

```

## Loading required package: carData

## Loading required package: zoo

##
## Attaching package: 'zoo'

## The following objects are masked from 'package:base':
##
##   as.Date, as.Date.numeric

##
## Please cite as:

##   Hlavac, Marek (2018). stargazer: Well-Formatted Regression and Summary Statistics Tables.

##   R package version 5.2.2. https://CRAN.R-project.org/package=stargazer

##
## Attaching package: 'dplyr'

## The following object is masked from 'package:car':
##
##   recode

## The following objects are masked from 'package:stats':
##
##   filter, lag

## The following objects are masked from 'package:base':
##
##   intersect, setdiff, setequal, union

##
## Attaching package: 'reshape2'

## The following object is masked from 'package:tidyr':
##
##   smiths

#read in the data
#data <- read.csv(file = 'H:/ROL/MIDS/W203 Stats/lab_3/crime_v2.csv') #Robert

data <- read.csv(file = '~/Desktop/W203/w203_lab3-master/crime_v2.csv') #Praveen

```

# 1. Introduction

Ryan's section

## 2. A Model Building Process

### Exploratory Data Analysis

We started by conducting exploratory data analysis. First, we read the original paper to get a better understanding of each variable. We defined the variables in the table below and grouped them into five groups in order to get a better handle on them.

```
crime_count <- c(1:25)
data_variables <- c("county", "year", "crmrte", "prbarr", "prbconv", "prbpris", "avgsen", "polpc", "density", "taxpc", "west", "central", "urban", "pctmin80", "wcon", "wtuc", "wtrd", "wfir", "wser", "wmfg", "wfed")
data_description <- c("county identifier", "1987", "crimes committed per person", "'probability' of arrest", "'probability' of conviction", "'probability' of prison sentence", "avg. sentence, days", "police per capita", "people per sq. mile", "tax revenue per capita", "=1 if in western N.C.", "=1 if in central N.C.", "=1 if in SMSA", "perc. minority, 1980", "weekly wage, construction", "wkly wge, trns, util, commun", "wkly wge, whlesle, retail trade", "wkly wge, fin, ins, real est", "wkly wge, service industry", "wkly wge, manufacturing", "wkly wge, fed employees")
data_group <- c("Control", "", "", "Deterrent", "Deterrent", "Deterrent", "Deterrent", "Deterrent", "Deterrent", "Demographic", "Demographic", "Region", "Region", "Urban", "Demographic", "Wages", "Wages", "Wages", "Wages", "Wages", "Wages", "Wages")
data_notes <- c("", "", "ratio of FBI index crimes to county population", "ratio of arrests to offenses", "ratio of convictions to arrests", "proportion of total convictions resulting in prison sentences", "average sentence in days", "country population divided by county land area", "dummy", "dummy", "dummy", "proportion of country population that is minority or nonwhite", "average weekly wage in that sector", "average weekly wage in that sector", "average weekly wage in that sector", "average weekly wage in that sector", "average weekly wage in that sector", "average weekly wage in that sector", "average weekly wage in that sector")
data_headers <- c("Variable", "Description", "Group", "Note")
data_table <- data.frame(data_variables, data_description, data_group, data_notes)
kable(data_table, col.names = data_headers, caption = "Descriptions and Groups of Variables")
```

Table 1: Descriptions and Groups of Variables

Variable	Description	Group	Note
county	county identifier	Control	
year	1987		
crmrte	crimes committed per person		ratio of FBI index crimes to county population
prbarr	'probability' of arrest	Deterrent	ratio of arrests to offenses
prbconv	'probability' of conviction	Deterrent	ratio of convictions to arrests
prbpris	'probability' of prison sentence	Deterrent	proportion of total convictions resulting in prison sentences
avgsen	avg. sentence, days	Deterrent	average sentence in days
polpc	police per capita	Deterrent	
density	people per sq. mile	Demographic	country population divided by county land area
taxpc	tax revenue per capita	Demographic	
west	=1 if in western N.C.	Region	dummy
central	=1 if in central N.C.	Region	dummy
urban	=1 if in SMSA	Urban	dummy
pctmin80	perc. minority, 1980	Demographic	proportion of country population that is minority or nonwhite
wcon	weekly wage, construction	Wages	average weekly wage in that sector
wtuc	wkly wge, trns, util, commun	Wages	average weekly wage in that sector
wtrd	wkly wge, whlesle, retail trade	Wages	average weekly wage in that sector
wfir	wkly wge, fin, ins, real est	Wages	average weekly wage in that sector
wser	wkly wge, service industry	Wages	average weekly wage in that sector
wmfg	wkly wge, manufacturing	Wages	average weekly wage in that sector
wfed	wkly wge, fed employees	Wages	average weekly wage in that sector

Variable	Description	Group	Note
wsta	wkly wge, state employees	Wages	average weekly wage in that sector
wloc	wkly wge, local gov emps	Wages	average weekly wage in that sector
mix	offense mix: face-to-face/other	Demographic	ratio of face-to-face crimes (robbery, assault, rape) to non-face
pctymle	percent young male	Demographic	proportion of country population that is male between 15 and

To get a better sense of the data set the summary function was run.

```
summary(data)
```

```
##      county      year      crmrte      prbarr
## Min.   : 1.0   Min.   :87   Min.   :0.005533   Min.   :0.09277
## 1st Qu.: 52.0   1st Qu.:87   1st Qu.:0.020927   1st Qu.:0.20568
## Median :105.0   Median :87   Median :0.029986   Median :0.27095
## Mean   :101.6   Mean   :87   Mean   :0.033400   Mean   :0.29492
## 3rd Qu.:152.0   3rd Qu.:87   3rd Qu.:0.039642   3rd Qu.:0.34438
## Max.   :197.0   Max.   :87   Max.   :0.098966   Max.   :1.09091
## NA's   :6      NA's   :6      NA's   :6      NA's   :6
##      prbconv      prbpris      avgsen      polpc
##           : 5      Min.   :0.1500   Min.   : 5.380   Min.   :0.000746
## 0.588859022: 2      1st Qu.:0.3648   1st Qu.: 7.340   1st Qu.:0.001231
## `           : 1      Median :0.4234   Median : 9.100   Median :0.001485
## 0.068376102: 1      Mean   :0.4108   Mean   : 9.647   Mean   :0.001702
## 0.140350997: 1      3rd Qu.:0.4568   3rd Qu.:11.420   3rd Qu.:0.001877
## 0.154451996: 1      Max.   :0.6000   Max.   :20.700   Max.   :0.009054
## (Other)     :86      NA's   :6      NA's   :6      NA's   :6
##      density      taxpc      west      central
## Min.   :0.00002   Min.   : 25.69   Min.   :0.0000   Min.   :0.0000
## 1st Qu.:0.54741   1st Qu.: 30.66   1st Qu.:0.0000   1st Qu.:0.0000
## Median :0.96226   Median : 34.87   Median :0.0000   Median :0.0000
## Mean   :1.42884   Mean   : 38.06   Mean   :0.2527   Mean   :0.3736
## 3rd Qu.:1.56824   3rd Qu.: 40.95   3rd Qu.:0.5000   3rd Qu.:1.0000
## Max.   :8.82765   Max.   :119.76   Max.   :1.0000   Max.   :1.0000
## NA's   :6      NA's   :6      NA's   :6      NA's   :6
##      urban      pctmin80      wcon      wtuc
## Min.   :0.00000   Min.   : 1.284   Min.   :193.6   Min.   :187.6
## 1st Qu.:0.00000   1st Qu.: 9.845   1st Qu.:250.8   1st Qu.:374.6
## Median :0.00000   Median :24.312   Median :281.4   Median :406.5
## Mean   :0.08791   Mean   :25.495   Mean   :285.4   Mean   :411.7
## 3rd Qu.:0.00000   3rd Qu.:38.142   3rd Qu.:314.8   3rd Qu.:443.4
## Max.   :1.00000   Max.   :64.348   Max.   :436.8   Max.   :613.2
## NA's   :6      NA's   :6      NA's   :6      NA's   :6
##      wtrd      wfir      wser      wmfgr
## Min.   :154.2   Min.   :170.9   Min.   : 133.0   Min.   :157.4
## 1st Qu.:190.9   1st Qu.:286.5   1st Qu.: 229.7   1st Qu.:288.9
## Median :203.0   Median :317.3   Median : 253.2   Median :320.2
## Mean   :211.6   Mean   :322.1   Mean   : 275.6   Mean   :335.6
## 3rd Qu.:225.1   3rd Qu.:345.4   3rd Qu.: 280.5   3rd Qu.:359.6
## Max.   :354.7   Max.   :509.5   Max.   :2177.1   Max.   :646.9
## NA's   :6      NA's   :6      NA's   :6      NA's   :6
##      wfed      wsta      wloc      mix
## Min.   :326.1   Min.   :258.3   Min.   :239.2   Min.   :0.01961
```

```
## 1st Qu.:400.2 1st Qu.:329.3 1st Qu.:297.3 1st Qu.:0.08074
## Median :449.8 Median :357.7 Median :308.1 Median :0.10186
## Mean :442.9 Mean :357.5 Mean :312.7 Mean :0.12884
## 3rd Qu.:478.0 3rd Qu.:382.6 3rd Qu.:329.2 3rd Qu.:0.15175
## Max. :598.0 Max. :499.6 Max. :388.1 Max. :0.46512
## NA's :6 NA's :6 NA's :6 NA's :6
## pctymle
## Min. :0.06216
## 1st Qu.:0.07443
## Median :0.07771
## Mean :0.08396
## 3rd Qu.:0.08350
## Max. :0.24871
## NA's :6
```

This function provides a high level view of each variable. Six rows have missing values for all variables. In addition, there is one duplicate row. Also the variable prbconv is loaded as a factor, so it needs to be converted to numeric. These issues are handled below to create the initial data set.

```
#eliminate N/A's (6 rows of NA were removed)
data_crmrte <- data[!is.na(data$crmerte),]

#remove duplicates (1 duplicate record was found)
data_crmrte <- data_crmrte %>% distinct()

#prbconv was defined as factor data type , we convert it to numeric
class(data_crmrte$prbconv)
```

```
## [1] "factor"
```

```
data_crmrte$prbconv <- as.numeric(as.character(data_crmrte$prbconv))
class(data_crmrte$prbconv)
```

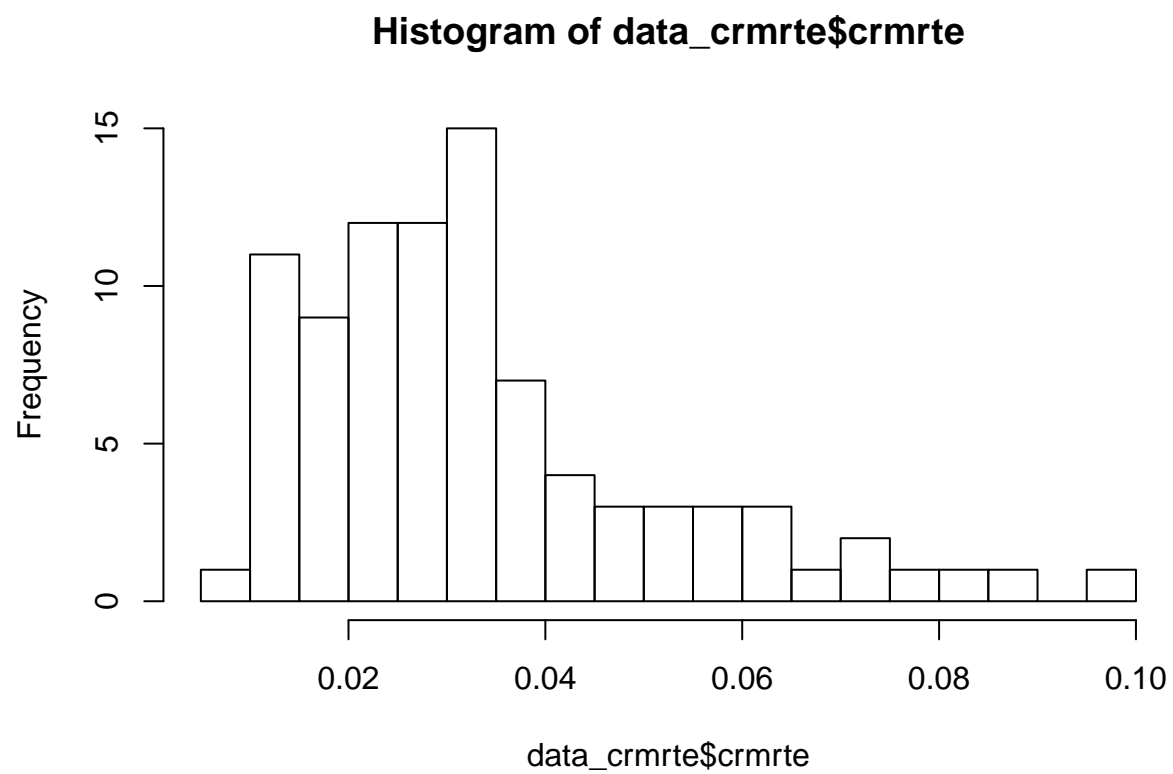
```
## [1] "numeric"
```

With 25 original variables in the data set the natural place to start is with the dependent variable, crmrte. To get a better sense of this variable, the distribution is graphed below.

```
quantile(data_crmrte$crmerte, c(0, .01, .05, .10, .25, .50, .75, .90, .95, .99, 1.0))
```

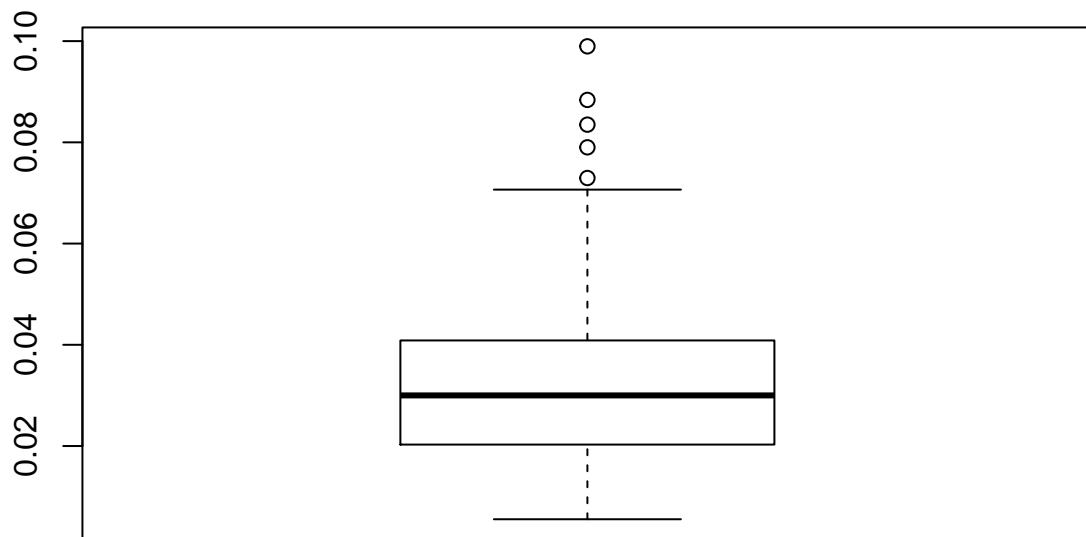
```
##          0%          1%          5%          10%          25%          50%          75%
## 0.00553320 0.01006330 0.01235660 0.01418007 0.02060425 0.03000200 0.04024925
##          90%          95%          99%         100%
## 0.06054659 0.07191830 0.08954881 0.09896590
```

```
hist(data_crmrte$crmerte,breaks=20)
```

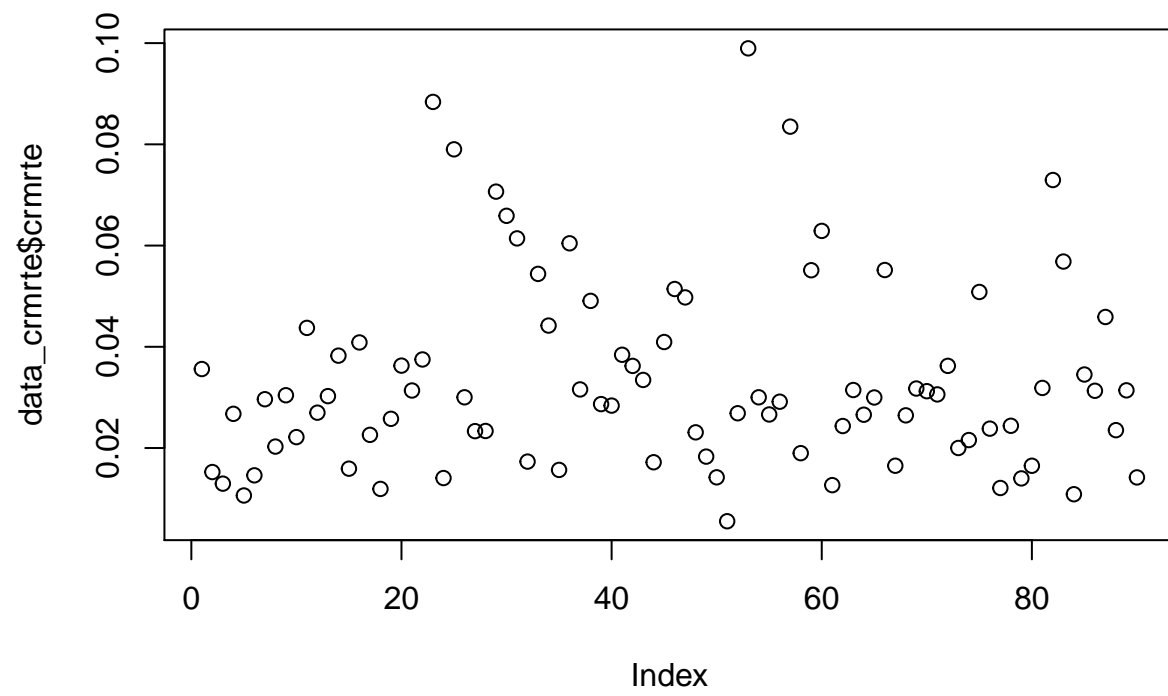


```
boxplot(data_crmrte$crmrte, main="Boxplot of crmrte")
```

**Boxplot of crrmrte**

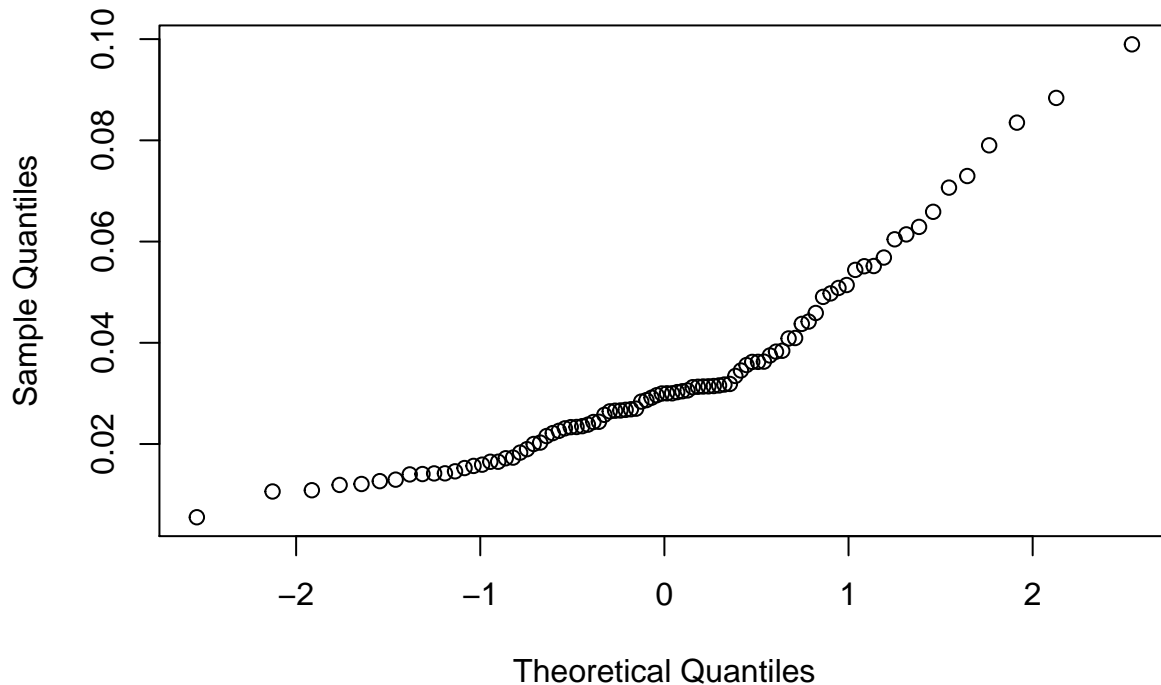


```
plot(data_crrmrte$crrmrte)
```



```
qqnorm(data_crmrte$crmrite)
```

## Normal Q-Q Plot



```
shapiro.test(data_crmrte$crmrte) # Shapiro-wilk test confirms non-normality
```

```
##
##  Shapiro-Wilk normality test
##
## data:  data_crmrte$crmrte
## W = 0.89162, p-value = 1.741e-06
```

### Outlier Analysis

There are several outliers in the variable crmrte and the distribution is right skewed. We have ninety observations so non-normality is not a top concern but this distribution is not perfectly normal. we analyse outliers for crime rate that are  $> 1.96 \times \text{Std-dev}$  from the mean crime rate ( $\sim 0.07$ )

The largest outliers (6 counties) on the right side of the distribution are examined. There are some insights:  
 1. 4 of out of the 6 outliers are in Urban areas  
 2. The average of density for the outlier is 3X the average density for the sample  
 3. Esp data ppt 53 which has the highest crime rate and is Urban

This is not very surprising as we expect urban areas with high density of population to have more crimes. we will consider the treatment of these outlier in later part of the report.

```
data_crmrte[data_crmrte$crmrte > 0.07,]
```

```
##      county year   crmrte  prbarr  prbconv  prbpris avgsen      polpc  density
## 23      51   87 0.0883849 0.155248 0.259833 0.407628  11.93 0.00190802 3.9345510
## 25      55   87 0.0790163 0.224628 0.207831 0.304348  13.57 0.00400962 0.5115089
```



```
## 29      63      87 0.0706599 0.133225 0.459216 0.363636 11.51 0.00237609 5.6744967
## 53     119      87 0.0989659 0.149094 0.347800 0.486183  7.13 0.00223135 8.8276520
## 57     129      87 0.0834982 0.236601 0.393413 0.415158  9.57 0.00255849 6.2864866
## 82     181      87 0.0729479 0.182590 0.343023 0.548023  7.06 0.00172948 1.5702811
##      taxpc west central urban pctmin80      wcon      wtuc      wtrd      wfir
## 23 35.69936    0      0      1 37.77920 283.6695 412.4720 213.7524 324.8357
## 25 119.76145    0      0      0  6.49622 309.5238 445.2762 189.7436 284.5933
## 29  50.19918    0      1      1 38.22300 349.3267 548.9865 238.9154 435.1107
## 53  75.67243    0      1      1 28.54600 436.7666 548.3239 354.6761 509.4655
## 57  67.67963    0      0      1 23.04410 315.5760 392.0999 220.4530 363.2880
## 82  27.59179    0      1      0 44.62830 244.8362 365.4716 279.2273 325.0271
##      wser  wmfgr  wfed  wsta  wloc      mix  pctymle
## 23 257.3344 441.72 433.94 367.34 333.71 0.10474319 0.14223780
## 25 221.3903 319.21 338.91 361.68 326.08 0.08437271 0.07613807
## 29 391.3081 646.85 563.77 415.51 362.58 0.07585382 0.09468981
## 53 354.3007 494.30 568.40 329.22 379.77 0.16869897 0.07916495
## 57 292.7027 464.49 548.49 421.36 319.08 0.07871422 0.08109921
## 82 213.5822 290.69 453.53 317.23 286.45 0.10003893 0.07977433
```

We also look at the lower range of outliers and find only data pt 51 (county 115) which has crime rate < 0.01. This outlier has some significant outlier effects and will be explored further later in the report.

