

Mathematics for Machine Learning

— Linear Regression

Problem Formulation & Parameter Estimation

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Credits for the resource

- The slides are based on the textbooks:
 - *Marc Peter Deisenroth, A. Aldo Faisal, and Cheng Soon Ong: Mathematics for Machine Learning. Cambridge University Press. 2020.*
 - *Arnold J. Insel, Lawrence E. Spence, Stephen H. Friedberg: Linear Algebra, 4th Edition. Prentice Hall. 2013.*
 - *Howard Anton, Chris Rorres, Anton Kaul: Elementary Linear Algebra, 12th Edition. Wiley. 2019.*
- We could partially refer to the monograph:
Francesco Orabona: A Modern Introduction to Online Learning.
<https://arxiv.org/abs/1912.13213>

Outline

- 1 Introduction
- 2 Problem Formulation
- 3 Parameter Estimation
 - Maximum Likelihood Estimation (MLE)
 - Overfitting in Linear Regression
 - Maximum A Posteriori Estimation (MAP)
 - MAP Estimation as Regularization
- 4 Bayesian Linear Regression

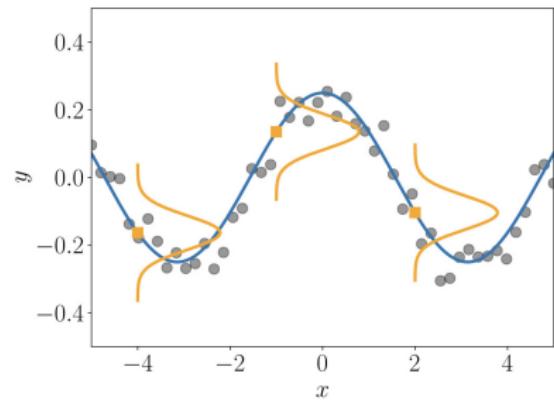
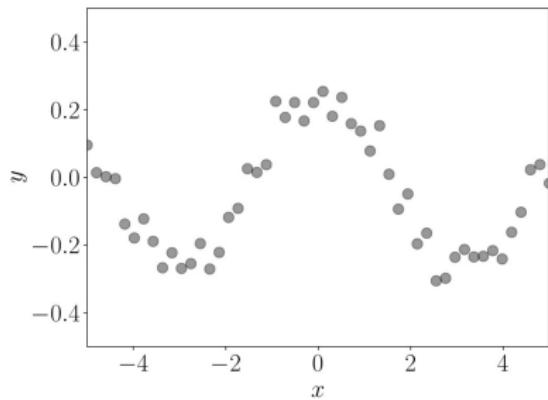
Linear Regression

Aim

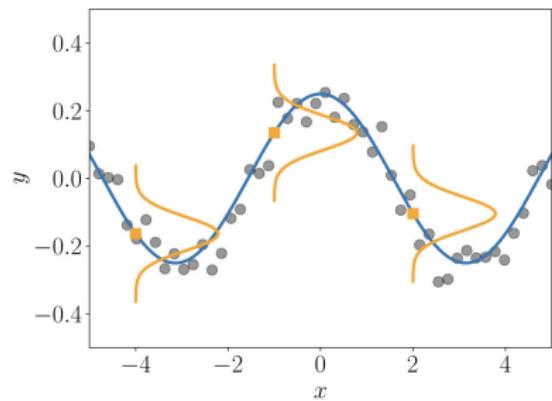
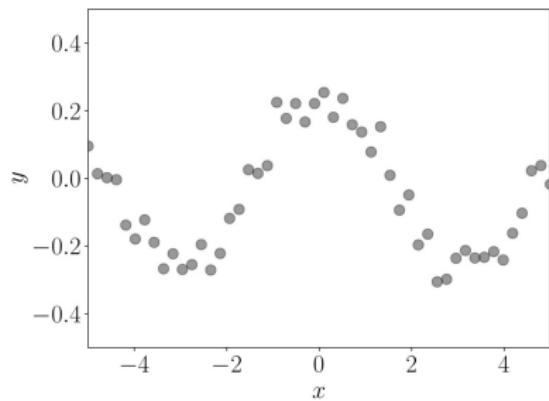
Find (or Infer) a function $f : \mathbb{R}^D \rightarrow \mathbb{R}$ which maps input $\mathbf{x} \in \mathbb{R}^D$ to the corresponding function values $f(\mathbf{x}) \in \mathbb{R}$.

- And we hope f to generalize well to unseen input.
- Training input: $\{\mathbf{x}_i\}_{i=1}^N$
- Assume the noisy observations $\{y_i\}_{i=1}^N$ for $y_i = f(\mathbf{x}_i) + \epsilon$, an i.i.d. random variable ϵ .
 - Consider zero-mean Gaussian noise throughout our discussions.

- Observe (noisy) function values $y_n = f(x_n) + \epsilon$.



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Applications of regression:

- Time series analysis, reinforcement learning, optimization, computer games, classification algorithms, etc.

Problems Involved in Regression

- Choice of the model and the parametrization.
 - Function classes, particular parametrization (e.g., degree of the polynomial)
- Finding good parameters.
 - Loss minimization w.r.t. different loss functions.
- Overfitting and model selection.
- Relationship b/w loss functions and parameter priors.
 - Probabilistic models.
- Uncertainty modeling.
 - We have limited amount of data.
 - The smaller the training set, the more important uncertainty modeling.
 - Equip model predictions with confidence bounds.

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Problem Formulation

- Because of observing noise, we adopt a probabilistic approach to explicitly model the noise using a **likelihood function**.
- Focus:** a regression problem with the likelihood function:

$$p(y | \mathbf{x}) = \mathcal{N}(y | f(\mathbf{x}), \sigma^2).$$

- $\mathbf{x} \in \mathbb{R}^D$: inputs.
- $y \in \mathbb{R}$: noisy function values (targets).
- The relationship between \mathbf{x} and y :

$$y = f(\mathbf{x}) + \epsilon,$$

for $\epsilon \sim \mathcal{N}(0, \sigma^2)$.

