

# VE401 Probabilistic Methods in Eng.

## RC#2 Discrete Distributions & Continuous Random Variables

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### Discrete Distributions

$X \sim \text{Bernoulli}(p)$

$$X : S \rightarrow \{0, 1\} \subset \mathbb{R}$$
$$f_X : \{0, 1\} \rightarrow \mathbb{R}, \quad f_X(x) = \begin{cases} 1 - p & \text{for } x = 0 \\ p & \text{for } x = 1 \end{cases} \quad (0 < p < 1)$$

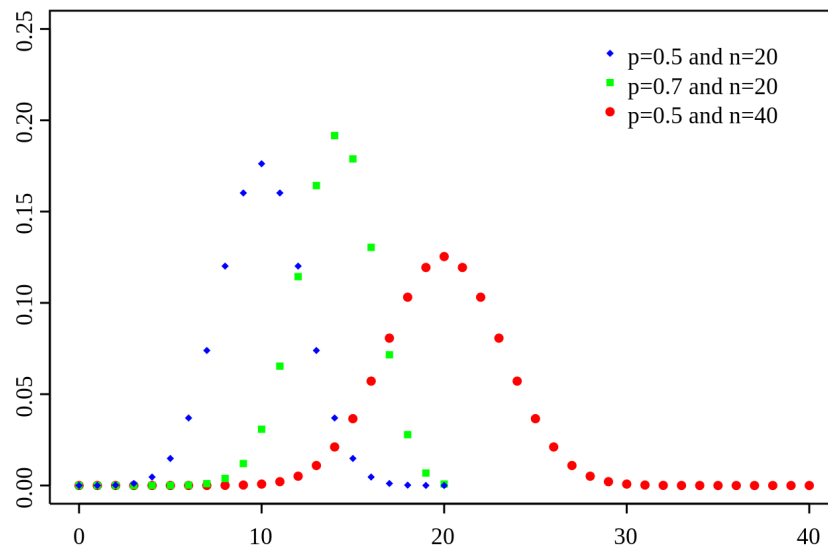
Then  $X$  is said to be a **Bernoulli random variable** or follow a Bernoulli distribution with parameter  $p$ .

**Binomial Distribution**  $X \sim B(n, p)$

Getting  $x$  success in  $n$  Bernoulli trials:

$$X : S \rightarrow \Omega = \{0, \dots, n\} \subset \mathbb{R} \quad (n \in \mathbb{N} \setminus \{0\})$$
$$f_X : \Omega \rightarrow \mathbb{R}, \quad f_X(x) = \binom{n}{x} p^x (1 - p)^{n-x} \quad (0 < p < 1)$$

Then  $X$  is said to be a **binomial random variable with parameters  $n$  and  $p$** .



## Geometric Distribution $X \sim \text{Geom}(p)$ .

Getting the first success on the  $x$ th Bernuolli trail:

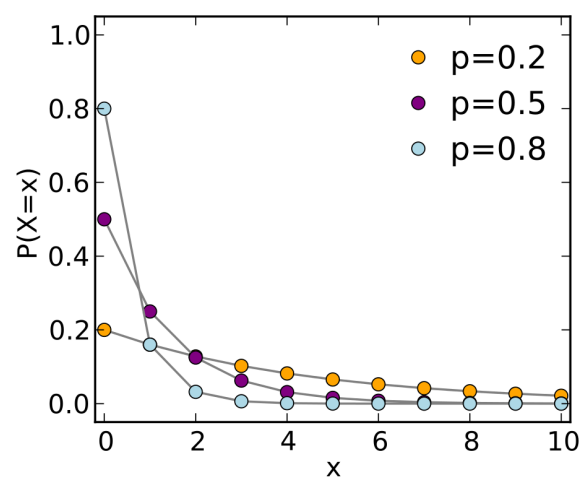
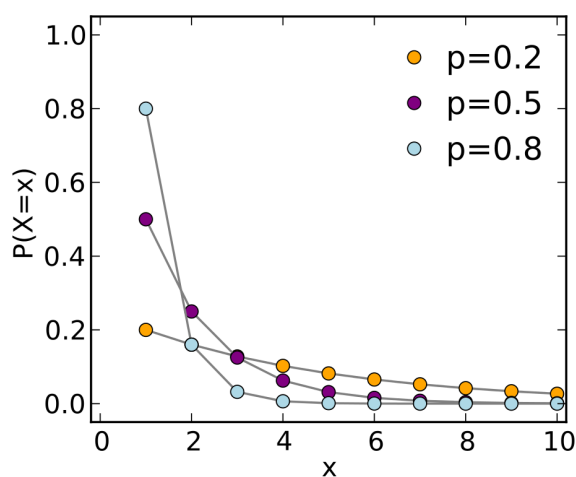
Let  $S$  be a sample space and

