

## Extension of BC. II

Let  $A_1, A_2, \dots$  be pairwise ind. and  $\sum_{k=1}^{\infty} P(A_k) = \infty$  (theorem 2.3.8 Durrett)

Then  $\frac{1}{\sum_{m=1}^n P(A_m)} \sum_{m=1}^n \mathbb{1}_{A_m} \xrightarrow[n \rightarrow \infty]{a.s.} 1$  much stronger BC II.

Example  $r_1, r_2, r_3, \dots$  are iid r symmetric Bernoulli.

$$P(r_{n+1}) = \frac{1}{2} = P(r_i = -1)$$

we look at  $\left\{ \sum_{k=1}^m r_{n_k} \right\}_{n_1 < n_2 < \dots < n_m}$  countable set

Claim : Pairwise Independence. But not Ind.

Consider

$$r_1, r_2, r_3 = r_1 \cdot r_2 \text{ not ind. for all 3.}$$

$$P(r_1=1, r_2=1, r_1 r_2 = -1) = 0. \quad 0 \neq \frac{1}{8} \therefore$$

$$P(r_1=1) = \frac{1}{2} = P(r_2=1) = P(r_1 r_2 = -1) = \frac{1}{8} \text{ not ind.}$$

$$P(r_2=1, r_1=-1) = P(r_1=1) P(r_1 r_2 = -1)$$

Proof of BC II Ext.

$$V\left(\sum_{m=1}^n \mathbb{1}_{A_m}\right) = \sum_{m=1}^n V(\mathbb{1}_{A_m}) = \sum_{m=1}^n P(A_m)(1 - P(A_m)) \leq \sum_{m=1}^n P(A_m) = E\sum_{m=1}^n \mathbb{1}_{A_m}$$

$$S_n = \sum_{m=1}^n \mathbb{1}_{A_m}. \text{ we get } V(S_n) \leq E(S_n)$$

Chebychev,

$$\frac{E(S_n)}{E(S_n)} \xrightarrow[n \rightarrow \infty]{P} 1$$

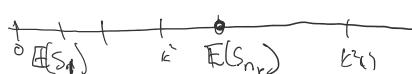
B.v.(1) subseq.  $\sum_{m=k}^n \mathbb{1}_{A_m} \xrightarrow{n \rightarrow \infty} 1$  with.

$$\text{Let } k^2 + 1 \geq E(S_{n_k}) \geq k^2, \quad k = 1, 2, \dots$$

why does  $\left\{ \mathbb{1}_{A_k} \right\}_{k \leq k^2}$  exist?

$$E(S_n) = \sum_{m=1}^n P(A_m)$$

$$E(S_{n+1}) = \sum_{m=1}^{n+1} P(A_m)$$



$$E(S_{n+1}) - E(S_n) = P(A_{n+1}) < 1$$

$E(S_n) \uparrow \infty$  ... law of large numbers study chebychev.

$$\mathbb{E}(S_{n+1}) - \mathbb{E}(S_n) = \mathbb{P}(A_{n+1}) < 1$$

$\mathbb{E}(S_n) \uparrow \infty$  study chebychev.

$$\mathbb{P}(|S_{n_k} - \mathbb{E}(S_{n_k})| > \delta(k^2 + 1)) \stackrel{\text{chebychev.}}{\leq} \frac{\mathbb{V}(S_{n_k})}{\delta^2(k^2 + 1)^2} \leq \frac{\mathbb{E}(S_{n_k}) k^2 + 1}{\delta^2(k^2 + 1)^2} = \frac{1}{\delta^2(k^2 + 1)}$$

$$\sum_{k=1}^{\infty} \mathbb{P}(|S_{n_k} - \mathbb{E}(S_{n_k})| > \delta(k^2 + 1)) \leq \frac{1}{\delta^2} \sum_{k=1}^{\infty} \frac{1}{k^2 + 1} < \infty$$

