

Recall: $|e^{ix} - 1| \leq |x| \wedge 2.$

$$|x| = \sum_{n=0}^{\infty} 1 \delta_0 \geq \left| \sum_{n=0}^{\infty} e^{iu} \delta_0 \right| = \left| \frac{e^{iu} - 1}{i} \right|$$

$$\left| \frac{a}{b} \right| = \frac{|a|}{|b|}, \quad |a| = |re^{i\theta}| = r$$

$$b = \sum_n e^{in\theta}$$

$$\left| \sum_{n=0}^{\infty} (e^{iu} - 1) \delta_u \right| \leq \sum_{n=0}^{\infty} |e^{iu} - 1| \delta_u \leq \sum_{n=0}^{\infty} |u| \delta_u \wedge \sum_{n=0}^{\infty} 2 \delta_u$$

$$\begin{aligned} & \left| \frac{e^{iu} - 1}{i} - 0 \right| \\ & |e^{iu} - (1 + iu)| \leq \frac{u^2}{2} \wedge 2|u| \\ & = \left| e^{iu} - 1 - iu \right| \end{aligned}$$

we need:

$$\left| e^{iu} - (1 + iu - \frac{u^2}{2}) \right| \leq \frac{|u|^3}{3!} \wedge u^2$$

now look at characteristic functions
Replace x with xt and take Expectation

$$\begin{aligned} |\mathbb{E}(e^{itx})| & \leq \mathbb{E} |e^{itx} - (1 + itx + \frac{x^2 t^2}{2} \epsilon^2)| \leq \mathbb{E} \frac{|t|^3 |x|^3}{3!} \wedge t^2 x^2 \\ |\psi_x(t) - (1 + it\mathbb{E}(x) + \frac{t^2 \mathbb{E}(x^2)}{2})| & \leq \mathbb{E} \left[\frac{|t|^3 |x|^3}{3!} \wedge x^2 t^2 \right] \quad \textcircled{*} \end{aligned}$$

If $\mathbb{E}(x) = 0$. Then $it\mathbb{E}(x) = 0$

Lindberg-Feller. LF

$$\{X_{n,m}\}_{1 \leq m \leq n}, \quad \mathbb{E}(X_{1,m}) = 0, \quad \{X_{n,m}\} \text{ pairwise ind. } \forall n \geq 1$$

$$\text{if } \sum_{m=1}^n \mathbb{E}(X_{n,m}^2) = 1, \quad n \geq 1 \quad \text{and} \quad \underbrace{\ln(\varepsilon) = \sum_{m=1}^n \mathbb{E}(X_{n,m}^2; |X_{n,m}| > \varepsilon)}_{\rightarrow 0} \quad \forall \varepsilon > 0$$

$$\text{then } S_n = \sum_{m=1}^n X_{n,m} \Rightarrow N(0, 1)$$

$$\text{Proof } Z_{n,m} = \psi_{X_{n,m}}(t), \quad \mathbb{E}(X_{n,m}^2) = \sigma_{n,m}^2$$

for $t \in \mathbb{R}$.

$$w_{n,m} = (-t^2 \sigma_{n,m}^2 / 2)$$

take Δ positive since
At the limit
 $t^2 \sigma_{n,m}^2 / 2$ is small

then using $\textcircled{*}$ $\forall \varepsilon > 0$

$$|Z_{n,m} - w_{n,m}| \leq \mathbb{E} \left[\frac{|t|^3 |X_{n,m}|^3}{3!} \wedge |X_{n,m}|^2 t^2 \right]$$

$$\leq t^2 \mathbb{E} \left(\frac{|t|^3 |X_{n,m}|^3}{6}; |X_{n,m}| < \varepsilon \right) + t^2 \mathbb{E}(X_{n,m}^2; |X_{n,m}| > \varepsilon)$$

since $|X_{n,m}| < \varepsilon$

$$\leq \frac{\varepsilon t^2}{6} \mathbb{E}(|X_{n,m}|^3) + t^2 \mathbb{E}(X_{n,m}^2; |X_{n,m}| > \varepsilon)$$

\therefore

$$\mathbb{E}(|X_{n,m}|^3) = 1 \text{ given.}$$

$$\therefore 1 \geq 1$$

Last time:

if $\varphi_{X_n}(t) \xrightarrow{n \rightarrow \infty} g(t)$, $\forall t \in \mathbb{R}$, and g is cont at $t=0$
then $\{X_n\}_{n \geq 1}$ is tinh.
Furthermore $g(t) = \varphi_Y(t)$, $t \in \mathbb{R}$ $\forall t$.

and $X_n \Rightarrow Y$

$$\begin{aligned} & \leq \frac{\epsilon t^2}{6} E\left[\left|t\left|X_{n,m}\right|^2\right] + t^2 E(X_{n,m}^2 | X_{n,m} > \epsilon) \\ \sum_{m=1}^n |Z_{n,m} - W_{n,m}| & \leq \frac{\epsilon |t|^3}{6} + t^2 \ln(\epsilon) \xrightarrow{n \rightarrow \infty} \frac{\epsilon |t|^3}{6} \\ & = 0 \end{aligned}$$

Recall: $\sum_{n=1}^{\infty} (1 + a_{n,m}) \xrightarrow{n \rightarrow \infty} e$

(i) $\sum_{m=1}^{\infty} a_{n,m} \xrightarrow{n \rightarrow \infty} 0$.

