

# How Prevalent is Transitivity-Failure in Bayesian Confirmation?

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It is a well-known fact among epistemologists and philosophers of science that transitivity can *fail* for Bayesian confirmation. That is, it is possible that  $A$  confirms  $B$  and  $B$  confirms  $C$  while  $A$  fails to confirm  $C$ . Still, there is a growing number of conditions in the literature under which this cannot happen, some of which are surprisingly weak. This raises the question how *prevalent* the phenomenon of transitivity-failure is: perhaps, Bayesian confirmation is transitive *in most cases*? The present paper aims at answering this and related questions with the help of a method known as Monte-Carlo integration. It also compares the prevalence of transitivity-failure with failures of adjacent inference patterns, some of which are known from non-monotonic reasoning and the logic of conditionals.

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## 1. Introduction

According to Bayesian confirmation theory, a proposition  $A$  *confirms* another proposition  $B$  relative to probability function  $P$  if and only if  $P(B|A) > P(B)$ . This conception of confirmation was popularized by Carnap ([1950]) but can be traced back to least Keynes ([1921]). And it is arguably *the* most influential conception of confirmation currently on the market (for overviews see Fitelson [2001] or Strevens [2017]). But it is a well-known fact among epistemologists and philosophers of science that confirmation, thus understood, is not transitive (see Suppes [1970]). That is, the following inference pattern is invalid for Bayesian confirmation:<sup>1</sup>

**Transitivity.** If  $A$  confirms  $B$  and  $B$  confirms  $C$ , then  $A$  confirms  $C$ .

For an instance of transitivity-failure, consider drawing a single card from a well-shuffled standard deck and let  $A$  state that the drawn card is a heart,  $B$  that it is red and  $C$  that it is a diamond (this example is due to Roche [2012a]). We then have:

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<sup>1</sup>Bayesian confirmation is often contrasted with *absolute* confirmation, i.e. sufficiently high conditional probability. I discuss transitivity-failure for absolute confirmation in section 4. For a recent alternative conception of confirmation see Smith's ([2016]) idea of *normic support*.

- $P(B|A) = 1 > P(B) = 1/2$
- $P(C|B) = 1/2 > P(C) = 1/4$
- $P(C|A) = 0 < P(C) = 1/4$

But the possibility of transitivity-failure does not mean that Bayesian confirmation is *never* transitive. In fact, there is a growing number of conditions in the literature under which transitivity is valid. For instance, Suppes ([1970]) pointed out that this is the case whenever  $P(B|A)$  and  $P(C|B)$  are maximal. And Hesse ([1970]) uncovered a condition that guarantees the transitivity of Bayesian confirmation *and* a restricted form of transitivity for absolute confirmation (see also Koscholke [2023]). Shogenji ([2003]) later showed that Bayesian confirmation is transitive whenever  $B$  screens-off  $A$  and  $C$  from each other (see also Reichenbach [1956] or Eells and Sober [1983]). And Roche ([2012b]) pointed out that screening-off can be weakened without compromising transitivity (see also Suppes [1986]). Finally, Atkinson and Peijnenburg ([2021ab]) recently presented a new condition that is even weaker than the aforementioned. In fact, their condition can be made *arbitrarily* weak and will still guarantee transitivity.

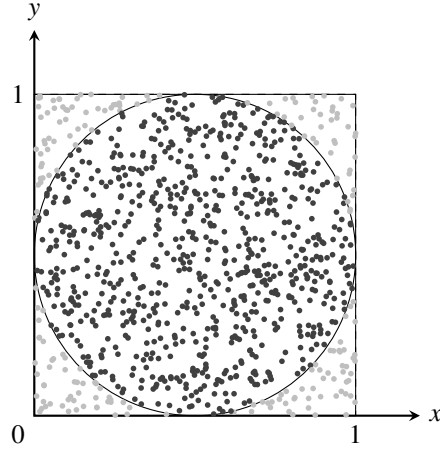
But if there are *so many* conditions under which Bayesian confirmation is transitive, some of which are surprisingly weak, one might wonder how *prevalent* instances of transitivity-failure actually are: could it be that Bayesian confirmation is transitive *in most cases*? In this paper, I wish to answer this question with the help of a method known as Monte-Carlo integration. I will start by briefly explaining this method in section 2 and continue by presenting the main results in section 3. In section 4, I present analogous results for absolute confirmation. And in section 5 to 7, I compare the prevalence of transitivity-failure in Bayesian confirmation with the failure of adjacent inference patterns. I conclude in section 8.<sup>2</sup>

