



MSC. DATA SCIENCE & ARTIFICIAL INTELLIGENCE

STATISTICAL INFERENCE - PRACTICE

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Final assignment

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1 Exercise 1

1.1 Question (a)

We compute $\mathbb{E}[y_i]$:

$$\begin{aligned}\mathbb{E}[y_i] &= \mathbb{E}[a + bx_i + \epsilon_i] && \text{(definition)} \\ &= \mathbb{E}[a] + \mathbb{E}[bx_i] + \mathbb{E}[\epsilon_i] && \text{(linearity of expectation)} \\ &= a + bx_i && (a, b, x_i \text{ deterministic, } \epsilon_i \text{ centered normal})\end{aligned}$$

We compute $\text{Var}(y_i)$:

$$\begin{aligned}\text{Var}[y_i] &= \mathbb{E}[y_i^2] - \mathbb{E}[y_i]^2 \\ &= \mathbb{E}[(a + bx_i + \epsilon_i)^2] - (a + bx_i)^2 \\ &= a^2 + b^2x_i^2 + \mathbb{E}[\epsilon_i^2] && \text{(linearity of expectation)} \\ &\quad + 2abx_i + 2a\mathbb{E}[\epsilon_i] + 2bx_i\mathbb{E}[\epsilon_i] && \text{and only } \epsilon_i \text{ random)} \\ &= (a^2 + b^2x_i^2 + 2abx_i) \\ &= \mathbb{E}[\epsilon_i^2] && (\epsilon_i \text{ centered normal)} \\ &= \mathbb{E}[\epsilon_i^2] - \underbrace{\mathbb{E}[\epsilon_i]^2}_{=0} \\ &= \text{Var}(\epsilon_i) \\ &= \sigma^2\end{aligned}$$

1.2 Question (b)

Our goal is to prove the following:

$$\mathbb{P}\left(\bigcap_{i=1}^N \{y_i = \nu_i\}\right) = \prod_{i=1}^N \mathbb{P}(\{y_i = \nu_i\}) \quad (1)$$

and we have that for a given $1 \leq i \leq N$:

$$\begin{aligned}\mathbb{P}(\{y_i = \nu_i\}) &= \mathbb{P}(\{a + bx_i + \epsilon_i = \nu_i\}) \\ &= \mathbb{P}(\{\epsilon_i = \nu_i - \underbrace{(a + bx_i)}_{\text{constant}}\})\end{aligned}$$

We note that for a given $1 \leq i \leq N$, a, b and x_i are constant, so we can rewrite (1) as:

$$\begin{aligned} \mathbb{P} \left(\bigcap_{i=1}^N \{y_i = \nu_i\} \right) &= \mathbb{P} \left(\bigcap_{i=1}^N \{\epsilon_i = \nu_i - (a + bx_i)\} \right) \\ &= \prod_{i=1}^N \mathbb{P}(\{\epsilon_i = \nu_i - (a + bx_i)\}) \end{aligned} \quad (2)$$

Now, the $\nu_i - (a + bx_i)$ in equation (2) are just deterministic values. Hence, by the fact that the ϵ_i 's, $1 \leq i \leq N$ are Independent and Identically Distributed (i.i.d.), we have that (2) holds.

Thus the y_i 's, $1 \leq i \leq N$ are i.i.d..

1.3 Question (c)

Let $g(y_i)$ denote the Probability Density Function (PDF) of y_i . We define the likelihood function of y_i , $1 \leq i \leq N$ as:

$$\mathcal{L}(\theta) := \prod_{i=1}^N g(y_i; \theta)$$

We also define the log-likelihood as:

$$\ell(\theta) := \log \mathcal{L}(\theta)$$

and since the y_i 's are i.i.d.:

$$\ell(\theta) = \log \prod_{i=1}^N g(y_i; \theta) = \sum_{i=1}^N \log g(y_i; \theta)$$

Moreover, we know that the PDF of a random variable $X \sim \mathcal{N}(\mu, \sigma^2)$ is given by:

$$f_{\mu, \sigma^2}(t) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left(-\frac{(t - \mu)^2}{2\sigma^2} \right)$$

We also know that for a given i , $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$. Therefore its PDF is given by:

$$f(t) := f_{0, \sigma^2}(t) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left(-\frac{t^2}{2\sigma^2} \right)$$

which we can rename to $f(t; \theta)$, to stress the fact that we have an influence on our parameters $\theta = (a, b) \in \mathbb{R}^2$, not over our observations. Thus by (2), we obtain that the log-likelihood becomes:

$$\begin{aligned}\ell(\theta) &= \sum_{i=1}^N \log f(\epsilon_i - (a + bx_i); \theta) \\ &= \sum_{i=1}^N \log \left[\frac{1}{\sqrt{2\pi\sigma^2}} \exp \left(-\frac{(\epsilon_i - (a + bx_i))^2}{2\sigma^2} \right) \right] \\ &= -\frac{N}{2\sigma^2} \log \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] \sum_{i=1}^N (\epsilon_i - (a + bx_i))^2\end{aligned}$$

and we want find \hat{a}_{ML} by maximising ℓ with respect to a . Therefore we get:

$$\begin{aligned}\hat{a}_{ML} &= \arg \max_a \ell(\theta) \\ &= \arg \max_a -\frac{N}{2\sigma^2} \log \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] \sum_{i=1}^N (\epsilon_i - (a + bx_i))^2 \\ &= \arg \min_a \sum_{i=1}^N (\epsilon_i - (a + bx_i))^2 \\ &= \arg \min_a \sum_{i=1}^N (\epsilon_i - bx_i - a)^2\end{aligned}$$

1.4 Question (d)

1.5 Question (e)

2 Exercise 2

2.1 Question (a)

2.2 Question (b)