

Randomized Algorithms

Coupon Collector's Problem

Joseph Chuang-Chieh Lin
Dept. CSIE, Tamkang University

Review

- Expectation of discrete random variables
- Linearity of expectation.
- Bernoulli and Binomial random variable

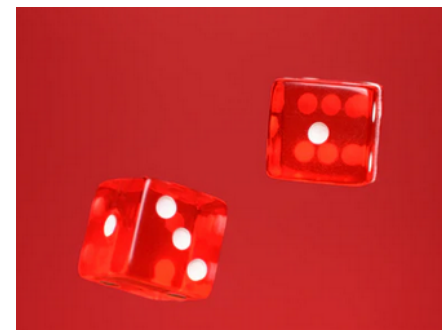
Expectation

- The expectation of a discrete random variable X , denoted by $\mathbf{E}[X]$, is

$$\mathbf{E}[X] = \sum_i i \cdot \Pr[X = i]$$

- Example: Let X denote the sum of of dices:

$$\mathbf{E}[X] = \frac{1}{36} \cdot 2 + \frac{2}{36} \cdot 3 + \frac{3}{36} \cdot 4 + \cdots + \frac{1}{36} \cdot 12 = 7.$$



Adam Fejes@adamfjs

Linearity of Expectation

- For any finite collection of discrete random variables X_1, X_2, \dots, X_n with finite expectations,

$$\mathbf{E} \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n \mathbf{E}[X_i].$$

- For any constant c and discrete random variable X ,

$$\mathbf{E}[cX] = c \cdot \mathbf{E}[X].$$

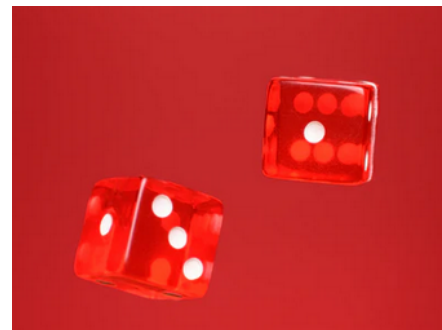
- Why is it useful?

Example

- Consider the dice-throwing example again.
 - X_1 : the outcome of die 1
 - X_2 : the outcome of die 2

$$\mathbf{E}[X_1] = \mathbf{E}[X_2] = \frac{1}{6} \cdot \sum_{j=1}^6 j = \frac{7}{2}$$

$$\mathbf{E}[X] = \mathbf{E}[X_1 + X_2] = 7.$$



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Bernoulli random variable

- Suppose we run an experiment that succeeds with probability p and fails with probability $1-p$.

$$Y = \begin{cases} 1 & \text{if the experiment succeeds,} \\ 0 & \text{otherwise.} \end{cases}$$

- Y : Bernoulli random variable.
 - or *indicator random variable*.

$$\mathbf{E}[Y] = 1 \cdot \Pr[Y = 1] + 0 \cdot \Pr[Y = 0] = \Pr[Y = 1] = p.$$



Binomial random variable

- A binomial random variable X with parameters n and p , denoted by $B(n, p)$, is defined as

$$\Pr[X = j] = \binom{n}{j} p^j (1 - p)^{n-j}.$$

for $j = 0, 1, 2, \dots, n$.

- Exercise: Show that $\sum_{j=0}^n \Pr[X = j] = 1$.

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X : the number of successful trials in the n experiments.


- Exercise: Show that $\sum_{j=0}^n \Pr[X = j] = 1$.

Binomial random variable (expectation)

- $$\begin{aligned}\mathbf{E}[X] &= \sum_{j=0}^n j \binom{n}{j} p^j (1-p)^{n-j} \\&= \sum_{j=0}^n j \frac{n!}{j!(n-j)!} p^j (1-p)^{n-j} \\&= \sum_{j=1}^n \frac{n!}{(j-1)!(n-j)!} p^j (1-p)^{n-j} \\&= np \sum_{j=1}^n \frac{(n-1)!}{(j-1)!((n-1)-(j-1))!} p^{j-1} (1-p)^{(n-1)-(j-1)} \\&= np \sum_{k=0}^{n-1} \frac{(n-1)!}{k!((n-1)-k)!} p^k (1-p)^{(n-1)-k} \\&= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^k (1-p)^{(n-1)-k} \\&= np.\end{aligned}$$

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$$(x+y)^n = \sum_{k=0}^n \binom{n}{k} x^k y^{n-k}.$$

Let's make it simpler!

