

HW 3 Solution

1.

$$E(X) = \sum_{k=1}^{\infty} k p (X=k) = c \sum_{k=1}^{\infty} p^k = \frac{cp}{1-p}$$

$$\begin{aligned} E(X^2) &= \sum_{k=1}^{\infty} k^2 p (X=k) = c \sum_{k=1}^{\infty} k p^k \\ &= cp \sum_{k=1}^{\infty} k p^{k-1} = \frac{cp}{(1-p)^2} \end{aligned}$$

$$VAR(X) = \frac{cp}{(1-p)^2} - \frac{c^2 p^2}{(1-p)^2} = \frac{cp(1-cp)}{(1-p)^2}$$

2. Let X be the number of children needed, from the 2nd child, to obtain different gender from the 1st baby. Then ~~is~~ $X-1$ is the $Geom(\frac{1}{2})$, hence $E(X) = 2$, hence the Expected total number of children = 3.

3. 20 drops per square inch per minute

of drops per 5 square inch per 3-second

$$= 100/20 = 5, \quad \lambda = 5$$

P (no rain drops per 5 inch² in 3 secs)

$$= \underset{\lambda=5}{P(k=0)} = \frac{e^{-5} \lambda^{k=0}}{k(k=0)!} = e^{-5} = 0.0067$$

4. a) let I_i be the indicator r.v. for the i th person being a mutual friend.

$$E\left(\sum_{i=1}^{1000} I_i\right) = 1000 E(I_i) = 1000 P(I_i = 1) \\ = 1000 \left(\frac{5}{1000}\right)^2 = 2.5$$

$$b) P(X=k) = \frac{\binom{50}{k} \binom{950}{50-k}}{\binom{1000}{50}} \quad X \sim \text{HGeom}(50, 950, 1000)$$

5. let I_i be the indicator r.v. for the i th box being empty
The # of empty boxes $= I_1 + \dots + I_N$

$$E\left(\sum_{i=1}^n I_i\right) = \sum_{i=1}^n E(I_i) = n \left(1 - \frac{1}{n}\right)^k$$

6. a) let N be the number of trials to see both success
 $(\underbrace{X=S, Y=S}_{P_1 P_2}), (X=S, Y=F), (X=F, Y=S), (X=F, Y=F)$

Hence, $N \sim \text{FS}(P_1 P_2)$ and $P(N=n) = P_1 P_2 (1 - P_1 P_2)^{n-1}$
 $E(N) = \frac{1}{P_1 P_2}$

- b) let T be the number of trials to see the first any success

$$(X=S, Y=S), (X=S, Y=F), (X=F, Y=S), (X=F, Y=F) \quad (1-P_1)(1-P_2)$$

$$T \sim \text{FS}(1 - (1-P_1)(1-P_2)) \quad E(T) = \frac{1}{1 - (1-P_1)(1-P_2)}$$

c) $P(T_X = T_Y) = \sum_{n=1}^{\infty} P(T_X = n | T_Y = n) P(T_Y = n)$
 $= \sum_{n=1}^{\infty} \frac{P}{2} P(T_X = n) P(T_Y = n) = \sum_{n=1}^{\infty} P_1^2 (1-P_1)^{2(n-1)}$
 $= \frac{1}{2 - P}$
 $1 = P(T_X < T_Y) + P(T_Y < T_X) + P(T_X = T_Y) = 2P(T_X < T_Y) + P(T_X = T_Y)$
 $\therefore P(T_X < T_Y) = \frac{1}{2} \cdot \left(1 - \frac{P}{2 - P}\right) = \frac{1 - P}{2 - P}$

7.

$$\begin{aligned}
 a) E(Xg(X)) &= e^{-\lambda} \sum_{k=0}^{\infty} k g(k) \frac{\lambda^k}{k!} \\
 &= \lambda e^{-\lambda} \sum_{k=1}^{\infty} g(k) \frac{\lambda^{k-1}}{(k-1)!} \\
 &= \lambda e^{-\lambda} \sum_{k=0}^{\infty} g(k+1) \frac{\lambda^k}{k!} \\
 &= \lambda E(g(X+1))
 \end{aligned}$$

$$\begin{aligned}
 b) E(X^3) &= E(X \cdot X^2) = \lambda E(X+1)^2 \\
 &= \lambda E(X^2) + 2\lambda E(X) + \lambda
 \end{aligned}$$

$$E(X) = X E(1) = \lambda$$

$$E(X^2) = \lambda E(X+1) = \lambda^2 + \lambda,$$

$$\therefore E(X^3) = \lambda(\lambda^2 + \lambda + 2\lambda + 1) = \lambda^3 + 3\lambda^2 + \lambda$$

8.

$$\begin{aligned}
 a) E(e^{tx}) &= e^{-\lambda} \sum_{k=0}^{\infty} e^{tk} \frac{\lambda^k}{k!} = e^{-\lambda} \sum \frac{(e^t \lambda)^k}{k!} \\
 &= e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t - 1)}
 \end{aligned}$$

$$\begin{aligned}
 b) E(e^{tx}) &= p \sum_{k=0}^{\infty} e^{tk} (1-p)^k = p \sum_{k=0}^{\infty} ((1-p)e^t)^k \\
 &= \frac{p}{1 - (1-p)e^t} \quad ((1-p)e^t < 1)
 \end{aligned}$$