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Research Article



Personality Maturation Around the World: A Cross-Cultural Examination of Social-Investment Theory

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Abstract

During early adulthood, individuals from different cultures across the world tend to become more agreeable, more conscientious, and less neurotic. Two leading theories offer different explanations for these pervasive age trends: Five-factor theory proposes that personality maturation is largely determined by genetic factors, whereas social-investment theory proposes that personality maturation in early adulthood is largely the result of normative life transitions to adult roles. In the research reported here, we conducted the first systematic cross-cultural test of these theories using data from a large Internet-based sample of young adults from 62 nations (N = 884,328). We found strong evidence for universal personality maturation from early to middle adulthood, yet there were significant cultural differences in age effects on personality traits. Consistent with social-investment theory, results showed that cultures with an earlier onset of adult-role responsibilities were marked by earlier personality maturation.

Keywords

personality development, Big Five, social investment, culture, adult development, cross-cultural differences, personality

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During the period of early adulthood, individuals across the world tend to become more agreeable, more conscientious, and less neurotic (Roberts, Wood, & Caspi, 2008). Two leading theories offer radically different explanations for these pervasive age trends. Five-factor theory argues that personality development is largely controlled by genetically determined biological influences, with environmental influences having only a negligible role (McCrae & Costa, 2008). In contrast, social-investment theory argues that investment in age-graded social roles, such as spouse, parent, and employee, is a key influence on personality development (Roberts, Wood, & Smith, 2005).

In the present research, we evaluated these two competing theories by testing whether cultural differences in personality development are associated with cultural differences in the normative timing of adult-role transitions. Social-investment theory predicts such an association,

whereas five-factor theory predicts no association. We tested for this association across 62 nations spread over five continents (America, Australia, Africa, Asia, and Europe).

Personality Maturation: A Pancultural Phenomenon?

A large body of research has suggested that normative personality development in early adulthood reflects growth in the direction of greater maturity. In terms of the Big Five model of personality (McCrae & John, 1992), both longitudinal and cross-sectional studies have shown

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that individuals tend to become more agreeable, more conscientious, and less neurotic from early adulthood through middle age (for reviews see, Roberts, Walton, & Viechtbauer, 2006; Roberts et al., 2008). This robust pattern has often been referred to as the *maturity principle* of personality development (Roberts et al., 2008). Less consistent results have been reported for extraversion and openness, which generally show rather modest age differences across the life span (Soto, John, Gosling, & Potter, 2011).

Most studies have used samples from Western societies, but there is some evidence to suggest that age trends are relatively similar in different cultures. Specifically, McCrae and colleagues (1999) showed relatively consistent age differences in Big Five self-reports across six nations. A subsequent study based on data from 50 nations revealed that all 50 were characterized by agerelated increases in conscientiousness (McCrae & Terracciano, 2005). Effects for extraversion and openness were also relatively consistent and suggested that levels of both traits are slightly higher in college-aged relative to middle-aged adults in almost every culture.

What can explain these consistent patterns across cultures? As noted above, two prominent theories—five-factor theory and social-investment theory—claim to illuminate this question.

Five-factor theory

Five-factor theory defines personality traits as basic tendencies representing "endogenous dispositions that follow intrinsic paths of development essentially independent of environmental influences" (McCrae et al., 2000, p. 173). Thus, both stability and change in traits should be under the control of biological factors that are mainly determined by genes. According to this position, a lack of cross-cultural differences in age trends in personality supports the assumption that the timing of personality development is a genetically determined human universal. To the extent that there are any cultural differences in age trends, they should be neither systematic nor related to any socio-cultural markers.

Social-investment theory

An alternative account for why normative personality changes are consistent across different cultures and occur predominantly during early adulthood has been elaborated within the framework of social-investment theory (Roberts et al., 2005). Social-investment theory proposes that age-graded life transitions in early adulthood—such as completing education, entering the labor force, marrying, and becoming a parent—stimulate personality maturation because they force individuals to invest in and

commit to new social roles. These roles are connected to culture-specific societal expectations at a given age. To the extent that these expectations can be formulated in terms of traits (e.g., an expectation to act in a conscientious way), transitional-role experiences are assumed to form a reward structure for personality maturation.

