

1 **wPresent-biased Meditators? The Effects of**  
2 **Mindfulness on Intertemporal Choice**

3 [Version: Figures included]

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25 **Present-biased Meditators? The Effects of**  
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28

29 **Abstract**

30 In this paper, we study the effects of mindfulness meditation on intertemporal decision-making.  
31 One essential aspect of mindfulness practice is to help people focus on the present moment.  
32 While this has been shown to have several positive consequences, it introduces the possibility  
33 that mindfulness meditators become more present-biased, which may have negative  
34 implications for decision-making. In three laboratory experiments and one longitudinal field  
35 study, we investigated the intertemporal decisions of people who engaged in mindfulness  
36 meditation. In our lab experiments, people were guided through short mindfulness exercises  
37 and then asked to make economic decisions with an intertemporal component. In our field  
38 study, participants completed an eight-week mindfulness course organized by an established  
39 mindfulness institute, and they faced intertemporal choices before and after their training. Our  
40 results show that mindfulness does not make people more present-biased or substantially affect  
41 their intertemporal decisions.

42

43 **Keywords:** Mindfulness, meditation, intertemporal choice, present-bias, impatience.

44

45 The analyses for all our experiments were conducted using STATA and R. Data is provided  
46 within the manuscript. The code, experimental instructions, and supplementary materials and  
47 analyses can be found in the following repository:

48 [https://osf.io/swg79/?view\\_only=c3bc543bedee4b48ad23b6d65002521e](https://osf.io/swg79/?view_only=c3bc543bedee4b48ad23b6d65002521e)

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57       “Breathing in, I calm body and mind. Breathing out, I smile. Dwelling in the present moment  
58       I know this is the only moment.”

59

— Thich Nhat Hanh

60

## 61       **Introduction**

62       Mindfulness is no longer confined to yoga studios and meditation retreats; it has become  
63       mainstream in boardrooms and executive training programs worldwide. Many Fortune 500  
64       companies — including industry giants like Google, Apple, Nike, and Goldman Sachs — have  
65       adopted mindfulness initiatives aimed at sharpening leaders' attention, reducing stress, and  
66       improving decision-making [1–2]. Yet, despite its widespread popularity and enthusiastic  
67       endorsements by millions of professionals, mindfulness practice has recently come under  
68       scrutiny for potential unintended consequences. Scholars and commentators have raised  
69       concerns that mindfulness, by enhancing focus on the present moment, might unintentionally  
70       encourage present-biased decision-making — leading people to prioritize immediate  
71       gratification at the expense of future benefits [3–6]. Given mindfulness's rapid and extensive  
72       integration into corporate cultures, rigorously investigating its impact on intertemporal choice  
73       is critical for both organizational effectiveness and individual well-being.

## 74       **General effects of mindfulness and potential concerns for**

### 75       **intertemporal decision-making**

76       Mindfulness, defined as “the awareness that arises through paying attention, on purpose,  
77       in the present moment, non-judgmentally” [7], has positive effects that have been well-  
78       established across many domains. Mindfulness interventions have been frequently applied in  
79       clinical contexts — when the group under study is diagnosed with or considered at risk of a

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

80 healthcare condition — to improve a variety of conditions, such as anxiety and personality  
81 disorders, substance abuse, stress, as well as chronic pain, to name just a few [8]. Several  
82 reviews have summarized the main findings and generally supported the clinical effectiveness  
83 of mindfulness (see [9–17]). This effectiveness continues to find support in recent research (see  
84 [16, 18–22]). Moreover, the positive effects of mindfulness are not limited to clinical contexts.  
85 Results from brain imaging studies also found a positive relationship between mindfulness and  
86 neural functions associated with well-being [23–25]. These results have been supported by  
87 studies that found mindfulness was positively associated with life satisfaction by reducing  
88 deviations from a balanced time perspective (DBTP) [21, 26]. Other studies found positive  
89 effects on immune system parameters [27–29].

90 Research on the effects of mindfulness on decision-making is much more limited, but  
91 studies have also found positive effects in important domains, such as ethical decisions [30–  
92 31], cooperative and pro-social decisions [32–33], and the sunk-cost bias [34]. Mindfulness  
93 was also associated with enhanced self-regulation and reduced aversion to waiting [35] and to  
94 lower risk-aversion [36]. Overall, while mindfulness received some criticism [37–39], the  
95 overarching conclusion is that it has desirable effects on mental health, well-being and decision-  
96 making.

97 While the practice of mindfulness may improve ethical, pro-social, and reduce risky  
98 decisions people make for their present self, could it actually hurt the intertemporal choices  
99 they make for their future self? Mindfulness practices are broadly conceived to focus meditators  
100 on the present moment, abstracting from concerns related to the past or the future. This can  
101 have beneficial effects on several well-being metrics but it might also make meditators discount  
102 the value of future outcomes more heavily and become more present-biased. A large literature  
103 in economics, psychology, and neuroscience has shown that when people made choices that  
104 involve intertemporal trade-offs — such as the ones related to saving, climate change  
105 mitigation, healthy eating, etc. —, they tend to hyperbolically discount the value of future  
106 rewards and display a drive for immediate gratification (see [40–41]). These tendencies have  
107 been linked to numerous detrimental behaviors, such as lower academic performance [42],

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

108 insufficient saving [43], unhealthy eating and obesity [44–45], and substance and alcohol use  
109 disorders [46], among others. They have also been linked to important life outcomes such as  
110 divorce rates [47] or suicide rates [48]. If, by focusing people on the present, mindfulness  
111 exacerbates these patterns, its practice could also carry negative consequences. Some existing  
112 research has linked mindfulness with patterns of substance abuse related to impulsivity [49].  
113 Given the increased prevalence of mindfulness across Western societies, these patterns could,  
114 in turn, have a substantial detrimental impact on the economy and society. Therefore, in the  
115 present paper, we aim to systematically study how mindfulness meditation affects intertemporal  
116 decision-making.

117 **Empirical evidence on the effects of mindfulness on**  
118 **intertemporal choice**

119 Several studies have examined the effects of mindfulness or related treatments on  
120 intertemporal decision-making in clinical and general populations — with very mixed results  
121 for both. At the clinical population level, several studies concluded that mindfulness can  
122 successfully reduce delay discounting by improving self-regulation, tolerance, and present-  
123 moment acceptance (see [19–20, 36, 50–52]). However, other studies found null effects [53–  
124 54]; and others argued that mindfulness can even increase discount rates depending on the  
125 population under study (see [55]). For instance, Rasmussen et al. [56] found that a mindful  
126 eating treatment increased the discounting of rewards for food-insecure women. These findings,  
127 which run contrary to the results of previous studies, are interpreted as mindfulness inducing  
128 them into a ‘survival mode’ by increasing their focus on present scarcity.

129 At the general population level, several studies found that mindfulness can reduce discount  
130 rates [35, 50, 57–62], but that this effect can be contingent upon framing [35] and whether the  
131 type of reward is primary (e.g. food) or monetary [51]. These contingent effects are consistent  
132 with other studies that found that delay discounting is decreased when time is framed in terms  
133 of calendar dates rather than days [63–64]; and studies that found that primary rewards are

134 discounted more steeply than monetary rewards [40, 65]. Other studies found no significant  
135 effects of mindfulness or related treatments on intertemporal choices [18, 51, 66]. Finally, some  
136 studies suggested that mindfulness induces higher discount rates when making financial  
137 decisions [56–57].