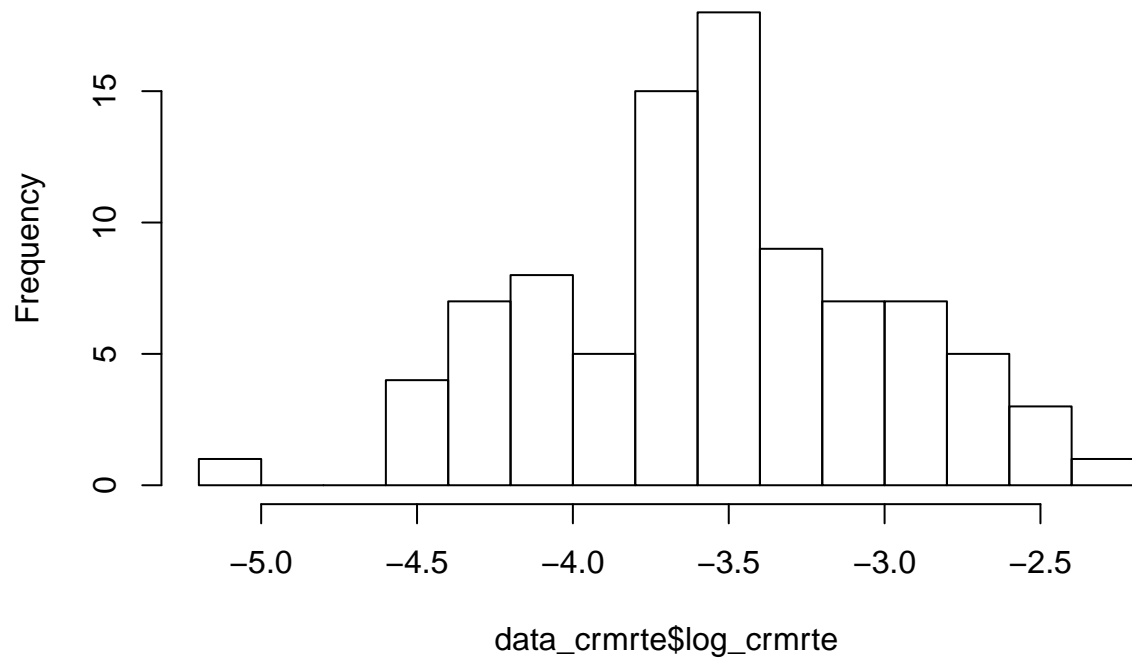
```
data_crmrte[data_crmrte$crmrte < 0.01,]
```

```
##      county year      crmrte prbarr prbconv prbpris avgsen      polpc      density
## 51      115      87 0.0055332 1.09091      1.5      0.5  20.7 0.00905433 0.3858093
##      taxpc west central urban pctmin80      wcon      wtuc      wtrd      wfir
## 51 28.1931    1      0      0 1.28365 204.2206 503.2351 217.4908 342.4658
##      wser  wmfgr  wfed  wsta  wloc mix  pctymle
## 51 245.2061 448.42 442.2 340.39 386.12 0.1 0.07253495
```

For campaign purposes, we want to predict crime. We want our candidate to be able to say that he or she can reduce crime in order to win votes. What is the most effective way to convey that? Using crime rate as it appears in the data set is using the level of crime rate and would suggest the following statement as a campaign slogan - “I can reduce crime to this rate by doing x, y, and z”. Transforming crime rate into the log of crime rate allows for the statement “I can reduce crime by n% by doing x, y, and z.” We find the latter more powerful and meaningful to voters since voters have no idea about the level of crime rates. In addition, we will show that the transformation of crime rate improves the normality and distribution of the variable, which will often reduce skew in the errors as well.

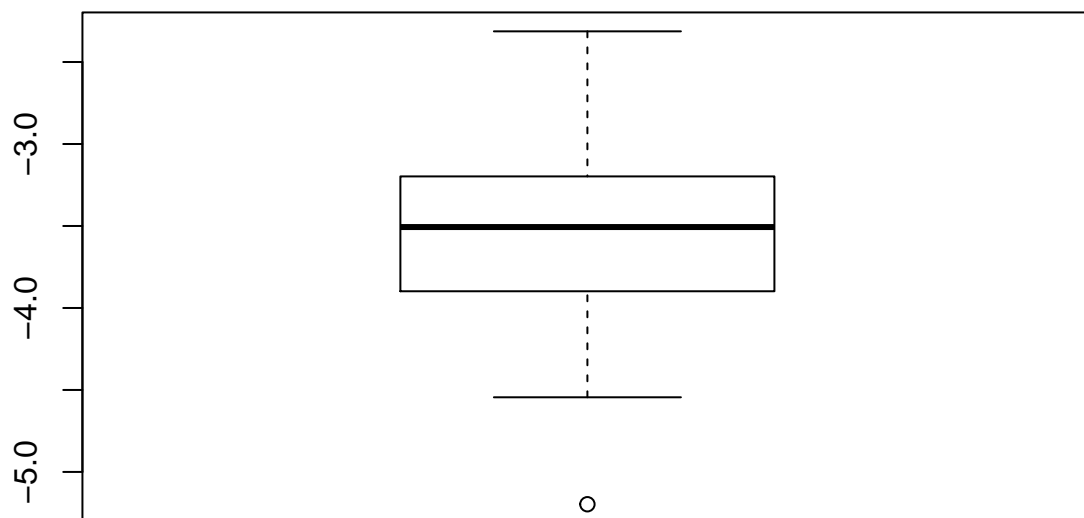
```
data_crmrte$log_crmrte <- log(data_crmrte$crmrte)
hist(data_crmrte$log_crmrte,breaks=20)
```

**Histogram of data\_crmrte\$log\_crmrte**

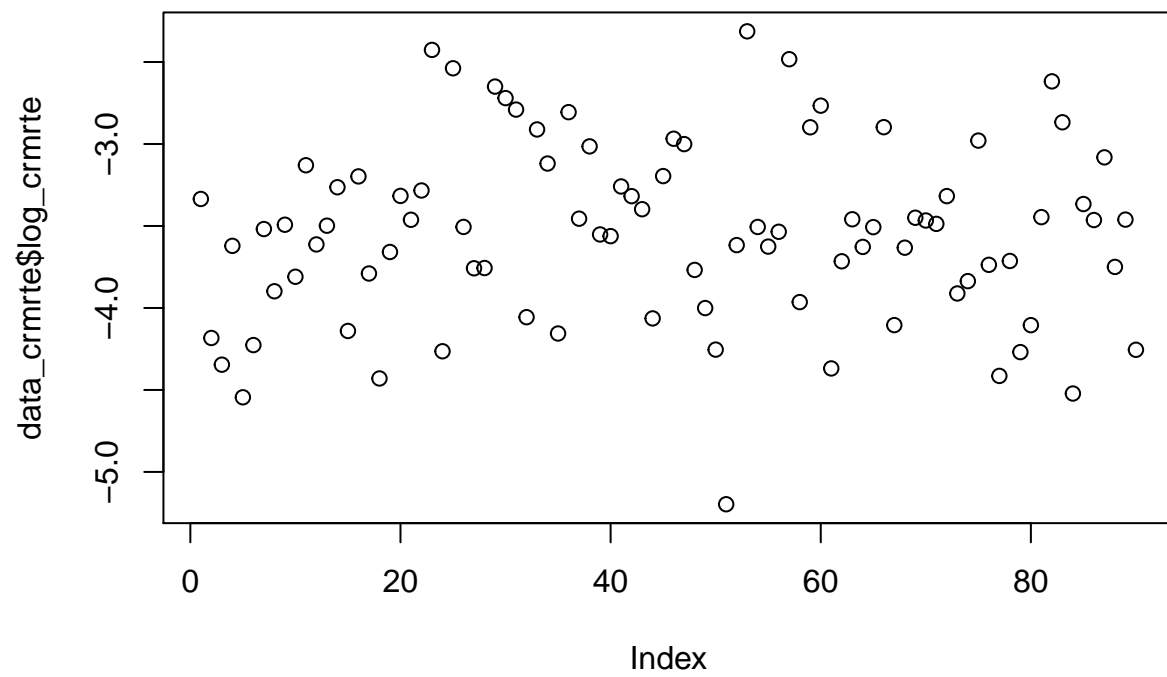


```
boxplot(data_crmrte$log_crmrte, main="Boxplot of log of crmrte")
```

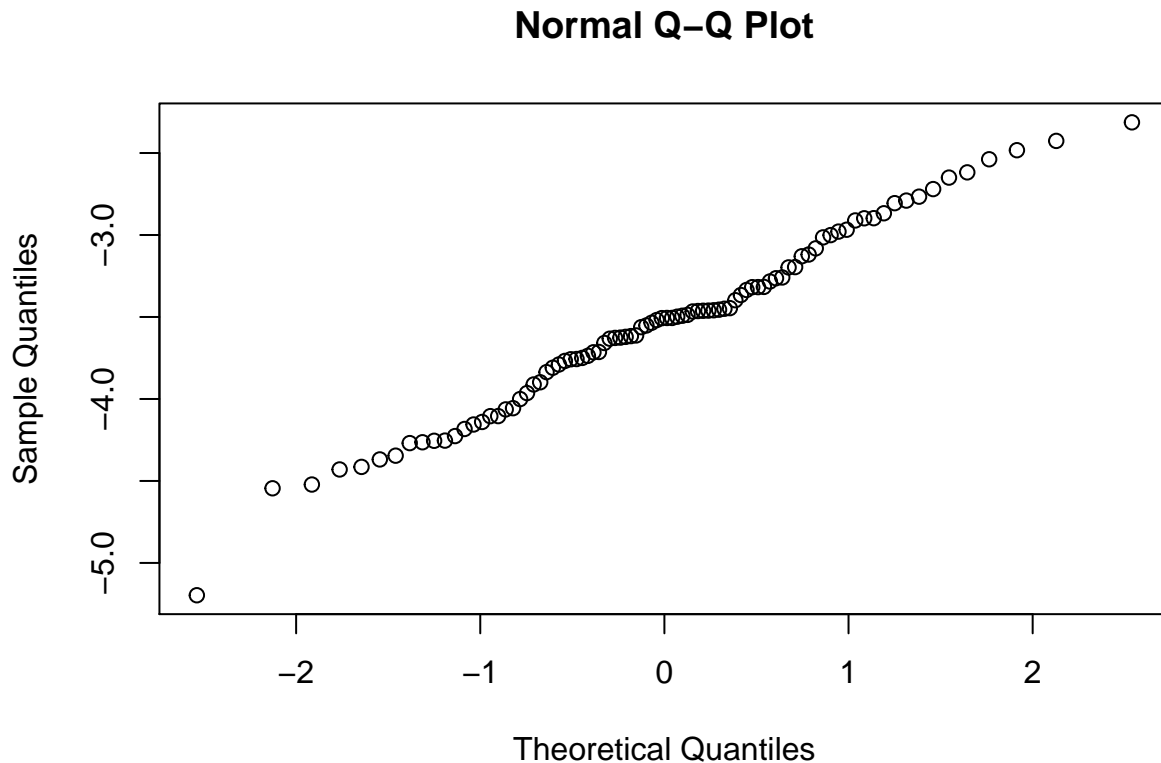
**Boxplot of log of crmrte**



```
plot(data_crmrte$log_crmrte)
```



```
qqnorm(data_crmrte$log_crmrte)
```



The histogram of the transformed crime rate is much more symmetrical and shows much less right skew. The box plot shows all of the outliers on the high end have been removed, though outlier 51 (countty 115) on the low end has been become more prominent.

The scatter plot looks much more normal, and the Q-Q plot is much closer to normal with the data points hugging the 45 degree line much more closely. Given the stronger argument for the political campaign and the benefits to normality we have chosen to model the tranformation of crime rate as opposed to crime rate.

### Groupings

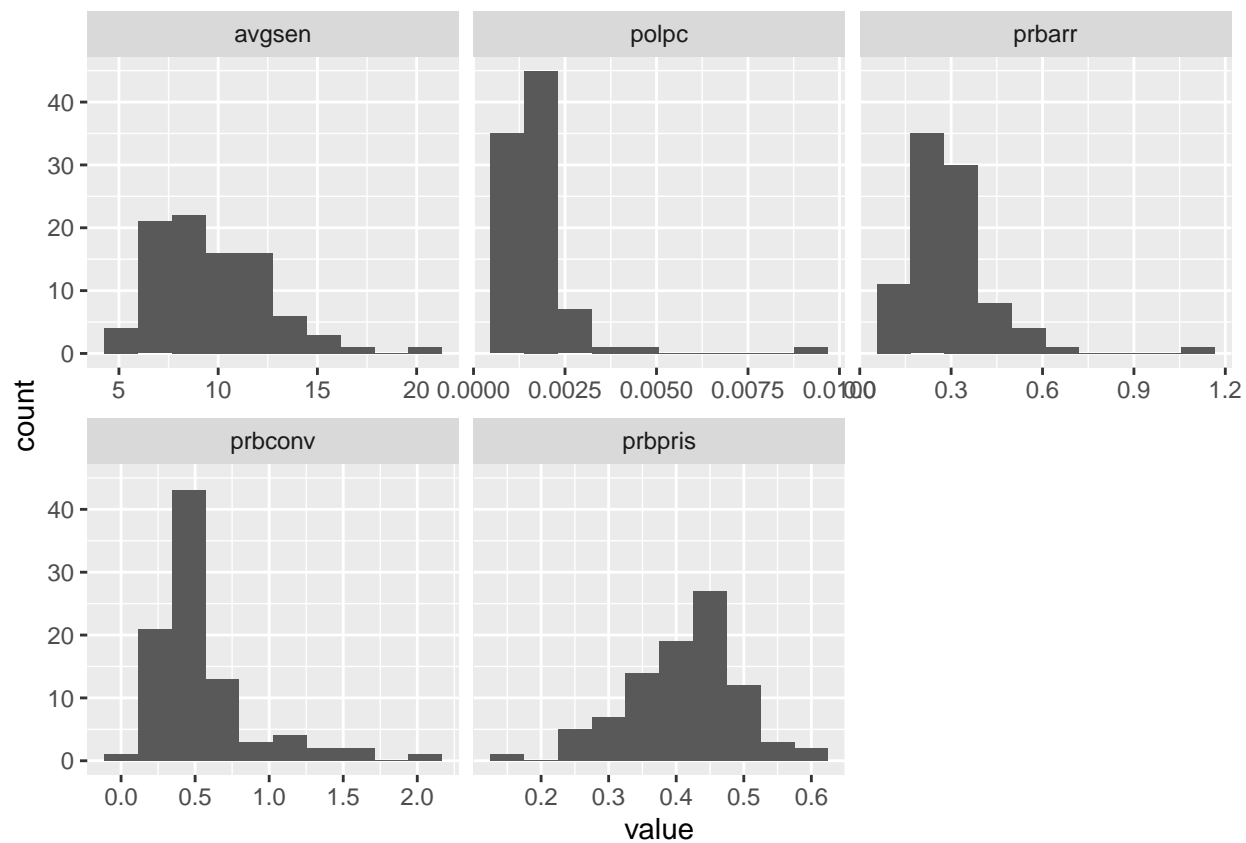
In order to digest the data in the data set we decided to group the variables into five groups: deterrent, wages, demographic, region, and urban. We performed exploratory data analysis on all of these variables.

The group is deterrent data. As cited in the original paper, these variables were hypothesized to reduce crime rate through disincentivizing crime. Essentially, as the probability of getting caught increases, criminals' desire to commit crimes decreases.

