

Statistical learning theory

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1 Proposed exercise lecture 02

1.1 Home Exercise

Demonstrate that the MAP (Maximum A Posteriori) classification rule is quadratic with respect to x^* . We have the following:

$$\hat{y}^* = \arg \max_{k=1,\dots,K} \mathbb{P}(Y = k|X = x^*) \quad (1)$$

We compute the probability:

$$\begin{aligned} \arg \max_{k=1,\dots,K} \mathbb{P}(Y = k|X = x^*) &\stackrel{\text{Bayes}}{=} \frac{\mathbb{P}(Y = k)\mathbb{P}(X = x^*|Y = k)}{\mathbb{P}(X = x^*)} \\ &\propto \mathbb{P}(Y = k)\mathbb{P}(X = x^*|Y = k) \\ &\propto \hat{\Pi}_k \mathcal{N}(x^*, \hat{\mu}_k, \hat{\Sigma}_k) \end{aligned}$$

$$\begin{aligned}
\hat{y}^* &= \arg \max_{k=1, \dots, K} \hat{\Pi}_k \mathcal{N}(x^*, \hat{\mu}_k, \hat{\Sigma}_k) \\
&= \arg \max_{k=1, \dots, K} \log [\hat{\Pi}_k] + \log [\mathcal{N}(x^*, \hat{\mu}_k, \hat{\Sigma}_k)] \\
&= \arg \max_{k=1, \dots, K} \log [\hat{\Pi}_k] + \log \left[\frac{1}{(2\pi)^{N/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} (x^* - \mu)^\top \Sigma^{-1} (x^* - \mu)} \right] \\
&= \arg \max_{k=1, \dots, K} \underbrace{\log [\hat{\Pi}_k]}_{\text{Constant w.r.t. } x^*} + \underbrace{\left[-\frac{1}{2} (x^* - \mu)^\top \Sigma^{-1} (x^* - \mu) \right]}_{\text{Quadratic in } x^*} \\
&\quad - \underbrace{\log [(2\pi)^{N/2} |\Sigma|^{1/2}]}_{\text{Constant w.r.t. } x^*}
\end{aligned}$$

The second term is Quadratic in x^* , the others are constant w.r.t. x^* , which completes the proof.

2 Proposed exercise lecture 05

2.1 In-class Exercise

Compute the derivative with respect to the means.

$$\begin{aligned}
\nabla_{\mu_k} \mathcal{L}(\theta) &= \nabla_{\mu_k} \sum_{n=1}^N \log \left[\sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma_k) \right] \\
&= \sum_{n=1}^N \frac{\pi_k \mathcal{N}(x_n | \mu_k, \Sigma_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma_k)} \\
&= \sum_{n=1}^N \frac{\pi_k \mathcal{N}(x_n | \mu_k, \Sigma_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(x_n | \mu_k, \Sigma_k)} \Sigma^{-1} (x_n - \mu_k)
\end{aligned}$$

3 Proposed exercise lecture 09

3.1 In-class Exercise

Write the observed log-likelihood.

$X_1 \sim \mathcal{N}(\mu_1, \Sigma_{11})$ observed, $X_2 \sim \mathcal{N}(\mu_2, \Sigma_{22})$ missing.

$$\begin{aligned}\mathcal{L}_{obs}(\theta, X) &= \int \mathcal{L}_{full}(\theta, X) dX^{mis} \\ &= \int \mathbb{P}(X_1, X_2, \theta) dX_2 \\ &= \int \mathbb{P}(X_1, X_2, \theta) dX_2 \\ &= \int \mathbb{P}(X_1) \mathbb{P}(X_2 | X_1) dX_2 \\ &= \mathbb{P}(X_1) \int \mathbb{P}(X_2 | X_1) dX_2 \\ &= \prod_{i=1}^n \mathbb{P}(X_{i1}) \int \mathbb{P}(X_2 | X_1) dX_2\end{aligned}$$

$$\begin{aligned}\log \mathbb{P}(X_{i1}) &= \frac{n}{2} \log(\sigma_{11}^2) - \frac{1}{2} \sum_{i=1}^n \frac{(X_{i1} - \mu_1)^2}{\sigma_{11}^2} \\ &= \frac{n}{2} \log(\sigma_{22} - \frac{\sigma_{21}^2}{\sigma_{11}}) - \frac{1}{2} \sum_{i=1}^n \frac{(X_{i2} - (\mu_2 + \frac{\sigma_{21}}{\sigma_{11}}(X_{i1} - \mu_1)))^2}{(\sigma_{22} - \frac{\sigma_{21}^2}{\sigma_{11}})^2}\end{aligned}$$