

Chapter 48 Outline

Network Modeling: Estimation, Inference, Comparison, and Selection

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Introduction:

1 Introduction

Do states go to war with the enemies of their enemies? Which international trade relationships are most likely to break down over the next decade? What would happen to the Statistical inference with network data permits researchers to test hypotheses about network generation, predict network structure, and simulate networks from probabilistic models. How would the world system of military alliances rewire if the United States were to exit the North Atlantic Treaty Organization? What these questions have in common is that they address systems of international relationships (conflict, trade, and alliances)—systems that can be represented as networks in which the states constitute the nodes, and the relationships constitute the edges between nodes. Inferential network analysis represents a methodological class that can be drawn upon to answer questions like these. Most statistical models used in the social sciences rely on the assumption that the observations in a dataset (or at least clusters of observation) are drawn independently from a common data generating process. However, when analyzing networks it is often inappropriate to assume that observations are independent ([Hafner-Burton, Kahler and Montgomery, 2009](#); [Harris, 2013](#)). Indeed, the network analyst is often interested in studying the ways in which observations depend upon each other ([Ogburn, 2018](#)) (e.g., do friends influence each others' choices to vote ([Bond et al., 2012](#)), do legislators reciprocate support for legislation ([Kirkland and Williams, 2014](#))).

Many, perhaps even most, of the empirical phenomena of interest to scholars of international relations can be represented as international networks. Paired with the availability of accessible software implementations, the conceptual appropriateness of inferential network analysis in international relations has led to widespread application over the last 10–15 years. Networks that have been studied through the use of inferential network analysis include, but are not limited to, trade ([Chu-Shore, 2010](#); [Chyzh, 2016](#); [Fagiolo and Mastrorillo, 2014](#); [Ward and Hoff, 2007](#)), conflict ([Cranmer and Desmarais, 2011](#); [Dorff and Ward, 2013](#); [Gallop, 2016](#); [Ward, Siverson and Cao, 2007](#)), alliances ([Cranmer, Desmarais and Menninga, 2012](#)), transnational terrorism ([Asal et al., 2016](#); [Bush and Bichler, 2015](#); [Desmarais and Cranmer, 2013](#); [Metternich et al., 2013](#)), sanctioning ([Cranmer, Heinrich and Desmarais, 2014](#); [Dorff and Minhas, 2017](#)), international governmental organizations ([Cao, 2012](#); [Davis and Pratt, 2017](#); [Lupu and Greenhill, 2017](#)), and non-governmental organizations ([Atouba and Shumate, 2015](#)). This research has led to several innovative findings. For example, [Kinne \(2018\)](#) finds that defense cooperation agreements (DCAs) between states are self-exciting in that they become more attractive as more states sign on, leading to patterns of triangle closure (i.e., friend of a friend is a friend) and preferential attachment (i.e., popular states gain more ties) in the formation of networks through DCAs. [Duque \(2018\)](#), in a study of international diplomatic networks, finds similar relational dynamics in that the diplomatic relations of a state with one partner state affects its relations with other partner states.

Methods for inferential network analysis come primarily in the form of probabilistic models for networks, which are fit to data using common estimation frameworks (e.g., maximum likelihood estimation,

Bayesian inference). Inferential network analysis is now far too broad of a methodological area to effectively review all of the models in any detail. In what follows we will review three modeling approaches in detail. These include the exponential random graph model (ERGM) (Cranmer and Desmarais, 2011), the latent space model (LSM) (Dorff, Minhas and Ward, 2016), and the stochastic block model (SBM) (Latouche, Birmelé and Ambroise, 2011). The first model, which we refer to as the latent space model (LSM), is a class of models in which each node is attributed with a position in a latent space of features. The second model, the exponential random graph model (ERGM) is a model that can be customized to represent networks with any quantifiable characteristic (e.g., a high number of reciprocated ties, a strong degree of clustering). The third model we discuss in detail is the stochastic blockmodel (SBM). The first two models, LSM and ERGM, are commonly used in international relations. The SBM, however, has seen little use in IR. We make the case that the SMB should be used in political science as an alternative to ERGM and SBM. After presenting a review of these three models, we provide a brief overview of several approaches that we did not review in detail, but of which IR scholars should be aware. After reviewing these models we present an application in which we replicate the analyses from using each of these models and focus in particular on the application of SBM, since SBM is new in IR.