### Deterrent Data

```
deterrent_data <- data_crmrte[,c('prbarr', 'prbconv', 'prbpris', 'crmrate',
                                'avgsen', 'polpc')]

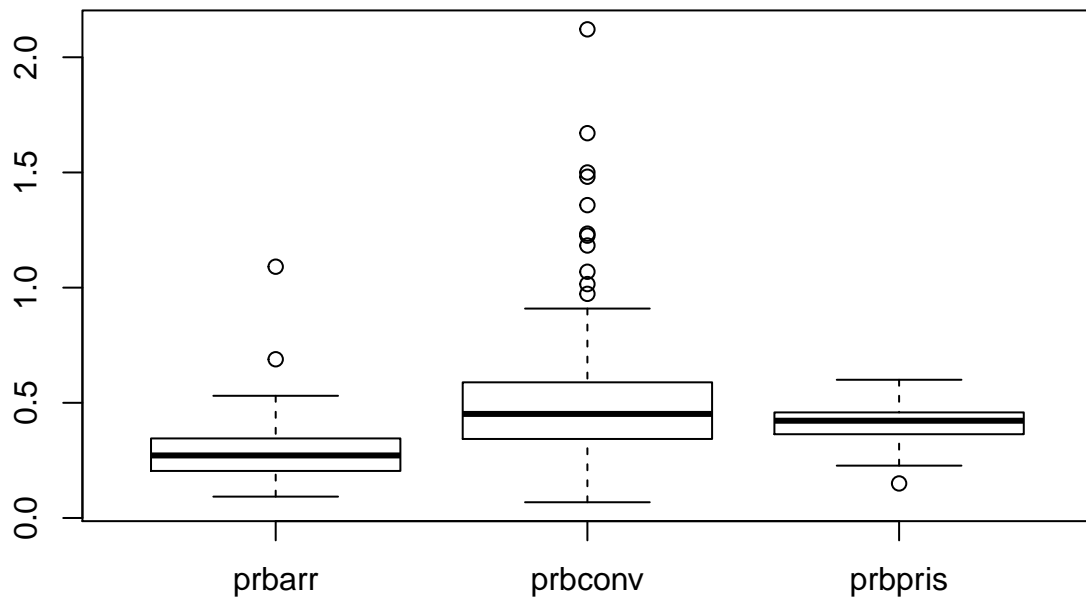
ggplot(gather(deterrent_data[,c('prbarr', 'prbconv', 'prbpris',
                                'avgsen', 'polpc')]), aes(value)) +
geom_histogram(bins = 10) + facet_wrap(~key, scales = 'free_x')
```



```
my_vars1 <- c("prbarr","prbconv","prbpris")
deterrent_data2 <- deterrent_data[my_vars1]
my_vars2 <- c("polpc")
deterrent_data3 <- deterrent_data[my_vars2]

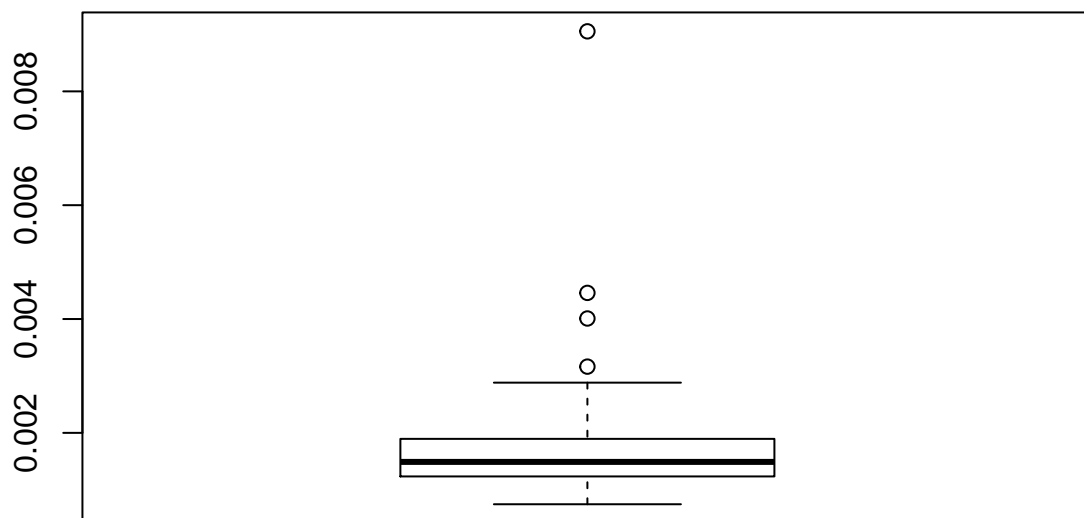
boxplot(deterrent_data2, main="Boxplot of prbarr, prbconv, prbpris")
```

**Boxplot of prbarr, prbconv, prbpris**



```
boxplot(deterrent_data3, main="Boxplot of polpc")
```

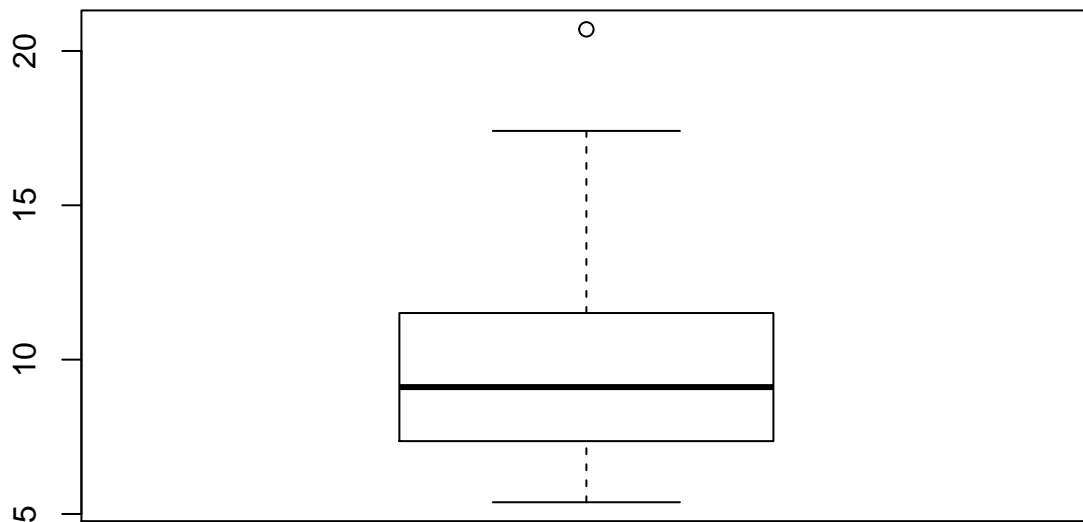
## Boxplot of polpc



```
boxplot(deterrent_data$avgsen, main="Boxplot of avgsen")
```



## Boxplot of avgsen

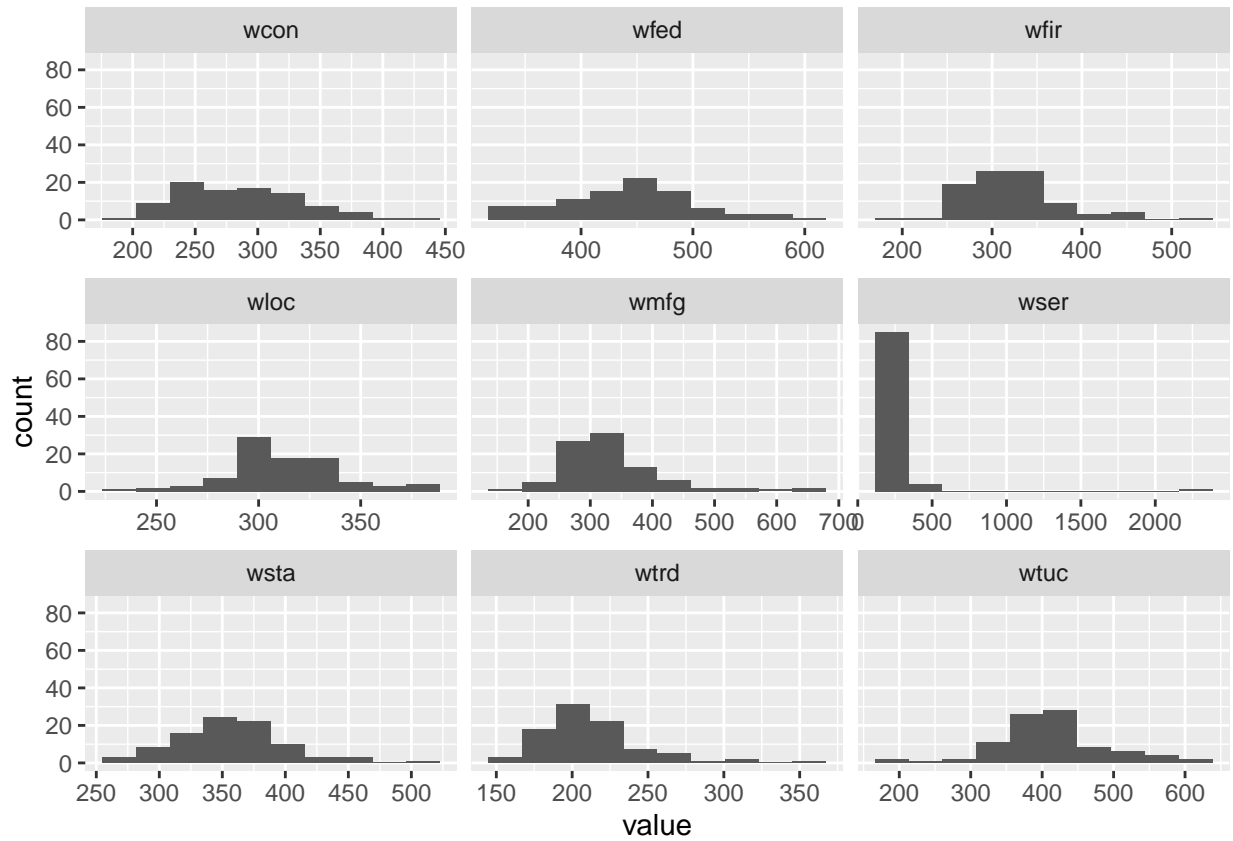


The first four histograms show right skew while prbpris shows left skew. The biggest outlier is observation 51. This observation has the lowest crime rate in the data set, obviously the highest polpc (police per capita), the highest avg sentence, the third highest prbconv, and the lowest pctmin80. This observation is likely to affect many of the regressions so it will need to be examined further. These variables are candidates to be transformed.

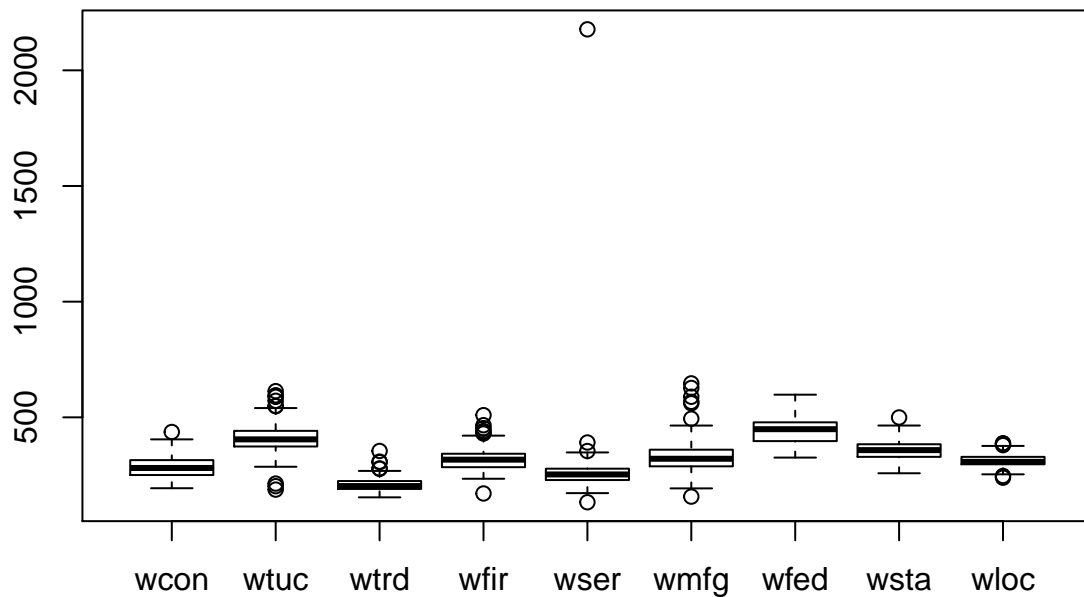
## Wages Data

```
#create a dataframe of just the wage variables
wages_data <- data_crmrte[,c('wcon','wtuc','wtrd','wfir', 'wser',
                             'wmfg','wfed', 'wsta', 'wloc')]

#plot histograms of just the wage variables
ggplot(gather(wages_data), aes(value)) +
  geom_histogram(bins = 10) +
  facet_wrap(~key, scales = 'free_x')
```



```
#generate boxplots of just the wage variables
boxplot(wages_data)
```



There is an obvious outlier for wser in data pt 84 (County 185) . The mean services wage across all the counties is \$275 ( with a std dev of 206) and 84 has wser of 2177 (~9sd from mean), which seems like a measurement or typographical error. The next highest average weekly wage in any sector is 646 versus the value of 2177. It is very possible that this data point might add measurement error and we will revisit this later.

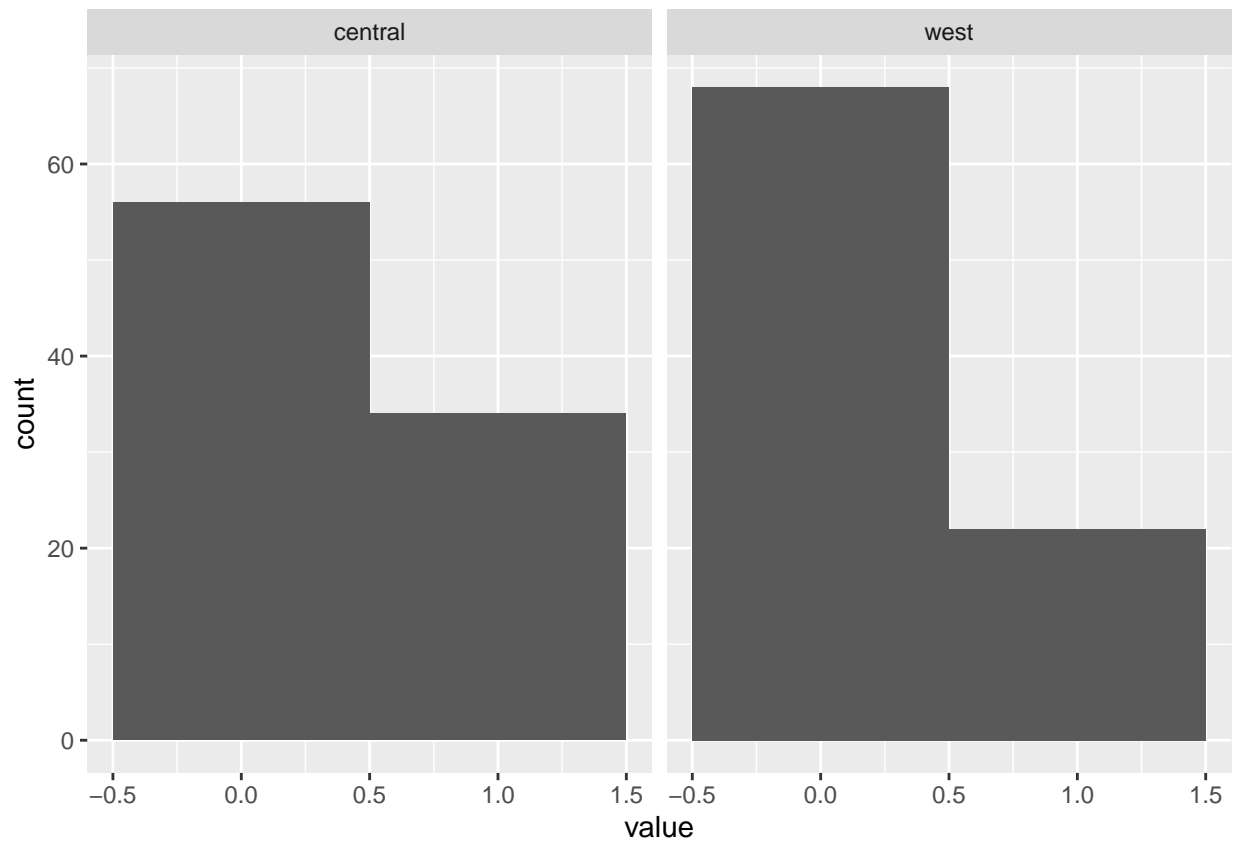
For now, we create an additional variable that is the median of all wage variables for each observation. If it conveys as much information, it has the benefit of increasing our degrees of freedom and removing the effect of the outlier.

```
data_crmrte$median_wage <- apply(data_crmrte[,c("wcon", "wtuc", "wtrd",
        "wfir", "wser", "wmfg",
        "wfed", "wsta", "wloc")],
        1, FUN=median, na.rm=TRUE)
```

Region Data\_\_

```
#create a dataframe of just the wage variables
dummies_data <- data_crmrte[,c('west', 'central')]

#plot histograms of just the dummy variables
ggplot(gather(dummies_data), aes(value)) +
  geom_histogram(bins = 2) +
  facet_wrap(~key)
```



```
#just a quick check that there is no overlap
region_check <- data_crmrte[which(data_crmrte$west == 1 && data_crmrte$central == 1)]

summary(region_check)
```

```
## < table of extent 0 x 0 >
```

The regions are broken up into central, west, and east. East is left out of the data set and its effect as the final level of the indicator variable will move to the intercept.

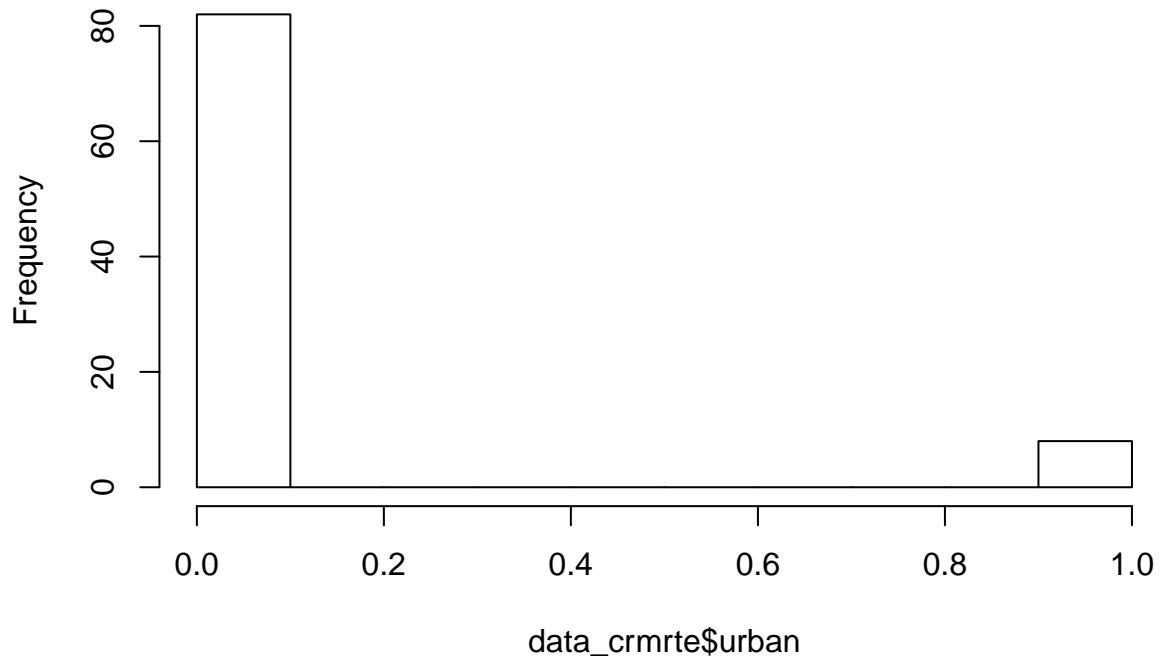
## Urban Data

```
#plot histograms of just the wage variables
sum(data_crmrte$urban) # There are only 8 Urban areas out of 90 counties
```

```
## [1] 8
```

```
hist(data_crmrte$urban)
```

## Histogram of data\_crmrte\$urban

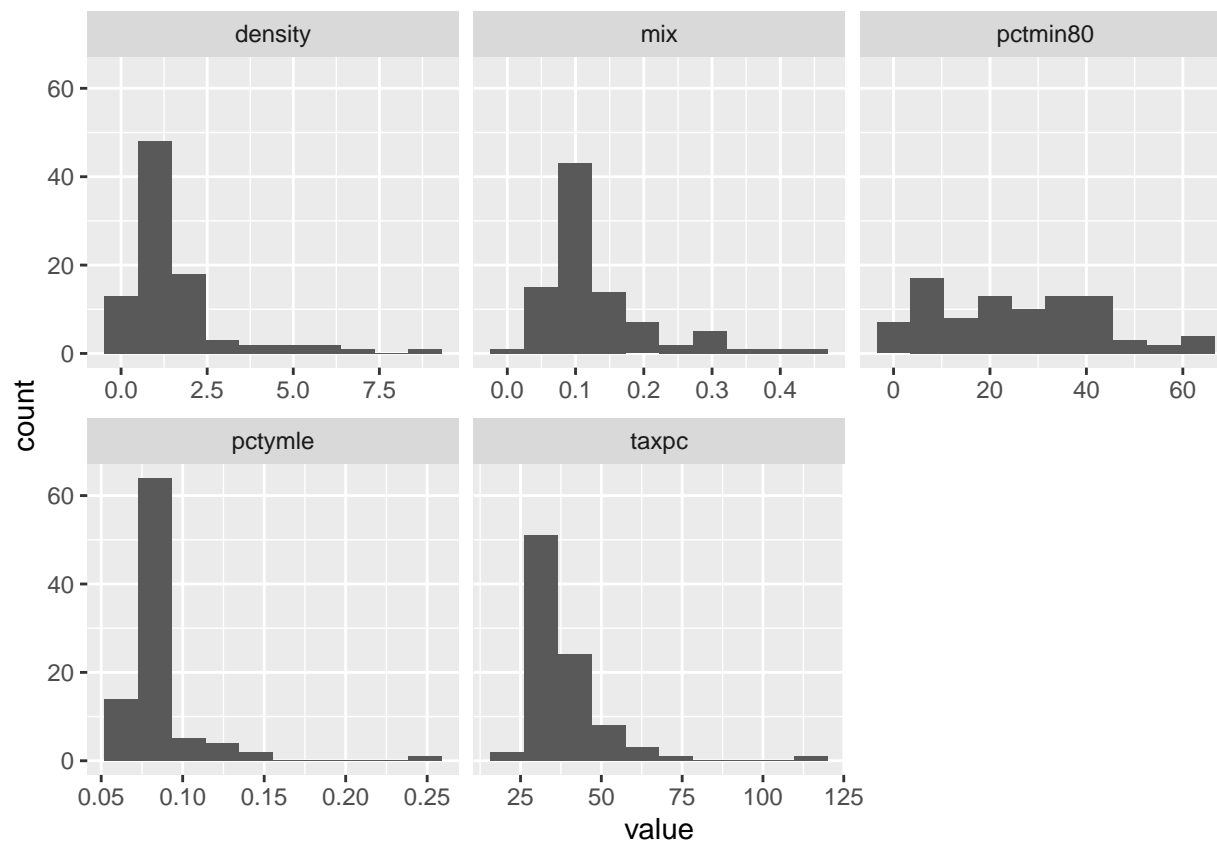


Urban did not fit into a great grouping so we left this variable on its own. A histogram shows that the state has relatively few urban counties, something to keep in mind when analyzing other variables such as density.