$$X : S \rightarrow \Omega = \mathbb{N} \setminus \{0\}$$

Let  $0 < p < 1$  and define the density function  $f_X : \mathbb{N} \setminus \{0\} \rightarrow \mathbb{R}$  given by

$$f_X(x) = (1 - p)^{x-1} p.$$

We say that  $X$  is a **geometric random variable with parameter  $p$** .



- moment generating function :  $q = 1 - p$

$$m_X : (-\infty, -\ln q) \rightarrow \mathbb{R}, \quad m_X(t) = \frac{pe^t}{1 - qe^t}$$

- Expectation & Variance

$$E[X] = \frac{1}{p} \quad \text{Var}[X] = \frac{q}{p^2}$$

## Pascal Distribution

obtain the  $r$ th success at  $x$ th trial:

$$\begin{aligned} X : S \rightarrow \Omega &= \mathbb{N} \setminus \{0, 1, \dots, r-1\} \\ &= \{r, r+1, r+2, \dots\} \end{aligned} \quad (r \in \mathbb{N} \setminus \{0\})$$

$$f_X : \Omega \rightarrow \mathbb{R}, \quad f_X(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad 0 < p < 1$$

is said to follow a **Pascal distribution with parameters  $p$  and  $r$** .

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$$m_X : (-\infty, -\ln q) \rightarrow \mathbb{R}, \quad m_X(t) = \frac{(pe^t)^r}{(1 - qe^t)^r}, \quad q = 1 - p$$

- Expectation & Variance

$$E[x] = \frac{r}{p} \quad \text{Var}[X] = \frac{rq}{p^2}$$

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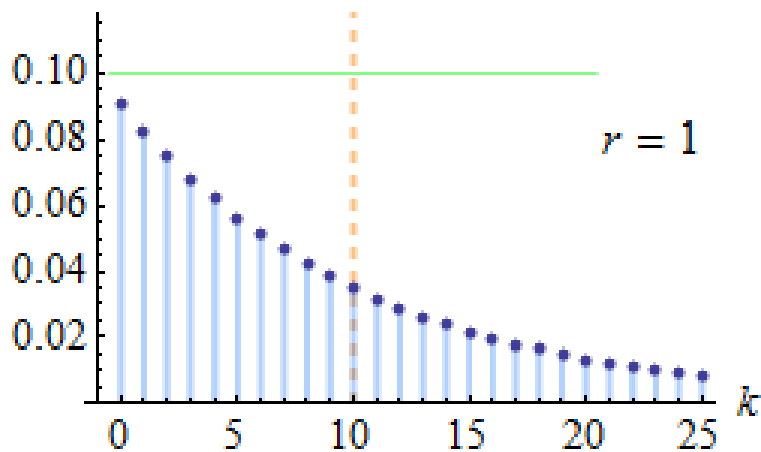
## Negative binomial distribution

Experience  $x$  failures before the  $r$  success:

$$X : S \rightarrow \Omega = \mathbb{N} \setminus \{0, 1, \dots, r-1\} = \{r, r+1, r+2, \dots\}, \quad r \in \mathbb{N} \setminus \{0\}$$

$$f_X : \Omega \rightarrow \mathbb{R}, \quad f_X(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad 0 < p < 1$$

is said to follow a **Pascal distribution with parameters  $p$  and  $r$** .



Show that the sum of two independent and identical geometric random variables follows a Pascal distribution with  $r = 2$ .