According to social-investment theory, the observed normative increases in maturity can be explained by the fact that a large portion of the population is engaged in these universal life tasks during the period of early adulthood (Bleidorn, 2012; Lodi-Smith & Roberts, 2007). That is, people across many different cultures may change in broadly similar ways because these universal life tasks are promoted by most modern cultures, and a majority of people in a majority of cultures go through similar life transitions at approximately the same ages. As a result, in examinations of broad developmental trends across broadly similar cultures, the predictions of five-factor theory and social-investment theory will converge.

The predictions of the two theories will diverge only in comparisons of cultures that are sufficiently diverse to differ in their cultural expectations regarding the normative timing of adult-role transitions. Specifically, socialinvestment theory would predict earlier personality maturation in cultures in which individuals take on adultrole responsibilities at an earlier age. That is, cultural differences in age trends in personality should be a function of culture-specific social clocks that define the normative timing of those transitions (Neugarten, Moore, & Lowe, 1965). For example, the average U.S. female marries (for the first time) at age 26, much later than the average female from India, who marries at age 20, but substantially earlier than the average Swedish female, who is already 32 on her wedding day (United Nations, Department of Economic and Social Affairs, Population Division, 2008). Social-investment theory would therefore predict earlier changes in maturity-related traits in India compared with the United States, which in turn should be marked by earlier changes than Sweden.

Cultural differences in the normative timing of adult-role transitions are also captured in the theory of emerging adulthood (Arnett, 2000). This theory places less emphasis on demographic role transitions, but it also views the attainment of adult status as a process of psychological maturation (Arnett, 2001). Building on the observation that this process has considerably slowed down in Western cultures, Arnett introduced the concept of *emerging adulthood* as a new life stage between adolescence and full-fledged adulthood in industrialized cultures that postpone entry into adult roles. Emerging adulthood is supposed to be a distinctive period in life that offers the most opportunity for explorations in the areas of love, work, and worldviews before one takes on enduring adult responsibilities. Accordingly, this period

should be marked by increases in emerging adults' openness to new ideas and experiences.

Combining the assumptions of social-investment theory and the theory of emerging adulthood, it can be hypothesized that cultures with a late transition to adult roles should be marked by delayed increases in emotional stability, agreeableness, and conscientiousness along with more pronounced increases in openness. Cultures with an earlier onset of adult-role responsibilities, on the other hand, should be characterized by accelerated increases in maturity-indicating traits but less pronounced increases in openness.

The Present Study

The present study was meant to extend previous work on cultural variation in age differences in personality by providing the first large-scale systematic test of the conflicting predictions of five-factor theory and social-investment theory. First, to confirm the well-established universal age trends in the direction of greater maturity, we examined the average effects of age on Big Five personality traits across 62 cultures. We then examined the degree of cross-cultural variance in these effects. Finally, we examined whether cross-cultural variation in patterns of change was associated with the normative timing of adult-role transitions (e.g., marriage, parenthood) across cultures.

Method

Participants and procedure

To provide a meaningful test of five-factor theory and social-investment theory, we needed to gather personality data from as broad a range of cultures as possible and from young adult participants in the age ranges in which the personality changes are hypothesized to occur. Thus, we made use of a large sample of participants who had provided personality and demographic information over the World Wide Web as part of the Gosling-Potter Internet Personality Project (for details, see Gosling, Vazire, Srivastava, & John, 2004). Potential respondents could find out about this noncommercial, advertisement-free Web site through several channels, including search engines and unsolicited links on other Web sites. After submitting their responses, participants were presented with a customized personality evaluation.

To capture the period of early adulthood, we focused on an age range of 16 to 40 years. Preliminary analyses had indicated that samples of at least 100 participants were necessary to estimate reliable age-group-specific means and systematic age trends. So, to ensure sufficiently large samples in a large number of countries, we

pooled participants within each country into five age groups of 5 years each (ages 16–20, 21–25, 26–30, 31–35, and 36–40) and selected cases only from those nations with at least 100 cases per age group. To increase the diversity of cultures in our analyses, we relaxed the criterion slightly for Kenya, Saudi Arabia, the United Arab Emirates, and Zimbabwe, which came very close to meeting the inclusion threshold. Finally, because of a tremendous overrepresentation of participants from the United States (51%), we selected a random subsample of cases from the United States (10%). As a result of the selection criteria, the initial pool of 3,116,516 participants was reduced to 884,328 participants (63% females, 37% males) from 62 nations. The list of selected nations and nation-specific sample characteristics are shown in Table 1.

