

1.31 Consider estimation in Linear Model  $Y = b_1 X_1 + b_2 X_2 + \epsilon$ ,  $0 \leq b_1, b_2 \leq 1$  for sample  $(Y_1, X_{11}, X_{21}), \dots, (Y_n, X_{1n}, X_{2n})$ . Errors  $\epsilon_i$  are independent and distributed to  $N(0, 1)$ . A noninformative prior is  $\pi(b_1, b_2) = \mathbb{I}_{[0,1]}(b_1) \mathbb{I}_{[0,1]}(b_2)$ .

a. Show that posterior means are given by  $i=1, 2$

$$E^\pi(b_i | y_1, \dots, y_n) = \frac{\int_0^1 \int_0^1 b_i \prod_{j=1}^n \phi(y_j - b_1 X_{1j} - b_2 X_{2j}) db_1 db_2}{\int_0^1 \int_0^1 \prod_{j=1}^n \phi(y_j - b_1 X_{1j} - b_2 X_{2j}) db_1 db_2}, \quad \phi \text{ density standard normal.}$$

Soln: Given  $\epsilon_i \sim N(0, 1)$ ,  $y_j - b_{1j} X_{1j} - b_{2j} X_{2j} \sim N(0, 1)$ .

$$E^\pi(b_i | y_1, \dots, y_n) = \int b_i \pi(b_1, b_2 | (X_{11}, X_{21}) \dots (X_{1n}, X_{2n})) db_1 db_2$$

Since  $\pi(b_1, b_2 | (X_{11}, X_{21}) \dots (X_{1n}, X_{2n})) \propto \pi(b_1, b_2) f((X_{11}, X_{21}) \dots (X_{1n}, X_{2n}) | b_1, b_2)$ , WLOG

$$\begin{aligned} \int b_i \pi(b_1, b_2 | (X_{11}, X_{21}) \dots (X_{1n}, X_{2n})) db_1 db_2 &= \int_0^1 \int_0^1 b_i k \pi(b_1, b_2) f((X_{11}, X_{21}) \dots (X_{1n}, X_{2n}) | b_1, b_2) db_1 db_2 \\ &= k \int_0^1 \int_0^1 b_i f((X_{11}, X_{21}) \dots (X_{1n}, X_{2n}) | b_1, b_2) \pi(b_1, b_2) db_1 db_2 \\ &= \frac{\int_0^1 \int_0^1 b_i f((X_{11}, X_{21}) \dots (X_{1n}, X_{2n}) | b_1, b_2) \pi(b_1, b_2) db_1 db_2}{\int_0^1 \int_0^1 f((X_{11}, X_{21}) \dots (X_{1n}, X_{2n}) | b_1, b_2) \pi(b_1, b_2) db_1 db_2}, \text{ where } k \text{ is marginal factor.} \end{aligned}$$

$$= \frac{\int_0^1 \int_0^1 b_i \prod_{j=1}^n \phi(y_j - b_{1j} X_{1j} - b_{2j} X_{2j}) \pi(b_1, b_2) db_1 db_2}{\int_0^1 \int_0^1 \prod_{j=1}^n \phi(y_j - b_{1j} X_{1j} - b_{2j} X_{2j}) \pi(b_1, b_2) db_1 db_2}$$

$$= \frac{\int_0^1 \int_0^1 b_i \prod_{j=1}^n \phi(y_j - b_{1j} X_{1j} - b_{2j} X_{2j}) db_1 db_2}{\int_0^1 \int_0^1 \prod_{j=1}^n \phi(y_j - b_{1j} X_{1j} - b_{2j} X_{2j}) db_1 db_2}$$

$$= \frac{\int_0^1 \int_0^1 \prod_{j=1}^n \phi(y_j - b_{1j} X_{1j} - b_{2j} X_{2j}) db_1 db_2}{\int_0^1 \int_0^1 \prod_{j=1}^n \phi(y_j - b_{1j} X_{1j} - b_{2j} X_{2j}) db_1 db_2} \quad \text{for noninformative priors (Proven).}$$