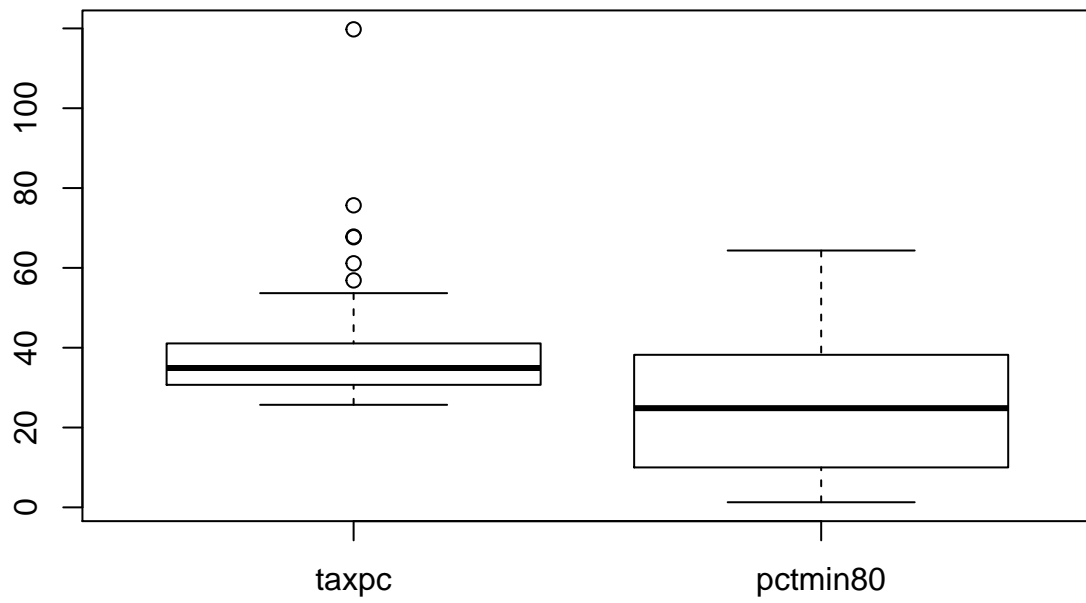
## Demographic Data

```
#create a dataframe of just the demographic variables
demographic_data <- data_crmrte[,c('density', 'taxpc', 'pctmin80',
                                   'mix', 'pctymle')]

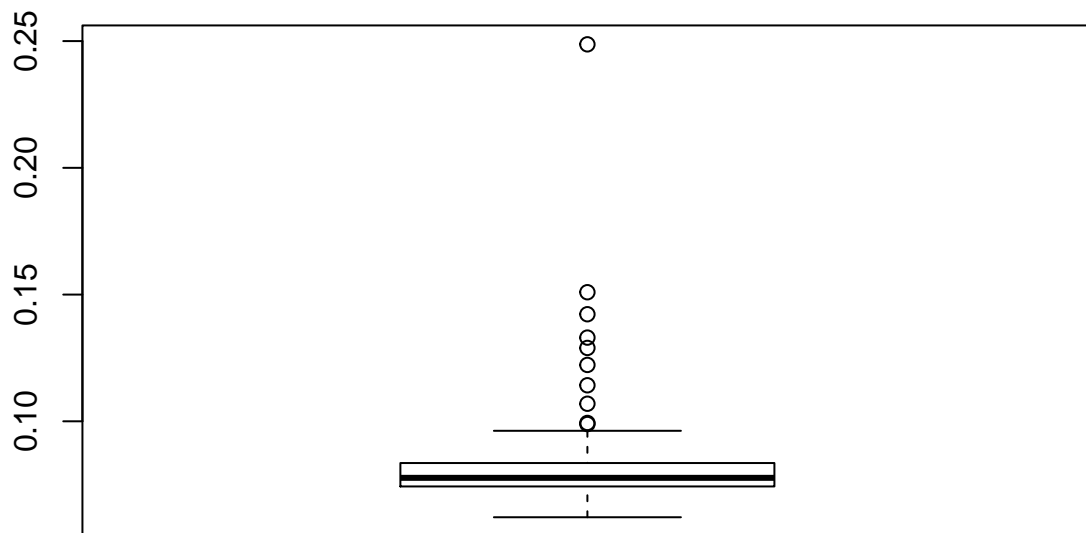
#plot histograms of just the demographic variables
ggplot(gather(demographic_data), aes(value)) +
  geom_histogram(bins = 10) +
  facet_wrap(~key, scales = 'free_x')
```



```
#Lots of skewed distributions above, particularly in pctymle and taxpc
#generate boxplots of just the demographic variables
demographic_data2 <- demographic_data[c("taxpc", "pctmin80")]
boxplot(demographic_data2)
```

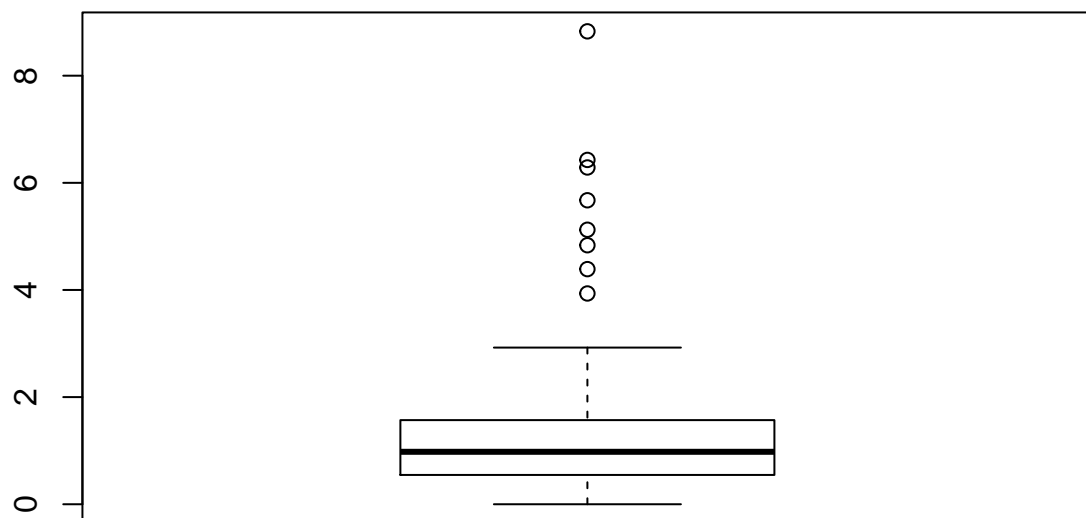


```
demographic_data3 <- demographic_data[c("pctymle")]  
boxplot(demographic_data3)
```

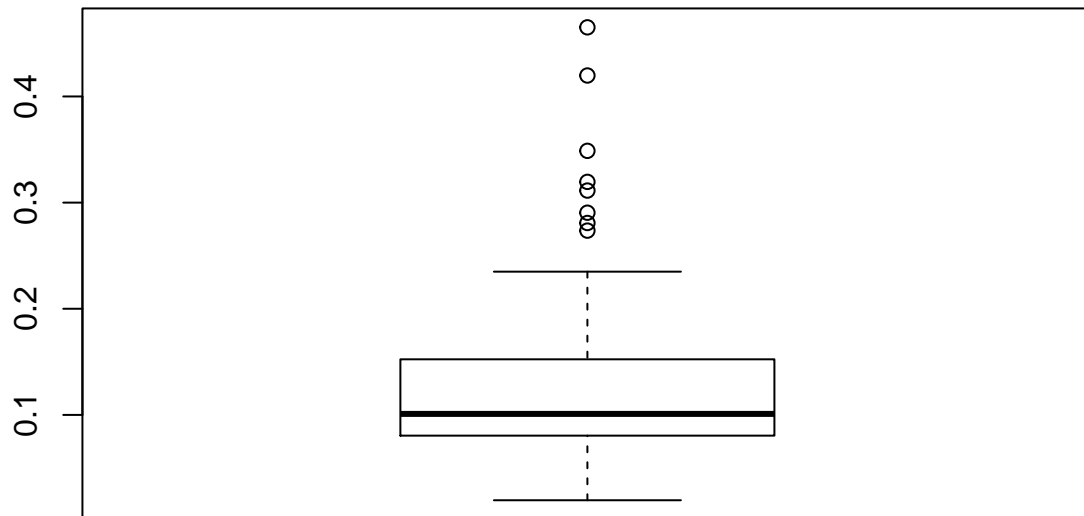


```
demographic_data4 <- demographic_data[c("density")]  
boxplot(demographic_data4)
```





```
demographic_data5 <- demographic_data[c("mix")]  
boxplot(demographic_data5)
```

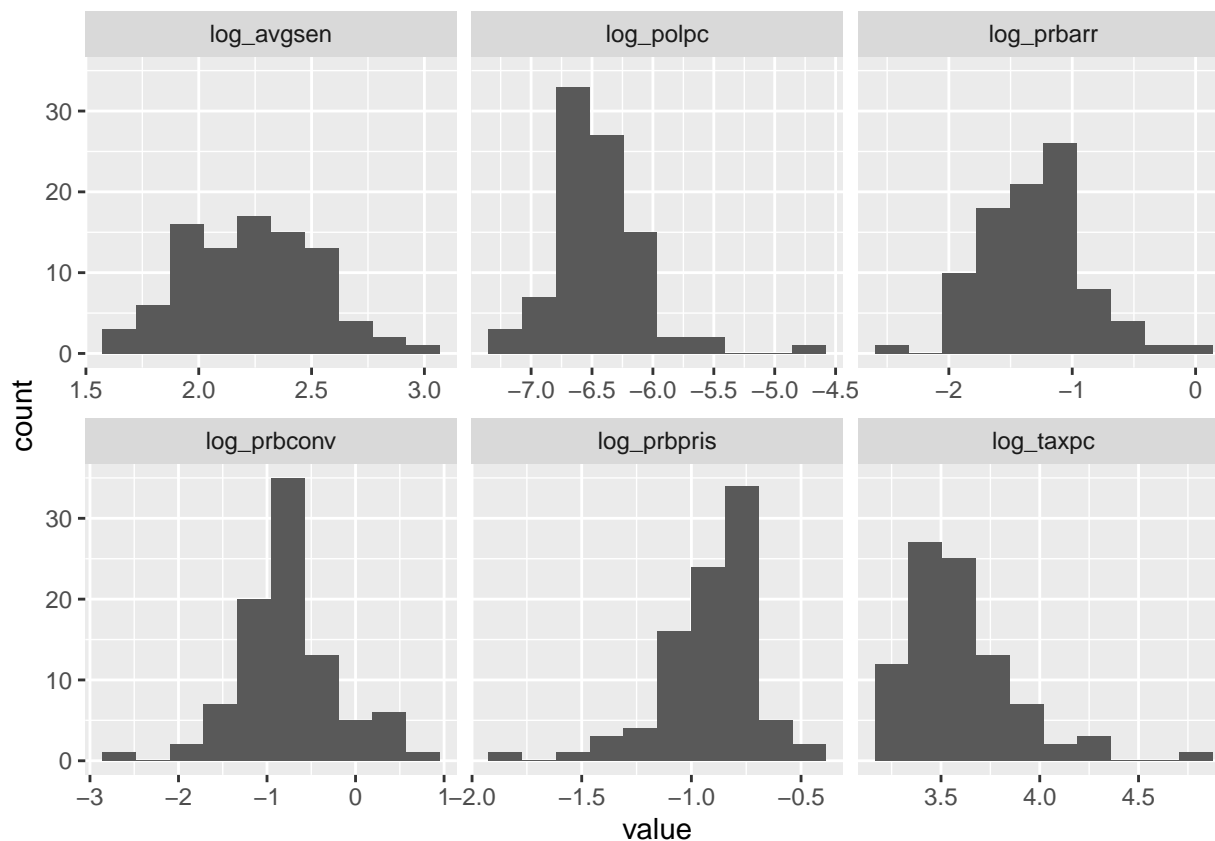


Once again we see a lot of right skewed distributions in the histograms and in the box plots.

After exploring all of the variables we decided to transform the other variables that are potentially under a politician's control - the deterrent variables. This gives us our final data set and so we can start running regressions.

```
data_crmrte$prbconv <- as.numeric(as.character(data_crmrte$prbconv))
data_crmrte$log_prbarr <- log(data_crmrte$prbarr)
data_crmrte$log_prbconv <- log(data_crmrte$prbconv)
data_crmrte$log_prbpris <- log(data_crmrte$prbpris)
data_crmrte$log_avgsen <- log(data_crmrte$avgsen)
data_crmrte$log_polpc <- log(data_crmrte$polpc)
data_crmrte$log_taxpc <- log(data_crmrte$taxpc)

#plot histograms of just the demographic variables
ggplot(gather(data_crmrte[,c('log_prbarr', 'log_prbconv', 'log_prbpris', 'log_avgsen', 'log_polpc', 'log_taxpc')],
  geom_histogram(bins = 10) +
  facet_wrap(~key, scales = 'free_x')
```



Though the distribution of the variables still exhibits skew, the skew does seem to be reduced.

### Log Tranformed Dependent Variable Comparison

In order to settle on the final data set we compare an all-in log-log model with an all-in log-linear to see which dependent variables are more suitable.

```
##### Initial Models #####
all_in_model <- lm(crmrte ~ prbarr + prbconv + prbpris
  + avgsen + polpc + density
  + taxpc + west + central + urban + pctmin80 + wcon
  + wtuc + wtrd + wfir + wser + wmfg
  + wfed + wsta + wloc
  + mix + pctymle,
  data = data_crmrte)
se.all_in_model = sqrt(diag(vcovHC(all_in_model)))
coeftest(all_in_model, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.3853e-02 3.0755e-02  0.4504 0.6538622
## prbarr      -5.1466e-02 1.5689e-02 -3.2805 0.0016467 **
## prbconv     -1.8633e-02 6.5853e-03 -2.8295 0.0061464 **
```

```
## prbpris      3.1727e-03  1.3586e-02  0.2335 0.8160642
## avgsgen     -3.9858e-04  5.5361e-04 -0.7200 0.4740570
## polpc       6.9679e+00  2.9536e+00  2.3591 0.0212406 *
## density     5.3314e-03  1.4895e-03  3.5793 0.0006464 ***
## taxpc       1.6240e-04  2.8408e-04  0.5717 0.5694537
## west        -2.5652e-03  4.4698e-03 -0.5739 0.5679579
## central     -4.2416e-03  3.7423e-03 -1.1334 0.2610725
## urban       -9.6498e-05  8.2752e-03 -0.0117 0.9907307
## pctmin80    3.2542e-04  1.3849e-04  2.3497 0.0217429 *
## wcon        2.3025e-05  3.2876e-05  0.7004 0.4861334
## wtuc        6.1914e-06  1.9862e-05  0.3117 0.7562178
## wtrd        2.8767e-05  8.7294e-05  0.3295 0.7427756
## wfir        -3.5455e-05  3.5699e-05 -0.9932 0.3242068
## wser        -1.7158e-06  9.9447e-05 -0.0173 0.9862856
## wmfg        -8.9675e-06  1.7469e-05 -0.5133 0.6094087
## wfed        2.9075e-05  3.7780e-05  0.7696 0.4442480
## wsta        -2.2302e-05  3.6828e-05 -0.6056 0.5468431
## wloc        1.4456e-05  8.5367e-05  0.1693 0.8660410
## mix         -1.8693e-02  2.2922e-02 -0.8155 0.4176761
## pctymle     1.0125e-01  4.7826e-02  2.1170 0.0379748 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(all_in_model)
```

```
## [1] -585.5858
```

```
BIC(all_in_model)
```

```
## [1] -525.5904
```

```
all_in_model_log_level <- lm(log_crmrte ~ prbarr + prbconv + prbpris
+ avgsgen + polpc + density
+ taxpc + west + central + urban
+ pctmin80 + wcon
+ wtuc + wtrd + wfir + wser + wmfg
+ wfed + wsta + wloc
+ mix + pctymle,
data = data_crmrte)
se.all_in_model_log_level = sqrt(diag(vcovHC(all_in_model_log_level)))
coeftest(all_in_model_log_level, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept) -4.0261e+00 8.4822e-01 -4.7466 1.128e-05 ***
## prbarr      -1.8891e+00 3.7955e-01 -4.9773 4.770e-06 ***
## prbconv     -6.5603e-01 1.7443e-01 -3.7611 0.0003579 ***
## prbpris     -9.3077e-02 3.9921e-01 -0.2332 0.8163542
## avgsgen     -7.8769e-03 1.6125e-02 -0.4885 0.6267962
## polpc       1.5484e+02 8.6523e+01  1.7895 0.0780510 .
```

```
## density      1.1653e-01  5.4037e-02  2.1566 0.0346326 *
## taxpc        3.3224e-03  7.2890e-03  0.4558 0.6500012
## west         -1.1492e-01  1.2509e-01 -0.9187 0.3615403
## central      -1.0078e-01  9.2053e-02 -1.0948 0.2775232
## urban        -1.6923e-01  2.2872e-01 -0.7399 0.4619535
## pctmin80     9.9770e-03  3.0480e-03  3.2733 0.0016833 **
## wcon         4.6001e-04  8.3564e-04  0.5505 0.5838140
## wtuc         1.0174e-04  6.0187e-04  0.1690 0.8662750
## wtrd         2.5964e-04  1.7638e-03  0.1472 0.8834136
## wfir        -1.1015e-03  1.1960e-03 -0.9210 0.3603557
## wser        -1.3142e-04  1.5060e-03 -0.0873 0.9307193
## wmfg        -2.0528e-04  5.1630e-04 -0.3976 0.6921878
## wfed         2.3405e-03  1.0820e-03  2.1632 0.0340968 *
## wsta        -1.1357e-03  8.9769e-04 -1.2651 0.2102213
## wloc         5.8983e-04  2.4003e-03  0.2457 0.8066400
## mix         -2.3924e-01  6.2632e-01 -0.3820 0.7036869
## pctymle      2.7706e+00  1.4330e+00  1.9334 0.0574191 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(all_in_model_log_level)
```

```
## [1] 21.354
```

```
BIC(all_in_model_log_level)
```

```
## [1] 81.34943
```

```
all_in_model_log_log <- lm(log_crmrte ~ log_prbarr + log_prbconv
+ log_prbpris + log_avgsen + log_polpc
+ density+ log_taxpc + west + central
+ urban + pctmin80 + wcon
+ wtuc + wtrd + wfir
+ wser + wmfg + wfed + wsta + wloc
+ mix + pctymle,
data = data_crmrte)
se.all_in_model_log_log = sqrt(diag(vcovHC(all_in_model_log_log)))
coeftest(all_in_model_log_log, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -3.36882669  2.97497990 -1.1324 0.261508
## log_prbarr   -0.52143620  0.16459898 -3.1679 0.002313 **
## log_prbconv  -0.33101341  0.15365522 -2.1543 0.034820 *
## log_prbpris  -0.06569465  0.19741379 -0.3328 0.740342
## log_avgsen   -0.19652151  0.18205821 -1.0794 0.284261
## log_polpc     0.29132794  0.27176129  1.0720 0.287567
## density       0.12320127  0.06040422  2.0396 0.045335 *
## log_taxpc     0.06158051  0.30979897  0.1988 0.843040
## west         -0.18453792  0.16353910 -1.1284 0.263174
```

```
## central      -0.10789292  0.09991865 -1.0798 0.284100
## urban        -0.14767055  0.26670745 -0.5537 0.581641
## pctmin80      0.00956927  0.00358175  2.6717 0.009466 **
## wcon          0.00078953  0.00090745  0.8701 0.387376
## wtuc          0.00010106  0.00075559  0.1337 0.894001
## wtrd          0.00029022  0.00177967  0.1631 0.870952
## wfir          -0.00108230  0.00125937 -0.8594 0.393186
## wser          -0.00042887  0.00096365 -0.4451 0.657718
## wmfg          -0.00014147  0.00061356 -0.2306 0.818343
## wfed          0.00224918  0.00136611  1.6464 0.104363
## wsta          -0.00102039  0.00106131 -0.9614 0.339787
## wloc          0.00017815  0.00261968  0.0680 0.945986
## mix           -0.44834658  0.77846459 -0.5759 0.566587
## pctymle       2.00755501  2.60186976  0.7716 0.443075
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(all_in_model_log_log)
```

```
## [1] 44.17803
```

```
BIC(all_in_model_log_log)
```

```
## [1] 104.1735
```

```
#Not comparing r-squared, just looking at significant variables
stargazer(all_in_model, all_in_model_log_level,
  all_in_model_log_log,
  type = "text", omit.stat = "f",
  se = list(se.all_in_model, se.all_in_model_log_level,
    se.all_in_model_log_log),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               crmrte      log_crmrte
##                               (1)         (2)         (3)
## -----
## prbarr                -0.051**   -1.889***
##                       (0.016)    (0.380)
##
## prbconv                -0.019**   -0.656***
##                       (0.007)    (0.174)
##
## prbpris                 0.003      -0.093
##                       (0.014)    (0.399)
##
## avgsen                 -0.0004     -0.008
##                       (0.001)    (0.016)
##
```

