SD-TSIA204 Statistics : linear models

Joseph Salmon

http://josephsalmon.eu Télécom ParisTech, Institut Mines-Télécom

Teachers

- Joseph Salmon (Assistant Professor) :
 - Positions : Paris Diderot-Paris 7, Duke University, Télécom ParisTech
 - Research themes: high dimensional statistics, aggregation, image processing, optimization
 - Email : joseph.salmon@telecom-paristech.fr
 - Office : E410
- Francois Portier (Assistant Professor) :
 - Positions : Université catholique de Louvain, Université de Rennes 1, Télécom ParisTech
 - Research themes: sparse regression, bootstrap, semi-parametric statistic, kernel smoothing
 - Email : fportier@enst.fr
 - Office : E 302

Teaching assistants

- Gower Robert (Assistant Professor): Optimization for machine learning, randomized numerical linear algebra, variable metric and quasi-Newton methods
- Giovanna Varni (Assistant Professor): Social Signal Processing (SSP), human-human/human-computer/human-robot interaction, expressive gesture, and interpersonal synchrony
- Garcia Alexandre (PhD Student): semi-supervised prediction for structured variables, sentiment analysis
- Korba Anna (PhD Student): ranking, preference learning, voting rules, recommender systems
- Ndiaye Eugene (PhD Student): sparse estimators, optimization with coordinates descent methods, safe screening rules

Grades: SD-TSIA204

- Practical 1 : 20% final grade
 - Problem statement available on Wednesday 06/12/2017 at 1:
 30 PM, deadline 23: 59 the same day.
 - Single accepted file : IPython Notebook
- Practical 2 : 20% final grade
 - Problem statement available on Wednesday 10/01/2018 at 3:
 15 PM, deadline 23: 59 the same day.
 - ► Single accepted file : IPython Notebook
- Final exam : 60% note finale
 - ▶ Date: 24/01/2018
 - Format : Quiz + exercises

<u>Rem</u>:Quiz questions samples are available on the course website (cf. Liste de questions section)

BEWARE : your practical should be personal not copy pasted from your neighbor!!!

Practical notation

Practicals are graded on a scale from 0 to 20, as follows

- scientific quality of answer 15 pts
- ▶ language/writing quality of answer (spelling, etc.) 2 pts
- indentation, PEP8 Style, useful comments in code, no/few warnings 2 pts
- no bug 1 pt (at least https://github.com/agramfort/check_notebook)
- one single .ipynb file expected, submitted on the "Site pédagogique" of the course; emailed work will receive a zero score and will not be graded

<u>Late</u>: **no Late work** work will be accepted, unless official reason accepted by Télécom ParisTech's administration; late work will receive a zero score and will not be graded

Bonus

1 pt out of the <u>final grade</u> could be rewarded for any useful contribution to improve this course (slides, codes, etc.)

Constraints:

- only the first contribution received on one aspect will be rewarded
- upload a .txt file for each proposition the "Propositions d'amélioration" section of the Site pédagogique
- detail precisely (code lines, slide page, etc.) the contribution proposed, and what it is improving/solving
- For typos: a minimum number of 5 typos is required to get the 1pt bonus

Outline of the course

```
Course 1 Joseph Salmon (08/11): Introduction
Course 2 François Portier (29/11): Least squares properties
Course 3 A. K./R. G./A. G./E. N./F. P./J. S. (06/12):
          Practical 1 Regression / Regularization
Course 4 Francois Portier (13/12) : CI / Bootstrap
Course 5 Francois Portier (20/12) : Ridge/Greedy/Lasso
Course 6 Joseph Salmon (10/01): PCA/SVD
Course 7 G. V./A. K./A. G./E. N./F. P. /J. S. (10/01) :
          Practical 2 (PCA/SVD)
Course 8 Joseph Salmon (17/01): GLM/Logistic regression
Course 9 (24/01): Final Exam
```

- Probability basis: probability, expectation, law of large number, Gaussian distribution, central limit theorem.
 Books: Foata et Fuchs (1996) (in French) or Murphy (2012, ch.1 and 2)
- Optimisation basis: (differential) calculus, convexity, first order conditions, gradient descent, Newton method Lecture: Boyd et Vandenberghe (2004), Bertsekas (1999)

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 Lecture: Golub et VanLoan (2013), Applied Numerical Computing par L. Vandenberghe

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Algorithmic aspects : some advice

Python installation : use Conda / Anaconda

Rem:you are on your own for this (or use the school machines)

Recommended tools: **Jupyter / IPython Notebook** (mandatory for your practical) **IPython** with a text editor *e.g.*,**Atom**, **Sublime Text**, **Visual Studio Code**, etc., for larger projects

- Python, Scipy, Numpy :
 - http://perso.telecom-paristech.fr/~gramfort/liesse_python/
- Pandas : http://pandas.pydata.org/
- scikit-learn : http://scikit-learn.org/stable/

<u>Rem</u>: for practicals, bring your own machine if you prefer but install your Python environment upfront

General advice

- Use version control system for your work : Git (e.g.,Bitbucket, Github, etc.)
- Use clean way of writing code/ presenting your code <u>Example</u>: **PEP8** for Python (use for instance **AutoPEP8**, <u>https://github.com/kenkoooo/jupyter-autopep8</u>)
- Learn from good examples :
 https://github.com/scikit-learn/,
 http://jakevdp.github.io/, etc.

Outline

Syllabus, grades, etc.

1D Least squares

Table of Contents

Syllabus, grades, etc.

