【計量ファイナンスA】

15. High-frequency SV モデル

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Introduction

Estimation of time-varying volatility

- 1. GARCH and Stochastic Volatility (SV) models
 - First, we must model the dynamics of volatility.
 - Then, volatility is estimated by estimating the model.
 - Volatility estimates depend on the model.
- 2. Realized volatility (RV)
 - Volatility is estimated using intraday high-frequency returns.
 - We need not model the dynamics of daily volatility.

Introduction

- ► We like to beat RV by modeling the dynamics of intraday high-frequency volatility.
- ► It is not straightforward to model the dynamics of intraday high-frequency volatility.
 - (1) Intraday seasonality
 - (2) Announcement effects
- ▶ We also take account of
 - (3) Negative correlation between return and volatility
 - (4) Jumps in return and volatility
 - (5) Fat-tail return distribution

Introduction

This research

- We extend daily SV models to intraday high-frequency SV models.
- develop a Bayesian method using MCMC for the analysis of intraday SV models.
- ▶ apply them to 5-min returns of Nikkei 225 stock index.
- show
 - (1) Intraday SV models fit the data better than intraday GARCH-type models.
 - (2) Intraday SV models perform better than daily RV models such as HAR and realized EGARCH models in one-day-ahead volatility forecasting.

Intraday high-frequency return

$$y_t = V_t \varepsilon_t + J_t Z_t^y$$
, $\varepsilon_t \sim N(0,1)$ or standardized $t(\nu)$

- $ightharpoonup V_t = ext{total volatility}$
- $ightharpoonup J_t Z_t^y = \text{jump component}$

t distribution

$$arepsilon_t = \sqrt{\lambda_t} z_t, \quad rac{
u-2}{\lambda_t} \sim \chi^2(
u), \quad z_t \sim N(0,1)$$

• We assume $\nu > 2$ for a finite variance.

Total volatility

$$V_t = X_t S_t A_t$$
 or $h_t = x_t + s_t + a_t$

- $ightharpoonup X_t = \mathsf{SV} \; \mathsf{part}$
- \triangleright $S_t = intraday seasonality$
- $ightharpoonup A_t = announcement effects$
- $h_t = \log(V_t^2), x_t = \log(X_t^2), s_t = \log(S_t^2), a_t = \log(A_t^2)$

SV part

$$x_{t+1} = \mu + \phi(x_t - \mu) + J_t Z_t^{\nu} + \eta_t$$

$$\begin{pmatrix} \varepsilon_t \\ \eta_t \end{pmatrix} \sim N(\mathbf{0}, \mathbf{S}), \quad \mathbf{S} = \begin{pmatrix} 1 & \rho\sigma \\ \rho\sigma & \sigma^2 \end{pmatrix}$$

Jump components

- ▶ Jump components $(J_t Z_t^y, J_t Z_t^v)$ are assumed to be coincident in returns and the volatility with the common jump indicator variable, $J_t \in \{0, 1\}$.
- ▶ Jump occurs with the probability $Pr[J_t = 1] = \kappa$.
- ▶ We assume

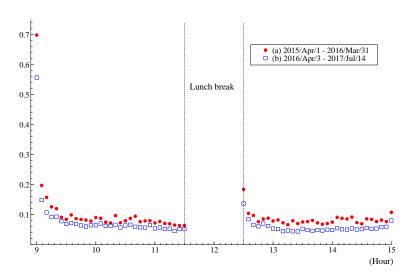
$$Z_t^y \sim N(\mu_y, \sigma_y^2), \quad Z_t^v \sim N(\mu_v, \sigma_v^2)$$

Intraday seasonality

- ▶ Trigonometric function: Harada and Watanabe (2009), Neely (2011), Tsuchida, Watanabe and Yoshiba (2016)
- ► Cubic smoothing spline: Stroud and Johaness (2014)
- Random walk: This paper

- ightharpoonup K =number of intraday returns during a day
- ightharpoonup N = number of days in the sample periods
- $ightharpoonup T = K \times N$: total sample size
- ► Tokyo stock exchange is open for 9:00-11:30 and for 12:30-15:00.
- ▶ We use 5-min returns when the market is open.
- ➤ We also use overnight (15:00-9:00) and lunch-time (11:30-12:30) returns.
- ▶ Then, K = 62.