## polpc	6.968*	154.835	
##	(2.954)	(86.523)	
##			
## log_prbarr			-0.521**
##			(0.165)
##			
## log_prbconv			-0.331*
##			(0.154)
##			
## log_prbpris			-0.066
##			(0.197)
##			
## log_avgsen			-0.197
##			(0.182)
##			
## log_polpc			0.291
##			(0.272)
##			
## density	0.005***	0.117*	0.123*
##	(0.001)	(0.054)	(0.060)
##			
## taxpc	0.0002	0.003	
##	(0.0003)	(0.007)	
##			
## log_taxpc			0.062
##			(0.310)
##			
## west	-0.003	-0.115	-0.185
##	(0.004)	(0.125)	(0.164)
##			
## central	-0.004	-0.101	-0.108
##	(0.004)	(0.092)	(0.100)
##			
## urban	-0.0001	-0.169	-0.148
##	(0.008)	(0.229)	(0.267)
##			
## pctmin80	0.0003*	0.010**	0.010**
##	(0.0001)	(0.003)	(0.004)
##			
## wcon	0.00002	0.0005	0.001
##	(0.00003)	(0.001)	(0.001)
##			
## wtuc	0.00001	0.0001	0.0001
##	(0.00002)	(0.001)	(0.001)
##			
## wtrd	0.00003	0.0003	0.0003
##	(0.0001)	(0.002)	(0.002)
##			
## wfir	-0.00004	-0.001	-0.001
##	(0.00004)	(0.001)	(0.001)
##			
## wser	-0.00000	-0.0001	-0.0004
##	(0.0001)	(0.002)	(0.001)
##			

```
## wmfg          -0.00001  -0.0002  -0.0001
##              (0.00002)  (0.001)  (0.001)
##
## wfed          0.00003   0.002*   0.002
##              (0.00004)  (0.001)  (0.001)
##
## wsta         -0.00002   -0.001   -0.001
##              (0.00004)  (0.001)  (0.001)
##
## wloc          0.00001   0.001   0.0002
##              (0.0001)  (0.002)  (0.003)
##
## mix          -0.019    -0.239   -0.448
##              (0.023)   (0.626)  (0.778)
##
## pctymle       0.101*    2.771    2.008
##              (0.048)   (1.433)  (2.602)
##
## Constant      0.014    -4.026*** -3.369
##              (0.031)   (0.848)  (2.975)
##
## -----
## Observations      90        90        90
## R2                0.855      0.854      0.812
## Adjusted R2       0.807      0.806      0.750
## Residual Std. Error (df = 67) 0.008      0.242      0.275
## =====
## Note:                *p<0.05; **p<0.01; ***p<0.001
```

```
# #r-squared comparison of final two models
# yhat_level_level <- predict(all_in_model)
#
# #get the coefficients
# for (b in coef(all_in_model_log_level))
# {
#   beta_log_level <- c(beta_log_level, b)
# }
# #calculate the predictions
# for (b in coef(all_in_model_log_level))
# {
#   beta_log_level <- c(beta_log_level, b)
# }
#
# data_crmrte$log_level_yhat <- exp(--3.36882669
#
#           -0.5214362*data_crmrte$log_prbarr
#           -0.33101341*data_crmrte$log_prbconv
#           -0.06569465*data_crmrte$log_prbpris
#           -0.19652151*data_crmrte$log_avgsen
#           +0.29132794*data_crmrte$log_polpc
#           +0.12320127*data_crmrte$density
#           +0.06158051*data_crmrte$log_taxpc
#           -0.18453792*data_crmrte$west
#           -0.10789292*data_crmrte$central
#           -0.14767055*data_crmrte$urban
```



```

# +0.00956927*data_crmrte$pctmin80
# +0.00078953*data_crmrte$wcon
# +0.00010106*data_crmrte$wtuc
# +0.00029022*data_crmrte$wtrd
# -0.0010823*data_crmrte$wfir
# -0.00042887*data_crmrte$wser
# -0.00014147*data_crmrte$wmfg
# +0.00224918*data_crmrte$wfed
# -0.00102039*data_crmrte$wsta
# +0.00017815*data_crmrte$wloc
# -0.44834658*data_crmrte$mix
# +2.00755501*data_crmrte$pctymle
# )
# r_squared_level_level <- cor(data_crmrte$crmte, yhat_level_level)
# r_squared_log_level <- cor(data_crmrte$crmte, data_crmrte$log_level_yhat)
# (r_squared_level_level)
# (r_squared_log_level)

```

## Model 1: Simple Model

In order to create a simple model we decided to build using a bottom up approach. We looked at a correlation matrix

Anyone know how to print this better?

```
cor(data_crmrte$log_crmrte,data_crmrte)
```

```
## Warning in cor(data_crmrte$log_crmrte, data_crmrte): the standard deviation is
## zero
```

```
##      county year   crmrte   prbarr   prbconv   prbpris   avgsen
## [1,] 0.02376789  NA 0.9415465 -0.4727669 -0.4468136 0.02147024 -0.04936931
##      polpc  density   taxpc    west  central   urban  pctmin80
## [1,] 0.0104058 0.6330234 0.3583234 -0.4143996 0.1847192 0.4914645 0.2329182
##      wcon   wtuc   wtrd   wfir   wser   wmfg   wfed
## [1,] 0.3937149 0.2014649 0.3937924 0.2932426 -0.113128 0.3075373 0.5233058
##      wsta   wloc   mix   pctymle log_crmrte median_wage log_prbarr
## [1,] 0.1697021 0.2885668 -0.1247344 0.2781547      1      0.454422 -0.4357539
##      log_prbconv log_prbpris log_avgsen log_polpc log_taxpc
## [1,] -0.3724961 0.06960729 0.02341717 0.2845396 0.3398432

```

/ In the above correlation matrix, focusing on the correlations between the log\_crmrte and all other variables, density has the highest correlation. This variable makes intuitive sense. As a single variable it might encompass a lot of other factors. Lower income people with more incentive to commit crimes tend to live in more highly populated areas. Below is the simple regression.

```

simple_regression_model <- lm(log_crmrte ~ density, data = data_crmrte)
se.simple_regression_model = sqrt(diag(vcovHC(simple_regression_model)))
coeftest(simple_regression_model, vcov = vcovHC)

```

```
##
## t test of coefficients:
##
##           Estimate Std. Error  t value Pr(>|t|)
## (Intercept) -3.869488   0.068563 -56.4366 < 2e-16 ***
## density      0.228298   0.030439  7.5003  4.8e-11 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(simple_regression_model)
```

```
## [1] 106.2991
```

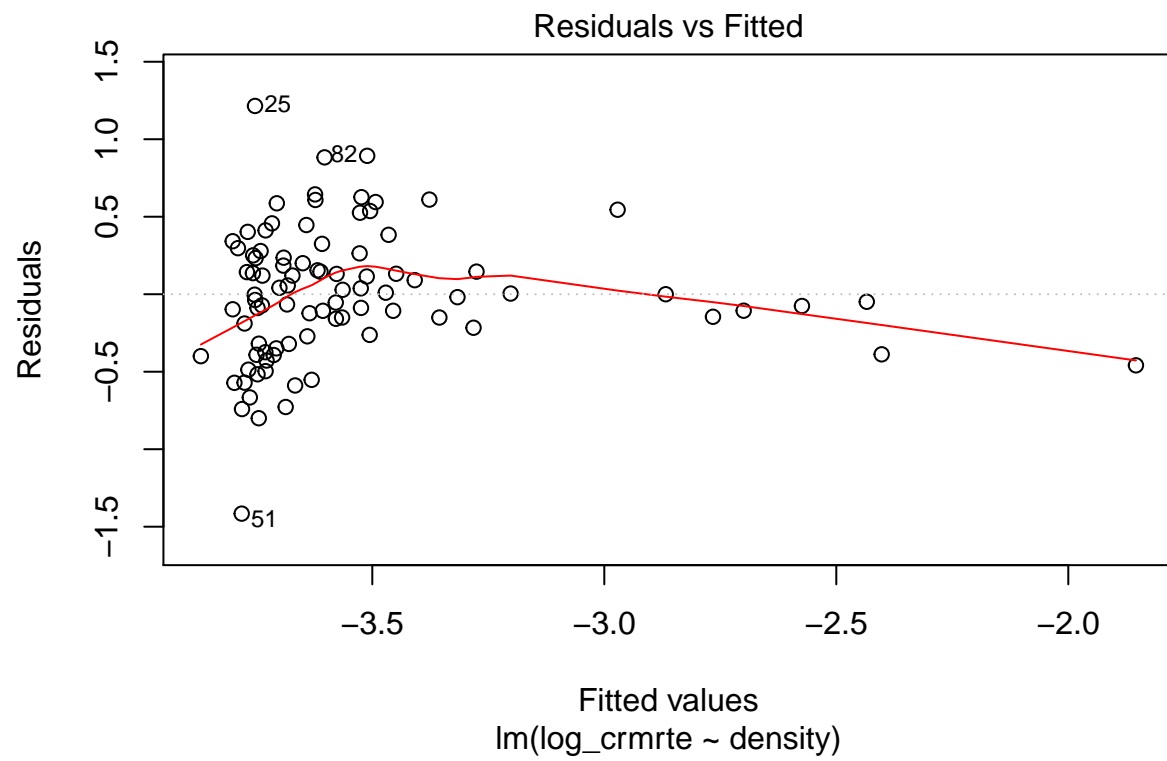
```
BIC(simple_regression_model)
```

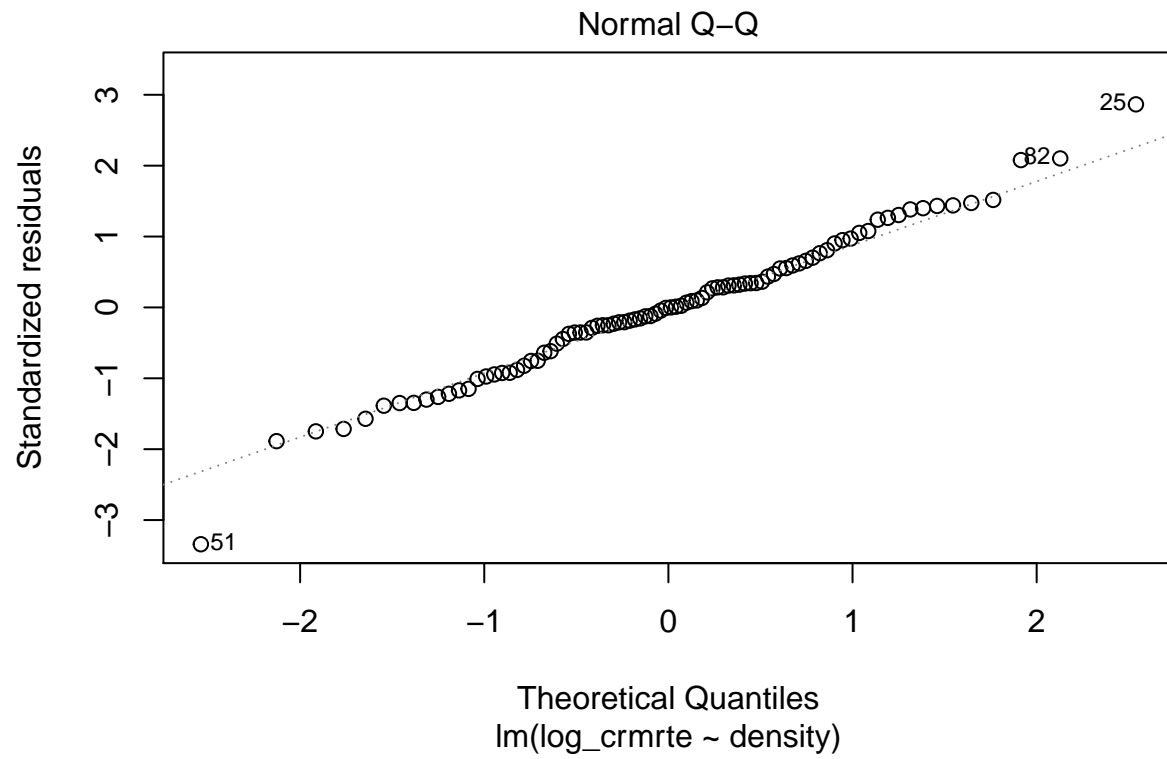
```
## [1] 113.7985
```

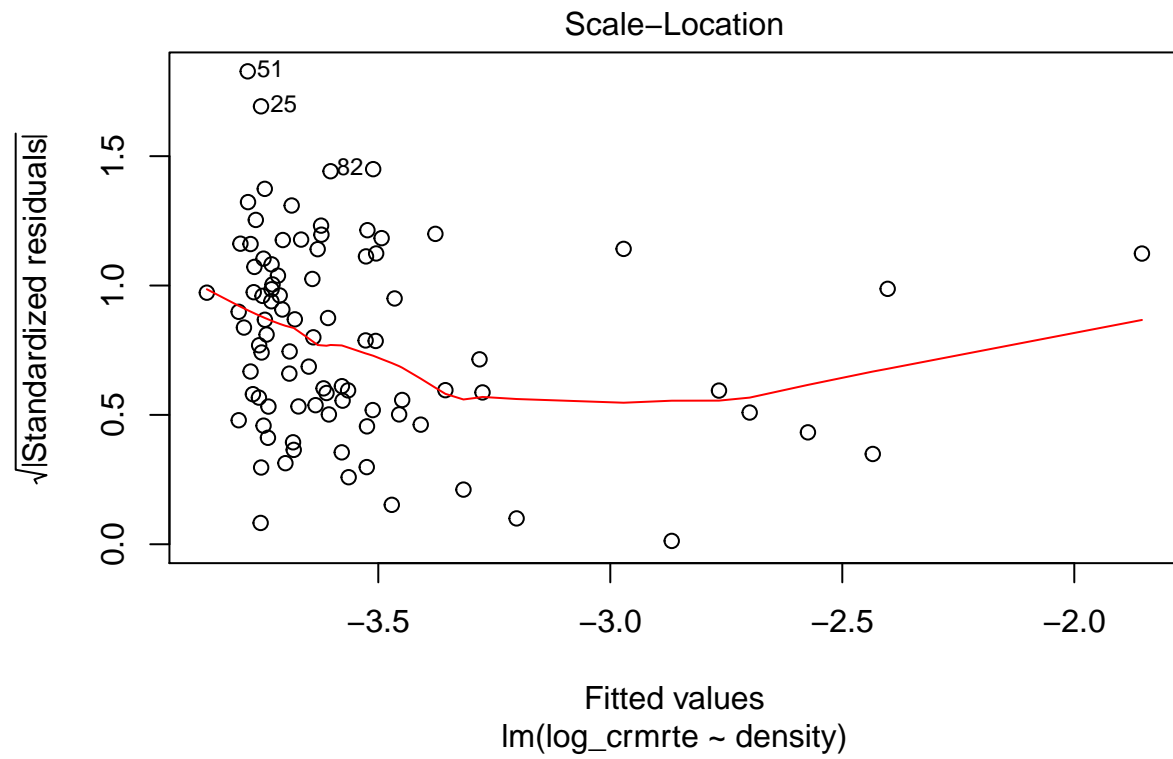
```
stargazer(simple_regression_model,
  type = "text", omit.stat = "f",
  se = list(se.simple_regression_model),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

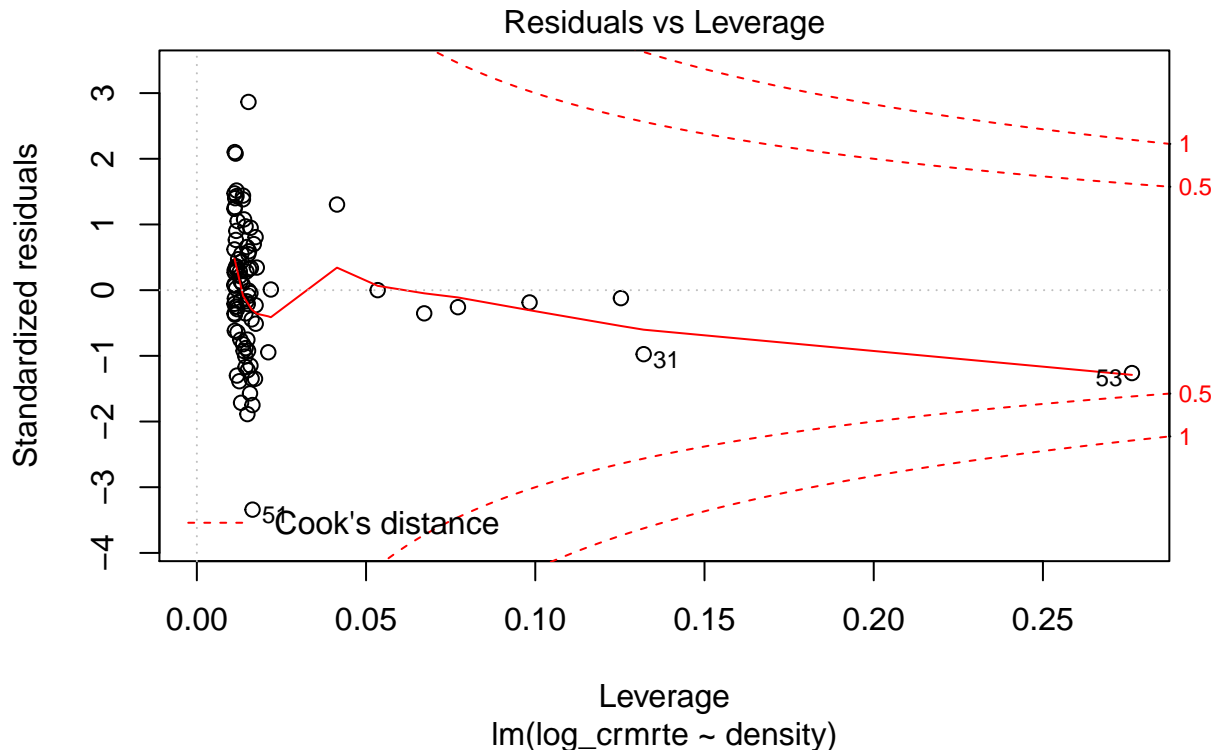
```
##
## =====
##                               Dependent variable:
##                               -----
##                               log_crmrte
## -----
## density                      0.228***
##                               (0.030)
##
## Constant                     -3.869***
##                               (0.069)
## -----
## Observations                  90
## R2                           0.401
## Adjusted R2                   0.394
## Residual Std. Error          0.427 (df = 88)
## =====
## Note:                        *p<0.05; **p<0.01; ***p<0.001
```

```
plot(simple_regression_model)
```









The variable density explains 40.1% of the variation in the log of crime rate. As density increases by 1 unit (as the county population divided by the county land area increases by 1%) crime increases by 22%.

### Model 2: Kitchen Sink Model

Still, we can do better in predicting the log crime rate than simply using one variable. We now examine a “kitchen sink” model. This model includes all of the variables in the data set except county (which has too many values to be a useful indicator variable) and year, which is a constant (1987). Below are the results.

```
all_in_model_log_level <- lm(log_crmrte ~ prbarr + prbconv + prbpris
+ avgsen + polpc + density
+ taxpc + west + central + urban
+ pctmin80 + wcon
+ wtuc + wtrd + wfir + wser + wmfg
+ wfed + wsta + wloc
+ mix + pctymle,
data = data_crmrte)
se.all_in_model_log_level = sqrt(diag(vcovHC(all_in_model_log_level)))
coeftest(all_in_model_log_level, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##          Estimate Std. Error t value Pr(>|t|)
## (Intercept) -4.0261e+00 8.4822e-01 -4.7466 1.128e-05 ***
## prbarr      -1.8891e+00 3.7955e-01 -4.9773 4.770e-06 ***
## prbconv     -6.5603e-01 1.7443e-01 -3.7611 0.0003579 ***
```

```
## prbpris      -9.3077e-02  3.9921e-01 -0.2332 0.8163542
## avgsen       -7.8769e-03  1.6125e-02 -0.4885 0.6267962
## polpc        1.5484e+02  8.6523e+01  1.7895 0.0780510 .
## density      1.1653e-01  5.4037e-02  2.1566 0.0346326 *
## taxpc        3.3224e-03  7.2890e-03  0.4558 0.6500012
## west         -1.1492e-01  1.2509e-01 -0.9187 0.3615403
## central      -1.0078e-01  9.2053e-02 -1.0948 0.2775232
## urban        -1.6923e-01  2.2872e-01 -0.7399 0.4619535
## pctmin80     9.9770e-03  3.0480e-03  3.2733 0.0016833 **
## wcon         4.6001e-04  8.3564e-04  0.5505 0.5838140
## wtuc         1.0174e-04  6.0187e-04  0.1690 0.8662750
## wtrd         2.5964e-04  1.7638e-03  0.1472 0.8834136
## wfir         -1.1015e-03  1.1960e-03 -0.9210 0.3603557
## wser         -1.3142e-04  1.5060e-03 -0.0873 0.9307193
## wmfg         -2.0528e-04  5.1630e-04 -0.3976 0.6921878
## wfed         2.3405e-03  1.0820e-03  2.1632 0.0340968 *
## wsta         -1.1357e-03  8.9769e-04 -1.2651 0.2102213
## wloc         5.8983e-04  2.4003e-03  0.2457 0.8066400
## mix          -2.3924e-01  6.2632e-01 -0.3820 0.7036869
## pctymle      2.7706e+00  1.4330e+00  1.9334 0.0574191 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(all_in_model_log_level)
```

```
## [1] 21.354
```

```
BIC(all_in_model_log_level)
```

```
## [1] 81.34943
```

```
stargazer(simple_regression_model, all_in_model_log_level,
  type = "text", omit.stat = "f",
  se = list(se.simple_regression_model, se.all_in_model_log_level),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               log_crmrte
##                               (1)         (2)
## -----
## prbarr                               -1.889***
##                               (0.380)
##
## prbconv                               -0.656***
##                               (0.174)
##
## prbpris                               -0.093
##                               (0.399)
##
```

## avgsen		-0.008
##		(0.016)
##		
## polpc		154.835
##		(86.523)
##		
## density	0.228***	0.117*
##	(0.030)	(0.054)
##		
## taxpc		0.003
##		(0.007)
##		
## west		-0.115
##		(0.125)
##		
## central		-0.101
##		(0.092)
##		
## urban		-0.169
##		(0.229)
##		
## pctmin80		0.010**
##		(0.003)
##		
## wcon		0.0005
##		(0.001)
##		
## wtuc		0.0001
##		(0.001)
##		
## wtrd		0.0003
##		(0.002)
##		
## wfir		-0.001
##		(0.001)
##		
## wser		-0.0001
##		(0.002)
##		
## wmfg		-0.0002
##		(0.001)
##		
## wfed		0.002*
##		(0.001)
##		
## wsta		-0.001
##		(0.001)
##		
## wloc		0.001
##		(0.002)
##		
## mix		-0.239
##		(0.626)
##		



```
## pctymle                2.771
##                (1.433)
##
## Constant             -3.869***   -4.026***
##                (0.069)   (0.848)
##
## -----
## Observations             90             90
## R2                      0.401           0.854
## Adjusted R2             0.394           0.806
## Residual Std. Error 0.427 (df = 88) 0.242 (df = 67)
## =====
## Note:                  *p<0.05; **p<0.01; ***p<0.001
```

Unsurprisingly, the r-squared of the “kitchen sink” model is substantially higher (85.4% vs. 40.1%). More importantly, the adjusted r-squared which accounts for the number of variables in the models, is also higher (80.6% vs 39.4%). Interestingly, density is no longer the variable with the highest statistical significance. The coefficients show the effect after all the other variables have been controlled for (partialled out). In the “kitchen sink” model prbarr and prbconv both have the lowest p-values.

### Model 3: Balanced Model

We took two approaches to building the balanced model. We used a bottom up approach that relied on both the correlation matrix and stepwise regression. We also used a top down approach that started with the “kitchen sink” model and excluded variables. Both methods are discussed below. Both approaches relied on our categories of variables to simplify the process.