1D Least squares

Introduction: visualization / Python

Modeling

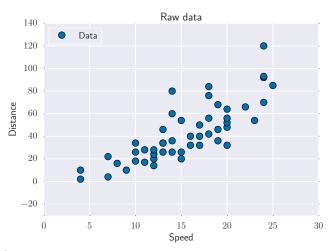
Mathematical Formulation

Centering - scaling

Likelihood

A 2D starting example

 $\frac{\text{Example}}{(n=50 \text{ measurements})} : \text{braking distance for cars as a function of speed}$



Dataset cars: https://stat.ethz.ch/R-manual/R-devel/library/datasets/html/cars.html https://stat.ethz.ch/R-manual/R-devel/library/datasets/html/cars.html

A 2D starting example

 $\frac{\text{Example}}{(n=50)} = \text{braking distance for cars as a function of speed}$



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Python command

```
import pandas as pd
import matplotlib.pyplot as plt
import sklearn.linear_model as lm
# Load data
url = 'cars.csv'
dat = pd.read csv(url)
y = dat['dist']
X = dat[['speed']] # sklearn needs X to have 2 dim.
skl_linmod = lm.LinearRegression(fit_intercept=False)
skl linmod.fit(X, y) # Fit regression model
fig = plt.figure(figsize=(8, 6))
plt.plot(X, y, 'o', label="Data")
plt.plot(X, skl_linmod.predict(X),
        label="OLS-sklearn-no-intercept")
plt.legend(loc='upper left')
plt.show()
```

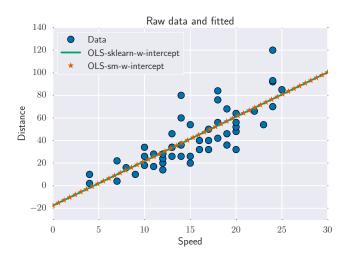
A 2D starting example : with an intercept

 $\underline{\underline{\mathsf{Example}}}$: braking distance for cars as a function of speed (n=50 measurements)



A 2D starting example : with an intercept

 $\underline{\underline{\text{Example}}} : \text{braking distance for cars as a function of speed } (n = 50)$



Python commands: with intercept

```
import statsmodels.api as sm
# data, fitted. etc
y = dat['dist']
X = dat[['speed']]
X = sm.add constant(X)
results = sm.OLS(y,X).fit()
# plot
fig, ax = plt.subplots(figsize=(8,6))
ax.plot(X['speed'], y, 'o', label="data")
ax.plot(X['speed'], results.fittedvalues,
       linewidth=3, label="OLS-sm-w-intercept")
ax.legend(loc='best')
```

<u>Alternative</u>: use lm.LinearRegression(fit_intercept=True)

Modeling I

Observations:
$$(y_i, x_i)$$
, for $i = 1, ..., n$

Linear model or linear regression hypothesis assume :

$$y_i \approx \theta_0^{\star} + \theta_1^{\star} x_i$$

- θ_0^{\star} : intercept (unknown)
- θ_1^{\star} : slope (unknown)

Rem: both parameters are unknown from the statistician

Definition

- y is an observation or a variable to explain
- x is a **feature** or a covariate

Notation interpretation

Example : dataset cars

- n = 50
- y_i : braking time for i-th car
- x_i : speed of *i*-th car
- y: the observation is the car's braking time
- x: the feature/covariate is the car's speed

Linear model / Linear regression hypothesis : assume that braking time is proportional to speed

Exo: use describe() from Pandas to get a rough data summary

Modeling II

Let us give a precise meaning to the sign \approx :

Probabilistic model

$$y_{i} = \theta_{0}^{\star} + \theta_{1}^{\star} x_{i} + \varepsilon_{i},$$

$$\varepsilon_{i} \stackrel{i.i.d}{\sim} \varepsilon, \text{ for } i = 1, \dots, n$$

$$\mathbb{E}(\varepsilon) = 0$$

where i.i.d. means "independent and identically distributed"

Interpretation

 $\varepsilon_i = y_i - \theta_0^\star - \theta_1^\star x_i$: represent the error between the theoretical model and the observations, represented by random variables ε_i centered (often referred to as white noise).

<u>Rem</u>: motivation for the random nature of the noise – measurement noise, transmission noise, in-population variability, etc.

Modeling III

$$y_i = \theta_0^{\star} + \theta_1^{\star} x_i + \varepsilon_i$$

Definition

We call

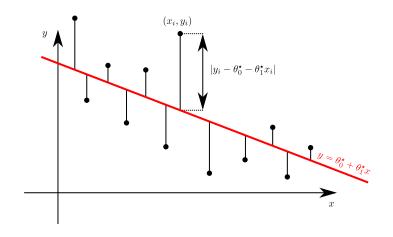
- intercept the scalar θ_0^{\star} (\blacksquare : ordonnée à l'origine)
- ▶ slope the scalar θ_1^* (■ : pente)

Goal

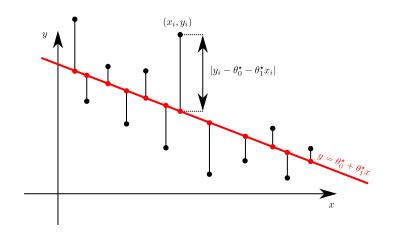
Estimate θ_0^\star and θ_1^\star (unknown) by $\hat{\theta}_0$ and $\hat{\theta}_1$ relying on observations (y_i, x_i) for $i = 1, \ldots, n$

<u>Rem</u>:The "hat" notation is classical in statistics for referring to estimators

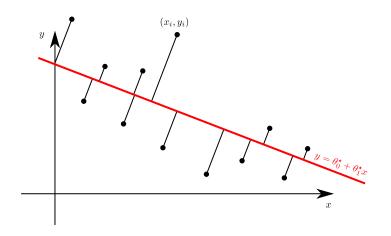
Least squares: visualization



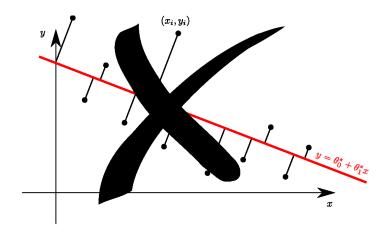
Least squares: visualization



(Total) Least squares : visualization



(Total) Least squares : visualization



Least squares: mathematical formulation

Definition

The **least squares** estimator is defined as :

$$(\hat{\theta}_0, \hat{\theta}_1) \in \underset{(\theta_0, \theta_1) \in \mathbb{R}^2}{\arg \min} \sum_{i=1}^n (y_i - \theta_0 - \theta_1 x_i)^2$$

