Basic Stats/Probability

Some definitions for everything below:

- \bullet X, Y, and Z are random variables
- a, b, c, d are constants
- Whenever you see μ_X or σ_X^2 , know that this is the mean and variance (respectively) of the random variable X
- \overline{X}_N is a sample mean from a sample of random variables X_n drawn from the same distribution.
- s_X^2 is the sample variance calculated from a sample of X.

Most of the rules follow almost immediately from the basic definitions, so you should test yourself by trying to prove each of them. It will also help you a lot if you know these rules like the back of your hand.

- $\mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]$
- $\mathbb{C}(X,Y) = \mathbb{E}[(X \mu_X)(Y \mu_Y)] = \mathbb{C}(Y,X)$
- $\mathbb{C}(a+bX,Y) = a\mathbb{C}(X,Y)$ $\Rightarrow \mathbb{C}(a+bX,c+dY) = ad\mathbb{C}(X,Y).$
- $\mathbb{C}(aX + bY, Z) = a\mathbb{C}(X, Z) + b\mathbb{C}(Y, Z)$
- $X \perp \!\!\! \perp Y \Rightarrow \mathbb{C}(X,Y) = 0$
- $\mathbb{V}[X+Y] = X + Y + 2\mathbb{C}(X,Y)$
- $\Rightarrow \mathbb{V}\left[\sum_{n} X_{n}\right] = \sum_{n} \left[X_{n}\right] \text{ if } X_{1} \perp X_{2} \dots \perp X_{N}$
- $\bullet \ \mathbb{E}[\overline{X}_N] = \mu_X$
- $\mathbb{V}[\overline{X}_N] = \frac{1}{N} \mathbb{V}[X]$ if $X_1 \perp \!\!\! \perp X_2 ... \perp \!\!\! \perp X_N$ (i.e. iid sample)
- $\mathbb{E}[X|Y] = 0 \Rightarrow \mathbb{E}[XY] = 0$
- $\mathbb{E}[X|Y] = a \Rightarrow \mathbb{C}(X,Y) = 0$
- If $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$, $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$ and Z = X + Y, then $Z \sim \mathcal{N}(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Z^2 + 2\sigma_{XY})$ where $\sigma_{XY} = \mathbb{C}(X, Y)$.
- When \overline{X}_N comes from an iid sample and $\mathbb{V}[X] < \infty$, $\lim_{N \to \infty} P\left[\frac{\overline{X}_N \mu_X}{s_X} \le z_{\alpha}\right] = 1 \alpha$ where z_{α} solves $P[Z \le z_{\alpha}] = 1 \alpha$ where $Z \sim \mathcal{N}(0, 1)$.

Now, let X be a random column vector, and let $X_1, X_2, ..., X_N$ be an iid sample of size N. Let $\mu_X = \mathbb{E}[X]$ and $\Sigma_X = \mathbb{V}[X] = \mathbb{E}[(X - \mu_X)(X - \mu_X)^T]$. Let C be a constant matrix.

- $\mathbb{V}[CX] = C\Sigma_X C^T$.
- $\overline{X} \to_p \mu_X$ (law of large numbers)
- $\sqrt{N}(\overline{X} \mu_X) \to_d (0, \Sigma_X)$ (central limit theorem)

OLS

Let $\mathbb{E}[Y_n|\boldsymbol{x}_n] = \boldsymbol{x}_n\beta$ where \boldsymbol{x}_n is a row vector. Let \boldsymbol{X} be N iid observations of \boldsymbol{x}_n stacked vertically and the same for \boldsymbol{Y} .

- OLS estimator: $\hat{\beta} = (\boldsymbol{X}^T \boldsymbol{X})^{-1} \boldsymbol{X}^T \boldsymbol{Y}$
- Variance for $\hat{\beta}$:

$$\mathbb{V}[\hat{\beta}] = \frac{1}{N} \mathbb{E}[\boldsymbol{x}_n^T \boldsymbol{x}_n]^{-1} \mathbb{E}[\boldsymbol{x}_n^T \boldsymbol{x}_n \epsilon_n^2] \mathbb{E}[\boldsymbol{x}_n^T \boldsymbol{x}_n]^{-1}$$

• And if $\mathbb{E}[\epsilon_n^2|\boldsymbol{x}_n] = \sigma^2$,

$$\mathbb{V}[\hat{eta}] = rac{1}{N}\mathbb{E}[oldsymbol{x}_n^Toldsymbol{x}_n]^{-1}\sigma^2$$

• To estimate $\mathbb{V}[\hat{\beta}]$ in general:

$$\hat{V}_{eta} = (oldsymbol{X}^Toldsymbol{X})^{-1} \sum_n oldsymbol{x}_n^Toldsymbol{x} \hat{\epsilon}_n^2 (oldsymbol{X}^Toldsymbol{X})^{-1}$$

• And if $\mathbb{E}[\epsilon_n^2|\boldsymbol{x}_n] = \sigma^2$:

$$\hat{V}_{\hat{\beta}} = \frac{1}{N} (\boldsymbol{X}^T \boldsymbol{X})^{-1} \sum_n \hat{\epsilon}_n^2$$