THE SKEW-NORMAL APPROXIMATION OF THE BINOMIAL DISTRIBUTION

Joyce Tipping

Truman State University

Spring 2011

DEFINITION (BINOMIAL)

Let X represent the number of successes in a sequence of n independent Bernoulli trials with probability of success p on each trial.

DEFINITION (BINOMIAL)

Let X represent the number of successes in a sequence of n independent Bernoulli trials with probability of success p on each trial.

X has discrete pdf

$$f_X(x) = \binom{n}{x} p^x q^{n-x}$$

DEFINITION (BINOMIAL)

Let X represent the number of successes in a sequence of n independent Bernoulli trials with probability of success p on each trial.

X has discrete pdf

$$f_X(x) = \binom{n}{x} p^x q^{n-x}$$

and cdf

$$F_X(x) = P(X \le x) = \sum_{k=0}^{x} f_X(k).$$

INTRODUCTION

The binomial cdf is easy to calculate for small $n\dots$

The binomial cdf is easy to calculate for small $n \dots$

But as *n* gets larger, it becomes increasingly difficult.

INTRODUCTION

The binomial cdf is easy to calculate for small $n \dots$

But as *n* gets larger, it becomes increasingly difficult.

For example ...

INTRODUCTION

When
$$n = 3$$
,

$$F(1) = \begin{pmatrix} 3 \\ 1 \end{pmatrix} p^1 q^2 + \begin{pmatrix} 3 \\ 0 \end{pmatrix} p^0 q^3$$

When n = 3,

$$F(1) = \begin{pmatrix} 3 \\ 1 \end{pmatrix} p^1 q^2 + \begin{pmatrix} 3 \\ 0 \end{pmatrix} p^0 q^3$$

When n = 25,

$$\begin{split} F(12) = & \binom{25}{12} \ \rho^{12} q^{13} + \binom{25}{11} \ \rho^{11} q^{14} + \binom{25}{10} \ \rho^{10} q^{15} + \binom{25}{9} \ \rho^{9} q^{16} \\ & + \binom{25}{8} \ \rho^{8} q^{17} + \binom{25}{7} \ \rho^{7} q^{18} + \binom{25}{6} \ \rho^{6} q^{19} + \binom{25}{5} \ \rho^{5} q^{20} \\ & + \binom{25}{4} \ \rho^{4} q^{21} + \binom{25}{3} \ \rho^{3} q^{22} + \binom{25}{2} \ \rho^{2} q^{23} + \binom{25}{1} \ \rho^{1} q^{24} \\ & + \binom{25}{0} \ \rho^{0} q^{25} \end{split}$$

A common technique is to use the normal distribution as an approximation:

$$F_X(x) \approx \Phi\left(\frac{x + 0.5 - \mu}{\sigma}\right),$$

where $\mu = np$, $\sigma = \sqrt{np(1-p)}$, and Φ is the standard normal cdf.

INTRODUCTION

A common technique is to use the normal distribution as an approximation:

$$F_X(x) \approx \Phi\left(\frac{x + 0.5 - \mu}{\sigma}\right),$$

where $\mu = np$, $\sigma = \sqrt{np(1-p)}$, and Φ is the standard normal cdf.

When does this work well? ...

A common technique is to use the normal distribution as an approximation:

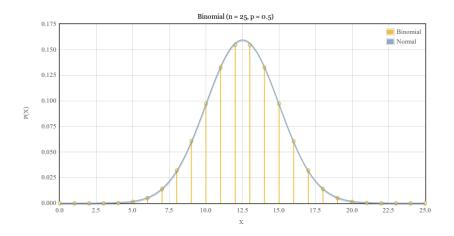
$$F_X(x) \approx \Phi\left(\frac{x + 0.5 - \mu}{\sigma}\right),$$

where $\mu = np$, $\sigma = \sqrt{np(1-p)}$, and Φ is the standard normal cdf.

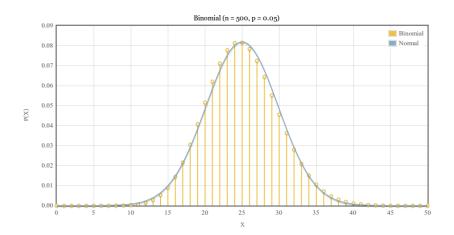
When does this work well? ... In a nutshell, when the binomial is symmetric.

INTRODUCTION

The binomial is symmetric when p = 0.5



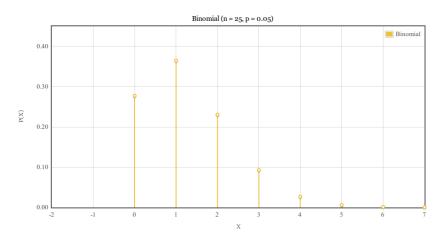
The binomial is symmetric when p = 0.5 or n is very large.



INTRODUCTION

However, when n is medium and p is extreme ...

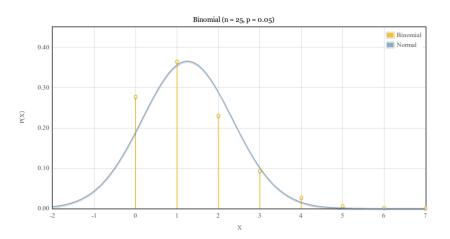
However, when n is medium and p is extreme ...



the binomial is very skewed ...



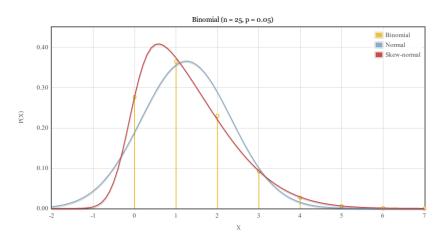
However, when n is medium and p is extreme ...



and the normal approximation doesn't work very well.



However, when n is medium and p is extreme ...



Introducing ... the skew-normal distribution.



OUTLINE

Today's itinerary:

1. Skew-Normal distribution – basic properties

OUTLINE

Today's itinerary:

- 1. Skew-Normal distribution basic properties
- 2. Method of Moments derive an approximation

OUTLINE

Today's itinerary:

- 1. Skew-Normal distribution basic properties
- 2. Method of Moments derive an approximation
- 3. Accuracy examine the accuracy of our approximation

DEFINITION (SKEW-NORMAL)

Let Y be a skew-normal distribution, with location parameter $\mu \in \mathbb{R}$, scale parameter $\sigma > 0$, and shape parameter $\lambda \in \mathbb{R}$. Then Y has pdf

$$f(x|\mu,\sigma,\lambda) = \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot \Phi\left(\frac{\lambda(x-\mu)}{\sigma}\right), \quad x \in \mathbb{R},$$

where ϕ is the standard normal pdf and Φ is the standard normal cdf.

We write $Y \sim SN(\mu, \sigma, \lambda)$.

DEFINITION (SKEW-NORMAL)

Let Y be a skew-normal distribution, with location parameter $\mu \in \mathbb{R}$, scale parameter $\sigma > 0$, and shape parameter $\lambda \in \mathbb{R}$. Then Y has pdf

$$f(x|\mu,\sigma,\lambda) = \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot \Phi\left(\frac{\lambda(x-\mu)}{\sigma}\right), \quad x \in \mathbb{R},$$

where ϕ is the standard normal pdf and Φ is the standard normal cdf.

We write $Y \sim SN(\mu, \sigma, \lambda)$.

Note: μ and σ are not intuitively related to the mean and variance.

So where did this funny-looking pdf come from? ...

So where did this funny-looking pdf come from? ...

LEMMA

If f_0 is a one-dimensional probability density function symmetric about 0, and G is a one-dimensional distribution function such that G' exists and is a density symmetric about 0, then

$$f(z) = 2 \cdot f_0(z) \cdot G\{w(z)\}$$
 $(-\infty < z < \infty)$

is a density function for any odd function $w(\cdot)$. (Lemma 1, Azzalini, 2005)

So where did this funny-looking pdf come from? ...

LEMMA

If f_0 is a one-dimensional probability density function symmetric about 0, and G is a one-dimensional distribution function such that G' exists and is a density symmetric about 0, then

$$f(z) = 2$$
 $f_0(z)$ $G\{w(z)\}$ $(-\infty < z < \infty)$

is a density function for any odd function $w(\cdot)$. (Lemma 1, Azzalini, 2005)

A pdf symmetric about 0 (kernel)

So where did this funny-looking pdf come from? ...