- ▶ it is also referred to as "ordinary least squares" (OLS)
- an original motivation for the squares is computational : first order conditions only require solving a linear system
- a solution always exists : minimizing a coercive continuous function (coercive : $\lim_{\|x\|\to+\infty} f(x) = +\infty$)

Rem: write $\alpha \in \arg \min$ » as long as you do not know if the solution is unique

Least square authorship (controversial)



Adrien-Marie Legendre : "Nouvelles méthodes pour la détermination des orbites des comètes", 1805



Carl Friedrich Gauss:
"Theoria Motus Corporum Coelestium
in sectionibus conicis solem
ambientium" 1809

Historical / robust detour

Definition

The least absolute deviation (LAD) estimator reads :

$$(\hat{\theta}_0, \hat{\theta}_1) \in \underset{(\theta_0, \theta_1) \in \mathbb{R}^2}{\arg \min} \sum_{i=1}^n |y_i - \theta_0 - \theta_1 x_i|$$

<u>Rem</u>: hard to compute without computer; requires an optimization solver for non-smooth function (or a Linear Programming solver)

Rem: more robust to outliers (■ : données aberrantes)

Least absolute deviation authorship



Ruđer Josip Bošković :"???", 1757

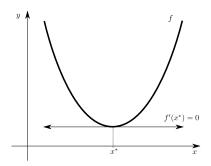


Pierre-Simon de Laplace "Traité de mécanique céleste", 1799

Local minimum: first order condition

Theorem: Fermat's rule

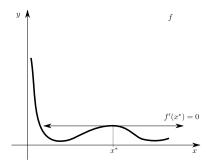
If f is differentiable, then at a local minimum x^* the gradient of f vanishes at x^* , i.e., $\nabla f(x^*)=0$.



Local minimum: first order condition

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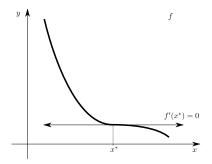


Rem: sufficient condition when f is convex!

Local minimum: first order condition

Theorem: Fermat's rule

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Rem: sufficient condition when f is convex!

Back to least squares

$$\hat{\boldsymbol{\theta}} = (\hat{\theta}_0, \hat{\theta}_1) \in \underset{(\theta_0, \theta_1) \in \mathbb{R}^2}{\operatorname{arg \, min}} \frac{1}{2} \sum_{i=1}^n (y_i - \theta_0 - \theta_1 x_i)^2$$

For least squares, minimize the function of two variables :

$$f(\theta_0, \theta_1) = f(\boldsymbol{\theta}) = \frac{1}{2} \sum_{i=1}^{n} (y_i - \theta_0 - \theta_1 x_i)^2$$

First order condition / Fermat's rule :

$$\begin{cases} \frac{\partial f}{\partial \theta_0}(\hat{\boldsymbol{\theta}}) = \sum_{i=1}^n (y_i - \hat{\theta}_0 - \hat{\theta}_1 x_i) = 0\\ \frac{\partial f}{\partial \theta_1}(\hat{\boldsymbol{\theta}}) = \sum_{i=1}^n (y_i - \hat{\theta}_0 - \hat{\theta}_1 x_i) x_i = 0 \end{cases}$$

Exo: Is f convex? help: a sum of convex functions is convex

Calculus continued

Usual mean notation :
$$\overline{x}_n=\frac{1}{n}\sum_{i=1}^n x_i$$
 and $\overline{y}_n=\frac{1}{n}\sum_{i=1}^n y_i$

With that, Fermat's rule states (dividing by n):

$$\begin{cases} \frac{\partial f}{\partial \hat{\theta}_0}(\hat{\boldsymbol{\theta}}) = \sum_{i=1}^n (y_i - \hat{\theta}_0 - \hat{\theta}_1 x_i) = 0 \\ \frac{\partial f}{\partial \hat{\theta}_1}(\hat{\boldsymbol{\theta}}) = \sum_{i=1}^n (y_i - \hat{\theta}_0 - \hat{\theta}_1 x_i) x_i = 0 \end{cases}$$

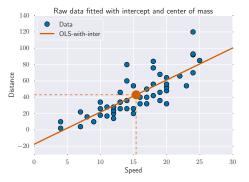
$$\Leftrightarrow$$

$$\begin{cases} \hat{\theta}_0 = \overline{y}_n - \hat{\theta}_1 \overline{x}_n & \text{(CNO1)} \\ \hat{\theta}_1 = \frac{\sum_{i=1}^n (x_i - \overline{x}_n)(y_i - \overline{y}_n)}{\sum_{i=1}^n (x_i - \overline{x}_n)^2} & \text{(CNO2)} \end{cases}$$

Exo: Prove that (CNO2) holds if and only if $\mathbf{x} = (x_1, \dots, x_n)^{\top}$ is non constant, *i.e.*, \mathbf{x} is not proportional to $\mathbf{1}_n = (1, \dots, 1)^{\top} \in \mathbb{R}^n$

Center of gravity and interpretation

(CNO1)
$$\Leftrightarrow (\overline{x}_n, \overline{y}_n) \in \{(x, y) \in \mathbb{R}^2 : y = \hat{\theta}_0 + \hat{\theta}_1 x\}$$



- $ightharpoonup \overline{speed} = 15.4$
- $\overline{dist} = 42.98$
- $\hat{\theta}_0 = -17.579095$ intercept (negative!)
- $\hat{\theta}_1 = 3.932409 \text{ slope}$

Physical interpretation: the cloud of points' center of gravity belongs to the (estimated) regression line

