calculates mean of population 1
mu1 <- ((pop1.prior.sampleSize*pop1.prior.mean)+(n1*mean(pop1)))/(pop1.prior.sampleSize+n1)
var1 <- simlist[i,2]/(sqrt(pop1.prior.sampleSize+n1))
simlist[i,4] <- rnorm(1,mu1,var1)

#same as above except uses population 2 parameters
mu2 <- ((pop2.prior.sampleSize*pop2.prior.mean)+(n2*mean(pop2)))/(pop2.prior.sampleSize+n2)
var2 <- simlist[i,3]/(sqrt(pop2.prior.sampleSize+n2))
simlist[i,5] <- rnorm(1,mu2,var2)

}
return (simlist)
}</pre>
```

Function that outputs useful results and plots

```
printResults <- function(simlist){</pre>
  #plots and gives means of parameters
  plot(simlist[,1], main = "Chain for p", type = "l", ylab = "P")
  plot(density(simlist[,1]), main = "Distribution of p")
  print("The estimated value of p is")
  print(mean(simlist[,1]))
  plot(density(simlist[,2]),main = "Distribution of SD of Pop1")
  print("The estimated SD of population 1 is")
  print(mean(simlist[,2]))
  plot(density(simlist[,3]),main = "Distribution of SD of Pop2")
  print("The estimated SD of population 2 is")
  print(mean(simlist[,3]))
  plot(density(simlist[,4]),main = "Distribution of Mean of Pop1")
  print("The estimated mean of population 1 is")
  print(mean(simlist[,4]))
  plot(density(simlist[,5]),main = "Distribution of Mean of Pop2")
  print("The estimated mean of population 2 is")
  print(mean(simlist[,5]))
```

This chunck runns a Gibbs sampler on data comprised of kids heights and adult heights

```
heightData = read.csv("height.csv")
adultHeight = heightData$Human.Height..University.of.Tuebingen..2015....cm.[heightData$Year > 1950]
kidHeight = read.csv("kidHeight.csv")
data = c(adultHeight,kidHeight$height)
plot(density(data))
truep <- length(kidHeight$height)/(length(adultHeight) + length(kidHeight$height))</pre>
```

```
pop1.trueMean <- mean(kidHeight$height)
pop1.trueSD <- sqrt(var(kidHeight$height))

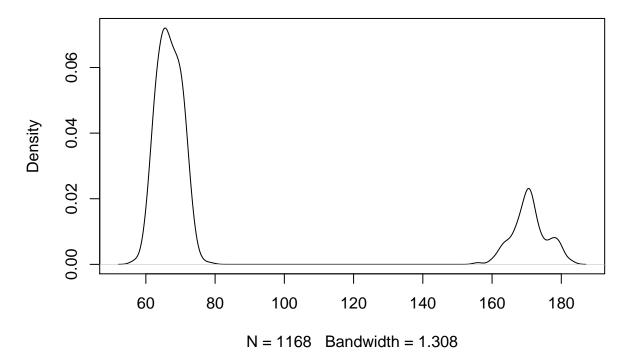
pop2.trueMean <- mean(adultHeight)
pop2.trueSD <- sqrt(var(adultHeight))

#initialize values for prior for population 1
pop1.prior.sampleSize <- 10
pop1.prior.trueMean <- 80
pop1.prior.trueSD <- 20

#initialize values for prior for population 2
pop2.prior.sampleSize <- 10
pop2.prior.trueMean <- 160
pop2.prior.trueSD <- 20

