Bayesian Inference in action

Deep-dive into Facebook Prophet

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Frequentist vs Bayesian inference

1 Introduction

- Objective vs Subjective
- Frequency vs Belief
- Point estimate vs Posterior distribution
- Confidence intervals vs Credible intervals

When to use Bayesian inference?

1 Introduction

- Prior knowledge available
- Point estimates will not suffice
- Hierarchical modeling

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The problem

- X_i 1 (heads) or 0 (tails)
- $\{X_1, \ldots, X_n\}$ a sample of n coin tosses
- Is the coin balanced?

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The frequentist model

- p the probability of getting 1
 - toss the coin an infinite number of times
 - the frequency of 1s is p
- $X_i \sim Bernoulli(p)$

$$P(X_i = x) = \begin{cases} p & \text{if } x = 1\\ 1 - p & \text{if } x = 0 \end{cases}$$

•
$$Y = \sum_{i=1}^{n} X_i \implies Y \sim B(n, p)$$

$$P(Y_i = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

Estimating p from a sample

- \widehat{p} function of $\{X_1, \ldots, X_n\}$ called an *estimator*
- $\mathcal{L}(p) = P(\{X_1, \dots, X_n\}|p)$ the *likelihood* function
- $\hat{p} = \underset{p}{\text{arg max}} \mathcal{L}(p)$ the maximum likelihood estimator (MLE)
- For the coin toss problem, the MLE is:

$$\widehat{p} = \frac{\sum_{i=1}^{n} X_i}{n}$$

What is a confidence interval?

- MLE is a *point estimate*
 - We know nothing about other possible values of p
 - We have no notion of confidence (belief, certainty)...
 - ...because p is not a random variable
- What about the c% confidence interval?
 - $-- \ \{ \{X_{1,1}, X_{1,2}, ..., X_{1,n_1}\}, \{X_{2,1}, X_{2,2}, ..., X_{2,n_2}\}, ... \}$
 - c% confidence interval for each sample
 - c% of the "c% confidence intervals" contain the true value of p
 - Confidence in our method, not in our estimate!

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Everything is random!

- Why is p not a random variable?
- Random variables have probability distribution
- Probability distribution give us *real* confidence intervals *credible intervals*
- p is not really random our belief in p is random!

Life before the sample... and after too!

- If p is a random variable:
 - P(p=x) probability distribution before the sample prior
 - $P(p=x|\{X_1,\ldots,X_n\})$ probability distribution after the sample **posterior**
- How do we calculate the posterior, given the sample and the prior?
 - $P({X_1, ..., X_n}|p)$ the likelihood w.r.t. the prior

The Bayes theorem

2 The mandatory coin toss problem

$$P(p = x | \{X_1, \dots, X_n\}) = \frac{P(p = x, \{X_1, \dots, X_n\})}{P(\{X_1, \dots, X_n\})}$$

$$= \frac{P(p = x)P(\{X_1, \dots, X_n\} | p = x)}{P(\{X_1, \dots, X_n\})}$$

$$posterior = \frac{prior \times likelihood}{evidence}$$

• $P(\{X_1, ..., X_n\})$ is difficult to calculate, but it is constant w.r.t. the sample **posterior** \propto **prior** \times **likelihood**

Now back to tossing coins...

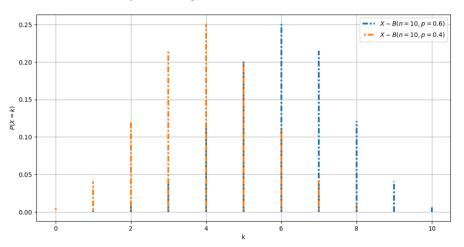
2 The mandatory coin toss problem

- Let's model $Y = \sum_{i=1}^{n} X_i$
- We already know the likelihood function:

$$P(Y = k|p) = \binom{n}{k} p^k (1-p)^{n-k}$$

• But how do we choose a prior?