LEMMA

If f_0 is a one-dimensional probability density function symmetric about 0, and G is a one-dimensional distribution function such that G' exists and is a density symmetric about 0, then

$$f(z) = 2 \cdot f_0(z) \cdot G\{w(z)\}$$
 $(-\infty < z < \infty)$

is a density function for any odd function w(.). (Lemma 1, Azzalini, 2005)

- A pdf symmetric about 0 (kernel)
- ► A cdf whose derivative is symmetric about 0

Basic properties:

$$E(Y) = \mu + b\delta\sigma$$

$$E(Y^2) = \mu^2 + 2b\delta\mu\sigma + \sigma^2$$

$$E(Y^3) = \mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3$$

$$Var(Y) = \sigma^2(1 - b^2\delta^2)$$

where
$$b = \sqrt{\frac{2}{\pi}}$$
 and $\delta = \frac{\lambda}{\sqrt{1 + \lambda^2}}$. (Pewsey, 2000)

$$f(x|\mu,\sigma,\lambda=0) = \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot \Phi(0)$$

$$f(x|\mu,\sigma,\lambda=0) = \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot \Phi(0)$$
$$= \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot 0.5$$

$$f(x|\mu,\sigma,\lambda=0) = \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot \Phi(0)$$
$$= \frac{2}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right) \cdot 0.5$$
$$= \frac{1}{\sigma} \cdot \phi\left(\frac{x-\mu}{\sigma}\right)$$

$$\begin{split} f(x|\mu,\sigma,\lambda &= 0) = \frac{2}{\sigma} \cdot \phi \left(\frac{x - \mu}{\sigma} \right) \cdot \Phi(0) \\ &= \frac{2}{\sigma} \cdot \phi \left(\frac{x - \mu}{\sigma} \right) \cdot 0.5 \\ &= \frac{1}{\sigma} \cdot \phi \left(\frac{x - \mu}{\sigma} \right) \\ &= \frac{1}{\sqrt{2\pi}\sigma} \cdot \exp\left(-\frac{(x - \mu)^2}{2\sigma^2} \right), \end{split}$$

What happens when $\lambda = 0$?

$$f(x|\mu,\sigma,\lambda=0) = \frac{2}{\sigma} \cdot \phi \left(\frac{x-\mu}{\sigma}\right) \cdot \Phi(0)$$

$$= \frac{2}{\sigma} \cdot \phi \left(\frac{x-\mu}{\sigma}\right) \cdot 0.5$$

$$= \frac{1}{\sigma} \cdot \phi \left(\frac{x-\mu}{\sigma}\right)$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \cdot \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right),$$

which is the pdf of the normal distribution (μ, σ) .

THE SKEW-NORMAL DISTRIBUTION: THE STANDARD SKEW-NORMAL

DEFINITION (STANDARD SKEW-NORMAL)

The $SN(0,1,\lambda)$ distribution is called the standard skew-normal and has pdf

$$f_{Z}(x|\lambda) = 2 \cdot \phi(x) \cdot \Phi(\lambda x), \quad x \in \mathbb{R}.$$

THE SKEW-NORMAL DISTRIBUTION: THE STANDARD SKEW-NORMAL

DEFINITION (STANDARD SKEW-NORMAL)

The $SN(0,1,\lambda)$ distribution is called the standard skew-normal and has pdf

$$f_Z(x|\lambda) = 2 \cdot \phi(x) \cdot \Phi(\lambda x), \quad x \in \mathbb{R}.$$

Similar to the normal and standard normal, $Z = \frac{Y - \mu}{\sigma}$ and $Y = \sigma Z + \mu$.

PROPERTY (1)

If $Z \sim SN(0, 1, \lambda)$, then $(-Z) \sim SN(0, 1, -\lambda)$.

PROPERTY (1)

If $Z \sim SN(0, 1, \lambda)$, then $(-Z) \sim SN(0, 1, -\lambda)$.

Short proof:

PROPERTY (1)

If
$$Z \sim SN(0, 1, \lambda)$$
, then $(-Z) \sim SN(0, 1, -\lambda)$.

Short proof:

$$f_{(-Z)}(z) = f_Z(-z)$$

Q.E.D.

PROPERTY (1)

If
$$Z \sim SN(0, 1, \lambda)$$
, then $(-Z) \sim SN(0, 1, -\lambda)$.

Short proof:

$$f_{(-Z)}(z) = f_Z(-z)$$

= $2 \cdot \phi(-z) \cdot \Phi(-\lambda z)$

Q.E.D.

PROPERTY (1)

If
$$Z \sim SN(0, 1, \lambda)$$
, then $(-Z) \sim SN(0, 1, -\lambda)$.

Short proof:

$$f_{(-Z)}(z) = f_Z(-z)$$

$$= 2 \cdot \phi(-z) \cdot \Phi(-\lambda z)$$

$$= 2 \cdot \phi(z) \cdot \Phi(-\lambda z),$$

 $Q.\mathcal{E}.\mathcal{D}.$

PROPERTY (1)

If
$$Z \sim SN(0, 1, \lambda)$$
, then $(-Z) \sim SN(0, 1, -\lambda)$.

Short proof:

$$f_{(-Z)}(z) = f_Z(-z)$$

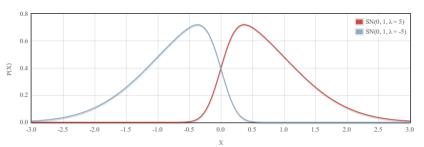
$$= 2 \cdot \phi(-z) \cdot \Phi(-\lambda z)$$

$$= 2 \cdot \phi(z) \cdot \Phi(-\lambda z),$$

which is the pdf of $SN(0, 1, -\lambda)$.

 $Q.\mathcal{E}.\mathcal{D}.$

Property 1: $-SN(0, 1, \lambda) \sim SN(0, 1, -\lambda)$



PROPERTY (2)

If $Z \sim SN(0,1,\lambda)$, then $Z^2 \sim \chi_1^2$ (chi-square with 1 degree of freedom).

PROPERTY (2)

If $Z \sim SN(0,1,\lambda)$, then $Z^2 \sim \chi_1^2$ (chi-square with 1 degree of freedom).

Short proof:

PROPERTY (2)

If $Z \sim SN(0,1,\lambda)$, then $Z^2 \sim \chi_1^2$ (chi-square with 1 degree of freedom).

Short proof:

Lemma 1 comes with a handy result, (Azzalini, 2005, page 161):

If $Y \sim f_0$ and $Z \sim f$, then $|Y| \stackrel{d}{=} |Z|$, where the notation $\stackrel{d}{=}$ denotes equality in distribution.

PROPERTY (2)

If $Z \sim SN(0,1,\lambda)$, then $Z^2 \sim \chi_1^2$ (chi-square with 1 degree of freedom).

Short proof:

Lemma 1 comes with a handy result, (Azzalini, 2005, page 161):

If $Y \sim f_0$ and $Z \sim f$, then $|Y| \stackrel{d}{=} |Z|$, where the notation $\stackrel{d}{=}$ denotes equality in distribution.

Let $X \sim N(0,1)$. Since $X^2 \sim \chi_1^2$ and $|X| \stackrel{d}{=} |Z|$, then $Z^2 \sim \chi_1^2$. $\mathcal{Q}.\mathcal{E}.\mathcal{D}.$

PROPERTY (3)

As $\lambda \to \pm \infty$, $SN(0, 1, \lambda)$ tends to the half normal distribution, $\pm |N(0, 1)|$.

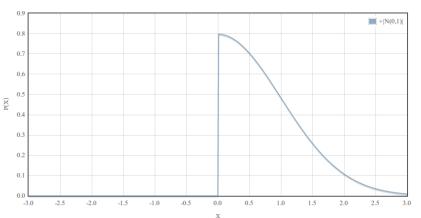
PROPERTY (3)

As $\lambda \to \pm \infty$, $SN(0,1,\lambda)$ tends to the half normal distribution, $\pm |N(0,1)|$.

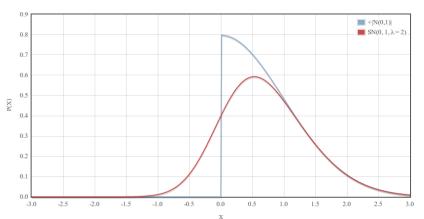
Let $X \sim |N(0,1)|$. Then

$$f_X(x) = egin{cases} 0 & \text{when } -\infty < x \leq 0 \ 2\phi & \text{when } 0 < x < \infty \end{cases}.$$

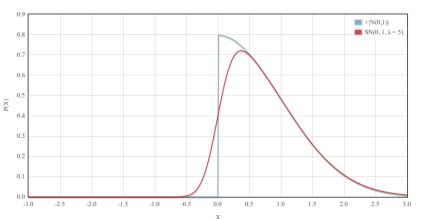
Property 3: $SN(0,1,\lambda) \rightarrow +|N(0,1)|$ as $\lambda \rightarrow \infty$:



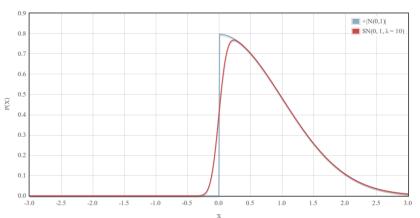
Property 3: $SN(0,1,\lambda) \rightarrow +|N(0,1)|$ as $\lambda \rightarrow \infty$:



