

Offense Pass Completion as a Predictor of NFL Score Differential*

Using offense pass completion to predict how successful an NFL team is at creating score advantages

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September 25, 2025

Measuring whether the quality of an NFL team's quarterback translates into more wins for a team. The quarterback's quality is measured through pass completion rate and the team's quality is measured through score differential. Overall a strong association is found between pass completion and score differential.

Introduction

In recent years, there has been an increase in the quarterback's pay as percent of salary cap. In a 5 year span, the number of quarterbacks taking "at least 10% of the cap"(Howe 2025) increased from 7 to 23. This leaves less room on a team's salary cap for the rest of the team. Crucial positions on the defense, and perhaps even on the offense may be left neglected. Is quarterback really that important? This paper tries to quantify whether a better quarterback can deliver a greater amount of wins using score-differential and pass completion rate.

Since an NFL team wins by scoring more than their opponent at the end of a game, this paper uses score-differential as a proxy for how well a team did during a game. Thus a better high score-differential should represent a more winning team. Offense pass completion percent is also used as a measure of a quarterback's quality.

*Project repository available at: <https://github.com/peteragao/MATH261A-project-template>.

Data

The `nfl-team-statistics.csv` dataset provided by SCORE Sports Data Repository (Yurko 2023) contains statistics about the regular season performance for each NFL team from 1999 to 2022. The data was collected using the `nflreadr` package (Ho and Carl 2025) in R.

The purpose of this paper is to observe the relationship between offense completion percentage and score-differential and as such these are the two relevant variables taken from the raw dataset provided by (Yurko 2023). A total of 765 observations were recorded and analyzed in this paper.

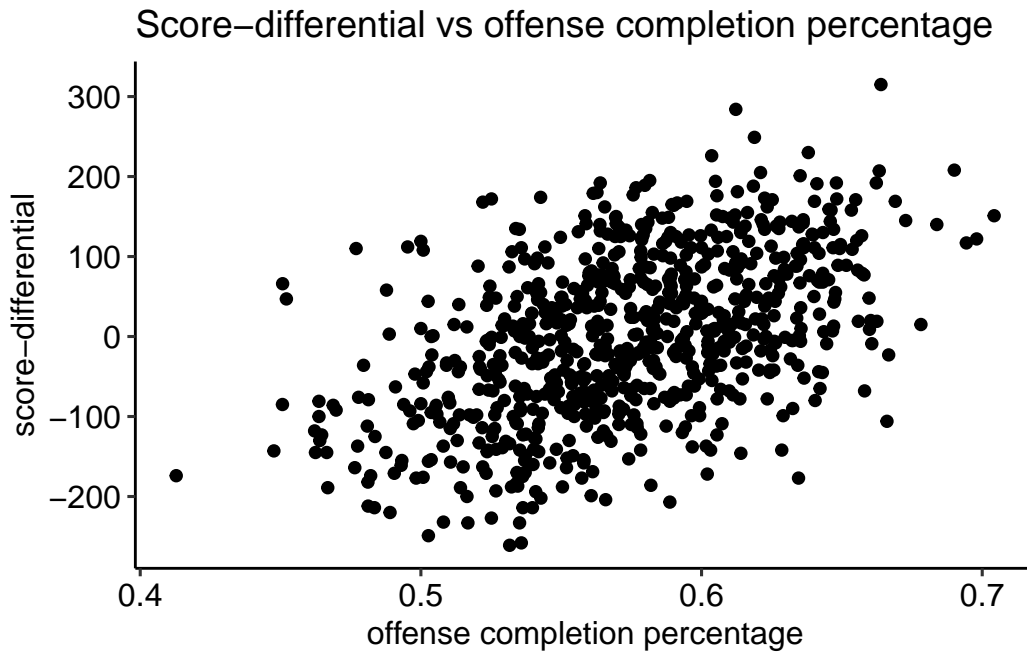


Figure 1: Scatter plot with offense completion percentage as the predictor values and score-differential as the outcome values

From Figure 1 we can see a correlation between the increase in offense completion percentage and score-differential. This paper will quantify this correlation.

offense_completion_percentage	score_differential
Min. :0.4128	Min. : -261
1st Qu.:0.5409	1st Qu.: -74
Median :0.5730	Median : 1
Mean :0.5732	Mean : 0
3rd Qu.:0.6061	3rd Qu.: 72
Max. :0.7043	Max. : 315

The table and Figure 1 above show a concentration of observations around 57% offense completion percentage with the 1st and 3rd quartile being within 3% of the mean. Additionally note the even distribution of score differential around the mean with the 1st and 3rd quartile 74 and 72 points from the mean respectively. Outliers are present in the data as seen in Figure 1.