Average intraday Nikkei 225 absolute 5-min returns



- ▶ β_k =intraday seasonality at intraday period k (k = 1, ..., K).
- $ightharpoonup H_{tk} = intraday-period indicator$

$$H_{tk} = \left\{ egin{array}{ll} 1 & ext{if time t corresponds to intraday period k} \\ 0 & ext{otherwise} \end{array}
ight.$$

 \triangleright Then, s_t is represented by

$$s_t = \sum_{k=1}^K H_{tk} \beta_k, \quad t = 1, \dots, T$$

▶ We assume that β_k follows the following random walk.

$$\beta_{k+1} = \beta_k + w_k, \quad w_k \sim N(0, c_k v_\beta^2), \quad k = 1, \dots, K - 1$$

 $\beta_1 \sim N(0, 100)$

- ▶ We set $c_k = 100$ for k = 1, 31, 32, 61 and $c_k = 1$ for other k.
- ▶ Identifying restriction: $\frac{1}{K} \sum_{k=1}^{K} \beta_k = 0$

Announcement effects

$$\begin{array}{lcl} \textbf{\textit{a}}_t & = & \sum_{j=1}^J \sum_{\ell=1}^L \textit{\textit{I}}_{jt\ell} \alpha_{j\ell} \\ \\ \alpha_{j,\ell+1} & = & \psi_j \alpha_{j\ell} + \zeta_{j\ell}, \quad \zeta_{j\ell} \, \sim \textit{\textit{N}}(\textbf{\textit{0}},\textit{\textit{v}}_{\alpha j}^2), \quad |\psi_j| < 1 \end{array}$$

- ightharpoonup J = 4 (GDP, IP, CPI, MPM)
- $ightharpoonup t_i^* = \text{announcement time of the } j \text{th variable}$
- ▶ $I_{jt\ell}$ = announcement-period indicator:

$$I_{jt\ell} = \left\{egin{array}{ll} 1 & ext{if } t \in \{t_j^*, t_j^*+1, \ldots, t_j^*+L-1\} \ 0 & ext{otherwise} \end{array}
ight.$$

L = 18 (i.e., 90 minutes)

Parameters

- 1. Parameters for SV: $\theta_X = (\phi, \sigma, \rho, \mu)$
- 2. Jump parameters: $\theta_J = (\kappa, \mu_y, \sigma_y, \mu_v, \sigma_v)$
- 3. Parameters for intraday seasonality and announcement effects: $\mathbf{v} = (v_{\beta}, v_{\alpha 1}, \dots, v_{\alpha J}), \ \psi = (\psi_{1}, \dots, \psi_{J})$
- 4. The degree of freedom for t distribution: ν

Latent variables

- 1. SV process: $x = (x_1, ..., x_T)$
- 2. Jump components:

$$oldsymbol{J} = (J_1, \dots, J_{T-1})$$

 $oldsymbol{Z}^{\varphi} = (Z_1^{\varphi}, \dots, Z_{T-1}^{\varphi}) ext{ for } \varphi = y, v$

3. Intraday seasonality and announcement effects:

$$m{eta} = (eta_1, \dots, eta_K)$$

 $m{lpha} = \{m{lpha}_1, \dots, m{lpha}_J\}$ where $m{lpha}_j = (lpha_{j1}, \dots, lpha_{jL})$

4. Latent variables for t distribution: $\lambda = (\lambda_1, \dots, \lambda_T)$

Bayesian analysis and computation

Estimation

- ▶ We develop a Bayesian estimation using MCMC.
- ▶ Parameters: $\boldsymbol{\theta} = \{\boldsymbol{\theta}_{\!X}, \boldsymbol{\theta}_{\!J}, \boldsymbol{v}, \boldsymbol{\psi}, \nu\}$
- ▶ Latent variables: $\Theta = \{x, J, Z^y, Z^v, \beta, \alpha, \lambda\}$
- We sample them from the joint posterior density $\pi(\theta, \Theta|y)$ using the Gibbs sampler.