b. Show an equivalent expression is  $\delta_i^\pi(y_1, \dots, y_n) = \frac{E^\pi(b_i \mathbb{I}_{[0,1]^2}(b_1, b_2) | y_1, \dots, y_n)}{p^\pi((b_1, b_2) \in [0,1]^2 | y_1, \dots, y_n)}$ , where the right-hand term is computed under the distribution  $\begin{pmatrix} b_1 \\ b_2 \end{pmatrix} \sim N_2\left(\begin{pmatrix} \hat{b}_1 \\ \hat{b}_2 \end{pmatrix}, (X^T X)^{-1}\right)$ , with  $(\hat{b}_1, \hat{b}_2)$  the unconstrained least square estimator of  $(b_1, b_2)$  and  $X = \begin{pmatrix} X_{11} & X_{21} \\ \vdots & \vdots \\ X_{1n} & X_{2n} \end{pmatrix}$ .

Soln: Given  $Y = b_1 X_1 + b_2 X_2 + \epsilon$ ,  $\delta_i^\pi(y_1, \dots, y_n) = \arg \min_{\delta} E(L(\delta, b_1, b_2) | y_1, \dots, y_n)$

$$= \min_{\delta} \int L(\delta, b_1, b_2) \pi(b_1, b_2 | y_1, \dots, y_n) db_1 db_2$$

For  $L(\delta, b_1, b_2)$  with  $(\hat{b}_1, \hat{b}_2)$  as the unconstrained Least Square Estimator, we want to minimise  $\epsilon_i$ ,  $\sum_{j=1}^n (y_j - b_{1j} X_{1j} - b_{2j} X_{2j})^2 = \sum_{j=1}^n (\epsilon_j)^2$ . In vector form, minimise  $\|y^T - Xb\|^2$ , given  $y = (y_1, \dots, y_n)$ ,  $X = \begin{pmatrix} X_{11} & X_{21} \\ \vdots & \vdots \\ X_{1n} & X_{2n} \end{pmatrix}$ ,  $b = \begin{pmatrix} b_1 \\ b_2 \end{pmatrix}$ . That is  $L(\delta, b_1, b_2) = \|y^T - Xb\|^2$ .

As loss function is quadratic loss,  $\delta^\pi(y_1, \dots, y_n) = E^\pi(b_i | y_1, \dots, y_n)$ , where  $b_i$  is independent parameter consideration given that it is under bivariate normal distribution  $\begin{pmatrix} b_1 \\ b_2 \end{pmatrix} \sim N_2\left(\begin{pmatrix} \hat{b}_1 \\ \hat{b}_2 \end{pmatrix}, (X^T X)^{-1}\right)$ .

By Conditional Expectation,  $E^\pi(b_i | y_1, \dots, y_n) = \frac{\int b_i \star \pi(b_1, b_2 | (X_{11}, X_{21}) \dots (X_{1n}, X_{2n})) db_1 db_2}{\int \pi(b_1, b_2 | (X_{11}, X_{21}) \dots (X_{1n}, X_{2n})) db_1 db_2} = \frac{E^\pi(b_i \mathbb{I}_{[0,1]^2}(b_1, b_2) | y_1, \dots, y_n)}{p^\pi((b_1, b_2) \in [0,1]^2 | y_1, \dots, y_n)}$

Generalisation is  $E(b_i | \text{under joint}) = \frac{E(b_i \text{ under joint dist.})}{P(b_i \text{ under joint dist.})}$ ,  $\rho \sigma_{b_1}, \sigma_{b_2} \neq 0$ .

$$\therefore \delta^\pi(y_1, \dots, y_n) = \frac{E^\pi(b_i \mathbb{I}_{[0,1]^2}(b_1, b_2) | y_1, \dots, y_n)}{p^\pi((b_1, b_2) \in [0,1]^2 | y_1, \dots, y_n)} \quad (\text{proven}).$$