```
base_forward = lm(log_crmrte ~ density,
                  data = data_crmrte)
forward_step = step(base_forward, scope = formula(all_in_model_log_level), direction = "forward")
```

```
## Start: AIC=-151.11
## log_crmrte ~ density
##
##           Df Sum of Sq    RSS    AIC
## + west      1   2.94477 13.116 -167.34
## + prbconv    1   2.59934 13.461 -165.00
## + prbarr     1   2.33206 13.729 -163.23
## + pctmin80   1   2.11480 13.946 -161.82
## + pctymle    1   1.14375 14.917 -155.76
## + wfed       1   0.92973 15.131 -154.48
## + taxpc      1   0.72379 15.337 -153.26
## + wser       1   0.53323 15.527 -152.15
## + wcon       1   0.38803 15.672 -151.31
## <none>              16.061 -151.11
## + avgscen    1   0.24653 15.814 -150.50
## + polpc      1   0.22427 15.836 -150.38
## + wfir       1   0.10688 15.954 -149.71
## + urban      1   0.06260 15.998 -149.46
## + central    1   0.05318 16.007 -149.41
## + mix        1   0.03960 16.021 -149.33
## + wmfgr      1   0.02854 16.032 -149.27
## + wsta       1   0.02520 16.035 -149.25
```

```

## + prbpris    1    0.02040 16.040 -149.22
## + wtrd       1    0.01224 16.048 -149.18
## + wtuc       1    0.00322 16.057 -149.13
## + wloc       1    0.00022 16.060 -149.11
##
## Step:  AIC=-167.34
## log_crmrte ~ density + west
##
##           Df Sum of Sq    RSS    AIC
## + prbconv   1   2.50648 10.609 -184.43
## + prbarr    1   1.68746 11.428 -177.73
## + pctymle   1   1.05265 12.063 -172.87
## + central   1   0.86332 12.252 -171.47
## + wser      1   0.68045 12.435 -170.13
## + wfed      1   0.57636 12.539 -169.38
## + taxpc     1   0.36020 12.756 -167.85
## <none>             13.116 -167.34
## + pctmin80   1   0.19632 12.919 -166.70
## + wcon       1   0.14814 12.968 -166.36
## + avgsgen    1   0.08265 13.033 -165.91
## + wmfg       1   0.07694 13.039 -165.87
## + wfir       1   0.06497 13.051 -165.79
## + mix        1   0.06018 13.056 -165.75
## + prbpris    1   0.04089 13.075 -165.62
## + wtuc       1   0.03304 13.083 -165.57
## + urban      1   0.03096 13.085 -165.55
## + polpc      1   0.02985 13.086 -165.54
## + wloc       1   0.01397 13.102 -165.44
## + wsta       1   0.00583 13.110 -165.38
## + wtrd       1   0.00529 13.111 -165.38
##
## Step:  AIC=-184.43
## log_crmrte ~ density + west + prbconv
##
##           Df Sum of Sq    RSS    AIC
## + prbarr     1   2.34643  8.2629 -204.92
## + wfed        1   0.83512  9.7742 -189.81
## + central     1   0.73021  9.8791 -188.84
## + mix         1   0.71874  9.8906 -188.74
## + pctymle     1   0.65915  9.9502 -188.20
## + pctmin80    1   0.32029 10.2890 -185.19
## + taxpc       1   0.26004 10.3493 -184.66
## + wmfg        1   0.24268 10.3666 -184.51
## <none>             10.6093 -184.43
## + wcon        1   0.13057 10.4787 -183.54
## + wtuc        1   0.08905 10.5203 -183.19
## + urban       1   0.04370 10.5656 -182.80
## + polpc       1   0.03042 10.5789 -182.69
## + wloc        1   0.02897 10.5803 -182.67
## + prbpris     1   0.02395 10.5854 -182.63
## + wser        1   0.00516 10.6042 -182.47
## + wtrd        1   0.00479 10.6045 -182.47
## + wsta        1   0.00284 10.6065 -182.45
## + wfir        1   0.00233 10.6070 -182.45

```

```

## + avgsen      1    0.00006 10.6093 -182.43
##
## Step:  AIC=-204.92
## log_crmrte ~ density + west + prbconv + prbarr
##
##           Df Sum of Sq    RSS    AIC
## + polpc      1    1.41147 6.8514 -219.78
## + wfed       1    0.81108 7.4518 -212.22
## + central    1    0.76023 7.5026 -211.61
## + pctmin80   1    0.70778 7.5551 -210.98
## + pctymle    1    0.30509 7.9578 -206.31
## + wmfgr      1    0.22113 8.0418 -205.37
## + taxpc      1    0.21216 8.0507 -205.26
## + wloc       1    0.19827 8.0646 -205.11
## <none>                8.2629 -204.92
## + avgsen     1    0.12246 8.1404 -204.27
## + wtuc       1    0.11422 8.1487 -204.18
## + mix        1    0.08199 8.1809 -203.82
## + wsta       1    0.04978 8.2131 -203.47
## + wcon       1    0.03148 8.2314 -203.27
## + wser       1    0.02647 8.2364 -203.21
## + wtrd       1    0.01465 8.2482 -203.08
## + urban      1    0.01199 8.2509 -203.05
## + wfir       1    0.00439 8.2585 -202.97
## + prbpris    1    0.00072 8.2622 -202.93
##
## Step:  AIC=-219.78
## log_crmrte ~ density + west + prbconv + prbarr + polpc
##
##           Df Sum of Sq    RSS    AIC
## + pctmin80   1    1.23128 5.6201 -235.61
## + central     1    0.63049 6.2209 -226.47
## + wfed        1    0.59242 6.2590 -225.92
## <none>                6.8514 -219.78
## + wsta       1    0.13668 6.7147 -219.59
## + pctymle    1    0.10414 6.7473 -219.16
## + wtuc       1    0.05144 6.8000 -218.46
## + wcon       1    0.04219 6.8092 -218.34
## + wmfgr      1    0.03521 6.8162 -218.25
## + mix        1    0.03490 6.8165 -218.24
## + avgsen     1    0.01955 6.8319 -218.04
## + wtrd       1    0.01788 6.8335 -218.02
## + urban      1    0.01590 6.8355 -217.99
## + wloc       1    0.00690 6.8445 -217.87
## + wfir       1    0.00240 6.8490 -217.81
## + prbpris    1    0.00027 6.8511 -217.79
## + taxpc      1    0.00020 6.8512 -217.78
## + wser       1    0.00014 6.8513 -217.78
##
## Step:  AIC=-235.61
## log_crmrte ~ density + west + prbconv + prbarr + polpc + pctmin80
##
##           Df Sum of Sq    RSS    AIC
## + wfed       1    0.53345 5.0867 -242.59

```

```

## + wsta      1    0.31439 5.3057 -238.79
## + wcon      1    0.22014 5.4000 -237.21
## + mix       1    0.19495 5.4252 -236.79
## + urban     1    0.14253 5.4776 -235.92
## + central   1    0.13315 5.4870 -235.77
## <none>      5.6201 -235.61
## + wtuc      1    0.12224 5.4979 -235.59
## + wtrd      1    0.09692 5.5232 -235.18
## + pctymle   1    0.07319 5.5469 -234.79
## + wloc      1    0.06955 5.5506 -234.73
## + wser      1    0.05762 5.5625 -234.54
## + wmfgr     1    0.04954 5.5706 -234.41
## + prbpris   1    0.01786 5.6023 -233.90
## + taxpc     1    0.00591 5.6142 -233.71
## + avgscn    1    0.00101 5.6191 -233.63
## + wfir      1    0.00073 5.6194 -233.62
##
## Step:  AIC=-242.59
## log_crmrte ~ density + west + prbconv + prbarr + polpc + pctmin80 +
##      wfed
##
##           Df Sum of Sq    RSS    AIC
## + wsta      1    0.36526 4.7214 -247.29
## + central   1    0.26095 4.8257 -245.33
## + pctymle   1    0.17155 4.9151 -243.67
## + wfir      1    0.15143 4.9353 -243.31
## <none>      5.0867 -242.59
## + urban     1    0.08107 5.0056 -242.03
## + taxpc     1    0.06932 5.0174 -241.82
## + wcon      1    0.05252 5.0342 -241.52
## + mix       1    0.05158 5.0351 -241.50
## + wser      1    0.03779 5.0489 -241.26
## + prbpris   1    0.02681 5.0599 -241.06
## + wtuc      1    0.02287 5.0638 -240.99
## + avgscn    1    0.01320 5.0735 -240.82
## + wmfgr     1    0.00108 5.0856 -240.61
## + wtrd      1    0.00084 5.0858 -240.60
## + wloc      1    0.00053 5.0862 -240.60
##
## Step:  AIC=-247.29
## log_crmrte ~ density + west + prbconv + prbarr + polpc + pctmin80 +
##      wfed + wsta
##
##           Df Sum of Sq    RSS    AIC
## + pctymle   1    0.269426 4.4520 -250.58
## + central   1    0.222728 4.4987 -249.64
## <none>      4.7214 -247.29
## + wfir      1    0.088681 4.6327 -247.00
## + mix       1    0.068999 4.6524 -246.62
## + prbpris   1    0.040292 4.6811 -246.06
## + urban     1    0.026383 4.6950 -245.80
## + wser      1    0.025477 4.6960 -245.78
## + taxpc     1    0.024825 4.6966 -245.77
## + wtrd      1    0.019294 4.7021 -245.66

```

```
## + wcon      1  0.018232 4.7032 -245.64
## + wmfg      1  0.008071 4.7134 -245.45
## + wloc      1  0.006051 4.7154 -245.41
## + avgssen   1  0.000694 4.7207 -245.31
## + wtuc      1  0.000097 4.7213 -245.29
##
## Step:  AIC=-250.58
## log_crmrte ~ density + west + prbconv + prbarr + polpc + pctmin80 +
##      wfed + wsta + pctymle
##
##           Df Sum of Sq    RSS    AIC
## + central  1  0.160121 4.2919 -251.88
## + taxpc    1  0.106562 4.3454 -250.76
## <none>      4.4520 -250.58
## + wfir     1  0.090730 4.3613 -250.43
## + mix      1  0.037862 4.4141 -249.35
## + prbpris  1  0.027609 4.4244 -249.14
## + wser     1  0.026573 4.4254 -249.12
## + wcon     1  0.020851 4.4312 -249.00
## + urban    1  0.016567 4.4354 -248.92
## + wmfg     1  0.009857 4.4421 -248.78
## + wloc     1  0.007999 4.4440 -248.74
## + wtrd     1  0.006269 4.4457 -248.71
## + avgssen  1  0.004132 4.4479 -248.66
## + wtuc     1  0.000726 4.4513 -248.60
##
## Step:  AIC=-251.88
## log_crmrte ~ density + west + prbconv + prbarr + polpc + pctmin80 +
##      wfed + wsta + pctymle + central
##
##           Df Sum of Sq    RSS    AIC
## <none>      4.2919 -251.88
## + wfir     1  0.074114 4.2178 -251.44
## + taxpc    1  0.065113 4.2268 -251.25
## + urban    1  0.043628 4.2483 -250.80
## + wcon     1  0.039905 4.2520 -250.72
## + avgssen  1  0.030713 4.2612 -250.52
## + mix      1  0.024668 4.2672 -250.40
## + wloc     1  0.019087 4.2728 -250.28
## + wmfg     1  0.010493 4.2814 -250.10
## + prbpris  1  0.007598 4.2843 -250.04
## + wtuc     1  0.003155 4.2887 -249.94
## + wtrd     1  0.002043 4.2898 -249.92
## + wser     1  0.001662 4.2902 -249.91
```

With the top down approach, we started with model 3 and looked to exclude variables that weren't as predictive. We ran hypothesis testing on all five groups, one group at a time.

```
#deterrent
linearHypothesis(all_in_model_log_level,
  c("prbarr = 0", "prbconv = 0", "prbpris = 0",
    "avgssen = 0", "polpc = 0"),
  vcov = vcovHC)
```

```
## Linear hypothesis test
##
## Hypothesis:
## prbarr = 0
## prbconv = 0
## prbpris = 0
## avgsgen = 0
## polpc = 0
##
## Model 1: restricted model
## Model 2: log_crmrte ~ prbarr + prbconv + prbpris + avgsgen + polpc + density +
##      taxpc + west + central + urban + pctmin80 + wcon + wtuc +
##      wtrd + wfir + wser + wmfg + wfed + wsta + wloc + mix + pctymle
##
## Note: Coefficient covariance matrix supplied.
##
##   Res.Df Df      F    Pr(>F)
## 1      72
## 2      67  5 6.0582 0.0001101 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
#wage
linearHypothesis(all_in_model_log_level,
                  c("wcon = 0", "wtuc = 0", "wtrd = 0",
                    "wfir = 0", "wser = 0", "wmfg = 0",
                    "wfed = 0", "wsta = 0", "wloc = 0"),
                  vcov = vcovHC)
```

```
## Linear hypothesis test
##
## Hypothesis:
## wcon = 0
## wtuc = 0
## wtrd = 0
## wfir = 0
## wser = 0
## wmfg = 0
## wfed = 0
## wsta = 0
## wloc = 0
##
## Model 1: restricted model
## Model 2: log_crmrte ~ prbarr + prbconv + prbpris + avgsgen + polpc + density +
##      taxpc + west + central + urban + pctmin80 + wcon + wtuc +
##      wtrd + wfir + wser + wmfg + wfed + wsta + wloc + mix + pctymle
##
## Note: Coefficient covariance matrix supplied.
##
##   Res.Df Df      F Pr(>F)
## 1      76
## 2      67  9 1.372 0.2185
```

```
#region
linearHypothesis(all_in_model_log_level,
                  c("west = 0", "central = 0"),
                  vcov = vcovHC)

## Linear hypothesis test
##
## Hypothesis:
## west = 0
## central = 0
##
## Model 1: restricted model
## Model 2: log_crmrte ~ prbarr + prbconv + prbpris + avgsgen + polpc + density +
##          taxpc + west + central + urban + pctmin80 + wcon + wtuc +
##          wtrd + wfir + wser + wmfg + wfed + wsta + wloc + mix + pctymle
##
## Note: Coefficient covariance matrix supplied.
##
##      Res.Df Df      F Pr(>F)
## 1         69
## 2         67  2 0.623 0.5394
```

```
#urban
linearHypothesis(all_in_model_log_level,
                  c("urban = 0"),
                  vcov = vcovHC)

## Linear hypothesis test
##
## Hypothesis:
## urban = 0
##
## Model 1: restricted model
## Model 2: log_crmrte ~ prbarr + prbconv + prbpris + avgsgen + polpc + density +
##          taxpc + west + central + urban + pctmin80 + wcon + wtuc +
##          wtrd + wfir + wser + wmfg + wfed + wsta + wloc + mix + pctymle
##
## Note: Coefficient covariance matrix supplied.
##
##      Res.Df Df      F Pr(>F)
## 1         68
## 2         67  1 0.5474 0.462
```

```
#demographic
linearHypothesis(all_in_model_log_level,
                  c("density = 0", "taxpc = 0", "pctmin80 = 0",
                    "mix = 0", "pctymle = 0"),
                  vcov = vcovHC)
```

```
## Linear hypothesis test
##
## Hypothesis:
```

```
## density = 0
## taxpc = 0
## pctmin80 = 0
## mix = 0
## pctymle = 0
##
## Model 1: restricted model
## Model 2: log_crmrte ~ prbarr + prbconv + prbpris + avgsgen + polpc + density +
##      taxpc + west + central + urban + pctmin80 + wcon + wtuc +
##      wtrd + wfir + wser + wmfgr + wfed + wsta + wloc + mix + pctymle
##
## Note: Coefficient covariance matrix supplied.
##
##      Res.Df Df      F    Pr(>F)
## 1      72
## 2      67  5 3.9627 0.003298 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

/ The hypothesis tests below show that of the five groups the only groups that are jointly significant are the deterrent data and the demographic data. These tests measure whether removing all the variables within a group reduces the r-squared by a statistically significant amount. We will re-run the models and compare.

```
balanced_model_top_1 <- lm(log_crmrte ~ prbarr + prbconv + prbpris
+ avgsgen + polpc + density
+ taxpc + pctmin80 + mix + pctymle,
data = data_crmrte)
se.balanced_model_top_1 = sqrt(diag(vcovHC(balanced_model_top_1)))
coeftest(balanced_model_top_1, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##      Estimate Std. Error t value Pr(>|t|)
## (Intercept) -3.3522918  0.3558217 -9.4213 1.467e-14 ***
## prbarr      -1.9627484  0.4014808 -4.8888 5.228e-06 ***
## prbconv     -0.7672158  0.1366862 -5.6130 2.846e-07 ***
## prbpris     -0.0764993  0.4732818 -0.1616 0.872005
## avgsgen     -0.0044749  0.0140406 -0.3187 0.750789
## polpc       176.1347220 82.5884550  2.1327 0.036056 *
## density      0.1135225  0.0351279  3.2317 0.001796 **
## taxpc        0.0020988  0.0055753  0.3764 0.707593
## pctmin80     0.0125062  0.0016215  7.7128 3.155e-11 ***
## mix         -0.7304967  0.5396416 -1.3537 0.179702
## pctymle      1.3832565  1.6211791  0.8532 0.396105
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(balanced_model_top_1)
```

```
## [1] 27.4514
```



```
BIC(balanced_model_top_1)
```

```
## [1] 57.44912
```

```
summary(balanced_model_top_1)
```

```
##
## Call:
## lm(formula = log_crmrte ~ prbarr + prbconv + prbpris + avgse +
##     polpc + density + taxpc + pctmin80 + mix + pctymle, data = data_crmrte)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.83872 -0.08696  0.03509  0.14723  0.50738
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  -3.352292   0.286701  -11.693  < 2e-16 ***
## prbarr        -1.962748   0.298386   -6.578 4.74e-09 ***
## prbconv       -0.767216   0.097473   -7.871 1.55e-11 ***
## prbpris       -0.076499   0.360363   -0.212 0.832432
## avgse        -0.004475   0.011615   -0.385 0.701072
## polpc         176.134722  43.331945    4.065 0.000113 ***
## density        0.113523   0.022249    5.102 2.26e-06 ***
## taxpc         0.002099   0.002537    0.827 0.410654
## pctmin80      0.012506   0.001763    7.095 4.91e-10 ***
## mix          -0.730497   0.413894   -1.765 0.081439 .
## pctymle       1.383256   1.311381    1.055 0.294728
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.2633 on 79 degrees of freedom
## Multiple R-squared:  0.7957, Adjusted R-squared:  0.7698
## F-statistic: 30.77 on 10 and 79 DF,  p-value: < 2.2e-16
```

```
stargazer(simple_regression_model, all_in_model_log_level, balanced_model_top_1,
  type = "text", omit.stat = "f",
  se = list(se.simple_regression_model, se.all_in_model_log_level, se.balanced_model_top_1),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               log_crmrte
##                               (1)         (2)         (3)
## -----
## prbarr                               -1.889***      -1.963***
##                               (0.380)      (0.401)
##
## prbconv                               -0.656***      -0.767***
##                               (0.174)      (0.137)
```

##			
## prbpris		-0.093	-0.076
##		(0.399)	(0.473)
##			
## avgsen		-0.008	-0.004
##		(0.016)	(0.014)
##			
## polpc		154.835	176.135*
##		(86.523)	(82.588)
##			
## density	0.228***	0.117*	0.114**
##	(0.030)	(0.054)	(0.035)
##			
## taxpc		0.003	0.002
##		(0.007)	(0.006)
##			
## west		-0.115	
##		(0.125)	
##			
## central		-0.101	
##		(0.092)	
##			
## urban		-0.169	
##		(0.229)	
##			
## pctmin80		0.010**	0.013***
##		(0.003)	(0.002)
##			
## wcon		0.0005	
##		(0.001)	
##			
## wtuc		0.0001	
##		(0.001)	
##			
## wtrd		0.0003	
##		(0.002)	
##			
## wfir		-0.001	
##		(0.001)	
##			
## wser		-0.0001	
##		(0.002)	
##			
## wmfg		-0.0002	
##		(0.001)	
##			
## wfed		0.002*	
##		(0.001)	
##			
## wsta		-0.001	
##		(0.001)	
##			
## wloc		0.001	
##		(0.002)	

```
##
## mix                -0.239          -0.730
##                  (0.626)          (0.540)
##
## pctymle            2.771           1.383
##                  (1.433)          (1.621)
##
## Constant          -3.869***        -4.026***        -3.352***
##                  (0.069)          (0.848)          (0.356)
##
## -----
## Observations      90              90              90
## R2                0.401           0.854           0.796
## Adjusted R2       0.394           0.806           0.770
## Residual Std. Error 0.427 (df = 88) 0.242 (df = 67) 0.263 (df = 79)
## =====
## Note:                *p<0.05; **p<0.01; ***p<0.001
```

Our adjusted r-squared has only fallen from 80.6% to 77.6% but we have dropped 12 variables. This is a much more parsimonious model. In order to double check wages, we decided to try to one more model that included just the median wage from all industries. The fundamental concept behind this is that the median could capture all opportunity for potential criminals, and it has the benefit of not being affected by the outlier in wser.