Bayesian analysis and computation

- 1. Parameters for SV: straightforward
- 2. Jump parameters: straightforward
- 3. Parameters for the intraday seasonal and announcement effects: straightforward
- 4. The degree of freedom for the Student-*t* distribution: Watanabe (2001)
- 5. SV: Omori and Watanabe (2008)
- 6. Jump components: straightforward
- 7. Intraday seasonal and announcement effects: Watanabe and Omori (2004)
- 8. Latent variables for t distribution: Watanabe (2001)

Data

- 5-min intraday returns of Nikkei 225 stock index.
- ▶ Total sample period: April 1, 2015—July 14, 2017
- Estimation: April 1, 2015–March 31,2016
- ► Forecasting: April 1, 2016–July 14, 2017

Priors

$$(\phi+1)/2 \sim B(20,1.5), \ \sigma^2 \sim IG(40,0.2), \ (\rho+1)/2 \sim B(1,1), \ \mu \sim N(-5,4), \ \nu \sim G(16,0.8)I[\nu>2], \ \kappa \sim B(1,500), \ \mu_y \sim N(0,1), \ \mu_v \sim N(1,1), \ \sigma_i^2 \sim IG(20,4) \ (i=y,v), \ v_\beta^2 \sim IG(10,1), \ v_{\alpha,j}^2 \sim IG(10,1), \ (\psi_j+1)/2 \sim B(20,1.5)$$

- 1. **SV:** normal distribution, no jumps
- 2. **SVt:** Student *t*-distribution, no jumps
- 3. **SVJ:** normal distribution, with jumps in return and volatility
- 4. **SVJt:** Student *t*-distribution, with jumps in return and volatility

Posterior estimates of the selected parameters for the high-frequency SV models (2015/Apr – 2016/Mar)

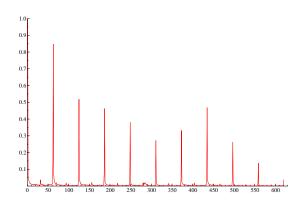
	SV	SVt	SVJ	SVJt
	0.9898 (0.0015)	0.9935 (0.0011)	0.9918 (0.0011)	0.9935 (0.0011)
ϕ	[0.9868, 0.9926]	[0.9911, 0.9956]	[0.9894, 0.9941]	[0.9913, 0.9955]
	5.0	16.5	12.6	20.1
	0.1558 (0.0070)	0.1213 (0.0069)	0.1082 (0.0068)	0.1090 (0.0068)
σ	[0.1411, 0.1694]	[0.1094, 0.1358]	[0.0956, 0.1224]	[0.0954, 0.1221]
	15.5	44.5	36.1	53.2
	-0.2279 (0.0301)	-0.2711 (0.0348)	-0.2989 (0.0384)	-0.2917 (0.0405)
ρ	[-0.2872, -0.1685]	[-0.3392, -0.2008]	[-0.3673, -0.2197]	[-0.3701, -0.2123]
	10.7	9.4	16.1	23.3

- ▶ The first row: posterior mean and standard deviation in parentheses.
- ▶ The second row: 95% credible interval in square brackets.
- ► The third row: inefficiency factor.