## 2. Monte-Carlo Integration

Monte-Carlo integration is a computational method developed by Metropolis and Ulam ([1949]) to solve integration problems that are too complex to be solved analytically (see Hammersley and Handscomb [1964] or Krauth [2006] for overviews).

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<sup>2</sup>An anonymous referee remarked that the formulation ‘could it be that Bayesian confirmation is transitive *in most cases*’ is somewhat misleading since whether or not a relation is transitive depends on what happens *in all cases*. To avoid confusion, the formulation should be ‘could it be that most cases in which  $A$  confirms  $B$  and  $B$  confirms  $C$  are also cases in which  $A$  confirms  $C$ ’. For ease of presentation, however, I will stick to the loose formulation above.



**Figure 1:** Approximating  $\pi$  with 1000 randomly sampled points.

**Table 1:** Stochastic truth table.

$A$	$B$	$C$	$P$
0	0	0	$p_1$
0	0	1	$p_2$
0	1	0	$p_3$
0	1	1	$p_4$
1	0	0	$p_5$
1	0	1	$p_6$
1	1	0	$p_7$
1	1	1	$p_8$

A well-known illustration of this method is the numerical approximation of  $\pi$  via integration over the area of a circle in a unit square, as shown in Figure 1. One starts by picking a random sample of  $n$  points from the unit square and then divides the number of points inside the circle by  $n$ . Since the area of a unit circle is  $\pi/4$ , one obtains an approximation of  $\pi$  by multiplying the previous result by 4. And the larger the random sample is, the closer the approximation will be to the real value. For example, for 1000 points one already obtains a circle area of around 79% of the unit square thus a  $\pi$  value of around 3.16.

Now, determining  $\pi$  numerically rather than analytically may seem like a mere math exercise. But it is very instructive, since our approach to determining the prevalence of transitivity-failure is almost identical. We did, however, not sample two-dimensional points from the Euclidean plane but eight-dimensional vectors from a hypervolume. The hypervolume represents the space of all probability functions

over an algebra generated by three variables. Thus, each vector  $(p_1, \dots, p_8)$  with  $p_i \geq 0$  and  $\sum p_i = 1$  is actually a probability function where each component is the probability assigned to the corresponding Boolean combination of our three variables  $A$ ,  $B$  and  $C$ . The exact correspondence is shown in Table 1, a so-called *stochastic truth-table* (for details see Fitelson [2008]). Since any such vector is either an instance of transitivity-failure or not, we can approximate the prevalence of transitivity-failure and transitivity-success by sampling as many vectors as possible. The results of this procedure are presented in the next section.<sup>3</sup>