## 138 **Limitations of previous research**

139 Research on the effects of mindfulness on intertemporal choice is still underexplored, and  
140 existing studies are limited in at least four ways.

141 First, most studies focus on very specific clinical populations and decision-making  
142 contexts. For instance, Hendrickson and Rasmussen [50–51] studied the effects of mindfulness  
143 to reduce impulsive food choices to curb obesity. Similarly, Rasmussen et al. [55] focused on  
144 food-insecure women, while Shead et al. [54] focused on regular gamblers, and Yao et al. [52]  
145 on gaming addiction. The relatively narrow focus of these studies — while important — makes  
146 it difficult to draw meaningful conclusions for the hundreds of millions of non-clinical  
147 mindfulness practitioners worldwide who make everyday decisions about how to invest their  
148 time and money (see [19–20, 36, 50–55]).

149 Second, the handful of studies that did examine the relationship between mindfulness  
150 and inter-temporal choice in the general population mainly did so using observational data. For  
151 instance, Duchêne et al. [66], Marcowski et al. [36], and Murphy & McKillop [20] did not  
152 induce mindfulness but estimated statistical relationships based on self-reported traits, using  
153 questionnaires like the Five Facet Mindfulness Questionnaire. This makes it impossible to  
154 establish causality , especially given socially desirable responding and the conceptual overlap  
155 between trait-mindfulness and inter-temporal choice questionnaires.

156 Third, regardless of the focus and methodological approach of the study, mindfulness  
157 manipulations or traits have often been mixed with other factors and interventions (see [36, 52–  
158 53, 60–62]). For instance, Yao et al. [52] analyzed the impact of a mindfulness intervention on  
159 decision impulsiveness to cure Internet gaming disorder. However, the authors used a

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

160 combination of classic therapy and mindfulness meditation as a treatment, which did not allow  
161 for the isolation of the effects of mindfulness. Similarly, Morrison et al. [60] and Morrison et  
162 al. [53] used acceptance-based training sessions as a treatment to investigate the effects of  
163 mindfulness on impulsivity. Although acceptance-based therapy is rooted in the philosophy  
164 that underpins the mindfulness movement, in the acceptance-based therapy used in the study,  
165 no formal mindfulness training was given to treated participants. Scholten et al. [61] also mixed  
166 mindfulness and acceptance-based training as a treatment, whereas Wang et al. [62] used a  
167 broad Buddhist practice intervention that includes some mindfulness but also other elements  
168 like chanting. Finally, observational research by Marcowski et al. [36] used a composite of  
169 trait mindfulness and psychological flexibility scores to create behavioral profiles and relate  
170 them to discounting.

171 Fourth, existing studies often exclusively used very short mindfulness treatments and  
172 focused on measuring the immediate effects of mindfulness on intertemporal decision-making  
173 (see [35, 54, 56–59]). For instance, Bazley et al. [56–57], Dixon et al. [58], and Errmann [35]  
174 used mindfulness video or audio inductions of just about 5 minutes, casting doubt on the  
175 conceptual validity of the manipulation and providing no insights into the medium- and long-  
176 term effects of mindfulness. To the best of our knowledge, only Alem et al. [18] has investigated  
177 the effects of a fully-fledged mindfulness program (conducted online for four weeks) on  
178 intertemporal choices — finding some patterns suggestive of reduced time discounting but no  
179 significant results.

180 **The present research**

181 We present four studies — three lab experiments and one longitudinal field experiment  
182 — that address the aforementioned gaps and improve upon the existing literature in five main  
183 ways.

184 First, in our filed study (Experiment 4), we used as a treatment the popular and intensive  
185 course in mindfulness known as the Mindfulness-Based Stress Reduction (MBSR) program,

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

186 designed by Jon Kabat-Zinn, which is conducted onsite and lasts for eight weeks. This improves  
187 on the most robust mindfulness manipulation used until now in the domain of intertemporal  
188 choice (see [18]), which lasted for four weeks and was conducted online.

189 Second, expanding on previous studies on the topic, our research combines both  
190 manipulation- and training-based approaches. This is important because we can compare the  
191 effects of momentary interventions that change the current state of the participants to the  
192 consequences of a longer program that reshapes behavior at a deeper level.

193 Third, we also combine our manipulation- and training-based approaches with the  
194 measurement of participants' trait mindfulness levels and experience in mindfulness practice.  
195 This allows us to analyze if either personal dispositions or experience over longer periods of  
196 time are associated with any particular effects.

197 Fourth, our research uses three different tasks to measure intertemporal choice behaviors.  
198 In Experiments 1 and 4, we used a reduced number of classic choices between smaller amounts  
199 of money sooner and larger amounts of money later, designed to measure both intertemporal  
200 preferences and time inconsistency. In Experiment 3, we implemented a more sophisticated  
201 design involving more choices that allows us to estimate model parameters related to time  
202 discounting and to preferences for immediate gratification. In Experiment 2, we employed  
203 hypothetical real-world scenarios to analyze intertemporal decisions in more realistic settings.  
204 This combination of tasks provides a relatively comprehensive analysis of inter-temporal  
205 decision making.

206 Fifth, in our research, we employ a mixture of hypothetical and real choices (i.e.,  
207 participants knew that in certain tasks, they could actually obtain the outcome that they chose  
208 at the proposed time). This provides a test of the effects of mindfulness on intertemporal choice  
209 across different incentive systems.

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213 **Experiment 1**

214 **Method**

215 We assigned 323 participants — in two different waves — (219 women, 104 men; *mean*  
216 *age* = 21 years, *range* = 18–51 years) to mindfulness, mind wandering, and control conditions,  
217 randomizing per session. Participants were students and local residents from the Universitat  
218 Pompeu Fabra Behavioral and Experimental Sciences Laboratory participant pool who  
219 responded to an advertisement offering 8€ for participation. Each participant sat in a semi-  
220 private cubicle within a laboratory. Our mindfulness and mind-wandering induction procedures  
221 drew on established methods [7, 34, 67–68]. All methods were carried out in accordance with  
222 relevant guidelines and regulations. The experimental protocols were approved by the Ethics  
223 Committee of Universitat Pompeu Fabra. Informed consent was obtained from all participants  
224 involved in the studies.

225 Participants listened to a 15-min audio-recorded induction created specifically for this  
226 research by a professional mindfulness-meditation instructor. Participants were led through a  
227 focused-breathing meditation exercise that instructed them to focus on the physical sensations  
228 of breath entering and leaving their body and repeatedly reminded them to focus on their  
229 experience of breathing. The content of the mind-wandering induction repeatedly instructed  
230 participants to think of whatever came to mind. Although this type of induction has commonly  
231 been used as a control condition in prior mindfulness experiments [34, 67–68] because it  
232 replicates a waking, baseline mental state [69], research has shown that such “forced” mind-  
233 wandering can be aversive (e.g., [70]). Therefore, going beyond prior studies, we also included  
234 a true control condition in which participants were not subjected to any manipulation, ensuring  
235 that observed mindfulness effects were not merely driven by negative reactions to the mind-  
236 wandering condition.