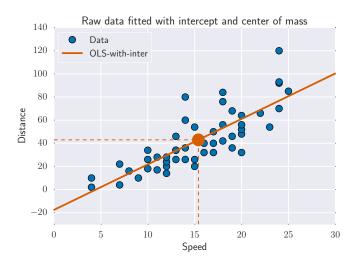
Vector formulation

$$\begin{split} & \underline{\text{Notation}}: \quad \mathbf{x} = (x_1, \dots, x_n)^\top \text{ and } \mathbf{y} = (y_1, \dots, y_n)^\top \\ & \qquad \qquad (\mathsf{CNO2}) \Leftrightarrow \hat{\theta}_1 = \frac{\sum_{i=1}^n (x_i - \overline{x}_n)(y_i - \overline{y}_n)}{\sum_{i=1}^n (x_i - \overline{x}_n)^2} \\ & \qquad \qquad (\mathsf{CNO2}) \Leftrightarrow \hat{\theta}_1 = \mathrm{corr}_n(\mathbf{x}, \mathbf{y}) \cdot \frac{\sqrt{\mathrm{var}_n(\mathbf{y})}}{\sqrt{\mathrm{var}_n(\mathbf{x})}} \\ & \qquad \qquad \text{where} \qquad & \mathrm{corr}_n(\mathbf{x}, \mathbf{y}) = \frac{\frac{1}{n} \sum_{i=1}^n (x_i - \overline{x}_n)(y_i - \overline{y}_n)}{\sqrt{\mathrm{var}_n(\mathbf{x})} \sqrt{\mathrm{var}_n(\mathbf{y})}} \\ & \qquad \qquad \text{and} \qquad & \qquad \mathrm{var}_n(\mathbf{z}) = \frac{1}{n} \sum_{i=1}^n (z_i - \overline{z}_n)^2 \text{ (for any } \mathbf{z} = (z_1, \dots, z_n)^\top) \end{aligned}$$

respectively empirical correlation empirical variances

Back to the cars example

Line slope :
$$\operatorname{corr}_n(\mathbf{x}, \mathbf{y}) \cdot \frac{\sqrt{\operatorname{var}_n(\mathbf{y})}}{\sqrt{\operatorname{var}_n(\mathbf{x})}} = 3.932409.$$



Centering

Centered model:

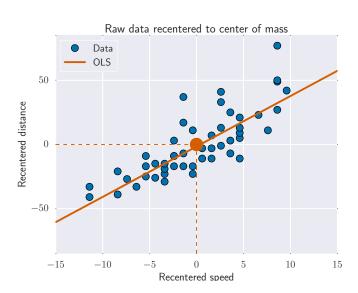
Write for any
$$i=1,\ldots,n: \begin{cases} x_i'=x_i-\overline{x}_n \\ y_i'=y_i-\overline{y}_n \end{cases} \Leftrightarrow \begin{cases} \mathbf{x}'=\mathbf{x}-\overline{x}_n\mathbf{1}_n \\ \mathbf{y}'=\mathbf{y}-\overline{y}_n\mathbf{1}_n \end{cases}$$

and $\mathbf{1}_n = (1, \dots, 1)^{\top} \in \mathbb{R}^n$, then solving the OLS with $(\mathbf{x}', \mathbf{y}')$ leads to

$$\begin{cases} \hat{\theta}'_0 = 0 \\ \hat{\theta}'_1 = \frac{\frac{1}{n} \sum_{i=1}^n x'_i y'_i}{\frac{1}{n} \sum_{i=1}^n x'_i^2} \end{cases}$$

<u>Rem</u>: equivalent to choosing the cloud of points' center of mass as origin, *i.e.*, $(\overline{x}'_n, \overline{y}'_n) = (0, 0)$

Centering (II)



Centering and interpretation

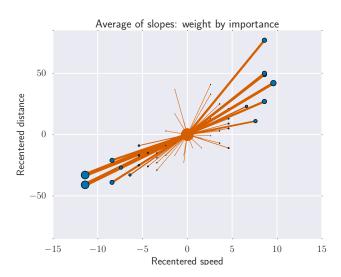
Consider the coefficient $\hat{\theta}_1'$ ($\hat{\theta}_0'=0$) for centered points \mathbf{y}', \mathbf{x}' , then :

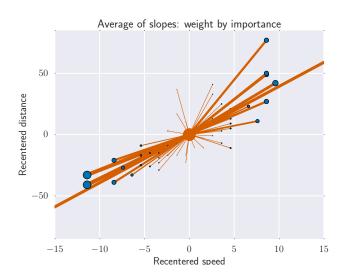
$$\widehat{\theta}_1' \in \operatorname*{arg\,min}_{\theta_1} \sum_{i=1}^n (y_i' - \theta_1 x_i')^2 = \operatorname*{arg\,min}_{\theta_1} \sum_{i=1}^n x_i'^2 \left(\frac{y_i'}{x_i'} - \theta_1\right)^2$$

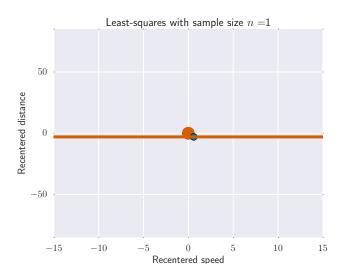
 $\underline{ \text{Interpretation}}: \widehat{\theta}_1' \text{ is a weighted average of the slopes } \tfrac{y_i'}{x_i'}$

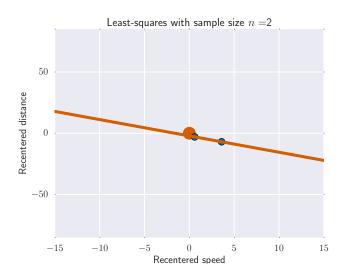
$$\widehat{\theta}'_1 = \frac{\sum_{i=1}^n x_i'^2 \frac{y_i'}{x_i'}}{\sum_{j=1}^n x_j'^2}$$

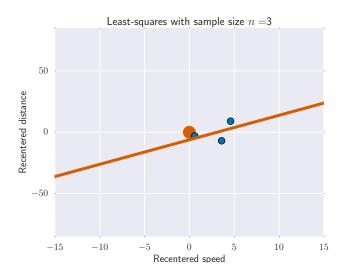
Influence of extreme points: weights proportional to x_i^2 ; connected to the **leverage** (\blacksquare : levier) effect