### 3. Transitivity-Failure in Bayesian Confirmation

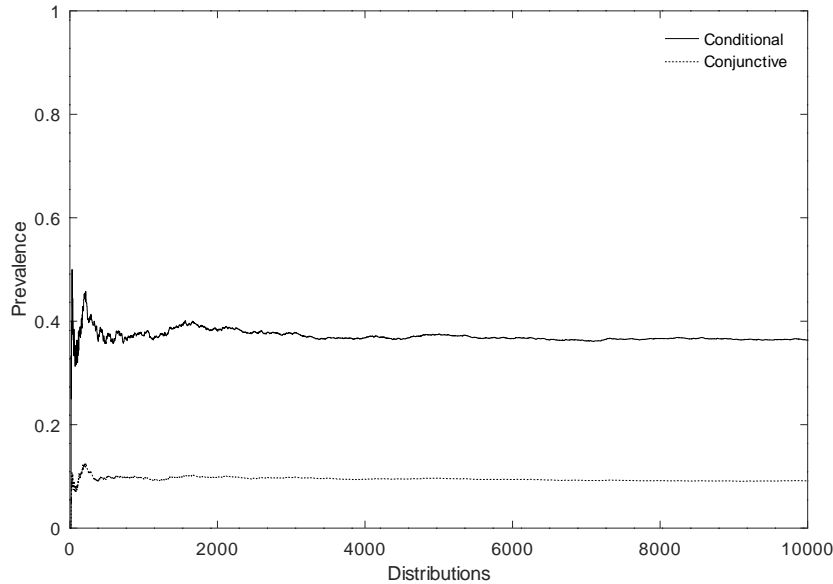
To ensure the numerical precision of our approximations, we sampled 10 million regular probability distributions from the set of all probability distributions on the algebra over three variables  $A$ ,  $B$  and  $C$ . These distributions were randomly generated using the Mersenne-Twister algorithm by Matsumoto and Nishimura ([1998]) which corresponds to sampling from a uniform Dirichlet distribution on the unit 8-simplex with concentration parameters  $\alpha_1, \dots, \alpha_8$  equal to 1 (see Kotz et al. [2000]). This sample was then used to approximate two types of prevalence values of transitivity-failure:

- **Conjunctive Prevalence:** the proportion of cases in which  $A$  confirms  $B$  and  $B$  confirms  $C$  and  $A$  fails to confirm  $C$ . Less formally, this quantity describes how prevalent instances of transitivity-failure are *unconditionally*.
- **Conditional Prevalence:** the prevalence of cases in which  $A$  fails to confirm  $C$  given that  $A$  confirms  $B$  and  $B$  confirms  $C$ . Less formally, this quantity describes how prevalent instances of transitivity-failure are *conditional on the* transitivity-antecedent already being satisfied.

As Figure 2 shows, the conjunctive prevalence of transitivity-failure quickly stabilized at around 9%. This means that instances of transitivity-failure in Bayesian confirmation are less prevalent than instances of transitivity-success. Notice, however, that the remaining 91% of transitivity-success include *trivial* cases where the

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<sup>3</sup>Monte-Carlo methods are well-established in the philosophical literature. Tentori et al. ([2007]) used them to test the agreement between Bayesian confirmation measures, Angere ([2007 2008]) and Koscholke et al. ([2018]) for probabilistic coherence measures. Wagner ([2013]) studied the prevalence of the so-called *corroboration paradox* and Douven ([2015]) investigated various closure principles for conditionals. Douven and Rott ([2018]) used it to find sets of informative beliefs according to different belief-rules. And recently, Fitelson and Crupi ([unpublished]) used them to study various types of Simpson-paradoxical phenomena.



**Figure 2:** Prevalence of transitivity-failure in Bayesian confirmation.

antecedent of transitivity is violated or the consequent is satisfied. The conjunctive prevalence of *non-trivial* instances of transitivity-success is around 16%.

The conditional prevalence stabilized at around 36%. This means that if  $A$  confirms  $B$  and  $B$  confirms  $C$ , inferring that  $A$  confirms  $C$  is more likely to be right than to be wrong, namely around 64%. Still, since this is a rather high chance of error, it is fair to conclude that Bayesian confirmation is not transitive *in most cases* (see also Bamber’s [2000] idea of entailment with near surety).