237 Then, participants in the first of the two waves of the experiment, in which we had a total  
238 of 129 participants (78 women, 41 men; *mean age* = 22 years; *age range* = 18–51), had to

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

239 complete three manipulation check items in order to advance in the experiment — something  
240 we did not include in the second wave and in the rest of the studies since we demonstrated the  
241 effectiveness of our manipulations in this first wave. The three items are included in the  
242 instructions in our online repository. The first two items were designed to test to which extent  
243 participants had been focused on the present moment or on the physical sensations in their  
244 bodies. The third item was designed to test to which extent participants had been mind-  
245 wandering. We measured the three items with a 5-point Likert scale.

246 **Table 1. Experiment 1 Task Description.**

Choice	Option A (sooner reward)	Option B (later reward)
1	€200 today	€220 in 4 weeks
2	€200 in 12 weeks	€220 in 16 weeks
3	€200 today	€250 in 4 weeks
4	€200 in 12 weeks	€250 in 16 weeks

247 Notes: In each decision, respondents chose between a sooner, smaller reward (Option A) and a later,  
248 larger reward (Option B). All amounts are nominal euros.

249 Finally, all participants made four hypothetical choices, presented in Table 1, between two  
250 assets. Asset A always offered a smaller amount of money (200€) sooner in time (either now  
251 or in 12 weeks). Asset B always gave a larger amount of money (either 220€ or 250€) at a later  
252 time (either in 4 or in 16 weeks). In two of the choices, asset A offered amounts only in the  
253 present moment, while in the other two choices, both Asset A and B offered delayed monetary  
254 rewards. Thus, participants made choices between smaller immediate rewards and larger later  
255 rewards or between smaller later rewards and larger even later rewards. Additionally, by  
256 offering the same monetary amounts with the same temporal distance between them at different  
257 moments in time, we can compare whether participants behaved in a time-consistent manner.  
258 Time inconsistency is one of the central phenomena in research on intertemporal decision  
259 making, and it typically takes the form of choosing the smaller sooner amount when it is close  
260 to the present (as in choices 1 and 3) and switching to the larger later reward when the same

261 options are delayed into the future (as in choices 2 and 4) (see [41]). The experiment also  
262 included other items, mostly related to risk aversion, ambiguity aversion, and loss aversion, not  
263 relevant to the current study.

264 **Results and discussion**

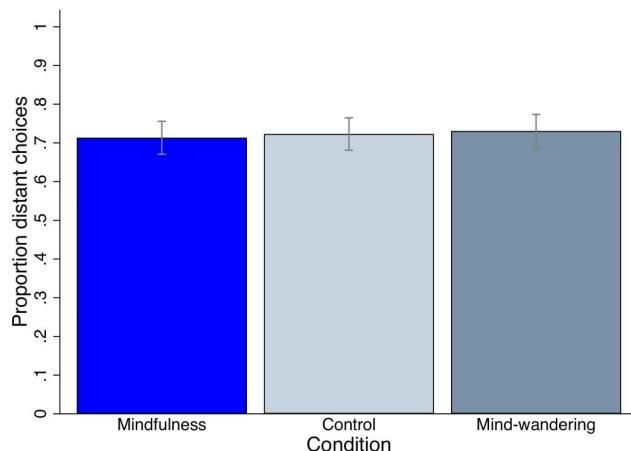
265       **Manipulation check.** Both our mindfulness and mind-wandering manipulations were  
266 effective. Participants in the mindfulness condition reported being significantly more focused  
267 on the present moment (*mean* = 3.64) than participants in the mind-wandering condition (*mean*  
268 = 2.61;  $t = 4.87$ ,  $p < .0001$ ) and the control condition (*mean* = 2.98;  $t = 3.00$ ,  $p = .0035$ ). There  
269 was no significant difference between the control and mind-wandering conditions ( $t = 1.41$ ,  $p$   
270 = .1618). Likewise, participants in the mindfulness condition reported being significantly more  
271 focused on their body sensations (*mean* = 3.73) than participants in the mind-wandering  
272 condition (*mean* = 1.95;  $t = 8.81$ ,  $p < .0001$ ) and control condition (*mean* = 2.30;  $t = 6.35$ ,  $p <$   
273 .0001). Again, there was no significant difference between the control and mind-wandering  
274 conditions ( $t = 1.36$ ,  $p = .1773$ ). Finally, participants in the mind-wandering condition reported  
275 more mind-wandering (*mean* = 3.47) than participants in the mindfulness condition (*mean* =  
276 2.82;  $t = 2.76$ ,  $p = .0072$ ) and control condition (*mean* = 2.83;  $t = 2.47$ ,  $p = .0154$ ). There was  
277 no significant difference between the mindfulness and control conditions ( $t = -.05$ ,  $p = .9599$ ).

278       **Time Preference.** As depicted in Figure 1, most participants preferred more money later  
279 over less money sooner, regardless of the experimental condition. The portion of participants  
280 who chose the distant alternative across the four choices in the mindfulness condition (71%)  
281 did not differ from the mind-wandering condition (73%),  $\chi^2$  (1,  $N = 848$ ) = .23,  $p = .6280$ , or  
282 the control condition (72%),  $\chi^2$  (1,  $N = 880$ ) = .06,  $p = .8073$ . The proportion of people  
283 selecting the distant alternative in the mind-wandering and control conditions did not  
284 statistically differ either,  $\chi^2$  (1,  $N = 856$ ) = .03,  $p = .8628$ .

285       A similar pattern emerged when focusing exclusively on the two choices that contrasted  
286 immediate and future rewards (Figure 2A), and the two choices that contrasted rewards in the

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

287 near vs. distant future (Figure 2B). Specifically, participants in the mindfulness condition chose  
288 the future rewards over immediate ones in similar proportions (69%) than participants in the  
289 mind-wandering condition (71%),  $\chi^2 (1, N = 424) = .13, p = .7208$ , and the control condition  
290 (70%),  $\chi^2 (1, N = 440) = .05, p = .8183$ . Likewise, participants in the mindfulness condition  
291 chose the more distant over the less distant future rewards in similar proportions (74%) than  
292 participants in the mind-wandering condition (75%),  $\chi^2 (1, N = 424) = .05, p = .8284$ , and the  
293 control condition (74%),  $\chi^2 (1, N = 440) = .00, p = .9968$ . In addition, the choice proportions  
294 of delayed vs. immediate rewards and more distant vs. less distant future rewards did not  
295 statistically differ between the mind-wandering and control conditions,  $\chi^2 (1, N = 428) = .00,$   
296  $p = .9755$  and  $\chi^2 (1, N = 428) = .01, p = .9146$ , respectively.



297

298 **Fig 1. Proportion of distant choices per condition.** This figure shows the proportion of distant  
299 choices combining the four intertemporal decisions as a function of the condition. Larger values  
300 indicate a greater proportion of distant choices. Error bars indicate the 95% confidence interval.