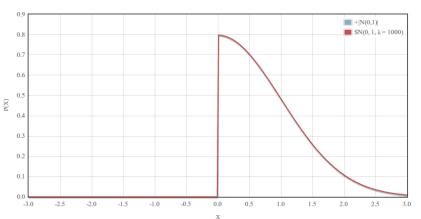
Property 3: $SN(0,1,\lambda) \rightarrow +|N(0,1)|$ as $\lambda \rightarrow \infty$:



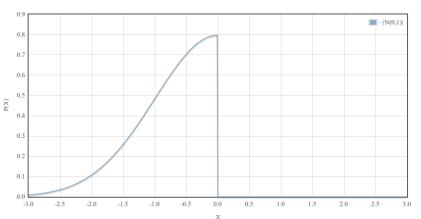
Property 3: $SN(0,1,\lambda) \rightarrow +|N(0,1)|$ as $\lambda \rightarrow \infty$:



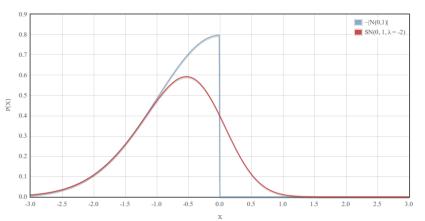
Property 3: $SN(0,1,\lambda) \rightarrow +|N(0,1)|$ as $\lambda \rightarrow \infty$:



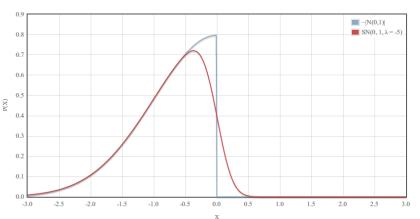
Property 3: $SN(0,1,\lambda) \rightarrow -|N(0,1)|$ as $\lambda \rightarrow -\infty$:



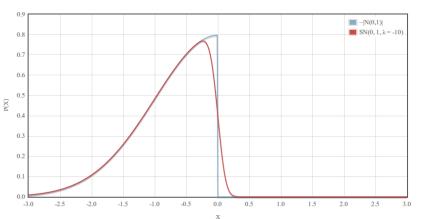
Property 3: $SN(0,1,\lambda) \rightarrow -|N(0,1)|$ as $\lambda \rightarrow -\infty$:



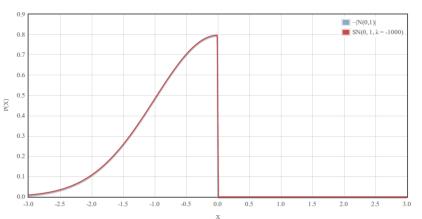
Property 3: $SN(0, 1, \lambda) \rightarrow -|N(0, 1)|$ as $\lambda \rightarrow -\infty$:



Property 3: $SN(0,1,\lambda) \rightarrow -|N(0,1)|$ as $\lambda \rightarrow -\infty$:



Property 3: $SN(0,1,\lambda) \rightarrow -|N(0,1)|$ as $\lambda \rightarrow -\infty$:



PROPERTY (4)

The moment generating function of $SN(0, 1, \lambda)$ is

$$M(t|\lambda) = 2 \cdot \Phi(\delta t) \cdot e^{t^2/2}$$
,

where
$$\delta = \frac{\lambda}{\sqrt{1+\lambda^2}}$$
 and $t \in (-\infty, \infty)$.

PROPERTY (4)

The moment generating function of $SN(0, 1, \lambda)$ is

$$M(t|\lambda) = 2 \cdot \Phi(\delta t) \cdot e^{t^2/2}$$

where
$$\delta = \frac{\lambda}{\sqrt{1+\lambda^2}}$$
 and $t \in (-\infty, \infty)$.

According to Equation 5 in Azzalini (2005), the mgf of $SN(\mu, \sigma, \lambda)$ is

$$M(t) = E\{e^{tY}\} = 2 \cdot \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right) \cdot \Phi(\delta \sigma t).$$

Our result follows.



Now we're going to derive the approximation!

Now we're going to derive the approximation!

Game plan:

Now we're going to derive the approximation!

Game plan:

1. Find the first three moments about the mean (central moments) of the binomial and the skew-normal.

Now we're going to derive the approximation!

Game plan:

1. Find the first three moments about the mean (central moments) of the binomial and the skew-normal.

What are central moments?:
$$E(X)$$
, $E([X - E(X)]^2)$, $E([X - E(X)]^3)$.

Now we're going to derive the approximation!

Game plan:

1. Find the first three moments about the mean (central moments) of the binomial and the skew-normal.

What are central moments?:
$$E(X)$$
, $E([X - E(X)]^2)$, $E([X - E(X)]^3)$.

2. Set them equal to each other.

Now we're going to derive the approximation!

Game plan:

- 1. Find the first three moments about the mean (central moments) of the binomial and the skew-normal.
 - What are central moments?: E(X), $E([X E(X)]^2)$, $E([X E(X)]^3)$.
- 2. Set them equal to each other.
- 3. Take *n* and *p* to be constants; solve for μ , σ , and λ .

Let's start with the binomial ...

Let $B \sim Bin(n, p)$.

Let $B \sim Bin(n, p)$.

The first two central moments of *B* are just the mean and variance:

Let
$$B \sim Bin(n, p)$$
.

The first two central moments of *B* are just the mean and variance:

$$E(B) = np$$

$$E([B - E(B)]^{2}) = Var(B) = np(1 - p)$$

The third one takes some elbow grease.

The third one takes some elbow grease.

First we'll need to find $E(B^2)$ and $E(B^3)$.

Recall that $Var(B) = E(B^2) - [E(B)]^2$.

Recall that
$$Var(B) = E(B^2) - [E(B)]^2$$
. Thus
$$E(B^2) = Var(B) + [E(B)]^2$$

Recall that
$$Var(B)=E(B^2)-[E(B)]^2$$
. Thus
$$E(B^2)=Var(B)+[E(B)]^2$$

$$=np(1-p)+n^2p^2$$

Recall that
$$Var(B) = E(B^2) - [E(B)]^2$$
. Thus
$$E(B^2) = Var(B) + [E(B)]^2$$
$$= np(1-p) + n^2p^2$$
$$= np - np^2 + n^2p^2.$$

We will get $E(B^3)$ via the third factorial moment, E[B(B-1)(B-2)].

$$E[B(B-1)(B-2)]$$

$$E[B(B-1)(B-2)] = \sum_{x=0}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$E[B(B-1)(B-2)] = \sum_{x=0}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$E[B(B-1)(B-2)]$$

$$= \sum_{x=0}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \frac{n!}{x! (n-x)!} p^{x} q^{n-x}$$

$$E[B(B-1)(B-2)]$$

$$= \sum_{x=0}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \frac{n!}{x! (n-x)!} p^{x} q^{n-x}$$

$$= \sum_{x=3}^{n} \frac{n!}{(x-3)! (n-x)!} p^{x} q^{n-x}$$

$$E[B(B-1)(B-2)]$$

$$= \sum_{x=0}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \left\{ \binom{n}{x} p^{x} q^{n-x} \right\}$$

$$= \sum_{x=3}^{n} x(x-1)(x-2) \cdot \frac{n!}{x! (n-x)!} p^{x} q^{n-x}$$

$$= \sum_{x=3}^{n} \frac{n!}{(x-3)! (n-x)!} p^{x} q^{n-x}$$

$$= \sum_{x=3}^{n} n(n-1)(n-2) p^{3} \cdot \frac{(n-3)!}{(x-3)! (n-x)!} p^{x-3} q^{n-x}$$

$$= \sum_{k=3}^{n} n(n-1)(n-2)p^{3} \cdot \frac{(n-3)!}{(x-3)!(n-x)!} p^{x-3}q^{n-x}$$

$$= \sum_{x=3}^{n} n(n-1)(n-2)p^{3} \cdot \frac{(n-3)!}{(x-3)!(n-x)!} p^{x-3}q^{n-x}$$

Let y = x - 3; then x = y + 3, and x = 3, $x = n \Rightarrow y = 0$, y = n - 3:

$$= \sum_{x=3}^{n} n(n-1)(n-2)p^{3} \cdot \frac{(n-3)!}{(x-3)!(n-x)!} p^{x-3}q^{n-x}$$

Let y = x - 3; then x = y + 3, and x = 3, $x = n \implies y = 0$, y = n - 3:

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! (n-(y+3))!} p^{y} q^{n-(y+3)}$$

$$= \sum_{x=3}^{n} n(n-1)(n-2)p^{3} \cdot \frac{(n-3)!}{(x-3)! (n-x)!} p^{x-3}q^{n-x}$$

Let y = x - 3; then x = y + 3, and x = 3, $x = n \implies y = 0$, y = n - 3:

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! (n-(y+3))!} p^{y} q^{n-(y+3)}$$

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! ((n-3)-y)!} p^{y} q^{(n-3)-y}$$

