

COSE 382 HW 7 Solutions

1. A fair 6-sided die is rolled once. Find the expected number of additional rolls needed to obtain a value at least as large as that of the first roll.

Solution:

Let X be the result of the first roll and Y be the number of rolls to obtain a value $\geq X$. The conditional distribution of Y given $X = k$ is $FS\left(\frac{7-k}{6}\right)$, i.e.,

$$\begin{aligned}P(Y | X = h) &\sim FS\left(\frac{7-k}{6}\right) \\E(Y) &= \sum_{k=1}^6 E(Y | X = k)P(X = k) = \sum_{k=1}^6 \frac{b}{7-k} \cdot \frac{1}{6} \\&= \sum_{k=1}^6 \frac{1}{k}\end{aligned}$$

2. Let X_1, X_2 be i.i.d., and let $\bar{X} = \frac{1}{2}(X_1 + X_2)$ be the sample mean. In many statistics problems, it is useful or important to obtain a conditional expectation given \bar{X} . As an example of this, find $E(w_1X_1 + w_2X_2 | \bar{X})$, where w_1, w_2 are constants with $w_1 + w_2 = 1$.

Solution:

$$2E(w_1X_1 + w_2X_2 | \bar{X}) = w_1E(X_1 | \bar{X}) + w_2E(X_2 | \bar{X})$$

Note that $E(X_1 | \bar{X}) = E(X_2 | \bar{X})$ by symmetry and

$$E(X_1 + X_2 | \bar{X}) = (X_1 + X_2)E(1 | \bar{X}) = (X_1 + X_2)$$

Hence, $E(X_1 | \bar{X}) = E(X_2 | \bar{X}) = \bar{X}$ Thus,

$$E(w_1X_1 + w_2X_2 | \bar{X}) = w_1\bar{X} + w_2\bar{X} = \bar{X}$$

3. Let X be the height of a randomly chosen adult man, and Y be his father's height, where X and Y have been standardized to have mean 0 and standard deviation 1. Suppose that (X, Y) is Bivariate Normal, with $X, Y \sim N(0, 1)$ and $\text{Corr}(X, Y) = \rho$.

- (a) Find a constant c (in terms of ρ) and an r.v. V such that $Y = cX + V$, with V independent of X .
- (b) Find a constant d (in terms of ρ) and an r.v. W such that $X = dY + W$, with W independent of Y .
- (c) Find $E(Y|X)$ and $E(X|Y)$.

Solution:

- a) Let $V := Y - cX$. To V and X be independent $\text{Corr}(V, X) = 0$ (since V and X are also bivariate)

$$\begin{aligned}\text{cor}(Y, X) &= E((Y - cX)X) \\ &= E(YX) - c(X^2) = 0 \\ &\Rightarrow c = \rho\end{aligned}$$

- b) By the same argument, $d = \rho$!

- c)

$$\begin{aligned}E(Y | X) &= E(\rho X + V | X) = \rho E(X | X) + E(V | X) \\ &= \rho X \\ E(X | Y) &= E(\rho Y + W | X) = \rho E(Y | Y) + E(W | Y) \\ &= \rho Y\end{aligned}$$

Solution:

4. Let X and Y be random variables with finite variances, and let $W = Y - E(Y|X)$. This is a residual: the difference between the true value of Y and the predicted value of Y based on X .

- (a) Compute $E(W)$ and $E(W|X)$.

- (b) Compute $\text{Var}(W)$, for the case that $W|X \sim N(0, X^2)$ with $X \sim N(0, 1)$.

Solution:

- a)

$$\begin{aligned}E(W) &= E(Y) - E(E(Y | X)) = EY - EY = 0 \\ E(W | X) &= E(Y | X) - E(E(Y | X) | X) \\ &= E(Y | X) - E(Y | X) = 0\end{aligned}$$

- b)

$$\begin{aligned}\text{Var}(W) &= \text{Var}(E(W | X)) + E(\text{Var}(W | X)) \\ &= \text{Var}(0) + E(X^2) = 1\end{aligned}$$

5. Show that if $E(Y|X) = c$ is a constant, then X and Y are uncorrelated.

Solution:

When $E(Y | X) = C$,

$$\begin{aligned}E(Y) &= E(c) = c \\ E(XY) &= E(E(XY | X)) = E(XE(Y | X)) \\ &= E(cX) = cE(X) = E(Y)E(X),\end{aligned}$$

6. In a national survey, a random sample of people are chosen and asked whether they support a certain policy. Assume that everyone in the population is equally likely to be surveyed at each step, and that the sampling is with replacement. Let n be the sample size, and let \hat{p} and p be the proportion of people who support the policy in the sample and in the entire population, respectively. Show that for every $c > 0$,

$$P(|\hat{p} - p| > c) \leq \frac{1}{4nc^2}$$

Solution:

Let $X \sim \text{Bin}(n, p)$, then $\hat{p} = \frac{X}{n}$

$$\begin{aligned} E(\hat{p}) &= \frac{E(X)}{n} = \frac{np}{n} = p \\ \text{Var}(\hat{p}) &= \frac{\text{Var}(X)}{n^2} = \frac{np(1-p)}{n^2} = \frac{p(1-p)}{n} \end{aligned}$$

By Chebyshev inequality

$$P(|\hat{p} - p| > c) \leq \frac{\text{Var}(\hat{p})}{c^2} = \frac{p(1-p)}{nc^2}$$

Since $p(1-p) \leq \frac{1}{4}$ for $\forall p \in (0, 1)$

$$P(|\hat{p} - p| > c) \leq \frac{1}{4nc^2}$$

7. For i.i.d. r.v.s X_1, \dots, X_n with mean μ and variance σ^2 , find a value of n which will ensure that there is at least a 99% chance that the sample mean will be within 2 standard deviations of the true mean μ .