## Poisson distribution

arrival in a continuous time interval:

Assumptions

1. Independence: If the intervals  $T_1, T_2$  doesn't overlap (except perhaps at one point), then the numbers of arrivals in these interval are independent of each other.
2. Constant rate of arrivals

Let  $k = \lambda t$ ,  $\lambda$  arriving rate,  $t$  time interval

$$f_X(x) = \frac{k^x e^{-k}}{x!}$$

is said to have a **Poisson distribution with parameter  $k$** .

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$$m_X : \mathbb{R} \rightarrow \mathbb{R}, \quad m_X(t) = e^{k(e^t - 1)}$$

- Expectation & Var

$$E[X] = k \quad \text{Var}[X] = k$$

- cumulative distribution function

$$F(x) = P[X \leq x] = \sum_{y=0}^{\lfloor x \rfloor} \frac{e^{-k} k^y}{y!}$$

## Poisson Approximation to the Binomial Distribution

Consider the density  $f$  of the binomial distribution,

$$f(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

Let  $k$  be fixed so that  $np = k$  and set  $p = k/n$ . Replace  $p$  by  $k/n$  everywhere in (\*) and then let  $n \rightarrow \infty$ . Use Stirling's formula<sup>1</sup> to show that for every  $x$ ,  $f(x) \rightarrow (k^x / x!) e^{-k}$ , the density of the Poisson distribution with parameter  $k$ .

<sup>1</sup> Stirling's formula states that

$$n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n \quad \text{as } n \rightarrow \infty$$

where  $f(n) \sim g(n)$  as  $n \rightarrow \infty$  means that  $\lim_{n \rightarrow \infty} \frac{f(n)}{g(n)} = 1$

## Continuous Random Variable

Let  $S$  be a sample space. A continuous random variable is a map  $X : S \rightarrow \mathbb{R}$  together with a function  $f_X : \mathbb{R} \rightarrow \mathbb{R}$  with the properties that

(i)  $f_X(x) \geq 0$  for all  $x \in \mathbb{R}$  and

(ii)  $\int_{-\infty}^{\infty} f_X(x) dx = 1$ .

The integral of  $f_X$  is interpreted as the probability that  $X$  assumes values  $x$  in a given range, i.e.,

$$P[a \leq X \leq b] = \int_a^b f_X(x) dx$$

The function  $f_X$  is called the **probability density function** (or just density) of the random variable  $X$ .

$$F_X(x) := P[X \leq x] = \int_{-\infty}^x f_X(y) dy$$

$F_X$  is called the **cumulative distribution function**. We can obtain the density  $f_X$  from  $F_X : f_X(x) = F'_X(x)$

$$E[X] := \int_{\mathbb{R}} x \cdot f_X(x) dx$$

$$E[\varphi \circ X] = \int_{-\infty}^{\infty} \varphi(x) \cdot f_X(x) dx$$

$$\text{Var}[X] := E[(X - E[X])^2] = E[X^2] - E[X]^2$$

$$m_X(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$$

- **median  $M_X$** : defined by  $P[X \leq M_X] = 0.5$ . aka the time where half of the components will have failed
- **mean  $E[X]$** .
- **mode  $x_0$** : the location of the maximum of  $f_X$  (if there is a unique maximum location). aka. the time with the greatest failure density, i.e., the time around which failure is most likely. For the exponential distribution,  $x_0 = 0$

## Transformation of random variables

1.3.13. Theorem. Let  $X$  be a continuous random variable with density  $f_X$ . Let  $Y = \varphi \circ X$ , where  $\varphi : \mathbb{R} \rightarrow \mathbb{R}$  is strictly monotonic and differentiable. The density for  $Y$  is then given by

$$f_Y(y) = f_X(\varphi^{-1}(y)) \cdot \left| \frac{d\varphi^{-1}(y)}{dy} \right| \quad \text{for } y \in \text{ran } \varphi$$

