

# Mathematics for Machine Learning

## — Expectation Maximization

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# Credits for the resource

- The slides are based on the textbooks:
  - *Marc Peter Deisenroth, A. Aldo Faisal, and Cheng Soon Ong: Mathematics for Machine Learning. Cambridge University Press. 2020.*
  - *Arnold J. Insel, Lawrence E. Spence, Stephen H. Friedberg: Linear Algebra, 4th Edition. Prentice Hall. 2013.*
  - *Howard Anton, Chris Rorres, Anton Kaul: Elementary Linear Algebra, 12th Edition. Wiley. 2019.*
- We could partially refer to the monograph:  
*Francesco Orabona: A Modern Introduction to Online Learning.*  
<https://arxiv.org/abs/1912.13213>

# Outline

- 1 Expectation Maximization (EM) Algorithm
- 2 Latent-Variable Perspective

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  - ∵ the complex dependency on the parameters.
- The likelihood approach suggests a simple iterative scheme for finding a solution to the parameters estimation problem.

# Expectation Maximization

- A general iterative scheme for learning parameters (MLE or MAP) in mixture models and latent-variable models.

Dempster et al. (1977)

Choose initial parameter values (i.e.,  $\mu_k, \Sigma_k, \pi_k$ ) and **alternate** between the following two steps until convergence:

- **E-step:** Evaluate the responsibilities  $r_{ik}$ 
  - It can be viewed as the **posterior probability** of data point  $i$  belonging to mixture component  $k$ .
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- Intuitive idea: the log-likelihood is increased after each step.

## EM algorithm for Estimating parameters of a GMM

- ① Initialize  $\mu_k, \Sigma_k, \pi_k$ .
- ② **E-step:** Evaluate  $r_{ik}$  for every data point  $\mathbf{x}_i$  using the current parameters:

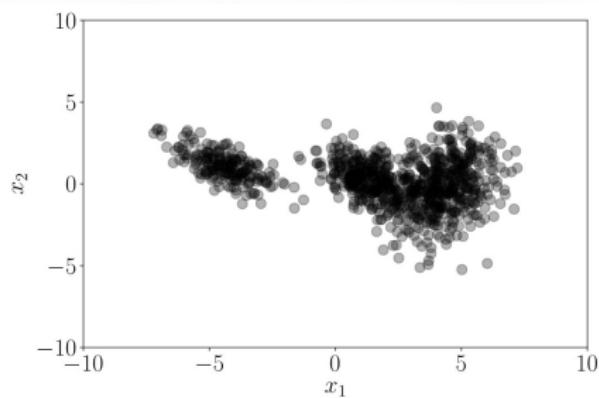
$$r_{ik} = \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \mu_k, \Sigma_k)}{\sum_j \pi_j \mathcal{N}(\mathbf{x}_i | \mu_j, \Sigma_j)}$$

- ③ **M-step:** Re-estimate parameters  $\mu_k, \Sigma_k, \pi_k$  using the current responsibilities  $r_{ik}$  from the E-step:

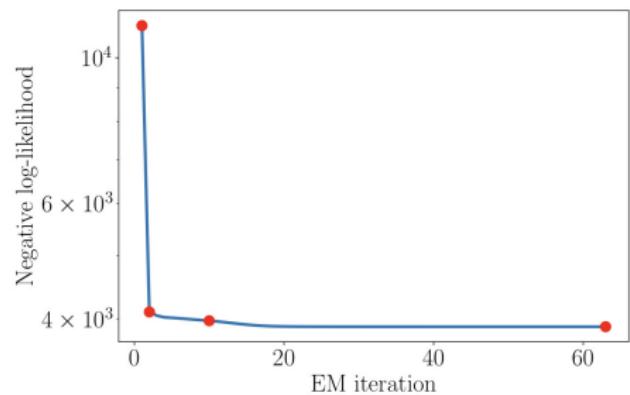
$$\mu_k = \frac{1}{N_k} \sum_{i=1}^N r_{ik} \mathbf{x}_i,$$

$$\Sigma_k = \frac{1}{N_k} \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \mu_k)(\mathbf{x}_i - \mu_k)^\top,$$

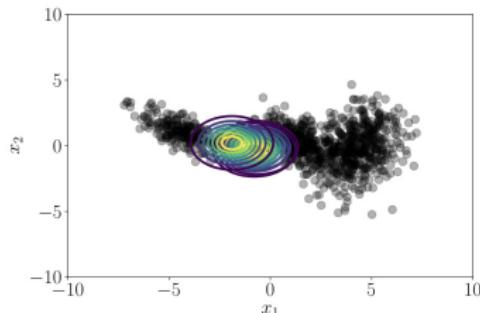
$$\pi_k = \frac{N_k}{N}.$$



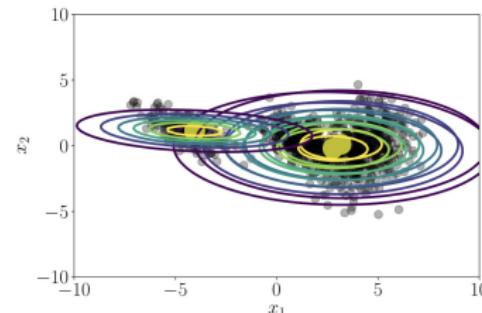
(a) Dataset.



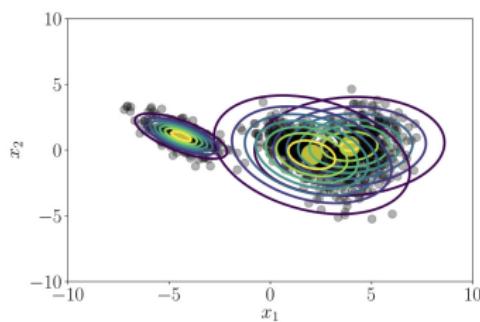
(b) Negative log-likelihood.



