#### Randomized Algorithms

#### Coupon Collector's Problem

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#### Review

- Expectation of discrete random variables
- Linearity of expectation.
- Bernoulli and Binomial random variable

#### Expectation

• The expectation of a discrete random variable X, denoted by  $\mathbf{E}[X]$ , is

$$\mathbf{E}[X] = \sum_{i} i \cdot \Pr[X = i]$$

• Example: Let *X* denote the sum of of dices:

$$\mathbf{E}[X] = \frac{1}{36} \cdot 2 + \frac{2}{36} \cdot 3 + \frac{3}{36} \cdot 4 + \dots + \frac{1}{36} \cdot 12 = 7.$$



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#### Linearity of Expectation

• For any finite collection of discrete random variables  $X_1, X_2, ..., X_n$  with finite expectations,

$$\mathbf{E}\left[\sum_{i=1}^{n} X_i\right] = \sum_{i=1}^{n} \mathbf{E}[X_i].$$

• For any constant c and discrete random variable *X*,

$$\mathbf{E}[cX] = c \cdot \mathbf{E}[X].$$

Why is it useful?

#### Example

- Consider the dice-throwing example again.
  - $X_1$ : the outcome of die 1
  - $X_2$ : the outcome of die 2

$$\mathbf{E}[X_1] = \mathbf{E}[X_2] = \frac{1}{6} \cdot \sum_{j=1}^{6} j = \frac{7}{2}$$

$$\mathbf{E}[X] = \mathbf{E}[X_1 + X_2] = 7.$$



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#### Bernoulli random variable

• Suppose we run an experiment that succeeds with probability p and fails with probability 1-p.

$$Y = \begin{cases} 1 & \text{if the experiment succeeds,} \\ 0 & \text{otherwise.} \end{cases}$$

- *Y*: Bernoulli random variable.
  - or indicator random variable.

$$\mathbf{E}[Y] = 1 \cdot \Pr[Y = 1] + 0 \cdot \Pr[Y = 0] = \Pr[Y = 1] = p.$$



#### Binomial random variable

• A binomial random variable X with parameters n and p, denoted by B(n, p), is defined as

$$\Pr[X = j] = \binom{n}{j} p^j (1-p)^{n-j}.$$

for j = 0, 1, 2, ..., n.

• Exercise: Show that  $\sum_{j=0}^{n} \Pr[X=j] = 1$ .

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X: the number of successful trials in the n experiments.

• Exercise: Show that  $\sum_{j=0}^{n} \Pr[X=j] = 1$ .

#### Binomial random variable (expectation)

$$\bullet \quad \mathbf{E}[X] = \sum_{j=0}^{n} j \binom{n}{j} p^{j} (1-p)^{n-j} \\
= \sum_{j=0}^{n} j \frac{n!}{j!(n-j)!} p^{j} (1-p)^{n-j} \\
= \sum_{j=1}^{n} \frac{n!}{(j-1)!(n-j)!} p^{j} (1-p)^{n-j} \\
= np \sum_{j=1}^{n} \frac{(n-1)!}{(j-1)!((n-1)-(j-1))!} p^{j-1} (1-p)^{(n-1)-(j-1)} \\
= np \sum_{k=0}^{n-1} \frac{(n-1)!}{k!((n-1)-k)!} p^{k} (1-p)^{(n-1)-k} \\
= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^{k} (1-p)^{(n-1)-k} \\
= np.$$

## Binomial random variable (expectation)

$$\begin{split} \bullet \quad & \mathbf{E}[X] \quad = \quad \sum_{j=0}^{n} j \binom{n}{j} p^{j} (1-p)^{n-j} \\ & = \quad \sum_{j=0}^{n} j \frac{n!}{j!(n-j)!} p^{j} (1-p)^{n-j} \\ & = \quad \sum_{j=1}^{n} \frac{n!}{(j-1)!(n-j)!} p^{j} (1-p)^{n-j} \\ & = \quad np \sum_{j=1}^{n} \frac{(n-1)!}{(j-1)!((n-1)-(j-1))!} p^{j-1} (1-p)^{(n-1)-(j-1)} \\ & = \quad np \sum_{k=0}^{n-1} \frac{(n-1)!}{k!((n-1)-k)!} p^{k} (1-p)^{(n-1)-k} \\ & = \quad np \sum_{k=0}^{n-1} \binom{n-1}{k} p^{k} (1-p)^{(n-1)-k} \\ & = \quad np. \end{split}$$

## Let's make it simpler!

- Denote a set of *n* Bernoulli random variables  $X_1, X_2, ..., X_n$ .
  - $X_i = 1$  if the *i*th trial is successful and 0 otherwise.
  - $X = X_1 + X_2 + \dots + X_n$

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$$\mathbf{E}[X] = \mathbf{E}\left[\sum_{i=1}^{n} X_i\right] = \sum_{i=1}^{n} \mathbf{E}[X_i] = np.$$

#### Geometric Distribution

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  - What's the distribution of the number of flips?

