

# Mathematics for Machine Learning

## — Gaussian Mixture Models

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## Credits for the resource

- The slides are based on the textbooks:
  - *Marc Peter Deisenroth, A. Aldo Faisal, and Cheng Soon Ong: Mathematics for Machine Learning. Cambridge University Press. 2020.*
  - *Howard Anton, Chris Rorres, Anton Kaul: Elementary Linear Algebra. Wiley. 2019.*
- We could partially refer to the monograph:  
*Francesco Orabona: A Modern Introduction to Online Learning.*  
<https://arxiv.org/abs/1912.13213>

# Outline

- 1 Introduction & Gaussian Mixture Model (GMM)
- 2 Parameter Learning via Maximum Likelihood
  - Updating the Means
  - Updating the Covariances
  - Updating the Mixture Weights

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## 2 Parameter Learning via Maximum Likelihood

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# Introduction

## Focus

- **Goal:** Density Estimation.
- Covering two important concepts:
  - Expectation maximization (EM).
  - Latent variable perspective.

# Motivation

- A straightforward way to represent data: Let them present themselves directly.
- **Issue:** The data might be *dirty* or too huge to show all of them.

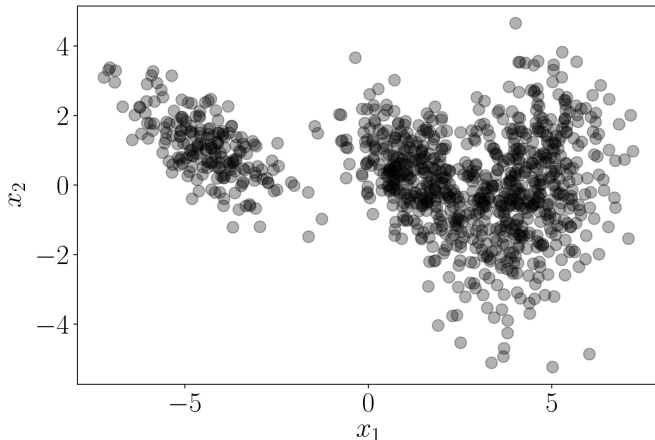
# Motivation

- A straightforward way to represent data: Let them present themselves directly.
- **Issue:** The data might be *dirty* or too huge to show all of them.

We want to represent the data compactly using a density from a parametric family, such as Gaussian or Beta distribution.

- Mean & variance.

# One Gaussian representation might not be meaningful.





# A Solution

- Consider **mixture models**:
  - A convex combination of  $K$  simple base distributions.
  - A distribution  $p(\mathbf{x})$ :

$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k p_k(\mathbf{x}),$$

$$0 \leq \pi_k \leq 1, \sum_{k=1}^K \pi_k = 1.$$

- $\pi_k$ : *mixture weights*.
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- $\pi_k$ : *mixture weights*.
- More expressive than a base distribution.
- **Gaussian mixture models (GMMs)**: the base distributions are Gaussians.

# Gaussian Mixture Model

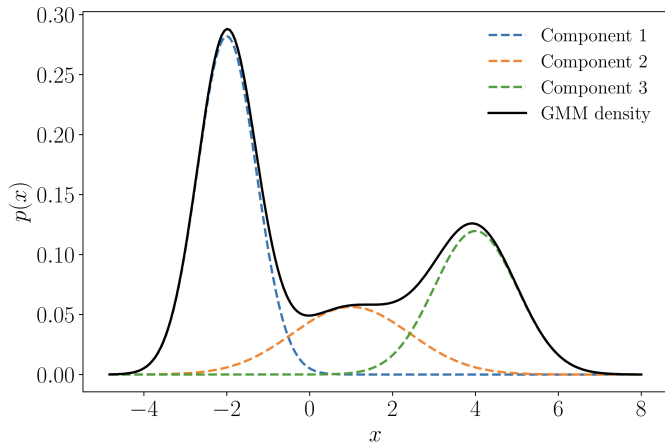
## Gaussian Mixture Model

A Gaussian mixture model is a density model where we combine a finite number of  $K$  Gaussian distributions  $\mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$  such that

$$p(\mathbf{x} \mid \boldsymbol{\theta}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$
$$0 \leq \pi_k \leq 1, \sum_{k=1}^K \pi_k = 1,$$

where  $\boldsymbol{\theta} := \{\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k, \pi_k \mid k = 1, \dots, K\}$ .