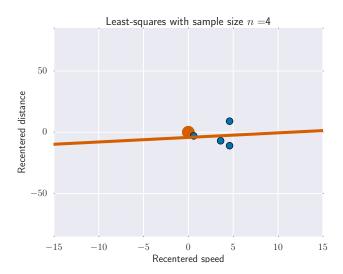


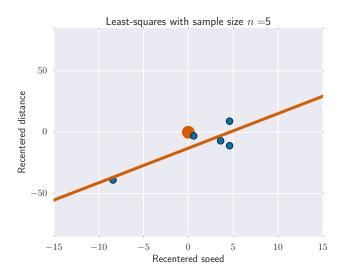


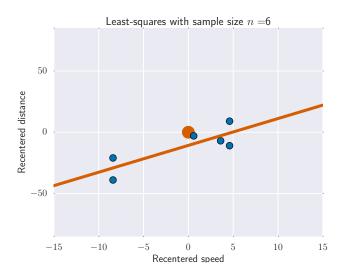


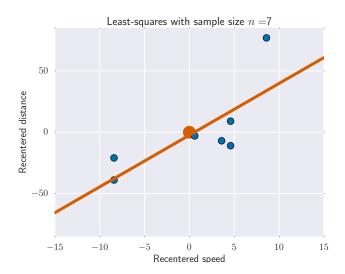


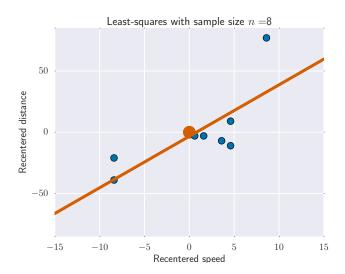


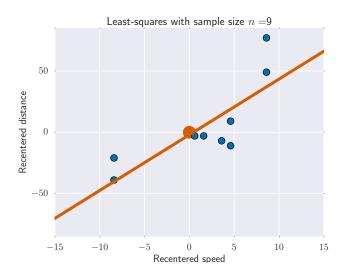


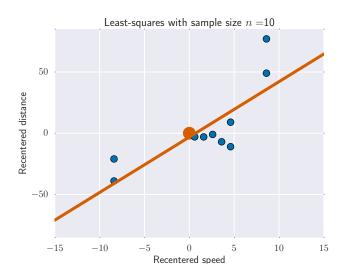


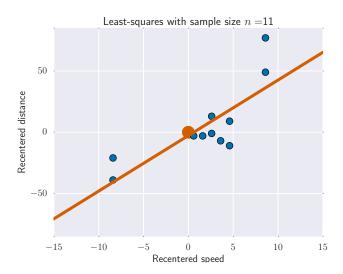


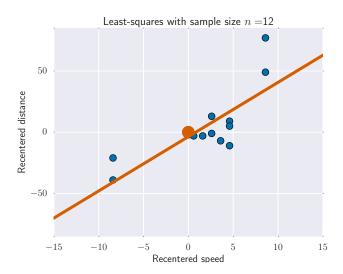


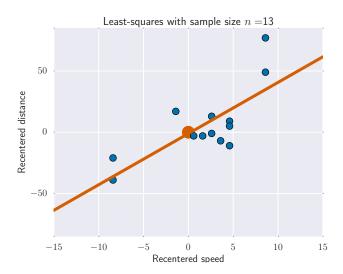


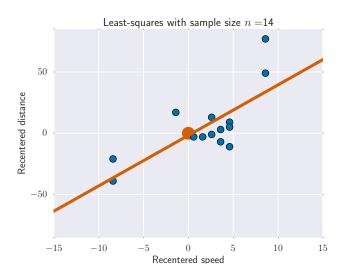


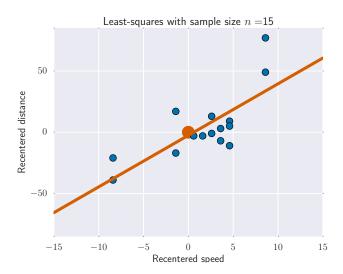


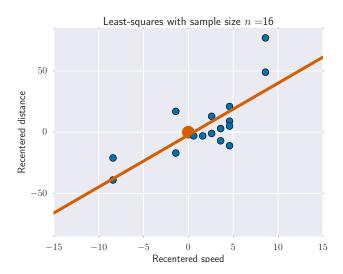


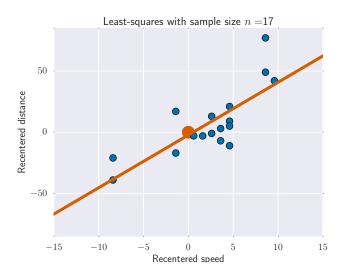


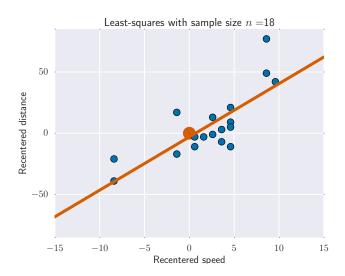


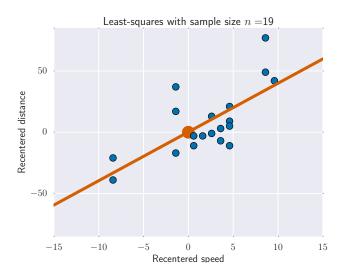


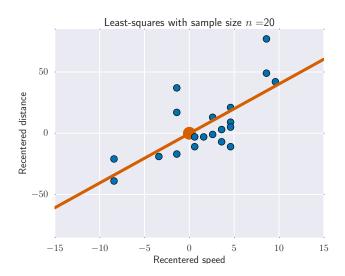


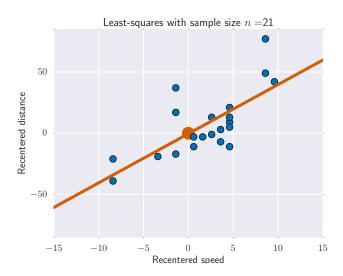


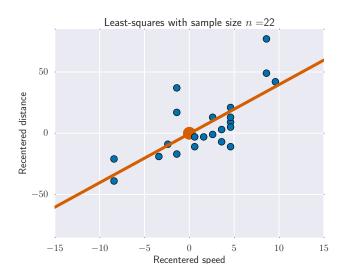


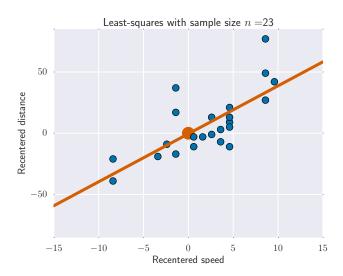


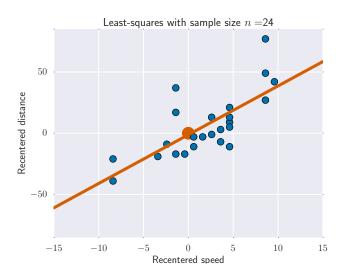


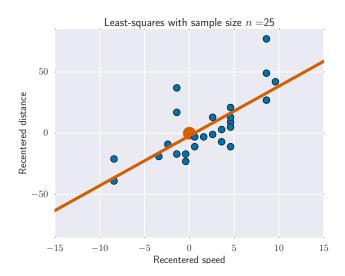


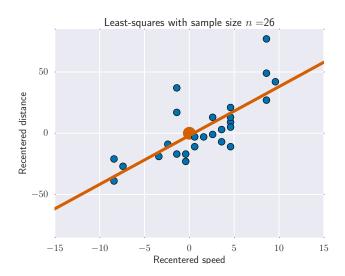




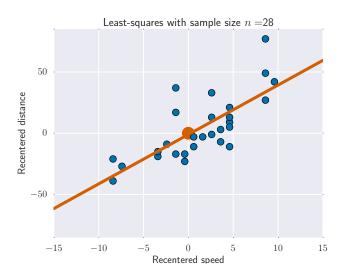


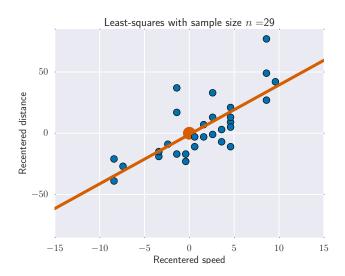


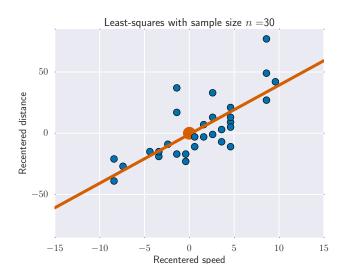


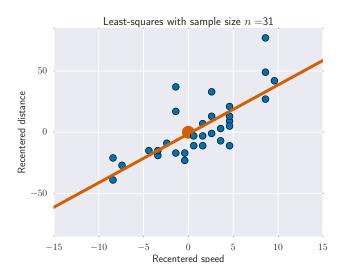




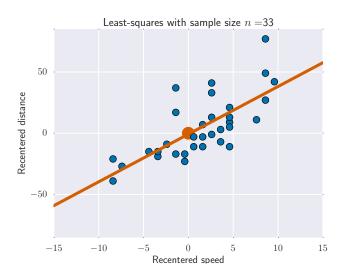


