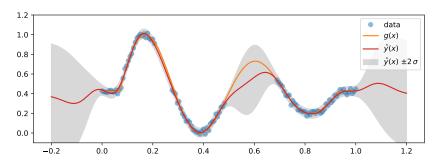
Introduction to Gaussian processes

Steve Schmerler

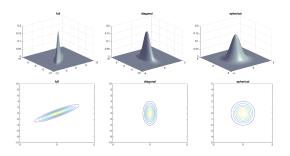
Helmholtz AI @HZDR

Motivation: Why GPs?



- ▶ interpolation or regression for low-dimensional problems ("smoothing device")
- predictive uncertainty
- building block for Bayesian optimization
- Bayesian stats and Gaussian process (GP) theory: understand uncertainty quantification (UQ) methods for neural networks (NNs)
- ▶ infinite width limits of NNs: neural network Gaussian process (NNGP) and the neural tangent kernel (NTK)
- two derivations: weight space, function space

Preliminaries: multivariate normal distribution



$$p(\boldsymbol{x}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\,\pi)^D\,|\boldsymbol{\Sigma}|}}\,\exp\left(-\frac{1}{2}\,(\boldsymbol{x}-\boldsymbol{\mu})^\top\,\boldsymbol{\Sigma}^{-1}\,(\boldsymbol{x}-\boldsymbol{\mu})\right)$$

$$\boldsymbol{\Sigma} = \begin{bmatrix} \sigma_1^2 & \operatorname{cov}[x_1, x_2] \\ \operatorname{cov}[x_1, x_2] & \sigma_2^2 \end{bmatrix}$$

Preliminaries: linear models

Linear model (parametric: $\dim {\pmb w} = D \neq N$, data set content "compressed" into ${\pmb w}$)

$$f(\boldsymbol{x}) = \boldsymbol{w}^{\intercal}\,\boldsymbol{x} = w_1\,x_1 + w_2\,x_2 + \cdots$$

$$f(\boldsymbol{x}) = \boldsymbol{w}^{\top} \left[1, \boldsymbol{x} \right] = w_0 + w_1 \, x_1 + w_2 \, x_2 + \cdots$$

Only regression models of the form

$$f:\mathbb{R}^D\to\mathbb{R}$$

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Only regression models of the form

$$f: \mathbb{R}^D \to \mathbb{R}$$

Data set

$$\begin{split} \mathcal{D} &= \{(\boldsymbol{x}_i, y_i)\}_{i=1}^N = (\mathbf{X}, \boldsymbol{y}) \\ \boldsymbol{x}_i &\in \mathcal{X} = \mathbb{R}^D \\ y_i &\in \mathcal{Y} = \mathbb{R} \\ \mathbf{X} &\in \mathbb{R}^{N \times D} \end{split}$$

Design matrix

$$\mathbf{X} = \overbrace{ egin{bmatrix} - oldsymbol{x}_1^ op & - oldsymbol{x}_2^ op & -$$

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Only regression models of the form

$$f: \mathbb{R}^D \to \mathbb{R}$$

Notation

$$\begin{array}{ll} \text{(noisy) data/target/label} & y \\ \text{model output (train)} & f = \boldsymbol{w}^\top \boldsymbol{x}, \ \boldsymbol{f} = \mathbf{X} \, \boldsymbol{w} \\ \text{model output (test)} & f_* = \boldsymbol{w}^\top \boldsymbol{x}_*, \ \boldsymbol{f}_* = \mathbf{X}_* \, \boldsymbol{w} \end{array}$$

Data set

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$$\mathbf{X} = \overbrace{\left[egin{array}{c} -oldsymbol{x}_1^ op & \ -oldsymbol{x}_2^ op & \ dots \ -oldsymbol{x}_N^ op & \ \end{array}
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Basis functions

Feature space mapping

$$oldsymbol{\phi}: \mathcal{X}
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f(x) is nonlinear in x but still linear in w

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Example: polynomial basis: $m{x} \in \mathbb{R}^2$, $m{\mathcal{F}} = \mathbb{R}^5$, $m{w}, m{\phi}(m{x}) \in \mathbb{R}^5$

$$\begin{split} & \boldsymbol{\phi}(\boldsymbol{x}) = [1, x_1, x_2, x_1^2, x_2^2] \\ & f(\boldsymbol{x}) = \boldsymbol{w}^\top \, \boldsymbol{\phi}(\boldsymbol{x}) = w_0 + w_1 \, x_1 + w_2 \, x_2 + w_3 \, x_1^2 + w_4 \, x_2^2 \end{split}$$

sklearn.preprocessing.PolynomialFeatures

Kernel function $\kappa:\mathcal{X}\times\mathcal{X}\to\mathbb{R}$ as similarity measure

- \blacktriangleright symmetric: $\kappa(\boldsymbol{x}_i,\boldsymbol{x}_j) = \kappa(\boldsymbol{x}_j,\boldsymbol{x}_i)$
- $lackbox{ positive: } \kappa({m x}_i,{m x}_j) \geq 0$

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Gram matrix $\mathbf{K} \in \mathbb{R}^{N imes N}$

$$K_{ij} = \kappa(\boldsymbol{x}_i, \boldsymbol{x}_j)$$

$$\mathbf{K} := \kappa(\mathbf{X}, \mathbf{X})$$

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Gram matrix $\mathbf{K} \in \mathbb{R}^{N \times N}$ "Kernel trick"

$$K_{ij} = \kappa(\boldsymbol{x}_i, \boldsymbol{x}_j) \qquad \qquad \kappa(\boldsymbol{x}_i, \boldsymbol{x}_j) = \left\langle \phi(\boldsymbol{x}_i), \phi(\boldsymbol{x}_j) \right\rangle \equiv \phi(\boldsymbol{x}_i)^\top \phi(\boldsymbol{x}_j)$$

 $\mathbf{K} := \kappa(\mathbf{X}, \mathbf{X})$

Rich theory (Reproducing kernel Hilbert space, Mercer's theorem, ...): no need to define ϕ explicitly, sufficient to define $\kappa(\cdot,\cdot)$, for certain κ we have $f(\boldsymbol{x}) = \sum_{i=1}^{\infty} w_i \, \phi_i(\boldsymbol{x})$