Methods

This paper will fit the simple linear regression model

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

to understand the relationship between the predictor variable, offense completion percentage, and the outcome variable, score-differential. In this model, X_i represents offense completion percentage of the i th observation and Y_i represents the score-differential of the i th outcome. β_0 represents the intercept coefficient, what we expect score-differential to be when the offense has a completion percentage of 0. β_1 represents the slope coefficient, what we expect the increase in score-differential will be for every percent increase in completion percentage. In this model, we assume the error term, ε_i to be random with mean 0 and finite variance σ^2 .

The parameters, β_0 and β_1 , can be estimated using the method of least squares which minimizes the sum of the squared deviations $Q = \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2$. This requires the assumption that the errors are uncorrelated and have equal variances with mean 0. From the OLS parameter estimates, the regression function can be estimated with $\hat{Y}_i = b_0 + b_1 X_i$. This gives the fitted values for case i . The fitted model created fits a linear line that best predicts score differential, from offense completion percentage.

I implemented this analysis using the R programming language (R Core Team 2025).

Results

Using the `lm()` function in fits our model using the ordinary least squares method. The results are shown below.

B0	B1
-618.0257	1078.1405

β_0 represents the intercept coefficient and has a value of -618.03 . Thus, score-differential can be predicted to be -618.03 when the offense has a pass completion percentage of 0%. β_1 represents the slope coefficient and has a value of 1078.14. When pass completion percentage goes to 100%, an increase of 1078.14 would be expected.

Revisualize the scatter plot from the beginning of the paper to better illustrate the linear association.

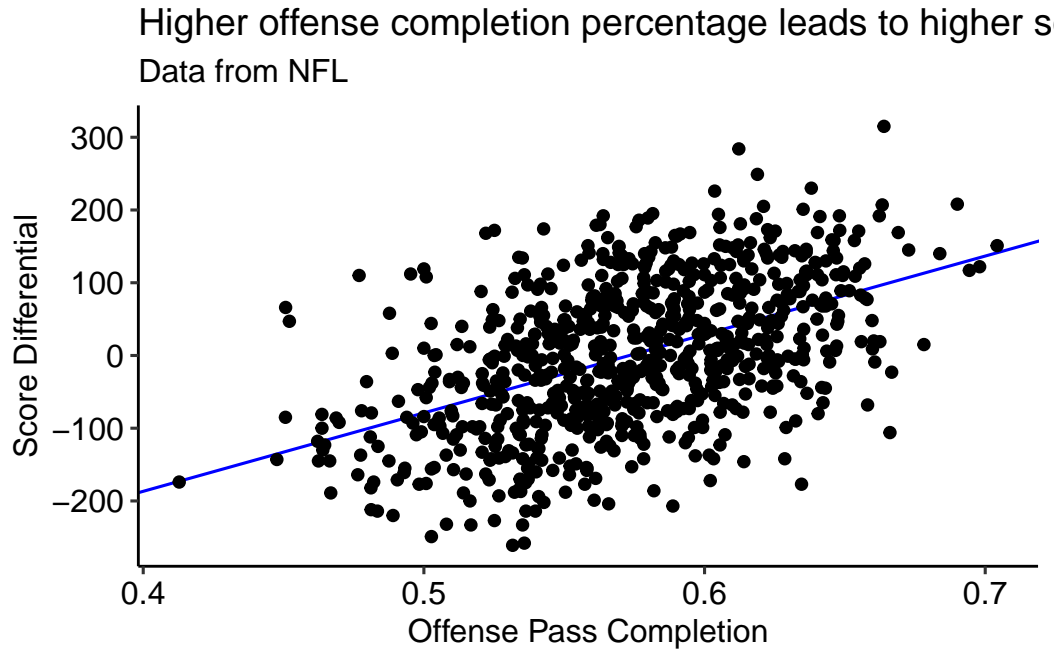


Figure 2

Call:

```
lm(formula = score_differential ~ offense_completion_percentage,
    data = nfl)
```

Residuals:

Min	1Q	Median	3Q	Max
-243.110	-64.738	-1.134	60.790	241.983

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-618.03	39.21	-15.76	<2e-16 ***
offense_completion_percentage	1078.14	68.18	15.81	<2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 87.62 on 763 degrees of freedom

Multiple R-squared: 0.2468, Adjusted R-squared: 0.2458

F-statistic: 250 on 1 and 763 DF, p-value: < 2.2e-16

In order to assess whether or not there is a linear association between X and Y , this paper defines the null and alternative hypotheses as $H_0 : \beta_1 = 0$ and $H_A : \beta_1 \neq 0$ respectively.

We can use a two sided t-test to test the null hypothesis, calculated by $t^* = b_1 / \widehat{se}(b_1)$, but the $lm()$ gives a P-value we can use. We use this to test at a 5% significant level and is shown below.

p-value
6.343907e-49

Discussion

The p-value achieved $6.343907e - 49 < 0.05$. This is much lower than our significance level, so we reject the null hypothesis. There is statistical evidence that there is a linear association between offense pass completion percentage and score differential.