## GMMs



$$p(x | \theta) = 0.5\mathcal{N}(x | -2, 0.5) + 0.2\mathcal{N}(x | 1, 2) + 0.3\mathcal{N}(x | 4, 1).$$

# Outline

## 1 Introduction & Gaussian Mixture Model (GMM)

## 2 Parameter Learning via Maximum Likelihood

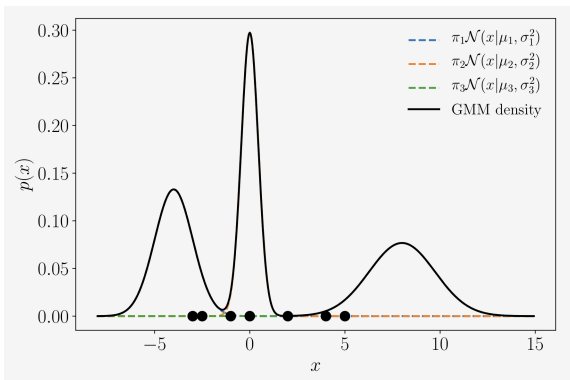
- Updating the Means
- Updating the Covariances
- Updating the Mixture Weights

# The Setting

- A dataset  $\mathcal{X} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$ , where each  $\mathbf{x}_i$  is drawn i.i.d. from an unknown distribution  $p(\mathbf{x})$ .
- Parameters:  $\theta := \{\mu_k, \Sigma_k, \pi_k \mid k = 1, \dots, K\}$ .

## Example of an Initial Setting

- $\mathcal{X} = \{-3, -2.5, -1, 0, 2, 4, 5\}$ .
- $K = 3$ .
- $p_1(x) = \mathcal{N}(x \mid -4, 1)$ ,  $p_2(x) = \mathcal{N}(x \mid 0, 0.2)$ ,  $p_3(x) = \mathcal{N}(x \mid 8, 3)$ .
- $\pi_1 = \pi_2 = \pi_3 = 1/3$ .



# The Likelihood

By the i.i.d. assumption, we have the factorized likelihood

$$p(\mathcal{X} | \theta) = \prod_{i=1}^N p(\mathbf{x}_i | \theta), \quad p(\mathbf{x}_i | \theta) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).$$

Then the log-likelihood is

$$\mathcal{L} := \log p(\mathcal{X} | \theta) = \sum_{i=1}^N \log p(\mathbf{x}_i | \theta) = \sum_{i=1}^N \log \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).$$



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- **Goal:** Find parameters  $\theta_{ML}^*$ .

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- We cannot obtain a closed-form solution here (except for  $K = 1$ , i.e., single Gaussian).
- We exploit an iterative scheme to find  $\theta_{ML}^*$ : the EM algorithm.
- **The key idea:** Update one model parameter at a time while keeping the others fixed.

Necessary conditions for a local optimum of  $\mathcal{L}$ :

$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}_k} = \mathbf{0}^\top &\iff \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} = \mathbf{0}^\top \\ \frac{\partial \mathcal{L}}{\partial \boldsymbol{\Sigma}_k} = \mathbf{0}^\top &\iff \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\Sigma}_k} = \mathbf{0}^\top \\ \frac{\partial \mathcal{L}}{\partial \pi_k} = 0 &\iff \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \pi_k} = 0.\end{aligned}$$

Applying the chain rule:

$$\frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\theta}}$$

and

$$\frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} = \frac{1}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}.$$

# Responsibilities: Facilitating our discussions

Responsibility of the  $k$ th mixture component for  $n$ th data point

$$r_{ik} := \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}.$$

- Note that

$$p(\mathbf{x}_i | \pi_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) = \pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

which is proportional to the likelihood.

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- High responsibility  $\implies$  The data point is plausible sample from that mixture component.



## Remark

$\mathbf{r}_i := [r_{i1}, \dots, r_{iK}]^\top \in \mathbb{R}^K$  is a normalized probability vector.

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- A *soft assignment* of  $\mathbf{x}_i$  to the  $K$  mixture component.