Measures

Personality. Big Five personality traits were measured by using the Big Five Inventory (BFI; John, Donahue, & Kentle, 1991). Participants provided self-ratings for all items, using 5-point scales from 1 (*strongly disagree*) to 5 (*strongly agree*). In the current sample, 56% of the participants completed the original English version, 32% the Spanish version (Benet-Martínez & John, 1998), ¹ 7% the German version (Rammstedt, 1997) and 5% the Dutch version (Denissen, Geenen, van Aken, Gosling, & Potter, 2008). All scales showed satisfactory to excellent reliabilities across age groups and nations (see Table 1).

BFI raw scores were transformed to the T-score metric (standard scores with a mean of 50 and standard deviation of 10). In addition to age and nation effects, we also controlled for gender effects when converting to T scores, following a procedure suggested by Soto et al. (2011). That is, we computed the overall mean for each BFI scale by first computing its mean in each of the 620 age, gender-, and nation-specific samples (5 age groups \times 2 genders \times 62 nations) and then averaging these 620 group means. Similarly, we computed the overall standard deviation of each scale as the square root of the pooled within-group variance term from a three-way analysis of variance. This produced a T-score distribution uninfluenced by nation, age, and gender differences.

Culture-level role-transition indices. To derive indices that captured the social norms regarding the timing of transitions into roles in marriage, parenthood, and the labor force (i.e., the roles typically considered to mark the transition to adulthood), we drew on secondary culture-level data for the nations studied. Specifically, to create indices of the normative timing of marriage and parenthood, we retrieved data from the United Nations Statistics Division (http://unstats.un.org) on the percentage of (a) male and (b) female teenagers (aged 15–19) ever