An Example of Linear Regression

- An example of linear regression:

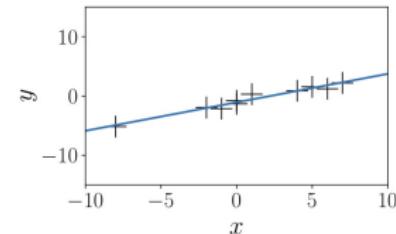
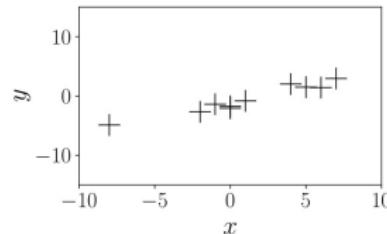
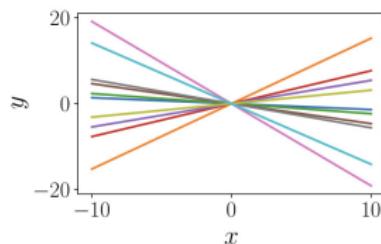
$$p(y \mid \mathbf{x}, \boldsymbol{\theta}) = \mathcal{N}(y \mid \mathbf{x}^T \boldsymbol{\theta}, \sigma^2).$$



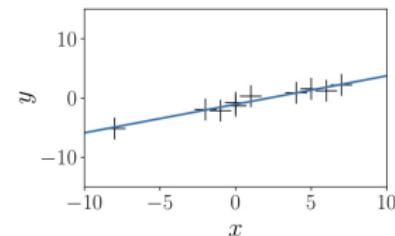
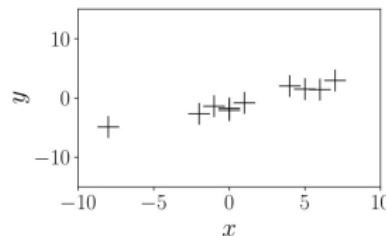
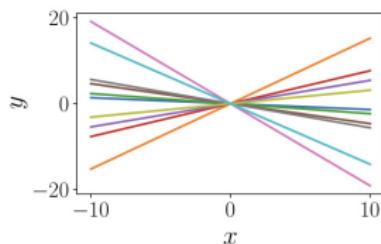
$$y = \mathbf{x}^T \boldsymbol{\theta} + \epsilon,$$

for $\epsilon \sim \mathcal{N}(0, \sigma^2)$.

- $\boldsymbol{\theta} \in \mathbb{R}^D$: the parameters we seek.
- ϵ : the only source of uncertainty.



- “Linear”: linear in the parameters.
 - Parameters: describing a function by a linear combination of input features.



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 - Parameters: describing a function by a linear combination of input features.
- Hence, $y = \phi^\top(\mathbf{x})\theta$ is also regarded as a linear regression (ϕ can be nonlinear).
 - A “feature” here is a representation $\phi(\mathbf{x})$ of the input \mathbf{x} .

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The Likelihood

- Given a training set $\mathcal{D} := \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_N, y_N)\}$, $\mathbf{x}_i \in \mathbb{R}^D$ and $y_i \in \mathbb{R}$ for $i = 1, \dots, N$.
- By the independence of the input, the likelihood factorizes:

$$p(\mathcal{Y} \mid \mathcal{X}, \boldsymbol{\theta}) = p(y_1, \dots, y_N \mid \mathbf{x}_1, \dots, \mathbf{x}_N, \boldsymbol{\theta})$$

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The likelihood and the factors $p(y_i \mid \mathbf{x}_i, \boldsymbol{\theta})$ are Gaussian due to the noise distribution.

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The likelihood and the factors $p(y_i \mid \mathbf{x}_i, \boldsymbol{\theta})$ are Gaussian due to the noise distribution.

- Goal:** Find optimal parameters $\boldsymbol{\theta}^* \in \mathbb{R}^D$.
- Then we can make predictions for an arbitrary test input \mathbf{x}_* and get target y_* with $p(y_* \mid \mathbf{x}_*, \boldsymbol{\theta}^*) = \mathcal{N}(y_* \mid \mathbf{x}_*^\top \boldsymbol{\theta}^*, \sigma^2)$.

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Maximum Likelihood Estimation (MLE)

Find parameters θ_{ML}

$$\theta_{ML} \in \arg \max_{\theta} p(\mathcal{Y} | \mathcal{X}, \theta).$$

Note:

- The likelihood $p(y | x, \theta)$ is NOT a probability distribution of θ .

Maximum Likelihood Estimation (MLE)

Find parameters θ_{ML}

$$\theta_{ML} \in \arg \max_{\theta} p(\mathcal{Y} | \mathcal{X}, \theta).$$

Note:

- The likelihood $p(y | x, \theta)$ is NOT a probability distribution of θ . It's a function of θ (might not be integrable w.r.t θ).
- However, it's a normalized probability distribution in y .

How to find the desired θ_{ML} ?

- ① Perform gradient ascent (or descent).

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- ① Perform gradient ascent (or descent).
- ② For linear regression, we can directly have a **closed-form** solution.
- ③ In practice, we do not maximize the likelihood directly. Instead, we apply the **negative log-likelihood**.
 - It does not suffer from **numerical underflow**.
 - The differentiation rules become simpler.

Maximize likelihood \Leftrightarrow Minimize negative log-likelihood

The negative log-likelihood

$$-\log p(\mathcal{Y} \mid \mathcal{X}, \theta) = -\log \prod_{i=1}^N p(y_i \mid \mathbf{x}_i, \theta)$$

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- ★ **Note:** the independence assumption on the training set applies here.

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- * **Note:** the independence assumption on the training set applies here.

$$\log p(y_i \mid \mathbf{x}_i, \theta) = -\frac{1}{2\sigma^2}(y_i - \mathbf{x}^\top \theta)^2 + \text{constant}_{\text{indepent of } \theta}.$$

Ignoring the constant terms, we obtain

$$\begin{aligned}\mathcal{L}(\theta) &:= \frac{1}{2\sigma^2} \sum_{i=1}^N (y_i - \mathbf{x}_i^\top \boldsymbol{\theta})^2 \\ &= \frac{1}{2\sigma^2} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^\top (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) = \frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|^2,\end{aligned}$$

where $\mathbf{X} := [\mathbf{x}_1, \dots, \mathbf{x}_N]^\top \in \mathbb{R}^{N \times D}$ and $\mathbf{y} := [y_1, \dots, y_N]^\top \in \mathbb{R}^N$.