Kernel function $\kappa: \mathcal{X} \times \mathcal{X} \to \mathbb{R}$ as similarity measure

- $\qquad \qquad \mathbf{ symmetric:} \ \kappa(\boldsymbol{x}_i,\boldsymbol{x}_j) = \kappa(\boldsymbol{x}_j,\boldsymbol{x}_i)$
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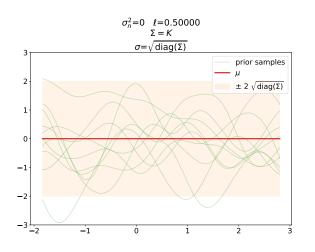
"Kernel trick"

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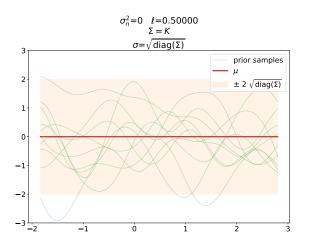
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$$\kappa(\boldsymbol{x}_i, \boldsymbol{x}_j) = \boldsymbol{x}_i^{\top} \boldsymbol{x}_j$$
 Linear / dot product kernel

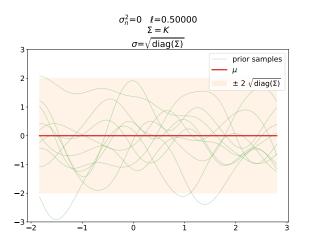
$$\kappa(\boldsymbol{x}_i, \boldsymbol{x}_j) = \exp\left(-\frac{\|\boldsymbol{x}_i - \boldsymbol{x}_j\|_2^2}{2\,\ell^2}\right) = \begin{cases} 1 & \boldsymbol{x}_i = \boldsymbol{x}_j \\ < 1 & \text{else} \end{cases} \quad \begin{array}{l} \text{Gaussian/RBF/"squared exponential"} \\ \text{kernel, characteristic length scale } \ell \end{cases}$$



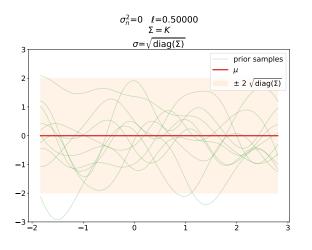
$$\boldsymbol{x} = x \in \mathbb{R}, \mathbf{X} \in \mathbb{R}^{N \times 1}$$



$$m{x} = x \in \mathbb{R}, \mathbf{X} \in \mathbb{R}^{N imes 1}$$
 GP prior for model $f = f(m{x}) = m{w}^{ op} \, m{\phi}(m{x})$ $p(m{w}) = \mathcal{N}(m{\theta}, m{\Sigma}_{m{w}})$ weight prior

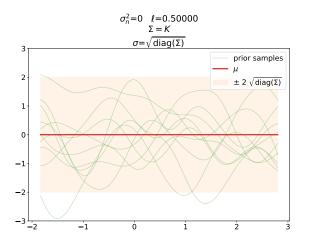


$$\begin{split} \boldsymbol{x} &= x \in \mathbb{R}, \mathbf{X} \in \mathbb{R}^{N \times 1} \\ \text{GP prior for model } f &= f(\boldsymbol{x}) = \boldsymbol{w}^\top \, \boldsymbol{\phi}(\boldsymbol{x}) \\ p(\boldsymbol{w}) &= \mathcal{N}(\boldsymbol{0}, \boldsymbol{\Sigma}_{\boldsymbol{w}}) \quad \text{weight prior} \\ p(\boldsymbol{f}|\mathbf{X}) &= \mathcal{N}(\boldsymbol{0}, \mathbf{K}) \end{split}$$

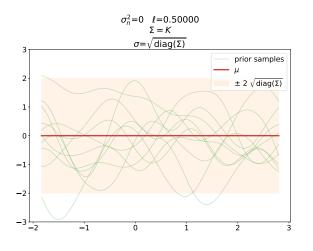


$1D\ {\rm example}$ where

$$m{x} = x \in \mathbb{R}, \mathbf{X} \in \mathbb{R}^{N imes 1}$$
 GP prior for model $m{f} = f(m{x}) = m{w}^{ op} \, m{\phi}(m{x})$ $p(m{w}) = \mathcal{N}(m{0}, m{\Sigma}_{m{w}})$ weight prior $p(m{f} | \mathbf{X}) = \mathcal{N}(m{0}, \mathbf{K})$ $\mathbb{E}[m{f}] = \mathbb{E}[m{\Phi} \, m{w}] = m{\Phi} \, \mathbb{E}[m{w}] = m{0}$



$$\begin{split} \boldsymbol{x} &= \boldsymbol{x} \in \mathbb{R}, \mathbf{X} \in \mathbb{R}^{N \times 1} \\ \text{GP prior for model } \boldsymbol{f} &= \boldsymbol{f}(\boldsymbol{x}) = \boldsymbol{w}^\top \, \boldsymbol{\phi}(\boldsymbol{x}) \\ \boldsymbol{p}(\boldsymbol{w}) &= \mathcal{N}(\boldsymbol{0}, \boldsymbol{\Sigma}_{\boldsymbol{w}}) \quad \text{weight prior} \\ \boldsymbol{p}(\boldsymbol{f} | \mathbf{X}) &= \mathcal{N}(\boldsymbol{0}, \mathbf{K}) \\ \mathbb{E}[\boldsymbol{f}] &= \mathbb{E}[\boldsymbol{\Phi} \, \boldsymbol{w}] = \boldsymbol{\Phi} \, \mathbb{E}[\boldsymbol{w}] = \boldsymbol{0} \\ \operatorname{cov}[\boldsymbol{f}] &= \mathbb{E}[(\boldsymbol{f} - \mathbb{E}[\boldsymbol{f}]) \, (\boldsymbol{f} - \mathbb{E}[\boldsymbol{f}])^\top] \\ &= \boldsymbol{\Phi} \, \boldsymbol{\Sigma}_{\boldsymbol{w}} \, \boldsymbol{\Phi}^\top =: \mathbf{K} \end{split}$$



