$$(x_i) = \sum_{i=1}^{n} e_i = \sum_{i=1}^{n} (y_i - \hat{y_i})$$

$$= \sum_{i=1}^{n} y_i - (\bar{y} + \sqrt{\frac{\sigma_{i+1}}{\sigma_{x}}} (x_i - \bar{x}))$$

$$\hat{y} = \bar{y} + r \sigma y = \overline{X} - \overline{X}$$

Plugging in X as x into the SLA gives of

A)
$$\sum_{i=1}^{n} e_{i} = 1^{T}e_{i} = 0$$
 $x^{T}(Y - X\hat{\theta}) = 0$
 $\hat{\theta} = (X^{T}X)^{T}X^{T}Y$
 $1^{T}e_{i} = 1^{T}(Y - \hat{Y})$
 $= 1^{T}X^{T}(I - X\hat{\theta})Y$
 $= 1^{T}X^{T}(I - Y(X^{T}X)^{T}X^{T})Y$
 $= e^{T}(X^{T} - X^{T}X(X^{T}X)^{T}X^{T})Y$
 $= e^{T}(X^{T} - X^{T})Y$
 $= 0$

b) $X^{T}e_{i} = X^{T}(Y - \hat{Y})$

- 0

a)
$$P(r) = \frac{1}{n} \sum_{i=1}^{n} (y_i - Y_{X_i})^2$$

 $P'(r) = -\frac{2}{n} \sum_{i=1}^{n} (y_i - Y_{X_i}) \times i$

$$-\frac{2}{N}\sum_{i=1}^{N}(y_{i}-Y_{i})\times i=0$$

$$P''(y) = \frac{\lambda}{\lambda y} - \frac{\lambda}{y} \sum_{i=1}^{n} (y_i - Y_{x_i})^2$$

$$= \frac{\lambda}{y} \sum_{i=1}^{n} (-x_i)^2$$

$$= \frac{\lambda}{y} \sum_{i=1}^{n} (x_i)^2 > 0$$

:- minimum

- a) Does not hold wlour an intercept term because wlour the column of 1's, the dox product bown 1 & e is not always D.
- b) tlolds true even wlonk an intercept term because even whom the column of 1's to eneate the matrix X, the dot product between \vec{x} is e is 0.
- c) This holds true even wlour an intercept term because \hat{y} whom the columns of is is the same things as the column secrov \hat{x} . Therefore, the therefore the dox product between \hat{y} and e will still be o.
- d) This will not hold wlout an intercept term because (\vec{x}, \vec{y}) will only be on the regression line if the intercept term exists in the SLR and is equal to \vec{y} .