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$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial \boldsymbol{\theta}} = \mathbf{0}^\top &\iff \boldsymbol{\theta}_{ML}^\top \mathbf{X}^\top \mathbf{X} = \mathbf{y}^\top \mathbf{X} \\ &\iff \boldsymbol{\theta}_{ML}^\top = \mathbf{y}^\top \mathbf{X} (\mathbf{X}^\top \mathbf{X})^{-1} \\ &\iff \boldsymbol{\theta}_{ML} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}.\end{aligned}$$

- * We use the positive definite property of $\mathbf{X}^\top \mathbf{X}$ if $\text{rank}(\mathbf{X}) = D$.

Remark

- We can get a global minimum because the Hessian $\nabla_{\theta}^2 \mathcal{L}(\theta) = \mathbf{X}^T \mathbf{X}$ is positive definite (for full rank \mathbf{X} ?).

MLE with Features

- Note that “linear” regression is linear in the “parameters”.
- We can perform an arbitrary **nonlinear** transformation $\phi(\mathbf{x})$ of the input \mathbf{x} , and then linearly combine these components.

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$$p(y | \mathbf{x}, \boldsymbol{\theta}) = \mathcal{N}(y | \phi^T(\mathbf{x})\boldsymbol{\theta}, \sigma^2).$$



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$$p(y | \mathbf{x}, \boldsymbol{\theta}) = \mathcal{N}(y | \phi^\top(\mathbf{x})\boldsymbol{\theta}, \sigma^2).$$

\iff

$$y = \phi^\top(\mathbf{x})\boldsymbol{\theta} + \epsilon = \sum_{k=0}^{K-1} \theta_k \phi_k(\mathbf{x}) + \epsilon$$

- $\phi : \mathbb{R}^D \rightarrow \mathbb{R}$ is a (nonlinear) transformation of the input \mathbf{x}
- $\phi_k : \mathbb{R}^D \rightarrow \mathbb{R}$: the k th feature vector of ϕ .

Polynomial Regression (Example)

Consider a regression problem $y = \phi^\top(x)\theta + \epsilon$, for $x \in \mathbb{R}$ and $\theta \in \mathbb{R}^K$. A polynomial transformation of x is often used as

$$\phi(x) = \begin{bmatrix} \phi_0(x) \\ \phi_1(x) \\ \vdots \\ \phi_{K-1}(x) \end{bmatrix} = \begin{bmatrix} 1 \\ x \\ x^2 \\ \vdots \\ x^{K-1} \end{bmatrix} \in \mathbb{R}^K.$$

- We lift the original one-dimensional input space into a K -dimensional feature space.
- We can model polynomials of degree $\leq K - 1$ as $f(x) = \sum_{k=1}^{K-1} \theta_k x^k = \phi^\top(x)\theta$, for $\theta = [\theta_0, \dots, \theta_{K-1}]^\top \in \mathbb{R}^K$ which contains the linear parameters θ_k .

For $\mathbf{x}_i \in \mathbb{R}^D$

We can also define a feature matrix as

$$\Phi := \begin{bmatrix} \phi^\top(\mathbf{x}_1) \\ \vdots \\ \phi^\top(\mathbf{x}_N) \end{bmatrix} = \begin{bmatrix} \phi_0(\mathbf{x}_1) & \cdots & \phi_{K-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \cdots & \phi_{K-1}(\mathbf{x}_2) \\ \vdots & & \vdots \\ \phi_0(\mathbf{x}_N) & \cdots & \phi_{K-1}(\mathbf{x}_N) \end{bmatrix} \in \mathbb{R}^{N \times K},$$

where $\Phi_{ij} = \phi_j(\mathbf{x}_i)$ and $\phi_j : \mathbb{R}^D \rightarrow \mathbb{R}$.

Example

Feature Matrix for Second-Order Polynomials

$$\Phi := \begin{bmatrix} 1 & x_1 & x_1^2 \\ 1 & x_2 & x_2^2 \\ \vdots & \vdots & \vdots \\ 1 & x_N & x_N^2 \end{bmatrix}.$$

With the feature matrix Φ :

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The negative log-likelihood can be written as

$$-\log p(\mathcal{Y} | \mathcal{X}, \theta) = \frac{1}{2\sigma^2} (\mathbf{y} - \Phi\theta)^\top (\mathbf{y} - \Phi\theta) + \text{constant}.$$

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- Both of them are independent of θ .

¹Requiring $\text{rank}(\Phi) = K$

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- Replacing \mathbf{X} by Φ .
- Both of them are independent of θ .
- Similarly, we have¹

$$\theta_{ML} = (\Phi^\top \Phi)^{-1} \Phi^\top \mathbf{y}.$$

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Estimating the Noise Variance (1/2)

- We can also use the principle of MLE to obtain that for σ^2_{ML} for the noise variance.

Estimating the Noise Variance (1/2)

- We can also use the principle of MLE to obtain that for σ_{ML}^2 for the noise variance.
- Write down the log-likelihood:

$$\begin{aligned}
 \log p(\mathcal{Y} | \mathcal{X}, \boldsymbol{\theta}, \sigma^2) &= \sum_{i=1}^N \log \mathcal{N}(y_i | \boldsymbol{\phi}^\top(\mathbf{x}_i)\boldsymbol{\theta}, \sigma^2) \\
 &= \sum_{i=1}^N \left(-\frac{1}{2} \log(2\pi) - \frac{1}{2} \log \sigma^2 - \frac{1}{2\sigma^2} (y_i - \boldsymbol{\phi}^\top(\mathbf{x}_i)\boldsymbol{\theta})^2 \right) \\
 &= -\frac{N}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^N (y_i - \boldsymbol{\phi}^\top(\mathbf{x}_i)\boldsymbol{\theta})^2 + \text{constant}
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 &= -\frac{N}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^N (y_i - \boldsymbol{\phi}^\top(\mathbf{x}_i)\boldsymbol{\theta})^2 + \text{constant}
 \end{aligned}$$

Let $s := \sum_{i=1}^N (y_i - \boldsymbol{\phi}^\top(\mathbf{x}_i)\boldsymbol{\theta})^2$.