1D example where

$$\boldsymbol{x} = x \in \mathbb{R}, \mathbf{X} \in \mathbb{R}^{N \times 1}$$

GP prior for model
$$f = f(\boldsymbol{x}) = \boldsymbol{w}^{\top} \boldsymbol{\phi}(\boldsymbol{x})$$

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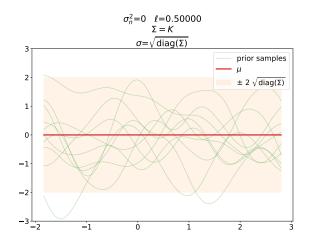
$$\begin{split} \mathbb{E}[\boldsymbol{f}] &= \mathbb{E}[\boldsymbol{\Phi} \, \boldsymbol{w}] = \boldsymbol{\Phi} \, \mathbb{E}[\boldsymbol{w}] = \boldsymbol{0} \\ & \cos[\boldsymbol{f}] &= \mathbb{E}[(\boldsymbol{f} - \mathbb{E}[\boldsymbol{f}]) \, (\boldsymbol{f} - \mathbb{E}[\boldsymbol{f}])^{\top}] \end{split}$$

$$= \mathbf{\Phi} \, \mathbf{\Sigma}_{w} \, \mathbf{\Phi}^{ op} =: \mathbf{K}$$

Covariance (kernel) function $\kappa(\cdot,\cdot)$

$$egin{aligned} & \pmb{K_{ij}} = \pmb{\phi}(\pmb{x}_i)^{\top} \, \pmb{\Sigma_w} \, \pmb{\phi}(\pmb{x}_j) =: \kappa(\pmb{x}_i, \pmb{x}_j) \ & \text{e.g.} \, \pmb{\Sigma_w} = \tau^2 \, \mathbf{I}_D
ightarrow \text{scaling factor in } \kappa \end{aligned}$$

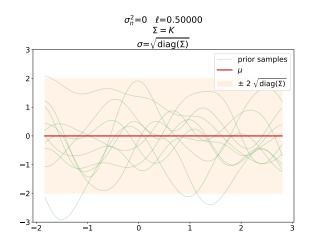
Function space view: the GP prior



The GP as a distribution over functions f

$$f \sim \mathcal{GP}(m(\cdot), \kappa(\cdot, \cdot))$$

Function space view: the GP prior



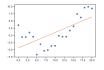
The GP as a distribution over functions f

$$f \sim \mathcal{GP}(m(\cdot), \kappa(\cdot, \cdot))$$

$$\begin{split} p(f_i|\boldsymbol{x}_i) &= \mathcal{N}(\boldsymbol{m}(\boldsymbol{x}_i), \kappa(\boldsymbol{x}_i, \boldsymbol{x}_i)) \\ p(\boldsymbol{f}|\mathbf{X}) &= \mathcal{N}(\boldsymbol{m}(\mathbf{X}), \mathbf{K}) \\ \mathbb{E}[f_i] &= m(\boldsymbol{x}_i) \\ \mathbb{E}[\boldsymbol{f}] &= \boldsymbol{m}(\mathbf{X}) \end{split}$$

$$\begin{split} & \cos[\textit{f}_i,\textit{f}_j] = \mathbb{E}[\left(\textit{f}_i - m(\boldsymbol{x}_i)\right)\left(\textit{f}_j - m(\boldsymbol{x}_j)\right)] \\ & =: \kappa(\boldsymbol{x}_i,\boldsymbol{x}_j) \\ & \cos[\textit{\textbf{f}}] = \mathbf{K} \end{split}$$

Likelihood







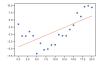
Model noise σ_n^2 in data y. $p(y|\boldsymbol{x}, \boldsymbol{w})$ interpretation:

- lackbox distribution $p(y|\ldots)$ over y
- ightharpoonup function of w

"The likelihood function reflects the data we expect to see for each setting of the parameters \boldsymbol{w} ."

$$p(y|\boldsymbol{x}, \boldsymbol{w}) = \mathcal{N}(\boldsymbol{w}^{\top} \boldsymbol{\phi}(\boldsymbol{x}), \sigma_n^2)$$

Likelihood







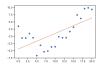
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$$egin{aligned} p(y|x, oldsymbol{w}) &= \mathcal{N}(oldsymbol{w}^{ op} \, \phi(x), \sigma_n^2) \ f &= oldsymbol{w}^{ op} \, \phi(x) \ y &= oldsymbol{w}^{ op} \, \phi(x) + \epsilon \ \epsilon &\sim \mathcal{N}(0, \sigma_n^2) \ oldsymbol{ heta} &= (oldsymbol{w}, oldsymbol{\xi}) \ oldsymbol{\xi} &= (\sigma_n^2) \quad ext{hyper parameters} \end{aligned}$$

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$$\begin{split} & p(y|\boldsymbol{x}, \boldsymbol{w}) = \mathcal{N}(\boldsymbol{w}^{\top} \, \phi(\boldsymbol{x}), \sigma_n^2) \\ & f = \boldsymbol{w}^{\top} \, \phi(\boldsymbol{x}) \\ & y = \boldsymbol{w}^{\top} \, \phi(\boldsymbol{x}) + \epsilon \\ & \epsilon \sim \mathcal{N}(0, \sigma_n^2) \\ & \boldsymbol{\theta} = (\boldsymbol{w}, \boldsymbol{\xi}) \\ & \boldsymbol{\xi} = (\sigma_n^2) \quad \text{hyper parameters} \end{split}$$

$$\begin{split} p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) &= \mathcal{N}(\boldsymbol{\Phi}\,\boldsymbol{w}, \sigma_n^2\,\mathbf{I}_N) \\ &= \prod_{i=1}^N p(y_i|\boldsymbol{x}_i, \boldsymbol{w}) \\ &= \prod_{i=1}^N \frac{1}{\sqrt{2\,\pi}\,\sigma_n}\,\exp\left(-\frac{(y_i - f_i)^2}{2\,\sigma_n^2}\right) \\ &= \frac{1}{\sqrt{(2\,\pi)^N}\,\sigma_n}\,\exp\left(-\frac{\boldsymbol{\epsilon}^\top\,\boldsymbol{\epsilon}}{2\,\sigma_n^2}\right) \end{split}$$

