

# Introduction to frequentist statistics and Bayesian inference

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**(HUST GW Summer School 2022, Lecture 1)**

# References

- Romano and Cornish, Living Reviews in Relativity article, 2017 (section 3)
- Rover, Messenger, Prix, “Bayesian versus frequentist upper limits,” PHYSTAT 2011 workshop
- Gregory, “Bayesian Logical data analysis”, 2005
- Howson and Urbach, “Scientific reasoning: the Bayesian approach”, 2006
- Helstrom, “Statistical theory of signal detection”, 1968
- Wainstein and Zubakov, “Extraction of signals from noise,” 1971

# Outline

1. Probabilistic inference (broadly defined)
2. Frequentist statistics
3. Bayesian inference
4. Exercises - worked examples

# Frequentist vs Bayesian “pre-test”

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- Answer 2: You interpret 90% as the long-term relative frequency with which the true mass of the NS lies in the set of intervals  $\{[\hat{M} - 0.02M_{\odot}, \hat{M} + 0.02M_{\odot}]\}$  where  $\{\hat{M}\}$  is the set of measured masses.

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- If you chose answer 1, then you are a Bayesian
- If you chose answer 2, then you are a frequentist

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- **Different from** mathematical deduction

# I. Probabilistic inference

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NOTE: For the frequentist definition, probabilities can only be assigned to propositions about outcomes of repeatable identical experiments (i.e., **random variables**), not to hypotheses or parameters describing the state of nature, which have fixed but unknown values

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- NOTE:  $P(X | Y)$  is the probability of  $X$  conditioned on  $Y$  (assuming  $Y$  is true)
- $P(X | Y) \neq P(Y | X)$  in general. Example  $X$ ="person is pregnant",  $Y$ ="person is female"

# Bayes' theorem (a simple consequence of the product rule!!)

A diagram showing the components of Bayes' theorem. The equation is  $P(H|D) = \frac{P(D|H)P(H)}{P(D)}$ . Arrows point from labels to parts of the equation: 'posterior' points to  $P(H|D)$ , 'likelihood' points to  $P(D|H)$ , 'prior' points to  $P(H)$ , and 'evidence' points to  $P(D)$ .

$$\begin{array}{c} \text{posterior} \\ \swarrow \\ P(H|D) \end{array} = \frac{\begin{array}{c} \text{likelihood} \\ \swarrow \\ P(D|H) \end{array} \begin{array}{c} \text{prior} \\ \swarrow \\ P(H) \end{array}}{\begin{array}{c} P(D) \\ \swarrow \\ \text{evidence} \end{array}}$$

where  $P(D) = P(D|H)P(H) + P(D|\bar{H})P(\bar{H})$

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“Learning from experience”: the probability of  $H$  being true (in light of new data) increases by the ratio of the probability of obtaining the new data  $D$  when  $H$  is true to the probability of obtaining  $D$  in any case

# Bayes' theorem (for parameters associated with a given hypothesis or model)

$$p(a | d, H) = \frac{p(d | a, H)p(a | H)}{p(d | H)}$$

where  $p(d | H) = \int da \, p(d | a, H)p(a | H)$

“marginalization” over  $a$

# Comparing frequentist & Bayesian inference

| Frequentist statistics  | Bayesian inference   |
|---|--|
| Probabilities are <b>long-run relative occurrences</b> of outcomes of repeatable expts —> can't be assigned to hypotheses | Probabilities are <b>degree of belief</b> —> can be assigned to hypotheses   |
| Usually start with a <b>likelihood function</b> $p(d H)$  | Same as frequentist  |
| Construct a <b>statistic</b> (some function of the data $d$ ) for parameter estimation or hypothesis testing              | Need to specify <b>priors</b> for parameters and hypotheses  |
| Calculate <b>sampling distribution</b> of the statistics (e.g., using time slide)   | Use <b>Bayes' theorem</b> to update degree of belief in a parameter or hypothesis                                      |
| Calculates <b>confidence intervals</b> (for parameter estimation) and <b>p-values</b> (for hypothesis testing)            | Construct <b>posteriors</b> (for parameter estimation) and <b>odds ratios (Bayes factors)</b> (for hypothesis testing) |

## II. Frequentist statistics

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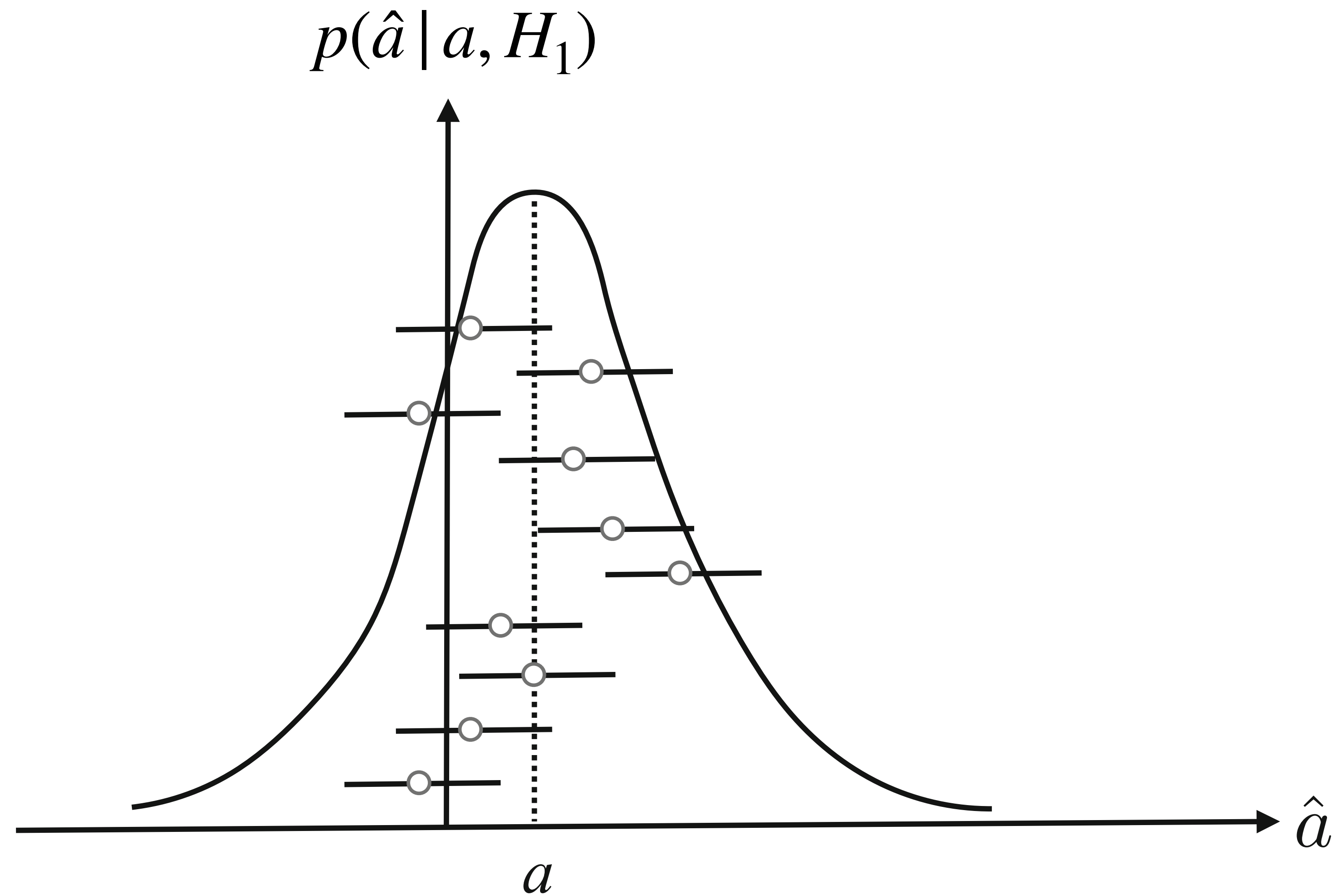
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- Statements like  $\text{Prob}(a - \Delta < \hat{a} < a + \Delta)$  make sense since  $\hat{a}$  is a random variable
- Statements like  $a = \hat{a} \pm \Delta$  with 90% confidence must be interpreted as statements about the **randomness of the intervals**—i.e., 90% is the long-term relative frequency with which the true value of the parameter lies in the set of intervals  $\{[\hat{a} - \Delta, \hat{a} + \Delta]\}$  where  $\{\hat{a}\}$  is the set of measured parameter estimates