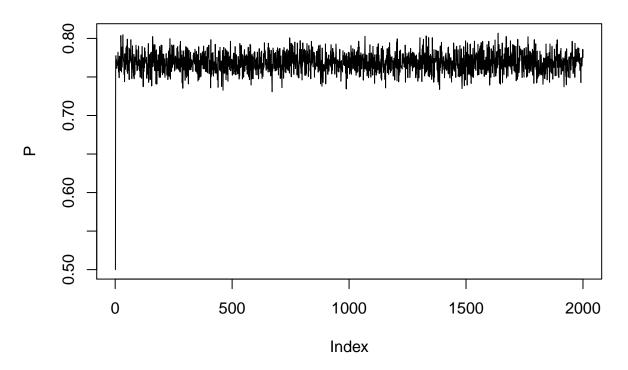
simlistHeight <- gibbs.MH(data,2000,0)</pre>
```

#### density.default(x = data)

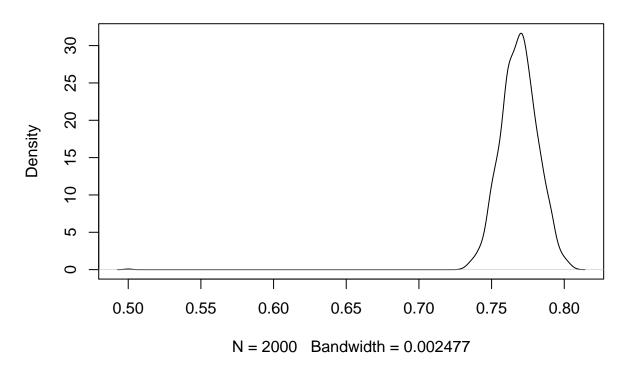


printResults(simlistHeight)

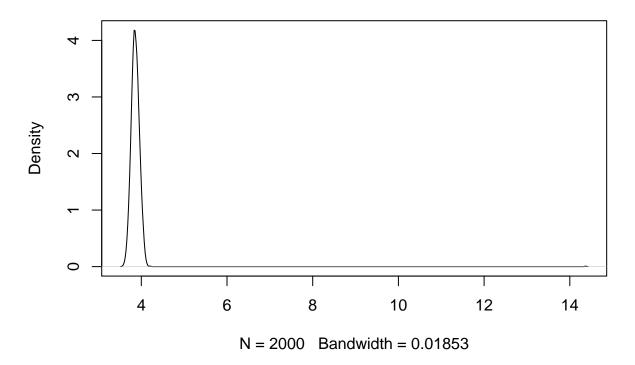
# Chain for p



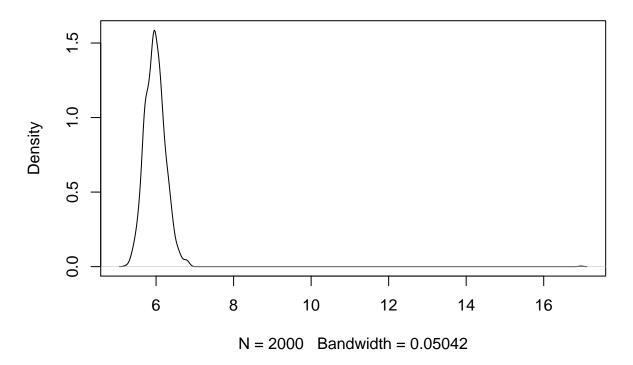
### Distribution of p



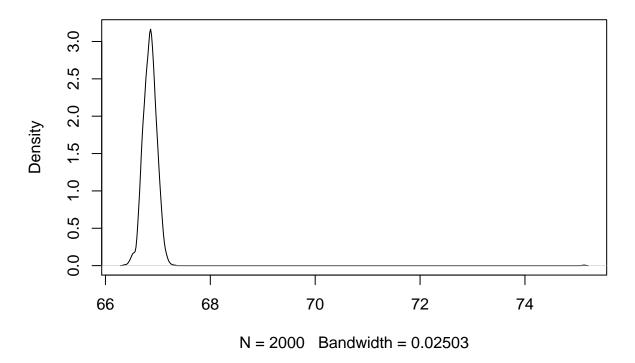
- ## [1] "The estimated value of p is" ## [1] 0.7688226



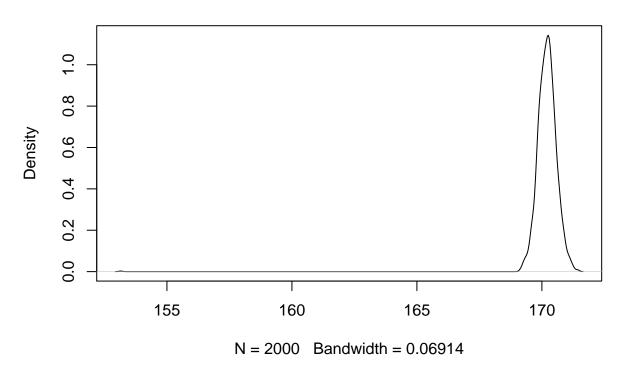
- ## [1] "The estimated SD of population 1 is"
  ## [1] 3.865781



<sup>## [1] &</sup>quot;The estimated SD of population 2 is" ## [1] 5.977574



- ## [1] "The estimated mean of population 1 is" ## [1] 66.85254



```
## [1] "The estimated mean of population 2 is"
## [1] 170.2034

print("The true value of p is")

## [1] "The true value of p is"

truep

## [1] 0.7688356

print("The true value of the SD for population 1 is")

## [1] "The true value of the SD for population 1 is"

sqrt(var(kidHeight$height))

## [1] 3.582918

print("The true value of the SD for population 2 is")

## [1] "The true value of the SD for population 2 is"

sqrt(var(heightData$Human.Height..University.of.Tuebingen..2015....cm.))