The binomial likelihood



A different perspective

2 The mandatory coin toss problem

- Three "values" in the likelihood function
 - Y number of 1s
 - n number of trials
 - p probability of getting 1
- Y is a random variable, n and p are fixed

$$P(Y = k; n, p) = \binom{n}{k} p^k (1 - p)^{n-k}$$

• What if p was not fixed (a random variable)?

$$P(p=x;n,k) \propto x^k (1-x)^{n-k}$$

Introducing the Beta prior

2 The mandatory coin toss problem

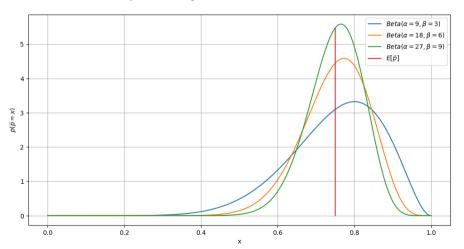
$$P(p=x;n,k) \propto x^k (1-x)^{n-k}$$

• How to replace \propto with =? Scale so that the probability density integrates to 1!

$$P(p = x; n, k) = \frac{x^{k}(1 - x)^{n - k}}{B(k + 1, n - k + 1)}$$

$$\alpha = k+1, \beta = n-k+1 \implies P(p=x;\alpha,\beta) = \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha,\beta)}$$

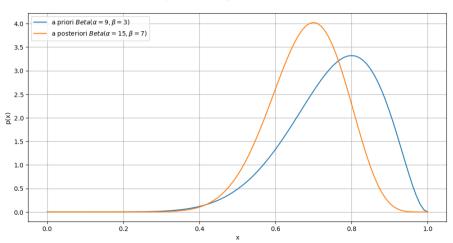
The Beta prior



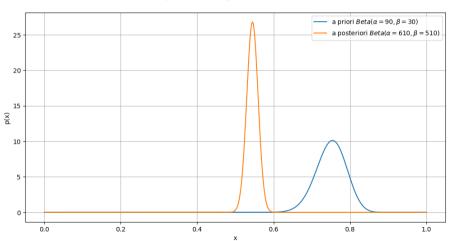
Calculating the posterior

- Conjugate priors
- Prior is $Beta(\alpha, \beta)$
- Likelihood is B(n, p)
- Sample is Y = k, n trials
- Posterior is $Beta(\alpha + k, \beta + n k)$

$Prior \rightarrow Posterior$



$Prior \rightarrow Posterior$



$Prior \rightarrow Posterior$

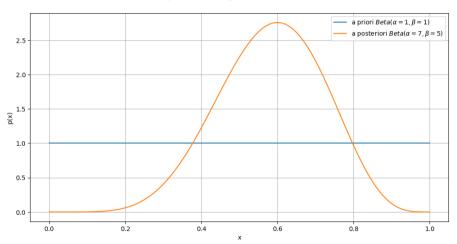


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Maximum a posteriori (MAP) estimate

- Point estimate
 - mode of the posterior distribution
- Regularized generalization of MLE
 - L1 regularization \Leftrightarrow Laplace prior + MAP
 - L2 regularization \Leftrightarrow Gaussian prior + MAP
- Fast, usually optimized with a gradient-based approach
 - L-BFGS-B

Markov chain Monte Carlo (MCMC) sampling 3 Bayesian inference techniques

- Explore full posterior
 - Markov assumption: the future is conditionally independent of the past, given the present
 - Monte Carlo: generate draws from the posterior
- Slow, hard to run in parallel
 - we must draw at step t before can draw at step t+1

MCMC algorithms

- Metropolis-Hastings (MH)
 - original, slow convergence
- NUTS
 - gradients used to detect high density regions

MCMC template

- 1. Calculate posterior of current point $P(p = x_{now}|Y)$
- 2. Calculate posterior of next point $P(p = x_{next}|Y)$
- 3. If $\frac{P(p=x_{next}|Y)}{P(p=x_{now}|Y)} > u; u \sim Uniform(0,1)$
 - 3.1 Add next point to trace
 - 3.2 Set current point to next point
- 4. Repeat until convergence

Variational Inference (VI)

- Middle-ground approach
- Analytical approximation to the true posterior
 - does not capture all essential characteristics
- Steps:
 - propose variational distribution
 - minimize KL divergence between true posterior and variational distribution
 - draw from variational distribution

VI algorithms

- ADVI
 - factorized Gaussian approximation
- FullRank ADVI:
 - ADVI + parameter correlations through a full-rank Gaussian