and

$$f_Y(y) = 0 \quad \text{for } y \notin \text{ran } \varphi .$$

## Exponential Distribution

The probability of  $x$  arrivals in the time interval  $[0, t]$

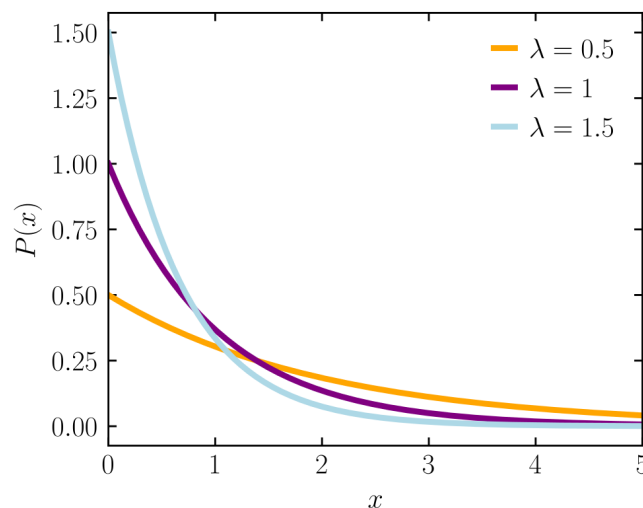
Time of the first arrival (continuous random variable)  $T$

$F_T(t)$  the cumulative distribution on the density of  $T$

$$f_T(t) = \frac{d}{dt} F_T(t) = \lambda e^{-\lambda t} \text{ (exponential distribution with } \beta = \lambda)$$

$$f_\beta(x) = \begin{cases} \beta e^{-\beta x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$$

$$E[X] = \frac{1}{\beta} \quad \text{Var}[X] = \frac{1}{\beta^2} \quad m_X(t) = \frac{\beta}{\beta - t}$$



## Gamma/Chi Square Distribution

The time of  $r$  arrivals

$$F_{Tj}(t) = \lambda e^{-\lambda t} \frac{(\lambda t)^{j-1}}{(j-1)!}$$

Let  $\alpha, \beta \in \mathbb{R}, \alpha, \beta > 0$ . A continuous random variable  $(X, f_{\alpha, \beta})$  with density

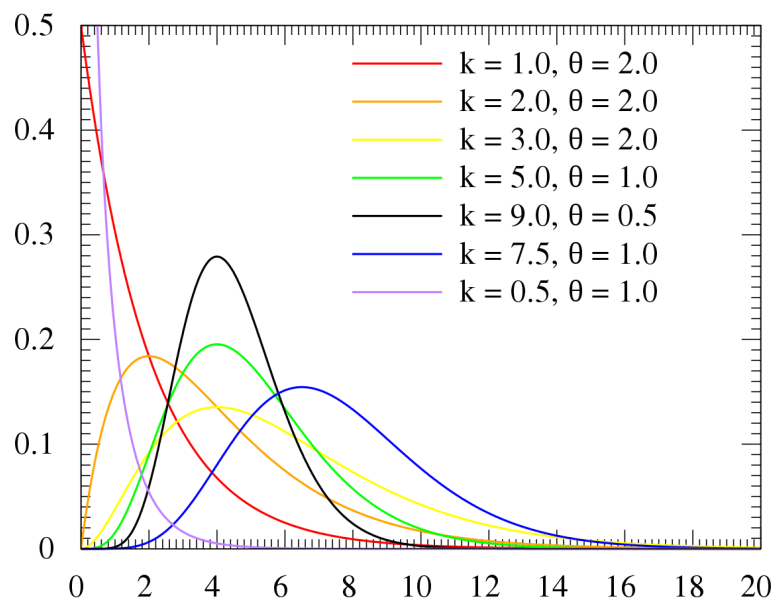
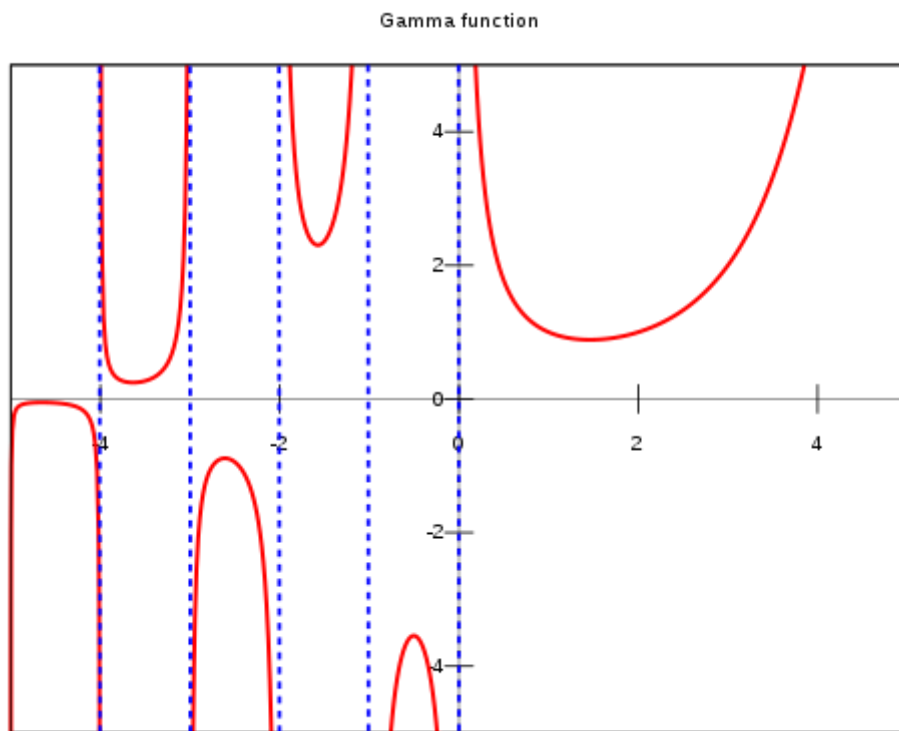
$$f_{\alpha, \beta}(x) = \begin{cases} \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}, & x > 0 \\ 0, & x \leq 0, \end{cases}$$