In this chapter we will review three of the most popular statistical models that have been developed to analyze network data, along with their extensions. These include the exponential random graph model (ERGM) (Cranmer and Desmarais, 2011), the latent space model (LSM) (Dorff, Minhas and Ward, 2016), and the stochastic block model (SBM) (Sweet, 2015). We will review the application of each method to networks in which ties are binary (e.g., are two legislators on at least one committee together), count (e.g., on how many committees do two legislators both serve), and continuous (e.g., how much time do two legislators spend together in committee meetings each week). The complexity and analytical intractability of the mathematical forms of these models render exact approaches to likelihood maximization infeasible. As such, estimation in these models proceeds by sampling from a Bayesian posterior and/or approximate methods of maximum likelihood (Nowicki and Snijders, 2001; Raftery et al., 2012; Van Duijn, Gile and Handcock, 2009). We will review these estimation methods, and how they are applied to the ERGM, LSM, and SBM. This chapter will be divided into three sections. In the first section we will give an overview of the most common models used in network analysis. In the second section we will discuss the approaches to estimation that are commonly used with these models. In the last section we will provide an application of all three models. We replicate the analysis in (Wojcik, 2017), which was originally completed using ERGM, and analyze the replication data using both LSM and SBM.

2 Model description

The three models we review in detail, the LSM, ERGM, and SBM share a common starting point in the case of dichotomous edges (i.e., the edge does or does not exist). Consider a research problem in which the researcher is interested in using covariates that can be measured on the directed pair (i, j) , where i is the index of the potential sender of an edge, and j is the index of the potential receiver of the edge, to model whether the edge (i, j) exists. Denote the edge indicator $y_{i,j} = 1$ if there is an edge from node i to node j , and 0 otherwise. Let $x_{i,j}$ be a covariate that can be measured on the directed dyad (e.g., if the nodes are states, the distance between their capital cities), or mapped to the dyad (e.g., $x_{i,j}$ is the GDP of the sender state, i). A standard approach to modeling the dyadic variable y as a function of x would be to estimate a logistic regression in which y is the dependent variable and x is the independent variable. The LSM, ERGM, and SBM all reduce to a standard regression model in which edge indicators (or values) are regressed on measured covariates in the case where the network components of the respective models do not contribute to the fit of the model (Lubbers and Snijders, 2007; Raftery et al., 2012; Sweet, 2015).

Specifically, logistic regression is a special case of LSM, ERGM, and SBM when modeling a dichotomous (i.e., absent or present) network.

In the case of each model, network structure is layered on top of the regression. The particular form of network structure added varies across the models. In the LSM, each node is represented by some number of continuous-valued features (e.g., coordinates in a Euclidean space), and the probability of a tie is given by some function of the features (e.g., the Euclidean distance between nodes). We refer to the network structure incorporated in the LSM framework as ‘selection’, since the function defined on the latent feature serves as a partner selection function for nodes in the network. The network component under ERGM is designed to capture the prevalence of subnetworks that are theoretically interesting or otherwise distinct. Examples of these subnetworks include triangles (i.e., triples of nodes (i, j, k) in which there is an edge connecting each pair of nodes), and mutual dyads (i.e., dyads in which there is an edge from i to j and from j to i). In the ERGM, subnetwork prevalence is controlled through the incorporation of dependence among the edges. For example, if a network is modeled to have a relatively high number of mutual dyads, the formation of an edge from i to j increases the likelihood that an edge will form from j to i . Since the basic building blocks of the network component of the ERGM reflect different forms of dependence, we refer to the network structure in the ERGM as ‘dependence’. The network structure of the SBM is build on a common interest in identifying the communities, clusters, or groups that define the main blocks of ties in the networks. Under the SBM’s network structure, the blocks of which two nodes are members determines whether those two nodes will form an edge. Since the network component of the SBM is focused on segmenting the nodes into a set of groups, we refer to the network structure in the SBM as segmentation.

2.1 Latent space modeling of networks

The notion of selection is deeply rooted in political science literature. For example, in the study of legislative networks, ties between legislators have been repeatedly found to be driven by partisanship, ideology, and the geographic proximity of districts ([Bratton and Rouse, 2011](#); [Clark and Caro, 2013](#); [Osei and Malang, 2018](#)). The probability of a tie between two legislators is a function of the features of the two legislators—primarily ideology and geography. The LSM reflects this very structure—the probability of a tie depending on a combination of functions of node (e.g., legislator) features—with exception being that the features inferred in the LSM are latent. That is, given one or more functions according to which the latent features affect the probability of an edge, in the LSM the nodes’ feature values are inferred to be those that best explain the pattern of edges observed in the data.

