


01-06

Monday, January 6, 2025 8:48 AM

Random vectors in \mathbb{R}^d , $d \geq 2$.

$$\underline{X} = \begin{pmatrix} x_1 \\ \vdots \\ x_d \end{pmatrix}$$


$$P(\underline{X} \in \prod_{i=1}^d [a_i, b_i]) = \prod_{i=1}^d P(x_i \in [a_i, b_i])$$

Examples

Binomial 2 outcomes

thus have more than 2.

$$\textcircled{1} \underline{X} \sim \text{multinomial}(n; p_1, \dots, p_d) \text{ where } \sum_{i=1}^d p_i = 1, p_i \geq 0$$

 d outcomes: $0, 1, \dots, d$

$$\underline{X} = (X_i)_{1 \leq i \leq d}, \quad X_i = \# \text{ of outcomes } i, \quad 1 \leq i \leq d.$$

$$X_i \in \{0, 1, \dots, n\}$$

$$\underline{X} = \sum_{k=1}^n Y_k, \quad \{Y_k\}_{1 \leq k \leq n} \text{ i.i.d.}$$

$$Y_k \in \left\{ \begin{pmatrix} 0 \\ \vdots \\ 0 \\ 1 \\ \vdots \\ 0 \end{pmatrix} \right\} \text{ if } i \text{ occurred in } k \text{ experiment.}$$

 $\therefore X_i$ is...

$$X_i \sim \text{Binomial}(n, p_i)$$

$$P(\underline{X} = \begin{pmatrix} n_1 \\ \vdots \\ n_d \end{pmatrix}) = \frac{n!}{\prod_{k=1}^d n_k!} \prod_{i=1}^d p_i^{n_i} \quad \left. \begin{array}{l} \text{CDF} \\ \text{of} \\ \text{Multinomial} \end{array} \right\}$$

$$\sum_{i=1}^d n_i = n.$$

People are interested in covariance matrix of \underline{X} .have d coordinates, i.e. $d \times d$ matrix.

$$\text{Covariance matrix } \Gamma(\underline{X}) = [\Gamma_{ij}]_{1 \leq i, j \leq d}.$$

$$\Gamma_{ij} = \text{Cov}(X_i, X_j)$$

$$= E(X_i X_j) - E(X_i)E(X_j)$$

$$\Gamma_{ii} = V(X_i) = np_i(1-p_i) \quad 1 \leq i \leq d.$$

$$\Gamma_{ij} = E(X_i X_j) - E(X_i)E(X_j)$$

$$X_i = \sum_{k=1}^n \delta_k, \quad \delta_k \sim \text{Ber}(p_i), \quad \{\delta_k\}_{1 \leq k \leq n} \text{ i.i.d.}$$

$$X_j = \sum_{k=1}^n \delta_k, \quad \delta_k \sim \text{Ber}(p_j), \quad \{\delta_k\}_{1 \leq k \leq n} \text{ i.i.d.}$$

$$\text{Cov}(X_i, X_j) = \sum_{k=1}^n \text{Cov}(\delta_k, \delta_k) = \sum_{k=1}^n \text{Cov}(\delta_k, \delta_k)$$

Bivariate linear so you can poll out.

$$\text{either } \delta_k = 1 \text{ or } \delta_k = 0.$$

$$\delta_k = 1 \text{ or } \delta_k = 0.$$

$$\delta_k = 1 \text{ or } \delta_k = 0.$$

$$\delta_k = 1 \text{ or } \delta_k = 0.$$

$$= n \text{Cov}(\delta_1, \delta_1) = n[E(\delta_1 \delta_1) - E(\delta_1)E(\delta_1)]$$

$$n[0 - p_i p_j]$$

Cov

$$\therefore \Gamma_{ij} = -np_i p_j.$$

$$\underline{X} \sim \text{multivariate normal}(\underline{\mu}, \Gamma) \quad \underline{\mu} \in \mathbb{R}^d$$

$$E(\underline{X}) = \underline{\mu}$$

$$\Gamma = \text{Cov}(\underline{X}, \underline{X})$$

$\mathbf{X} \sim \text{Multivariate normal}(\mu, \Gamma)$

$$\Gamma = [\text{cov}(\mathbf{X}_i, \mathbf{X}_j)]_{1 \leq i, j \leq d}$$

$$\begin{aligned} \mu &\in \mathbb{R}^d \\ E(\mathbf{X}) &= \mu \\ E(\mathbf{X}) &= \begin{pmatrix} E(X_1) \\ \vdots \\ E(X_d) \end{pmatrix} \end{aligned}$$