[pdf of Bin(n-3, p) summed over its domain] = 1

$$= \sum_{x=3}^{n} n(n-1)(n-2)p^{3} \cdot \frac{(n-3)!}{(x-3)! (n-x)!} p^{x-3}q^{n-x}$$
Let $y = x - 3$; then $x = y + 3$, and $x = 3$, $x = n \Rightarrow y = 0$, $y = n - 3$:
$$= n(n-1)(n-2)p^{3} \cdot \sum_{x=0}^{n-3} \frac{(n-3)!}{(n-3)!} p^{y}q^{n-(y+3)}$$

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! (n-(y+3))!} p^{y} q^{n-(y+3)}$$

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! ((n-3)-y)!} p^{y} q^{(n-3)-y}$$
[pdf of $Bin(n-3,p)$ summed over its domain] = 1

$$= n(n-1)(n-2)p^3$$

$$= \sum_{x=3}^{n} n(n-1)(n-2)p^{3} \cdot \frac{(n-3)!}{(x-3)! (n-x)!} p^{x-3}q^{n-x}$$

Let y = x - 3; then x = y + 3, and x = 3, $x = n \Rightarrow y = 0$, y = n - 3:

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! (n-(y+3))!} p^{y} q^{n-(y+3)}$$

$$= n(n-1)(n-2)p^{3} \cdot \sum_{y=0}^{n-3} \frac{(n-3)!}{y! ((n-3)-y)!} p^{y} q^{(n-3)-y}$$
[pdf of $Bin(n-3,p)$ summed over its domain] = 1

$$= n(n-1)(n-2)p^3$$

= $n^3p^3 - 3n^2p^3 + 2np^3$

$$\boldsymbol{E[B(B-1)(B-2)]}$$

$$E[B(B-1)(B-2)] = E[B^3 - 3B^2 + 2B]$$

$$E[B(B-1)(B-2)]$$
= $E[B^3 - 3B^2 + 2B]$
= $E(B^3) - 3E(B^2) + 2E(B)$

$$\begin{split} &E[B(B-1)(B-2)]\\ &= E\left[B^3 - 3B^2 + 2B\right]\\ &= E(B^3) - 3E(B^2) + 2E(B)\\ &= E(B^3) - 3(np - np^2 + n^2p^2) + 2np \end{split}$$

$$E[B(B-1)(B-2)]$$

$$= E[B^3 - 3B^2 + 2B]$$

$$= E(B^3) - 3E(B^2) + 2E(B)$$

$$= E(B^3) - 3(np - np^2 + n^2p^2) + 2np$$

$$= E(B^3) - 3np + 3np^2 - 3n^2p^2 + 2np$$

$$E[B(B-1)(B-2)]$$

$$= E[B^3 - 3B^2 + 2B]$$

$$= E(B^3) - 3E(B^2) + 2E(B)$$

$$= E(B^3) - 3(np - np^2 + n^2p^2) + 2np$$

$$= E(B^3) - 3np + 3np^2 - 3n^2p^2 + 2np$$

$$= E(B^3) + 3np^2 - 3n^2p^2 - np$$

Left side: $E(B^3) + 3np^2 - 3n^2p^2 - np$

Right side: $n^3p^3 - 3n^2p^3 + 2np^3$

Left side: $E(B^3) + 3np^2 - 3n^2p^2 - np$

Right side: $n^3p^3 - 3n^2p^3 + 2np^3$

Set them equal and solve for $E(B^3)$:

Left side:
$$E(B^3) + 3np^2 - 3n^2p^2 - np$$

Right side: $n^3p^3 - 3n^2p^3 + 2np^3$

Set them equal and solve for $E(B^3)$:

$$E(B^3) + 3np^2 - 3n^2p^2 - np = n^3p^3 - 3n^2p^3 + 2np^3$$

Left side:
$$E(B^3) + 3np^2 - 3n^2p^2 - np$$

Right side: $n^3p^3 - 3n^2p^3 + 2np^3$

Set them equal and solve for $E(B^3)$:

$$\begin{split} E(B^3) + 3np^2 - 3n^2p^2 - np &= n^3p^3 - 3n^2p^3 + 2np^3 \\ \Rightarrow & E(B^3) = n^3p^3 - 3n^2p^3 + 2np^3 - 3np^2 + 3n^2p^2 + np \end{split}$$

$$E\left([B-E(B)]^3\right)$$

$$E([B - E(B)]^3)$$
= $E(B^3 - 3B^2E(B) + 3B[E(B)]^2 - [E(B)]^3)$

$$E([B - E(B)]^3)$$

$$= E(B^3 - 3B^2E(B) + 3B[E(B)]^2 - [E(B)]^3)$$

$$= E(B^3) - 3E(B^2)E(B) + 3E(B)[E(B)]^2 - [E(B)]^3$$

$$E([B - E(B)]^{3})$$

$$= E(B^{3} - 3B^{2}E(B) + 3B[E(B)]^{2} - [E(B)]^{3})$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 3E(B)[E(B)]^{2} - [E(B)]^{3}$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 2[E(B)]^{3}$$

$$\begin{split} &E\left([B-E(B)]^3\right) \\ &= E\left(B^3 - 3B^2E(B) + 3B[E(B)]^2 - [E(B)]^3\right) \\ &= E(B^3) - 3E(B^2)E(B) + 3E(B)[E(B)]^2 - [E(B)]^3 \\ &= E(B^3) - 3E(B^2)E(B) + 2[E(B)]^3 \\ &= (n^3p^3 - 3n^2p^3 + 2np^3 - 3np^2 + 3n^2p^2 + np) \\ &- 3(np - np^2 + n^2p^2)(np) + 2(np)^3 \end{split}$$

$$E([B-E(B)]^{3})$$

$$= E(B^{3} - 3B^{2}E(B) + 3B[E(B)]^{2} - [E(B)]^{3})$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 3E(B)[E(B)]^{2} - [E(B)]^{3}$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 2[E(B)]^{3}$$

$$= (n^{3}p^{3} - 3n^{2}p^{3} + 2np^{3} - 3np^{2} + 3n^{2}p^{2} + np)$$

$$- 3(np - np^{2} + n^{2}p^{2})(np) + 2(np)^{3}$$

$$= p^{3}p^{3} - 3n^{2}p^{3} + 2np^{3} - 3np^{2} + 3n^{2}p^{2} + np$$

$$- 3n^{2}p^{2} + 3n^{2}p^{3} - 3n^{3}p^{3} + 2n^{3}p^{3}$$

Now we can (finally!) compute the third central moment:

$$E([B-E(B)]^{3})$$

$$= E(B^{3} - 3B^{2}E(B) + 3B[E(B)]^{2} - [E(B)]^{3})$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 3E(B)[E(B)]^{2} - [E(B)]^{3}$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 2[E(B)]^{3}$$

$$= (n^{3}p^{3} - 3n^{2}p^{3} + 2np^{3} - 3np^{2} + 3n^{2}p^{2} + np)$$

$$- 3(np - np^{2} + n^{2}p^{2})(np) + 2(np)^{3}$$

$$= p^{3}p^{3} - 3n^{2}p^{3} + 2np^{3} - 3np^{2} + 3n^{2}p^{2} + np$$

$$- 3n^{2}p^{2} + 3n^{2}p^{3} - 3n^{3}p^{3} + 2n^{3}p^{3}$$

$$= 2np^{3} - 3np^{2} + np$$

Now we can (finally!) compute the third central moment:

$$E([B-E(B)]^{3})$$

$$= E(B^{3} - 3B^{2}E(B) + 3B[E(B)]^{2} - [E(B)]^{3})$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 3E(B)[E(B)]^{2} - [E(B)]^{3}$$

$$= E(B^{3}) - 3E(B^{2})E(B) + 2[E(B)]^{3}$$

$$= (n^{3}p^{3} - 3n^{2}p^{3} + 2np^{3} - 3np^{2} + 3n^{2}p^{2} + np)$$

$$- 3(np - np^{2} + n^{2}p^{2})(np) + 2(np)^{3}$$

$$= n^{3}p^{3} - 3n^{2}p^{3} + 2np^{3} - 3np^{2} + 3n^{2}p^{2} + np$$

$$- 3n^{2}p^{2} + 3n^{2}p^{3} - 3n^{3}p^{3} + 2n^{3}p^{3}$$

$$= 2np^{3} - 3np^{2} + np$$

$$= np(p-1)(2p-1)$$

Let's restate our results:

$$E(B) = np,$$

$$E([B - E(B)]^2) = np(1 - p),$$

$$E([B - E(B)]^3) = np(p - 1)(2p - 1)$$

Now lets move on to skew-normal ...

Let $Y \sim SN(\mu, \sigma, \lambda)$.

Let $Y \sim SN(\mu, \sigma, \lambda)$.

Again, the first and second central moments of *Y* are the mean and variance.

Let
$$Y \sim SN(\mu, \sigma, \lambda)$$
.

