

Lindeberg-Feller CLT

Triangular Array $\{X_{n,k}\}_{1 \leq k \leq n, n \geq 1}$
 $E(X_{n,k}) = 0, \sum_{k=1}^n E(X_{n,k}^2) = 1, \{X_{n,k}\}_{1 \leq k \leq n}$ row i.i.d.

on Friday:

 $\{X_k\}_{k \geq 1}$ i.i.d. $E(X_k) = 0, E(X_k^2) = 1,$

$X_{n,k} = \frac{X_k}{\sqrt{n}}, 1 \leq k \leq n.$

to get above. "how can we generalize.
 - Remove i.i.d. constraint.

Really need Lindeberg condition.

If $L_n(\epsilon) = \sum E(X_{n,k}^2; |X_{n,k}| > \epsilon) \xrightarrow{n \rightarrow \infty} 0 \forall \epsilon > 0$
 then $S_n = \sum X_n \sim N(0, 1)$ ($S_n = \sum_{k=1}^n X_{n,k}$)
 or is it?

$\cancel{\text{a. } E(\sum_{k=1}^n X_{n,k}^2; |\sum_{k=1}^n X_{n,k}| > \sqrt{n}\epsilon) \xrightarrow{n \rightarrow \infty} 0}$

Because $EY^2 \rightarrow 1$ this goes to 0. As $n \rightarrow \infty$.Lindeberg Condition $\Rightarrow \max_{1 \leq k \leq n} E(X_{n,k}^2) \xrightarrow{n \rightarrow \infty} 0$ (D) Part (i)

$\max_{1 \leq k \leq n} E(X_{n,k}^2) = \limsup E(X_{n,k}^2) \xrightarrow{n \rightarrow \infty} \epsilon^2 \rightarrow 0 \quad \forall \epsilon > 0$

$E(X_{n,k}^2) = E(X_{n,k}^2; |X_{n,k}| > \epsilon) + E(X_{n,k}^2; |X_{n,k}| \leq \epsilon)$

$\leq L_n(\epsilon) + \epsilon^2$

By Lindeberg

0

$E(X_{n,k}^2) = \sigma_{n,k}^2 \quad \text{sd}(X_{n,k}) = \sigma_{n,k}$

 $f \in C_B^{(3)} \mathbb{R} \quad \exists z_i \text{ i.i.d. } z \sim N(0, 1)$

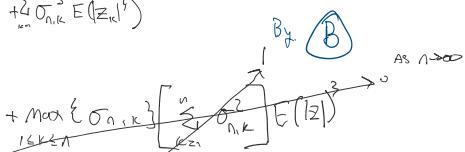
$|E(f(\sum_{k=1}^n X_{n,k})) - f(\sum_{k=1}^n \sigma_{n,k} z_k)| \quad (\oplus)$

$\sum_{k=1}^n \sigma_{n,k} z_k \sim N(0, 1),$ Replace one by one
like Friday

$(A) \leq \sum_{k=1}^n |E f(T_{n,k} + X_{n,k}) - f(T_{n,k} + \sigma_{n,k} z_k)|$

$$\begin{aligned} T_{n,k} &= \sum_{m=1}^{k-1} X_{n,m} + \sum_{m=k+1}^n \sigma_{n,m} z_m \\ &\quad \underbrace{\text{Before } X_{n,k}}_{\text{After } T_{n,k}} \quad \underbrace{\text{All less than } 1}_{\text{Can replace}} \\ &\leq C \sum_{k=1}^n E(X_{n,k}^2; |X_{n,k}| > \epsilon) + \sum_{k=1}^n E(|X_{n,k}|^2; |X_{n,k}| \leq \epsilon) + \sum_{k=1}^n \sigma_{n,k}^3 E(|z_k|^3) \end{aligned}$$

$\leq C L_n(\epsilon) + \epsilon$



$\xrightarrow{n \rightarrow \infty} \epsilon$

Example where $L_n(\epsilon) \xrightarrow{n \rightarrow \infty} 0 \quad \forall \epsilon > 0$

$\sum_{k=1}^n E(|X_{n,k}|^{2+\delta}) \xrightarrow{n \rightarrow \infty} 0 \quad \delta > 0$