 Table 1. Sample Characteristics

Nation	N	Female respondents (%)	CRT index (z score)		BFI scales (Cronbach's α) ^a				
			Family	Job	Neurot- icism	Extra- version	Openness	Agree- ableness	Con- scientiousness
Argentina	55,030	73	0.75	-0.04	.7781	.7880	.75–.77	.69–.76	.75–.77
Australia	34,004	62	-1.04	-1.51	.8486	.8788	.7580	.7881	.8384
Austria	6,026	60	-0.60	0.23	.8387	.8790	.7980	.7275	.8486
Belgium	8,721	62	-1.07	-1.71	.8488	.8587	.7679	.7376	.8185
Bolivia	3,231	66	1.26	0.39	.7581	.79–.82	.7380	.4583	.73–.83
Brazil	3,682	49	1.23	0.93	.8082	.8085	.68–.75	.64–.74	.75–.78
Canada	74,602	63	-0.46	-1.77	.83–.85	.86–.87	.76–.81	.7880	.8183
Chile	25,134	72	0.44	0.31	.78–.82	.79–.83	.76–.78	.66–.77	.76–.79
China	10,103	65	-0.10	1.14	.78–.81	.72–.78	.73–.75	.66–.72	.73–.77
Colombia	17,367	71	1.21	0.28	.77–.82	.77–.82	.74–.76	.6974	.74–.77
Costa Rica	2,826	70	0.66	-0.16	.7780	.75–.83	.67–.79	.70–.84	.73–.78
Croatia	1,774	66	-0.61	0.80	.84–.88	.77–.85	.63–.70	.72–.78	.76–.83
Cuba	1,075	74	1.64	0.59	.78–.82	.72–.81	.72–.80	.64–.74	.70–.85
Denmark	3,492	49	-1.31	-0.95	.82–.86	.83–.85	.74–.79	.74–.79	.79–.84
		76				.70–.80		.74–.79	.66–.76
Dominican Republic	3,589		1.53	0.20	.78–.80		.73–.77		
Ecuador	3,649	69	2.00	0.00	.74–.81	.75–.81	.75–.79	.68–.76	.74–.76
Egypt	2,206	65	0.30	0.57	.80–.85	.7176	.65–.81	.6079	.75–.79
El Salvador	1,990	69	1.53	0.39	.77–.81	.76–.82	.7077	.6779	.71–.76
Finland	6,618	66	-1.13	-0.89	.82–.86	.83–.85	.71–.75	.71–.77	.79–.82
France	4,279	54	-1.29	-0.94	.8086	.82–.85	.75–.81	.70–.76	.79–.83
Germany	56,723	58	-1.27	-1.56	.83–.87	.86–.88	.7880	.71–.73	.83–.85
Great Britain	79,091	58	-0.07	-1.46	.8386	.86–.87	.7681	.7880	.8384
Greece	3,288	64	-0.66	0.11	.7882	.75–.81	.67–.73	.65–.74	.78–.82
Guatemala	2,872	67	1.36	1.22	.76–.84	.7681	.7577	.6481	.69–.77
Hong Kong	4,942	66	-1.20	0.17	.8285	.8082	.7380	.6774	.7781
Hungary	1,251	61	-1.01	-0.07	.8187	.8085	.6774	.7276	.7683
India	26,691	49	1.14	0.43	.8183	.7980	.6469	.6972	.7883
Indonesia	3,846	60	0.39	-0.24	.7783	.7681	.6472	.6573	.7580
Iran	1,787	59	0.51	0.74	.8083	.7279	.6674	.6574	.7379
Ireland	9,677	59	-1.12	-0.71	.8286	.8386	.7880	.7678	.8385
Israel	2,038	56	-0.50	-1.57	.8589	.7982	.6571	.7078	.7985
Italy	3,716	57	-1.21	0.39	.8185	.7884	.7176	.6674	.7682
Jamaica	1,361	76	-1.04	-0.47	.7585	.7983	.6876	.6177	.7986
Japan	2,506	69	-1.00	-0.48	.7782	.7781	.7180	.7481	.7478
Kenya	1,354	61	1.25	0.70	.7483	.7682	.67–.77	.7579	.7983
Malaysia	9,397	67	-0.29	1.87	.79–.85	.7679	.6473	.6976	.78–.81
Mexico	67,492	66	1.47	1.12	.7781	.7680	.76–.77	.7074	.7476
The Netherlands	44,493	63	-1.31	-2.40	.83–.87	.8485	.76–.79	.7273	.8182
New Zealand	10,189	65	0.47	-1.04	.83–.86	.86–.87	.76–.80	.7779	.82–.83
Nigeria	1,645	56	1.71	0.48	.72–.83	.72–.76	.66–.73	.68–.77	.79–.84
Norway	10,900	59	-1.41	-1.20	.81–.85	.81–.85	.74–.78	.70–.74	.76–.80
Pakistan	4,830	50	0.53	2.21	.79–.82	.74–.80	.59–.73	.64–.72	.72–.82
Peru	11,084	68	0.28	-0.76	.75–.81	.79–.83	.74–.78	.64–.76	.72–.82
Philippines		08 75			.75–.81	.79–.83	./4–./8 .67–.73	.71–.75	.7/79 .7680
* *	20,328	/5 62	0.47	0.47					
Poland	3,762		-0.43	0.31	.82–.85	.82–.85	.65–.75	.74–.77	.75–.85
Portugal	1,749	52	-0.22	0.55	.80–.87	.75–.87	.75–.77	.7077	.75–.83
Romania	4,080	68	-0.34	0.94	.81–.85	.8384	.7477	.71–.75	.77–.82
Russia	2,630	67	0.27	-0.07	.80–.85	.79–.82	.6874	.66–.75	.78–.81

(continued)

Table 1. (continued)

Nation	N	Female respondents (%)	CRT index (z score)		BFI scales (Cronbach's α) ^a				
			Family	Job	Neurot- icism	Extra- version	Openness	Agree- ableness	Con- scientiousness
			1 anny	,00					
Saudi Arabia	1,657	56	-0.19	1.48	.7585	.6880	.6577	.6471	.7781
Serbia	1,679	61	-0.37	0.67	.8287	.8284	.5672	.6372	.7886
Singapore	8,648	64	-0.87	1.02	.8187	.8284	.7378	.7478	.8084
South Africa	6,011	66	-0.08	0.88	.8184	.8486	.7278	.7477	.8283
South Korea	2,896	50	-1.08	-0.27	.6577	.7380	.7177	.6872	.7177
Spain	80,916	70	-0.98	-0.88	.8086	.8284	.7880	.6671	.7880
Sweden	10,173	57	-1.42	-0.81	.8286	.8486	.7278	.7477	.8083
Switzerland	6,609	58	-1.11	-0.51	.8285	.8587	.7879	.6674	.8184
Turkey	1,802	55	0.32	0.71	.7883	.7985	.6271	.6573	.7884
United Arab Emirates	2,295	64	0.01	0.38	.7885	.7581	.6677	.6977	.7183
United States of America	83,381	64	0.44	-2.02	.8385	.8787	.7881	.8080	.8283
Uruguay	3,674	73	1.02	0.30	.7782	.7581	.7581	.6164	.7578
Venezuela	10,709	74	1.51	-0.39	.7781	.7881	.7376	.6577	.7376
Zimbabwe	758	65	1.09	1.90	.8186	.7882	.6977	.7477	.8084