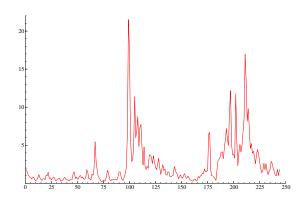
	SV	SVt	SVJ	SVJt
	-4.6803 (0.1228)	-4.6400 (0.1542)	-5.4439 (0.1762)	-4.9691 (0.1757)
μ	[-4.9192, -4.4400]	[-4.9405, -4.3284]	[-5.8133, -5.1185]	[-5.3179, -4.6423]
	0.5	1.0	22.0	30.3
	0.2325 (0.0201)	0.2328 (0.0212)	0.2430 (0.0217)	0.2317 (0.0209)
v_{β}	[0.1967, 0.2745]	[0.1946, 0.2806]	[0.2046, 0.2880]	[0.1944, 0.2764]
	0.9	0.9	1.4	1.0
	0.3241 (0.0521)	0.3225 (0.0522)	0.3240 (0.0512)	0.3222 (0.0515)
$v_{\alpha 1}$	[0.2417, 0.4470]	[0.2404, 0.4467]	[0.2423, 0.4414]	[0.2385, 0.4395]
	0.9	0.6	1.1	1.7
	0.7074 (0.1414)	0.7056 (0.1389)	0.7082 (0.1401)	0.7071 (0.1436)
ψ_1	[0.4223, 0.9430]	[0.4138, 0.9518]	[0.4231, 0.9563]	[0.4150, 0.9512]
	0.8	0.7	0.7	0.4
		10.889 (0.9258)		12.851 (1.3137)
ν		[9.2183, 12.784]		[10.759, 15.874]
		43.1		57.2

	SVJ	SVJt
	0.0091 (0.0018)	0.0017 (0.0008)
κ	[0.0057, 0.0130]	[0.0005, 0.0034]
	24.5	69.8
	-0.0134 (0.0407)	-0.1258 (0.1187)
μ_y	[-0.0970, 0.0632]	[-0.3744, 0.1048]
	9.4	31.7
	0.6391 (0.1198)	1.2841 (0.2916)
$\mu_{ m v}$	[0.4256, 0.8921]	[0.7553, 1.8865]
	15.4	47.3
	0.3301 (0.0255)	0.4155 (0.0410)
σ_y	[0.2853, 0.3821]	[0.3445, 0.5039]
	17.7	10.0
	0.4898 (0.0540)	0.4753 (0.0562)
$\sigma_{ m v}$	[0.3927, 0.6068]	[0.3787, 0.5988]
	17.7	7.3

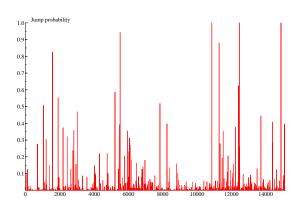
Posterior mean of total volatility V_t^2 (5-min in the first 10 days)



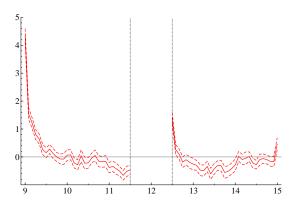
Posterior mean of total volatility V_t^2 (daily)



Posterior probability of jump (5-min)

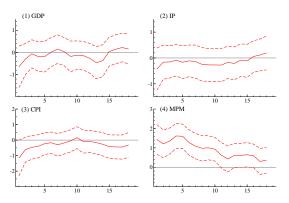


Posterior mean of parameters $(\beta_1, \ldots, \beta_K)$ for the intraday seasonality



▶ Posterior means (solid) and 95% credible intervals (dashed).

Announcement effects $\alpha_{j\ell}$ for GDP, industrial production (IP), consumer price index (CPI) and the Bank of Japan's monetary policy meeting (MPM)



Posterior means (solid) and 95% credible intervals (dashed).