- Denote a set of n Bernoulli random variables X_1, X_2, \dots, X_n .
 - $X_i = 1$ if the i th trial is successful and 0 otherwise.
 - $X = X_1 + X_2 + \dots + X_n$

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 - Compute $\mathbf{E}[X]$ using linearity of expectation:

$$\mathbf{E}[X] = \mathbf{E} \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n \mathbf{E}[X_i] = np.$$

Geometric Distribution

- Imagine: *flip a coin until it lands on a head.*
 - What's the distribution of the number of flips?

Geometric Distribution

- Imagine: *flip a coin until it lands on a head.*
 - What's the distribution of the number of flips?
- Definition. A geometric random variable X with parameter p is

$$\Pr[X = n] = (1 - p)^{n-1}p.$$

for $n = 1, 2, \dots$

- Exercise. Show that $\sum_{n \geq 1} \Pr[X = n] = 1.$

Memoryless

- Let X be a geometric random variable X with parameter $p > 0$.
- For any $n, k > 0$, $\Pr[X = n + k \mid X > k] = \Pr[X = n]$.

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- Let X be a geometric random variable X with parameter $p > 0$.
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• *Proof.*

$$\begin{aligned}\Pr[X = n + k \mid X > k] &= \frac{\Pr[(X = n + k) \cap (X > k)]}{\Pr[X > k]} \\ &= \frac{\Pr[X = n + k]}{\Pr[X > k]} \\ &= \frac{(1 - p)^{n+k-1}p}{\sum_{i=k}^{\infty} (1 - p)^i p} \\ &= \frac{(1 - p)^{n+k-1}p}{(1 - p)^k} \\ &= (1 - p)^{n-1}p = \Pr[X = n].\end{aligned}$$

The mean of a geometric r.v. $X(p)$

$$\begin{aligned}\mathbf{E}[X] &= \sum_{j=1}^{\infty} j \Pr[X = j] & \mathbf{E}[X] &= \sum_{i=1}^{\infty} (1-p)^{i-1} = \frac{1}{1-(1-p)} = \frac{1}{p}. \\ &= \sum_{j=1}^{\infty} \sum_{i=1}^j \Pr[X = j] \\ &= \sum_{i=1}^{\infty} \sum_{j=i}^{\infty} \Pr[X = j] \\ &= \sum_{i=1}^{\infty} \Pr[X \geq i].\end{aligned}$$

$$\Pr[X \geq i] = \sum_{k=i}^{\infty} (1-p)^{k-1} p = (1-p)^{i-1}$$

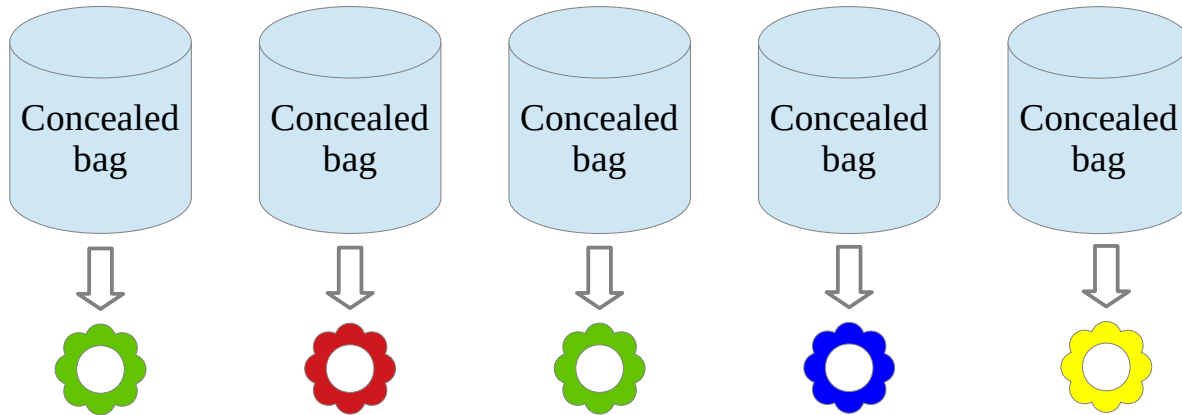
Coupon Collector's Problem



▲一名網友在臉書社團「爆廢公社公開版」表示，多年前為了收集連鎖便利商店 7-11 的一款贈品，花了不少金額，還拿到許多重複的款式，貼文引發 3 千多名網友共鳴。（圖／翻攝自爆廢公社公開版）