Note: CRT Index = cultural-role-transition indices (z scores).

married, (c) teenage birth rates (number of annual live births per 1,000 women aged 15–19), and mean age at first marriage for (d) males and (e) females. To create indices of culture-level normative timing of completion of education and entry into the labor force, we retrieved data from the UNESCO Institute for Statistics (http://www.uis.unesco.org) on (f) years of compulsory education and (g) the percentage of the population that had completed tertiary education (i.e., received a bachelor's degree or higher).

We ran a principal component analysis with Promax rotation on these variables across the 62 cultures, which resulted in two moderately correlated factors (r=.45) that explained 77% of the variance among cultures. The first component was marked by the first five variables selected to indicate the transition to marriage and parenthood; we therefore labeled this component the family index. The second component was marked by the variables selected to indicate the transition to the labor force, so we labeled it the job index. Figure 1 shows each culture's position on the two components as indicated by the regression-based factor scores (see also Table 1). Higher values point to an earlier normative timing of family and job transitions.

Analyses

We used multilevel modeling techniques (using Mplus 6.11; Muthén & Muthén, 1998) to account for the nested

structure of the data. Level 1 represented variation among individuals within nations, and Level 2 represented the variation among nations.

In a first step, we ran a series of random-coefficient regression models (Hox, 2002) to examine the average age effects on the five BFI scales across all nations. In addition to age, we included gender at Level 1 to account for the widely established gender differences in the Big Five (e.g., Schmitt, Realo, Voracek, & Allik, 2009). In these models, the intercept and slope coefficients were allowed to vary on Level 2 to capture the cross-cultural variation in BFI means as well as in age and gender effects. In a second step, we extended these models to intercept-andslope-as-outcome models. That is, we included the two culture-level role-transition indicators (family and job) as explanatory variables at Level 2 to test whether these could explain cultural differences in age trends in personality traits. (See the Supplemental Material, available online, for details of the multilevel analyses, measurement invariance, and auxiliary statistics.)

Results

Random-coefficient regression models

Table 2 shows the results of the random-coefficient regression models. Given the coding of both explanatory Level 1 variables (age and gender), the estimated fixed parameters can be interpreted as referring to the expected outcomes for a male individual aged 16 to 20.

^aThe table shows the range of Cronbach's alphas across the five age groups for each nation.

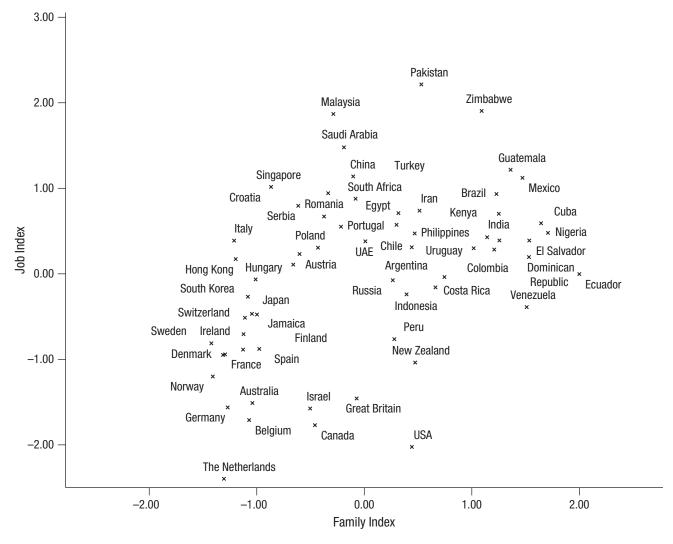


Fig. 1. Scatter plot showing the relation between regression-based factor scores on the family- and job-role-transition indices for 62 nations (see Table 1 for exact *z* scores). UAE = United Arab Emirates; USA = United States of America.

