

Incorporating Random effect into GP Sparse Approximation

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1 GP linear mixed model

1.1 Model specifications

The model specifications remain the same except that the random effect term has been added.

$$\begin{aligned}y|\theta &\sim \mathcal{N}(Z\alpha + A\beta, \gamma^2 I_n), (n \times n) \\ \alpha|\sigma^2 &\sim \mathcal{N}\left(0, \frac{\sigma^2}{m} I_{2m}\right), (2m \times 1) \\ \beta &\sim \mathcal{N}(0, \Sigma_\beta), (s \times 1) \\ \lambda &\sim \mathcal{N}(\mu_\lambda, \Sigma_\lambda), (d \times 1) \\ \sigma &\sim \text{half-Cauchy}(A_\sigma) \\ \gamma &\sim \text{half-Cauchy}(A_\gamma)\end{aligned}$$

where A is the design matrix for the random effects and β is the parameter vector of the random effects.

1.2 Lower bound

$$p(y, \theta) = \mathcal{N}(y|Z\alpha + A\beta, \gamma^2 I_n) \mathcal{N}\left(\alpha \middle| 0, \frac{\sigma^2}{m} I_{2m}\right) \mathcal{N}(\beta|0, \Sigma_\beta) \mathcal{N}(\lambda, \mu_\lambda, \Sigma_\lambda) \text{HC}(A_\sigma) \text{HC}(A_\gamma)$$

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$$\begin{aligned}
(1) \mathbb{E} [\log p(y|\theta)] &= -\frac{n}{2} \log(2\pi) - \frac{n}{2} \mathbb{E} [\log \gamma^2] \\
&\quad - \frac{1}{2} \mathbb{E} \left[\frac{1}{\gamma^2} \right] \left\{ y^T y - 2 (\mathbb{E} [Z] m_\alpha + A m_\beta)^T y + \text{Tr} (\mathbb{E} [Z^T Z] S_\alpha) + m_\alpha^T \mathbb{E} [Z^T Z] m_\alpha \right. \\
&\quad \left. + 2 m_\beta^T A^T \mathbb{E} [Z] m_\alpha + \text{Tr} (A^T A S_\beta) + m_\beta^T A^T A m_\beta \right\} \\
(2) \mathbb{E} [\log p(\alpha|\sigma)] &= -m \log(2\pi) - m \mathbb{E} [\log \sigma^2] + m \log m - \frac{m}{2} \mathbb{E} \left[\frac{1}{\sigma^2} \right] \text{Tr} (S_\alpha + m_\alpha m_\alpha^T) \\
(3) \mathbb{E} [\log p(\beta)] &= -\frac{s}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_\beta| - \frac{1}{2} \left\{ \text{Tr} (\Sigma_\beta^{-1} S_\beta) + m_\beta^T \Sigma_\beta^{-1} m_\beta \right\} \\
(4) \mathbb{E} [\log p(\lambda)] &= -\frac{d}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_\lambda| - \frac{1}{2} (m_\lambda - \mu_\lambda)^T \Sigma_\lambda^{-1} (m_\lambda - \mu_\lambda) - \frac{1}{2} \text{Tr} (\Sigma_\lambda^{-1} S_\lambda) \\
(5) \mathbb{E} [\log p(\sigma)] &= \log(2A_\sigma) - \log \pi - \mathbb{E} [\log (A_\sigma^2 + \sigma^2)] \\
(6) \mathbb{E} [\log p(\gamma)] &= \log(2A_\gamma) - \log \pi - \mathbb{E} [\log (A_\gamma^2 + \gamma^2)] \\
(1) \mathbb{E} [\log q(\alpha)] &= -m \log(2\pi) - \frac{1}{2} \log |S_\alpha| - m \\
(2) \mathbb{E} [\log q(\lambda)] &= -\frac{d}{2} \log(2\pi) - \frac{1}{2} \log |S_\lambda| - \frac{d}{2} \\
(3) \mathbb{E} [\log q(\beta)] &= -\frac{s}{2} \log(2\pi) - \frac{1}{2} \log |S_\beta| - \frac{s}{2} \\
(4) \mathbb{E} [\log q(\sigma)] &= -C_\sigma \frac{\mathcal{H}(2m, C_\sigma, A_\sigma^2)}{\mathcal{H}(2m-2, C_\sigma, A_\sigma^2)} - \log \mathcal{H}(2m-2, C_\sigma, A_\sigma^2) - 2m \mathbb{E} [\log \sigma] - \mathbb{E} [\log (A_\sigma^2 + \sigma^2)] \\
(5) \mathbb{E} [\log q(\gamma)] &= -C_\gamma \frac{\mathcal{H}(n, C_\gamma, A_\gamma^2)}{\mathcal{H}(n-2, C_\gamma, A_\gamma^2)} - \log \mathcal{H}(n-2, C_\gamma, A_\gamma^2) - n \mathbb{E} [\log \gamma] - \mathbb{E} [\log (A_\gamma^2 + \gamma^2)] \\
\mathcal{L} &= (1) + (2) + (3) + (4) + (5) + (6) - (1) - (2) - (3) - (4) - (5)
\end{aligned}$$

2 GP Logistic Model

2.1 Model specifications

For logistic models, we first postulate a link function, $g(\cdot)$ for the predictors.