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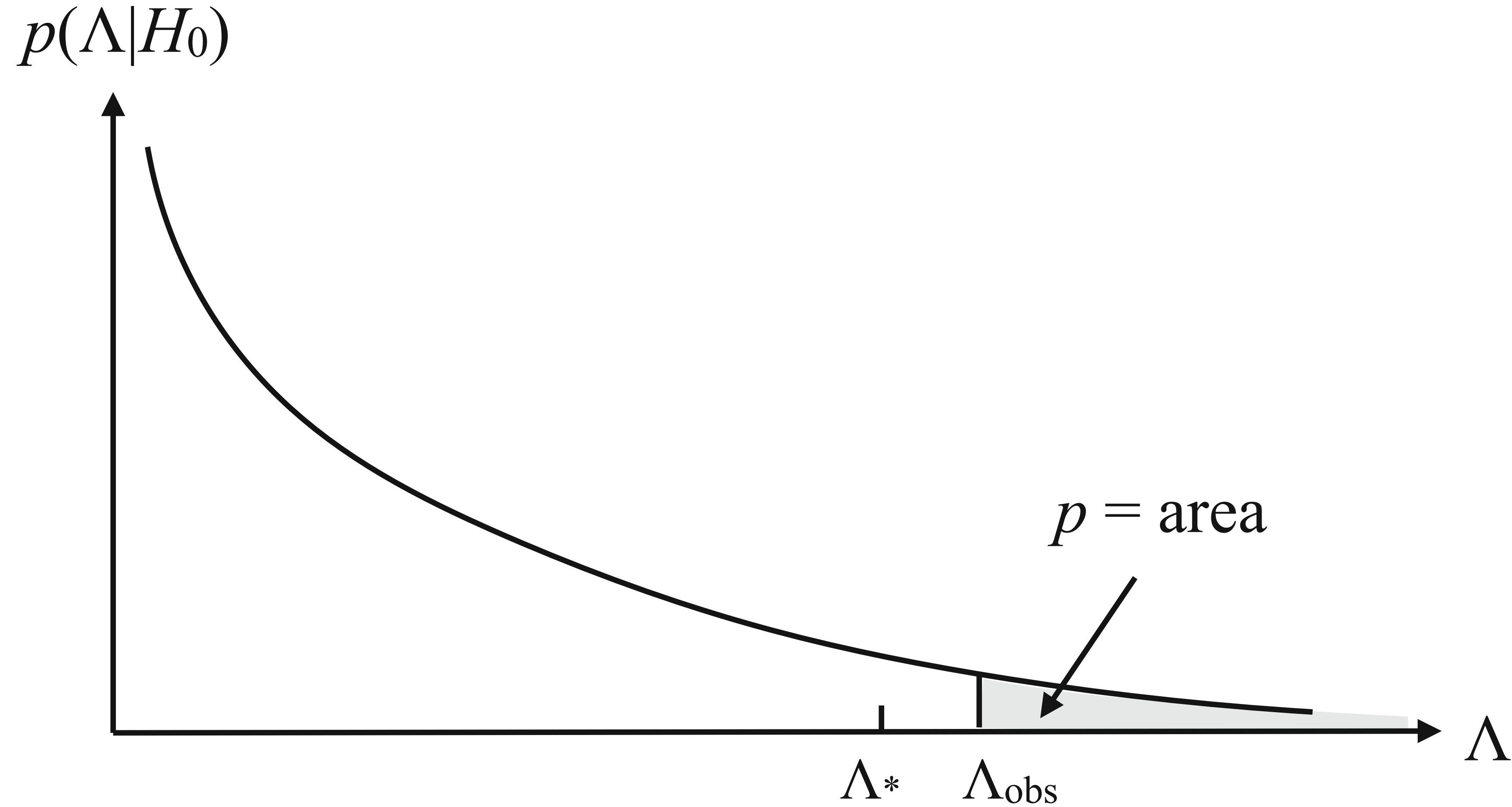
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- So you construct a **test statistic**  $\Lambda$  and calculate its sampling distributions  $p(\Lambda | H_0)$  and  $p(\Lambda | a, H_1)$  conditioned on  $H_0$  and  $H_1$
- If the observed value of  $\Lambda$  lies far out in the tail for the null distribution,  $p(\Lambda | H_0)$ , you reject  $H_0$  (accept  $H_1$ ) at the  $p \times 100\%$  level where  $p = \text{Prob}(\Lambda > \Lambda_{\text{obs}} | H_0)$  is the so-called  **$p$ -value**

# Frequentist p-value



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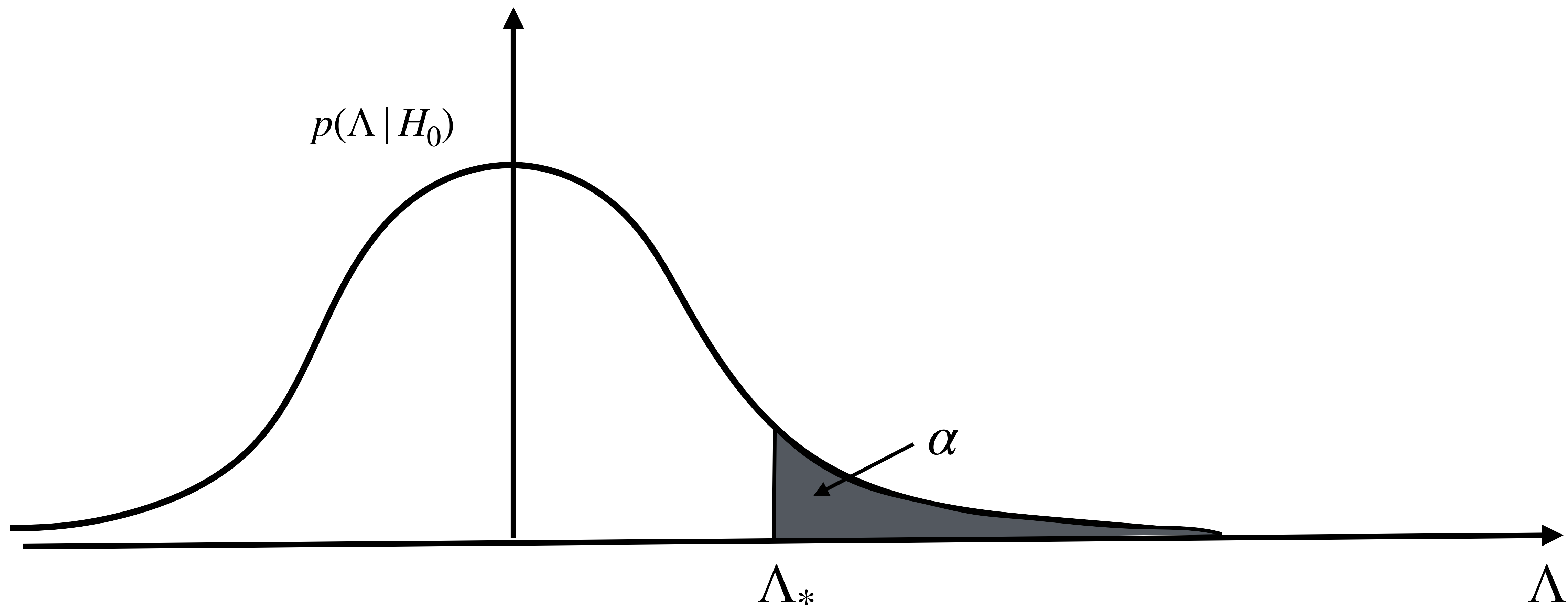
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- Different test statistics are **judged according to their false alarm and false dismissal probabilities**
- In GW data analysis, one typically sets the false alarm probability to some acceptably low level (e.g., 1 in 1000), then finds the test statistic that minimizes the false dismissal probability for fixed false alarm probability (called the **Neyman-Pearson criterion**)

