Log-Gaussian Cox Process for London crime data

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Cox Process

Cox process is a natural choice for an environmentally driven point process. (Diggle et al., 2013)

Definition

Cox process is defined by two postulates:

- 1. $\Lambda = \{\Lambda(x) : x \in \mathbf{R}^2\}$ is a nonnegative-valued stochastic process;
- 2. conditional on the realisation $\Lambda(x) = \lambda(x) : x \in \mathbb{R}^2$, the point process is an inhomogeneous Poisson process with intensity $\lambda(x)$.

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Log-Gaussian Cox Process

ightharpoonup Cox process with intensity driven by a Gaussian Process f(x):

$$\Lambda(\boldsymbol{x}) = \exp{(f(\boldsymbol{x}))}.$$

- Tractability of multivariate Normal distribution carries over to the associated Cox process.
- A common approach for analysis is to introduce a grid over the domain X.

Field inference - Laplace Approximation

Flaxman et al. (2015)

▶ Approximate the posterior distribution of the Gaussian Process by:

$$p(\mathbf{f}|\mathbf{y},X) \approx \mathcal{N}\left(\hat{\mathbf{f}}, -\left(\nabla \nabla \Psi(\mathbf{f})|_{\hat{\mathbf{f}}}\right)^{-1}\right),$$

where $\Psi(\mathbf{f}) := \log p(\mathbf{f}|\mathbf{y}, X) \stackrel{\text{const}}{=} \log p(\mathbf{y}|\mathbf{f}) + \log p(\mathbf{f}|X)$ is unormalised log posterior.

Newton's method to find $\hat{\mathbf{f}}$.

Field inference - Newton Optimisation

Flaxman et al. (2015)

► The Newton optimisation step:

$$\mathbf{f}^{\mathsf{new}} \leftarrow \mathbf{f}^{\mathsf{old}} - (\nabla \nabla \Psi)^{-1} \, \nabla \Psi$$

▶ $\nabla\nabla\Psi$ and $\nabla\Psi$ require inverting the covariance matrix of the GP:

$$\nabla \Psi(\mathbf{f}) = \nabla \log p(\mathbf{y}|\mathbf{f}) - K^{-1}\mathbf{f}$$
$$\nabla \nabla \Psi(\mathbf{f}) = -\mathbf{W} - \mathbf{K}^{-1},$$

where $W \coloneqq -\nabla\nabla \log p(\mathbf{y}|\mathbf{f})$.

Hyperparameters - Marginal Likelihood

Flaxman et al. (2015)

- ▶ Not the whole story: $p(\mathbf{f}|X)$ should be $p(\mathbf{f}|X, \boldsymbol{\theta})$, where \boldsymbol{K} depends on $\boldsymbol{\theta}$.
- Marginal log-likelihood:

$$\log p(\mathbf{y}|X, \boldsymbol{\theta}) = \log \int \exp \left[\Psi(\mathbf{f})\right] d\mathbf{f}$$
$$\approx \log p(\mathbf{y}|\hat{\mathbf{f}}) - \frac{1}{2} \mathbf{f}^{\top} K^{-1} \mathbf{f} - \frac{1}{2} \log |\mathbf{I} + \mathbf{K} \mathbf{W}|$$

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Computation

Flaxman et al. (2015)

The computatios require matrix inverse, matrix determinant, and matrix-vector multiplications:

- \triangleright Conjugate gradient for inverting K, exploiting Kronecker structure
- ▶ Determinant approximation due to Fiedler (1971):

$$\log |\boldsymbol{I} + \boldsymbol{K} \boldsymbol{W}| = \log \left(|\boldsymbol{K} + \boldsymbol{W}^{-1}| |\boldsymbol{W}| \right)$$

$$\leq \log \left\{ \prod_{i} \left(e_i + W_{ii}^{-1} \right) \prod_{i} W_{ii} \right\}$$

$$= \sum_{i} \log \left(1 + e_i W_{ii} \right),$$

where e_1, \ldots, e_n are sorted eigenvalues of K.

▶ Matrix-vector multiplication are efficient due to Kronecker structure.

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Separable kernels, Kronecker methods

Flaxman et al. (2015)

► Separable kernel functions:

$$k((x_1, y_1), (x_2, y_2)) = k_1(x_1, x_2)k_2(y_1, y_2)$$

► On a regular grid, we get:

$$K = K_1 \otimes K_2$$

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Experiment

Spatial model with isotropic Matérn covariance function:

▶ Dataset used: 2016 data

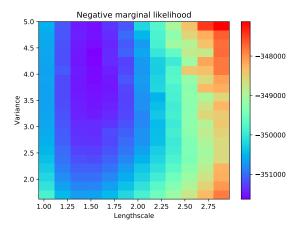
► Crime types: Burglary, Theft from the person

▶ Grid: 117x91, one cell is an area of 500m by 500m.