$$\begin{aligned}
y &= g^{-1}(\eta) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \gamma^2 I_n) \\
g(\mathbb{E}[y]) &= \eta \\
\eta &= f(x) + A\beta \\
g^{-1}(x) &= e^x / (1 + e^x) \\
f &\sim \mathcal{GP}(m(\cdot), \kappa(\cdot, \cdot))
\end{aligned}$$

Since y is a Bernoulli random variable, $\mathbb{E}[y] = \mathbb{P}(y=1)$. Without loss of generality, we will assume the mean function to be zero and that the Gaussian process has a sparse approximation representation as in Tan & Nott. Therefore,

$$f(x) \approx \sum_{r=1}^m \left[a_r \cos \left\{ (s_r \odot x)^T \lambda \right\} + b_r \sin \left\{ (s_r \odot x)^T \lambda \right\} \right]$$

and this could further be represented in matrix form which reduces this to a linear model.

$$f(x) = Z\alpha$$

$$\begin{aligned}\eta &= Z\alpha + A\beta \\ y &= \exp \{ (Z\alpha + A\beta) - \log (\mathbf{1} + \exp \{ Z\alpha + A\beta \}) \} + \epsilon\end{aligned}$$

Every scalar function applied to a vector or a matrix is done so elementwise. We think of the following priors:

$$\begin{aligned}\alpha | \sigma &\sim \mathcal{N} \left(0, \frac{\sigma^2}{m} I_{2m} \right) \\ \beta &\sim \mathcal{N} (\mu_\beta, \Sigma_\beta) \\ \lambda &\sim \mathcal{N} (\mu_\lambda, \Sigma_\lambda) \\ \sigma &\sim \text{half-Cauchy} (A_\sigma) \\ \gamma &\sim \text{half-Cauchy} (A_\gamma) \\ \theta &= (\alpha, \beta, \lambda, \sigma, \gamma)\end{aligned}$$

$$\begin{aligned}\log p(y, \theta) &= y^T (Z\alpha + A\beta) - \mathbf{1}_n^T \log (\mathbf{1}_n + \exp \{ Z\alpha + A\beta \}) - \left(m + \frac{s+d}{2} \right) \log (2\pi) - m \log \sigma^2 + m \log m \\ &\quad - \frac{m}{2\sigma^2} \alpha^T \alpha - \frac{1}{2} \log |\Sigma_\beta| - \frac{1}{2} (\beta - \mu_\beta)^T \Sigma_\beta^{-1} (\beta - \mu_\beta) - \frac{1}{2} \log |\Sigma_\lambda| - \frac{1}{2} (\lambda - \mu_\lambda)^T \Sigma_\lambda^{-1} (\lambda - \mu_\lambda) \\ &\quad + \log (2A_\sigma) + \log (2A_\gamma) 2 \log \pi - \log (A_\sigma^2 + \sigma^2) - \log (A_\gamma^2 + \gamma^2)\end{aligned}$$

Because $-\mathbf{1}_n^T \log (\mathbf{1}_n + \exp \{ Z\alpha + A\beta \})$ is analytically intractable for expectation which is essentially integration, we come up with the following approximation:

$$\begin{aligned}-\log (1 + e^x) &= \max_{\xi \in \mathbb{R}} \left\{ B(\xi) x^2 - \frac{1}{2} x + C(\xi) \right\}, \quad \forall x \in \mathbb{R} \\ B(\xi) &= -\tanh (\xi/2) / (4\xi) \\ C(\xi) &= \xi/2 - \log (1 + e^\xi) + \xi \tanh (\xi/2) / 4\end{aligned}$$

then

$$\begin{aligned}-\mathbf{1}_n^T \log \{ \mathbf{1}_n^T + \exp (Z\alpha + A\beta) \} &\geq \mathbf{1}_n^T \left\{ B(\xi) \odot (Z\alpha + A\beta)^2 - \frac{1}{2} (Z\alpha + A\beta) + C(\xi) \right\} \\ &= (Z\alpha + A\beta)^T \text{Dg} \{ B(\xi) \} (Z\alpha + A\beta) - \frac{1}{2} \mathbf{1}_n^T (Z\alpha + A\beta) + \mathbf{1}_n^T C(\xi),\end{aligned}$$

where $\xi = (\xi_1, \dots, \xi_n)$.

$$\begin{aligned}\log \underline{p}(y, \theta; \xi) &= y^T (Z\alpha + A\beta) + (Z\alpha + A\beta)^T \text{Dg} \{ B(\xi) \} (Z\alpha + A\beta) - \frac{1}{2} \mathbf{1}_n^T (Z\alpha + A\beta) + \mathbf{1}_n^T C(\xi) \\ &\quad - \left(m + \frac{s+d}{2} \right) \log (2\pi) - m \log \sigma^2 + m \log m - \frac{m}{2\sigma^2} \alpha^T \alpha - \frac{1}{2} \log |\Sigma_\beta| - \frac{1}{2} (\beta - \mu_\beta)^T \Sigma_\beta^{-1} (\beta - \mu_\beta) \\ &\quad - \frac{1}{2} \log |\Sigma_\lambda| - \frac{1}{2} (\lambda - \mu_\lambda)^T \Sigma_\lambda^{-1} (\lambda - \mu_\lambda) + \log (2A_\sigma) + \log (2A_\gamma) 2 \log \pi - \log (A_\sigma^2 + \sigma^2) \\ &\quad - \log (A_\gamma^2 + \gamma^2)\end{aligned}$$