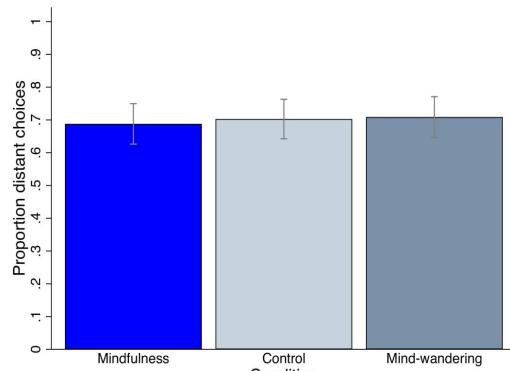
301 **Time consistency.** In Experiment 1, participants could potentially be time inconsistent  
302 two times, one comparing choices 1 and 2 and another comparing choices 3 and 4. To analyze  
303 if there were differences between conditions in this respect, we gave participants a time-  
304 consistency score of 1 if their choices were fully time-consistent and zero otherwise. We then  
305 tested whether the proportions of people with a score of 1 were significantly different across  
306 conditions. Figure 3 shows that the portion of participants choosing in a fully consistent way  
307 in the mindfulness condition (67%) was not significantly different from the proportion in the

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

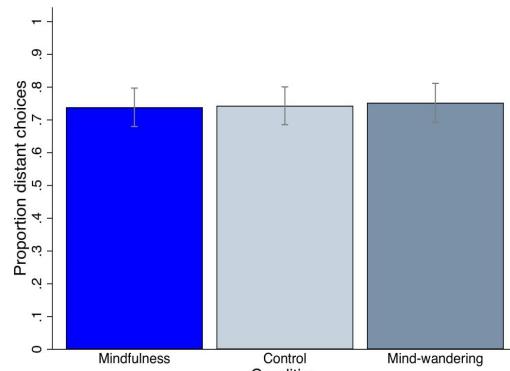
308 mind-wandering condition (65%)  $\chi^2 (1, N = 212) = .02, p = .8803$ , or the control condition  
309 ( $72\%$ )  $\chi^2 (1, N = 220) = .46, p = .4995$ . The difference in the proportion of fully consistent  
310 people between the mind-wandering and control conditions was also not statistically  
311 significant  $\chi^2 (1, N = 214) = .92, p = .3373$ .

312

A



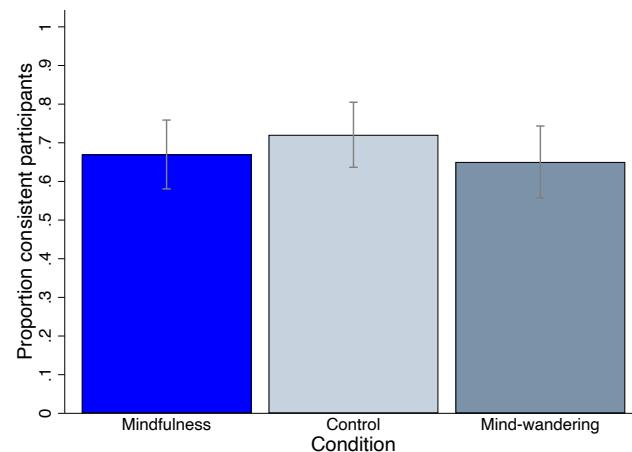
B



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314 **Figure 2. Proportion of distant choices per condition for choices between immediate vs. future**  
315 **rewards (Panel A) and near vs. distant future rewards (Panel B).** Error bars indicate the 95%  
316 confidence interval.

317 Overall, Experiment 1 shows that a 15-min mindfulness meditation does not increase the  
318 preference for short-term gratification using a typical task to study intertemporal decision  
319 making. This is true both when considering differences in choices involving all types of delays  
320 and when breaking down decisions between immediate vs. future alternatives and near vs.  
321 distant future alternatives. Moreover, mindfulness does not change the consistency of people's  
322 time preferences. Finally, all these null results were robust to including participants' age and  
323 gender as control variables (see Table S1 in the Supplementary Materials, located in our online  
324 repository).



325

326      **Figure 3. Proportion of time consistent participants per condition.** This figure shows the  
327      proportion of time consistent participants. We define a participant as time consistent if they chose  
328      the same asset selections in the first and second choices and in the third and fourth choices. Large  
329      numbers indicate a greater proportion of time-consistent choices. Error bars indicate the 95%  
330      confidence interval.

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339 **Experiment 2**

340 **Method**

341 To investigate the effects of mindfulness beyond the classical intertemporal tasks found in  
342 the decision-making literature, we decided to use additional tasks inspired by real-world  
343 scenarios. In Experiment 2, our procedures were generally the same as in Experiment 1. We  
344 assigned 140 participants (83 women, 57 men; *mean* age = 21 years; age range = 18–37) to a  
345 mindfulness, a mind-wandering, or a control condition, randomized per session. Participants in  
346 the mindfulness and mind-wandering conditions completed the same mindfulness or mind-  
347 wandering induction procedures as in Experiment 1 and all procedures were approved by the  
348 Ethics Committee of Universitat Pompeu Fabra and conducted in accordance with the same  
349 ethical standards described for Experiment 1.

350 Next, all participants completed three hypothetical resource allocation scenarios.  
351 Participants were first asked to imagine that they had won €10,000 and to indicate how much  
352 of this money they would allocate to immediate hedonic spending and how much they would  
353 save in a checking account for a rainy day. Next, participants were told to imagine that they  
354 earned €2,000 net per month (a decent salary at the time for our Spanish undergraduate  
355 participants) and indicate how they would allocate this monthly income between saving in a  
356 checking account, pension plan, rent and home expenses, leisure, food, and other expenses (all  
357 summing to €2,000). Finally, participants were asked to imagine a scenario where they needed  
358 to buy a car but had only half the necessary funds. They were presented with two options:  
359 borrow the remaining amount from a bank and pay interest, or wait two years until they had  
360 saved enough, avoiding any interest payments (see the instructions in the online repository for  
361 full details).

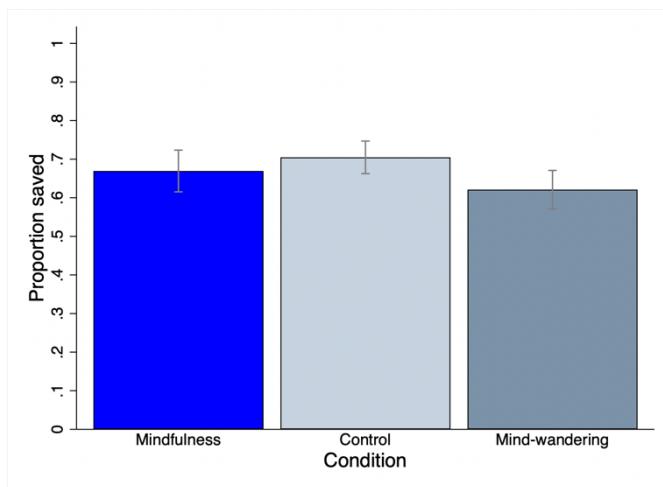
362 These three tasks capture in different ways the trade-off between smaller short-term  
363 rewards and larger longer-term ones. In the first task, allocating more money to savings  
364 indicates a stronger preference for the long-term rewards, which is typically associated with

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

365 less time discounting. In the second task, the options that represent longer-term rewards are  
366 saving and investing in a pension plan. In the third one, the longer-term option is waiting to  
367 avoid paying interest. These tasks were not directly incentivized; participants received a fixed  
368 payment for participation. The experiment also included other items, mostly related to risk  
369 aversion, ambiguity aversion, and loss aversion, not relevant to the current study.