Bayes' rule

Bayesian inference: infer posterior distribution over weights (i.e. models) $p(w|\mathbf{X},y)$ by using training data (\mathbf{X},y)

$$\underbrace{ p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y})}_{\text{weight posterior}} = \underbrace{ \underbrace{ p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})}_{\text{p}(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})} \underbrace{ p(\boldsymbol{w})}_{\text{p}(\boldsymbol{w})} }_{\text{marginal likelihood or evidence}} = \underbrace{ p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w})}_{\text{p}(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w}}$$

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$$\frac{ p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y}) }{ p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y}) } = \frac{ \overbrace{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \quad p(\boldsymbol{w}) }^{\text{likelihood} \quad \text{weight prior}} { p(\boldsymbol{y}|\mathbf{X}) } = \frac{ p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) }{ \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w} }$$
 marginal likelihood or evidence

More compact notation

$$p(\boldsymbol{w}|\mathcal{D}) = \frac{p(\mathcal{D}|\boldsymbol{w}) p(\boldsymbol{w})}{p(\mathcal{D})} = \frac{p(\mathcal{D}|\boldsymbol{w}) p(\boldsymbol{w})}{\int p(\mathcal{D}|\boldsymbol{w}) p(\boldsymbol{w}) d\boldsymbol{w}}$$

In simple cases, inference can be performed analytically, e.g. for a Gaussian likelihood.

Bayesian model averaging (BMA)

$$\langle w \rangle = \int w p(w) dw$$

 $\langle f(w) \rangle = \int f(w) p(w) dw$

Bayesian model averaging (BMA)

$$p(\mathbf{f_*}|\mathbf{X_*},\mathbf{X},\mathbf{y}) = \int \underbrace{p(\mathbf{f_*}|\mathbf{X_*},\mathbf{w})}_{\text{likelihood}} \underbrace{p(\mathbf{w}|\mathbf{X},\mathbf{y})}_{\text{weight posterior}} \, \mathrm{d}\mathbf{w} = \mathcal{N}(\boldsymbol{\mu_*},\boldsymbol{\Sigma_*}) \qquad \frac{\langle w \rangle = \int w \, p(w) \, \mathrm{d}w}{\langle f(w) \rangle = \int f(w) \, p(w) \, \mathrm{d}w}$$

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Predictive mean μ_* and cov. Σ_*

$$\mu_* = \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \mathbf{I}_N \right)^{-1} \mathbf{y}$$
$$= \mathbf{K}_* \alpha$$

$$\begin{split} \mathbf{K}_* &= \kappa(\mathbf{X}_*, \mathbf{X}) \\ \mathbf{K}_{**} &= \kappa(\mathbf{X}_*, \mathbf{X}_*) \end{split}$$

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Predictive mean μ_* and cov. Σ_*

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Bayesian model averaging (BMA)

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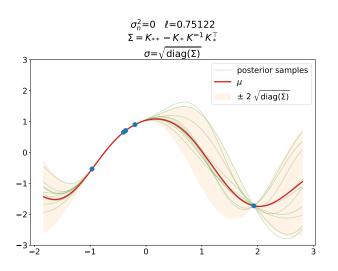
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$$\begin{aligned} \boldsymbol{\mu}_* &= \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \mathbf{I}_N \right)^{-1} \, \boldsymbol{y} \\ &= \mathbf{K}_* \, \boldsymbol{\alpha} \\ \boldsymbol{\mu}_* &= \sum_{j=1}^N \alpha_j \, \kappa(\boldsymbol{x}_*, \boldsymbol{x}_j) \\ \boldsymbol{\Sigma}_* &= \mathbf{K}_{**} - \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \mathbf{I}_N \right)^{-1} \, \mathbf{K}_*^\top \\ \mathbf{K}_* &= \kappa(\mathbf{X}_*, \mathbf{X}) \\ \mathbf{K}_{**} &= \kappa(\mathbf{X}_*, \mathbf{X}_*) \end{aligned}$$

Non-parametric model: $\mu = \mathbf{K} \, lpha$

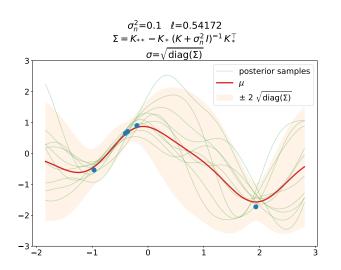
- $\mathbf{K} \in \mathbb{R}^{N imes N}$ contains info about whole training inputs $\mathbf{X} \in \mathbb{R}^{N imes D}$
- $lackbox{lack}$ weights $oldsymbol{lpha} \in \mathbb{R}^N$ contain info about $(\mathbf{X},oldsymbol{y})$
- lackbox large data sets (large N) make vanilla GPs costly

Posterior predictive with $\sigma_n^2 = 0$



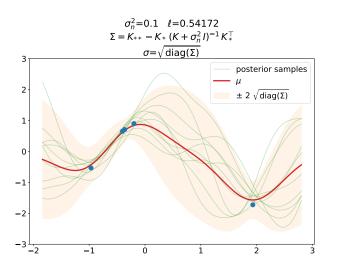
$$\begin{split} p(\mathbf{f}_*|\mathbf{X}_*, \mathbf{X}, \mathbf{y}) &= \mathcal{N}(\boldsymbol{\mu}_*, \boldsymbol{\Sigma}_*) \\ \boldsymbol{\mu}_* &= \mathbf{K}_* \, \mathbf{K}^{-1} \, \mathbf{y} \\ &= \mathbf{K}_* \, \boldsymbol{\alpha} \\ \boldsymbol{\mu}_* &= \sum_j \alpha_j \, \kappa(\boldsymbol{x}_*, \boldsymbol{x}_j) \\ \boldsymbol{\Sigma}_* &= \mathbf{K}_{**} - \mathbf{K}_* \mathbf{K}^{-1} \, \mathbf{K}_*^\top \end{split}$$