In line with the maturity principle and consistent with prior research, results showed the strongest age effects on neuroticism, agreeableness, and conscientiousness and only moderate effects on extraversion and openness. Specifically, neuroticism decreased about 2 *T*-score points, agreeableness increased about 3 *T*-score points, and conscientiousness increased nearly 8 *T*-score points from teenage years to middle age. According to Cohen's (1988) guidelines for interpreting effect sizes, the age differences in neuroticism and agreeableness represent small effects, whereas the difference of 8 *T*-score points in conscientiousness can be considered a large effect.

The random parameters indicated that the magnitude of the age effects varied significantly among nations. This result justified the use of further analyses to test whether these cultural differences in age trends were related to cultural norms regarding the timing of adult-role transitions. For the sake of completeness, Table 2 also shows the gender differences in personality across cultures. In line with previous cross-cultural findings (e.g., Schmitt et al., 2009), on average, females scored higher on neuroticism, extraversion, agreeableness, and conscientiousness, whereas males had higher scores on openness. Also consistent with prior research, the magnitude of these sex differences varied substantially across cultures.

Intercept-and-slope-as-outcome models

We extended our initial models by including the two culture-level variables indicating the normative timing of family-role and job-role transitions as explanatory variables at Level 2. The top portion of Table 3 shows the direct cross-level effects of the two culture-level variables on mean levels for Big Five factors. The family index

Table 2. Results From Random-Coefficient Regression Models: Pre	edicting the Effects of Age and Gender on
Big Five Inventory (BFI) Scales Within and Between Nations	
nr.	vr 1

		BFI scale							
Parameter	Neuroticism	Extraversion	Openness	Agreeableness	Conscientiousness				
		Fixed 6	effects						
Intercept	49.12	48.73	50.10	48.16	45.70				
Level 1									
Age	-0.56***	0.22**	0.26***	0.76***	1.91***				
Gender	4.08***	1.48***	-1.22***	1.14***	1.16***				
		Random p	arameters						
Level 1		Î							
σ^2 intercept	100.93***	99.62***	101.84***	93.96***	99.52***				
Level 2									
σ^2 intercept	2.20***	0.95***	4.35***	3.88***	2.84***				
σ^2 age	0.09***	0.08***	0.16***	0.08***	0.10***				
σ² gender	0.64***	0.56***	0.67***	0.56***	0.78***				

Note: At Level 1, N = 884,328 individuals; at Level 2, N = 62 nations. σ^2 intercept, σ^2 age, and σ^2 gender = Level 1 and Level 2 variance components for mean trait-level, age, and gender effects, respectively. **p < .01. ***p < .001.

showed positive main effects on openness and conscientiousness. The job index showed a positive main effect on neuroticism and a negative main effect on extraversion.

It is important to note that we had no strong hypotheses regarding these main effects. Rather, our main focus was on the cross-level interactions between the two role-transition indices and age (Age × Family and Age × Job) because these parameters should illuminate the question of whether cultural norms regarding the timing of adult-role transitions moderate age trends in personality. There was only one significant cross-level interaction involving the family index, which indicated that age-related increases in openness tended to be less pronounced in nations with an earlier onset of family-role responsibilities. We found a similar marginally significant effect for the job index.

We estimated the proportional reduction of cross-cultural variance in age effects in comparison with the random-coefficient regression models as a baseline, a measure analogous to the R^2 statistic used in ordinary regression analyses (Hox, 2002; see the Auxiliary Statistics section of the Supplemental Material for more details on this multilevel effect-size measure). Accordingly, the two culture-level predictors together explained 25% of the cross-cultural variance in age effects on openness. We found three further cross-level interactions between the job index and age that were all in line with the predictions of social-investment theory. Specifically, there were two positive cross-level interactions suggesting accelerated increases in agreeableness and conscientiousness in cultures with an earlier normative timing of job-role

transitions. These effects explained 13% and 20% of the cross-cultural variation in age differences in agreeableness and conscientiousness, respectively. In addition, there was a negative cross-level interaction that explained about 11% of the cross-cultural variance in age differences in neuroticism. This effect implied more pronounced age-related decreases in neuroticism in cultures with earlier job-role transitions. Figure 2 illustrates this effect by showing the age trends in neuroticism implied by the model for the cultures with the highest scores (Pakistan), average scores (Argentina), and the lowest scores (The Netherlands) on the job index.