# False alarm, false dismissal probabilities

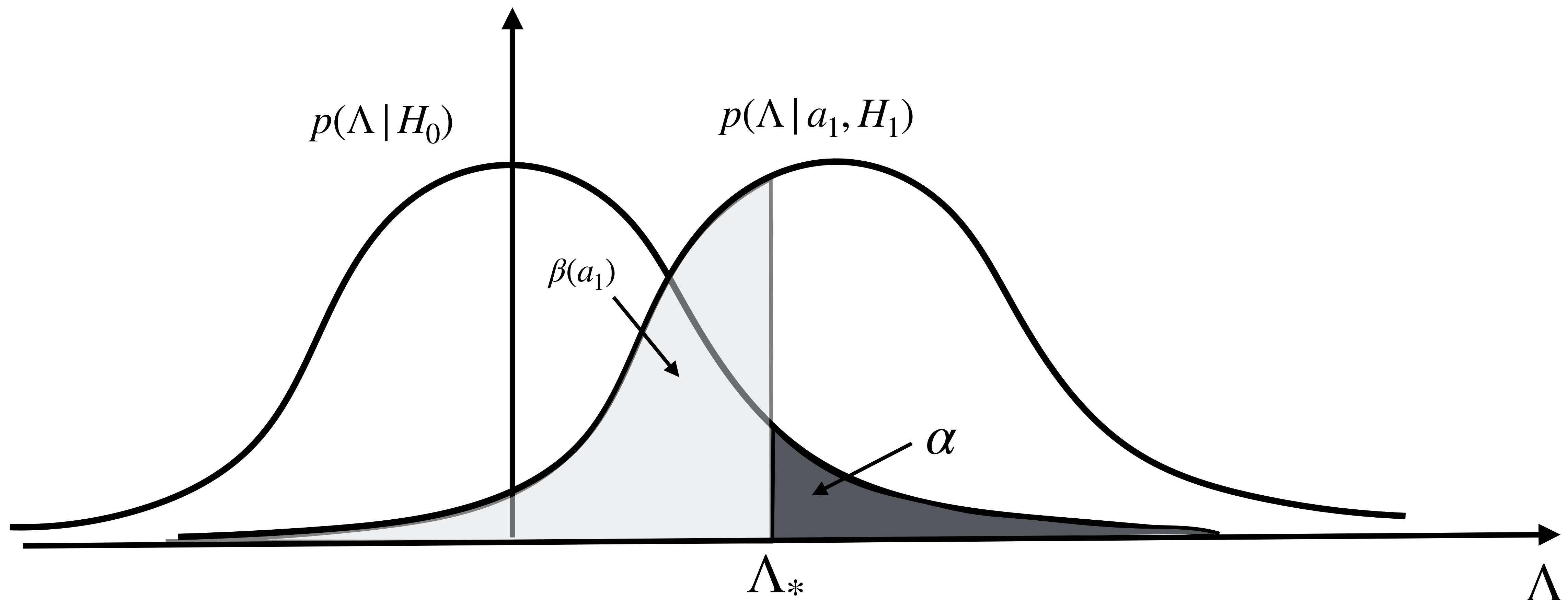
- $\alpha$  is the false alarm probability (refers to  $H_0$ ), e.g., 10%