Posterior predictive with $\sigma_n^2 > 0$



$$\begin{split} p(\mathbf{f}_*|\mathbf{X}_*,\mathbf{X},\mathbf{y}) &= \mathcal{N}(\boldsymbol{\mu}_*,\boldsymbol{\Sigma}_*) \\ \boldsymbol{\mu}_* &= \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N\right)^{-1} \, \boldsymbol{y} \\ &= \mathbf{K}_* \, \boldsymbol{\alpha} \\ \boldsymbol{\mu}_* &= \sum_j \alpha_j \, \kappa(\boldsymbol{x}_*,\boldsymbol{x}_j) \\ \boldsymbol{\Sigma}_* &= \mathbf{K}_{**} - \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N\right)^{-1} \, \mathbf{K}_*^\top \end{split}$$

Posterior predictive with $\sigma_n^2 > 0$



$$\begin{split} & \boldsymbol{\mu_*} = \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N \right)^{-1} \, \boldsymbol{y} \\ & = \mathbf{K}_* \, \boldsymbol{\alpha} \\ & \boldsymbol{\mu_*} = \sum_j \alpha_j \, \kappa(\boldsymbol{x}_*, \boldsymbol{x}_j) \\ & \boldsymbol{\Sigma_*} = \mathbf{K}_{**} - \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N \right)^{-1} \, \mathbf{K}_*^\top \end{split}$$

Data noise σ_n^2 : transform interpolation \to regression, same effect as a regularization term in NN training

 $p(\mathbf{f}_*|\mathbf{X}_*,\mathbf{X},\mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}_*,\boldsymbol{\Sigma}_*)$

Function space view of GPs: the joint

We rewrite the prior: Divide data into "train" ${\it f}$ and "test/prediction" ${\it f}_*$

$$\begin{split} M &= N + N_* \\ \mathbf{X} \in \mathbb{R}^{M \times D} &\to \left(\mathbf{X} \in \mathbb{R}^{N \times D}, \mathbf{X}_* \in \mathbb{R}^{N_* \times D} \right) \\ \mathbf{f} \in \mathbb{R}^M &\to \left(\mathbf{f} \in \mathbb{R}^N, \mathbf{f}_* \in \mathbb{R}^{N_*} \right) \end{split}$$

Function space view of GPs: the joint

We rewrite the prior: Divide data into "train" f and "test/prediction" f_*

$$\begin{split} M &= N + N_* \\ \mathbf{X} &\in \mathbb{R}^{M \times D} \rightarrow \left(\mathbf{X} \in \mathbb{R}^{N \times D}, \mathbf{X}_* \in \mathbb{R}^{N_* \times D} \right) \\ \boldsymbol{f} &\in \mathbb{R}^M \rightarrow \left(\boldsymbol{f} \in \mathbb{R}^N, \boldsymbol{f}_* \in \mathbb{R}^{N_*} \right) \end{split}$$

write the prior as joint over concat. $(\mathbf{f}, \mathbf{f}_*)$

$$\begin{bmatrix} \boldsymbol{f} \\ \boldsymbol{f}_* \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \boldsymbol{m}(\mathbf{X}) \\ \boldsymbol{m}(\mathbf{X}_*) \end{bmatrix}, \begin{bmatrix} \operatorname{cov}[\boldsymbol{f}] & \mathbf{K}_*^\top \\ \mathbf{K}_* & \operatorname{cov}[\boldsymbol{f}_*] \end{bmatrix} \right)$$

$$\sim \mathcal{N} \left(\begin{bmatrix} \boldsymbol{m}(\mathbf{X}) \\ \boldsymbol{m}(\mathbf{X}_*) \end{bmatrix}, \begin{bmatrix} \mathbf{K} & \mathbf{K}_*^\top \\ \mathbf{K}_* & \mathbf{K}_{**} \end{bmatrix} \right)$$

$$\sim p(\boldsymbol{f}, \boldsymbol{f}_* | \mathbf{X}, \mathbf{X}_*)$$

Function space view of GPs: the joint

We rewrite the prior: Divide data into "train" f and "test/prediction" f_*

$$\begin{split} M &= N + N_* \\ \mathbf{X} \in \mathbb{R}^{M \times D} &\to \left(\mathbf{X} \in \mathbb{R}^{N \times D}, \mathbf{X}_* \in \mathbb{R}^{N_* \times D} \right) \\ \boldsymbol{f} \in \mathbb{R}^M &\to \left(\boldsymbol{f} \in \mathbb{R}^N, \boldsymbol{f}_* \in \mathbb{R}^{N_*} \right) \end{split}$$

write the prior as joint over concat. $(\mathbf{f}, \mathbf{f}_*)$

$$\begin{bmatrix} \boldsymbol{f} \\ \boldsymbol{f}_* \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \boldsymbol{m}(\mathbf{X}) \\ \boldsymbol{m}(\mathbf{X}_*) \end{bmatrix}, \begin{bmatrix} \operatorname{cov}[\boldsymbol{f}] & \mathbf{K}_*^\top \\ \mathbf{K}_* & \operatorname{cov}[\boldsymbol{f}_*] \end{bmatrix} \right)$$

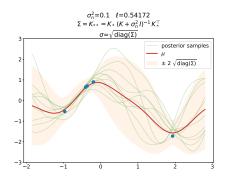
$$\sim \mathcal{N} \left(\begin{bmatrix} \boldsymbol{m}(\mathbf{X}) \\ \boldsymbol{m}(\mathbf{X}_*) \end{bmatrix}, \begin{bmatrix} \mathbf{K} & \mathbf{K}_*^\top \\ \mathbf{K}_* & \mathbf{K}_{**} \end{bmatrix} \right)$$

$$\sim p(\boldsymbol{f}, \boldsymbol{f}_* | \mathbf{X}, \mathbf{X}_*)$$

For noisy $y = f + \epsilon$, we have

$$\begin{split} \begin{bmatrix} \boldsymbol{y} \\ \boldsymbol{f}_* \end{bmatrix} &\sim \mathcal{N} \left(\begin{bmatrix} \boldsymbol{m}(\mathbf{X}) \\ \boldsymbol{m}(\mathbf{X}_*) \end{bmatrix}, \begin{bmatrix} \operatorname{cov}[\boldsymbol{y}] & \mathbf{K}_*^\top \\ \mathbf{K}_* & \operatorname{cov}[\boldsymbol{f}_*] \end{bmatrix} \right) \\ &\sim \mathcal{N} \left(\begin{bmatrix} \boldsymbol{m}(\mathbf{X}) \\ \boldsymbol{m}(\mathbf{X}_*) \end{bmatrix}, \begin{bmatrix} \mathbf{K} + \sigma_n^2 \, \mathbf{I}_N & \mathbf{K}_*^\top \\ \mathbf{K}_* & \mathbf{K}_{**} \end{bmatrix} \right) \\ &\sim p(\boldsymbol{y}, \boldsymbol{f}_* | \mathbf{X}, \mathbf{X}_*) \end{split}$$