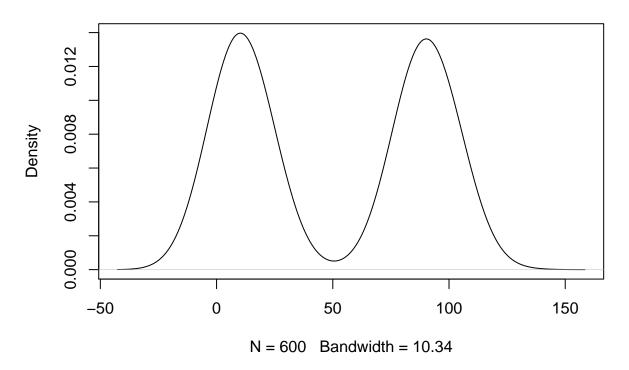
## [1] 4.638687

print("The true value of the mean of population 1 is")
```

## [1] "The true value of the mean of population 1 is"

```
mean(kidHeight$height)
## [1] 66.76069
print("The true value of the mean of population 2 is")
## [1] "The true value of the mean of population 2 is"
mean(heightData$Human.Height..University.of.Tuebingen..2015....cm.)
## [1] 167.6473
This chunk runs a combination Gibbs sampler/Metropolis-Hastings algorithm on simulated normal data
obs1 <- 300
obs2 <- 300
#initialize values for population 1
pop1.trueMean <- 10</pre>
pop1.trueSD <- 10
#initialize values for population 2
pop2.trueMean <- 90</pre>
pop2.trueSD <- 10
#creates data, then binds together to create normal mixture model
samp1 <- rnorm(obs1,pop1.trueMean,pop1.trueSD)</pre>
samp2 <- rnorm(obs2,pop2.trueMean,pop2.trueSD)</pre>
data <- c(samp1,samp2)</pre>
plot(density(data))
```

### density.default(x = data)

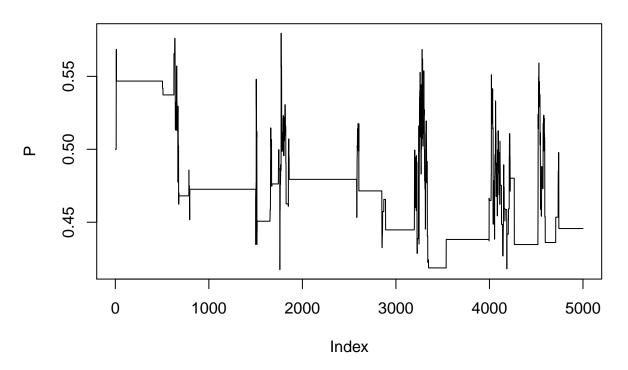


```
#initialize values for prior for population 1
pop1.prior.sampleSize <- 10
pop1.prior.trueMean <- 20
pop1.prior.trueSD <- 20

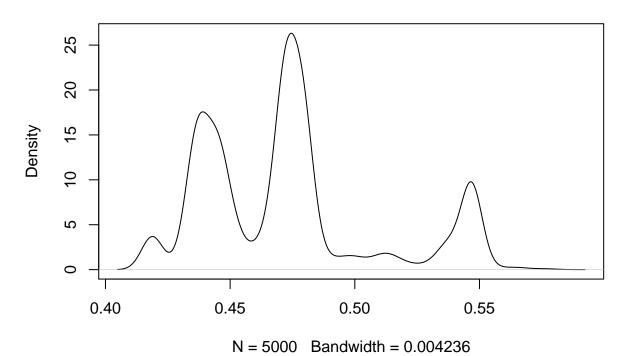
#initialize values for prior for population 2
pop2.prior.sampleSize <- 10
pop2.prior.trueMean <- 80
pop2.prior.trueSD <- 20