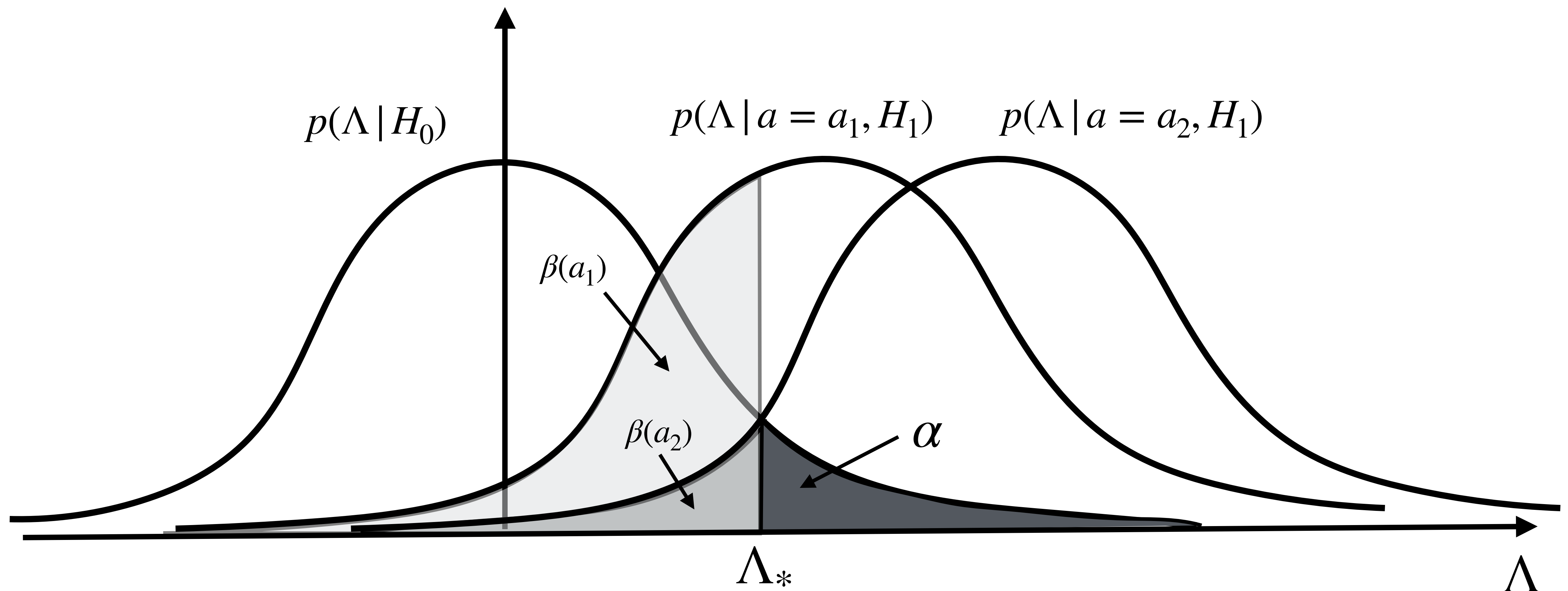
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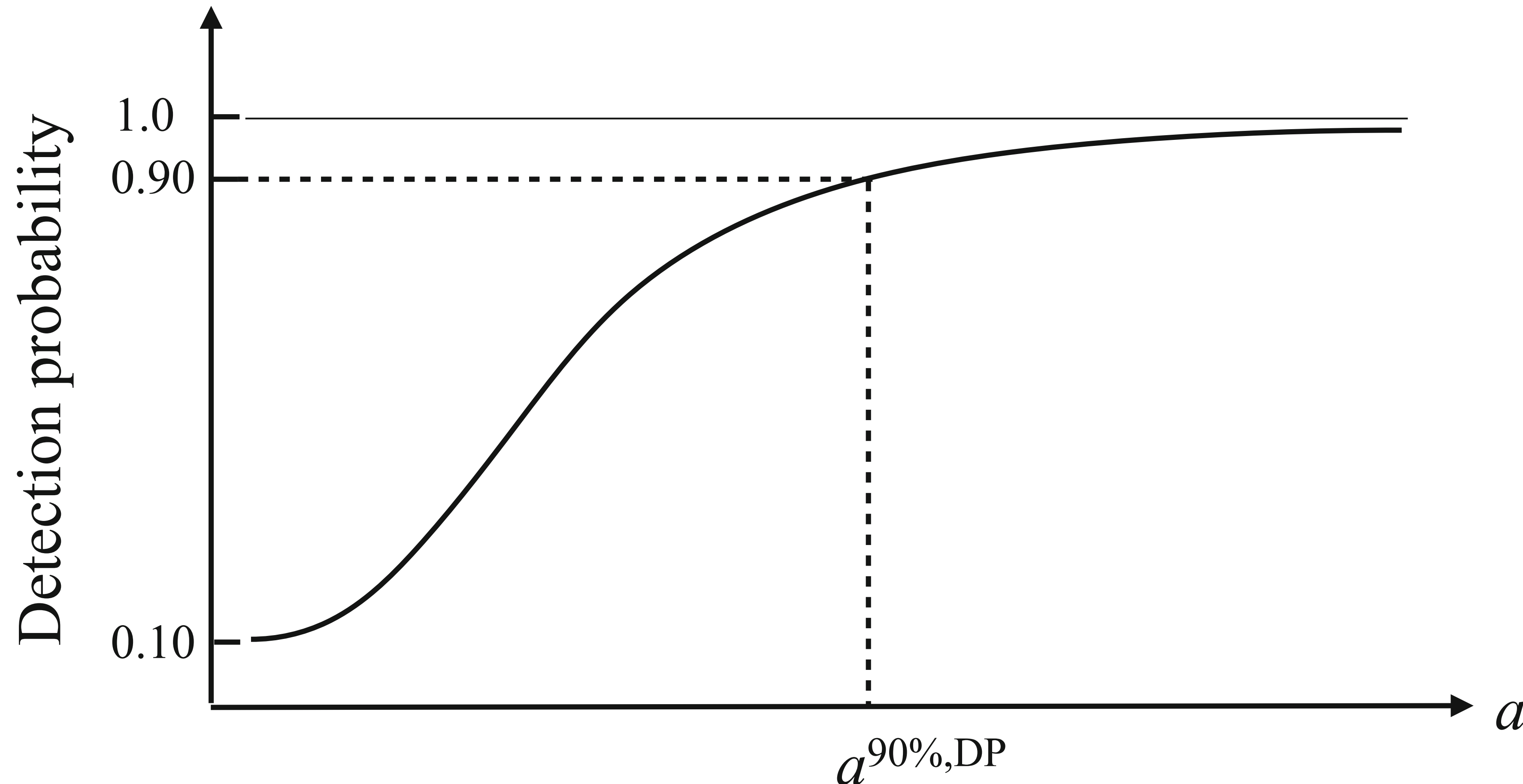
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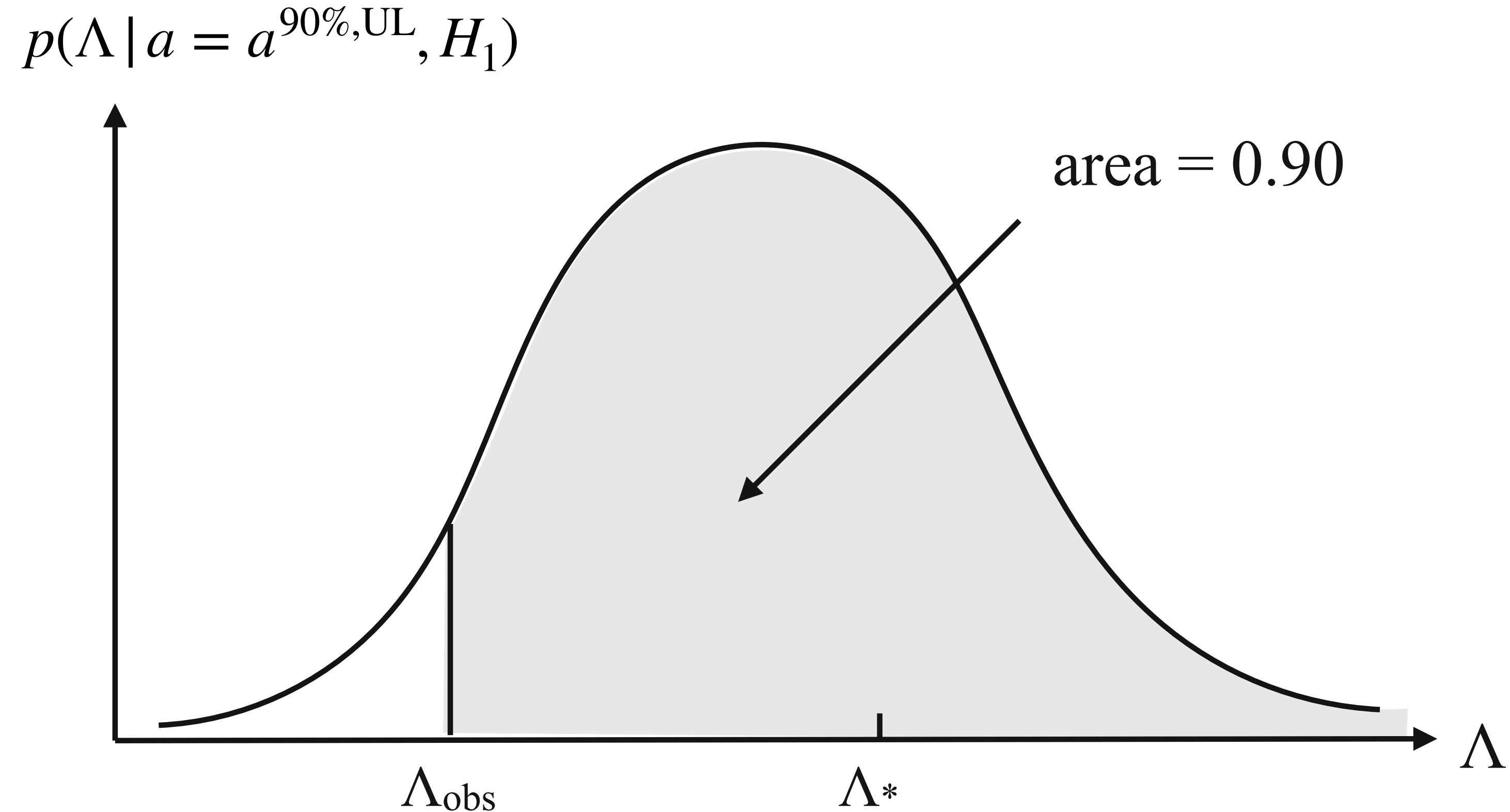
# Detection probability