Under the LSM,

$$\Pr(y_{i,j} = 1) = \text{logit} \left[\beta_0 + \sum_{p=1}^P \beta_p x_{i,j}^{(p)} + d(\mathbf{z}_i, \mathbf{z}_j) \right],$$

where $\text{logit}(a) = 1/(1 + \exp(-a))$, β_0 is an intercept that controls the overall density of the network, $\sum_{p=1}^P \beta_p x_{i,j}^{(p)}$ models the effect of measured covariates on the probability of an edge forming, \mathbf{z}_i and \mathbf{z}_j are latent coordinates of nodes i and j , and $d()$ is a distance function that maps the coordinates into a metric space. In the LSM, the \mathbf{z} are inferred as parameters. They are latent attributes that capture the structure of the network, but are not observed. Example functions that have been used for $d()$ include the negated Euclidean distance (e.g., [Mahmood and Sismeiro, 2017](#)),

$$d(\mathbf{z}_i, \mathbf{z}_j) = -\sqrt{\sum_{k=1}^K \left(z_i^{(k)} - z_j^{(k)} \right)^2},$$

and the bilinear metric (e.g., [Vance, Archie and Moss, 2009](#)),

$$d(\mathbf{z}_i, \mathbf{z}_j) = \sum_{k=1}^K z_i^{(k)} \times z_j^{(k)}.$$

The latent variables account for network structure that is not modeled effectively using the observed covariates.

Selection of the particular distance function used can be based on a few different factors. First, the researcher may have a substantive reason to prefer one distance function over another. For example, the Euclidean function models the effect of latent variables such that the log-odds of a tie is inversely proportional to the straight-line distance between two nodes in the latent space whereas under the bilinear function the log-odds of a tie is inversely proportional to the angle between the two coordinate vectors. The researcher may consider one distance function to be most appropriate in their application. Second, one function may fit the data most effectively. Based on information theoretic measures such as the Bayesian Information Criterion (BIC) ([Wang, 2009](#)) or visual assessments of model fit, the researcher may determine that one of the distance function provides the best fit to the data. Third, if the network is very large, computational properties of estimation may make the use of one of the distance functions more feasible. If the Markov Chain Monte Carlo approach to Bayesian inference in the LSM converges much faster under one distance function than another, the researcher may choose to go with the model that converges faster.

2.2 The exponential random graph model

Theories of dependence in political science have been developing rapidly in recent years. Example findings include that there is a strong tendency towards rapid retaliation in the issuance of international economic sanctions ([Cranmer, Heinrich and Desmarais, 2014](#)), the finding that bilateral preferential trade agreement networks form triads and/or four-cycles to share costs among larger partner groups ([Milewicz et al., 2018](#)), and influence networks among policy actors are hierarchical—exhibiting a tendency towards transitive tie formation ([Christopoulos and Ingold, 2015](#)). The ERGM opens up a new class of hypotheses that can be tested relative to regression models. In addition to studying how covariates effect the absence/presence of edges, the ERGM permits researchers to test hypotheses about how the edges effect each other (e.g., is the enemy of an enemy a friend)? The key components in using an ERGM are to develop hypotheses regarding the ways in which edges depend on each other, and conceptualize these dependencies in terms of subnetwork structures such as dyads and triads.

Under the ERGM, the probability (likelihood function) of observing the entire network \mathbf{Y} is

$$\mathcal{P}(\mathbf{Y}, \boldsymbol{\theta}) = \frac{\exp\{\boldsymbol{\theta}'\mathbf{Y}\}}{\sum_{\mathbf{Y}^* \in \mathcal{Y}} \exp\{\boldsymbol{\theta}'\mathbf{Y}^*\}},$$

where $\mathbf{h}(\mathbf{Y})$ is the vector of network statistics selected to specify the model, $\boldsymbol{\theta}$ is the vector of effects of the network statistics on the probability of observing a particular configuration of the network, $\exp\{\boldsymbol{\theta}'\mathbf{h}(\mathbf{Y})\}$ is the positive weight associated with the respective network configuration, \mathcal{Y} is the set of all possible network configurations, and $\sum_{\mathbf{Y}^* \in \mathcal{Y}} \exp\{\boldsymbol{\theta}'\mathbf{Y}^*\}$ is the normalizing constant that assures that the probability distribution over networks sums to one. The ERGM is specified through the selection of $\mathbf{h}(\mathbf{Y})$, and is very flexible. This flexibility comes at two costs when it comes to application. First, since \mathcal{Y} is so large— $2^{n*(n-1)}$, where n is the number of nodes in the network, for a directed network with no self-ties—the likelihood function cannot be computed exactly, which means estimation involves either approximation of the likelihood or method of moments criteria ([He and Zheng, 2016](#); [Hummel, Hunter and Handcock, 2012](#);

Schmid and Desmarais, 2017). Second, the researcher may specify a model that cannot be effectively fit to the observed data, the consequence of which is model degeneracy. Model degeneracy is a condition under which the model does not converge to a model that can effectively approximate even the number of edges in the network, the result of which is a model that places nearly all of the probability on either the completely full or completely empty network (Schweinberger, 2011). An active methodological literature has made substantial progress towards solving these two problems (e.g., DeMuse, Larcomb and Yin, 2018), and state-of-the-art open-source software implements many of these solutions (Hunter et al., 2008).