Solution:

We have to find n such that

$$P(|\bar{X}_n - \mu| > 2\sigma) \leq 0.01$$

By Chebyshev inequality.

$$P(|\bar{X}_n - \mu| > 2\sigma) \leq \frac{\text{Var}(\bar{X}_n)}{(2\sigma)^2} = \frac{\frac{\sigma^2}{n}}{4\sigma^2} = \frac{1}{4n}$$

8. Let X and Y be i.i.d. positive r.v.s, and let $c > 0$. For each part below, fill in the appropriate equality or inequality symbol. If no relation holds in general, write “?”.

(a) $E(X) \quad \sqrt{E(X^2)}$

(b) $P(X > c) \quad E(X^3)/c^3$

(c) $E(X^3) \quad \sqrt{E(X^2)E(X^4)}$

(d) $P(|X + Y| > 2) \quad \frac{1}{10}E((X + Y)^4)$

(e) $E(Y|X) \quad E(Y|X + 3)$

(f) $P(X + Y > 2) \quad (EX + EY)/2$

Solution:

a) $E(X)$ v.s. $\sqrt{E(X^2)}$ equivalently $E^2(X)$ v.s. $E(X^2)$.

Since $\text{Var}(X) = E(X^2) - E^2(X) \geq 0$,

$$E(X^2) \geq E^2(X) \Rightarrow \sqrt{E(X^2)} \geq E(X)$$

b)

$$P(X > c) = P(X^3 > c^3) = P(|X^3| > c^3)$$

By Markov inequality,

$$P(X^3 > c^3) \leq \frac{E(X^3)}{c^3}$$

(c)

$$E(X^3) = E(X^2 \cdot X) = |E(X^2 \cdot X)|$$

By Cauchy-Schwartz

$$\begin{aligned} |E(X^2 \cdot X)| &\leq \sqrt{E(X^4) \cdot E(X^2)} \\ E(X^3) &\leq \sqrt{E(X^4) \cdot E(X^2)} \end{aligned}$$

d) $P(|X + Y| > 2) = P((X + Y)^4 \geq 16)$ By Markov inequality

$$P((X + Y)^4 \geq 16) \leq \frac{E((X + Y)^4)}{16} \leq \frac{1}{10} E((X + Y)^4)$$

e) $E(Y | X) = E(Y | X + 3)$

f)

$$\begin{aligned} P(X + Y > 2) &= P(|X + Y| \geq 2) \\ P(|X + Y| \geq 2) &\leq \frac{E(X + Y)}{2} = \frac{E(X) + E(Y)}{2} \end{aligned}$$

9. Consider i.i.d. $\text{Pois}(\lambda)$ r.v.s X_1, X_2, \dots . The MGF of X_j is $M(t) = e^{\lambda(e^t - 1)}$.

(a) Find the MGF $M_n(t)$ of the sample mean $\bar{X}_n = \frac{1}{n} \sum_{j=1}^n X_j$

(b) Find the limit of $M_n(t)$ as $n \rightarrow \infty$

Solution:

a)

$$\begin{aligned} M_n(t) &= E(e^{t\bar{X}_n}) = E\left(e^{\frac{t}{n}(X_1 + \dots + X_n)}\right) \\ &= \left(E\left(e^{\frac{t}{n}X_1}\right)\right)^n = e^{n\lambda(e^{t/n} - 1)} \end{aligned}$$

b)

$$\lim_{n \rightarrow \infty} M_n(t) = e^{t\lambda} \text{ since } \bar{X}_n \rightarrow \lambda \text{ as } n \rightarrow \infty$$

10. Let $Y = e^X$, with $X \sim \text{Expo}(3)$.

(a) Find the mean and variance of Y .

(b) For Y_1, \dots, Y_n i.i.d. with the same distribution as Y , what is the approximate distribution of the sample mean $\bar{Y}_n = \frac{1}{n} \sum_{j=1}^n Y_j$ when n is large?

Solution:

a)

$$\begin{aligned} E(Y) &= \int_0^\infty e^x (3e^{-3x}) dx = \frac{3}{2} \\ E(Y^2) &= \int_0^\infty e^{2x} (3e^{-3x}) dx = 3 \\ E(Y) &= \frac{3}{2}, \text{Var}(Y) = 3 - \frac{9}{4} = 3/4 \end{aligned}$$

b) By Central Limit Theorem, $\bar{Y}_n \simeq N\left(\frac{3}{2}, \frac{3}{4n}\right)$ for large n