Again, the first and second central moments of *Y* are the mean and variance.

$$E(Y) = \mu + b\delta\sigma$$
$$Var(Y) = \sigma^{2}(1 - b^{2}\delta^{2})$$

$$E([Y - E(Y)]^3)$$

$$E([Y - E(Y)]^3)$$

= $E(Y^3) - 3E(Y^2)E(Y) + 2[E(Y)]^3$

$$E([Y - E(Y)]^{3})$$
= $E(Y^{3}) - 3E(Y^{2})E(Y) + 2[E(Y)]^{3}$
= $(\mu^{3} + 3b\delta\mu^{2}\sigma + 3\mu\sigma^{2} + 3b\delta\sigma^{3} - b\delta^{3}\sigma^{3})$
 $-3(\mu^{2} + 2b\delta\mu\sigma + \sigma^{2})(\mu + b\delta\sigma) + 2(\mu + b\delta\sigma)^{3}$

$$\begin{split} &E([Y-E(Y)]^3) \\ &= E(Y^3) - 3E(Y^2)E(Y) + 2[E(Y)]^3 \\ &= (\mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3) \\ &- 3(\mu^2 + 2b\delta\mu\sigma + \sigma^2)(\mu + b\delta\sigma) + 2(\mu + b\delta\sigma)^3 \\ &= \mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3 - 3\mu^3 - 3b\delta\mu^2\sigma \\ &- 6b\delta\mu^2\sigma - 6b^2\delta^2\mu\sigma^2 - 3\mu\sigma^2 - 3b\delta\sigma^3 + 2\mu^3 + 6b\delta\mu^2\sigma \\ &+ 6b^2\delta^2\mu\sigma^2 + 2b^3\delta^3\sigma^3 \end{split}$$

$$\begin{split} &E([Y-E(Y)]^3) \\ &= E(Y^3) - 3E(Y^2)E(Y) + 2[E(Y)]^3 \\ &= (\mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3) \\ &- 3(\mu^2 + 2b\delta\mu\sigma + \sigma^2)(\mu + b\delta\sigma) + 2(\mu + b\delta\sigma)^3 \\ &= \mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3 - 3\mu^3 - 3b\delta\mu^2\sigma \\ &- 6b\delta\mu^2\sigma - 6b^2\delta^2\mu\sigma^2 - 3\mu\sigma^2 - 3b\delta\sigma^3 + 2\mu^3 + 6b\delta\mu^2\sigma \\ &+ 6b^2\delta^2\mu\sigma^2 + 2b^3\delta^3\sigma^3 \\ &= 2b^3\delta^3\sigma^3 - b\delta^3\sigma^3 \end{split}$$

$$\begin{split} &E([Y-E(Y)]^3) \\ &= E(Y^3) - 3E(Y^2)E(Y) + 2[E(Y)]^3 \\ &= (\mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3) \\ &- 3(\mu^2 + 2b\delta\mu\sigma + \sigma^2)(\mu + b\delta\sigma) + 2(\mu + b\delta\sigma)^3 \\ &= \mu^3 + 3b\delta\mu^2\sigma + 3\mu\sigma^2 + 3b\delta\sigma^3 - b\delta^3\sigma^3 - 3\mu^3 - 3b\delta\mu^2\sigma \\ &- 6b\delta\mu^2\sigma - 6b^2\delta^2\mu\sigma^2 - 3\mu\sigma^2 - 3b\delta\sigma^3 + 2\mu^3 + 6b\delta\mu^2\sigma \\ &+ 6b^2\delta^2\mu\sigma^2 + 2b^3\delta^3\sigma^3 \\ &= 2b^3\delta^3\sigma^3 - b\delta^3\sigma^3 \\ &= b\delta^3\sigma^3(2b^2 - 1) \end{split}$$

Our results, restated:

$$E(Y) = \mu + b\delta\sigma \qquad = \mu + \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}}$$

$$E([Y - E(Y)]^2) = \sigma^2 (1 - b^2 \delta^2) \qquad = \sigma^2 \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1 + \lambda^2}\right)$$

$$E([Y - E(Y)]^3) = b\delta^3 \sigma^3 (2b^2 - 1) = \sigma^3 \sqrt{\frac{2}{\pi}} \left(\frac{\lambda}{\sqrt{1 + \lambda^2}}\right)^3 \left(\frac{4}{\pi} - 1\right)$$

We're finally ready to derive our approximation!

$$np = \mu + \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}}$$
 (1a)

$$np = \mu + \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}}$$
 (1a)

$$np(1-p) = \sigma^2 \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1 + \lambda^2} \right) \tag{1b}$$

$$np = \mu + \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}}$$
 (1a)

$$np(1-p) = \sigma^2 \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1 + \lambda^2} \right)$$
 (1b)

$$np(p-1)(2p-1) = \sigma^3 \sqrt{\frac{2}{\pi}} \left(\frac{\lambda}{\sqrt{1+\lambda^2}}\right)^3 \left(\frac{4}{\pi} - 1\right)$$
 (1c)

To get λ , divide the cube of (1b) by the square of (1c):

To get λ , divide the cube of (1b) by the square of (1c):

$$\frac{\sigma^6 \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1 + \lambda^2}\right)^3}{\sigma^6 \cdot \frac{2}{\pi} \left(\frac{\lambda}{\sqrt{1 + \lambda^2}}\right)^6 \left(\frac{4}{\pi} - 1\right)^2} = \frac{n^3 p^3 (1 - p)^3}{n^2 p^2 (p - 1)^2 (2p - 1)^2}$$

To get λ , divide the cube of (1b) by the square of (1c):

$$\frac{\sigma^{6} \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^{2}}{1 + \lambda^{2}}\right)^{3}}{\sigma^{6} \cdot \frac{2}{\pi} \left(\frac{\lambda}{\sqrt{1 + \lambda^{2}}}\right)^{6} \left(\frac{4}{\pi} - 1\right)^{2}} = \frac{n^{3} p^{3} (1 - p)^{3}}{n^{2} p^{2} (p - 1)^{2} (2p - 1)^{2}}$$

$$\Rightarrow \frac{\left(1 - \frac{2}{\pi} \cdot \frac{\lambda^{2}}{1 + \lambda^{2}}\right)^{3}}{\frac{2}{\pi} \left(\frac{\lambda^{2}}{1 + \lambda^{2}}\right)^{3} \left(\frac{4}{\pi} - 1\right)^{2}} = \frac{n p (1 - p)}{(1 - 2p)^{2}} \tag{2}$$

Equation (2) can be solved for λ^2 .

Equation (2) can be solved for λ^2 .

Then take λ to be

$$\lambda = \{ \text{sign of } (1 - 2p) \} \sqrt{\lambda^2}.$$

Equation (2) can be solved for λ^2 .

Then take λ to be

$$\lambda = \{ \text{sign of } (1 - 2p) \} \sqrt{\lambda^2}.$$

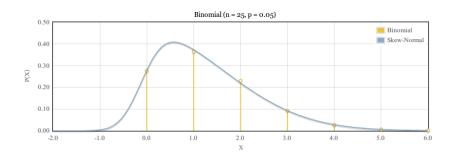
Why?

Equation (2) can be solved for λ^2 .

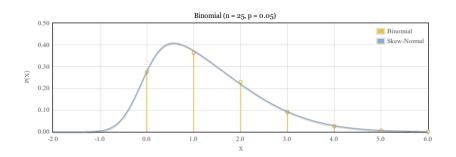
Then take λ to be

$$\lambda = \{ \text{sign of } (1 - 2p) \} \sqrt{\lambda^2}.$$

Why? Recall Property 3 ...



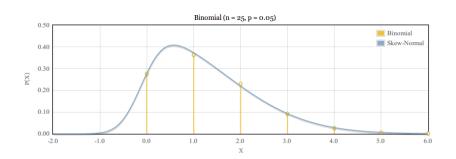
When p < 0.5:



When p < 0.5:

► The binomial skews right (weight shifts left) and approaches $+|N(0,1)| \longrightarrow \lambda$ is positive

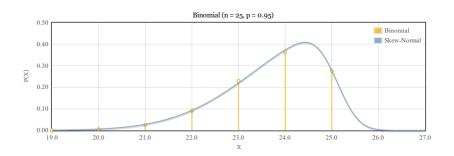




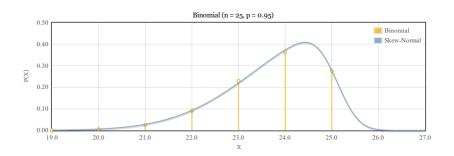
When *p*< 0.5:

- ► The binomial skews right (weight shifts left) and approaches $+|N(0,1)| \longrightarrow \lambda$ is positive
- ▶ (1-2p) is positive





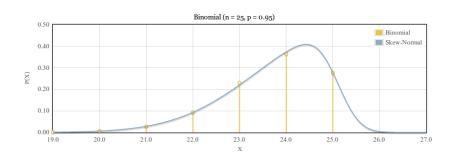
When p > 0.5:



When p > 0.5:

► The binomial skews left (weight shifts right) and approaches $-|N(0,1)| \longrightarrow \lambda$ is negative





When p > 0.5:

- ▶ The binomial skews left (weight shifts right) and approaches $-|N(0,1)| \longrightarrow \lambda$ is negative
- ▶ (1-2p) is negative



Once you have λ , solve for σ and then μ .