RESULT: Unfortunately, though it was much better, it was still not predictive.

```
balanced_model_top_2 <- lm(log_crmrte ~ prbarr + prbconv + prbpris
+ avgsen + polpc + density
+ taxpc + pctmin80 + mix + pctymle
+ median_wage,
data = data_crmrte)
se.balanced_model_top_2 = sqrt(diag(vcovHC(balanced_model_top_2)))
coeftest(balanced_model_top_2, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##          Estimate Std. Error t value Pr(>|t|)
## (Intercept) -3.7018614  0.6241512 -5.9310 7.799e-08 ***
## prbarr      -1.9379154  0.4176289 -4.6403 1.381e-05 ***
## prbconv     -0.7650603  0.1360985 -5.6214 2.826e-07 ***
## prbpris     -0.1114352  0.4509889 -0.2471 0.805487
## avgsen      -0.0046181  0.0137443 -0.3360 0.737770
## polpc       167.9127417  85.2311288  1.9701 0.052378 .
## density      0.0977596  0.0343625  2.8450 0.005671 **
## taxpc        0.0020705  0.0056750  0.3648 0.716214
## pctmin80     0.0123893  0.0016202  7.6467 4.534e-11 ***
## mix         -0.5896970  0.5806401 -1.0156 0.312961
## pctymle      1.5670818  1.9109317  0.8201 0.414680
## median_wage  0.0011536  0.0014133  0.8162 0.416848
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(balanced_model_top_2)
```

```
## [1] 28.07462
```

```
BIC(balanced_model_top_2)
```

```
## [1] 60.57215
```

```
stargazer(simple_regression_model, all_in_model_log_level, balanced_model_top_1, balanced_model_top_2,
  type = "text", omit.stat = "f",
  se = list(se.simple_regression_model, se.all_in_model_log_level,
    se.balanced_model_top_1, se.balanced_model_top_2),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               log_crmrte
##                               (1)         (2)         (3)         (4)
## -----
## prbarr                        -1.889***      -1.963***      -1.938***
##                               (0.380)        (0.401)        (0.418)
##
## prbconv                      -0.656***      -0.767***      -0.765***
##                               (0.174)        (0.137)        (0.136)
##
## prbpris                      -0.093         -0.076         -0.111
##                               (0.399)        (0.473)        (0.451)
##
## avgsen                      -0.008         -0.004         -0.005
##                               (0.016)        (0.014)        (0.014)
##
## polpc                       154.835        176.135*       167.913*
##                               (86.523)        (82.588)        (85.231)
##
## density                      0.228***      0.117*         0.114**        0.098**
##                               (0.030)        (0.054)        (0.035)        (0.034)
##
## taxpc                       0.003         0.002         0.002
##                               (0.007)        (0.006)        (0.006)
##
## west                        -0.115
##                               (0.125)
##
## central                     -0.101
##                               (0.092)
##
## urban                       -0.169
##                               (0.229)
##
## pctmin80                    0.010**        0.013***       0.012***
```

```

##                                (0.003)          (0.002)          (0.002)
##
## wcon                          0.0005
##                                (0.001)
##
## wtuc                          0.0001
##                                (0.001)
##
## wtrd                          0.0003
##                                (0.002)
##
## wfir                          -0.001
##                                (0.001)
##
## wser                          -0.0001
##                                (0.002)
##
## wmfg                          -0.0002
##                                (0.001)
##
## wfed                          0.002*
##                                (0.001)
##
## wsta                          -0.001
##                                (0.001)
##
## wloc                          0.001
##                                (0.002)
##
## mix                           -0.239          -0.730          -0.590
##                                (0.626)          (0.540)          (0.581)
##
## pctymle                       2.771          1.383          1.567
##                                (1.433)          (1.621)          (1.911)
##
## median_wage                   0.001
##                                (0.001)
##
## Constant                      -3.869***      -4.026***      -3.352***      -3.702***
##                                (0.069)        (0.848)        (0.356)        (0.624)
##
## -----
## Observations                   90             90             90             90
## R2                            0.401           0.854           0.796           0.799
## Adjusted R2                   0.394           0.806           0.770           0.770
## Residual Std. Error 0.427 (df = 88) 0.242 (df = 67) 0.263 (df = 79) 0.263 (df = 78)
## =====
## Note:                          *p<0.05; **p<0.01; ***p<0.001

```

Three of the five groups have been eliminated, with only the deterrent and demographic groups remaining. We will use step wise regression to evaluate.

```
base_backward = lm(log_crmrte ~ prbarr + prbconv + prbpris
                  + avgscen + polpc + density
                  + taxpc + pctmin80 + mix + pctymle,
                  data = data_crmrte)
backward_step = step(base_backward, scope = formula(base_backward), direction = "backward")
```

```
## Start: AIC=-229.96
## log_crmrte ~ prbarr + prbconv + prbpris + avgscen + polpc + density +
## taxpc + pctmin80 + mix + pctymle
##
##           Df Sum of Sq  RSS   AIC
## - prbpris    1    0.0031 5.4786 -231.91
## - avgscen    1    0.0103 5.4858 -231.79
## - taxpc      1    0.0474 5.5229 -231.18
## - pctymle    1    0.0771 5.5526 -230.70
## <none>                        5.4755 -229.96
## - mix        1    0.2159 5.6914 -228.48
## - polpc      1    1.1452 6.6207 -214.87
## - density    1    1.8044 7.2799 -206.32
## - prbarr     1    2.9989 8.4744 -192.65
## - pctmin80   1    3.4894 8.9649 -187.59
## - prbconv    1    4.2940 9.7695 -179.85
##
## Step: AIC=-231.91
## log_crmrte ~ prbarr + prbconv + avgscen + polpc + density + taxpc +
## pctmin80 + mix + pctymle
##
##           Df Sum of Sq  RSS   AIC
## - avgscen    1    0.0091 5.4877 -233.76
## - taxpc      1    0.0529 5.5316 -233.04
## - pctymle    1    0.0815 5.5601 -232.58
## <none>                        5.4786 -231.91
## - mix        1    0.2221 5.7007 -230.33
## - polpc      1    1.1489 6.6275 -216.77
## - density    1    1.8111 7.2897 -208.20
## - prbarr     1    2.9964 8.4750 -194.64
## - pctmin80   1    3.5028 8.9814 -189.42
## - prbconv    1    4.2956 9.7742 -181.81
##
## Step: AIC=-233.76
## log_crmrte ~ prbarr + prbconv + polpc + density + taxpc + pctmin80 +
## mix + pctymle
##
##           Df Sum of Sq  RSS   AIC
## - taxpc      1    0.0545 5.5422 -234.87
## - pctymle    1    0.0789 5.5667 -234.47
## <none>                        5.4877 -233.76
## - mix        1    0.2139 5.7016 -232.32
## - polpc      1    1.2273 6.7150 -217.59
## - density    1    1.8088 7.2965 -210.12
## - prbarr     1    3.0183 8.5060 -196.31
## - pctmin80   1    3.5470 9.0348 -190.89
## - prbconv    1    4.3218 9.8095 -183.48
```

```
##
## Step: AIC=-234.87
## log_crmrte ~ prbarr + prbconv + polpc + density + pctmin80 +
##      mix + pctymle
##
##           Df Sum of Sq      RSS      AIC
## - pctymle   1    0.0525  5.5947 -236.02
## <none>                        5.5422 -234.87
## - mix       1    0.2148  5.7571 -233.44
## - polpc     1    1.7384  7.2806 -212.31
## - density   1    1.8905  7.4328 -210.45
## - prbarr    1    3.5982  9.1404 -191.84
## - pctmin80  1    3.6798  9.2220 -191.04
## - prbconv   1    4.8508 10.3931 -180.28
##
## Step: AIC=-236.02
## log_crmrte ~ prbarr + prbconv + polpc + density + pctmin80 +
##      mix
##
##           Df Sum of Sq      RSS      AIC
## <none>                        5.5947 -236.02
## - mix       1    0.2319  5.8267 -234.36
## - density   1    1.8553  7.4500 -212.24
## - polpc     1    1.9402  7.5349 -211.22
## - pctmin80  1    3.7524  9.3471 -191.83
## - prbarr    1    3.9930  9.5877 -189.54
## - prbconv   1    5.3975 10.9922 -177.24

balanced_model_top_3 <- lm(log_crmrte ~ mix + density
                        + polpc + pctmin80
                        + prbarr + prbconv,
                        data = data_crmrte)
se.balanced_model_top_3 = sqrt(diag(vcovHC(balanced_model_top_3)))
coeftest(balanced_model_top_2, vcov = vcovHC)

##
## t test of coefficients:
##
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept) -3.7018614  0.6241512 -5.9310 7.799e-08 ***
## prbarr      -1.9379154  0.4176289 -4.6403 1.381e-05 ***
## prbconv     -0.7650603  0.1360985 -5.6214 2.826e-07 ***
## prbpris     -0.1114352  0.4509889 -0.2471  0.805487
## avgsen      -0.0046181  0.0137443 -0.3360  0.737770
## polpc       167.9127417 85.2311288  1.9701  0.052378 .
## density      0.0977596  0.0343625  2.8450  0.005671 **
## taxpc        0.0020705  0.0056750  0.3648  0.716214
## pctmin80     0.0123893  0.0016202  7.6467 4.534e-11 ***
## mix         -0.5896970  0.5806401 -1.0156  0.312961
## pctymle      1.5670818  1.9109317  0.8201  0.414680
## median_wage  0.0011536  0.0014133  0.8162  0.416848
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
AIC(balanced_model_top_3, k=2)
```

```
## [1] 21.39003
```

```
stargazer(simple_regression_model, all_in_model_log_level, balanced_model_top_1, balanced_model_top_3,
  type = "text", omit.stat = "f",
  se = list(se.simple_regression_model, se.all_in_model_log_level,
    se.balanced_model_top_1, se.balanced_model_top_3),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               log_crmrte
##                               (1)         (2)         (3)         (4)
## -----
## prbarr                        -1.889***      -1.963***      -2.100***
##                               (0.380)        (0.401)        (0.436)
##
## prbconv                      -0.656***      -0.767***      -0.809***
##                               (0.174)        (0.137)        (0.126)
##
## prbpris                      -0.093         -0.076
##                               (0.399)        (0.473)
##
## avgsen                      -0.008         -0.004
##                               (0.016)        (0.014)
##
## polpc                        154.835        176.135*        190.549**
##                               (86.523)        (82.588)        (71.937)
##
## density                      0.228***      0.117*        0.114**        0.113***
##                               (0.030)        (0.054)        (0.035)        (0.027)
##
## taxpc                        0.003         0.002
##                               (0.007)        (0.006)
##
## west                        -0.115
##                               (0.125)
##
## central                      -0.101
##                               (0.092)
##
## urban                        -0.169
##                               (0.229)
##
## pctmin80                    0.010**        0.013***      0.013***
##                               (0.003)        (0.002)        (0.001)
##
## wcon                        0.0005
##                               (0.001)
```



```

##
## wtuc                0.0001
##                    (0.001)
##
## wtrd                0.0003
##                    (0.002)
##
## wfir               -0.001
##                    (0.001)
##
## wser              -0.0001
##                    (0.002)
##
## wmfg              -0.0002
##                    (0.001)
##
## wfed               0.002*
##                    (0.001)
##
## wsta              -0.001
##                    (0.001)
##
## wloc               0.001
##                    (0.002)
##
## mix                -0.239      -0.730      -0.745
##                    (0.626)      (0.540)      (0.474)
##
## pctymle            2.771      1.383
##                    (1.433)      (1.621)
##
## Constant           -3.869***   -4.026***   -3.352***   -3.197***
##                    (0.069)      (0.848)      (0.356)      (0.237)
##
## -----
## Observations        90          90          90          90
## R2                  0.401        0.854        0.796        0.791
## Adjusted R2         0.394        0.806        0.770        0.776
## Residual Std. Error 0.427 (df = 88) 0.242 (df = 67) 0.263 (df = 79) 0.260 (df = 83)
## =====
## Note:                                     *p<0.05; **p<0.01; ***p<0.001

```

The difference between the backward and forward model is that the backward model chooses variables for exclusion based on comparing significance while the forward model looks for significance in inclusion. We also used the f-tests (hypothesis tests) to give the backward stepwise regression a head start.

The backward stepwise regression yielded a more reasonable model so that is the model we are choosing for our balanced model. This model strikes a nice balance between parsimony and explanatory power. The variables included are prbarr, prbconv, polpc, density, pctmin80, and mix. Six out of the original 24 independent variables are included. The adjust r-squared is only 3% lower (77.6% vs. 80.6%). It includes a blend of actionable items for the campaign in the deterrent data as well as demographic variables that perhaps can focus the campaign's efforts.

## Praveen's section

### 3. An Assessment of the CLM Assumptions

We choose our balanced model for the complete assessment of all 6 classical linear model assumptions.

Assumptions	Ways Assumption Fails	Diagnostic	Conseq of Failed Assump	Solution
<b>Assumption MLR.1 Linear in Parameters</b> The model in the population can be written as $y = \beta_0 + \beta_1x_1 + \beta_2x_2 + \dots + \beta_kx_k + u$ [3.31] where $\beta_0, \beta_1, \dots, \beta_k$ are the unknown parameters (constants) of interest and $u$ is an unobserved random error or disturbance term.	Fit linear model to nonlinear data	plot of observed vs. predicted values or plot of residuals vs. predicted values	Bad predictions, particularly out of range of the sample data	Apply a nonlinear transformation to either independent or dependent variables; add another regressor that is a nonlinear function of another variable; add a possibly omitted variable
<b>Assumption MLR.2 Random Sampling</b> We have a random sample of $n$ observations, $(x_{1i}, x_{2i}, \dots, x_{ki}, y_i); i = 1, 2, \dots, n$ , following the population model in Assumption MLR.1.	1.) Clustering (researchers can only access a limited number) 2.) Autocorrelation common for time series data	1.) Use knowledge of where data comes from 2.) Durbin Watson Statistic	1.) Observing less variation than actually exists in the population; betas still unbiased but estimates are much less precise	1.) Use clustered standard errors 2.) No simple fix for serial correlation - use time series model
<b>Assumption MLR.3 No Perfect Collinearity</b> In the sample (and therefore in the population), none of the independent variables is constant, and there are no exact linear relationships among the independent variables.	Extremely high or perfect multicollinearity (assumption only rules out perfect multicollinearity); often from lagged variables of another variable, a shared common time trend, or variables that capture similar phenomena	Correlation matrix (though difficult for several variables); VIF's (multicollinearity likely between 5 and 10, problem > 10)	When variables are highly correlated but not perfectly collinear, OLS works but estimates will be much less precise. R-squared may be high but t-stats are low; regression becomes sensitive to small changes in specification and adding or removing a variable changes betas a lot; you might get nonsensical coefficient signs and magnitudes; confidence intervals might be very wide	Drop redundant variables
<b>Assumption MLR.4 Zero Conditional Mean</b> The error $u$ has an expected value of zero given any values of the independent variables. In other words, $E(u x_1, x_2, \dots, x_k) = 0$ [3.36]	The error exhibits a pattern that is not in a fairly constant band around zero or it shows a pattern that results in nonzero errors for different $x$ 's	For one variable, plot residuals vs predictor (should see flat average line around zero). For multiple regression, plot residuals vs. fitted values (predicted values). Should again see a flat band or line. Also use domain knowledge on any omitted variables.	Endogeneity is a violation of zero-conditional mean and results in OLS coefficients that are biased and inconsistent. If the explanatory variables are uncorrelated with the error term they are exogenous	1.) Change the functional form (log of independent or dependent variable, $x$ and $x^2$ , etc.); might lose interpretability though. 2.) Adding new variables. 3.) Decide we can't meet zero conditional mean but we can meet exogeneity. If we satisfy the first three assumptions and exogeneity $Cov(x, u) = 0$ for all $x$ , dependent variables then OLS estimators are consistent (unbiased as $n \rightarrow \infty$ )

Figure 1: MLR 1-4'

**MLR.1: The model is linear in parameters ( and the error term)**

```
balanced_model_top_3 <- lm(log_crmrte ~ mix + density
+ polpc + pctmin80
+ prbarr + prbconv,
data = data_crmrte)
```

**MLR.2: Random sampling**

First thing to note is that we are dealing with a single cross-section (1987) of a multi-year panel data.

Secondly this is observational data and not experiental so perfect random sampling is hard to achieve.

CORNWELL – TRUMBULL (1994) specifically state they choose panel data because cross – section data were not able to capture the real effect of the crime rate on several independent regressors.