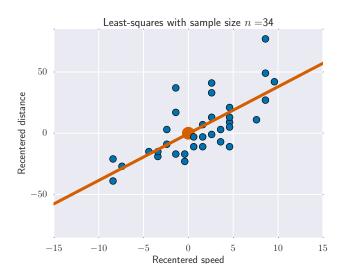


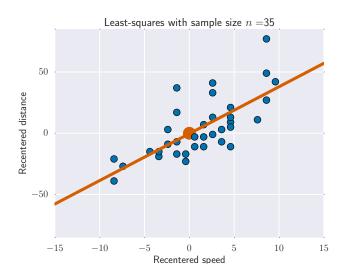


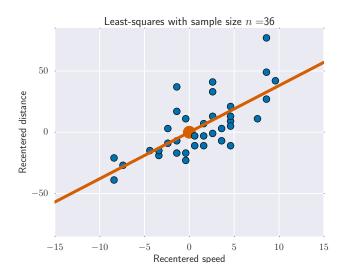


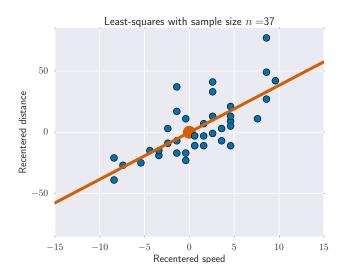


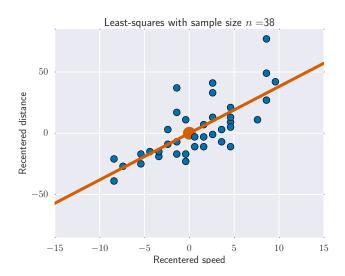


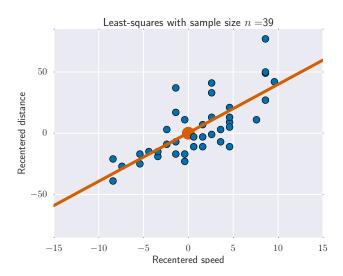


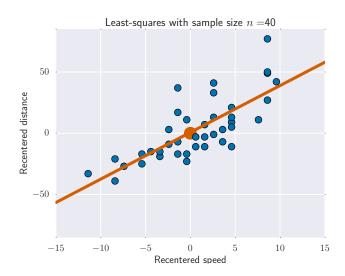


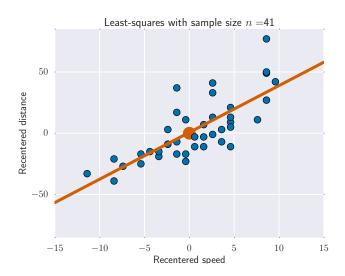


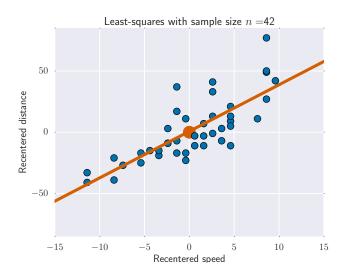


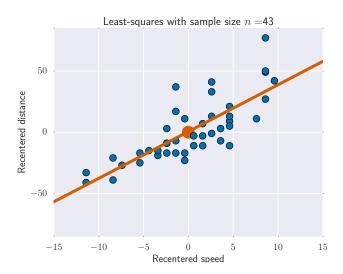


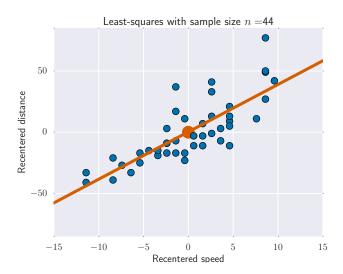


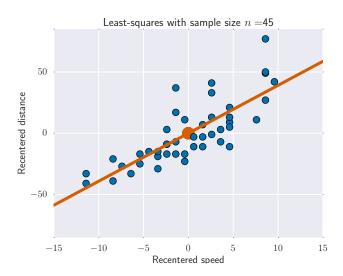


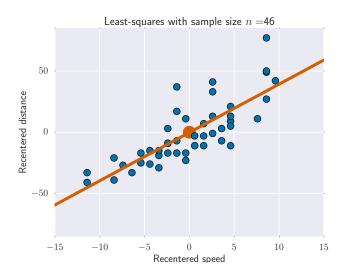


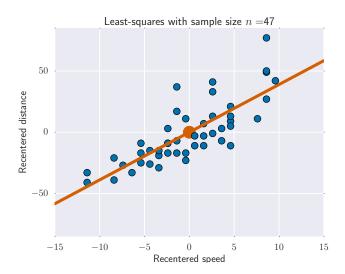


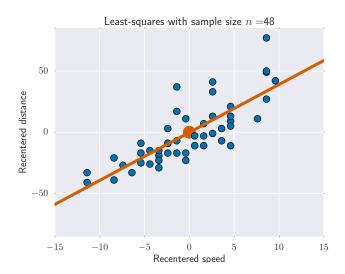




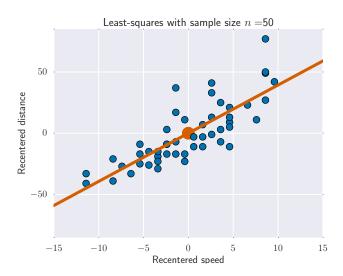












Centering + scaling

Centered-scaled model:

$$\forall i = 1, \dots, n : \begin{cases} x_i'' = (x_i - \overline{x}_n) / \sqrt{\operatorname{var}_n(\mathbf{x})} \\ y_i'' = (y_i - \overline{y}_n) / \sqrt{\operatorname{var}_n(\mathbf{y})} \end{cases} \Leftrightarrow \begin{cases} \mathbf{x}'' = \frac{\mathbf{x} - x_n \mathbf{1}_n}{\sqrt{\operatorname{var}_n(\mathbf{x})}} \\ \mathbf{y}'' = \frac{\mathbf{y} - \overline{y}_n \mathbf{1}_n}{\sqrt{\operatorname{var}_n(\mathbf{y})}} \end{cases}$$