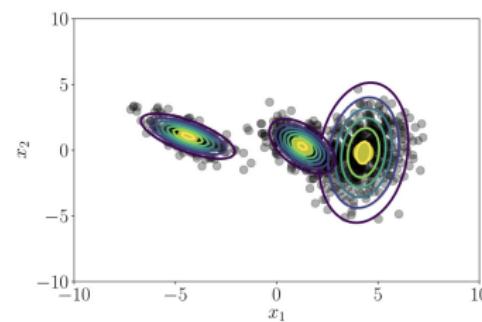
(c) EM initialization.



(d) EM after one iteration.



(e) EM after 10 iterations.



(f) EM after 62 iterations.

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# Latent-Variable Perspective

- View the GMM from the perspective of a **discrete latent variable** model.
- The latent variable  $z$  can attain only a **finite** set of values.

# A View of Generative Process

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- Define  $\mathbf{z} := [z_1, z_2, \dots, z_K]^\top \in \mathbb{R}^K$  as a vector consisting of **exactly one 1 and  $K - 1$  many 0s**.
  - One-hot encoding.
  - $\mathbf{z} = [z_1, z_2, z_3]^\top = [0, 1, 0]^\top \Rightarrow$  the 2nd mixture component is selected.

## Prior on the latent variable

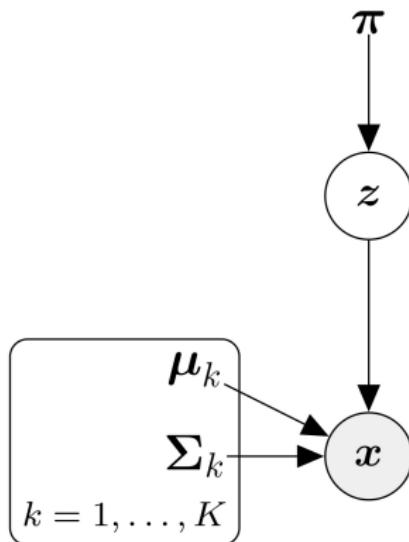
- When the variables  $z_k$  are unknown, we can place a prior distribution on  $\mathbf{z}$  in practice:

$$p(\mathbf{z}) = \boldsymbol{\pi} = [\pi_1, \pi_2, \dots, \pi_K]^\top, \sum_{k=1}^K \pi_k = 1,$$

where the  $k$ th entry  $\pi_k = p(z_k = 1)$  describes the prob. that the  $k$ th mixture component generated data point  $\mathbf{x}$ .

# Sampling from a GMM

Ancestral sampling.



## A Simple Sampling Procedure

- ① Sample  $z^{(i)} \sim p(\mathbf{z})$ .
- ② Sample  $\mathbf{x}^{(i)} \sim p(\mathbf{x} \mid z^{(i)} = 1)$ .

# Sampling from a GMM

The joint distribution

$$p(\mathbf{x}, z_k = 1) = p(\mathbf{x} | z_k = 1)p(z_k = 1) = \pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k),$$

for  $k = 1, \dots, K$ . So, we have

$$p(\mathbf{x}, \mathbf{z}) = \begin{bmatrix} p(\mathbf{x}, z_1 = 1) \\ p(\mathbf{x}, z_2 = 1) \\ \vdots \\ p(\mathbf{x}, z_K = 1) \end{bmatrix} = \begin{bmatrix} \pi_1 \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1) \\ \pi_2 \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2) \\ \vdots \\ \pi_K \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K), \end{bmatrix}$$

which fully specifies the probabilistic model.

# Likelihood $p(\mathbf{x} | \boldsymbol{\theta})$ in a latent-variable model

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- Summing out (marginalizing out) all latent variables from  $p(\mathbf{x}, \mathbf{z})$ :

$$p(\mathbf{x} | \boldsymbol{\theta}) = \sum_{\mathbf{z}} p(\mathbf{x} | \boldsymbol{\theta}, \mathbf{z}) p(\mathbf{z} | \boldsymbol{\theta})$$

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$$\boldsymbol{\theta} := \{\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \pi_k : k = 1, 2, \dots, K\}.$$

- There is only one single nonzero entry in each  $\mathbf{z}$ , so there are **only  $K$  possible configurations** of  $\mathbf{z}$ .

So, the desired marginal distribution is

$$p(\mathbf{x} \mid \boldsymbol{\theta}) = \sum_{k=1}^K p(\mathbf{x} \mid \boldsymbol{\theta}, z_k = 1) p(z_k = 1 \mid \boldsymbol{\theta}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

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The latent variable model with latent indicators  $z_k$  is an equivalent way of thinking about a Gaussian mixture model.

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★ The responsibility of the  $k$ th mixture component for  $\mathbf{x}$ !

## Extending to a Full Dataset (1/2)

- Consider a dataset of  $N$  data points  $\mathcal{X} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$ .
- Assume that every data point  $\mathbf{x}_i$  possesses its own latent variable

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## Extending to a Full Dataset (2/2)

Consider the posterior distribution  $p(z_{ik} = 1 | \mathbf{x}_i)$  by applying Bayes' theorem:

$$\begin{aligned} p(z_{ik} = 1 | \mathbf{x}_i) &= \frac{p(\mathbf{x}_i | z_{ik} = 1)p(z_{ik} = 1)}{\sum_{j=1}^K p(\mathbf{x}_i | z_{ij} = 1)p(z_{ij} = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \end{aligned}$$

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- Now, we see that the responsibilities have a mathematically justified interpretation as posterior probabilities.

# Discussions