- $\gamma(a) \equiv 1 - \beta(a)$  is the fraction of the time that the test statistic  $\Lambda$  **correctly identifies the presence of a signal** with amplitude  $a$



# Frequentist upper limits

- If  $\Lambda_{\text{obs}} < \Lambda_*$  one often sets an UL on the amplitude  $a$  of the signal
- $a^{90\%,\text{UL}}$  is the value of  $a$  for which  $\text{Prob}(\Lambda \geq \Lambda_{\text{obs}} | a = a^{90\%,\text{UL}}, H_1) = 0.90$



# III. Bayesian inference

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- If the posterior distribution depends on several parameters, you can obtain the posterior for one parameter by **marginalizing** over the others,

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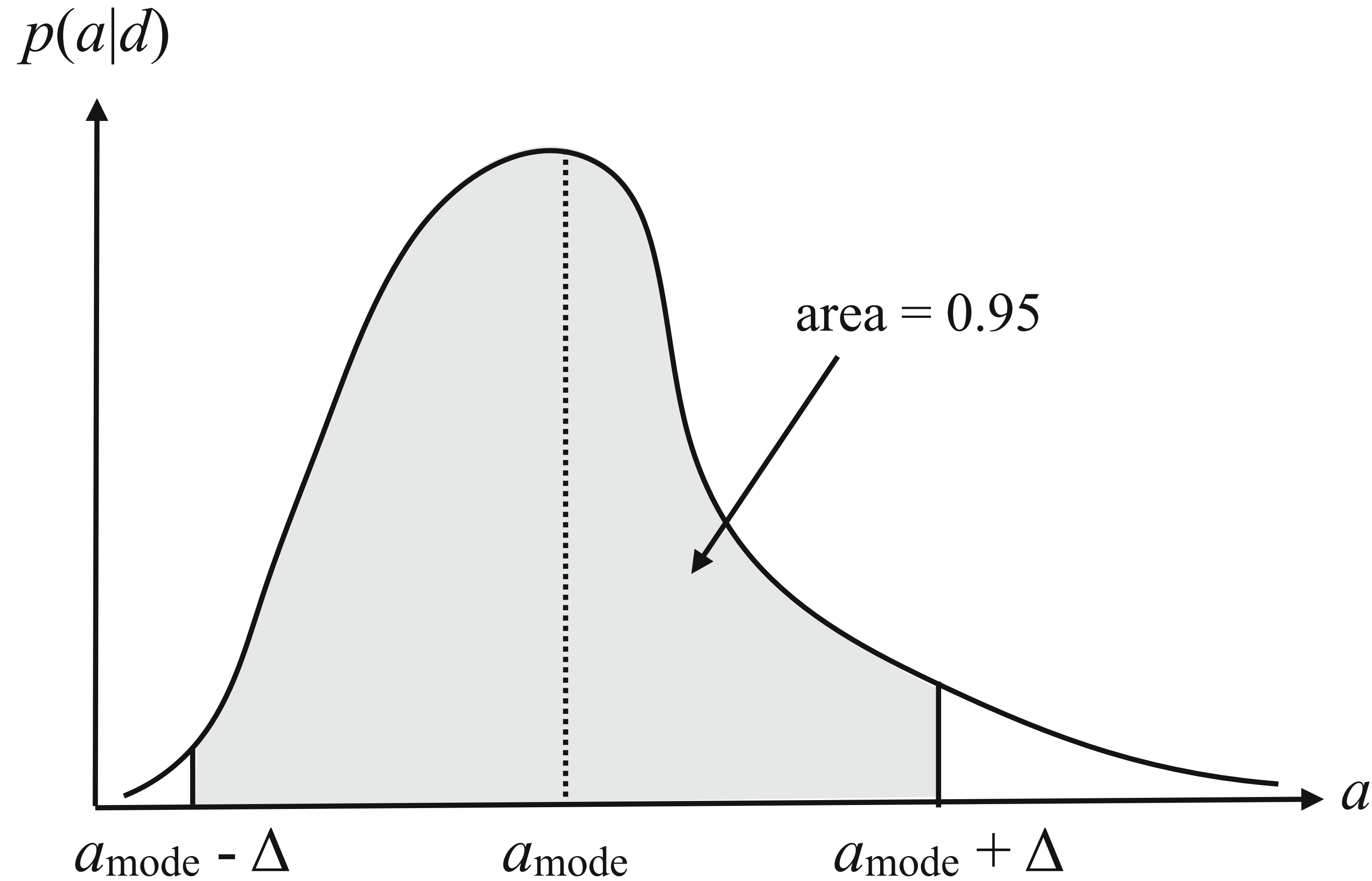
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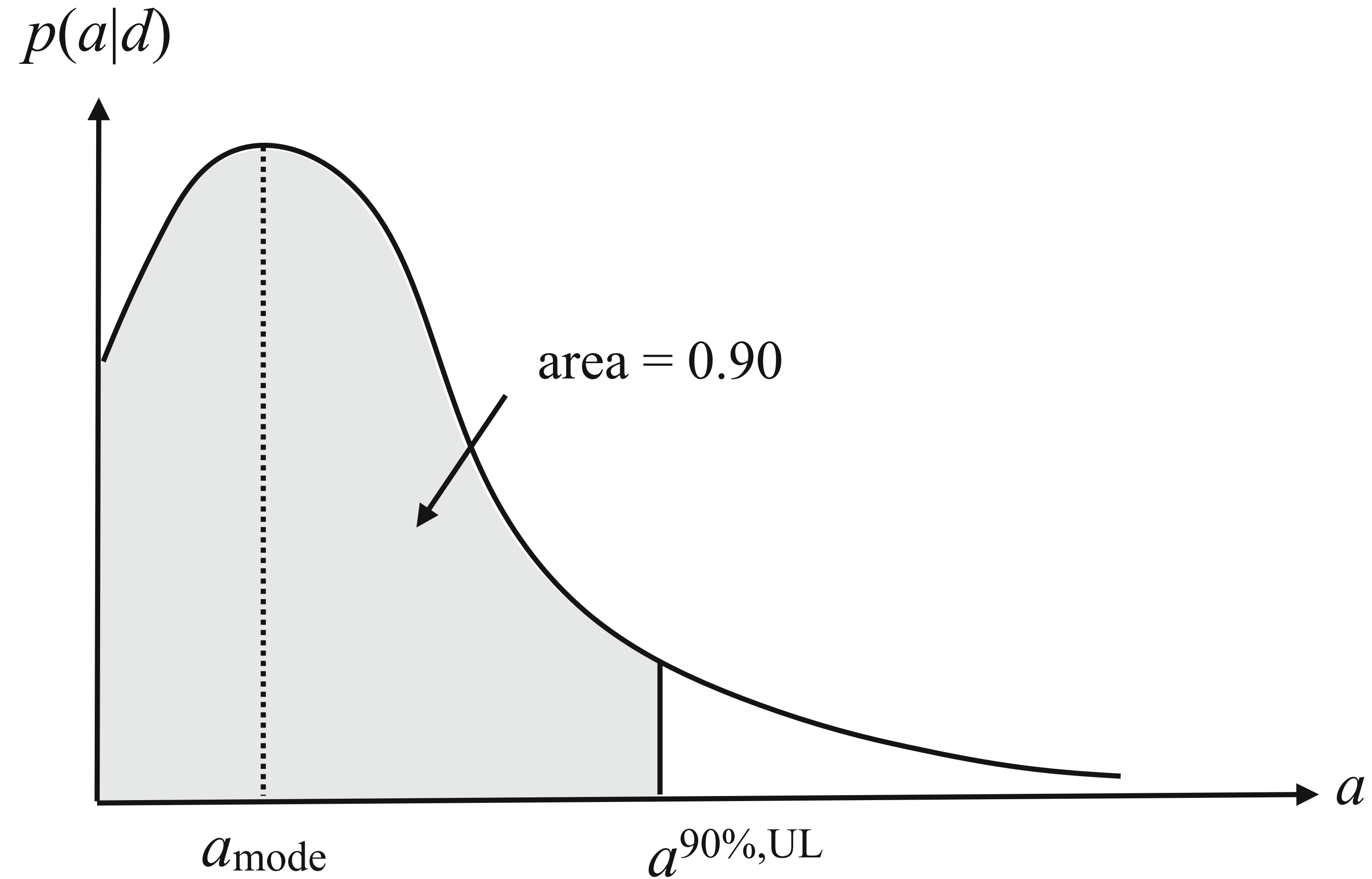
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- A Bayesian **credible interval** or **upper limit** defined in terms of the area under the posterior distribution