Solving OLS with $(\mathbf{x}'', \mathbf{y}'')$ then

$$\begin{cases} \widehat{\theta}_0'' = 0 \\ \widehat{\theta}_1'' = \frac{1}{n} \sum_{i=1}^n x_i'' y_i'' \end{cases}$$

<u>Rem</u>: equivalent to choosing the points cloud center of mass as origin and normalize x and y to have unit <u>empirical norm</u> $\|\cdot\|_n$:

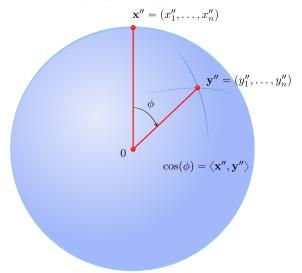
$$\|\mathbf{x}''\|_n^2 = \frac{1}{n} \sum_{i=1}^n (x_i'')^2 = 1$$
$$\|\mathbf{y}''\|_n^2 = \frac{1}{n} \sum_{i=1}^n (y_i'')^2 = 1$$

Centering + scaling



Correlation interpretation (centered-scaled case)

Example : n = 3 and $\|\mathbf{x}''\|_n^2 = \|\mathbf{y}''\|_n^2 = 1$



When/why preprocessing?

Centering ${\bf y}$ or using an intercept (or adding a constant feature) is equivalent

Rem:for sparse (\blacksquare : creux) cases centering y adding a constant feature could be preferred

Scaling features is important:

- if you want to <u>interpret</u> the coefficients' amplitude in regression (better solution : t-tests)
- if you want to <u>penalize</u> or <u>regularize</u> coefficients (*cf.*Lasso, Ridge, etc.) a single scale is needed
- for <u>computing</u> reasons (*e.g.*, store scaling to improve efficiency, improve matrix conditioning, etc.)