Once you have λ , solve for σ and then μ .

$$np(1-p) = \sigma^2 \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1 + \lambda^2}\right) \quad \Rightarrow \quad \sigma = \sqrt{\frac{np(1-p)}{1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1 + \lambda^2}}}$$

Once you have λ , solve for σ and then μ .

$$np(1-p) = \sigma^2 \left(1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1+\lambda^2} \right) \quad \Rightarrow \quad \sigma = \sqrt{\frac{np(1-p)}{1 - \frac{2}{\pi} \cdot \frac{\lambda^2}{1+\lambda^2}}}$$

$$np = \mu + \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}} \quad \Rightarrow \quad \mu = np - \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}}$$

When p = 0.5, the right hand side of (2)

$$\frac{\left(1-\frac{2}{\pi}\cdot\frac{\lambda^2}{1+\lambda^2}\right)^3}{\frac{2}{\pi}\left(\frac{\lambda^2}{1+\lambda^2}\right)^3\left(\frac{4}{\pi}-1\right)^2}=\frac{np(1-p)}{(1-2p)^2}$$

is undefined.

When p = 0.5, the right hand side of (2)

$$\frac{\left(1-\frac{2}{\pi}\cdot\frac{\lambda^2}{1+\lambda^2}\right)^3}{\frac{2}{\pi}\left(\frac{\lambda^2}{1+\lambda^2}\right)^3\left(\frac{4}{\pi}-1\right)^2}=\frac{np(1-p)}{(1-2p)^2}$$

is undefined.

Uh-oh.

Fortunately, our intuition rescues us:

Fortunately, our intuition rescues us:

When p = 0.5, the binomial is symmetric, so λ should equal 0.

Fortunately, our intuition rescues us:

When p = 0.5, the binomial is symmetric, so λ should equal 0.

This takes us back to the usual normal approximation:

Fortunately, our intuition rescues us:

When p = 0.5, the binomial is symmetric, so λ should equal 0.

This takes us back to the usual normal approximation:

$$\sigma = \sqrt{\frac{np(1-p)}{1 - \frac{2}{\pi} \cdot \frac{0^2}{1 + 0^2}}} = \sqrt{\frac{np(1-p)}{1}} = \sqrt{np(1-p)}$$

Fortunately, our intuition rescues us:

When p = 0.5, the binomial is symmetric, so λ should equal 0.

This takes us back to the usual normal approximation:

$$\sigma = \sqrt{\frac{np(1-p)}{1 - \frac{2}{\pi} \cdot \frac{0^2}{1 + 0^2}}} = \sqrt{\frac{np(1-p)}{1}} = \sqrt{np(1-p)}$$

$$\mu = np - \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{0}{\sqrt{1 + 0^2}} = np - 0 = np$$

Unfortunately, though better than the normal approximation, our skew-normal method isn't universal.

To be able to solve for λ , we must restrict n and p by the following equation:

To be able to solve for λ , we must restrict n and p by the following equation:

$$np(1-p) \ge (1-2p)^2$$
 (3)

To be able to solve for λ , we must restrict n and p by the following equation:

$$np(1-p) \ge (1-2p)^2$$
 (3)

From (3), we can answer two questions:

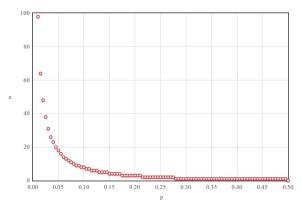
One: Given *p*, what is the least *n* necessary?

One: Given *p*, what is the least *n* necessary?

$$n \geq \frac{(1-2p)^2}{p(1-p)}$$

One: Given *p*, what is the least *n* necessary?

$$n \geq \frac{(1-2p)^2}{p(1-p)}$$



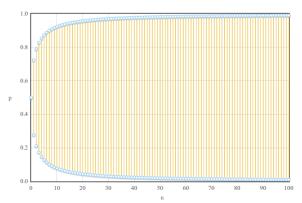
Two: Given n, what is the range of possible p's?

Two: Given n, what is the range of possible p's?

$$\frac{1}{2} - \frac{1}{2}\sqrt{\frac{n}{n+4}} \leq p \leq \frac{1}{2} + \frac{1}{2}\sqrt{\frac{n}{n+4}}$$

Two: Given n, what is the range of possible p's?

$$\frac{1}{2} - \frac{1}{2} \sqrt{\frac{n}{n+4}} \ \leq \ p \ \leq \ \frac{1}{2} + \frac{1}{2} \sqrt{\frac{n}{n+4}}$$



We have an approximation!!

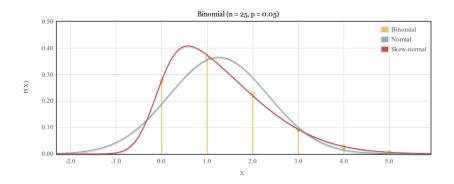
We have an approximation!!

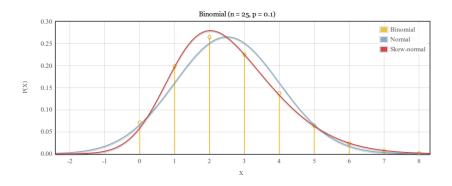
But is it more accurate? ...

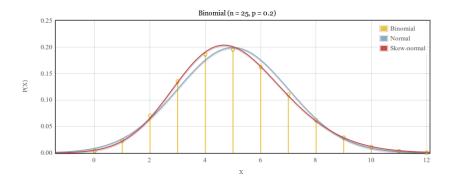
We have an approximation!!

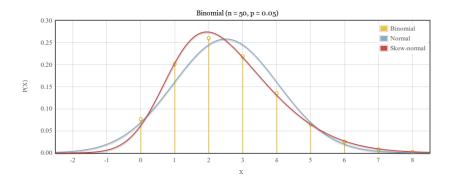
But is it more accurate? ... Answer: Yes!

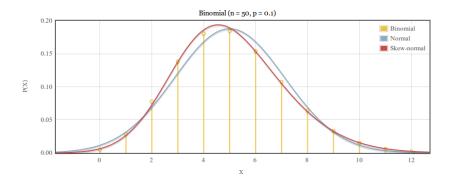
The easiest way of gauging accuracy is by visual inspection.

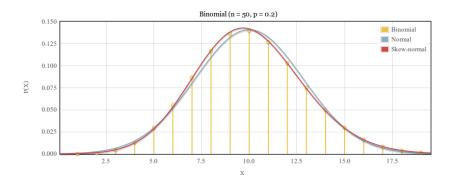


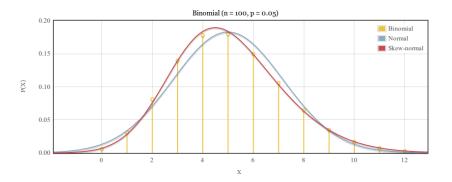


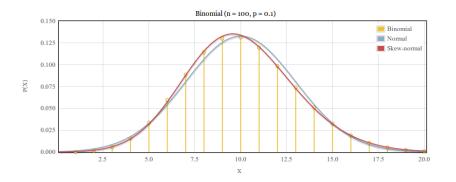


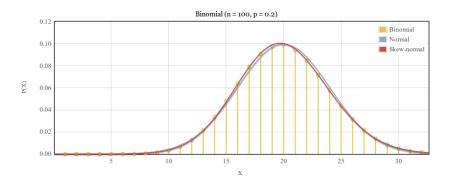












A more numerical way of gauging accuracy is the maximal absolute error (*MABS*), defined as

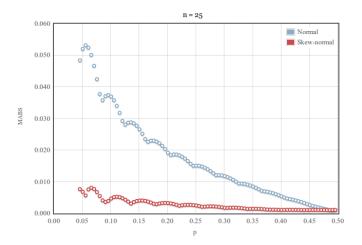
$$\mathsf{MABS}(\textit{n}, \textit{p}) \ = \ \max_{\textit{k} \in \{0, 1, \dots, n\}} \left| \textit{F}_{\textit{B}(\textit{n}, \textit{p})}(\textit{k}) - \textit{F}_{\mathsf{appr}(\textit{n}, \textit{p})}(\textit{k} + 0.5) \right|.$$

A more numerical way of gauging accuracy is the maximal absolute error (*MABS*), defined as

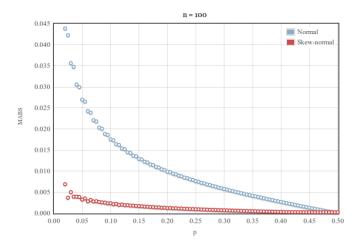
$$\mathsf{MABS}(n,p) \ = \ \max_{k \in \{0,1,\dots,n\}} \left| F_{B(n,p)}(k) - F_{\mathsf{appr}(n,p)}(k+0.5) \right|.$$

(In English: The biggest vertical distance between the binomial cdf and the approximation curve.)