# Bayesian credible interval



# Bayesian credible upper limit



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$$\frac{p(H_1 | d)}{p(H_0 | d)} = \frac{p(d | H_1)}{p(d | H_0)} \frac{p(H_1)}{p(H_0)}$$

posterior odds      **Bayes factor**  $\mathcal{B}_{10}(d)$       prior odds  
(ratio of marginalized likelihoods or “evidences”)

# Relating Bayes factors and maximum-likelihood ratios



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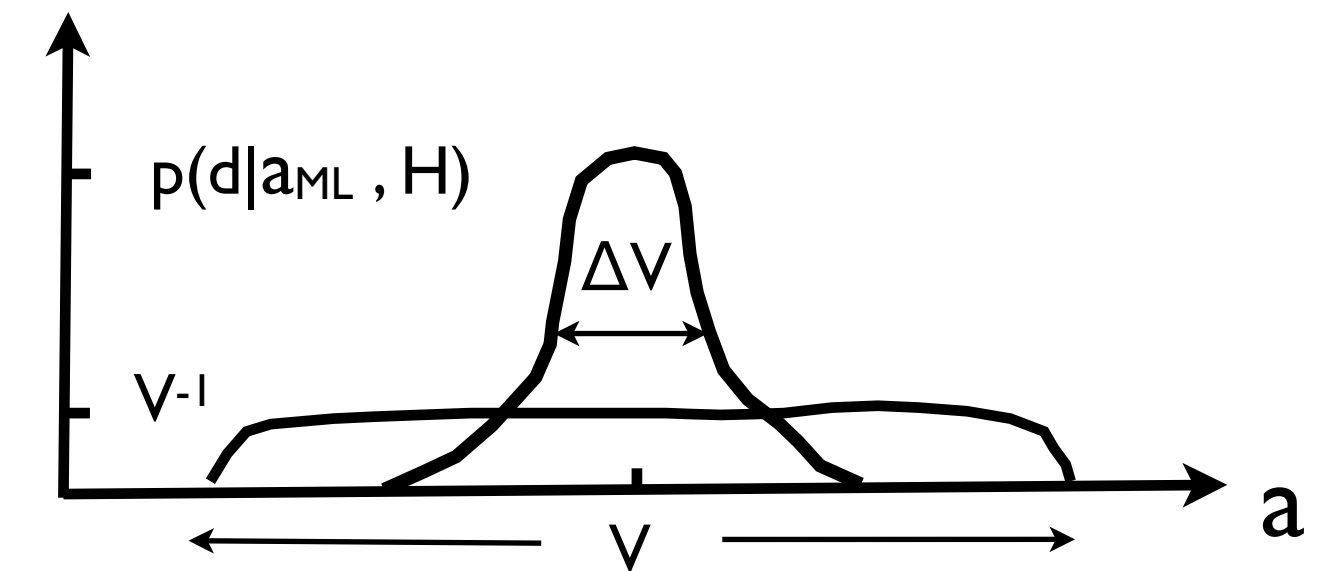
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$$p(d|H) \simeq p(d|a_{\text{ML}}, H)p(a_{\text{ML}}|H)\Delta a = \mathcal{L}_{\text{ML}}(d|H)\Delta V/V$$



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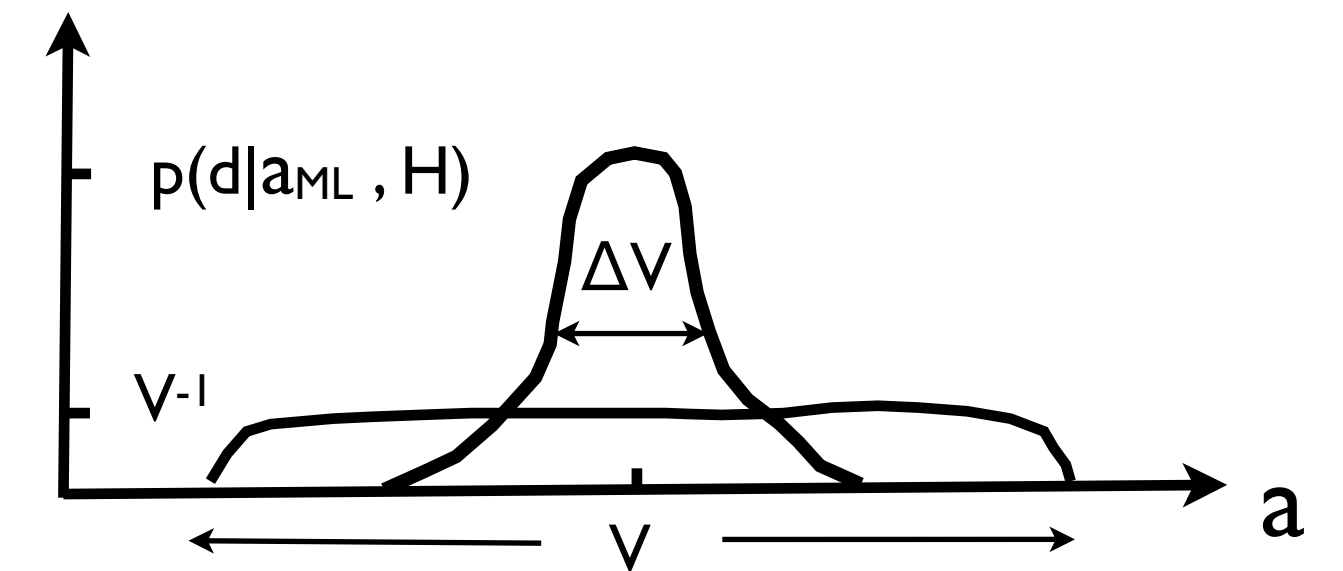
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- Bayes factor:

$$\mathcal{B}_{10}(d) \equiv \frac{p(d|H_1)}{p(d|H_0)} = \frac{\int da_1 \, p(d|a_1, H_1)p(a_1|H_1)}{\int da_0 \, p(d|a_0, H_0)p(a_0|H_0)} \simeq \Lambda_{\text{ML}}(d) \frac{\Delta V_1/V_1}{\Delta V_0/V_0}$$