Analysis below reaffirms the assumptions made in this paper.

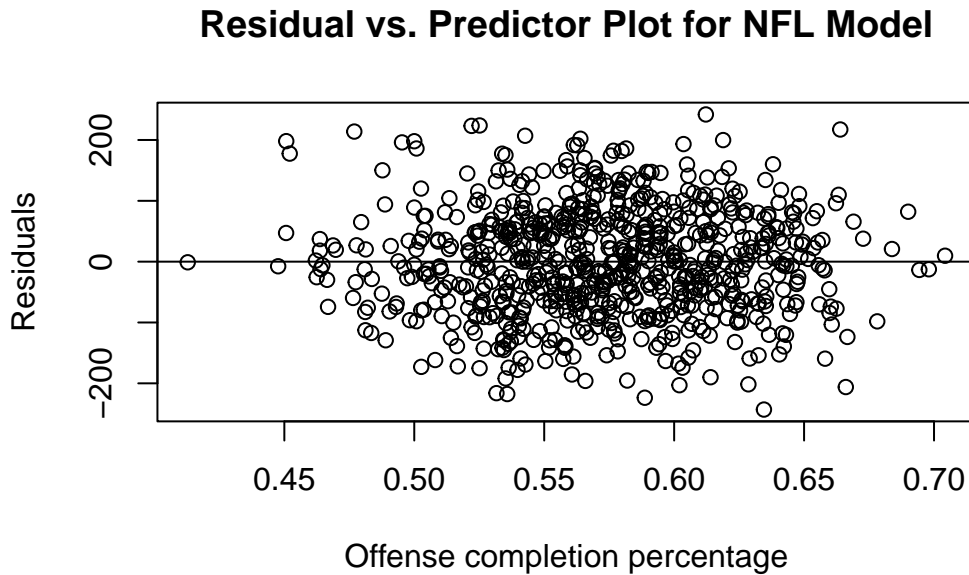


Figure 3: A residual vs predictor plot for NFL model using offense pass completion as a predictor.

The Figure 3 figure shows random scatter around 0, which indicates that the residuals have a constant variance. This is important for showing linearity and homoscedasticity.

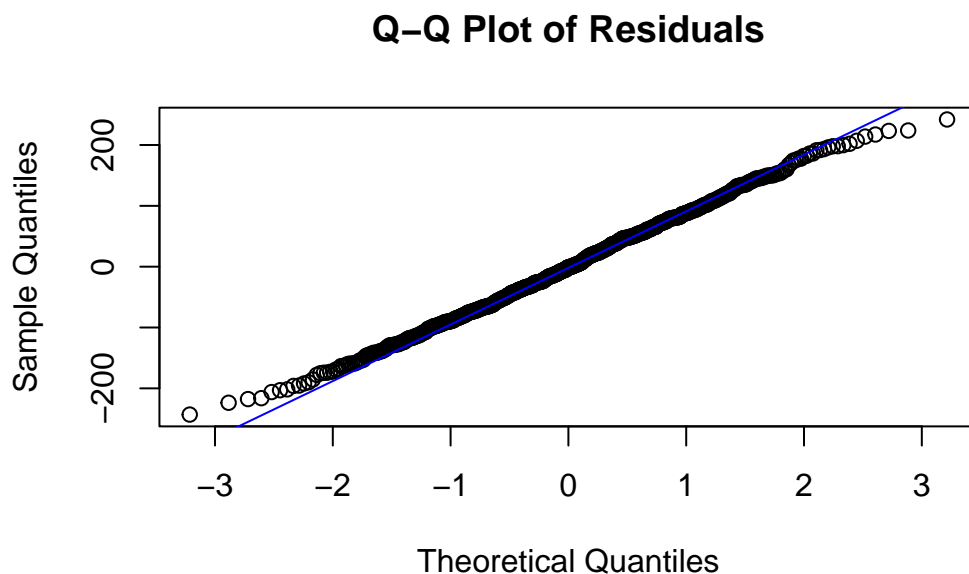


Figure 4: A residual vs predictor plot for NFL model using offense pass completion as a predictor.

The quantile-quantile plot shown in Figure 4 checks if the data is normally distributed. From Figure 4 slight deviation is seen from the tails at the ends of the regression line which would indicate outliers in the data, but it does not appear to be very influential in the dataset.

While the regression itself shows a strong association between the two variables used, other assumptions made in this paper pose questions about the conclusions made. This paper uses pass completion as a proxy for a judge for quarterbacks, but this neglects the quality of the receiver. Additionally score differential can be influenced by the quality of the defense, masking the true influence of the quarterback.

References

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- Howe, Jeff. 2025. "Why Explosion of NFL QB Contracts Has Drastically Altered How Teams Build Rosters." <https://www.nytimes.com/athletic/6576049/2025/08/28/nfl-qb-contracts-mahomes-allen-prescott-burrow/>.

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