Estimating the Noise Variance (2/2)

- The partial derivative w.r.t. σ^2 :

$$\frac{\partial \log p(\mathcal{Y} | \mathcal{X}, \boldsymbol{\theta}, \sigma^2)}{\partial \sigma^2} = -\frac{N}{2\sigma^2} + \frac{1}{\sigma^4}s = 0$$
$$\iff \frac{N}{2\sigma^2} = \frac{s}{2\sigma^4}.$$

Thus,

$$\sigma_{ML}^2 = \frac{s}{N} = \frac{1}{N} \sum_{i=1}^N (y_i - \boldsymbol{\phi}^\top(\mathbf{x}_i) \boldsymbol{\theta})^2.$$

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Evaluating the Quality of the Model

- We can evaluate the quality of the model by computing the error/loss.
- Given that σ^2 is not a free model parameter, we can ignore that term by scaling by $1/\sigma^2$ and derive a squared-error function $\|\mathbf{y} - \Phi\boldsymbol{\theta}\|^2$.

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- Given that σ^2 is not a free model parameter, we can ignore that term by scaling by $1/\sigma^2$ and derive a squared-error function $\|\mathbf{y} - \Phi\boldsymbol{\theta}\|^2$.
- To compare the errors of datasets with **different sizes** and **the same scale**, we often use the root-mean squared error (RMSE):

Evaluating the Quality of the Model

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- Model selection:

Evaluating the Quality of the Model

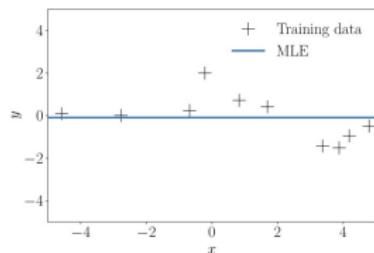
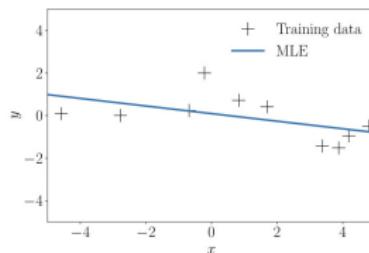
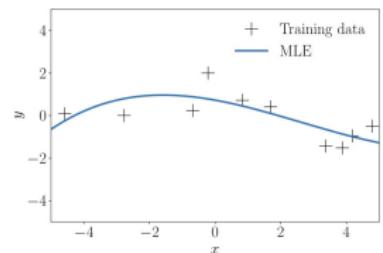
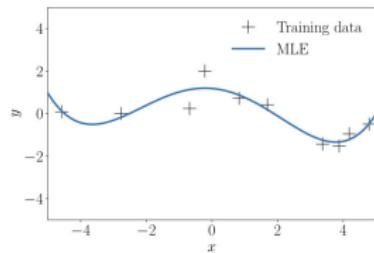
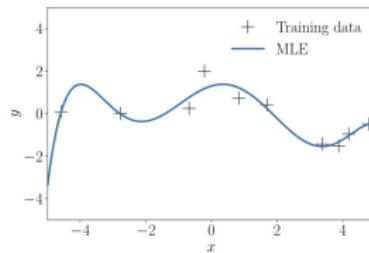
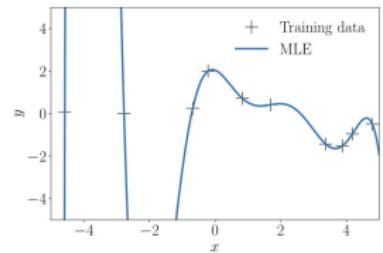
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- Model selection: determine the best degree of the polynomial.
 - Brute-force searching and enumerate all reasonable polynomial degrees M .

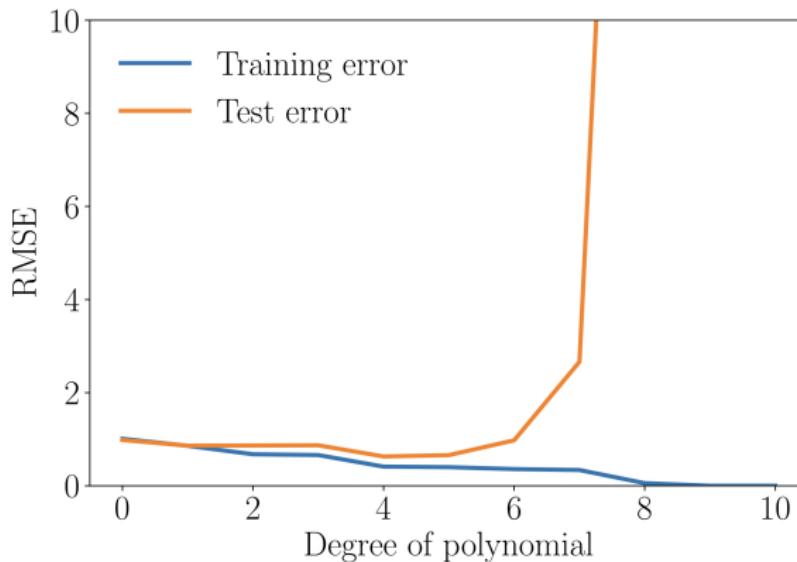
Parameter Estimation

Overfitting in Linear Regression

(a) $M = 0$ (b) $M = 1$ (c) $M = 3$ (d) $M = 4$ (e) $M = 6$ (f) $M = 9$

Goal: a good generalization by making *accurate* predictions for new unseen data.

- A separate test set comprising 200 data points generated using exactly the same procedure used to generate the training set.



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- MLE is prone to overfitting.
- **Experience:** The parameter values becomes relatively large when the model is overfitting.
- To mitigate the effect of huge parameter values, we place a **prior distribution** $p(\theta)$ on the parameters.
- **Rough idea:** Encode the parameter values that are plausible before seeing any data.
 - For example, a Gaussian prior $p(\theta) = \mathcal{N}(\mathbf{0}, \mathbf{I})$.

Maximum a Posteriori Estimation (1/5)

- Once a dataset $(\mathcal{X}, \mathcal{Y})$ is available, we seek parameters that maximize the posterior distribution $p(\theta | \mathcal{X}, \mathcal{Y})$ instead of maximizing the likelihood.

$$p(\theta | \mathcal{X}, \mathcal{Y}) = \frac{p(\mathcal{Y} | \mathcal{X}, \theta)p(\theta)}{p(\mathcal{Y} | \mathcal{X})}.$$

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- The prior will have an effect on the parameter vector.
- θ_{MAP} : the maximizer of the above posterior (i.e., the MAP estimate).