The authors identify that the time-series component of the panel data is able to identify specific characteristics of county heterogeneity, which is correlated with the criminal justice variables.

```
se.balanced_model_top_3 = sqrt(diag(vcovHC(balanced_model_top_3)))
coeftest(balanced_model_top_3, vcov = vcovHC)
```

```
##
## t test of coefficients:
##
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -3.1966143  0.2371960 -13.4767 < 2.2e-16 ***
## mix         -0.7447751  0.4742158  -1.5705  0.120094
## density      0.1134850  0.0265199   4.2792 4.997e-05 ***
## polpc       190.5494683 71.9365456   2.6489  0.009666 **
## pctmin80     0.0127752  0.0014745   8.6643 3.067e-13 ***
## prbarr       -2.0998396  0.4356245  -4.8203 6.398e-06 ***
## prbconv      -0.8094922  0.1261063  -6.4191 8.051e-09 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### MLR.3: No perfect multicollinearity

Multicollinearity refers to a situation in which two or more explanatory variables in a multiple regression model are highly linearly related. We have perfect multicollinearity if, for example as in the equation above, the correlation between two independent variables is equal to 100% or negative 100%.

As seen from the correlation matrix below, there is no perfect multicollinearity in the model but we observe some meaningful correlations between (Prbarr, mix) and (Prbarr,density)

These linear relationships among the X's don't invlaidate the MLR but they lower precision and increase the std-errors in the mdoel

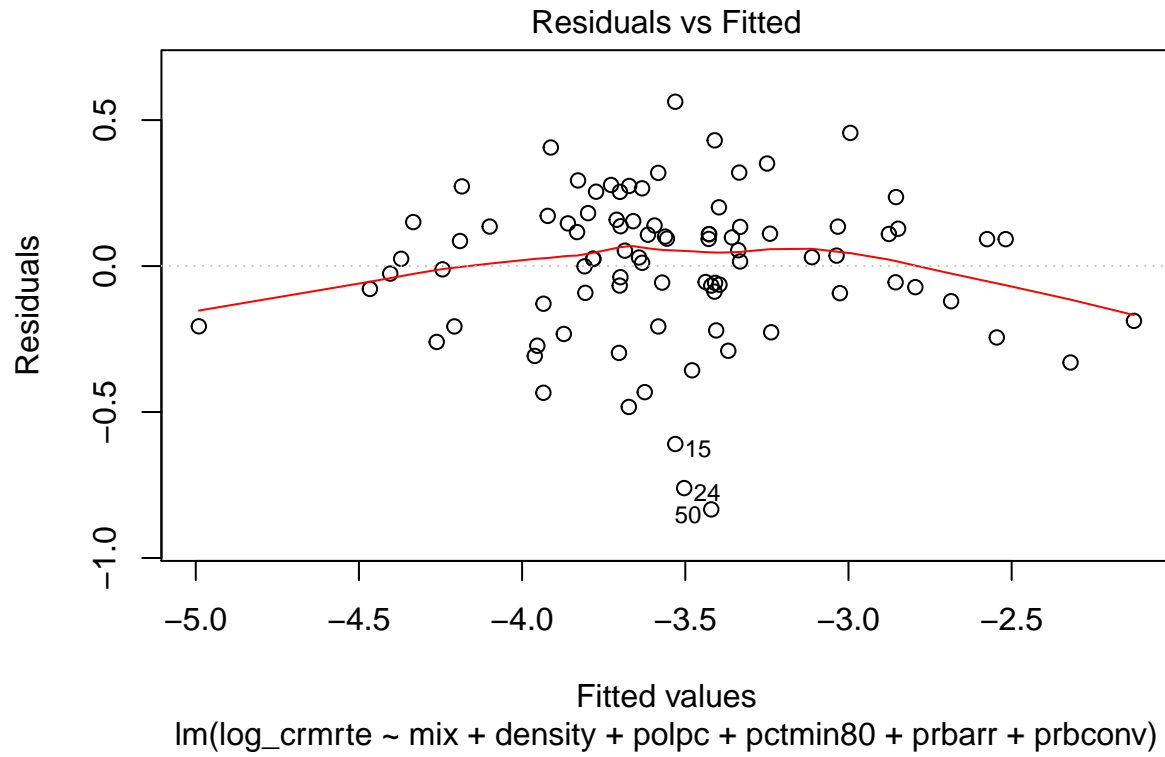
```
balanced_model <- c( "mix", "density","polpc", "pctmin80","prbarr","prbconv")
balanced_model_data <- data_crmrte[balanced_model]
round(cor(balanced_model_data)*100,2) # correlations displayed as % for convenience
```

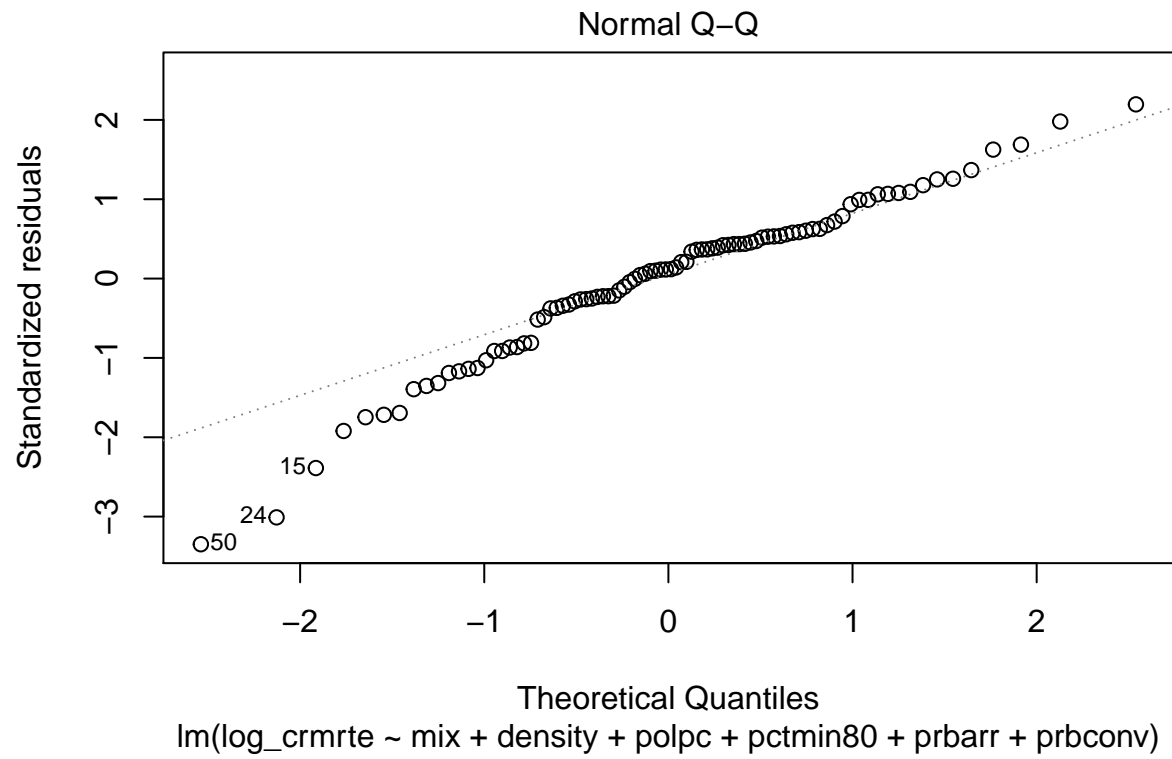
```
##           mix density  polpc pctmin80 prbarr prbconv
## mix       100.00 -13.69   2.41   20.12  41.29 -30.43
## density  -13.69  100.00  15.91   -7.46 -30.27 -22.67
## polpc     2.41   15.91 100.00  -16.91  42.60  17.19
## pctmin80 20.12   -7.46 -16.91  100.00   4.91   6.25
## prbarr    41.29 -30.27  42.60    4.91 100.00  -5.58
## prbconv  -30.43 -22.67  17.19    6.25 -5.58 100.00
```

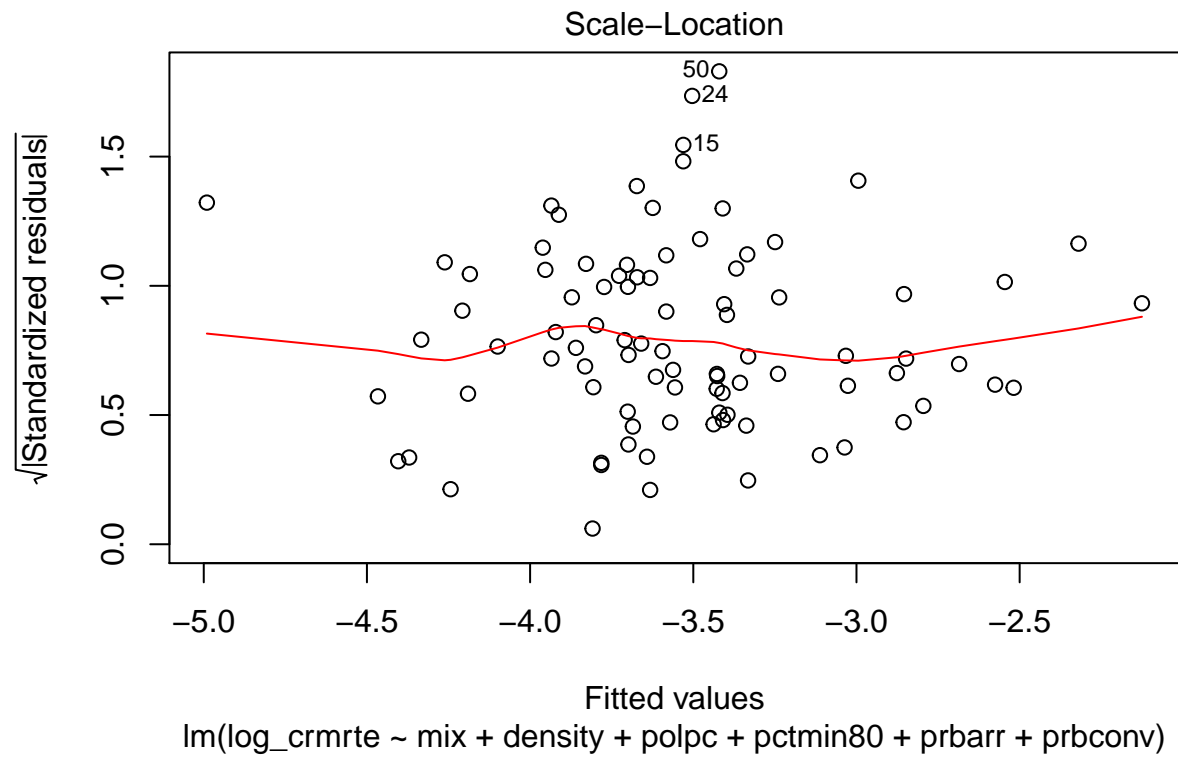
### MLR.4: Zero Conditional Mean / exogeneity

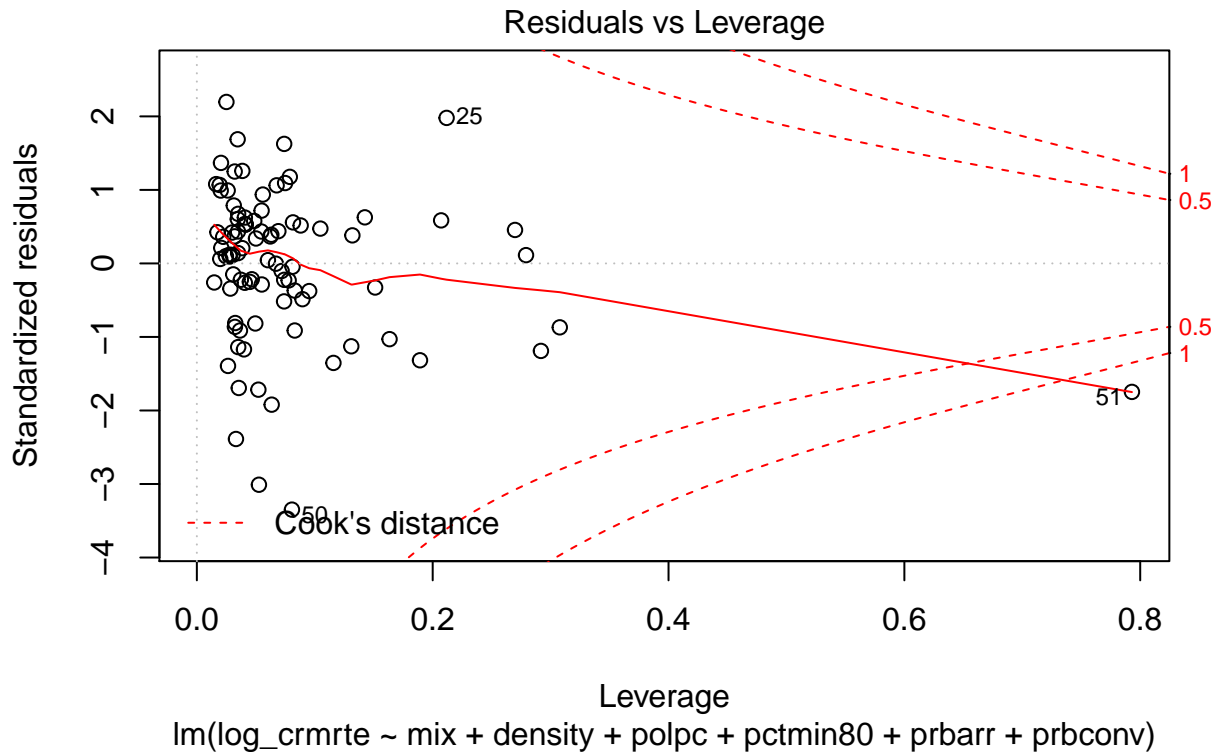
ZCM is best analysed by studying the regression plots of the residuals. Let's start by looking at the regression plots of the balanced model

```
plot(balanced_model_top_3)
```









#### CLM assumptions analysis from plots

- Plot 1. The residuals vs. fitted plot indicates that the zero conditional mean assumption is NOT perfectly satisfied but the red line is close enough to zero, a big improvement compared to some of the other models we tested. The non-uniform thickness of the residuals indicates possible heteroskedasticity.
- Plot 2. The Q-Q plot shows that the residuals are not perfectly normally distributed, but the log transform of the crime rate improved the distribution
- Plot 3. The scale location plot indicated the presence of heteroskedasticity especially in the middle where the thickness of the band varies and outliers such as '50' and '24' are generating large standardized residuals
- Plot 4. The residuals vs leverage plot shows some of the outliers we had discussed earlier ( 51, 25, 84) but the most significantly outlier is 51 ( having high leverage and Cook's distance >1). This outlier 51 significantly affects our model estimate and might be worth removing from the data to improve model accuracy.

```
round(cor(balanced_model_top_3$residuals, balanced_model_data)*100,5)
```

```
##      mix density polpc pctmin80 prbarr prbconv
## [1,]  0         0      0         0      0      0
```

Finally we check the correlation between the X's and the errors in the model to ensure there is no endogeneity in the model. There is a more extensive discussion of omitted variable and their implication on model endogeneity in section 5 of the report.

<b>Assumption MLR.5</b> <b>Homoskedasticity</b> The error $u$ has the same variance given any values of the explanatory variables. In other words, $\text{Var}(u_j   x_1, \dots, x_k) = \sigma^2$ .	Variance of error term is not constant across $x$ values (increasing, decreasing, increasing then decreasing, etc.)	1.) Residuals vs fitted value plot. 2.) Scale location plot 3.) Breusch-Pagan Test (sensitive to sample size)	Difficult to gauge true variance of errors; confidence intervals too wide or too narrow; confidence intervals will be too narrow for out-of-sample predictions if increasing variance	Calculate heteroskedasticity robust standard errors (White standard errors). If accompanied by zero conditional mean violation then there may be an exponential or log relationship in data.
<b>Assumption MLR.6</b> <b>Normality</b> The population error $u$ is independent of the explanatory variables $x_1, x_2, \dots, x_k$ and is normally distributed with zero mean and variance $\sigma^2$ ; $u \sim \text{Normal}(0, \sigma^2)$ .	Often if $y$ variable is skewed errors will be skewed as well.	Histogram of residuals, Q-Q plot, Shapiro Wilk test (though this test doesn't tell you how large deviations from normality are)	Difficult to gauge whether model coefficients are significantly different from zero and calculate confidence intervals; outliers can have outsized effect on parameter estimates (OLS minimizes squared error)	Technically OK if sample size is large enough; try in order: 1.) rely on asymptotic properties of OLS 2.) for small datasets transform $y$ variable 3.) If residuals vs fitted value plot shows curvature, violation of both normal errors and zero conditional mean so try quadratic or additional predictor 4.) Use bootstrapping
<b>THEOREM 3.1</b> <b>UNBIASEDNESS OF OLS</b> Under Assumptions MLR.1 through MLR.4,				
$E(\hat{\beta}_j) = \beta_j, j = 0, 1, \dots, k, \quad [3.37]$ for any values of the population parameter $\beta_j$ . In other words, the OLS estimators are unbiased estimators of the population parameters.				

Figure 2: MLR 5 & 6'

## MLR.5: Homoskedasticity

Homoscedasticity describes a situation in which the error term has the same variance across all values of the independent variables.

The regression plots indicate the presence of some heteroskedasticity in the errors let's test if they are statistically significant using the Breusch-Pagan test.

- The Breusch-Pagan test below allows us to test for heteroskedasticity under the

$H_0$  : There is homoskedasticity

```
bptest(balanced_model_top_3)
```

```
##
## studentized Breusch-Pagan test
##
## data:  balanced_model_top_3
## BP = 8.6648, df = 6, p-value = 0.1933
```

From the BP test, surprisingly we find p-value is not statistically significant, therefore we fail to reject  $H_0$  : There is homoskedasticity

However, we will still choose to be more conservative and use HC consistent std-errors (Huber-white Std-errors) using `coeftest` function from the `sandwich` package in R. This conservative approach we have taken throughout this report in our model selection process in choosing regressors for different models

## MLR.6: Normality of the error term

Often, if the  $Y$  variable is skewed, the error terms will be skewed as well.

We can check the normality using the Q-Q plot to visualize the distribution of residuals.

We saw in the earlier section that the crime rate has some positive skew, but we were able to reduce the skew by applying log transform to the crime rate.



We can also run a Shapiro - wilk test for normality

```
shapiro.test(balanced_model_top_3$residuals)
```

```
##
## Shapiro-Wilk normality test
##
## data:  balanced_model_top_3$residuals
## W = 0.96118, p-value = 0.008754
```

The p-value is significant, therefore we reject  $H_0$ : Errors are normal

The non-normality of the residuals is statistically significant for this model.

There is some negative skew from outlier 51 in the transformed variable but However since we have  $n > 30$  under CLT we have OLS estimators are normally distributed.

## 4. A Regression Table

The results were displayed in stargazer using HC standard errors as part of model selection

- This sections has been fully covered under section 2 of the report
- We hav include statistical F-tests besides the standard t-tests for regression coefficients to check model validity.
- Additionally the practical significance of the model variable chosen have also been discussed in detail
- Below is the summary of the regression models and the AIC & BIC scores which provides a parsimony adjusted measure of fit

```
stargazer(simple_regression_model, all_in_model_log_level, balanced_model_top_1, balanced_model_top_3,
  type = "text", omit.stat = "f",
  se = list(se.simple_regression_model, se.all_in_model_log_level,
    se.balanced_model_top_1, se.balanced_model_top_3),
  star.cutoffs = c(0.05, 0.01, 0.001))
```

```
##
## =====
##                               Dependent variable:
##                               -----
##                               log_crmrte
##                               (1)         (2)         (3)         (4)
## -----
## prbarr                        -1.889***    -1.963***    -2.100***
##                               (0.380)      (0.401)      (0.436)
##
## prbconv                      -0.656***    -0.767***    -0.809***
##                               (0.174)      (0.137)      (0.126)
##
## prbpris                      -0.093       -0.076
##                               (0.399)      (0.473)
##
## avgsen                      -0.008       -0.004
##                               (0.016)      (0.014)
```

##				
## polpc		154.835	176.135*	190.549**
##		(86.523)	(82.588)	(71.937)
##				
## density	0.228***	0.117*	0.114**	0.113***
##	(0.030)	(0.054)	(0.035)	(0.027)
##				
## taxpc		0.003	0.002	
##		(0.007)	(0.006)	
##				
## west		-0.115		
##		(0.125)		
##				
## central		-0.101		
##		(0.092)		
##				
## urban		-0.169		
##		(0.229)		
##				
## pctmin80		0.010**	0.013***	0.013***
##		(0.003)	(0.002)	(0.001)
##				
## wcon		0.0005		
##		(0.001)		
##				
## wtuc		0.0001		
##		(0.001)		
##				
## wtrd		0.0003		
##		(0.002)		
##				
## wfir		-0.001		
##		(0.001)		
##				
## wser		-0.0001		
##		(0.002)		
##				
## wmfg		-0.0002		
##		(0.001)		
##				
## wfed		0.002*		
##		(0.001)		
##				
## wsta		-0.001		
##		(0.001)		
##				
## wloc		0.001		
##		(0.002)		
##				
## mix		-0.239	-0.730	-0.745
##		(0.626)	(0.540)	(0.474)
##				
## pctymle		2.771	1.383	
##		(1.433)	(1.621)	

```
##
## Constant          -3.869***      -4.026***      -3.352***      -3.197***
##                   (0.069)        (0.848)        (0.356)        (0.237)
##
## -----
## Observations      90             90             90             90
## R2                 0.401         0.854         0.796         0.791
## Adjusted R2       0.394         0.806         0.770         0.776
## Residual Std. Error 0.427 (df = 88) 0.242 (df = 67) 0.263 (df = 79) 0.260 (df = 83)
## =====
## Note:                                     *p<0.05; **p<0.01; ***p<0.001
```

```
AIC(simple_regression_model)
```

```
## [1] 106.2991
```

```
AIC(all_in_model_log_level)
```

```
## [1] 21.354
```

```
AIC(balanced_model_top_1)
```

```
## [1] 27.4514
```

```
AIC(balanced_model_top_3)
```

```
## [1] 21.39003
```

```
BIC(simple_regression_model)
```

```
## [1] 113.7985
```

```
BIC(all_in_model_log_level)
```

```
## [1] 81.34943
```

```
BIC(balanced_model_top_1)
```

```
## [1] 57.44912
```

```
BIC(balanced_model_top_3)
```

```
## [1] 41.38851
```

## Parsimony adjusted model performance

Though AIC and BIC are both Maximum Likelihood estimate driven and penalize free parameters in an effort to combat overfitting, they do so in ways that result in significantly different behavior. Lets look at one commonly presented version of the methods (which results from stipulating normally distributed errors and other well behaving assumptions):

$AIC = -2\ln(\text{likelihood}) + 2k$ , and  $BIC = -2\ln(\text{likelihood}) + \ln(N)k$ ,

where:  $k$  = model degrees of freedom (  $K=2$  is default for OLS)  $N$  = number of observations

The quick explanation is: - AIC is best for prediction as it is asymptotically equivalent to cross-validation.  
- BIC is best for explanation as it allows consistent estimation of the underlying data generating process.

When  $N$  is large the two models will produce quite different results. Then the BIC applies a much larger penalty for complex models, and hence will lead to simpler models than AIC for very large  $N$ .

So we check both IC for our model and in both cases a lower value implies a better parsimony adjusted outcome.

```
AIC(simple_regression_model)
```

```
## [1] 106.2991
```

```
AIC(all_in_model_log_level)
```

```
## [1] 21.354
```

```
AIC(balanced_model_top_1)
```

```
## [1] 27.4514
```

```
AIC(balanced_model_top_3)
```

```
## [1] 21.39003
```

```
BIC(simple_regression_model)
```

```
## [1] 113.7985
```

```
BIC(all_in_model_log_level)
```

```
## [1] 81.34943
```

```
BIC(balanced_model_top_1)
```

```
## [1] 57.44912
```

```
BIC(balanced_model_top_3)
```

```
## [1] 41.38851
```