Discussion

The present study provides the first systematic cross-cultural examination of the conflicting predictions generated by the two leading theories on personality development. In line with both five-factor theory and social-investment theory, results revealed strong evidence for the assumption that personality maturation during early adulthood is a universal phenomenon: Across 62 cultures from all over the world, neuroticism tended to decrease and extraversion,² openness, agreeableness, and conscientiousness tended to increase from early to middle adulthood. Yet there were slight but significant cultural differences in age effects on personality traits, which could be partly explained by cultural differences in the normative timing of adult-role transitions in the domains of work and family.

Table 3. Results From Intercept-and-Slope-as-Outcome Models: Predicting the Effects of Age, Gender, and the Normative Timing of Social-Role Transitions on Big Five Inventory (BFI) Scales Within and Between Nations

	BFI scale								
Parameter	Neuroticism	Extraversion	Openness	Agreeableness	Conscientiousness				
		Fixed effects							
Intercept	48.64	48.95	50.20	47.95	45.78				
Level 1									
Age	-0.51***	0.19**	0.33***	0.70***	1.84***				
Gender	4.28***	1.70***	-1.29***	1.37***	1.39***				
Level 2									
BFI score regressed on family index	0.10	0.12	1.44***	0.28	0.58**				
BFI score regressed on job index	0.70***	-0.33*	-0.30	0.27	-0.18				
Age × Family	0.03	0.01	-0.14**	-0.01	0.04				
Age × Job	-0.10*	0.05	-0.10^{\dagger}	0.09^{\dagger}	0.10*				
		Random paramet	ers						
Level 1									
σ^2 intercept	100.93***	99.62***	101.84***	93.96***	99.52***				
Level 2									
σ^2 intercept	1.60***	0.85***	2.65***	3.41***	2.29***				
σ^2 age	0.08***	0.08***	0.12***	0.07***	0.08***				
σ^2 gender	0.64***	0.56***	0.67***	0.56***	0.78***				

Note: At Level 1, N = 884,328 individuals; at Level 2, N = 62 nations; σ^2 intercept, σ^2 age, and σ^2 gender = Level 1 and Level 2 variance components for mean trait-level, age, and gender effects, respectively. "Family index" and "job index" refer to cultural indicators for the normative timing of family- and job-role transitions, respectively.

Our findings largely support the assumptions of social-investment theory. That is, the overall pattern of age trends across cultures seems not to be fully determined by a universal genetic maturation program. Rather, our findings suggest that universal life tasks during early adulthood—such as finding a partner, starting a family, and establishing one's career—are also important catalysts for personality maturation (Roberts et al., 2005).

Strikingly, transitions to the labor force turned out to be more important than family-role transitions. In fact, the expected pattern of an earlier onset of age trends in neuroticism, agreeableness, and conscientiousness was found only for cultures with an earlier normative onset of jobrole responsibilities. The lack of effects of cultural norms regarding family-role transitions is especially notable given that psychologists (Lodi-Smith & Roberts, 2007), sociologists (Hogan & Astone, 1986), and anthropologists (Schlegel, 1995) generally consider marriage and parenthood to be the most pervasive markers of adult status. Interestingly, laypeople themselves—young as well as middle-aged adults—view occupational status and career, not marriage and parenthood, as the key markers of adult status (Arnett, 2001). Perhaps the family role is marked by more freedom and less strict behavioral expectations than is the job role. In other words, the job role is probably characterized by well-defined and relatively strict demands providing a strong reward structure for personality maturation (Roberts, Caspi, & Moffitt, 2003). The family role, on the other hand, might afford more leeway, resulting in more variance at the individual level, which might in turn have diluted potential differences at the cultural level.

Cultural norms regarding the timing of both familyand job-role transitions had an effect on age trends in openness. Consistent with the theory of emerging adulthood (Arnett, 2000), our results revealed pronounced age-related increases in openness for those cultures with a late onset of family- and job-role responsibilities. This finding supports the idea of emerging adulthood as a period of prolonged identity exploration that allows young adults to broaden their horizons and to try out different lifestyles before they take over the relatively enduring responsibilities of adult life.