(ii) $\sup_n \sum_{m=1}^{\infty} |a_{n,m}| < \infty$

(iii) $\lim_{1 \leq m \leq n} |a_{n,m}| \xrightarrow{n \rightarrow \infty} 0$

Another Result:

$$\left| \sum_{i=1}^n z_i - \sum_{i=1}^n w_i \right| \leq \sum_{i=1}^n |z_i - w_i|, \text{ if } |z_i| \leq 1, |w_i| \leq 1 \quad \forall i \in n.$$

$$|Z_{n,m}| \leq |W_{n,m}| \leq 1, \quad n \geq N$$

$$\lim_{n \rightarrow \infty} \left| \sum_{m=1}^n Z_{n,m} - \sum_{m=1}^n \left(1 - \frac{t^2 \sigma_{n,m}^2}{2}\right) \right| \leq \lim_{n \rightarrow \infty} \sum_{m=1}^n |Z_{n,m} - W_{n,m}| = 0$$

$$\begin{aligned} \lim_{n \rightarrow \infty} \left| \Psi_{S_n}(t) - \sum_{m=1}^n \left(1 - \frac{t^2 \sigma_{n,m}^2}{2}\right) \right| &= 0 & \left\{ \sum_{m=1}^n \left(1 - \frac{\sigma_{n,m}^2}{2}\right) \xrightarrow{t \rightarrow 0} \frac{-t^2}{e^2} \right\} &= \Psi_{N(0,1)}(t) \\ \Rightarrow \Psi_{S_n}(t) &\xrightarrow{t \in \mathbb{R}} e^{-t^2/2} \end{aligned}$$

Gives Result $S_n \Rightarrow N(0,1)$.

truncation of CLT.

Example 1

Using truncation.

$$\{X_k, X_{k+1}\}_{k \geq 1} \text{ iid.}, \quad X_k \stackrel{d}{=} X \text{ symmetric.}$$

$$P(|X| \geq x) = x^{-2}, \quad x \geq 1 \quad P(|X| \leq 1) = 0$$

$$E(X) = 0 \quad (\text{from symmetric}).$$

$$E(X^2) \geq \int_{x=1}^{\infty} P(|X| \geq x) \cdot x dx = \int_1^{\infty} 2x^{-2} x dx = 2 \int_1^{\infty} x^{-1} dx = \infty$$

Want truncation.

$$Y_{n,m} = X_m \mathbf{1}_{|X_m| \leq \sqrt{n \log \log n}}, \quad C_n = \sqrt{n \log \log n}$$

$$\text{if } n = 7 \quad e^7 \\ \log \log 7 = e^7$$

$$Y_{n,m} = X_m \mathbb{1}_{|X_m| \leq \sqrt{n \log \log n}}, \quad C_n = \sqrt{n \log \log n} \quad \log \log 4 = e^k$$

$$S_n = \sum_{m=1}^n Y_{n,m}, \quad T_n \geq Y_{n,m}$$

$$\textcircled{1} \quad P(S_n \neq T_n) \leq \sum_{m=1}^n P(Y_{n,m} \neq X_m) = n \cdot P(|X| > C_n) = \frac{n}{C_n^2} = \frac{1}{\log \log n} \xrightarrow[n \rightarrow \infty]{} 0$$

$$E(Y_{n,m}^2) = \int_1^{C_n} y^2 \cdot f_{Y_{n,m}}(y) dy = \int_1^{C_n} y^2 \frac{2}{y^3} dy = \log(C_n) = \frac{1}{2}(\log(n) + \log \log(n))$$

$$\text{Note } \frac{\delta}{\partial x} x^2 = \frac{2}{y^3}$$

$$E(Y_{n,m}) \sim \log(n), \quad E(Y_{n,m}) = 0.$$

$$\text{Var}(T_n) \equiv n \log(n)$$

L-F theorem.

$$\frac{T_n}{\sqrt{n \log n}} \Rightarrow N(0,1).$$

$$\Rightarrow \frac{S_n}{\sqrt{n \log(n)}} \Rightarrow N(0,1)$$

$$S_n = T_n + (S_n - T_n)$$

Divide by $\sqrt{n \log n}$

Paul Levy. Result Domain of attraction of $N(0,1)$

$$\{x_n\} \text{ i.i.d. } \exists a_n, b_n \text{ s.t. } \frac{S_n - a_n}{b_n} \Rightarrow N(0,1)$$

$$\text{i.f.r. } \frac{y^2 \cdot P(|X| > y)}{E(X^2; |X| > y)} \xrightarrow[y \rightarrow \infty]{} 0$$