#### 4. Transitivity-Failure in Absolute Confirmation

Transitivity-failure occurs not only in *Bayesian* but also in *absolute* confirmation. On the latter conception, a proposition  $A$  confirms another proposition  $B$  relative to probability function  $P$  if and only if  $P(B|A) > \theta$ , where  $\theta$  is a threshold for high probability, presumably at least  $1/2$  (see Carnap [1950]). For an example of transitivity-failure for  $\theta = 1/2$ , consider again sampling from a well-shuffled standard deck of cards. This time, however, let  $A$  state that the drawn card is a Two, Three, Four, Five, or Six,  $B$  that the card is a Four, Five, Six, Seven, or Eight, and  $C$  that the card is a Six, Seven, Eight, Nine, or Ten (this example is due to Roche [2017]). Here we have the following probabilities (and for even stronger

**Table 2:** Transitivity-failure in absolute confirmation.

Threshold $\theta$	Conjunctive	Conditional
.5	0.05905960	0.23635440
.7	0.01328240	0.28435759
.9	0.00025400	0.30814024
.95	0.00001690	0.29911504
.99	0.00000000	0.00000000

counterexamples see Douven [2011]):

- $P(B|A) = 3/5 > \theta$
- $P(C|B) = 3/5 > \theta$
- $P(C|A) = 1/5 < \theta$

Still, there are also conditions under which absolute confirmation *is* transitive. For instance, drawing on Hesse’s ([1970]) work, Roche ([2017]) pointed out that transitivity holds for absolute confirmation whenever  $A$  is not negatively relevant for  $C$  conditional on  $B$  and the two conditional probabilities  $P(B|A)$  and  $P(C|B)$  are above the more demanding threshold for *super-high* probability  $\sqrt[3]{\theta}$ . One might therefore also be interested in how prevalent transitivity-failure is in absolute confirmation.

The results for different thresholds are shown in Table 2. It turns out that, just as in Bayesian confirmation, transitivity-failure in absolute confirmation is less prevalent than transitivity-success: the conjunctive prevalence ranges from 6% to 0.001690%, while the corresponding conjunctive prevalence of *non-trivial* instances of transitivity-success ranges from 19% to 0.0039600%. The latter can be calculated from the values shown above. Importantly, the conditional prevalence of transitivity-failure is also lower than the corresponding prevalence of transitivity-success. This means that transitivity-style inferences are more likely to be right than to be wrong for absolute confirmation, namely around 70-76%. And interestingly, this also means that transitivity-style inferences are more likely to be wrong for *Bayesian* than for *absolute* confirmation, namely around 36% versus 24-30%.

## 5. Varieties of Transitivity

Transitivity-failure means that whenever  $A$  confirms  $B$  and  $B$  confirms  $C$ , it does not follow that  $A$  confirms  $C$ . But one might be tempted to infer that  $A$  *in conjunction*

with  $B$  confirms  $C$ . That is, one might think that the following inference pattern, also discussed by Douven ([2011]), is valid for Bayesian confirmation:

**Conjunctive Transitivity.** If  $A$  confirms  $B$  and  $B$  confirms  $C$ , then the conjunction  $A \wedge B$  confirms  $C$ .

Similarly, one might be tempted to infer that  $A$  confirms  $C$  *conditional on*  $B$ , that is  $P(C|A \wedge B) > P(C|B)$ . To the best of my knowledge, the resulting inference pattern has not yet been discussed in the literature:

**Conditional Transitivity.** If  $A$  confirms  $B$  and  $B$  confirms  $C$ , then  $A$  confirms  $C$  conditional on  $B$ .

But our initial card example shows that both patterns are invalid for Bayesian confirmation. If  $A$  is that the drawn card is a heart,  $B$  that it is red and  $C$  that it is a diamond, we have:

- $P(B|A) = 1 > P(B) = 1/2$
- $P(C|B) = 1/2 > P(C) = 1/4$
- $P(C|A \wedge B) = 0 < P(C) = 1/4 < P(C|B) = 1/2$

One might also wonder about the validity of another transitivity-like pattern known from non-monotonic reasoning (see Kraus et al. [1990]):

**Cumulative Transitivity.** If  $A$  confirms  $B$  and  $A \wedge B$  confirms  $C$ , then  $A$  confirms  $C$ .

But this pattern is also invalid for Bayesian confirmation. To see this, consider the following variation of our card example for absolute confirmation. Let  $A$  state that the drawn card is a Two, Three, Four, Five, or Six,  $B$  that the card is a Four, Five, Six, Seven, or Eight, and  $C$  that the card is a Five, Six, Seven, Eight or Nine. We then have:

- $P(B|A) = 3/5 > P(B) = 5/13$
- $P(C|A \wedge B) = 2/3 > P(C) = 5/13$
- $P(C|A) = 2/5 < P(C) = 5/13$

**Table 3:** Failure-prevalence for other inference patterns. Here,  $>$  abbreviates Bayesian confirmation, a subscript indicates conditional confirmation. For an investigation of some of these patterns for absolute confirmation see Hawthorne ([1996]) or Hawthorne ([2007]).

Label	Pattern	Conjunctive	Conditional
Transitivity	If $A > B$ and $B > C$ , then $A > C$	0.090	0.359
Conjunctive Transitivity	If $A > B$ and $B > C$ , then $A \wedge B > C$	0.056	0.226
Conditional Transitivity	If $A > B$ and $B > C$ , then $A >_B C$	0.111	0.443
Cumulative Transitivity	If $A > B$ and $A \wedge B > C$ , then $A > C$	0.056	0.225
Agglomeration	If $B > A$ and $B > C$ , then $B > A \wedge C$	0.025	0.101
Cautious Monotonicity	If $B > A$ and $B > C$ , then $B \wedge C > A$	0.057	0.226
Rational Monotonicity	If $B > A$ and $B \not> \neg C$ , then $B \wedge C > A$	0.057	0.226
Corroboration	If $A > B$ and $C > B$ , then $A >_C B$ and $C >_A B$	0.092	0.368
Amalgamation	If $A > B$ and $C > B$ , then $A \vee C > B$	0.025	0.100

Interestingly, while each of the aforementioned transitivity-like patterns is invalid, there are significant differences between them: failures of conjunctive and cumulative transitivity are less prevalent than failures of conditional transitivity or standard transitivity; and somewhat surprisingly, conditional transitivity has the highest prevalence of failure. The exact values are shown in Table 3.<sup>4</sup>

## 6. Agglomeration and Monotonicity

On the Bayesian conception of confirmation, the antecedent of transitivity is equivalent to  $B$  confirming  $A$  and  $C$  individually. And Reichenbach ([1956]) famously showed that if additionally  $B$  screens-off  $A$  and  $C$  from each other, it follows that  $B$  confirms the conjunction  $A \wedge C$ . This corresponds to the following well-known inference pattern (see Leitgeb [2013] for the label):

**Agglomeration.** If  $B$  confirms  $A$  and  $B$  confirms  $C$ , then  $B$  confirms their conjunction  $A \wedge C$ .

And the antecedent of transitivity and of agglomeration are identical to the antecedent of the following weak form of monotonicity (see Gabbay [1985]):

**Cautious Monotonicity.** If  $B$  confirms  $A$  and  $B$  confirms  $C$ , then  $B \wedge C$  confirms  $A$ .

<sup>4</sup>Freund et al. ([1991]) also discuss a variation of transitivity with a strengthened antecedent. When adapted to confirmation, it reads:

**Weak Transitivity.** If  $A$  confirms  $B$ ,  $B$  confirms  $C$  and  $B$  does not confirm  $\neg A$ , then  $A$  confirms  $C$ .

But for Bayesian confirmation, the added conjunct is already implied by the first conjunct of the antecedent.



And cautious monotonicity is closely related to another weak form of monotonicity known from the logic of conditionals (see Lewis [1973]):

**Rational Monotonicity.** If  $B$  confirms  $A$  and  $B$  does not confirm  $\neg C$ , then  $B \wedge C$  confirms  $A$ .