370 **Results and discussion**

371 Figure 4 gives the average proportion of money allocated to longer-term options per  
372 condition across scenarios 1 and 2 (i.e., the amounts allocated to the long term divided by the  
373 total available amount). Because participants could allocate any amount along a continuous  
374 scale (rather than making binary choices), and these allocation proportions showed non-normal  
375 distributions, we used Wilcoxon rank-sum tests (Mann–Whitney U) to compare conditions.  
376 Mindful participants allocated a statistically similar proportion to longer-term alternatives ( $M$   
377 = .669) than did mind-wandering participants ( $M$  = .621;  $W$  = 1237,  $p$  = .1345) and control  
378 participants ( $M$  = .705;  $W$  = 841,  $p$  = .3054). The difference between mind-wandering  
379 participants and control participants was significant this time ( $W$  = 1599,  $p$  = .0044), with mind-  
380 wandering resulting in favoring the long term less than in the control condition.

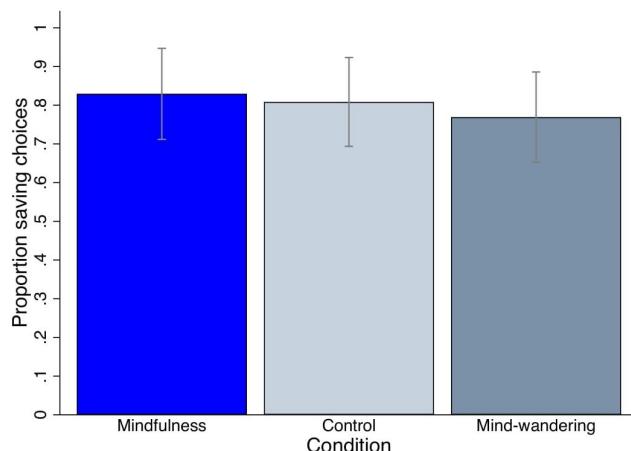


381  
382 **Figure 4. Proportion of money allocated to longer-term options per condition in tasks 1 and**  
383 **2.** This figure shows the average proportion of money allocated to the longer-term options as a  
384 function of the condition. Larger values indicate greater allocations to the long term. Error bars  
385 indicate the 95% confidence interval.

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

386       Figure 5 shows the proportion of participants who chose the waiting alternative per  
387       condition in scenario 3. In the mindfulness condition, the proportion of participants deciding to  
388       wait (83%) was not significantly different from that of mind-wandering participants (77%)  $\chi^2$   
389       (1,  $N = 93$ ) = .51,  $p = .476$ , or the control condition (81%)  $\chi^2$  (1,  $N = 88$ ) = .06,  $p = .801$ . The  
390       difference in the proportion of people selecting the waiting alternative between the mind-  
391       wandering and control conditions was also not significant  $\chi^2$  (1,  $N = 99$ ) = .23,  $p = .633$ .

392       The results of Experiment 2 replicate the findings from Experiment 1 using more  
393       naturalistic scenarios: A 15-min mindfulness meditation does not increase people's propensity  
394       to pursue immediate gratification over longer-term rewards. These results hold true even when  
395       controlling for age, gender, and previous meditation practice (see Table S2 in Supplementary  
396       Materials). However, Experiments 1 and 2 relied on hypothetical decisions, which may not  
397       fully reflect actual decision-making behavior. In Experiment 3, we address this limitation by  
398       examining whether mindfulness affects intertemporal choices in consequential, incentive-  
399       compatible contexts. Going beyond basic choice scenarios, we use a more comprehensive  
400       measure of intertemporal choice, allowing us to separately measure how people choose between  
401       delayed rewards and how strongly they favor immediate rewards over delayed ones (present  
402       bias).



403

404       **Figure 5. Proportion of participants who chose the waiting alternative per condition in task 3.**  
405       This figure shows the proportion of waiting decisions per condition in task 3. Large numbers  
406       indicate more waiting. Error bars indicate the 95% confidence interval.

407 **Experiment 3**

408 **Method**

409 We assigned 389 participants (263 women, 126 men; *mean* age = 21 years; age range =  
410 18–51) to mindfulness, mind-wandering, and control conditions, randomizing per session.  
411 Participants in the mindfulness and mind-wandering conditions completed the same  
412 mindfulness or mind-wandering induction procedure as in Experiments 1 and 2. Again,  
413 participants in the control condition were not subjected to any procedure and all procedures  
414 were approved by the Ethics Committee of Universitat Pompeu Fabra and conducted in  
415 accordance with the same ethical standards described for Experiment 1.

416 Next, all participants made 42 choices as in McClure et al. [71] between receiving smaller  
417 cash amounts (between 5€ and 34€) earlier (immediately, two weeks from the day of the  
418 experiment, or four weeks from the day of the experiment) and larger cash amounts (between  
419 7€ and 43€) later (two, four, or six weeks, respectively, from the day of the experiment). The  
420 full list of the 42 choices is available in the instructions in our online repository. Participants  
421 made choices between smaller immediate rewards and larger later rewards, or between smaller  
422 later rewards and larger even later rewards. We incentivized participants to express their true  
423 preferences by randomly selecting 1 out of every 25 participants to realize one of their choices,  
424 paying the preferred alternative for a randomly selected choice pair. The payment was made at  
425 the chosen moment and for the chosen amount using electronically sent Amazon.com gift  
426 certificates. The experiment also included other items, mostly related to risk aversion,  
427 ambiguity aversion, and loss aversion, not relevant to the current study.

428 This task allows us to estimate the parameters of a model that distinguishes between  
429 two types of processes, as captured by the following quasi-hyperbolic discounting function  
430 (Laibson [72], O'Donoghue and Rabin [73]):

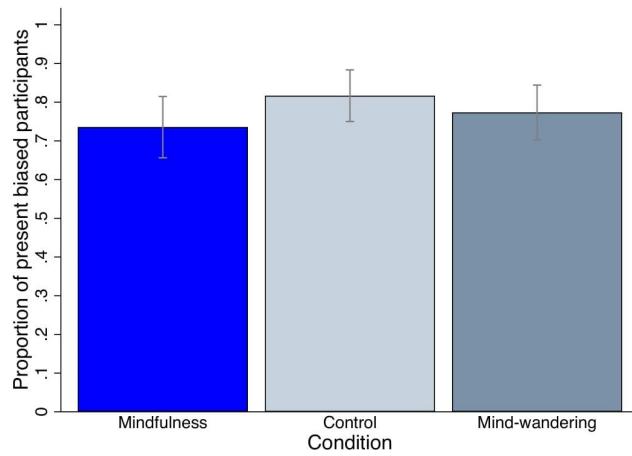
431 
$$D(t) = \begin{cases} 1 & \text{if } t = 0 \\ \beta \delta^t & \text{if } t > 0 \end{cases}$$

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

432 One of the parameters of the function ( $\delta$ ) captures time discounting when choosing  
433 between delayed rewards, assuming exponential discounting. The other parameter ( $\beta$ ) captures  
434 what is known as present bias: the additional discounting occurring between rewards in the  
435 present and any delayed rewards. This bias occurs when  $0 < \beta < 1$ , and it is also associated with  
436 time-inconsistent behavior.

437 **Results and discussion**

438 To test whether the parameters differ between the three conditions, we estimated them by  
439 fitting each participant's choices to the quasi-hyperbolic discounting function using maximum-  
440 likelihood estimation, constraining  $\beta$  and  $\delta$  to be between 0 and 1. This approach generated two  
441 parameters per participant, allowing us to compare the distribution of  $\beta$  and  $\delta$  across conditions.