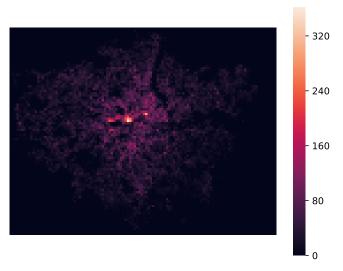
▶ Two parameters inferred: lengthscale(ℓ), marginal variance (σ^2)

Burglary - inferred parameters

Inferred parameters: $\ell = 1.45$, and $\sigma^2 = 4.32$



Burglary - counts

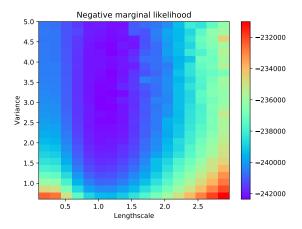


Burglary - latent field

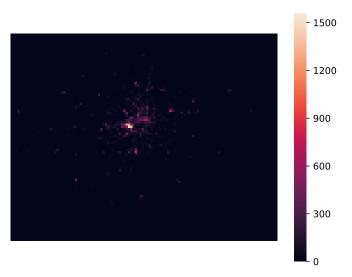


Theft from the person - inferred parameters

Inferred parameters: $\ell = 1.11$, and $\sigma^2 = 2.80$



Theft from the person - counts



Theft from the person - latent field



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Time component, and predictions

Possible options are:

► A kernel with period of 12 months for seasonal variation (Flaxman, 2014):

$$k_P(t, t') = \exp\left(-\frac{2\sin^2\left(\frac{(t-t')\pi}{12}\right)}{\ell^2}\right)$$

▶ Spectral mixture kernel with *Q* components (Flaxman et al., 2015):

$$k(\tau) = \sum_{q=1}^{Q} w_q \exp\left(-2\pi^2 \tau^2 v_q\right) \cos\left(2\pi\tau\mu_q\right)$$

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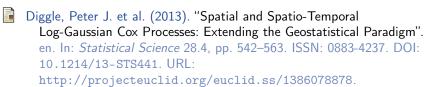
Stochastic PDEs

- ► Finite Element Method to solve SPDEs as described in Lindgren, Rue, and Lindström (2011).
- ► Sigrist, Künsch, and Stahel (2015) solve transport-diffusion SPDE using spectral methods on a grid.

More on this from Seppo.

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Matérn Covariance Function

$$k(r) = \frac{2^{1-\nu}}{\Gamma(\nu)} \left(\frac{\sqrt{2\nu}r}{\ell}\right)^{\nu} K_{\nu} \left(\frac{\sqrt{2\nu}r}{\ell}\right)$$

We use $\nu=2.5$.

Inference Newton step details

Extra slides 28

Kronecker Algebra

Extra slides 29

Incomplete grids

Wilson et al. (2014)

We have that $y_i \sim \operatorname{Poisson}(f_i)$. For the points of the grid that are not in the domain, we let $y_i \sim \mathcal{N}(f_i, \epsilon^{-1})$ and $\epsilon \to 0$. Hence,

$$p(\mathbf{y}|\mathbf{f}) = \prod_{i \in \mathcal{D}} \frac{\left(e^{\mathbf{f}_i}\right)^{\mathbf{y}_i} e^{-e^{\mathbf{f}_i}}}{\mathbf{y}_i!} \prod_{i \notin \mathcal{D}} \frac{1}{\sqrt{2\pi\epsilon^{-1}}} e^{\frac{-\epsilon(\mathbf{y}_i - \mathbf{f}_i)^2}{2}}$$

The log-likelihood is thus:

$$\sum_{i \in \mathcal{D}} \left[\mathsf{y}_i \mathsf{f}_i - \exp(f_i) + \mathsf{const} \right] - \frac{1}{2} \sum_{i \notin \mathcal{D}} \epsilon(\mathsf{y}_i - \mathsf{f}_i)^2$$

We now take the gradient of the \log of the likelihood as

$$\nabla \log p(\mathbf{y}|\mathbf{f})_i = \begin{cases} \mathbf{y}_i - \exp(\mathbf{f}_i), & \text{if } i \in \mathcal{D} \\ \epsilon(\mathbf{y}_i - \mathbf{f}_i), & \text{if } i \notin \mathcal{D} \end{cases}$$

and the hessian of the log-likelihood as

$$\nabla\nabla \log p(\mathbf{y}|\mathbf{f})_{ii} = \begin{cases} -\exp(\mathsf{f}_i), & \text{if } i \in \mathcal{D} \\ -\epsilon & \text{if } i \notin \mathcal{D} \end{cases}.$$