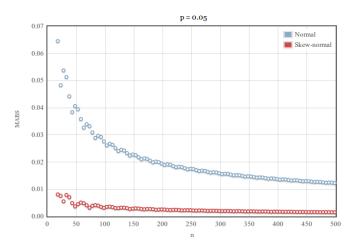
MABS as a function of p, with fixed n = 25:



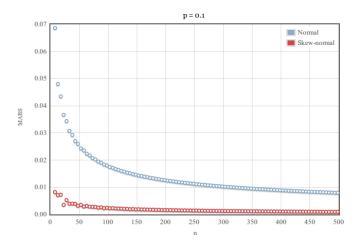
MABS as a function of p, with fixed n = 100:



MABS as a function of n, with fixed p = 0.05:



MABS as a function of n, with fixed p = 0.1:



RESOURCES

A few resources you might find helpful \dots

RESOURCES

Estimations of $SN(\mu, \sigma, \lambda)$ for Bin(n, p)

				n		
		25	50	100	250	500
	0.05	(-0.11, 1.74, 4.56)	(0.79, 2.30, 2.54)	(2.85, 3.06, 1.86)	(9.58, 4.52, 1.38)	(21.32, 6.11, 1.15)
	0.10	(0.89, 2.20, 2.31)	(2.97, 2.94, 1.74)	(7.44, 3.94, 1.40)	(21.53, 5.88, 1.10)	(45.62, 8.01, 0.94)
	0.15	(2.02, 2.49, 1.79)	(5.32, 3.34, 1.43)	(12.25, 4.51, 1.19)	(33.77, 6.77, 0.96)	(70.30, 9.27, 0.82)
	0.20	(3.23, 2.67, 1.50)	(7.76, 3.61, 1.24)	(17.18, 4.89, 1.04)	(46.18, 7.39, 0.85)	(95.18, 10.16, 0.74)
	0.25	(4.49, 2.79, 1.29)	(10.28, 3.78, 1.09)	(22.20, 5.15, 0.93)	(58.71, 7.83, 0.76)	(120.22, 10.80, 0.67)
	0.30	(5.80, 2.85, 1.12)	(12.86, 3.88, 0.95)	(27.31, 5.32, 0.82)	(71.34, 8.12, 0.68)	(145.39, 11.24, 0.60)
	0.35	(7.17, 2.86, 0.96)	(15.50, 3.92, 0.83)	(32.49, 5.39, 0.72)	(84.09, 8.28, 0.60)	(170.70, 11.50, 0.53)
	0.40	(8.59, 2.83, 0.80)	(18.23, 3.89, 0.70)	(37.76, 5.39, 0.61)	(96.96, 8.32, 0.51)	(196.18, 11.60, 0.45)
р	0.45	(10.12, 2.73, 0.61)	(21.08, 3.79, 0.53)	(43.21, 5.29, 0.47)	(110.07, 8.23, 0.40)	(221.93, 11.54, 0.35)
Ρ	0.50	(12.50, 2.50, 0.00)	(25.00, 3.54, 0.00)	(50.00, 5.00, 0.00)	(125.00, 7.91, 0.00)	(250.00, 11.18, 0.00)
	0.55	(14.88, 2.73, -0.61)	(28.92, 3.79, -0.53)	(56.79, 5.29, -0.47)	(139.93, 8.23, -0.40)	(278.07, 11.54, -0.35)
	0.60	(16.41, 2.83, -0.80)	(31.77, 3.89, -0.70)	(62.24, 5.39, -0.61)	(153.04, 8.32, -0.51)	(303.82, 11.60, -0.45)
	0.65	(17.83, 2.86, -0.96)	(34.50, 3.92, -0.83)	(67.51, 5.39, -0.72)	(165.91, 8.28, -0.60)	(329.30, 11.50, -0.53)
	0.70	(19.20, 2.85, -1.12)	(37.14, 3.88, -0.95)	(72.69, 5.32, -0.82)	(178.66, 8.12, -0.68)	(354.61, 11.24, -0.60)
	0.75	(20.51, 2.79, -1.29)	(39.72, 3.78, -1.09)	(77.80, 5.15, -0.93)	(191.29, 7.83, -0.76)	(379.78, 10.80, -0.67)
	0.80	(21.77, 2.67, -1.50)	(42.24, 3.61, -1.24)	(82.82, 4.89, -1.04)	(203.82, 7.39, -0.85)	(404.82, 10.16, -0.74)
	0.85	(22.98, 2.49, -1.79)	(44.68, 3.34, -1.43)	(87.75, 4.51, -1.19)	(216.23, 6.77, -0.96)	(429.70, 9.27, -0.82)
	0.90	(24.11, 2.20, -2.31)	(47.03, 2.94, -1.74)	(92.56, 3.94, -1.40)	(228.47, 5.88, -1.10)	(454.38, 8.01, -0.94)
	0.95	(25.11, 1.74, -4.56)	(49.21, 2.30, -2.54)	(97.15, 3.06, -1.86)	(240.42, 4.52, -1.38)	(478.68, 6.11, -1.15)

RESOURCES

All values in my project were computed using a Python library, which is freely available online:

http://github.com/joycetipping/skew-normal-capstone/

BIBLIOGRAPHY I

- B. C. Arnold and G. D. Lin. Characterizations of the skew-normal and generalized chi distributions. *Sankhy: The Indian Journal of Statistics*, 66(4):593–606, 2004.
- Adelchi Azzalini. A class of distributions which includes the normal ones. *Scandinavian Journal of Statistics*, 12: 171–178, 1985.
- Adelchi Azzalini. The skew-normal distribution and related multivariate families. *Scandinavian Journal of Statistics*, 32: 159–188, 2005.
- Lee J. Bain and Max Engelhardt. *Introduction to Probability and Mathematical Statistics*. Duxbury, 2nd edition, 1992.
- Ching-Hui Chang, Jyh-Jiuan Lin, Nabendu Pal, and Miao-Chen Chiang. A note on improved approximation of the binomial distribution by the skew-normal distribution. *American Statistical Association*, 62(2):167–170, May 2008.



BIBLIOGRAPHY II

- A. K. Gupta, T. T. Nguyen, and J. A. T. Sanqui. Characterization of the skew-normal distribution. *Annals of the Institute of Statistical Mathematics*, 56:351–360, 2004.
- A. O'Hagan and Tom Leonard. Bayes estimation subject to uncertainty about parameter constraints. *Biometrika*, 63: 201–203, 1976.
- Arthur Pewsey. Problems of inference for azzalini's skew-normal distribution. *Journal of Applied Statistics*, 27(7): 859–870, 2000.
- Martin Schader and Friedrich Schmid. Two rules of thumb for the approximation of the binomial distribution by the normal distribution. *American Statistical Association*, 43(1):23–24, February 1989.

Appendix

Calculating a skew-normal approximation

CALCULATING A SKEW-NORMAL APPROXIMATION

Although easier with a computer program, calculating estimates for μ , σ , and λ by hand is perfectly possible.

CALCULATING A SKEW-NORMAL APPROXIMATION

Although easier with a computer program, calculating estimates for μ , σ , and λ by hand is perfectly possible.

Here's an example!

By far the biggest battle is λ .

By far the biggest battle is λ .

By far the biggest battle is λ .

$$\left(\frac{1+\lambda^2}{\lambda^2} - \frac{2}{\pi}\right)^3 \left(\frac{\pi^3}{2(4-\pi)^2}\right) = \frac{np(1-p)}{(1-2p)^2} \tag{4}$$

By far the biggest battle is λ .

$$\underbrace{\left(\frac{1+\lambda^{2}}{\lambda^{2}}-\frac{2}{\pi}\right)^{3}\left(\frac{\pi^{3}}{2(4-\pi)^{2}}\right)}_{f(\lambda)} = \frac{np(1-p)}{(1-2p)^{2}} \tag{4}$$

By far the biggest battle is λ .

$$\underbrace{\left(\frac{1+\lambda^2}{\lambda^2} - \frac{2}{\pi}\right)^3 \left(\frac{\pi^3}{2(4-\pi)^2}\right)}_{f(\lambda)} = \underbrace{\frac{np(1-p)}{(1-2p)^2}}_{k_{n,p}} \tag{4}$$

CALCULATING A SKEW-NORMAL APPROXIMATION: FINDING λ

By far the biggest battle is λ .

We'll use a simplified version of equation (2):

$$\underbrace{\left(\frac{1+\lambda^2}{\lambda^2} - \frac{2}{\pi}\right)^3 \left(\frac{\pi^3}{2(4-\pi)^2}\right)}_{f(\lambda)} = \underbrace{\frac{np(1-p)}{(1-2p)^2}}_{k_{n,p}} \tag{4}$$

The closed form solution to (4) is pretty hideous, so we'll take a numerical approach.

CALCULATING A SKEW-NORMAL APPROXIMATION: FINDING λ

By far the biggest battle is λ .

