

HW 4 solution

1.

a) from $F(x) \geq 0$, $f(x) \geq 0$
 $g(x) \geq 0$

$$\int g(x) dx = 2 \int F(x) f(x) dx, \quad \text{let } y = F(x)$$

then $dy = f(x) dx$
 $F(-\infty) = 0, F(\infty) = 1$
 $0 \leq y \leq 1$

$$= 2 \int_0^1 y dy = 1$$

b) $h(x) \geq 0$,

$$\int_{-\infty}^{\infty} h(x) dx = \frac{1}{2} \int_{-\infty}^{\infty} f(x) dx + \frac{1}{2} \int_{-\infty}^{\infty} f(x) dx = 1$$

2. $0 \leq R \leq 1$.

a) Let $U \sim \text{Unif}(0, 1)$, and $X = \min(U, 1-U)$ ($0 \leq x \leq \frac{1}{2}$)

$$P(R \leq r) = P(X \leq r(1-X)) = P\left(X \leq \frac{r}{1+r}\right)$$

$$P(X \leq x) = 1 - P(X > x) = 1 - P(\min(U, 1-U) > x)$$

$$= 1 - P(U > x \text{ and } 1-U > x)$$

$$= 1 - P(U > x \text{ and } 1-x > U)$$

$$= 1 - P(x < U < 1-x) = 2x \quad (0 \leq x \leq \frac{1}{2})$$

$$\therefore P(R \leq r) = \frac{2r}{1+r}$$

$$F_R(r) = \frac{2r}{1+r} \quad 0 \leq r \leq 1$$

$$f_R(r) = \frac{2(1+r) - 2r}{(1+r)^2} = \frac{2}{(1+r)^2}$$

b) $E(R) = 2 \int_0^1 \frac{r}{(1+r)^2} dr = 2 \int_1^2 \frac{(t-1)}{t^2} dt$

$$= 2 \int_1^2 \frac{1}{t} dt - 2 \int_1^2 \frac{1}{t^2} dt = 2 \ln 2 - 1$$

$$b) \quad L = \min(T_1, \dots, T_n)$$

$$P(L > t) = P(T_1 > t, \dots, T_n > t) = e^{-n\lambda t}$$

$$F_L(t) = 1 - e^{-n\lambda t}$$

$$E(L) = \frac{1}{n\lambda}, \quad \text{var}(L) = \frac{1}{(n\lambda)^2}$$

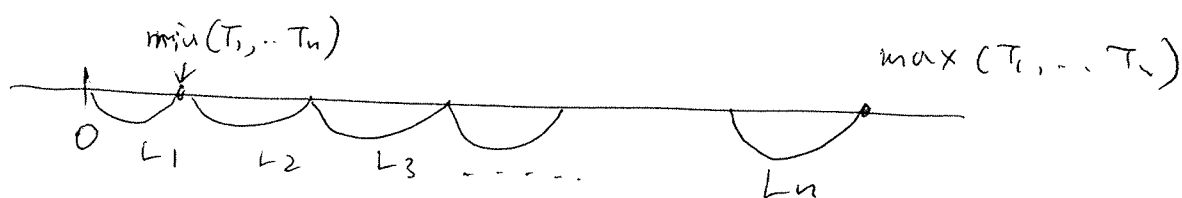
$$c) \quad \text{let } L_1 = \min(T_1, \dots, T_n)$$

$$L_2 = \min(T_1, \dots, T_n \text{ except } L_1) - L_1$$

$$L_3 = \min(T_1, \dots, T_n \text{ except } L_1, L_2 + L_1) - L_2$$

$$L_i = \min(T_1, \dots, T_n \text{ except } \{L_1, L_1 + L_2, \dots, L_1 + L_2 + \dots + L_{i-1}\}) - L_{i-1}$$

$$L_n = \max(T_1, \dots, T_n) - L_{n-1}$$



Notice that $L_1 \sim \text{Expo}(\lambda n)$

$$L_2 \sim \text{Expo}(\lambda(n-1)) \leftarrow \text{minimum among } (n-1) \text{ Expo.}$$

$$\vdots$$

$$L_i \sim \text{Expo}(\lambda(n-i+1))$$

$$L_n \sim \text{Expo}(\lambda)$$

Furthermore, L_i and L_j are independent ($i \neq j$)

Hence

$$L_1 + \dots + L_n = M$$

$$E(M) = \sum_{i=1}^n E(L_i) = \frac{1}{\lambda} \sum_{i=1}^n \frac{1}{i}$$

$$\text{Var}(M) = \sum_{i=1}^n \text{Var}(L_i) = \frac{1}{\lambda^2} \sum_{i=1}^n \frac{1}{i^2}$$

6. The MGF of $U \sim \text{Unif}(0,1)$

$$E(e^{tU}) = \int_0^1 e^{tu} du = \frac{1}{t} (e^t - 1)$$

Let $X = U_1 + \dots + U_{60}$,

$$E(e^{tX}) = (E(e^{tU}))^{60} = \frac{(e^t - 1)^{60}}{t^{60}}$$

$$\begin{aligned} 7. \quad M_L(t) &= E(e^{t(X-Y)}) = E(e^{tX}) \cdot E(e^{-tY}) \\ &= \left(\frac{1}{1-t}\right) \left(\frac{1}{1+t}\right) = \frac{1}{1-t^2} \end{aligned}$$

$$\begin{aligned} M_{\text{Laplace}}(t) &= \frac{1}{2} \int_{-\infty}^0 e^{tw} e^{+w} dw + \frac{1}{2} \int_0^{\infty} e^{tw} e^{-w} dw \\ &= \frac{1}{2(1+t)} + \frac{1}{2(1-t)} = \frac{1}{1-t^2} \end{aligned}$$

Hence $M_L(t) = M_{\text{Laplace}}(t) \Rightarrow f(x) = \frac{1}{2} e^{-|x|}$

$$\begin{aligned} 8. \quad M_W(t) &= E(e^{t(X^2+Y^2)}) = E(e^{tX^2} \cdot e^{tY^2}) \\ &= E(e^{tX^2}) \cdot E(e^{tY^2}) \\ &= (1-2t)^{-1} = \frac{1}{1-2t} \end{aligned}$$

Let $X \sim \text{Expo}(\frac{1}{2})$

$$M_X(t) = \frac{1}{1-2t} \quad \therefore \quad W \sim \text{Expo}(\frac{1}{2})$$

9. a) CDF of $Y = X^3$,

$$P(Y \leq y) = P(X^3 \leq y) = P(X \leq \sqrt[3]{y})$$

$$= 1 - e^{-y^{1/3}}$$

$$P(Y > s+t | Y > s) = \frac{P(Y > s+t)}{P(Y > s)} = \frac{e^{-(s+t)^{1/3}}}{e^{-s^{1/3}}}$$

$$\neq e^{-t^{1/3}} = P(Y > t)$$

b) From $E(X^n) = n!$,

$$E(Y^n) = E(X^{3n}) = (3n)!$$

$$c) E(e^{tY}) = E(e^{tX^3}) = \int_0^\infty e^{tx^3 - x} dx$$

$$\text{for all } t > 0 \quad \int_0^\infty e^{tx^3 - x} dx \rightarrow \infty$$