MAP vs MCMC vs VI

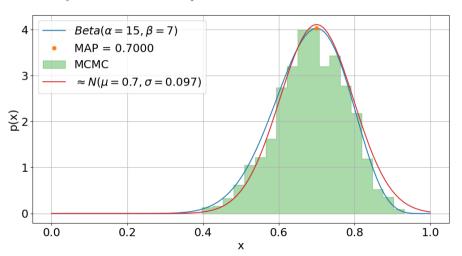


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Types of priors

- Informative priors
 - Derived from expert knowledge
 - No formal method to get to them
- Generic weakly informative priors
- Weakly informative priors
 - Principled regularization
- Vague priors
- Flat priors
 - Frequentist approach + posterior exploration

So what is the Normal distribution?

$$P(X = x; \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- $(x \mu)^2$ we are modeling the square distance from what we think our parameter should be!
- $2\sigma^2$ controls the spread of the bell shaped curve!
- $\frac{1}{\sqrt{2\pi\sigma^2}}$ a scaling monstrocity to make this integrate to 1!

What happens when we use a Normal prior?

$$\begin{array}{l} \mathbf{posterior} \propto \mathbf{prior} \times \mathbf{likelihood} \\ \mathbf{posterior} \propto \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \times \mathbf{likelihood} \\ \mathbf{logposterior} \propto -\frac{1}{2\sigma^2} (x-\mu)^2 + \mathbf{loglikelihood} \end{array} \label{eq:posterior}$$

- The scaling factor does not matter!
- If $\mu = 0$, this is L2 regularization with $\lambda = \frac{1}{2\sigma^2}$

$$\mathcal{L}_{L2} = \mathcal{L} + \lambda ||x||_2^2$$

So what is the Laplace distribution?

$$P(X = x; \mu, b) = \frac{1}{2b} e^{-\frac{|x-\mu|}{b}}$$

- $|x \mu|$ we are modeling the absolute distance from what we think our parameter should be!
- b controls the spread of the bell shaped curve!
- $\frac{1}{2b}$ scaling to make this integrate to 1!

What happens when we use a Laplace prior?

4 Let's talk about priors

$$\begin{array}{l} \mathbf{posterior} \propto \mathbf{prior} \times \mathbf{likelihood} \\ \mathbf{posterior} \propto \frac{1}{2b} e^{-\frac{|x-\mu|}{b}} \times \mathbf{likelihood} \\ \\ \mathbf{logposterior} \propto -\frac{1}{b} |x-\mu| + \mathbf{loglikelihood} \end{array} \tag{$$} / \log \\ \\ \end{array}$$

• If $\mu = 0$, this is L1 regularization with $\lambda = \frac{1}{b}$

$$\mathcal{L}_{L1} = \mathcal{L} + \lambda ||x||_1$$

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What is a time series?

- Given a set of time-points $T = \{t^{(k)} : k \in \mathbb{N}\} \subseteq \mathbb{R}$
- And a function $y: T \to \mathbb{R}$
- A time series is the set:

$$X = \{y(t) : t \in T\}$$

The Facebook Prophet model

5 Facebook Prophet

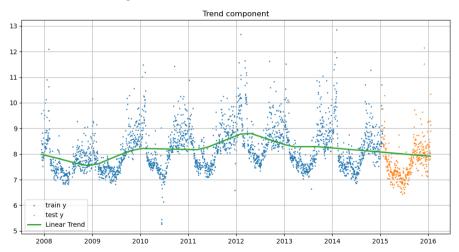
• Generalized additive model:

$$y(t; \boldsymbol{\gamma}, \boldsymbol{\theta}) = g(t; \boldsymbol{\gamma}) + \sum_{j=1}^{m} s(p^{(j)}, t; \boldsymbol{\theta}^{(j)}) + \epsilon^{(t)}$$

- g trend parametrized with γ
- s seasonality parametrized with $\theta^{(j)}$ for period $p^{(j)}$

$$P = \{p^{(j)} : j \in \mathbb{N}, 1 \le j \le m\} \subseteq \mathbb{R}$$
$$\forall k, j \in \mathbb{N}, j \le m$$
$$s(p^{(j)}, t_k) = s(p^{(j)}, t_k - p^{(j)})$$