We'll use a simplified version of equation (2):

$$\underbrace{\left(\frac{1+\lambda^{2}}{\lambda^{2}}-\frac{2}{\pi}\right)^{3}\left(\frac{\pi^{3}}{2(4-\pi)^{2}}\right)}_{f(\lambda)}=\underbrace{\frac{np(1-p)}{(1-2p)^{2}}}_{k_{n,p}}$$
(4)

The closed form solution to (4) is pretty hideous, so we'll take a numerical approach.

Our goal is to find λ such that $f(\lambda)$ is within a certain margin of error (*e*) of $k_{n,p}$.



Recall that the sign of λ is determined independently of the value. Thus, we can consider only $\lambda>0$.

Recall that the sign of λ is determined independently of the value. Thus, we can consider only $\lambda > 0$.

It is possible to show, by taking its derivative, that f is monotonically decreasing for positive λ .

CALCULATING A SKEW-NORMAL APPROXIMATION: FINDING λ

Recall that the sign of λ is determined independently of the value. Thus, we can consider only $\lambda > 0$.

It is possible to show, by taking its derivative, that f is monotonically decreasing for positive λ .

This convenient fact allows us to find lower and upper bounds for λ and repeatedly bisect our interval until we are within e of $k_{n,p}$.

For this demonstration, we will take n = 25 and p = 0.1.

For this demonstration, we will take n = 25 and p = 0.1.

Since we're doing this by hand, we'll take our error margin *e* to be a modest 0.01.

Step 1: Find $k_{n,p}$.

Step 1: Find $k_{n,p}$.

Our value:
$$k_{n,p} = \frac{25 \cdot 0.1 \cdot 0.9}{(1 - 2 \cdot 0.1)^2} = 3.5156.$$

Step 1: Find $k_{n,p}$.

Our value:
$$k_{n,p} = \frac{25 \cdot 0.1 \cdot 0.9}{(1 - 2 \cdot 0.1)^2} = 3.5156.$$

Step 2: Find a and b such that $f(a) > k_{n,p} > f(b)$.

Step 1: Find $k_{n,p}$.

Our value:
$$k_{n,p} = \frac{25 \cdot 0.1 \cdot 0.9}{(1 - 2 \cdot 0.1)^2} = 3.5156.$$

Step 2: Find a and b such that $f(a) > k_{n,p} > f(b)$.

Our values: a = 1, b = 3.

Step 3: Repeatedly bisect (a, b) until f(c) is within e of $k_{n,p}$.

Step 3: Repeatedly bisect (a, b) until f(c) is within e of $k_{n,p}$.

Calculate
$$c = \frac{a+b}{2}$$
.

Step 3: Repeatedly bisect (a, b) until f(c) is within e of $k_{n,p}$.

Calculate
$$c = \frac{a+b}{2}$$
.

▶ If $f(c) \le k_{n,p} - 0.01$, we need a small value of c, so we take our new interval to be (a, c).

Step 3: Repeatedly bisect (a, b) until f(c) is within e of $k_{n,p}$.

Calculate
$$c = \frac{a+b}{2}$$
.

- ▶ If $f(c) \le k_{n,p} 0.01$, we need a small value of c, so we take our new interval to be (a, c).
- ▶ If $f(c) \ge k_{n,p} + 0.01$, we need a larger value of c, so we take our new interval to be (c, b).

Step 3: Repeatedly bisect (a, b) until f(c) is within e of $k_{n,p}$.

Calculate
$$c = \frac{a+b}{2}$$
.

- ▶ If $f(c) \le k_{n,p} 0.01$, we need a small value of c, so we take our new interval to be (a, c).
- ▶ If $f(c) \ge k_{n,p} + 0.01$, we need a larger value of c, so we take our new interval to be (c, b).

Repeat this step until f(c) is within e of $k_{n,p}$, or more precisely $k_{n,p} - 0.01 < f(c) < k_{n,p} + 0.01$.

(Step 3)

The following table shows our iterations:

Iteration a b c f(c) $f(c) \le k_{n,p} - 0.01$ $f(c) \ge k_{n,p} + 0.01$

(Step 3)

Iteration	а	b	С	f(c)	$f(c) \leq k_{n,p} - 0.01$	$f(c) \geq k_{n,p} + 0.01$
1	2.00	3.000	2.5000	3.0164	True	False

(Step 3)

Iteration	а	ь	С	f(c)	$f(c) \leq k_{n,p} - 0.01$	$f(c) \geq k_{n,p} + 0.01$
1	2.00	3.000	2.5000	3.0164	True	False
2	2.00	2.500	2.2500	3.7129	False	True

(Step 3)

Iteration	а	b	С	f(c)	$f(c) \leq k_{n,p} - 0.01$	$f(c) \geq k_{n,p} + 0.01$
1	2.00	3.000	2.5000	3.0164	True	False
2	2.00	2.500	2.2500	3.7129	False	True
3	2.25	2.500	2.3750	3.3252	True	False

(Step 3)

Iteration	а	ь	С	f(c)	$f(c) \leq k_{n,p} - 0.01$	$f(c) \geq k_{n,p} + 0.01$
1	2.00	3.000	2.5000	3.0164	True	False
2	2.00	2.500	2.2500	3.7129	False	True
3	2.25	2.500	2.3750	3.3252	True	False
4	2.25	2.375	2.3125	3.5076	False	False

(Step 3)

The following table shows our iterations:

Iteration	а	b	С	f(c)	$f(c) \leq k_{n,p} - 0.01$	$f(c) \geq k_{n,p} + 0.01$
1	2.00	3.000	2.5000	3.0164	True	False
2	2.00	2.500	2.2500	3.7129	False	True
3	2.25	2.500	2.3750	3.3252	True	False
4	2.25	2.375	2.3125	3.5076	False	False

We take the last value of c: 2.3125.

Step 5: Find the sign of (1 - 2p).

Step 5: Find the sign of (1 - 2p).

Our $p = 0.1 \Rightarrow (1 - 2 \cdot 0.1) = 0.8 \Rightarrow$ positive.

Step 5: Find the sign of (1 - 2p).

Our
$$p = 0.1 \Rightarrow (1 - 2 \cdot 0.1) = 0.8 \Rightarrow$$
 positive.

Step 6: Final answer: {sign of (1 - 2p)} λ

Step 5: Find the sign of (1 - 2p).

Our $p = 0.1 \Rightarrow (1 - 2 \cdot 0.1) = 0.8 \Rightarrow$ positive.

Step 6: Final answer: {sign of (1-2p)} λ

Our final answer: $\lambda = 2.3125$.

Once we have λ , we can easily find σ :

Once we have λ , we can easily find σ :

$$\sigma = \sqrt{\frac{np(1-p)}{1-\frac{2}{\pi}\cdot\frac{\lambda^2}{1+\lambda^2}}} = \sqrt{\frac{25\cdot 0.1\cdot 0.9}{1-\frac{2}{\pi}\cdot\frac{2.3125^2}{1+2.3125^2}}} = 2.2029.$$

And with λ and σ , we can also find μ :

And with λ and σ , we can also find μ :

$$\mu = np - \sigma \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{\lambda}{\sqrt{1 + \lambda^2}}$$

$$= 25 \cdot 0.1 - 2.2029 \cdot \sqrt{\frac{2}{\pi}} \cdot \frac{2.3125}{\sqrt{1 + 2.3125^2}}$$

$$= 0.8867.$$

Here's the bibliography again \dots

BIBLIOGRAPHY I

- B. C. Arnold and G. D. Lin. Characterizations of the skew-normal and generalized chi distributions. *Sankhy: The Indian Journal of Statistics*, 66(4):593–606, 2004.
- Adelchi Azzalini. A class of distributions which includes the normal ones. *Scandinavian Journal of Statistics*, 12: 171–178, 1985.
- Adelchi Azzalini. The skew-normal distribution and related multivariate families. *Scandinavian Journal of Statistics*, 32: 159–188, 2005.
- Lee J. Bain and Max Engelhardt. *Introduction to Probability and Mathematical Statistics*. Duxbury, 2nd edition, 1992.
- Ching-Hui Chang, Jyh-Jiuan Lin, Nabendu Pal, and Miao-Chen Chiang. A note on improved approximation of the binomial distribution by the skew-normal distribution. *American Statistical Association*, 62(2):167–170, May 2008.



BIBLIOGRAPHY II

- A. K. Gupta, T. T. Nguyen, and J. A. T. Sanqui. Characterization of the skew-normal distribution. *Annals of the Institute of Statistical Mathematics*, 56:351–360, 2004.
- A. O'Hagan and Tom Leonard. Bayes estimation subject to uncertainty about parameter constraints. *Biometrika*, 63: 201–203, 1976.
- Arthur Pewsey. Problems of inference for azzalini's skew-normal distribution. *Journal of Applied Statistics*, 27(7): 859–870, 2000.
- Martin Schader and Friedrich Schmid. Two rules of thumb for the approximation of the binomial distribution by the normal distribution. *American Statistical Association*, 43(1):23–24, February 1989.