<u>Rem</u>: in practice centering/scaling is useful for **estimation** not so much for **prediction** (see next courses)

Centering with Python

Use centering classes from sklearn, see preprocessing: http://scikit-learn.org/stable/modules/preprocessing.html

```
from sklearn import preprocessing
scaler = preprocessing.StandardScaler().fit(X)
print(np.isclose(scaler.mean_, np.mean(X)))
print(np.array_equal(scaler.std_, np.std(X)))
print(np.array_equal(scaler.transform(X),
                   (X - np.mean(X)) / np.std(X))
print(np.array_equal(scaler.transform([26]),
                   (26 - np.mean(X)) / np.std(X)))
```

Rem:most valuable with pipeline

http://scikit-learn.org/stable/modules/pipeline.html

Definitions

Prediction

We call **prediction** function the function that associates an estimation of the variable of interest to a new sample. For least squares the prediction is given by :

$$\operatorname{pred}(x_{n+1}) = \hat{\theta}_0 + \hat{\theta}_1 x_{n+1}$$

Rem: often written \hat{y}_{n+1} (implicit dependence on x_{n+1})

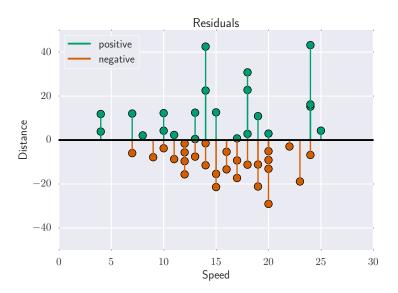
Residual

The **residual** : difference between observations and predicted values

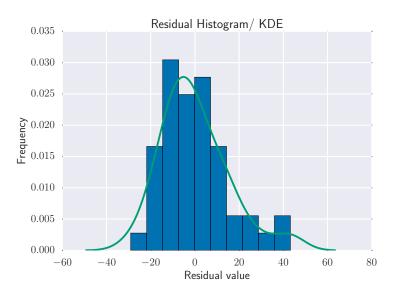
$$r_i = y_i - \text{pred}(x_i) = y_i - \hat{y}_i = y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i)$$

Rem: observable estimate of the unobservable statistical error

Residuals (on cars)



Residual histograms



Reminder: $r_i = y_i - \operatorname{pred}(x_i) = y_i - \hat{y}_i = y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i)$

Properties

Residuals are **centered** : $\frac{1}{n} \sum_{i=1}^{n} r_i = 0$

Reminder:
$$r_i = y_i - \operatorname{pred}(x_i) = y_i - \hat{y}_i = y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i)$$

Properties

Residuals are **centered**:
$$\frac{1}{n} \sum_{i=1}^{n} r_i = 0$$

$$\frac{1}{n}\sum_{i=1}^{n}r_{i} = \frac{1}{n}\sum_{i=1}^{n}(y_{i} - \operatorname{pred}(x_{i}))$$

Reminder:
$$r_i = y_i - \operatorname{pred}(x_i) = y_i - \hat{y}_i = y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i)$$

Properties

Residuals are **centered**:
$$\frac{1}{n} \sum_{i=1}^{n} r_i = 0$$

$$\frac{1}{n} \sum_{i=1}^{n} r_i = \frac{1}{n} \sum_{i=1}^{n} (y_i - \text{pred}(x_i))$$
$$= \frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{y}_i)$$

Reminder:
$$r_i = y_i - \operatorname{pred}(x_i) = y_i - \hat{y}_i = y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i)$$

Properties

Residuals are **centered**:
$$\frac{1}{n}\sum_{i=1}^{n}r_{i}=0$$

$$\frac{1}{n} \sum_{i=1}^{n} r_i = \frac{1}{n} \sum_{i=1}^{n} (y_i - \text{pred}(x_i))$$

$$= \frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{y}_i)$$

$$= \frac{1}{n} \sum_{i=1}^{n} (y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i))$$

Reminder:
$$r_i = y_i - \operatorname{pred}(x_i) = y_i - \hat{y}_i = y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i)$$

Properties

Residuals are **centered**:
$$\frac{1}{n}\sum_{i=1}^{n}r_{i}=0$$

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$$= \frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{y}_i)$$

$$= \frac{1}{n} \sum_{i=1}^{n} (y_i - (\hat{\theta}_0 + \hat{\theta}_1 x_i))$$

$$= \overline{y}_n - (\hat{\theta}_0 + \hat{\theta}_1 \overline{x}_n) = 0$$

Least squares motivation

- Computing advantage : computationally heavy methods avoided before computers (e.g., iterative methods)
- Theoretical advantage : least square analysis easy under simple hypothesis

<u>Rem</u>: for another noise model and/or to limit outliers influence one can solve (see *e.g.*,QuantReg in statsmodels)

$$\hat{\boldsymbol{\theta}} = (\hat{\theta}_0, \hat{\theta}_1) \in \underset{(\theta_0, \theta_1) \in \mathbb{R}^2}{\operatorname{arg \, min}} \sum_{i=1}^n |y_i - \theta_0 - \theta_1 x_i|$$

Gaussian likelihood

Reminder: univariate probability density function (pdf)

We write $Y \sim \mathcal{N}(\mu, \sigma^2)$, for a random variable with pdf $\varphi_{\mu,\sigma}(y) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y-\mu)^2}{2\sigma^2}\right)$

Assume : $y_i \sim \mathcal{N}(\theta_0^{\star} + \theta_1^{\star} x_i, \sigma^2)$, i.e., $\varepsilon_i \sim \mathcal{N}(0, \sigma^2)$, then the most likely couple (θ_0, θ_1) based on the observations is maximizing the pdf of (y_1, \dots, y_n)

Under an independence hypothesis, this is achieved by solving :

$$(\hat{\theta}_0, \hat{\theta}_1) \in \operatorname*{arg\,max}_{(\theta_0, \theta_1) \in \mathbb{R}^2} \prod_{i=1}^n \left(\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_i - \theta_0 - \theta_1 x_i)^2}{2\sigma^2} \right) \right)$$

Exo: check that this is equivalent to the least squares formulation

Discussion: toward multivariate cases

Physical laws (or your driving school memories) would lead to rather pick a **quadratic** model instead of a **linear** one : the OLS can be applied by choosing x_i^2 as features instead of x_i :



Web sites and books to go further

Datascience in general : Blog + videos by Jake Vanderplas http://jakevdp.github.io/

```
Homework for next lesson : watch the following videos
http://jakevdp.github.io/blog/2017/03/03/
reproducible-data-analysis-in-jupyter/
```

- A few notebooks of OLS with statsmodels
- McKinney (2012) about Python for statistics
- ► Lejeune (2010) about linear models (in French)
- Regression course by B. Delyon (in French, more technical)

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