Bikeshare and Traffic Fatalities: Evidence from the Boston Metropolitan Area

Joshua Rosen

Outline

- ► Background / Research Question
- Data
- Previous Approach
- Updated Approach
- Next Steps

Background and Research Question

- Lightweight micromobility systems, such as dock-to-dock bikeshare and dockless scooters are ubiquitous in a growing number of cities, but the impacts have not been well studied
- Between 2011 and 2014, the Boston metropolitan area launched and began an initial expansion of its first dock-to-dock bikeshare system
- ▶ Does the introduction of bikeshare have a (negative) impact on city-level traffic injuries?

Brief literature review

- ▶ Cloud, Heb, Kasinger (2022) operationalize the staggered rollout of e-scooters in 93 European cities to a roughly 10% increase in reported traffic accidents after e-scooters were introduced. Effects were largest in cities with limited cycling infrastructure, and no effects were identified in cities with existing infrastructure.
- Fisherman, Schepers (2016) find that the introduction of a bikeshare system is associated with a reduction in cycling injury risk.

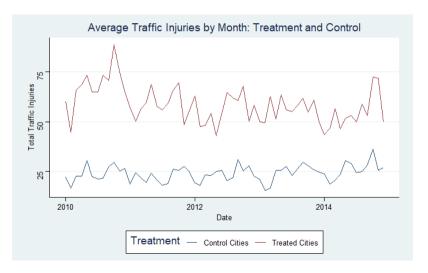
Data

Aggregated and merged month/city level data for the following variables:

- Traffic injuries
- Aggregate bikeshare trips

Data includes 7 cities: 4 cities that did not receive bikeshares during the treatment period, and 3 cities that received bikeshare (staggered treatment dates)

Trends



Decreasing trend absent of intervention (not quite clear in this plot)

Previous Empirical Framework

► Goal: identify the causal effect of bikeshare on traffic injuries using bayesian instrumental variables with random effects

Refresher on Instrumental variables:

$$x_i = \gamma + z_i \delta + \eta_i$$

$$y_i = \alpha + \hat{x}_i \beta + \epsilon_i$$

Where the two equations represent the first and second stages respectively. Here, α and γ are intercepts, and δ and β are regression coefficients. We also note that in the second stage, x_i is replaced with \hat{x}_i to indicate the fitted value of our endogenous treatment variable generated by estimating the first stage with instrument z_i . Critically, the errors η_i and ϵ_i are correlated, and we are thus unable to only estimate the second stage regression.

First stage specification

$$\begin{aligned} \textit{Trips}_{i} &= \phi + \nu_{0,cm[i]} + \gamma_{0} T_{i} + (\gamma_{1} + \nu_{1,m[i]}) T_{i} + (\delta_{0} + \nu_{2,c[i]}) T_{i} \times \\ & \textit{After}_{i} + (\delta_{1} + \nu_{3,m[i]}) T_{i} \times \textit{After}_{i} \\ & \nu_{0,cm[i]} | \sigma_{\nu_{0}} \quad \textit{i.i.d} \quad \textit{N}(0, \sigma_{\nu_{0}}^{2}) \\ & \nu_{1,m[i]} | \sigma_{\nu_{2}} \quad \textit{i.i.d} \quad \textit{N}(0, \sigma_{\nu_{2}}^{2}) \\ & \nu_{2,c[i]} | \sigma_{\nu_{3}} \quad \textit{i.i.d} \quad \textit{N}(0, \sigma_{\nu_{3}}^{2}) \\ & \nu_{3,c[i]} | \sigma_{\nu_{1}} \quad \textit{i.i.d} \quad \textit{N}(0, \sigma_{\nu_{4}}^{2}) \end{aligned}$$

Where $\nu_{c[i]}$ and $\nu_{m[i]}$ allow the relationship between T_i and $T_i \times After_i$ to vary by city-level and month-level, respectively

Second stage specification

(6)
$$y_{i} = \alpha + \eta_{0,mc[i]} + (\beta_{1} + \eta_{1,c[i]}) Trips_{i} + (\beta_{2} + \eta_{2,m[1]}) Trips_{i}$$

$$\eta_{0,cm[i]} | \sigma_{\eta_{0}} \quad i.i.d \quad N(0, \sigma_{\eta_{0}}^{2})$$

$$\eta_{1,c[i]} | \sigma_{\eta_{1}} \quad i.i.d \quad N(0, \sigma_{\eta_{1}}^{2})$$

$$\eta_{2,m[i]} | \sigma_{\eta_{2}} \quad i.i.d \quad N(0, \sigma_{\eta_{2}}^{2})$$

▶ Goal: capture city-specific $\eta_{1,c[i]}$ and month-specific $n_{2,m[i]}$ deviations in the relationship between the instrumented total bike trips and the outcome variable total traffic injuries

Current Empirical Framework

- Updated Goal: Utilize seasonal ARIMA for Bayesian time series
- Compare the predicted traffic injuries post bikeshare introduction to the observed traffic injuries. If no other conditions change at intervention point t_i, then we can attribute the forecast error to be the causal effect of bikeshare on traffic injuries.
- Forecast error: the difference between an observed value and its forecast (measured using MAE, RMSE, etc.)

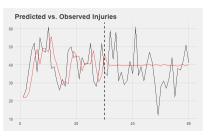
$$e_{j,T+h} = y_{T+h} - \hat{y}_{T+h|T}$$

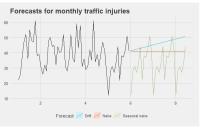
Reference Materials

- ► Largely relying on the 'bayesforecast' library in R
- 'bayesforecast' allows users to fit Bayesian time series models using (r)stan for full Bayesian inference

Methods matter!

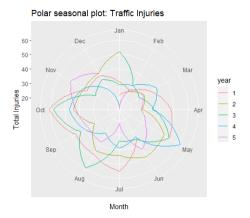
ARIMA vs. Non-ARIMA





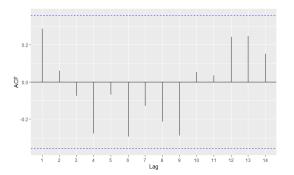
Data (ctd)

- ► A look at Cambridge data
- ▶ 60 months (5 years); 30 months before and after intervention



 Interestingly, here there is no clear seasonal trend (before or after treatment)

Data (ctd)



Correlogram: Autocorrelation Function

- Plotting the autocorrelation function r_k indicates that that monthly variation in traffic injuries is essentially random and is uncorrelated with previous months
- No trend or seasonality: this is white noise (stationary), so timing is predictable

Simple specification

Naive model / Random walk (forecast is equal to the last observed value):

$$\epsilon_{j,t} = y_{j,t} - \hat{y}_{j,t}$$
 where $\hat{y}_{j,t} = y_{j,t-1}$ and $\epsilon \sim N(0, \sigma^2)$

ARIMA specification

Seasonal ARIMA model

General model: ARIMA(p, d, q)(P, D, Q)s

▶ We might estimate a model:

$$y_{j,t} = c + \phi_1 y_{j,t-1} + \dots + \phi_{j,t-p} + \dots + \theta_1 \varepsilon_{j,t-1} + \dots + \theta_q \varepsilon_{j,t-q} + \varepsilon_{j,t}$$

Next Steps

- Set priors
- ► Set ARIMA parameters
- ▶ Dive into hierarchical time series