Function space view of GPs: posterior predictive



Transform the joint $p(\boldsymbol{y}, \boldsymbol{f}_*|\mathbf{X}, \mathbf{X}_*)$ into the posterior predictive $p(\boldsymbol{f}_*|\mathbf{X}_*, \mathbf{X}, \boldsymbol{y})$ by conditioning on $(\mathbf{X}, \boldsymbol{y})$ ("training data").

$$\begin{split} p(\mathbf{f}_*|\mathbf{X}_*,\mathbf{X},\mathbf{y}) &= \mathcal{N}(\boldsymbol{\mu}_*,\boldsymbol{\Sigma}_*) \\ \boldsymbol{\mu}_* &= \mathbf{m}(\mathbf{X}_*) + \mathbf{K}_* \, \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N\right)^{-1} \, \left(\mathbf{y} - \mathbf{m}(\mathbf{X})\right) \\ \boldsymbol{\Sigma}_* &= \mathrm{cov}[\mathbf{f}_*] = \mathbf{K}_{**} - \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N\right)^{-1} \, \mathbf{K}_*^\top \end{split}$$

Same result as the posterior predictive obtained from Bayes' rule + model averaging. Here we also have a mean function $m(\cdot) \neq 0$.

$$p(y|\mathbf{X}) = \int p(y|\mathbf{X}, w) p(w) dw$$

Bayes' rule

$$\frac{ \underset{p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{w})}{\text{weight posterior}} = \underbrace{ \frac{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})}{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})} \underbrace{p(\boldsymbol{w})}_{\text{marginal likelihood or evidence}}^{\text{weight prior}} }_{\text{marginal likelihood or evidence}} = \frac{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w})}{\int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w}}$$

$$p(y|\mathbf{X}) = \int p(y|\mathbf{X}, w) p(w) dw$$

Marginal likelihood as function of hyperparameters $\boldsymbol{\xi} = (\ell, \sigma_n^2)$.

Bayes' rule

$$\frac{\underset{p(\boldsymbol{w}|\mathbf{X},\boldsymbol{w})}{\text{weight posterior}}}{p(\boldsymbol{w}|\mathbf{X},\boldsymbol{y})} = \frac{\frac{\underset{p(\boldsymbol{y}|\mathbf{X},\boldsymbol{w})}{p(\boldsymbol{y}|\mathbf{X},\boldsymbol{w})} \underbrace{p(\boldsymbol{w})}}{p(\boldsymbol{y}|\mathbf{X})}}{\underbrace{\frac{p(\boldsymbol{y}|\mathbf{X})}{p(\boldsymbol{w})}}}$$

$$= \frac{p(\boldsymbol{y}|\mathbf{X},\boldsymbol{w}) p(\boldsymbol{w})}{\int p(\boldsymbol{y}|\mathbf{X},\boldsymbol{w}) p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w}}$$

$$p(\boldsymbol{y}|\mathbf{X}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) p(\boldsymbol{w}) d\boldsymbol{w}$$

Marginal likelihood as function of hyperparameters $\boldsymbol{\xi} = (\boldsymbol{\ell}, \sigma_n^2)$. Because of $\boldsymbol{f} = \boldsymbol{\Phi} \, \boldsymbol{w}, \ \boldsymbol{\int} \cdots \mathrm{d} \boldsymbol{w} \to \boldsymbol{\int} \cdots \mathrm{d} \boldsymbol{f}$

$$p(\boldsymbol{y}|\mathbf{X}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{f}) \, p(\boldsymbol{f}|\mathbf{X}) \, \mathrm{d}\boldsymbol{f}$$
$$p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{f}, \boldsymbol{\xi}) \, p(\boldsymbol{f}|\mathbf{X}, \boldsymbol{\xi}) \, \mathrm{d}\boldsymbol{f}$$

Bayes' rule

$$\frac{\mathbf{w} \text{ eight posterior}}{p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y})} = \frac{\overbrace{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})}^{\text{ likelihood}} \underbrace{p(\boldsymbol{y}|\mathbf{X})}^{\text{ weight prior}} \underbrace{p(\boldsymbol{y}|\mathbf{X})}_{\text{marginal likelihood or evidence}} \\ = \frac{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w})}{\int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w}}$$

$$p(\boldsymbol{y}|\mathbf{X}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) p(\boldsymbol{w}) d\boldsymbol{w}$$

Marginal likelihood as function of hyperparameters $\boldsymbol{\xi} = (\boldsymbol{\ell}, \sigma_n^2)$. Because of $\boldsymbol{f} = \boldsymbol{\Phi} \, \boldsymbol{w}, \ \boldsymbol{\int} \cdots \mathrm{d} \boldsymbol{w} \to \boldsymbol{\int} \cdots \mathrm{d} \boldsymbol{f}$

$$p(\boldsymbol{y}|\mathbf{X}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{f}) \, p(\boldsymbol{f}|\mathbf{X}) \, \mathrm{d}\boldsymbol{f}$$
$$p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{f}, \boldsymbol{\xi}) \, p(\boldsymbol{f}|\mathbf{X}, \boldsymbol{\xi}) \, \mathrm{d}\boldsymbol{f}$$

Bayes' rule

$$\frac{p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y})}{p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y})} = \frac{\overbrace{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})}^{\text{likelihood}} \underbrace{p(\boldsymbol{w})}^{\text{weight prior}}}{\underbrace{p(\boldsymbol{y}|\mathbf{X})}^{\text{polity}}}$$
marginal likelihood or evidence
$$= \frac{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w})}{\int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w}}$$