Maximum a Posteriori Estimation (2/5)

The log-transformation of the posterior:

$$\log p(\theta | \mathcal{X}, \mathcal{Y}) = \log p(\mathcal{Y} | \mathcal{X}, \theta) + \log p(\theta) + \text{constant}$$

The constant is independent of θ .

We can see that the MAP estimate is a compromise between the prior and the data-dependent likelihood.

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We minimize the negative log-posterior w.r.t. θ :

$$\theta_{MAP} \in \arg \min_{\theta} \{-\log p(\mathcal{Y} | \mathcal{X}, \theta) - \log p(\theta)\}.$$

Maximum a Posteriori Estimation (3/5)

$$\boldsymbol{\theta}_{MAP} \in \arg \min_{\boldsymbol{\theta}} \{-\log p(\mathcal{Y} | \mathcal{X}, \boldsymbol{\theta}) - \log p(\boldsymbol{\theta})\}.$$

The gradient:

$$-\frac{d \log p(\boldsymbol{\theta} | \mathcal{X}, \mathcal{Y})}{d \boldsymbol{\theta}} = -\frac{d \log p(\mathcal{Y} | \mathcal{X}, \boldsymbol{\theta})}{d \boldsymbol{\theta}} - \frac{d \log p(\boldsymbol{\theta})}{d \boldsymbol{\theta}}.$$

Assume the Gaussian prior $p(\boldsymbol{\theta}) = \mathcal{N}(\mathbf{0}, b^2 \mathbf{I})$. We have

$$-\log p(\boldsymbol{\theta} | \mathcal{X}, \mathcal{Y}) = \frac{1}{2\sigma^2} (\mathbf{y} - \Phi \boldsymbol{\theta})^\top (\mathbf{y} - \Phi \boldsymbol{\theta}) + \frac{1}{2b^2} \boldsymbol{\theta}^\top \boldsymbol{\theta} + \text{constant}$$

Maximum a Posteriori Estimation (4/5)

$$-\log p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y}) = \frac{1}{2\sigma^2} (\mathbf{y} - \Phi\boldsymbol{\theta})^\top (\mathbf{y} - \Phi\boldsymbol{\theta}) + \frac{1}{2b^2} \boldsymbol{\theta}^\top \boldsymbol{\theta} + \text{constant}$$

Hence, the gradient of the log-posterior w.r.t. $\boldsymbol{\theta}$ is

$$-\frac{d \log p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y})}{d\boldsymbol{\theta}} = \frac{1}{\sigma^2} (\boldsymbol{\theta}^\top \Phi^\top \Phi - \mathbf{y}^\top \Phi) + \frac{1}{b^2} \boldsymbol{\theta}^\top.$$

Setting the gradient to $\mathbf{0}^\top$ to get $\boldsymbol{\theta}_{MAP}$:

Maximum a Posteriori Estimation (5/5)

$$\begin{aligned} & \frac{1}{\sigma^2}(\boldsymbol{\theta}^\top \boldsymbol{\Phi}^\top \boldsymbol{\Phi} - \mathbf{y}^\top \boldsymbol{\Phi}) + \frac{1}{b^2} \boldsymbol{\theta}^\top = \mathbf{0}^\top \\ \iff & \boldsymbol{\theta}^\top \left(\frac{1}{\sigma^2} \boldsymbol{\Phi}^\top \boldsymbol{\Phi} + \frac{1}{b^2} \mathbf{I} \right) - \frac{1}{\sigma^2} \mathbf{y}^\top \boldsymbol{\Phi} = \mathbf{0}^\top \\ \iff & \boldsymbol{\theta}^\top \left(\boldsymbol{\Phi}^\top \boldsymbol{\Phi} + \frac{\sigma^2}{b^2} \mathbf{I} \right) = \mathbf{y}^\top \boldsymbol{\Phi} \\ \iff & \boldsymbol{\theta}^\top = \mathbf{y}^\top \boldsymbol{\Phi} \left(\boldsymbol{\Phi}^\top \boldsymbol{\Phi} + \frac{\sigma^2}{b^2} \mathbf{I} \right)^{-1}. \end{aligned}$$

Finally, we have

Maximum a Posteriori Estimation (5/5)

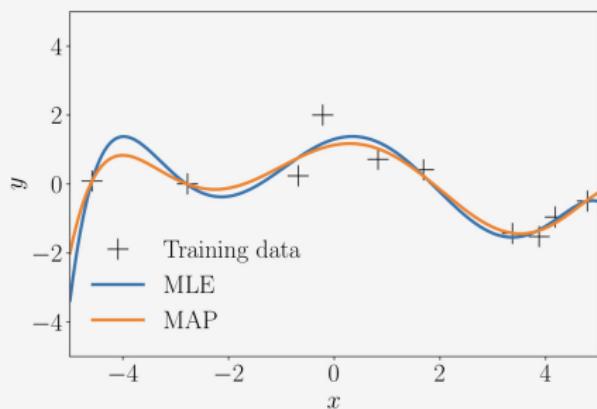
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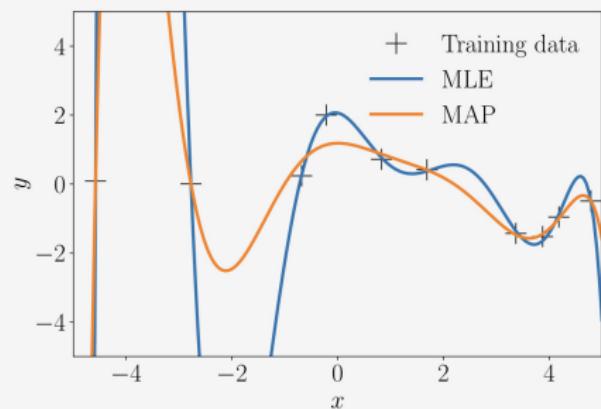
$$\boldsymbol{\theta}_{MAP} = \left(\Phi^\top \Phi + \frac{\sigma^2}{b^2} \mathbf{I} \right)^{-1} \Phi^\top \mathbf{y}.$$

Parameter Estimation

Maximum A Posteriori Estimation (MAP)



(a) Polynomials of degree 6.



(b) Polynomials of degree 8.

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Motivation (I)

- Mitigate the effect of overfitting by **penalizing the amplitude of the parameters by means of regularization**.
- Consider the regularized least squares:

$$\underbrace{\|\mathbf{y} - \Phi\boldsymbol{\theta}\|^2}_{\text{for fitting data}} + \underbrace{\lambda\|\boldsymbol{\theta}\|_2^2}_{\text{regularizer}}$$

for the regularization parameter $\lambda \geq 0$.