Visual representation of Facebook Prophet



Visual representation of Facebook Prophet

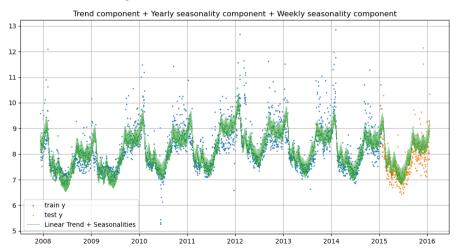
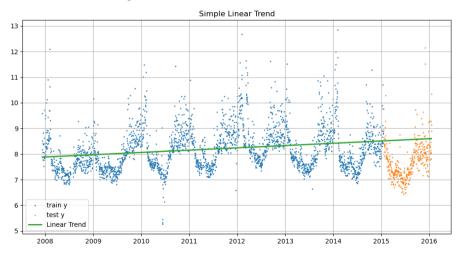


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The simple form - Linear Trend



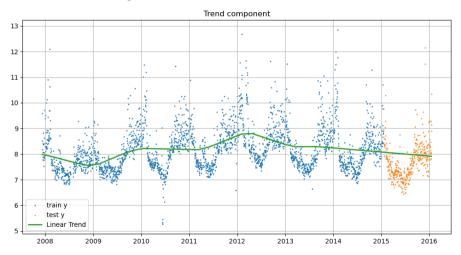
The simple form - Linear Trend

5 Facebook Prophet

• Two parameters: slope and intercept

$$w \sim \mathcal{N}(0,5)$$
$$b \sim \mathcal{N}(0,5)$$
$$\gamma = \begin{pmatrix} w & b \end{pmatrix}$$
$$g(t; \gamma) = wt + b$$

Introducing change points



Introducing change points

- Points in which w can chage
- $S \in \mathbb{N}$ number of change points
- $\mathbf{s} \in T^S$ column vector of change points
 - Set at equal distance
- $\boldsymbol{\delta} \in \mathbb{R}^S$ change in w at points \mathbf{s}
 - Random vector
 - $-\delta \sim Laplace(0, 0.05)$

The slope at every time-point

- $\mathbf{t} = \begin{pmatrix} t^{(1)} & t^{(2)} & \dots & t^{(|T|)} \end{pmatrix}$ row vector representation of T
- $\mathbf{w} \in \mathbb{R}^{|T|}$ column vector representation of the slope at every time-point

$$w_k = w + \sum_{l=1}^{t^{(k)} > s_l} \delta_l$$

Vectorizing the slope calculation

5 Facebook Prophet

• $\mathbf{a}_k \in \{0,1\}^S$ - row vector

$$a_k^{(l)} = \begin{cases} 1 & t^{(k)} > s_l \\ 0 & t^{(k)} \le s_l \end{cases} \implies w_k = w + \mathbf{a}_k \boldsymbol{\delta}$$

• $\mathbf{A} \in \mathbb{R}^{|T| \times S}$ - matrix, each row is \mathbf{a}_k

$$\mathbf{w} = w + \mathbf{A}\boldsymbol{\delta}$$

What about the intercept?

5 Facebook Prophet

• $\mathbf{b} \in \mathbb{R}^{|T|}$ - column vector representation of the intercept at every time-point

$$g(t; \gamma) = \begin{cases} wt + b & t < s_1 \\ (w + \delta_1)t + b_1 & t \in [s_1, s_2) \\ \dots & \\ (w + \sum_{l=1}^m \delta_l)t + b_m & t \in [s_m, s_{m+1}) \\ \dots & \\ (w + \sum_{l=1}^S \delta_l)t + b_S & t \ge s_S \end{cases}$$

Find b!

5 Facebook Prophet

• The trend is continuous if:

$$(w + \sum_{l=1}^{m-1} \delta_l)s_m + b_{m-1} = (w + \sum_{l=1}^{m} \delta_l)s_m + b_m$$

• Solve for b_m :

$$b_m = b_{m-1} - \delta_m s_m$$

• Special case:

$$ws_1 + b = (w + \delta_1)s_1 + b_1 \implies b_1 = b - \delta_1 s_1$$

Vectorize the intercept calculation

5 Facebook Prophet

• Recursively apply

$$b_1 = b - \delta_1 s_1; \qquad b_m = b_{m-1} - \delta_m s_m$$

$$\implies b_m = b - \sum_{l=1}^m \delta_l s_l = b - \sum_{l=1}^{t^{(k)} > s_l} \delta_l s_l$$

• Now we can use A

$$b_m = b - \mathbf{a}_m(\boldsymbol{\delta} \odot \mathbf{s}) \implies \mathbf{b} = b - \mathbf{A}(\boldsymbol{\delta} \odot \mathbf{s})$$