Claim $\ln(\varepsilon) \leq \varepsilon^{-\delta} \sum_{k=1}^{\infty} E(|X_{n,k}|^{2+\delta}) \xrightarrow{n \rightarrow \infty} 0$

$$\varepsilon^\delta \ln(\varepsilon) \leq \sum_{k=1}^{\infty} E(|X_{n,k} - \varepsilon^\delta|; |X_{n,k}| > \varepsilon) \stackrel{\text{not by } \varepsilon^\delta}{=} \varepsilon^\delta$$

$$|X_{n,k}|^\delta > \varepsilon^\delta \quad \text{replace by.}$$

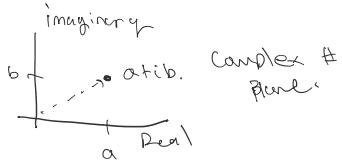
$$\leq \sum_{k=1}^{\infty} E(|X_{n,k}|^{2+\delta}; |X_{n,k}| > \varepsilon) \xrightarrow{n \rightarrow \infty} 0 \quad \text{By Assumption.}$$

Characteristic Function (Fourier transform) Ch 3.3.

Def Let X be r.v.

$$\psi_X(t) = E(e^{itX}) \quad t \in \mathbb{R}$$

$$e^{itx} = \cos(tx) + i \sin(tx)$$



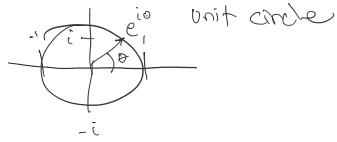
$$|a+ib| = \sqrt{a^2+b^2} \Rightarrow |a+ib|^2 = a^2+b^2 = (a+ib)(a-ib) = a^2+b^2$$

$$\text{Conjugate: } \overline{a+ib} = a-ib$$

$$e^{i\theta} = \cos(\theta) + i \sin(\theta)$$

$$|\psi^{i\theta}|^2 = \cos^2(\theta) + \sin^2(\theta) = 1$$

$$|\psi^{i\theta}| = 1.$$



| Def magnitude

| Def conjugate

| Important
Identity

Moment Generating Function (MGF)

$$\psi_X(t) = E(e^{tx})$$

| Relation to MGF.

Properties

$$(a) \psi(0) = 1$$

$$(b) \widehat{\psi(t)} = \psi(-t) \quad \text{conjugate}$$

$$(c) \text{ If } X \perp\!\!\!\perp Y \text{ then } \psi_{X+Y}(t) = \psi_X(t)\psi_Y(t) \quad ; \quad t \in \mathbb{R} \quad | \text{ theorem 3.3.2}$$

| Part of theorem 3.3.1

Characteristic function of a vector.
 $\psi_{(X,Y)}(s,t) \in e^{i(sX+ty)}, \quad s,t \in \mathbb{R}^c$

$$\text{If } X, Y \text{ r.v. } \Leftrightarrow \psi_{(X,Y)}(s,t) = \psi_X(s) \cdot \psi_Y(t)$$

| $X \perp\!\!\!\perp Y$

Example

X\Y	0	1	2	$P(X=x)$	$P(X+Y=k)$	$P(Y=k)$
0	$\frac{1}{3}$	0	$\frac{2}{3}$	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{6}$
1	$\frac{2}{3}$	$\frac{1}{3}$	0	$\frac{1}{3}$	$\frac{1}{3}$	$\frac{1}{3}$
2	0	$\frac{1}{3}$	$\frac{2}{3}$	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{6}$

$P(X+Y=k) = \sum_{x,y} P(X=x, Y=y) = \sum_{x,y} P(X=x)P(Y=y)$

$X = \begin{pmatrix} 0 \\ 1 \\ 2 \end{pmatrix}, \quad Y = \begin{pmatrix} 0 \\ 1 \\ 2 \end{pmatrix}$

$$\begin{array}{c|ccc}
 & 1 & 1 & 0 \\
 \begin{array}{c} 1 \\ 2 \end{array} & \left. \begin{array}{ccc} 2/9 & 1/9 & 0 \\ 0 & 2/9 & 1/9 \end{array} \right\} & \left. \begin{array}{c} 1/3 \\ 1/9 \end{array} \right\} & P(X+Y=k) \quad 1/9 \quad 1/9 \quad 1/9 \\
 \hline
 P(X+Y) & 1/3 & 1/3 & 1/3
 \end{array}$$

Rewritten $\psi_{X+Y}^{(t)} = \psi_X^{(t)} \psi_Y^{(t)}$

$$\begin{array}{c}
 X \stackrel{\text{D}}{=} Y \\
 \text{SAME Distribution.}
 \end{array}$$