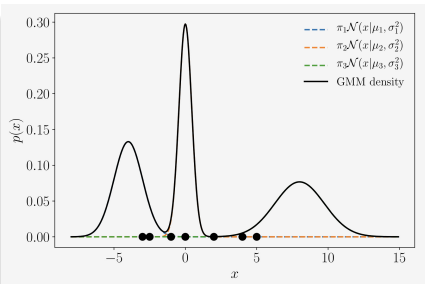
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- A *soft assignment* of  $\mathbf{x}_i$  to the  $K$  mixture component.
  - Similar idea: softmax functions.

## Example (responsibilities of the previous example)

$$\begin{bmatrix} 1.0 & 0.0 & 0.0 \\ 1.0 & 0.0 & 0.0 \\ 0.057 & 0.943 & 0.0 \\ 0.001 & 0.999 & 0.0 \\ 0.0 & 0.066 & 0.934 \\ 0.0 & 0.0 & 1.0 \\ 0.0 & 0.0 & 1.0 \end{bmatrix} \in \mathbb{R}^{N \times K}.$$



- Try to compute it by yourselves.

# Update of the GMM Means

## Theorem [Update of the Means]

The update of the mean parameters  $\mu_k$ ,  $k = 1, \dots, K$ , of the GMM is given by

$$\mu_k^{new} = \frac{\sum_{i=1}^N r_{ik} \mathbf{x}_i}{\sum_{i=1}^N r_{ik}}.$$

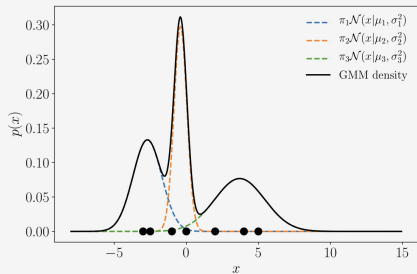
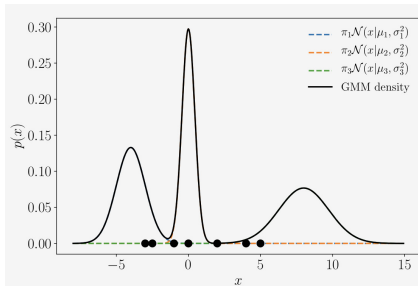
$$\begin{aligned}
 \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} &= \sum_{j=1}^K \pi_j \frac{\partial \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}{\partial \boldsymbol{\mu}_k} = \pi_k \frac{\partial \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\partial \boldsymbol{\mu}_k} \\
 &= \pi_k (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \boldsymbol{\Sigma}_k^{-1} \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).
 \end{aligned}$$

$$\begin{aligned}
 \frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}_k} &= \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} = \sum_{i=1}^N \frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\mu}_k} \\
 &= \sum_{i=1}^N (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \boldsymbol{\Sigma}_k^{-1} \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \\
 &= \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \boldsymbol{\Sigma}_k^{-1}.
 \end{aligned}$$

Solving  $\frac{\partial \mathcal{L}(\boldsymbol{\mu}_k^{new})}{\partial \boldsymbol{\mu}_k} = \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \boldsymbol{\mu}_k^{new})^\top \boldsymbol{\Sigma}_k^{-1} = \mathbf{0}^\top$ :

$$\begin{aligned} \sum_{i=1}^N r_{ik} \mathbf{x}_i &= \sum_{i=1}^N r_{ik} \boldsymbol{\mu}_k^{new} \\ \iff \boldsymbol{\mu}_k^{new} &= \frac{\sum_{i=1}^N r_{ik} \mathbf{x}_i}{\sum_{i=1}^N r_{ik}} = \frac{1}{N_k} \sum_{i=1}^N r_{ik} \mathbf{x}_i, \end{aligned}$$

where  $N_k := \sum_{i=1}^N r_{ik}$ .



- $\mu_1 : -4 \rightarrow -2.7$ .
- $\mu_2 : 0 \rightarrow -0.4$ .
- $\mu_3 : 8 \rightarrow 3.7$ .



## Remark

- $r_{ik}$  is a function of  $\pi_j, \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j$  for all  $j = 1, \dots, K$ .
- Hence the updates depend on all parameters of the GMM.