- The 2-norm $\|\cdot\|_2$ can be replaced by other types of norm.

Motivation (II)

- The regularizer $\lambda\|\theta\|_2^2$ can be seen as a negative log-Gaussian prior.
- The Gaussian prior $p(\theta) = \mathcal{N}(\mathbf{0}, b^2 \mathbf{I})$, so the negative log-Gaussian prior is

$$-\log p(\theta) = \frac{1}{2b^2}\|\theta\|_2^2 + \text{constant}$$

hence we have $\lambda = \frac{1}{2b^2}$.

Minimizing the regularized least-squares loss function yields

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This is identical to the MAP estimate for $\lambda = \frac{\sigma^2}{b^2}$.

- σ^2 : the noise variance
- b^2 : the variance of the Gaussian prior $p(\boldsymbol{\theta}) = \mathcal{N}(\mathbf{0}, b^2 \mathbf{I})$.

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From MAP to Bayesian Linear Regression

- So far we have used **point estimates** of the parameters θ :
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 - Make predictions by **integrating over all plausible parameter values**.
- This will give us not only a prediction, but also a **measure of uncertainty**.

Bayesian Linear Regression Model (1/2)

- Recall the feature-based linear regression model

$$y = \phi^\top(\mathbf{x})\boldsymbol{\theta} + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \sigma^2).$$

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where

- $\Phi \in \mathbb{R}^{N \times K}$: feature matrix, i -th row $\boldsymbol{\phi}^\top(\mathbf{x}_i)$,
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- The **likelihood** is Gaussian:

$$p(\mathcal{Y} | \mathcal{X}, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{y} | \Phi\boldsymbol{\theta}, \sigma^2 \mathbf{I}_N).$$

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$$p(\theta) = \mathcal{N}(\theta | \mathbf{m}_0, \mathbf{S}_0),$$

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- This prior encodes our **a priori beliefs** about likely parameter values.
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- Because both prior and likelihood are Gaussian, the posterior will also be Gaussian (Exercise).

Remark: Product of Gaussian Densities

- Consider two D -dimensional Gaussians in **the same variable x :**

$$\mathcal{N}(x | \mathbf{a}, \mathbf{A}) \text{ and } \mathcal{N}(x | \mathbf{b}, \mathbf{B}).$$

- Their product is proportional to another Gaussian:

$$\mathcal{N}(x | \mathbf{a}, \mathbf{A}) \mathcal{N}(x | \mathbf{b}, \mathbf{B}) = \rho \mathcal{N}(x | \mathbf{c}, \mathbf{C}),$$

where

$$\mathbf{C} = (\mathbf{A}^{-1} + \mathbf{B}^{-1})^{-1}, \quad \mathbf{c} = \mathbf{C}(\mathbf{A}^{-1}\mathbf{a} + \mathbf{B}^{-1}\mathbf{b}),$$

and the scaling constant ρ is

$$\rho = (2\pi)^{-D/2} \det(\mathbf{A} + \mathbf{B})^{-1/2} \exp\left(-\frac{1}{2}(\mathbf{a} - \mathbf{b})^\top (\mathbf{A} + \mathbf{B})^{-1}(\mathbf{a} - \mathbf{b})\right).$$

- The constant ρ can itself be written as a Gaussian density with “inflated” covariance $\mathbf{A} + \mathbf{B}$:

$$\rho = \mathcal{N}(\mathbf{a} | \mathbf{b}, \mathbf{A} + \mathbf{B}) = \mathcal{N}(\mathbf{b} | \mathbf{a}, \mathbf{A} + \mathbf{B}).$$

Posterior over Parameters (1/2)

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- The exponent is a **quadratic function** of $\theta \Rightarrow$ the posterior is Gaussian.

Posterior over Parameters (2/2)

Posterior of θ (proof skipped)

The posterior is

$$p(\theta | \mathcal{X}, \mathcal{Y}) = \mathcal{N}(\theta | \mathbf{m}_N, \mathbf{S}_N),$$

where the posterior covariance and mean are given by

$$\mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \frac{1}{\sigma^2} \Phi^\top \Phi,$$

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- \mathbf{S}_N : the prior uncertainty \mathbf{S}_0 + information from the data.
- \mathbf{m}_N : a **precision-weighted average** of prior mean and data evidence.

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- Fewer data or very noisy data (large σ^2) \Rightarrow posterior is closer to the prior.
- For a Gaussian posterior, the **MAP estimate** and the **posterior mean** coincide, but
 - MAP uses only the mode θ_{MAP} ,
 - Bayesian prediction uses the *full posterior* $p(\theta | \mathcal{X}, \mathcal{Y})$.

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- It reflects what we expect *before* seeing any training data.

Posterior Predictive Distribution (1/2)

- After observing the dataset \mathcal{D} , we use the posterior to make predictions:

$$p(y_* \mid \mathbf{x}_*, \mathcal{X}, \mathcal{Y}) = \int p(y_* \mid \mathbf{x}_*, \boldsymbol{\theta}) p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y}) d\boldsymbol{\theta}.$$

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- Again, both distributions inside the integral are Gaussian \Rightarrow the **posterior predictive** is Gaussian.
- Intuitively:
 - we average predictions over all plausible parameter values,
 - weighted by how probable they are under the posterior.

Posterior Predictive Distribution (2/2)

Posterior predictive of y_*

The predictive distribution at a new input \mathbf{x}_* is

$$p(y_* \mid \mathbf{x}_*, \mathcal{X}, \mathcal{Y}) = \mathcal{N}(y_* \mid \mu_*(\mathbf{x}_*), \sigma_*^2(\mathbf{x}_*)) ,$$

with mean

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Posterior Predictive Distribution (2/2)

Posterior predictive of y_*

The predictive distribution at a new input \mathbf{x}_* is

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- The first term in σ_*^2 quantifies parameter uncertainty.
- The second term σ^2 is the measurement noise.

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- Bayesian linear regression therefore yields
 - a **mean prediction** $\mu_*(\mathbf{x}_*)$ and
 - a **credible interval** (e.g., mean $\pm 2\sigma_*(\mathbf{x}_*)$).
- This is very useful for model assessment and decision making.