• <u>Definition</u>. A geometric random variable *X* with parameter *p* is

$$\Pr[X = n] = (1 - p)^{n - 1} p.$$

for 
$$n = 1, 2, ...$$

• Exercise. Show that  $\sum_{n\geq 1} \Pr[X=n] = 1$ .

#### Memoryless

- Let *X* be a geometric random variable *X* with parameter p > 0.
- For any n, k > 0,  $\Pr[X = n + k \mid X > k] = \Pr[X = n]$ .

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- Let *X* be a geometric random variable *X* with parameter p > 0.
- For any n, k > 0,  $\Pr[X = n + k \mid X > k] = \Pr[X = n]$ .

• Proof. 
$$\Pr[X = n + k \mid X > k] = \frac{\Pr[(X = n + k) \cap (X > k)]}{\Pr[X > k]}$$

$$= \frac{\Pr[X = n + k]}{\Pr[X > k]}$$

$$= \frac{(1 - p)^{n+k-1}p}{\sum_{i=k}^{\infty} (1 - p)^{i}p}$$

$$= \frac{(1 - p)^{n+k-1}p}{(1 - p)^{k}}$$

$$= (1 - p)^{n-1}p = \Pr[X = n].$$

## The mean of a geometric r.v. X(p)

$$\mathbf{E}[X] = \sum_{j=1}^{\infty} j \Pr[X = j] \qquad \mathbf{E}[X] = \sum_{i=1}^{\infty} (1-p)^{i-1} = \frac{1}{1-(1-p)} = \frac{1}{p}.$$

$$= \sum_{j=1}^{\infty} \sum_{i=1}^{j} \Pr[X = j]$$

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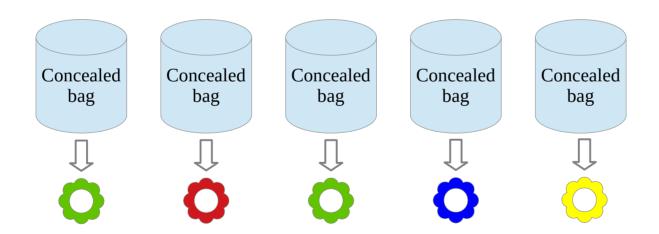
$$= \sum_{i=1}^{\infty} \Pr[X \ge i].$$

$$\Pr[X \ge i] = \sum_{k=i}^{\infty} (1-p)^{k-1} p = (1-p)^{i-1}$$



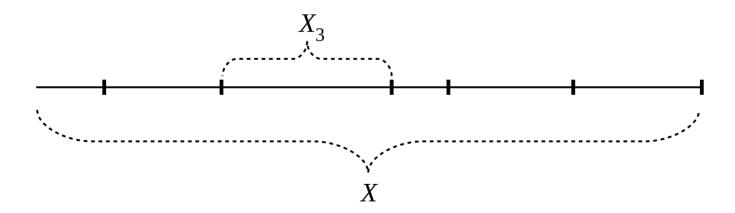
▲一名網友在臉書社團「爆廢公社公開版」表示,多年前為了收集連鎖便利商店 7-11 的一款贈品,花了不少金額,還拿到許多重複的款式,貼文引發 3 千多名網友共鳴。(圖/翻攝自爆廢公社公開版)

- Have you already got all of them (totally n types)?
- Have you ever thought about how much you should pay for them?



• Each bag is chosen independently and uniformly at random from the *n* possibilities.

- Let *X* be the number of bags bought until every type of coupon is obtained.
- Let  $X_i$  be the number of bags bought while you had already got exactly i-1 different coupons.



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- Let *X* be the number of bags bought until every type of coupon is obtained.
- Let  $X_i$  be the number of bags bought while you had already got exactly i-1 different coupons.
  - Geometric random variables?!
  - What about  $X = \sum_{i=1}^{n} X_i$ ?