442  
443 **Figure 6. Proportion of present-biased participants per condition.** This figure shows the  
444 proportion of individuals displaying present bias ( $\beta$  parameters lower than 1) as a function of the  
445 condition. Error bars indicate the 95% confidence interval.

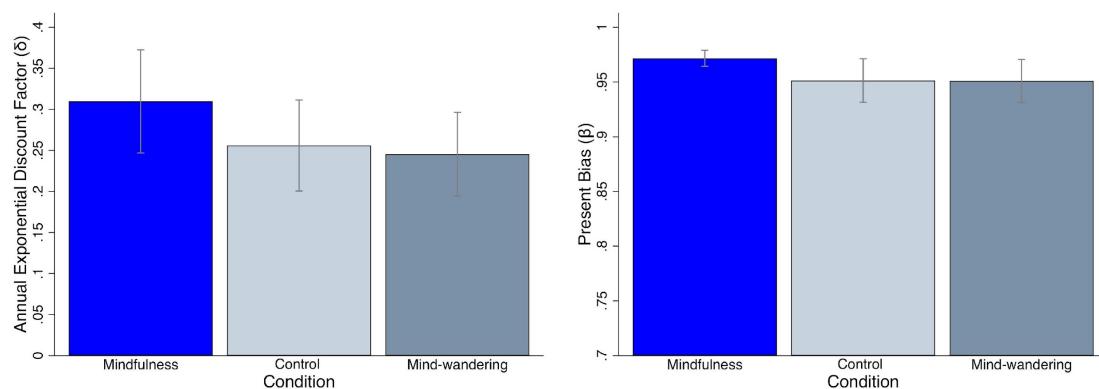
446 To begin with, Figure 6 presents the proportion of participants who showed present bias  
447 (by displaying a  $\beta$  lower than 1) in the different conditions.  $\beta$  was statistically equally likely to  
448 be lower than 1 among mindfulness participants (74%) than among mind-wandering  
449 participants (77%)  $\chi^2 (1, N = 258) = .51, p = .476$ , and also not significantly different than the  
450 share of present-biased participants in the control condition (82%)  $\chi^2 (1, N = 252) = 2.40, p$

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

451 = .121. The difference in the proportion of people with present bias between the mind-  
452 wandering and control conditions was also not significant  $\chi^2(1, N=268) = .76, p = .383$ .

453 Figure 7 shows the mean values for both parameters ( $\beta$  and  $\delta$ ) in the three conditions. When  
454 we compare conditions focusing on discounting in choosing between delayed rewards ( $\delta$ ), we  
455 observe that mindful participants (*mean*  $\delta = .310$ , *median*  $\delta = .154$ ) discounted equally as mind-  
456 wandering participants (*mean*  $\delta = .245$ , *median*  $\delta = .104$ ;  $W = 9238, p = .1128$ ). Due to the  
457 markedly skewed distributions of the intertemporal choice parameters  $\beta$  and  $\delta$  (see Figures 7  
458 and 8), we used Mann–Whitney U instead of parametric t-tests to compare conditions in this  
459 case. The results show there were also no differences in terms of  $\delta$  between mindful participants  
460 and control participants (*mean*  $\delta = .256$ , *median*  $\delta = .126$ ;  $W = 8757, p = .1506$ ) or between  
461 mind-wandering and control participants ( $W = 9045, p = .9115$ ).

462 Focusing on the degree of present bias ( $\beta$ ), we see that mindful participants (*mean*  $\beta =$   
463  $.972$ , *median*  $\beta = .989$ ) did not significantly differ from mind-wandering participants (*mean*  $\beta$   
464 =  $.951$ , *median*  $\beta = .987$ ;  $W = 8629, p = .5669$ ), or control participants (*mean*  $\beta = .951$ , *median*  
465  $\beta = .993$ ;  $W = 8282, p = .5357$ ). The level of present bias in mind-wandering (*mean*  $\beta = .951$ ,  
466 *median*  $\beta = .987$ ) and control participants did not significantly differ either ( $W = 8581, p = 1$ ).

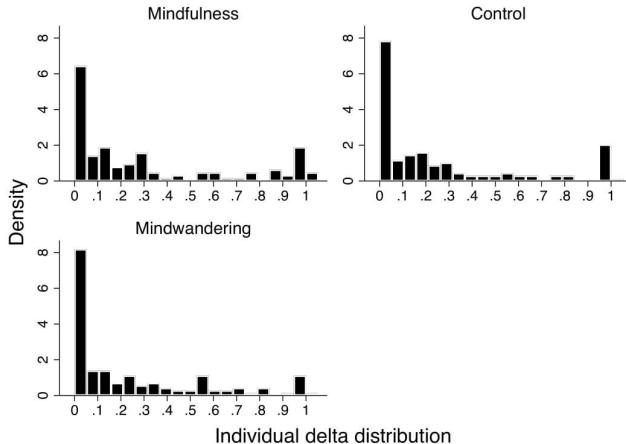


467  
468 **Figure 7. Average individual discount factor ( $\delta$ ) (left graph) and present bias ( $\beta$ ) (right  
469 graph) as a function of condition.** Large numbers indicate more patience. Error bars indicate the  
470 95% confidence interval.

471 To further test whether the distributions of individual parameters of participants in the  
472 three conditions differed, we conducted a series Kolmogorov-Smirnov tests. As depicted in  
473 Figure 8, the distribution of the  $\delta$  parameters did not differ between the mindfulness and control

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

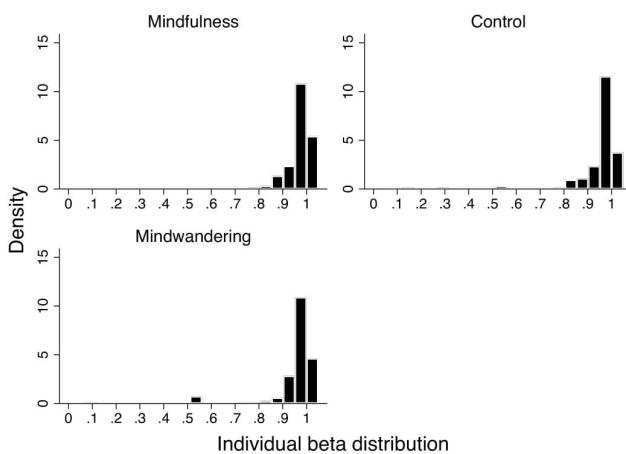
474 conditions ( $D(252) = .11; p = .359$ ), the mindfulness and the mind-wandering conditions  
475 ( $D(258) = .11; p = .367$ ), or between the control and mind-wandering conditions ( $D(268) = .07;$   
476  $p = .900$ ).



477

478 **Figure 8. Histogram of  $\delta$  per condition.** This figure shows the histogram plot of the discount factor  
479 ( $\delta$ ), plotted as a function of the condition.

480 Similarly, as depicted in Figure 9, Kolmogorov-Smirnov tests for the distribution of the  $\beta$   
481 parameters showed no difference between the mindfulness and control conditions ( $D(252) =$   
482  $.09; p = .664$ ), the mindfulness and the mind-wandering conditions ( $D(258) = .06; p = .956$ ), or  
483 between the control and mind-wandering conditions ( $D(268) = .10; p = 0.476$ ).