Example: Polynomial Regression Revisited

- Recall the polynomial feature maps

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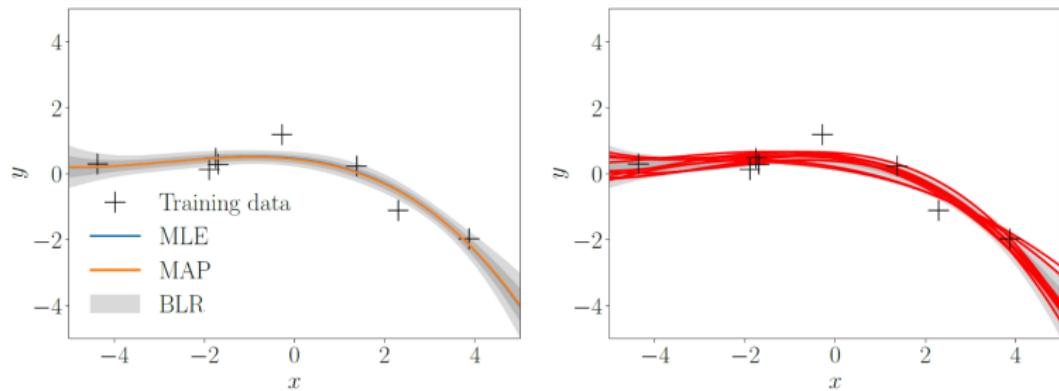
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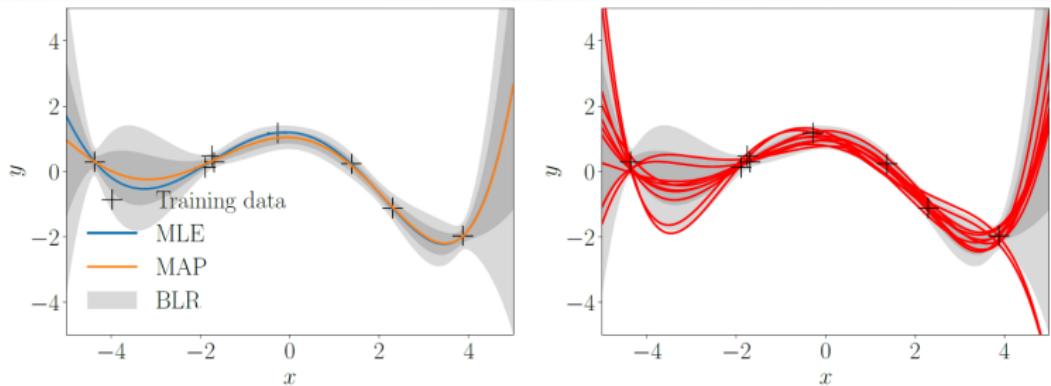
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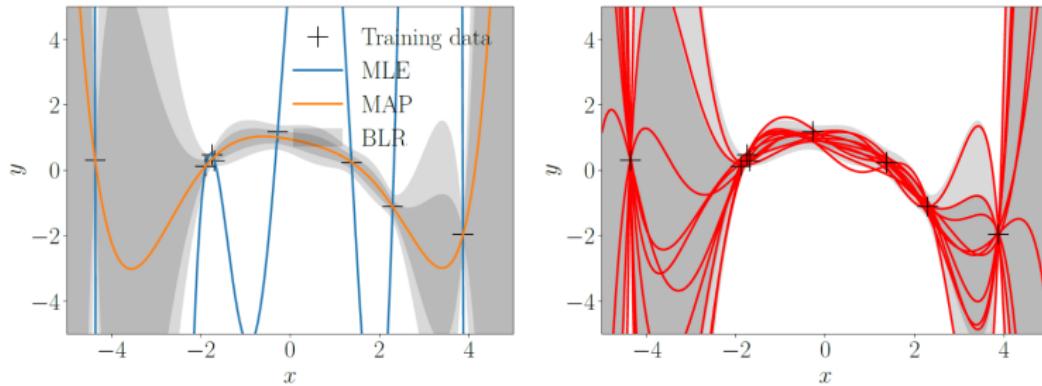
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- In contrast, Bayesian linear regression
 - produces a **distribution** over curves,
 - with narrow uncertainty bands where data are dense,
 - and wide bands where there are few or no observations.
- This behaviour matches the qualitative picture in the textbook for Bayesian linear regression.



(a) Posterior distribution for polynomials of degree $M = 3$ (left) and samples from the posterior over functions (right).



(b) Posterior distribution for polynomials of degree $M = 5$ (left) and samples from the posterior over functions (right).



(c) Posterior distribution for polynomials of degree $M = 7$ (left) and samples from the posterior over functions (right).

Marginal Likelihood (1/2)

- The **marginal likelihood** (or **model evidence**) is

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- It measures how well the model (including the prior) explains the data.
- In Bayesian linear regression, the integral can be computed in closed form.

Marginal Likelihood (2/2)

Closed-form marginal likelihood

Using Gaussian identities, we obtain

$$p(\mathcal{Y} | \mathcal{X}) = \mathcal{N}\left(\mathbf{y} | \Phi \mathbf{m}_0, \Phi \mathbf{S}_0 \Phi^\top + \sigma^2 \mathbf{I}_N\right).$$

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- Different model choices (e.g., different polynomial degrees) lead to different marginal likelihoods.
- The marginal likelihood automatically trades off **data fit** (i.e., likelihood) and **model complexity** (i.e., prior).

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 - variance $\sigma_*^2(\mathbf{x}_*) = \phi^\top(\mathbf{x}_*) \mathbf{S}_N \phi(\mathbf{x}_*) + \sigma^2$.
- Bayesian linear regression improves over MLE and MAP by
 - quantifying parameter and predictive uncertainty,
 - enabling principled model comparison via the marginal likelihood.

Discussions

Theorem 9.1 (Parameter Posterior)

Consider the linear regression model with Gaussian noise

$$\mathbf{y} \mid \boldsymbol{\theta}, \mathcal{X} \sim \mathcal{N}(\boldsymbol{\Phi}\boldsymbol{\theta}, \sigma^2 \mathbf{I}),$$

and Gaussian prior on the parameters

$$\boldsymbol{\theta} \sim \mathcal{N}(\mathbf{m}_0, \mathbf{S}_0),$$

where $\boldsymbol{\Phi}$ is the design matrix, \mathbf{m}_0 the prior mean and \mathbf{S}_0 the prior covariance.