• When exactly *i*−1 coupons have been collected, the probability of obtaining a new one is

$$p_i = 1 - \frac{i-1}{n}$$

•  $X_i$  is a geometric random variable, so

$$\mathbf{E}[X_i] = \frac{1}{p_i} = \frac{n}{n-i+1}.$$

## Coupon Collector's Problem (contd.)

• 
$$\mathbf{E}[X] = \mathbf{E}\left[\sum_{i=1}^{n} X_{i}\right]$$

$$= \sum_{i=1}^{n} \mathbf{E}[X_{i}]$$

$$= \sum_{i=1}^{n} \frac{n}{n-i+1}$$

$$= \sum_{i=1}^{n} \frac{1}{i}.$$

$$\sum_{k=1}^{n} \frac{1}{k} \ge \int_{x=1}^{n} \frac{1}{x} dx = \ln n$$

$$\sum_{k=2}^{n} \frac{1}{k} \le \int_{x=1}^{n} \frac{1}{x} dx = \ln n$$

## Coupon Collector's Problem (contd.)

• So, you are about to buy  $n \ln n + \Theta(n)$  bags for collecting all the coupons (stickers)!

## On conditional expectation

• There are terminologies which may confusing you.

## Conditional Expectation

• Definition.

$$\mathbf{E}[Y \mid Z = z] = \sum_{y} y \Pr[Y = y \mid Z = z].$$

- Example of two dices.
  - $X_1$ : the number showing on the first die
  - $X_2$ : the number showing on the second die
  - $X = X_1 + X_2$  $\mathbf{E}[X \mid X_1 = 2] = \sum_{x=3}^{8} x \cdot \frac{1}{6} = \frac{11}{2}.$

#### Conditional Expectation (contd.)

• <u>Lemma</u>. For any random variables *X* and *Y*,

$$\mathbf{E}[X] = \sum_{y} \Pr[Y = y] \mathbf{E}[X \mid Y = y].$$

• Proof. 
$$\sum_{y} \Pr[Y = y] \cdot \mathbf{E}[X \mid Y = y] = \sum_{y} \Pr[Y = y] \cdot \sum_{x} x \Pr[X = x \mid Y = y]$$

$$= \sum_{x} \sum_{y} x \Pr[X = x \mid Y = y] \cdot \Pr[Y = y]$$

$$= \sum_{x} \sum_{y} x \Pr[X = x \cap Y = y]$$

$$= \sum_{x} x \Pr[X = x]$$

$$= \sum_{x} x \Pr[X = x]$$

## Conditional Expectation

• <u>Lemma</u>. For any finite collection of discrete random variables  $X_1$ ,  $X_2$ , ...,  $X_n$  with finite expectations and for any random variable Y,

$$\mathbf{E}\left[\sum_{i=1}^{n} X_i \mid Y=y\right] = \sum_{i=1}^{n} \mathbf{E}[X_i \mid Y=y].$$

## Conditional Expectation (contd.)

- A weird definition.
- $\mathbf{E}[Y \mid Z]$  : regarded as a **random variable** f(Z).
  - It takes on the value  $\mathbf{E}[Y \mid Z = z]$  when Z = z.
- In the previous example,

$$\mathbf{E}[X \mid X_1] = \sum_{x} x \cdot \Pr[X = x \mid X_1] = \sum_{X_1 + 1}^{X_1 + 6} x \cdot \frac{1}{6} = X_1 + \frac{7}{2}.$$

So it makes sense that

$$\mathbf{E}[\mathbf{E}[X \mid X_1]] = \mathbf{E}\left[X_1 + \frac{7}{2}\right] = \frac{7}{2} + \frac{7}{2} = 7.$$

## Conditional Expectation (contd.)

• Theorem.  $\mathbf{E}[Y] = \mathbf{E}[\mathbf{E}[Y \mid Z]].$ 

• Proof.

$$\mathbf{E}[\mathbf{E}[Y \mid Z]] = \sum_{z} \mathbf{E}[Y \mid Z = z] \cdot \Pr[Z = z] = \mathbf{E}[Y].$$

# Assignment 02

1. Let  $X \sim$  binomial distribution B(n, p) be a random variable. Show that

$$\sum_{j=0}^{n} \Pr[X=j] = 1.$$

2. Let  $X \sim$  geometric distribution G(p) be a random variable. Show that

$$\sum_{n>1} \Pr[X=n] = 1.$$

3. Give a practical example of coupon collector's problem and provide your analysis.