484

485 **Figure 9. Histogram of  $\beta$  per condition.** This figure shows the histogram plot of the present bias  
486 parameter ( $\beta$ ), plotted as a function of the condition.

487 Taken together, these results suggest that while we observe a small average trend in the  
488 individual parameter estimations, whereby mindfulness slightly decreased the desire to get

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

489 something sooner, we did not observe significant differences between the three experimental  
490 conditions in our tests.

491 In sum, Experiment 3 replicates the null effect of mindfulness on intertemporal choice  
492 using incentivized decisions and a more sophisticated estimation approach. On possibility,  
493 however, is that our short-term mindfulness induction may not have been powerful enough to  
494 fundamentally shift underlying decision-making patterns. To address this potential limitation,  
495 Experiment 4 moves beyond brief experimental interventions, exploring whether a longer,  
496 structured mindfulness training — the classic eight-week Mindfulness-Based Stress Reduction  
497 (MBSR) program — can meaningfully influence intertemporal choices.

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508 **Experiment 4**

509 **Method**

510 We partnered with the largest provider of the mindfulness-based stress reduction program  
511 (MBSR) in Spain to conduct a field experiment testing the effects of mindfulness on  
512 intertemporal choices. The MBSR program is a therapeutic intervention developed by Jon  
513 Kabat-Zinn that involves weekly group classes and daily mindfulness exercises to practice at  
514 home over an 8-week period [7]. MBSR teaches people how to increase mindfulness through  
515 yoga and meditation [7], and it is the most popular standardized program on mindfulness  
516 worldwide [74].

517 We used a within-subject design, measuring participants' intertemporal choices at two  
518 points in time: first, prior to beginning the mindfulness course during an initial orientation  
519 session, and second, in the final session — specifically, after a guided meditation midway  
520 through the two-hour session. Participation was voluntary, with no direct monetary  
521 compensation. However, participants who completed both assessments received a mindfulness  
522 book authored by the director of the center, as a token of appreciation. A total of 57 participants  
523 (32 women, 19 men; *mean* age = 43 years; age range = 23–69) completed the pre-treatment and  
524 post-treatment waves (the attrition in the second wave was 35%). All methods were carried out  
525 in accordance with relevant guidelines and regulations. The experimental protocols were  
526 approved by the Ethics Committee of Universitat Pompeu Fabra. Informed consent was  
527 obtained from all participants involved in the studies.

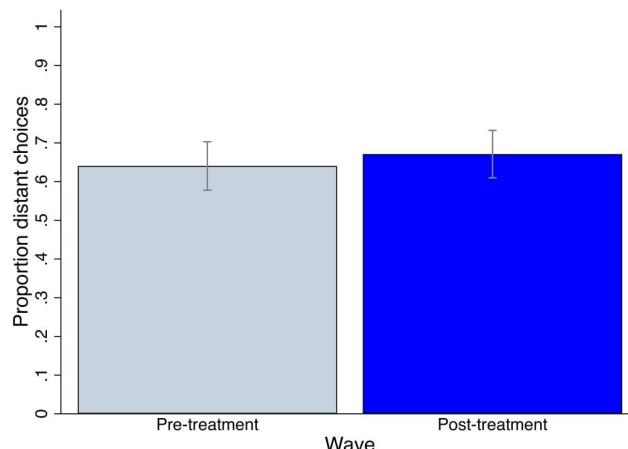
528 For practical reasons, we used the short 4-item measure of intertemporal choice from  
529 Experiment 1 in a booklet format (see the online repository). We also elicited trait mindfulness  
530 using the Mindful Attention Awareness Scale (MAAS), developed by Brown and Ryan (2003).  
531 The MAAS is a 15-item self-report questionnaire designed to assess an individual's  
532 dispositional tendency to attend to and be aware of present-moment experiences in daily life.  
533 Respondents rate items on a 6-point Likert scale ranging from 1 ("almost always") to 6 ("almost

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL CHOICE

534 never”), with higher average scores indicating greater trait mindfulness. The MAAS has been  
535 widely validated and is commonly used in psychological and behavioral research to quantify  
536 individual differences in mindful attention and awareness. In addition, we asked participants  
537 two questions related to their prior mindfulness meditation experience: whether they had ever  
538 practiced mindfulness before the program, and for how many months they had practiced  
539 mindfulness. The experiment also included other items, mostly related to risk aversion,  
540 ambiguity aversion, and loss aversion, not relevant to the current study.

541 **Results and discussion**

542 Figure 10 shows the proportion of participants who chose the distant options combining  
543 the four items in Experiment 4. In the pre-treatment wave, the portion of participants selecting  
544 the distant alternatives (64%) was not significantly different from that of the participants in the  
545 post-treatment wave (67%)  $\chi^2 (1, N = 456) = .35, p = .5543$ .



546

547 **Figure 10. Proportion of distant choices per wave.** This figure shows the proportion of distant  
548 choices combining the four intertemporal decisions as a function of the wave. Larger values indicate  
549 a greater proportion of distant choices. Error bars indicate the 95% confidence interval.

550 Table 2 reports the results from three OLS regression models further examining the  
551 impact of mindfulness training and trait mindfulness on the number of distant intertemporal  
552 choices participants made before and after the eight-week MBSR program. The dependent  
553 variable is the number of distant (i.e., delayed reward) options selected across eight binary

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

554 decisions — four in the pre-treatment wave and four in the post-treatment wave. The primary  
555 independent variables include a post-treatment indicator ( $Post_i$ ), a trait mindfulness score  
556 ( $MAAS_i$ ), and their interaction. Additional covariates include our prior meditation experience  
557 variables, age and gender. Standard errors are clustered at the participant level and are robust  
558 to heteroskedasticity.

559 In the model in column 1, which includes no controls or fixed effects, the coefficient  
560 for  $Post_i$  was not statistically significant ( $b = 0.880$ ,  $SE = 1.471$ ,  $p = .552$ ), suggesting that the  
561 average number of distant choices did not increase following the mindfulness course. Similarly,  
562 neither the main effect of trait mindfulness ( $b = 0.00295$ ,  $SE = 0.0206$ ,  $p = .887$ ) nor its  
563 interaction with the post-treatment indicator ( $b = -0.0108$ ,  $SE = 0.0234$ ,  $p = .645$ ) were  
564 statistically significant, indicating no meaningful association between mindfulness traits and  
565 intertemporal choice, nor evidence of moderation by trait mindfulness. The model explains very  
566 little of the variance in the dependent variable ( $R^2 = 0.005$ ).

567 The model in column 2 adds individual-level covariates, including two indicators for  
568 prior mindfulness practice, age, and gender. The post-treatment coefficient remained  
569 statistically non-significant ( $b = 0.925$ ,  $SE = 1.409$ ,  $p = .514$ ), as did the interaction term ( $Post_i$   
570  $\times MAAS_i$ ;  $b = -0.0107$ ,  $SE = 0.0220$ ,  $p = .628$ ). Notably, the number of months of prior  
571 meditation practice emerged as a positive and statistically significant predictor of distant  
572 choices ( $b = 0.00380$ ,  $SE = 0.00177$ ,  $p = .036$ ), suggesting that individuals with longer-standing  
573 mindfulness experience are more inclined toward delayed gratification, regardless of the formal  
574 course. Other covariates, including gender ( $b = 0.558$ ,  $SE = 0.407$ ,  $p = .176$ ) and age ( $b =$   
575  $0.00454$ ,  $SE = 0.0207$ ,  $p = .828$ ), were not significantly associated with the outcome. The  
576 inclusion of these controls slightly improved model fit ( $R^2 = 0.048$ ).