Coupon Collector's Problem

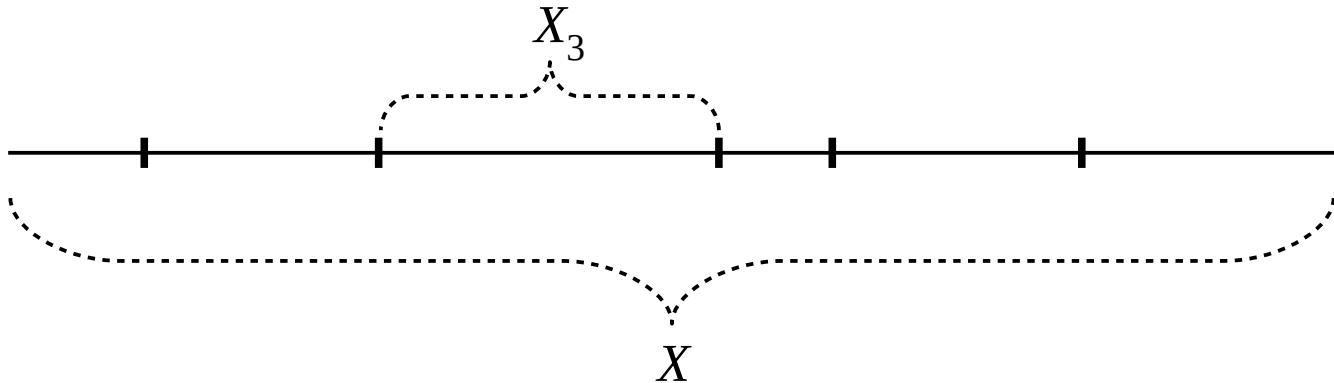
- Have you already got all of them (totally n types)?
- Have you ever thought about how much you should pay for them?



- Each bag is chosen independently and uniformly at random from the n possibilities.

Coupon Collector's Problem

- Let X be the number of bags bought until every type of coupon is obtained.
- Let X_i be the number of bags bought while you had already got exactly $i-1$ different coupons.



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- Let X be the number of bags bought until every type of coupon is obtained.
- Let X_i be the number of bags bought while you had already got exactly $i-1$ different coupons.
 - Geometric random variables?!
 - What about $X = \sum_{i=1}^n X_i$?

Coupon Collector's Problem

- When exactly $i-1$ coupons have been collected, the probability of obtaining a new one is

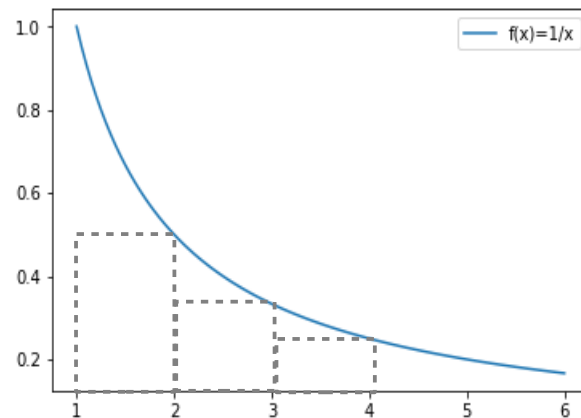
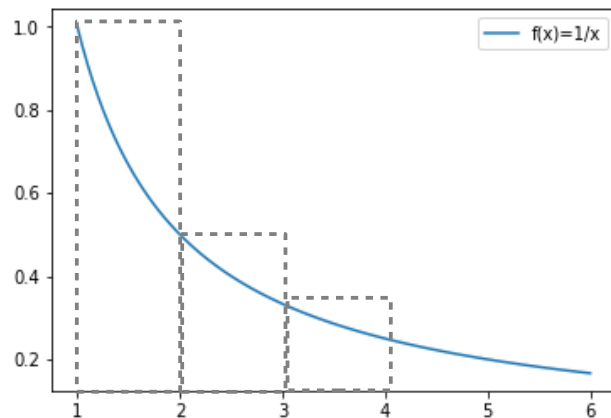
$$p_i = 1 - \frac{i-1}{n}$$