simlistSim <- gibbs.MH(data,5000,1)
printResults(simlistSim)</pre>
```

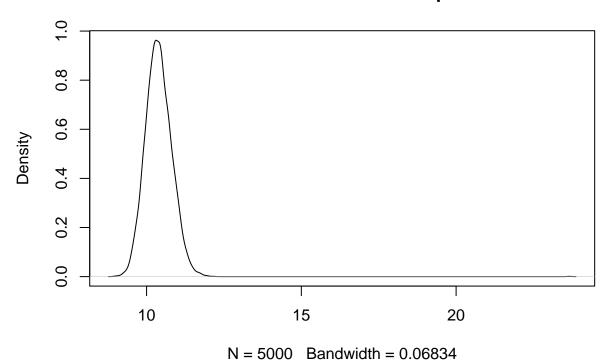
# Chain for p



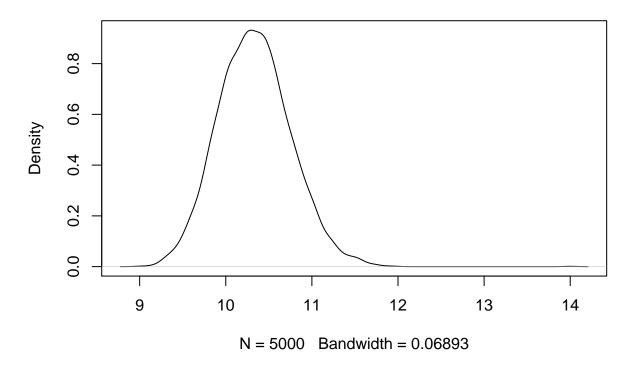
### Distribution of p



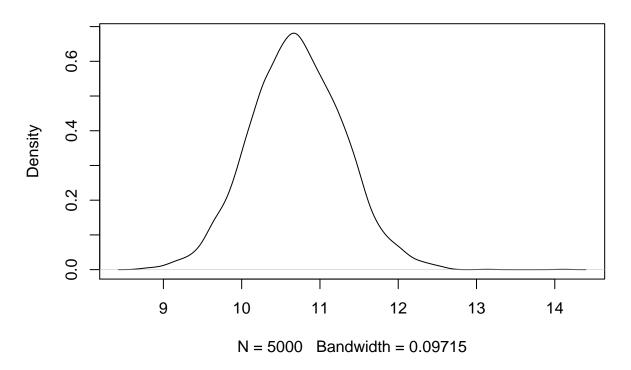
## [1] "The estimated value of p is" ## [1] 0.472849



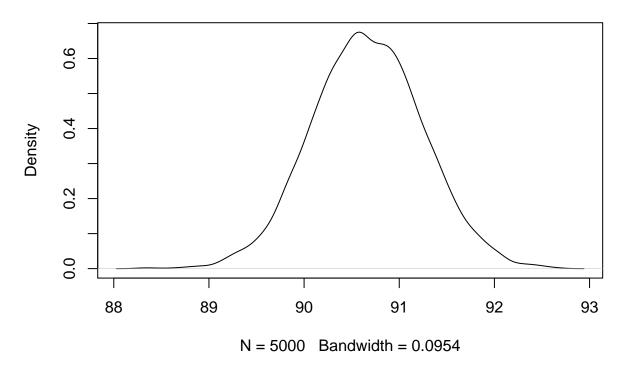
- ## [1] "The estimated SD of population 1 is" ## [1] 10.39227



- ## [1] "The estimated SD of population 2 is" ## [1] 10.33467



- ## [1] "The estimated mean of population 1 is" ## [1] 10.7051



```
## [1] "The estimated mean of population 2 is"
## [1] 90.66554

print("The true value of p is")

## [1] "The true value of p is"

obs1/(obs1+obs2)

## [1] 0.5

print("The true value of the SD for population 1 is")

## [1] "The true value of the SD for population 1 is"

pop1.trueSD

## [1] 10

print("The true value of the SD for population 2 is")

## [1] "The true value of the SD for population 2 is"

pop2.trueSD

## [1] 10

print("The true value of the mean of population 1 is")
```

## [1] "The true value of the mean of population 1 is"

```
pop1.trueMean

## [1] 10

print("The true value of the mean of population 2 is")

## [1] "The true value of the mean of population 2 is"

pop2.trueMean

## [1] 90
```