Limitations

The current study used a large sample and advanced analytic methods, but the findings must still be considered in light of some important limitations. First, our Internet-based samples were not representative of the general populations of the nations examined. However, it

 $^{^{\}dagger}p < .10. *p < .05. **p < .01. ***p < .001.$

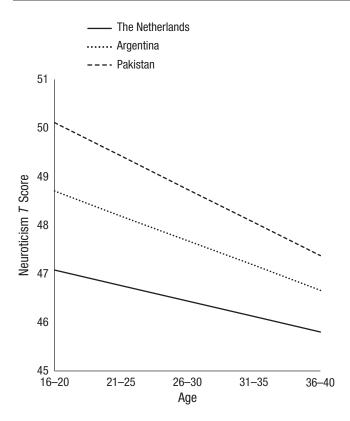


Fig. 2. Neuroticism as a function of age, as implied by the model. Results are shown for the cultures with the highest factor scores (Pakistan), average factor scores (Argentina), and the lowest factor scores (The Netherlands) on the job index (cf. Table 1).

should be noted that Internet-based samples are much more diverse and more representative than are the convenience samples commonly used in social-science research (Gosling et al., 2004), and similar results are typically obtained across Internet and non-Internet samples, especially with regard to personality variables (Srivastava, John, Gosling, & Potter, 2003). More importantly, the representativeness is likely to be poorest for developing countries in which Internet penetration is lowest. As a result, the samples obtained in developing countries are likely to be the least representative with regard to the role-transition indices associated with those countries. This restriction would tend to diminish the effects reported here, making our findings conservative estimates of the underlying effects.

Second, the cross-sectional nature of the design raises the possibility that some of the observed age trends might reflect birth-cohort effects. However, the data were collected over an 8-year period (2001–2009). Therefore, each specific age group included members of several different birth years (e.g., the oldest age group comprised 36- to 40-year-old individuals who were born between 1961 and 1973), which provided a degree of generalizability across cohorts (Soto et al., 2011). Two further

aspects of the present findings argue against the influence of cohort effects. First, the observed age trends are consistent with findings from longitudinal studies, which are not susceptible to cohort effects (Roberts et al., 2006). Second, the observed effects of cultural norms regarding the timing of role transitions are difficult to explain as a result of cohort effects.

A further limitation can be seen in the use of secondary data to develop the two culture-level predictors. These variables are only indirect markers for cultural norms regarding the timing of adult-role transitions. More proximal indicators might show even stronger effects.

Finally, cultures that differ in their onset of family- and job-role transitions might differ in numerous other ways, too. In fact, both indices were negatively correlated with nations' per capita income. However, there is neither a theoretical nor an empirical foundation for prosperity (or lack of prosperity) as an underlying source of personality maturation. But there are both theoretical and empirical reasons to view social-role transitions as key drivers of personality maturation in early adulthood (Roberts et al., 2005).

Conclusion

Our results suggest that personality maturation in early adulthood can be observed in different cultures across the world. Cultural differences in age effects can be attributed to cultural differences in the normative timing of the social-role transitions that mark entry into adulthood. These findings clearly support social-investment theory. Although genetically shaped maturation processes certainly play a role in personality development, they apparently do not tell the whole story. Rather, the majority of young adults in the majority of cultures seem to show similar age trends in personality because they to go through similar role transitions at roughly the same ages. The investment in these age-graded social institutions, such as marriage, parenthood, and work, can therefore be seen as one of the key driving mechanisms of personality change in early adulthood.

Author Contributions

W. Bleidorn developed the study concept. S. D. Gosling, P. J. Rentfrow, and J. Potter collected the data. W. Bleidorn analyzed the data, and T. A. Klimstra and J. J. A. Denissen helped with interpreting the results. W. Bleidorn drafted the manuscript, and S. D. Gosling, P. J. Rentfrow, J. J. A. Denissen, and T. A. Klimstra provided critical revisions. All authors approved the final version of the manuscript for submission.

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Declaration of Conflicting Interests

The authors declared that they had no conflicts of interest with respect to their authorship or the publication of this article.

Supplemental Material

Additional supporting information may be found at http://pss.sagepub.com/content/by/supplemental-data

Notes

- 1. Because of a programming error, we had to exclude one item from the Spanish agreeableness scale ("starts quarrels with others"; reverse-coded). However, the reliability of the Spanish agreeableness scale was only marginally affected and remained in the range of the original Spanish version (α = .64; Benet-Martínez & John, 1998).
- 2. Prior research suggested that two subcomponents of extraversion—social dominance and social vitality—demonstrate different developmental patterns (Roberts et al., 2006). In light of these differences, we ran all analyses for the two BFI extraversion facet scales measuring assertiveness and activity (Soto & John, 2009). Random-coefficient regression models revealed small but significant positive age effects on both facets. Also, we found no significant cross-level effects involving the job or family indices on either assertiveness or activity.

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