577 Column 3 introduces participant fixed effects to account for unobserved, time-invariant  
578 heterogeneity at the individual level, while omitting the additional controls. In this within-  
579 subject specification, the post-treatment effect shrinks substantially and remains non-  
580 significant ( $b = 0.919$ ,  $SE = 1.858$ ,  $p = .623$ ). The MAAS score shows a larger, though still  
581 statistically insignificant, association with distant choices ( $b = 0.0251$ ,  $SE = 0.0217$ ,  $p = .253$ ),

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

582 and the interaction term remains insignificant ( $b = -0.0152$ ,  $SE = 0.0289$ ,  $p = .601$ ). The fixed  
583 effects specification substantially increases explanatory power, with the model explaining 87%  
584 of the variance in the outcome ( $R^2 = 0.87$ ), consistent with the inclusion of subject-level  
585 intercepts.

586 **Table 2. Mindfulness Training and Mindfulness Trait Effects on Number of Distant  
587 Choices.**

<b>Dependent Variable</b>	<b>Number of Distant Choices<sub>i</sub></b>		
<b><i>Post<sub>i</sub></i></b>	0.880 (0.60)	0.925 (0.66)	0.919 (0.49)
<b><i>MAAS<sub>i</sub></i></b>	0.003 (0.14)	-0.002 (-0.08)	0.025 (1.16)
<b><i>Post<sub>i</sub> * MAAS<sub>i</sub></i></b>	-0.011 (-0.46)	-0.011 (-0.49)	-0.015 (-0.53)
<b><i>Practice Mindfulness<sub>i</sub></i></b>		-0.227 (-0.53)	
<b><i>Months Meditation Practice<sub>i</sub></i></b>		0.004* (2.15)	
<b><i>Age<sub>i</sub></i></b>		0.005 (0.22)	
<b><i>Gender<sub>i</sub></i></b>		0.558 (1.37)	
<b>Constant</b>	2.403** (2.12)	2.341 (1.62)	1.197 (0.96)
<b>Controls</b>	No	Yes	No
<b>Participant Fixed Effects</b>	No	No	Yes
<b>R<sup>2</sup></b>	0.006	0.048	0.870
<b>Observations</b>	110	110	110

588 Notes: The dependent variable in the regression models above is *Number of Distant Choices<sub>i</sub>*, which  
589 is the number of choices of asset B that participant *i* made in the 8 intertemporal choices of  
590 Experiment 4 in the pre-treatment (four) and post-treatment (four) waves. The independent variables  
591 include: *Post<sub>i</sub>*, an indicator variable that takes a value of one when participant *i* was in the second  
592 wave and zero when in the first wave; *MAAS<sub>i</sub>*, the cumulative score of participant *i* on the mindful  
593 attention awareness scale (MAAS) in the pre-treatment wave; the interaction term *Post<sub>i</sub> \* MAAS<sub>i</sub>*;  
594 *Practice Mindfulness<sub>i</sub>*, an indicator variable that takes a value of one if participant *i* had practiced  
595 meditation before the start of the course; *Months of Meditation Practice<sub>i</sub>*, which is the number of  
596 months of meditation practice stated in the pre-treatment wave of participant *i*. Standard errors are  
597 robust to heteroskedasticity. *t*-statistics are in parentheses. \*\*\*, \*\*, \* indicate significance at the  
598 1%, 5% and 10% level, respectively. Asset B always offered a higher € amount, but in a more distant  
599 moment in time than the other alternative.

600 Overall, across all three specifications, there is no evidence that participation in the  
601 mindfulness program influenced intertemporal choice behavior. Trait mindfulness as measured  
602 by the MAAS does not predict distant choices nor interacts meaningfully with the intervention.  
603 The only statistically significant finding is the positive relationship between months of prior  
604 meditation practice and patient decision-making, once the covariates are introduced in the  
605 second regression. This may suggest that accumulated meditation practice over longer periods  
606 of time is associated with a greater propensity to delay gratification in intertemporal trade-offs.  
607 However, given the nonsignificant differences in all our other mindfulness indicators, it is  
608 unclear whether this is a reliable result or a spurious correlation.

609 **General Discussion**

610 There has been growing concern both inside and outside academia about potential  
611 deleterious effects of mindfulness for decision-makers, particularly in corporate environments.  
612 Our research provides the most comprehensive investigation into the relationship between  
613 mindfulness and intertemporal choice conducted to date. Across four studies — including non-  
614 incentivized and incentivized decisions, hypothetical real-life scenarios, and a longitudinal field  
615 intervention — we found no evidence that mindfulness makes people more present-biased or  
616 impatient. While some minor differences in the opposite direction emerged in Experiment 3  
617 (see also the regressions reported in the Supplementary Materials), they did not generalize  
618 across studies, outcomes, or methods. Importantly, even a structured, eight-week Mindfulness-  
619 Based Stress Reduction (MBSR) program did not produce significant changes in intertemporal  
620 decision-making.

621 These findings suggest that the extensively documented well-being benefits associated  
622 with cultivating awareness of the present moment do not appear to compromise people's ability  
623 to make patient decisions regarding the future. This is particularly reassuring given the  
624 widespread adoption of mindfulness-based interventions in corporate and educational settings

Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

625 (see [1,75]), where the balance between immediate and delayed gratification is essential to  
626 success (see, for example, [76–77]).

627 Additionally, our findings contribute to the broader theoretical understanding of  
628 intertemporal decision-making by suggesting that temporal attention and temporal valuation  
629 might be more separable than previously assumed. Specifically, our research implies that  
630 enhancing present-focused attention through mindfulness practices does not necessarily  
631 influence how individuals value immediate versus delayed rewards. This dissociation aligns  
632 with neuroscientific evidence highlighting separate neural networks for attentional control (e.g.,  
633 the dorsal frontoparietal network [79–80]) and valuation processes (e.g., the ventromedial  
634 prefrontal cortex and ventral striatum [81–82]). Recent research further emphasizes flexible  
635 interactions among these networks rather than uniform responses to attentional manipulation  
636 alone [83].

637 Nonetheless, these results should be interpreted with care. Although we controlled for  
638 basic demographic variables (age, gender) and previous meditation experience, it is possible  
639 that specific individual differences (e.g., baseline impulsivity, cognitive control ability),  
640 contextual factors (e.g., stress, reward framing), and mindfulness components (e.g., awareness  
641 of breath vs. awareness of thoughts and feelings) could moderate the insignificant effect of  
642 mindfulness on intertemporal decision-making.

643 Overall, our research consistently demonstrates that mindfulness does not inherently  
644 promote present-biased choices. In this sense, our findings suggest that organizations and  
645 individuals can confidently adopt mindfulness practices without concerns about undermining  
646 long-term decision-making.

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Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
CHOICE

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Running head: THE EFFECTS OF MINDFULNESS ON INTERTEMPORAL  
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