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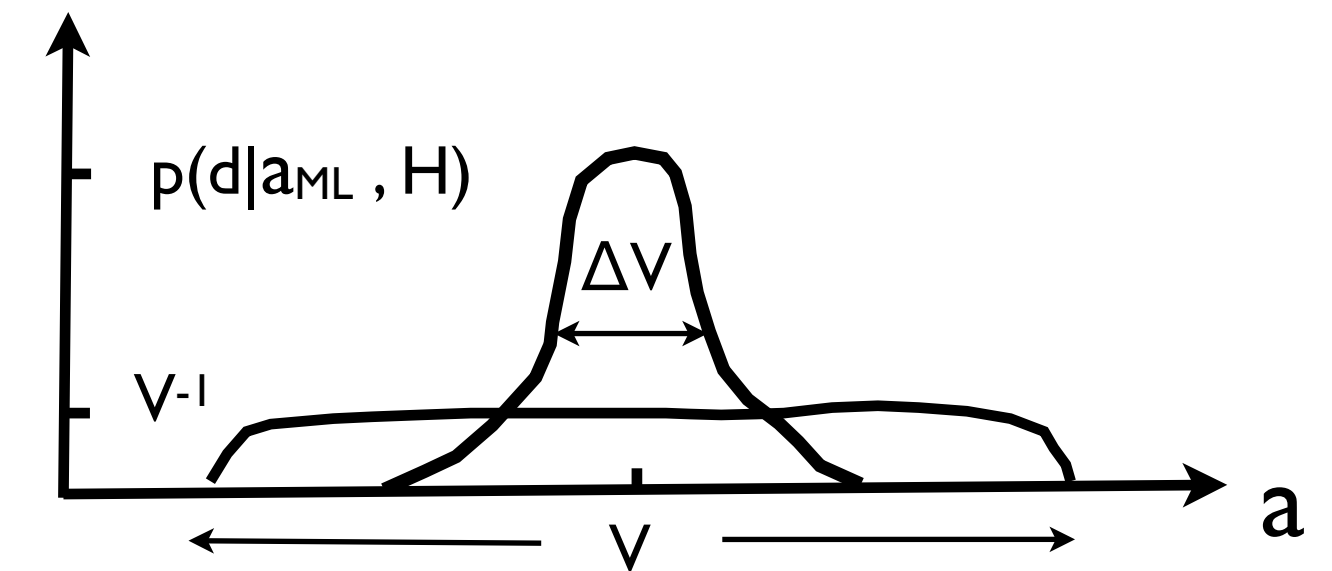
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- The  $\Delta V/V$  factors penalize hypotheses that uses more parameter space volume  $V$  than necessary to fit the data  $\Delta V$  (**Occam's penalty factor**)



# Significance of Bayes factor values

approximately equal to the  
squared SNR of the data



| $\mathcal{B}_{\alpha\beta}(d)$ | $2 \ln \mathcal{B}_{\alpha\beta}(d)$ | Evidence for model $\mathcal{M}_\alpha$ relative to $\mathcal{M}_\beta$ |
|--------------------------------|--------------------------------------|---|
| $<1$                           | $<0$                                 | Negative (supports model $\mathcal{M}_\beta$ )                          |
| 1–3                            | 0–2                                  | Not worth more than a bare mention                                      |
| 3–20                           | 2–6                                  | Positive  |
| 20–150                         | 6–10                                 | Strong  |
| $>150$                         | $>10$                                | Very strong   |

Adapted from Kass and Raftery (1995)

## IV. Exercises / worked examples

# 1. Practical application of Bayes' theorem

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- Suppose on your last visit to the doctor's office you took a test for some rare disease. This type of disease occurs in only 1 out of 10,000 people, as determined by a random sample of the population. The test that you took is rather effective in that it can correctly identify the presence of the disease 95% of the time, but it gives false positives 1% of the time.



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- Suppose the test came up positive. What is the probability that you have the disease?

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- Information:

$$P(H) = 0.0001 \quad P(\bar{H}) = 0.9999$$

$$P(+ | H) = 0.95 \quad P(+ | \bar{H}) = 0.01$$

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- Calculate:

$$P(H | +) = \frac{P(+ | H)P(H)}{P(+)}$$

$$\begin{aligned} P(+) &= P(+ | H)P(H) + P(+ | \bar{H})P(\bar{H}) \\ &= 0.95 \times 0.0001 + 0.01 \times 0.9999 \\ &\approx 0.01 \end{aligned}$$