Final piece-wise linear trend form

5 Facebook Prophet

• All parameters

$$\gamma = \begin{pmatrix} w & b & \delta_1 & \dots & \delta_S \end{pmatrix}$$

• $\mathbf{g} \in \mathbb{R}^{|T|}$ - column vector representation of the linear trend at every time-point

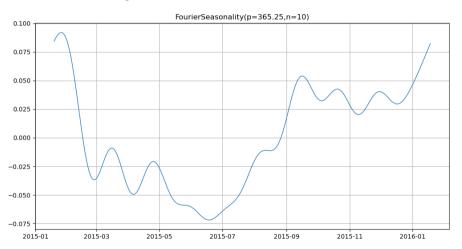
$$\mathbf{g} = \mathbf{w} \odot \mathbf{t}^T + \mathbf{b}$$
$$= (w + \mathbf{A}\boldsymbol{\delta}) \odot \mathbf{t}^T + b - \mathbf{A}(\boldsymbol{\delta} \odot \mathbf{s})$$

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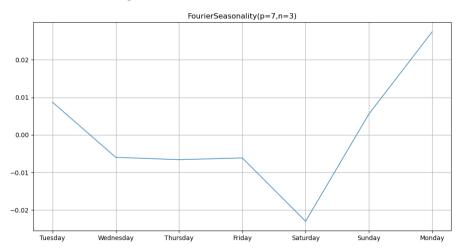
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Visual representation of seasonalities



Visual representation of seasonalities



The Fourier series

- Used for approximating smooth functions
- C_i number of sine and cosine terms for period $p^{(j)}$
- $\theta^{(j)}$ column vector of Fourier coefficients

$$\boldsymbol{\theta}^{(j)} = \begin{pmatrix} \beta_1^{(j)} & \beta_2^{(j)} & \dots & \beta_{2C_j-1}^{(j)} & \beta_{2C_j}^{(j)} \end{pmatrix}^T$$
$$s(p^{(j)}, t; \boldsymbol{\theta}^{(j)}) = \sum_{l=1}^{C_j} (\beta_{2l-1} cos(\frac{2\pi t}{p^{(j)}}) + \beta_{2l} sin(\frac{2\pi t}{p^{(j)}}))$$

Let's vectorize this!

- $\mathbf{f}_{k}^{(j)} \in \mathbb{R}^{2C_{j}}$ the sine and cosine terms for point $t^{(k)} \in T$ $\mathbf{f}_{k}^{(j)} = \left(cos(\frac{2\pi t^{(k)}}{p^{(j)}}), sin(\frac{2\pi t^{(k)}}{p^{(j)}}), ..., cos(\frac{2\pi C_{j} t^{(k)}}{p^{(j)}}), sin(\frac{2\pi C_{j} t^{(k)}}{p^{(j)}})\right)$
- $\mathbf{F}^{(j)} \in \mathbb{R}^{|T| \times 2C_j}$ matrix, each row is $\mathbf{f}_k^{(j)}$
- $\mathbf{fs}^{(j)} \in \mathbb{R}^{|T|}$ column vector representation of the Fourier seasonality at every time-point

$$\mathbf{f}\mathbf{s}^{(j)} = \mathbf{F}^{(j)}\boldsymbol{\theta}^{(j)}$$

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Why is the noise normal?

- The trend and seasonality model the systematic structure in the data
- The residual unsystematic, idiosyncratic noise—measurement error, unmodeled influences, natural variability
- ullet The sum of these independent residual factors is normal CLT
- Variants:
 - Student's t heavy tails
 - Conditional variance heteroscedasticity
 - Poisson, Negative Binomial count data

A hyper-prior for noise

$$y \sim \mathcal{N}(\mathbf{g} + \sum_{j=1}^{m} \mathbf{f}\mathbf{s}^{(j)}, \sigma)$$

- Modeling noise is a side effect of modeling the outcome
- σ is also a random variable

$$\sigma \sim \mathcal{HN}(0, 0.5)$$

The complete model

$$\begin{aligned} w &\sim \mathcal{N}(0, 5) \\ b &\sim \mathcal{N}(0, 5) \\ \boldsymbol{\delta} &\sim Laplace(0, 0.05) \\ \boldsymbol{\theta}^{(1)} &\sim \mathcal{N}(0, 10) \\ \boldsymbol{\theta}^{(2)} &\sim \mathcal{N}(0, 10) \\ \sigma &\sim \mathcal{H}\mathcal{N}(0, 0.5) \\ y &\sim \mathcal{N}(\mathbf{g} + \mathbf{f}\mathbf{s}^{(1)} + \mathbf{f}\mathbf{s}^{(2)}, \sigma) \end{aligned}$$