is said to follow an **gamma distribution with parameters  $\alpha$  and  $\beta$** .

Here

$$\Gamma(\alpha) = \int_0^{\infty} z^{\alpha-1} e^{-z} dz, \quad \alpha > 0$$

is the Euler gamma function. ( $n! = \Gamma(n+1)$ )



- $E[X] = \alpha\beta \quad \text{Var}X = \alpha\beta^2$

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$$m_X : (-\infty, 1/\beta) \rightarrow \mathbb{R}, \quad m_X(t) = (1 - \beta t)^{-\alpha}$$

Let  $\gamma \in \mathbb{N}$ . A continuous random variable  $(\chi_\gamma^2, f_X)$  with density



$$f_{\gamma}(x) = \begin{cases} \frac{1}{\Gamma(\gamma/2)2^{\alpha}} x^{\gamma/2-1} e^{-x/2}, & x > 0 \\ 0, & x \leq 0 \end{cases}$$

is said to follow a **chi-squared distribution with  $\gamma$  degrees of freedom**. chi-squared distribution is simply that of a **gamma random variable with  $\beta = 2$  and  $\alpha = \gamma/2$** .

$$E[\chi_{\gamma}^2] = \gamma, \quad \text{Var}[\chi_{\gamma}^2] = 2\gamma$$

## Stock Control Problems

[https://www.jstor.org/stable/3007410?seq=7#metadata\\_info\\_tab\\_contents](https://www.jstor.org/stable/3007410?seq=7#metadata_info_tab_contents)

In order to ensure a high standard of serviceability without incurring unnecessary aircraft delay, an airline adopts the following replacement policy for its repairable aircraft components:

1. An unserviceable component is removed from an aircraft and sent for inspection to repair.
2. An immediate demand for a serviceable replacement is made to the stores. If available, the replacement is at once fitted to the aircraft.
3. When the previously unserviceable component has been repaired to be serviceable, it is placed in the stores.

Suppose that the components are delivered serviceable to stores in the same order as they are removed from aircraft. We denote that

- "Demand": a component for repair.
- "Lead time" with parameter  $k$ : the time interval between the removal of the component and its entry into stores, i.e., the time interval for the store to deliver an order of  $k$  specified components.

What is the probability  $p_{rk}$  of exactly  $r$  demands during the lead time when the parameter value is  $k$ ?