$$\text{if } \mathbf{X}^D = \bar{\mu} + A \bar{\mathbf{Z}}$$

where $A \in M_{d \times d}$ ← matrix

$$A = [a_{ij}]_{1 \leq i, j \leq d}$$

$$\bar{\mathbf{Z}} = \begin{pmatrix} Z_1 \\ Z_2 \\ \vdots \\ Z_d \end{pmatrix} \quad \{Z_i\} \text{ i.i.d. } N(0, 1)$$

what happens to covariance matrix

$$\bar{\mathbf{Y}} \in \mathbb{R}^d, E(\mathbf{Y}) = 0$$

$$\Gamma(\mathbf{Y}) = \begin{pmatrix} Y_1 \\ \vdots \\ Y_d \end{pmatrix} (Y_1 \dots Y_d) = \begin{bmatrix} \text{dxd.} \\ Y_i Y_j \\ \text{1} \leq i, j \leq d \end{bmatrix}$$

$$\Gamma(\mathbf{Y}) = E(\mathbf{Y} \cdot \mathbf{Y}^t)$$

$$\Gamma = E((\mathbf{X} - \mu)(\mathbf{X} - \mu)^t)$$

$$= A \bar{\mathbf{Z}} (A \bar{\mathbf{Z}})^t$$

$$= A \bar{\mathbf{Z}} \bar{\mathbf{Z}}^t A^t$$

$$A E(\underbrace{\bar{\mathbf{Z}} \bar{\mathbf{Z}}^t}_{\mathbf{I}}) A^t$$

$$= A A^t$$

$A A^t$ - is symmetric matrix.

transpose

$$(AB)^t = B^t A^t$$

\mathbf{X} is R.V. is Multinormal

$$\left\{ \begin{aligned} \text{iff } t \cdot \mathbf{X} &= [t, \mathbf{X}] \text{ is } N(\mu, \sigma^2), \forall t \in \mathbb{R}^d \\ t \cdot \mathbf{X} &= \sum_{i=1}^d t_i X_i \end{aligned} \right.$$

Multivariate & Multinormals are 2 important examples.
How to characterize the distributions

Let $\mathbf{X} \in \mathbb{R}^d$

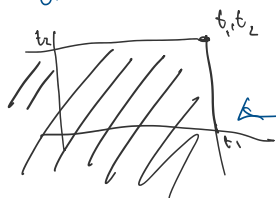
$$F_{\mathbf{X}}(t_1, \dots, t_d) = P(X_1 \leq t_1, \dots, X_d \leq t_d), \mathbf{Z} \in \mathbb{R}^d$$

you characterize the distribution w/ def

$$t_1, t_2$$

calc covariance Γ
of Multivariate normal(μ, Γ)

you characterize the distribution w/ det



The important char is not
That Even though Book says

Distribution of \mathbf{X} is characterized by $\{t \cdot \mathbf{X}\}_{t \in \mathbb{R}^d}$

Namely, $P(t \cdot \mathbf{X} \leq u)$ $u \in \mathbb{R}$

this looks like half plane



$\{t \cdot \mathbf{X} \leq 0\}$

in \mathbb{R}^d .

if $\|t\|_2 = 1$ Norm $\sum_{i=1}^d t_i^2 = 1$

$$\varphi_{t, \mathbf{X}}(1) = E e^{i(t \cdot \mathbf{X})} \quad \text{it (like I multiplied } t \cdot \text{ by } U.$$

Since for all t we can consider U part of t .

$$\varphi_{\mathbf{X}}(t) = \varphi_{t, \mathbf{X}}(1) = E e^{i(t \cdot \mathbf{X})}, \quad t \in \mathbb{R}^d.$$

Claim this is the characteristic function of \mathbf{X} .

$\mathbf{X} \in \mathbb{R}^d, t \in \mathbb{R}^d$.

Inversion Formula

Proved similar to 1D

no char. func. in infinite dim.

Problem use $\varphi_{\mathbf{X}}(t)$ to calculate $P(\mathbf{X} \in A)$ where

$$A = \bigcup_{i=1}^n (a_i, b_i)$$

the boundary.

$$P(\mathbf{X} \in J_A)$$

\mathbf{X} subinterval ind. uniform.

$$Y = \mathbf{X} - U.$$

$$U \sim \text{Uniform}(A)$$

$$f_U(t) = \frac{1}{\text{Volume}(A)}, \quad t \in A.$$

Where $U \neq \mathbf{X}$ ind.

if $\mathbf{X} \notin A$ then 0.

φ is bounded integrable even if \mathbf{X} is ind.

$$f_{\mathbf{X}}(\vec{0}) = P(\mathbf{X} \in A).$$