

# VE401 Probabilistic Methods in Eng.

## # Final Review

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if you want to edit this note, you can find it here <https://github.com/joydddd/VE401-2021SP-notes>

Mathematica code is also available on github.

## Categorical Test

$$f_{X_1 X_2 \dots X_k}(x_1, \dots, x_k) = \frac{n!}{x_1! \dots x_k!} p_1^{x_1} \dots p_k^{x_k}$$

$p_1, \dots, p_k \in (0, 1), n \in \mathbb{N} \setminus \{0\}$  is said to have a multinomial distribution with parameters  $n$  and  $p_1, \dots, p_k$ .

1. The (marginal) expectations of the individual random variables  $X_i$  are given by

$$E[X_i] = np_i, \quad i = 1, \dots, k.$$

2.  $\text{Var}[X_i] = np_i(1 - p_i), i = 1, \dots, k,$

3.  $\text{Cov}[X_i, X_j] = -np_i p_j, 1 \leq i < j \leq k.$

## Pearson Chi-squared Goodness of Fit Test

Test if the data follows **multinomial distribution with parameters  $(p_0, p_1, p_2, \dots)$**

$$H_0 : p_i = p_{i_0}, \quad i = 1, \dots, k$$
$$X_{k-1}^2 = \sum_{i=1}^k \frac{(X_i - np_{i_0})^2}{np_{i_0}}$$

We reject  $H_0$  at significance level  $\alpha$  if  $X_{k-1}^2 > \chi_{\alpha, k-1}^2$ .

**Cochran's Rule:** make sure the chi-squared approximation is appropriate

$E[X_i] = np_i \geq 1$  for all  $i = 1, \dots, k,$

$E[X_i] = np_i \geq 5$  for 80% of all  $i = 1, \dots, k,$

Especially if the  $p_i$  are not known roughly beforehand, care needs to be taken to ensure that the sample size  $n$  is sufficiently large so that these criteria can apply.

## Categorical Test on Discrete Distribution

Test if the data follows **a particular discrete distribution with m parameters estimated from the given data**

Divide our data into categories, use the discrete distribution to estimate the parameters, calculate the expected count of each category.

$$X_{k-1-m}^2 = \sum_1^k \frac{(O_i - E_i)^2}{E_i}$$

# Independence of Category

Test if the row and column categorizations are independent,

$$H_0 : p_{ij} = p_{i \cdot} \cdot p_{\cdot j}$$

We can now compare the observed frequencies  $O_{ij}$  in the  $(i, j)$  th cell to the expected frequencies  $E_{ij}$ .

$$X^2_{(r-1)(c-1)} = \sum_{i=1}^r \sum_{j=1}^c \frac{(O_{ij} - E_{ij})^2}{E_{ij}}$$
$$E_{ij} = n \cdot \hat{p}_{ij} = \frac{n_{i \cdot} \cdot n_{\cdot j}}{n}$$

We reject  $H_0$  if the value of  $X^2_{(r-1)(c-1)}$  exceeds the critical value of the corresponding chi-squared distribution.

## Linear Regression

### Simple Linear Regression

$Y \mid x = \beta_0 + \beta_1 x + E$  where  $E[E] = 0$

$E$ : remainder.  $E[E] = 0$  is guaranteed because  $b_0, b_1$  are unbiased estimators.

$\beta_0, \beta_1$ : true parameter of linear relationship.

$b_0, b_1$ : estimator for  $\beta_0, \beta_1$ .

$B_0, B_1$ : statistics for estimators  $b_0, b_1$

### Least Square Method

For each measurement  $y_i$  find residual  $e_i$

$$Y_i = b_0 + b_1 x_i + e_i$$

error sum of squares:

$$SS_E := e_1^2 + e_2^2 + \dots + e_n^2 = \sum_{i=1}^n (y_i - (b_0 + b_1 x_i))^2$$

least-squares estimates for  $\beta_0$  and  $\beta_1$  is determined by minimizing this sum of squares

$$S_{xx} := \sum_{i=1}^n (x_i - \bar{x})^2, S_{yy} := \sum_{i=1}^n (y_i - \bar{y})^2,$$
$$S_{xy} := \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}).$$
$$b_0 = \bar{y} - b_1 \bar{x}, \quad b_1 = \frac{S_{xy}}{S_{xx}}$$
$$SS_E = S_{yy} - b_1 S_{xy}$$

with confidence interval

$$B_1 \pm t_{\alpha/2, n-2} \frac{S}{\sqrt{S_{xx}}}, \quad B_0 \pm t_{\alpha/2, n-2} \frac{S \sqrt{\sum x_i^2}}{\sqrt{n S_{xx}}}$$

## Significance Test

We say that a regression is significant if there is statistical evidence that the slope  $\beta_1 \neq 0$ .

We reject

$$H_0 : \beta_1 = 0$$

at significance level  $\alpha$  if the statistic

$$T_{n-2} = \frac{B_1}{S/\sqrt{S_{xx}}}$$

satisfies  $|T_{n-2}| > t_{\alpha/2, n-2}$ .

## Test for Correlation

### Coefficient

$$R^2 := \frac{SS_T - SS_E}{SS_T} = \frac{S_x^2}{S_x S_{yy}} = \hat{\rho}_{XY}^2 \text{ (from Paired T-test)}$$

The coefficient  $R^2$  expresses the proportion of the total variation in  $Y$  that is explained by the linear model

Let  $(X, Y)$  follow a bivariate normal distribution with correlation coefficient  $\rho \in (-1, 1)$ . Let  $R$  be the estimator for  $\rho$ . Then

$$H_0 : \rho = 0$$

is rejected at significance level  $\alpha$  if

$$\left| \frac{R\sqrt{n-2}}{\sqrt{1-R^2}} \right| > t_{\alpha/2, n-2}$$

## Lack of Fit Test

Test if the linear regression model is appropriate. Need multiple y data at each x data point!

$$SS_E = SS_{E,pe} + SS_{E,lf}$$

pr: pure error, lf: lack of fitting (X & Y are not linearly related)

$$SS_{E,pe} := \sum_{i=1}^k \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_i)^2$$

$H_0$  : the linear regression model is appropriate

is rejected at significance level  $\alpha$  if the test statistic

$$F_{k-2, n-k} = \frac{SS_{E, lf} / (k-2)}{SS_{E, pe} / (n-k)}$$

satisfies  $F_{k-2, n-k} > f_{\alpha, k-2, n-k}$ .

## Prediction

$Y \mid x = \beta_0 + \beta_1 x + E$  where  $E[E] = 0$

$100(1 - \alpha)\%$  Prediction interval for  $Y \mid x$  :

$$\widehat{Y \mid x} \pm t_{\alpha/2, n-2} S \sqrt{1 + \frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}}$$

$100(1 - \alpha)\%$  prediction interval for  $\mu_Y \mid x$  :

$$\widehat{\mu_Y \mid x} \pm t_{\alpha/2, n-2} S \sqrt{\frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}}$$