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Theorem 9.1 (Parameter Posterior)

The posterior over parameters is Gaussian:

$$p(\boldsymbol{\theta} | \mathcal{X}, \mathcal{Y}) = \mathcal{N}(\boldsymbol{\theta} | \mathbf{m}_N, \mathbf{S}_N),$$

$$\text{with } \mathbf{S}_N = (\mathbf{S}_0^{-1} + \sigma^{-2} \boldsymbol{\Phi}^\top \boldsymbol{\Phi})^{-1}, \quad \mathbf{m}_N = \mathbf{S}_N (\mathbf{S}_0^{-1} \mathbf{m}_0 + \sigma^{-2} \boldsymbol{\Phi}^\top \mathbf{y}),$$

where the subscript N indicates dependence on the full training set.

Proof of Theorem 9.1 (1/3)

- By Bayes' rule, up to a normalizing constant,

$$p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y}) \propto p(\mathbf{y} \mid \mathcal{X}, \boldsymbol{\theta}) p(\boldsymbol{\theta}).$$

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$$p(\boldsymbol{\theta}) = \mathcal{N}(\boldsymbol{\theta} \mid \mathbf{m}_0, \mathbf{S}_0).$$

- Work in log-space (and drop constants independent of $\boldsymbol{\theta}$):

$$\log p(\mathbf{y} \mid \mathcal{X}, \boldsymbol{\theta}) = -\frac{1}{2\sigma^2} (\mathbf{y} - \Phi\boldsymbol{\theta})^\top (\mathbf{y} - \Phi\boldsymbol{\theta}) + \text{const},$$

$$\log p(\boldsymbol{\theta}) = -\frac{1}{2} (\boldsymbol{\theta} - \mathbf{m}_0)^\top \mathbf{S}_0^{-1} (\boldsymbol{\theta} - \mathbf{m}_0) + \text{const.}$$

Proof of Theorem 9.1 (2/3) – completing the squares

- Summing the two log terms (still ignoring constants):

$$\begin{aligned}\log p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y}) &= -\frac{1}{2\sigma^2}(\mathbf{y} - \Phi\boldsymbol{\theta})^\top(\mathbf{y} - \Phi\boldsymbol{\theta}) \\ &\quad - \frac{1}{2}(\boldsymbol{\theta} - \mathbf{m}_0)^\top \mathbf{S}_0^{-1}(\boldsymbol{\theta} - \mathbf{m}_0) + \text{const.}\end{aligned}$$

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- Summing the two log terms (still ignoring constants):

$$\begin{aligned}\log p(\theta | \mathcal{X}, \mathcal{Y}) &= -\frac{1}{2\sigma^2}(\mathbf{y} - \Phi\theta)^\top(\mathbf{y} - \Phi\theta) \\ &\quad - \frac{1}{2}(\theta - \mathbf{m}_0)^\top \mathbf{S}_0^{-1}(\theta - \mathbf{m}_0) + \text{const.}\end{aligned}$$

- Expand the quadratic terms in θ :

$$(\mathbf{y} - \Phi\theta)^\top(\mathbf{y} - \Phi\theta) = \mathbf{y}^\top \mathbf{y} - 2\mathbf{y}^\top \Phi\theta + \theta^\top \Phi^\top \Phi\theta,$$

$$(\theta - \mathbf{m}_0)^\top \mathbf{S}_0^{-1}(\theta - \mathbf{m}_0) = \theta^\top \mathbf{S}_0^{-1}\theta - 2\mathbf{m}_0^\top \mathbf{S}_0^{-1}\theta + \mathbf{m}_0^\top \mathbf{S}_0^{-1}\mathbf{m}_0.$$

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- Expand the quadratic terms in $\boldsymbol{\theta}$:

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$$(\boldsymbol{\theta} - \mathbf{m}_0)^\top \mathbf{S}_0^{-1}(\boldsymbol{\theta} - \mathbf{m}_0) = \boldsymbol{\theta}^\top \mathbf{S}_0^{-1}\boldsymbol{\theta} - 2\mathbf{m}_0^\top \mathbf{S}_0^{-1}\boldsymbol{\theta} + \mathbf{m}_0^\top \mathbf{S}_0^{-1}\mathbf{m}_0.$$

- Collect all terms that depend on $\boldsymbol{\theta}$:

$$\begin{aligned}\log p(\boldsymbol{\theta} \mid \mathcal{X}, \mathcal{Y}) &= -\frac{1}{2} \left[\boldsymbol{\theta}^\top (\sigma^{-2} \Phi^\top \Phi + \mathbf{S}_0^{-1}) \boldsymbol{\theta} \right. \\ &\quad \left. - 2(\sigma^{-2} \Phi^\top \mathbf{y} + \mathbf{S}_0^{-1} \mathbf{m}_0)^\top \boldsymbol{\theta} \right] + \text{const.}\end{aligned}$$

Proof of Theorem 9.1 (3/3)

- A multivariate Gaussian $\mathcal{N}(\boldsymbol{\theta} \mid \mathbf{m}_N, \mathbf{S}_N)$ has log-density (up to an additive constant)

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- Compare this with the quadratic form obtained on the previous slide:

$$\boldsymbol{\theta}^\top (\sigma^{-2} \Phi^\top \Phi + \mathbf{S}_0^{-1}) \boldsymbol{\theta} \quad \text{and} \quad (\sigma^{-2} \Phi^\top \mathbf{y} + \mathbf{S}_0^{-1} \mathbf{m}_0)^\top \boldsymbol{\theta}.$$

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- Matching the coefficients of the quadratic and linear terms in $\boldsymbol{\theta}$ gives

$$\mathbf{S}_N^{-1} = \sigma^{-2} \Phi^\top \Phi + \mathbf{S}_0^{-1} \implies \mathbf{S}_N = (\mathbf{S}_0^{-1} + \sigma^{-2} \Phi^\top \Phi)^{-1},$$

$$\mathbf{m}_N^\top \mathbf{S}_N^{-1} = (\sigma^{-2} \Phi^\top \mathbf{y} + \mathbf{S}_0^{-1} \mathbf{m}_0)^\top \implies \mathbf{m}_N = \mathbf{S}_N (\mathbf{S}_0^{-1} \mathbf{m}_0 + \sigma^{-2} \Phi^\top \mathbf{y}).$$