But all three inference patterns are invalid for Bayesian confirmation. To see this, simply reconsider our initial card example where  $A$  is that the drawn card is a heart,  $B$  that it is red and  $C$  that it is a diamond. We then have:

- $P(A|B) = 1/2 > P(A) = 1/4$
- $P(C|B) = 1/2 > P(C) = 1/4$
- $P(A \wedge C|B) = P(A \wedge C) = 0$

Despite their invalidity, however, each of the three patterns has a lower prevalence of failure than transitivity. In fact, agglomeration-failure has the lowest prevalence of all investigated patterns. See Table 3 for the precise values.

## 7. Corroboration and Amalgamation

On the Bayesian conception of confirmation, the transitivity-antecedent is also equivalent to  $A$  and  $C$  individually confirming  $B$ . And Cohen ([1977]) showed that if additionally two weak conditional independence conditions hold, it follows that  $A$  and  $C$  *together* make  $B$  more likely than they do *individually*. This corresponds to the following inference pattern, also discussed by Wagner ([2013]):<sup>5</sup>

**Corroboration.** If  $A$  confirms  $B$  and  $C$  confirms  $B$ , then  $A$  confirms  $B$  conditional on  $C$  and  $C$  confirms  $B$  conditional on  $A$ .

Whenever  $A$  and  $C$  confirm  $B$  individually, one might also be tempted to infer that their disjunction  $A \vee C$  confirms  $B$ . This corresponds to the following inference pattern (see Smith [2016] for the label):

**Amalgamation.** If  $A$  confirms  $B$  and  $C$  confirms  $B$ , then their disjunction  $A \vee C$  confirms  $B$ .

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<sup>5</sup>For critical comments on Cohen's work see O'Neill ([1982]), Falk ([1986]) and Schlesinger ([1988]), but see also Cohen ([1982]), Cohen ([1986]), Cohen ([1991]) and Wagner ([1991]) for replies. Atkinson and Peijnenburg ([unpublished]) provide a nice overview.

But again, both aforementioned patterns are invalid for Bayesian confirmation. Our initial card example yields the following probabilities:

- $P(B|A) = 1 > P(B) = 1/2$
- $P(B|C) = 1 > P(B) = 1/2$
- $P(B|A \wedge C)$  is undefined and cannot be higher than  $P(B|A)$  or  $P(B|C)$
- $P(B|A \vee C) = P(B) = 1/2$

Interestingly, while amalgamation-failure is less prevalent than transitivity-failure, corroboration-failure is significantly more prevalent than the latter. Arguably, this is due to the fact that the corroboration-consequent is a conjunction and thus more difficult to satisfy. The exact numbers are again shown in Table 3.<sup>6</sup>

## 8. Conclusion

Bayesian confirmation may not be transitive, but the phenomenon of transitivity-failure is much less common than one might expect: the results above show that it is less prevalent than trivial as well as non-trivial instances of transitivity-success. Still, this does not mean that Bayesian confirmation is transitive *in most cases*: there is a non-negligible chance of error in transitivity-style inferences. And transitivity-failure is also more prevalent than failures of many adjacent inference patterns from non-monotonic reasoning and the logic of conditionals.

The above results can also be taken as an inspiration to further investigate the logic of Bayesian confirmation quantitatively rather than qualitatively. Scholars like Pearl ([1991]) have complained that the Bayesian notion of confirmation “in itself is too weak to yield interesting inferences” ([Pearl 1991, 160]). But this is only true if interesting inferences are required to be classically valid. The quantitative picture is much more complex. Then, as we have seen above, the notion of Bayesian confirmation yields many interesting inferences with varying degrees of validity. And there are probably more than the ones examined here.<sup>7</sup>

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<sup>6</sup>The MATLAB code underlying these results is available upon request.

<sup>7</sup>Do the results presented here mean that it is *rational* to infer, for example, the conclusion that *A* confirms *C* from the premises that *A* confirms *B* and that *B* confirms *C*? No. As the literature on the normativity of logic shows, it is already difficult enough to formulate rationality principles for *valid* inferences (see Steinberger [2017]). Accordingly, we should be even more careful when thinking about the implications the results might have for the rationality of *invalid* inferences. Thanks to an anonymous referee for pushing me to say a bit more about this.

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