The complete model

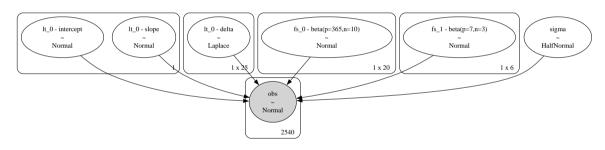


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Problem description

- Your store is selling winter equipment
- You offer a new product for three months and you need to forecast demand for the next year
- You know from your experience with the other products you sell that:
 - Your store is gaining traction (sales show an increasing trend)
 - There is a strong yearly seasonal effect you sell more in autumn and winter
- Our problem is harder
 - Our store was only selling one product
 - Multiple new products that we have sold for three months we need to forecast one year demand for each one of them
- How to use what you already know when modeling your 3 month samples?

The traditional Bayesian approach

- We can use Facebook Prophet with *informative* priors
- If we know that the trend is increasing, we can switch from $w \sim \mathcal{N}(0,5)$ to $w \sim \mathcal{N}(2,5)$
 - Or $w \sim \mathcal{N}(1,5)$?
- But what to do about the yearly seasonality?

The hierarchical Facebook Prophet

6 Future research

- Let's look at the slope
- Let w be the **shared** slope
- Let w_i be the **individual** slopes

$$w \sim \mathcal{N}(0,5)$$
$$w_i \sim \mathcal{N}(w, \frac{5}{\text{shrinkage}})$$

• Partial pooling

Posterior \rightarrow prior

- Complete Bayesian setup
 - prior \rightarrow posterior \rightarrow prior $\rightarrow \dots$
- $\operatorname{prior} \to \operatorname{posterior}$
 - Bayes theorem
- posterior \rightarrow prior
 - Keep distribution shape, estimate parameters from posterior sample
 - MvN approximation of the posterior as a prior
- Fit Prophet on the large time series to get posterior
- Use posterior as prior for Prophet when fitting short time series

Why does it work?

- Empirical priors
- Transfer learning + principled regularization
- The loc of the posterior transfer learning
- The scale of the posterior uncertainty estimate, principled regularization
- We also do standard regularization by adding "potentials"
 - But let's keep that discussion for another time

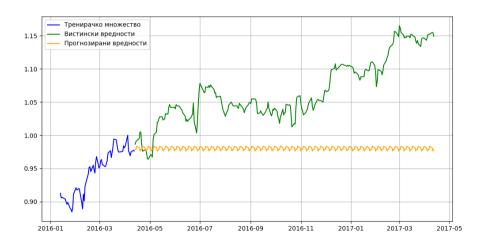
Combined approach

- Fit Prophet on the large time series to get posterior
- Use posterior as prior for the shared parameters in the hierarchical prophet
- Result: **shrinkage** does not need to be tuned!
- Maybe we don't even need to change the loc of the shared parameters?

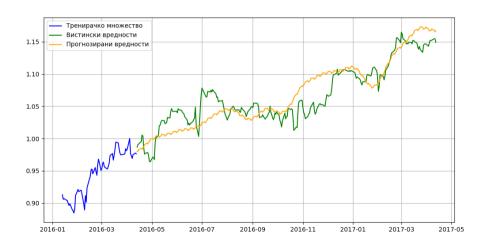
Results

	\mathbf{MSE}	\mathbf{RMSE}	\mathbf{MAE}	\mathbf{MAPE}	+
Prophet	0.1147	0.2696	0.2335	0.2421	/
Timeseers	0.0505	0.1801	0.1535	0.1541	36.35%
Vangja hierarchical	0.0413	0.1640	0.1424	0.1384	10.19%
Vangja a posteriori	0.0380	0.1583	0.1342	0.1324	4.34%
Vangja combined	0.0348	0.1493	0.1284	0.1297	2.04%

Results



Results



Q&A

Thank you for listening! Your feedback will be highly appreciated!