- X_i is a **geometric random variable**, so

$$\mathbf{E}[X_i] = \frac{1}{p_i} = \frac{n}{n-i+1}.$$

Coupon Collector's Problem (contd.)

$$\begin{aligned}
 \bullet \mathbf{E}[X] &= \mathbf{E}\left[\sum_{i=1}^n X_i\right] \\
 &= \sum_{i=1}^n \mathbf{E}[X_i] \\
 &= \sum_{i=1}^n \frac{n}{n-i+1} \\
 &= n \cdot \sum_{i=1}^n \frac{1}{i}.
 \end{aligned}$$



$$\sum_{k=1}^n \frac{1}{k} \geq \int_{x=1}^n \frac{1}{x} dx = \ln n$$

$$\sum_{k=2}^n \frac{1}{k} \leq \int_{x=1}^n \frac{1}{x} dx = \ln n$$

$$\rightarrow H(n) = \sum_{i=1}^n \frac{1}{i} = \ln n + \Theta(1).$$

Coupon Collector's Problem (contd.)

- So, you are about to buy $n \ln n + \Theta(n)$ bags for collecting all the coupons (stickers)!

On conditional expectation

- There are terminologies which may confusing you.

Conditional Expectation

- Definition.

$$\mathbf{E}[Y \mid Z = z] = \sum_y y \Pr[Y = y \mid Z = z].$$

- Example of two dices.
 - X_1 : the number showing on the first die
 - X_2 : the number showing on the second die
 - $X = X_1 + X_2$

$$\mathbf{E}[X \mid X_1 = 2] = \sum_{x=3}^8 x \cdot \frac{1}{6} = \frac{11}{2}.$$

Conditional Expectation (contd.)

- Lemma. For any random variables X and Y ,

$$\mathbf{E}[X] = \sum_y \Pr[Y = y] \mathbf{E}[X \mid Y = y].$$

- *Proof*.
$$\begin{aligned} \sum_y \Pr[Y = y] \cdot \mathbf{E}[X \mid Y = y] &= \sum_y \Pr[Y = y] \cdot \sum_x x \Pr[X = x \mid Y = y] \\ &= \sum_x \sum_y x \Pr[X = x \mid Y = y] \cdot \Pr[Y = y] \\ &= \sum_x \sum_y x \Pr[X = x \cap Y = y] \\ &= \sum_x x \Pr[X = x] \\ &= \mathbf{E}[X]. \end{aligned}$$

Conditional Expectation

- Lemma. For any finite collection of discrete random variables X_1, X_2, \dots, X_n with finite expectations and for any random variable Y ,

$$\mathbf{E} \left[\sum_{i=1}^n X_i \mid Y = y \right] = \sum_{i=1}^n \mathbf{E}[X_i \mid Y = y].$$

Conditional Expectation (contd.)

- A weird definition.
- $\mathbf{E}[Y \mid Z]$: regarded as a **random variable** $f(Z)$.
 - It takes on the value $\mathbf{E}[Y \mid Z = z]$ when $Z = z$.
- In the previous example,

$$\mathbf{E}[X \mid X_1] = \sum_x x \cdot \Pr[X = x \mid X_1] = \sum_{X_1+1}^{X_1+6} x \cdot \frac{1}{6} = X_1 + \frac{7}{2}.$$

- So it makes sense that

$$\mathbf{E}[\mathbf{E}[X \mid X_1]] = \mathbf{E}\left[X_1 + \frac{7}{2}\right] = \frac{7}{2} + \frac{7}{2} = 7.$$

Conditional Expectation (contd.)

- Theorem. $\mathbf{E}[Y] = \mathbf{E}[\mathbf{E}[Y \mid Z]]$.

- *Proof*.

$$\mathbf{E}[\mathbf{E}[Y \mid Z]] = \sum_z \mathbf{E}[Y \mid Z = z] \cdot \Pr[Z = z] = \mathbf{E}[Y].$$