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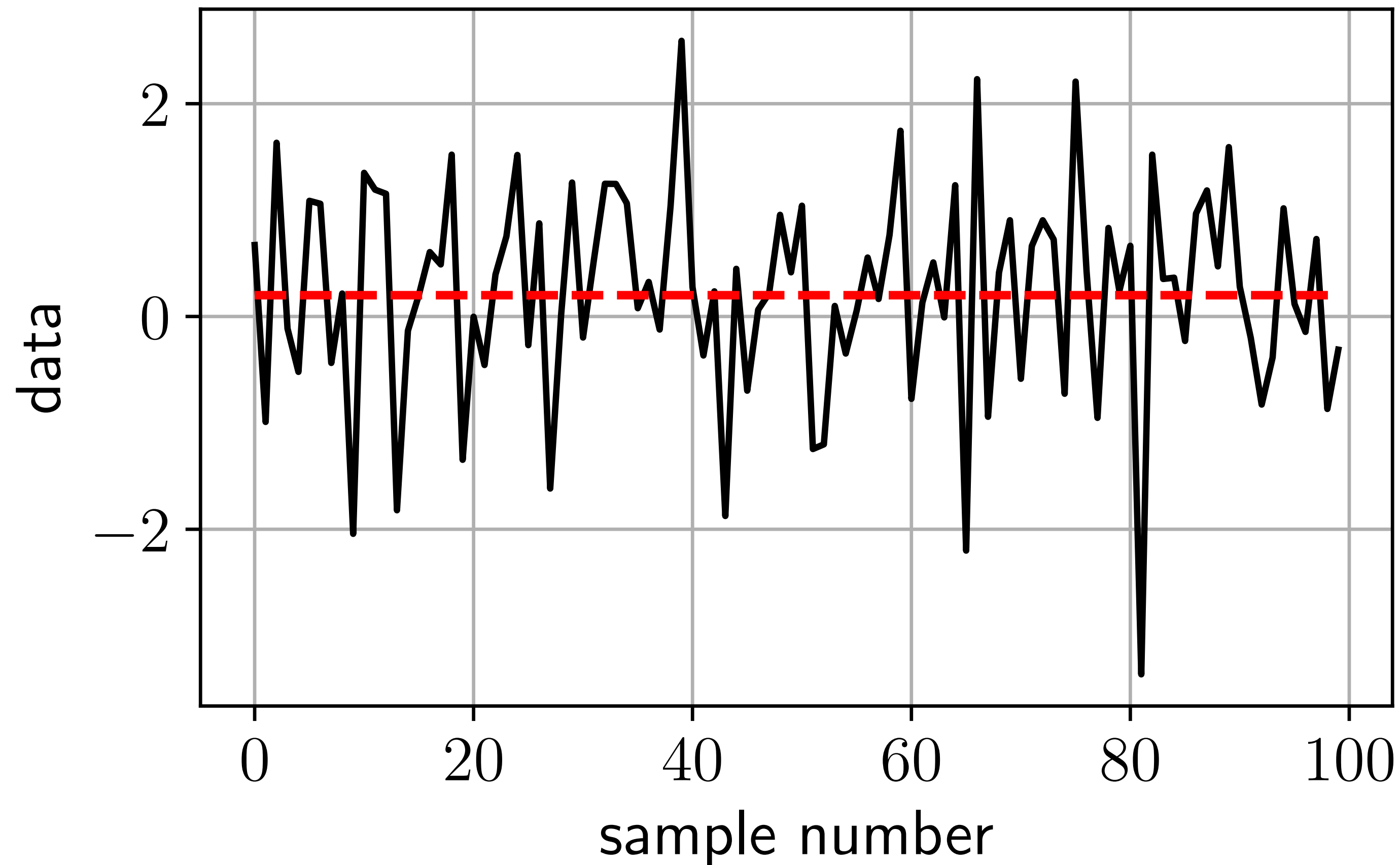
$$P(H | +) = \frac{P(+ | H)P(H)}{P(+)}$$

$$\begin{aligned} P(+) &= P(+ | H)P(H) + P(+ | \bar{H})P(\bar{H}) \\ &= 0.95 \times 0.0001 + 0.01 \times 0.9999 \\ &\approx 0.01 \end{aligned}$$

- Final result:

$$P(H | +) \approx 0.0095 \approx 0.01$$

## 2. Comparing frequentist and Bayesian analyses for a constant amplitude signal in white noise



# Key formulae

Likelihoods functions:

$$p(d | \mathcal{M}_0) = \left( \frac{1}{\sqrt{2\pi}\sigma} \right)^N \exp \left[ -\frac{1}{2\sigma^2} \sum_{i=1}^N d_i^2 \right]$$

$$p(d | a, \mathcal{M}_1) = \left( \frac{1}{\sqrt{2\pi}\sigma} \right)^N \exp \left[ -\frac{1}{2\sigma^2} \sum_{i=1}^N (d_i - a)^2 \right]$$

Prior:

$$p(a | \mathcal{M}_1) = \frac{1}{a_{\max}}$$

Parameter choices:

$$N = 100, \quad \sigma = 1, \quad 0 \leq a \leq a_{\max}, \quad a_0 = \text{true value}$$



# Key formulae

Maximum-likelihood estimator:

$$\hat{a} \equiv a_{\text{ML}}(d) = \frac{1}{N} \sum_{i=1}^N d_i \equiv \bar{d} \quad \sigma_{\hat{a}}^2 = \frac{\sigma^2}{N}$$

Useful identity:

$$\sum_{i=1}^N (d_i - a)^2 = \sum_i d_i^2 - N\hat{a}^2 + N(a - \hat{a})^2 = N (\text{Var}[d] + (a - \hat{a})^2)$$

Likelihood function (in terms of ML estimator):

$$p(d | a, \mathcal{M}_1) = \left( \frac{1}{\sqrt{2\pi}\sigma} \right)^N \exp \left[ -\frac{\text{Var}[d]}{2\sigma_{\hat{a}}^2} \right] \exp \left[ -\frac{(a - \hat{a})^2}{2\sigma_{\hat{a}}^2} \right]$$

Evidence:

$$p(d | \mathcal{M}_1) = \frac{\exp \left[ -\frac{\text{Var}[d]}{2\sigma_{\hat{a}}^2} \right] \left[ \text{erf} \left( \frac{a_{\text{max}} - \hat{a}}{\sqrt{2}\sigma_{\hat{a}}} \right) + \text{erf} \left( \frac{\hat{a}}{\sqrt{2}\sigma_{\hat{a}}} \right) \right]}{2a_{\text{max}} \left( \sqrt{2\pi}\sigma \right)^{N-1} \sqrt{N}}$$

Posterior distribution:

$$p(a | d, \mathcal{M}_1) = \frac{1}{\sqrt{2\pi}\sigma_{\hat{a}}} \exp \left[ -\frac{(a - \hat{a})^2}{2\sigma_{\hat{a}}^2} \right] 2 \left[ \text{erf} \left( \frac{a_{\text{max}} - \hat{a}}{\sqrt{2}\sigma_{\hat{a}}} \right) + \text{erf} \left( \frac{\hat{a}}{\sqrt{2}\sigma_{\hat{a}}} \right) \right]^{-1}$$

# Key formulae

Bayes factor:

$$\mathcal{B}_{10}(d) = \exp \left[ \frac{\hat{a}^2}{2\sigma_{\hat{a}}^2} \right] \left( \frac{\sqrt{2\pi}\sigma_{\hat{a}}}{a_{\max}} \right) \frac{1}{2} \left[ \operatorname{erf} \left( \frac{a_{\max} - \hat{a}}{\sqrt{2}\sigma_{\hat{a}}} \right) + \operatorname{erf} \left( \frac{\hat{a}}{\sqrt{2}\sigma_{\hat{a}}} \right) \right] \simeq \exp \left[ \frac{\hat{a}^2}{2\sigma_{\hat{a}}^2} \right] \left( \frac{\sqrt{2\pi}\sigma_{\hat{a}}}{a_{\max}} \right)$$

Maximum likelihood ratio statistic:

$$\Lambda_{\text{ML}}(d) = \exp \left( \frac{\hat{a}^2}{2\sigma_{\hat{a}}^2} \right)$$

Frequentist test statistic:

$$\Lambda(d) \equiv 2 \ln \Lambda_{\text{ML}}(d) = \frac{\hat{a}^2}{\sigma_{\hat{a}}^2} = \left( \frac{\sqrt{N}d}{\sigma} \right)^2 \equiv \rho^2$$

Sampling distributions of the test statistic:

$$p(\Lambda | \mathcal{M}_0) = \frac{1}{\sqrt{2\pi\Lambda}} e^{-\Lambda/2}$$

$$p(\Lambda | a, \mathcal{M}_1) = \frac{1}{\sqrt{2\pi\Lambda}} \frac{1}{2} \left[ e^{-\frac{1}{2}(\sqrt{\Lambda} - \sqrt{\lambda})^2} + e^{-\frac{1}{2}(\sqrt{\Lambda} + \sqrt{\lambda})^2} \right]$$

$$\lambda = \langle \rho \rangle^2 = \frac{Na^2}{\sigma^2}$$

**See romano\_notes1.pdf and romano\_code1.ipynb for solutions**