► This result is consistent with

(1) Harada and Watanabe (2009): Yen/Dollar exchange rate

(2) Tsuchida, Watanabe and Yoshiba (2016): Japanese government bond futures

BIC for intraday SV and GARCH models (2015/Apr/1 – 2016/Mar/31)

Model	BIC	Ranking
SV	-30090	3
SVJ	-29723	4
SVt	-31016	1
SVJt	-30327	2
GARCH	-26275	9
GJR	-26341	8
EGARCH	-26264	10
GARCHt	-27315	7
GJRt	-27379	5
EGARCHt	-27371	6

Models for daily RV

► HAR model (Corsi 2009)

$$\log RV_{t}$$

$$= c + b_{d} \log RV_{t-1} + b_{w} \log RV_{t-1}^{(w)} + b_{m} \log RV_{t-1}^{(m)} + u_{t}$$

where

$$RV_{t-1}^{(w)} = \frac{1}{5} \sum_{i=1}^{5} RV_{t-i}, \quad RV_{t-1}^{(m)} = \frac{1}{22} \sum_{i=1}^{22} RV_{t-i}$$

Realized EGARCH model (Hansen and Huang 2016)

$$egin{array}{lcl} R_t & = & \sigma_t z_t, & z_t \sim {\it N}(0,1) \ & \log \sigma_{
m t}^2 & = & \omega + eta \ln(\sigma_{t-1}^2) + au_1 z_t + au_2(z_t^2-1) + \gamma v_{t-1} \ & \log {
m RV}_t & = & \xi + \log \sigma_{
m t}^2 + \delta_1 z_t + \delta_2(z_t^2-1) + v_t \ & v_t \sim {\it N}(0,\sigma_{
m v}^2) \end{array}$$

 Realized stochastic volatility model (Takahashi, Omori and Watanabe 2009, Takahashi, Watanabe and Omori 2016)

Out-of-sample forecast performance

- We compare the predictive ability of one-day-ahead volatility of our model with those of the HAR and realized EGARCH models.
- ► Following Patton (2011), we use the following loss functions.

$$\begin{aligned} \mathrm{MSE} &=& \frac{1}{N_1} \sum_{\tau = N_0 + 1}^{N_0 + N_1} \left(\widehat{\sigma}_{\tau}^2 - \sigma_{\tau}^2 \right)^2 \\ \mathrm{QLIKE} &=& \frac{1}{N_1} \sum_{\tau = N_0 + 1}^{N_0 + N_1} \left(\frac{\sigma_{\tau}^2}{\widehat{\sigma}_{\tau}^2} + \log \widehat{\sigma}_{\tau}^2 \right) \end{aligned}$$

where σ_{τ}^2 and $\hat{\sigma}_{\tau}^2$ denote the true volatility and the volatility forecast for day τ , respectively.

- Since the true volatility is unobserved, we must use the proxy.
- ▶ Patton (2011) shows that the MSE and QLIKE are robust loss functions in the sense that they lead to the same ranking as the one when the true volatility is used if the proxy is the unbiased estimator of the true volatility.
- ➤ We calculate the realized kernel (Barndorff-Nielsen et al 2008, 2009) using 1-min returns when the market is open.
- We convert it to the volatility estimates for 24 hours including non-trading hours such as lunch-time and overnight following Hansen and Lunde (2005).

One-day ahead volatility forecasting performance (2016/Apr/1 - 2017/July/14)

	140=	011175
Model	MSE	QLIKE
SV	10.85 [3]	1.381 [4]
SVJ	10.80 [2]	1.352 [3]
SVt	11.04 [4]	1.402 [5]
SVJt	10.32 [1]	1.305 [1]
HAR	11.34 [5]	1.309 [2]
REGARCH	11.60 [6]	1.417 [6]

Ranking is in brackets.

Conclusion

- We extend daily SV models to intraday high-frequency SV models.
- 2. develop an MCMC Bayesian method for the analysis of intraday high-frequency SV models.
- 3. apply them to 5-min returns of Nikkei 225 stock index.
- 4. show
 - (1) Intraday SV models fit the data better than intraday GARCH-type models.
 - (2) Intraday SV models perform better than daily RV models such as HAR and realized EGARCH models in one-day-ahead volatility forecasting.