→ BC. I. know

BC. I. says

$$\mathbb{P}\left(\frac{|S_{n_k} - \mathbb{E}(S_{n_k})|}{k^2 + 1} > \delta \text{ i.o.}\right) = 0 \quad \text{i.o.}$$

$$\Rightarrow \frac{\mathbb{E}(S_{n_k}) [S_{n_k} - \mathbb{E}(S_{n_k})]}{k^2 + 1} \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 0$$

$$\frac{k^2}{k^2 + 1} \leq \frac{\mathbb{E}(S_{n_k})}{k^2 + 1} \leq 1 \quad \Rightarrow \quad \frac{S_{n_k}}{\mathbb{E}(S_{n_k})} \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 1$$

"How do we go from subseq to seq?"

$$\text{Let } n_{k_l} \leq n \leq n_{k_l+1}$$

$$\frac{S_{n_k}}{\mathbb{E}(S_{n_k})} \frac{S_n}{\mathbb{E}(S_n)} \leq \frac{S_{n_{k_l+1}}}{\mathbb{E}(S_{n_{k_l}})} \quad S_n \text{ is always smaller than } S_{n_k}$$

$$\begin{aligned} &= \frac{S_{k+1}}{\mathbb{E}(S_{n_{k+1}})} \cdot \frac{\mathbb{E}(S_{n_{k+1}})}{\mathbb{E}(S_{n_k})} \\ &\quad \downarrow \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 1 \quad \downarrow \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 1 \\ &\left[ \frac{\mathbb{E}(S_{n_k})}{\mathbb{E}(S_{n_{k+1}})} \right] \cdot \frac{S_{n_k}}{\mathbb{E}(S_{n_k})} = \\ &\quad \downarrow \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 1 \quad \downarrow \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 1 \end{aligned}$$

What is ratio.  
This shows it is between  
two ratios that converge to 1

$$S_{n_{k+1}} \geq S_{n_k} \quad \therefore 1 \leq \frac{\mathbb{E}(S_{n_{k+1}})}{\mathbb{E}(S_{n_k})} \leq \frac{(k+1)^2 + 1}{k^2} \xrightarrow[k \rightarrow \infty]{\text{a.s.}} 1$$

which proves BCII ext.

Record Wave:

Let  $\{x_i, \bar{x}_i\}_{i=1}^{\infty}$  iid.  $\mathbb{P}(X=x) = 0 \quad \forall x \in \mathbb{R}$

$$A_k = \left\{ X_k > \max_{1 \leq i \leq k-1} \{x_i\} \right\} \quad \text{given some } k \text{ is the highest.}$$

$$\text{Prob}(A_{1:k}) = \mathbb{P}(X_{1:k} = k)$$

$$= \frac{(k-1)!}{k!} = \frac{1}{k}$$

Random Permutation on  $\{1, 2, \dots, k\}$ .  
 $X_k : X_1, X_2, \dots, X_k$

#Permutations is  $k!$

they were pairwise independent.  $P(\sigma) = \frac{1}{k!}$

Let  $j < k$

$$P(A_k \cap A_j) = P(x_k=k, x_j \geq \max\{x_1, \dots, x_{j-1}\})$$

$$= \frac{(j-1)! (k-j)! \binom{k-1}{j}}{k!} = \frac{1}{k_j}$$

↑ how many times before, ↓ from  $\frac{1}{k_j}$

The Probability of Events A  
uncorrelated

$A_k, A_j$  are FND.

Now use the Borel-Cantelli Lemma (Extension of BCL)

$$\frac{R_n}{\sum_{k=1}^n k} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} 1$$

$$\frac{R_n}{(\log n)^{\frac{1}{2}}} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} 1$$

$0 \leq x_i \leq M \quad \{x_i\}_{i \geq 1} \text{ uncorrelated.}$

$$\sum_{k=1}^{\infty} E(X_k) = \infty$$

Then:

$$\frac{S_n}{E(S_n)} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} 1$$

Imp  $\rightarrow$  pairwise

$$\frac{(k-1)!}{j!(k-1-j)!}$$

$$\binom{k-1}{k-1-j} \cdot \binom{k-1-k+1+j}{k-1-j}$$

trick: Replace  $x_i = \frac{X_i}{M}, 0 \leq y_i \leq 1$