# Update of the GMM Covariances

## Theorem [Update of the Covariances]

The update of the covariance parameters  $\Sigma_k$ ,  $k = 1, \dots, K$ , of the GMM is given by

$$\Sigma_k^{new} = \frac{1}{N_k} \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \boldsymbol{\mu}_k)(\mathbf{x}_i - \boldsymbol{\mu}_k)^\top,$$

where

$$r_{ik} := \frac{\pi_k \mathcal{N}(\mathbf{x}_i \mid \boldsymbol{\mu}_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i \mid \boldsymbol{\mu}_j, \Sigma_j)}.$$

and  $N_k := \sum_{i=1}^N r_{ik}$ .

$$\frac{\partial \mathcal{L}}{\partial \Sigma_k} = \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \Sigma_k} = \sum_{i=1}^N \frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \Sigma_k}$$

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$$\begin{aligned} \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \Sigma_k} &= \frac{\partial}{\partial \Sigma_k} \left( \pi_k (2\pi)^{-\frac{D}{2}} \det(\Sigma_k)^{-\frac{1}{2}} \exp \left( -\frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \Sigma_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) \right) \right) \\ &= \pi_k (2\pi)^{-\frac{D}{2}} \left[ \frac{\partial}{\partial \Sigma_k} \det(\Sigma_k)^{-\frac{1}{2}} \exp \left( -\frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \Sigma_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) \right) \right. \\ &\quad \left. + \det(\Sigma_k)^{-\frac{1}{2}} \frac{\partial}{\partial \Sigma_k} \exp \left( -\frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \Sigma_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) \right) \right] \end{aligned}$$

Note that

$$\begin{aligned} \frac{\partial}{\partial \Sigma_k} \det(\Sigma_k)^{-\frac{1}{2}} &= -\frac{1}{2} \det(\Sigma_k)^{-\frac{1}{2}} \Sigma_k^{-1}, \\ \frac{\partial}{\partial \Sigma_k} (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \Sigma_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) &= -\Sigma_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \Sigma_k^{-1} \end{aligned}$$

$$\frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\Sigma}_k} = \pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \cdot \left[ -\frac{1}{2} (\boldsymbol{\Sigma}_k^{-1} - \boldsymbol{\Sigma}_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \boldsymbol{\Sigma}_k^{-1}) \right]$$

Thus,

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial \boldsymbol{\Sigma}_k} &= \sum_{i=1}^N \frac{\partial \log p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\Sigma}_k} = \sum_{i=1}^N \frac{1}{p(\mathbf{x}_i | \boldsymbol{\theta})} \frac{\partial p(\mathbf{x}_i | \boldsymbol{\theta})}{\partial \boldsymbol{\Sigma}_k} \\ &= \sum_{i=1}^N \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} \\ &\quad \cdot \left[ -\frac{1}{2} (\boldsymbol{\Sigma}_k^{-1} - \boldsymbol{\Sigma}_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \boldsymbol{\Sigma}_k^{-1}) \right] \\ &= -\frac{1}{2} \sum_{i=1}^N r_{ik} (\boldsymbol{\Sigma}_k^{-1} - \boldsymbol{\Sigma}_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \boldsymbol{\Sigma}_k^{-1}) \\ &= -\frac{1}{2} \left( \sum_{i=1}^N r_{ik} \right) \boldsymbol{\Sigma}_k^{-1} + \frac{1}{2} \boldsymbol{\Sigma}_k^{-1} \left( \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \boldsymbol{\mu}_k) (\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \right) \boldsymbol{\Sigma}_k^{-1}. \end{aligned}$$

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Setting  $\frac{\partial \mathcal{L}}{\partial \Sigma_k} = \mathbf{0}^\top$ , we have

$$N_k \Sigma_k^{-1} = \Sigma_k^{-1} \left( \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \boldsymbol{\mu}_k)(\mathbf{x}_i - \boldsymbol{\mu}_k)^\top \right) \Sigma_k^{-1}$$