(negative) log marginal likelihood

$$\ln p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi}) = -\frac{1}{2} \left[\underbrace{\boldsymbol{y}^\top \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N\right)^{-1} \, \boldsymbol{y}}_{\text{model fit}} + \underbrace{\ln |\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N|}_{\text{model complexity}} + N \, \ln(2 \, \pi) \right]$$

$$p(\boldsymbol{y}|\mathbf{X}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \, p(\boldsymbol{w}) \, \mathrm{d}\boldsymbol{w}$$

Marginal likelihood as function of hyperparameters $\boldsymbol{\xi} = (\boldsymbol{\ell}, \sigma_n^2)$. Because of $\boldsymbol{f} = \boldsymbol{\Phi} \, \boldsymbol{w}, \ \boldsymbol{\int} \cdots \mathrm{d} \boldsymbol{w} \to \boldsymbol{\int} \cdots \mathrm{d} \boldsymbol{f}$

$$p(\boldsymbol{y}|\mathbf{X}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{f}) \, p(\boldsymbol{f}|\mathbf{X}) \, \mathrm{d}\boldsymbol{f}$$
$$p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi}) = \int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{f}, \boldsymbol{\xi}) \, p(\boldsymbol{f}|\mathbf{X}, \boldsymbol{\xi}) \, \mathrm{d}\boldsymbol{f}$$

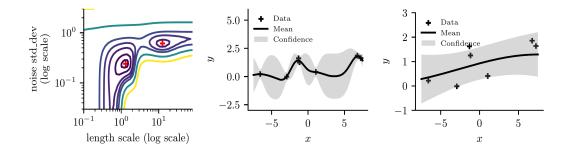
Bayes' rule

$$\begin{split} \frac{\text{weight posterior}}{p(\boldsymbol{w}|\mathbf{X}, \boldsymbol{y})} &= \frac{\overbrace{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w})}^{\text{likelihood}} \underbrace{p(\boldsymbol{w})}^{\text{weight prior}} \\ &= \frac{p(\boldsymbol{y}|\mathbf{X})}{p(\boldsymbol{w})} \\ &= \frac{p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \ p(\boldsymbol{w})}{\int p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{w}) \ p(\boldsymbol{w}) \ \mathrm{d}\boldsymbol{w}} \end{split}$$

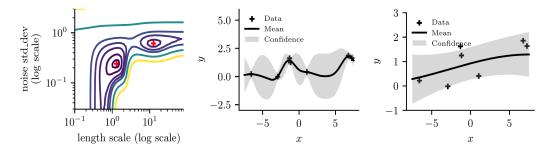
(negative) log marginal likelihood

$$\begin{split} & \ln p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi}) = -\frac{1}{2} \left[\underbrace{\boldsymbol{y}^\top \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N \right)^{-1} \, \boldsymbol{y}}_{\text{model fit}} + \underbrace{\ln \left| \mathbf{K} + \sigma_n^2 \, \mathbf{I}_N \right|}_{\text{model complexity}} + N \, \ln(2 \, \pi) \right] \\ & \boldsymbol{\xi}^* = \arg \max_{\boldsymbol{\xi}} \ln p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi}) = \arg \min_{\boldsymbol{\xi}} (-\ln p(\boldsymbol{y}|\mathbf{X}, \boldsymbol{\xi})) \end{split}$$

Multiple minima of negative log marginal likelihood



Multiple minima of negative log marginal likelihood



Multiple minima: explain data in different ways

- lacktriangle small length scale ℓ , flexible model, low variance $\sigma_n^2 o$ good model fit but complex model
- large length scale ℓ , "stiff"/low curvature model, high variance $\sigma_n^2 \to$ worse model fit but low model complexity

Relation to uncertainty quantification

Different kinds of uncertainty:

- $lackbox{ }$ aleatoric / data uncertainty: σ_n^2

Relation to uncertainty quantification

Different kinds of uncertainty:

- $\begin{array}{c} \bullet \ \ \text{epistemic} \ / \ \text{model uncertainty: weight posterior} \ p(\boldsymbol{w}|\mathbf{X},\boldsymbol{y}) \ \text{and} \\ & \cos[\boldsymbol{f}_*] = \boldsymbol{\Sigma}_* = \mathbf{K}_{**} \mathbf{K}_* \left(\mathbf{K} + \sigma_n^2 \, \mathbf{I}_N\right)^{-1} \, \mathbf{K}_*^\top \\ \end{array}$
- ightharpoonup aleatoric / data uncertainty: σ_n^2

Distinction between "noise-free/noiseless prediction" of $\emph{\textbf{f}}_*$ when using Σ_*

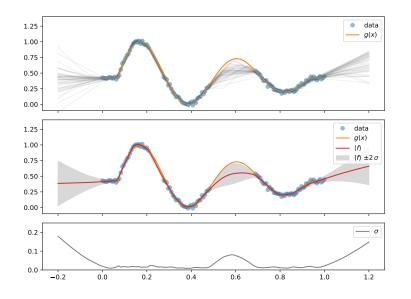
$$p(\mathbf{f}_*|\mathbf{X}_*,\mathbf{X},\mathbf{y}) = \mathcal{N}(\boldsymbol{\mu}_*,\boldsymbol{\Sigma}_*)$$

and "noisy predictions" of $oldsymbol{y}_*$ where we use $oldsymbol{\Sigma}_* + \sigma_n^2 \, \mathbf{I}_N$

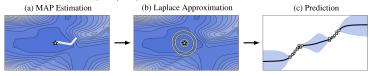
$$p(\boldsymbol{y}_*|\boldsymbol{\mathrm{X}}_*,\boldsymbol{\mathrm{X}},\boldsymbol{y}) = \mathcal{N}(\boldsymbol{\mu}_*,\boldsymbol{\Sigma}_* + \sigma_n^2\,\boldsymbol{\mathrm{I}}_N)$$

In both cases we have the same μ_*

Approximate $p(\boldsymbol{w}|\mathcal{D})$: NN ensembles for UQ



Approximate $p(\boldsymbol{w}|\mathcal{D})$: Laplace approximation for UQ



Post-processing step after NN training (= MAP estimate): $w^* = \arg\min_{w} - \ln p(w|\mathcal{D})$

$$-\ln p(\boldsymbol{w}|\mathcal{D}) = -\ln \left(\frac{p(\mathcal{D}|\boldsymbol{w})\,p(\boldsymbol{w})}{p(\mathcal{D})}\right) = \underbrace{\frac{\text{NN loss }L(\boldsymbol{w}) = \text{NLL+regularizer}}{-\ln p(\mathcal{D}|\boldsymbol{w}) - \ln p(\boldsymbol{w})} + \ln p(\mathcal{D})$$

With gradient $g = \nabla L|_{{\bm w}^*}$, Hessian ${\bf H} = \partial^2 L|_{{\bm w}^*}$ and ${\bm h} = {\bm w} - {\bm w}^*$, approximate loss to 2nd order

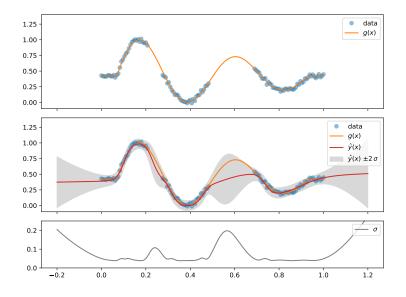
$$L(\boldsymbol{w}) pprox L(\boldsymbol{w}^*) + \underbrace{\boldsymbol{g}^{\top}_{-\boldsymbol{\theta}}}_{-\boldsymbol{\theta}} + \frac{1}{2} \boldsymbol{h}^{\top} \mathbf{H} \boldsymbol{h}$$

Approximate posterior probability distribution over $oldsymbol{w}$ (i.e. over models)

$$p(\boldsymbol{w}|\mathcal{D}) \approx \mathcal{N}(\boldsymbol{w}^*, \boldsymbol{\Sigma})$$
 where $\boldsymbol{\Sigma} = \mathbf{H}^{-1}$

E. Daxberger et al. Laplace Redux – Effortless Bayesian Deep Learning. Version 3. 2022. URL: http://arxiv.org/abs/2106.14806 (visited on 02/05/2023).

Approximate $p(\boldsymbol{w}|\mathcal{D})$: Laplace approximation for UQ

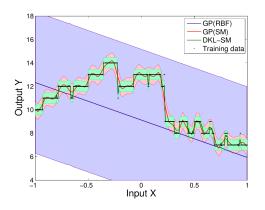


Kernel learning with NNs

(deep) kernel learning: more flexible kernels via NNs: use base kernel + NN features:

$$\kappa(\boldsymbol{x}_i, \boldsymbol{x}_j) = \exp\left(-\frac{\|\boldsymbol{h}_{\boldsymbol{\gamma}}(\boldsymbol{x}_i) - \boldsymbol{h}_{\boldsymbol{\gamma}}(\boldsymbol{x}_j)\|_2^2}{2\,\ell^2}\right)$$

with $h_{\gamma}(x_i)$ an NN embedding ("feature extractor") and γ the NN parameters (weights, biases), optimize $\boldsymbol{\xi}=(\gamma,\ell,\sigma_n^2)$ jointly using log marginal likelihood



A. G. Wilson et al. "Deep Kernel Learning". In: *Proc. 19th Int. Conf. Artif. Intell. Stat.* Artificial Intelligence and Statistics. PMLR, 2016, pp. 370–378.

Software

- https://scikit-learn.org, uses numpy
 - ▶ sklearn.gaussian_process.GaussianProcessRegressor
 - sklearn.kernel_ridge.KernelRidge
- https://gpytorch.ai: *PyTorch*-based, lots of advanced features, approximate methods for scaling GPs, API flexible but complex, GPU support via *PyTorch*
 - ▶ Define a mean function, since $f \sim \mathcal{G}P(m(\cdot), \kappa(\cdot, \cdot))$, so far we had $m(\cdot) = 0$
 - GPs are non-parametric models, $\mathbf{K} \in \mathbb{R}^{N \times N}$, $\dim \alpha = N$, Cholesky decomposition for $\alpha = \left(\mathbf{K} + \sigma_n^2 \mathbf{I}_N\right)^{-1} y$ is $\mathcal{O}(N^3)$, lots of approximate methods, such as KISS-GP (a.k.a. SKI = structured kernel interpolation) for improved scaling
 - lacktriangle variational GPs for approximate inference of $p(w|\mathcal{D})$, e.g. for non-Gaussian likelihoods
 - ▶ GP theory is for $f: \mathbb{R}^D \to \mathbb{R}$, *GPyTorch* supports multi-output GPs for $f: \mathbb{R}^D \to \mathbb{R}^M$
- ▶ https://github.com/dfm/tinygp: basic (educational) code, GPU support via JAX
- https://github.com/JaxGaussianProcesses, similar to tinygp but more features, GPU support via JAX
- https://github.com/SheffieldML/GPy, uses numpy + Cython
- Laplace approximation: https://github.com/AlexImmer/Laplace, PyTorch

Resources

- ► The Book: C. E. Rasmussen and C. K. I. Williams. Gaussian Processes for Machine Learning. MIT Press, 2006 (http://gaussianprocess.org/gpml)
- K. P. Murphy. Probabilistic Machine Learning: An introduction. MIT Press, 2022, K. P. Murphy. Probabilistic Machine Learning: Advanced Topics (draft version). MIT Press, 2022 (https://probml.github.io/pml-book)
- M. Kanagawa et al. Gaussian Processes and Kernel Methods: A Review on Connections and Equivalences. 2018. URL: http://arxiv.org/abs/1807.02582
- shameless plug: https://elcorto.github.io/gp_playground
- ▶ UQ in classification problems: P. Steinbach et al. "Machine Learning State-of-the-Art with Uncertainties". In: ICLR (2022)
- ▶ J. Gawlikowski et al. A Survey of Uncertainty in Deep Neural Networks. 2022. URL: http://arxiv.org/abs/2107.03342 (visited on 07/12/2022)