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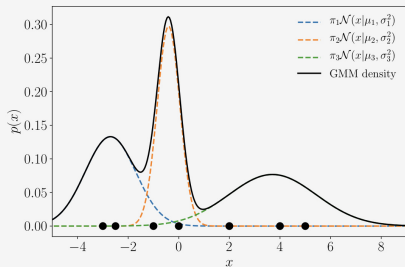
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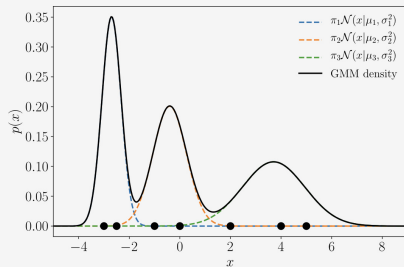
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Hence,

$$\Sigma_k^{\text{new}} = \frac{1}{N_k} \sum_{i=1}^N r_{ik} (\mathbf{x}_i - \boldsymbol{\mu}_k)(\mathbf{x}_i - \boldsymbol{\mu}_k)^\top.$$



(a) GMM density and individual components prior to updating the variances.



(b) GMM density and individual components after updating the variances.

- $\sigma_1^2 : 1 \rightarrow 1.44$ .
- $\sigma_2^2 : 0.2 \rightarrow 0.44$ .
- $\sigma_3^2 : 3 \rightarrow 1.53$ .

# Update of the GMM Mixture Weights

## Theorem [Update of the Mixture Weights]

The update of the mixture weights of the GMM is given by

$$\pi_k^{\text{new}} = \frac{N_k}{N}, \quad k = 1, \dots, K.$$

- $N$ : the number of data points.
- $N_k := \sum_{i=1}^N r_{ik}$ .

- We account for the constraint  $\sum_k \pi_k = 1$ .
  - Using Lagrange multipliers.

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- The Lagrangian:

$$\begin{aligned}\mathfrak{L} &= \mathcal{L} + \lambda \left( \sum_{k=1}^K \pi_k - 1 \right) \\ &= \sum_{i=1}^N \log \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_i \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) + \lambda \left( \sum_{k=1}^K \pi_k - 1 \right).\end{aligned}$$

Obtain the partial derivative of  $\mathcal{L}$  w.r.t.  $\pi_k$ :

$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial \pi_k} &= \sum_{i=1}^N \frac{\mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} + \lambda \\&= \frac{1}{\pi_k} \sum_{i=1}^N \frac{\pi_k \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} + \lambda \\&= \frac{N_k}{\pi_k} + \lambda,\end{aligned}$$

and the partial derivative w.r.t.  $\lambda$  is

$$\frac{\partial \mathcal{L}}{\partial \lambda} = \sum_{k=1}^K \pi_k - 1.$$

Now we have

$$\frac{\partial \mathcal{L}}{\partial \pi_k} = \frac{N_k}{\pi_k} + \lambda$$

$$\frac{\partial \mathcal{L}}{\partial \lambda} = \sum_{k=1}^K \pi_k - 1$$

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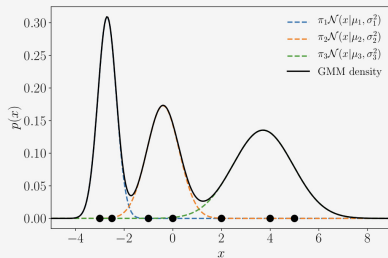
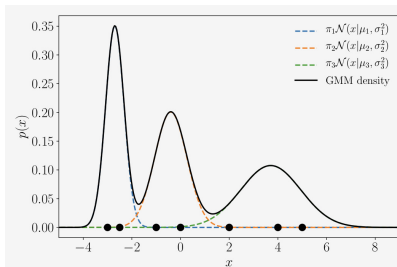
$$\frac{\partial \mathcal{L}}{\partial \pi_k} = \frac{N_k}{\pi_k} + \lambda$$
$$\frac{\partial \mathcal{L}}{\partial \lambda} = \sum_{k=1}^K \pi_k - 1$$

Setting both to  $\mathbf{0}^\top$  we have

$$\pi_k = -\frac{N_k}{\lambda}$$
$$1 = \sum_{k=1}^K \pi_k = -\sum_{k=1}^K \frac{N_k}{\lambda} = -\frac{N}{\lambda}$$

So  $\lambda = -N \implies \pi_k^{new} = \frac{N_k}{N}$ .





- $\pi_1 : \frac{1}{3} \rightarrow 0.29$ .
- $\pi_2 : \frac{1}{3} \rightarrow 0.29$ .
- $\pi_3 : \frac{1}{3} \rightarrow 0.42$ .

# Discussions