The ERGM has been used to study several different forms of dependence in political networks—structural patterns that go beyond the effects of covariates, of which ERGM is uniquely capable of modeling. These dependence patterns concern the ways in which ties in the network are related to each other, and are modeled through the specification of corresponding network statistics in $\mathbf{h}(\mathbf{Y})$. Peng et al. (2016) apply ERGM to study reciprocity in the online communication networks among members of the United States Congress. Osei and Malang (2018) use ERGM to analyze the degree to which political discussion is characterized by triadic closure (i.e., the friend of a friend is a friend) among members of the Ghanaian legislature. Gerber, Henry and Lubell (2013) use ERGMs to study the degree to which ties beget ties (i.e., tie formation is driven by popularity dynamics) in organizational collaboration networks surrounding local regional planning in California. Guerrero, Mcallister and Wilson (2015) analyze a multilevel collaboration network among organizations working on agricultural conservation in Australia. The levels include individual property owners, sub-regional organizations (e.g., local governments) and supra-regional organizations (e.g., NGOs). They use ERGM to study the prevalence of nodes that bridge multiple levels in the network. Dependence patterns such as the ones cited in this paragraph represent an exciting and important component of the function of political systems, and a class of structural patterns that can only be studied through the use of inferential network analysis.

2.3 The stochastic blockmodel

Concepts in political science related to group and system polarity (e.g., Baldassarri and Bearman, 2007; Cranmer, Menninga and Mucha, 2015) fall within the orbit of the SBM formulation. Under the SBM there is a finite number of node types or ‘blocks’, and each block is defined by the probabilities according to which nodes within the block form ties with each other, and with nodes in other blocks. If there are n nodes in the network, the SBM reduces the complexity of interactions among those n nodes, to interactions among and within a set of b blocks, where b is typically much smaller than n .

2.4 Additional methods in brief

The LSM, ERGM, and SBM reflect the core components

3 Application of ERGM, LSM, and SBM

[John]: [Background summary of Wojcik paper] For the application we replicate a recent study on the determining factors for when actors form directed connections in Brazilian legislative networks (Wojcik, 2017). The author argues that despite high levels of personalism in Latin American politics, political parties are equally effective in determining actor behavior, indicating that there are higher levels of party discipline than previously thought. By addressing this problem from a network framework, the author offers new evidence that the relationships between legislators is interdependent.

[John]: [summary of replication/extension] In addition to replicating the ERGM models in the article, we also fit the three models using latent space models and stochastic block models. We use the R software

package `ergm` to replicate the ERGM models, `latentnet` for the LSMs, and `CIDnetworks` for the SBMs. We fit models using 2 to 5 blocks for the SBMs and report the model results that has the lowest DIC. In general we find that results from all the models match the general conclusions of the paper. Below I explain the data from the study, the model specifications, and the results in more detail.

3.1 Data

[John]: [One paragraph reviewing how wojcik gathered dependent variable network] To collect the network data the author distributed an interactive digital survey to office managers from the Brazilian equivalent of US House of Representatives. The three different types of connections surveyed were communication on legislative issues, socialization about non-legislative issues, and information gathering to resolve questions about legislative proposals and procedures. The author also collected biographic information on those that completed the survey to look for any identifiable biases and compared traits of sample to that of the entire Congress, finding no identifiable differences. The survey had a 25% response rate and missing data was filled in using results from a pen and paper survey from 2013.

[John]: [One paragraph providing descriptive statistics on the dependent variable network, perhaps include a network plot]

Table 1. Summary Statistics: Dependent Variables

Network	Measurement	Global Value	Mean	s.d.	Max
Communication	Edge Density	0.008	NA	NA	NA
Communication	Degree Distribution	NA	3.307	3.020	18.000
Communication	Centrality	0.072	204.588	448.319	3009.852
Communication	Reciprocity	0.255	NA	NA	NA
Communication	Transitivity	0.108	NA	NA	NA
Social	Edge Density	0.006	NA	NA	NA
Social	Degree Distribution	NA	2.482	2.254	11.000
Social	Centrality	0.025	103.633	222.925	1009.333
Social	Reciprocity	0.203	NA	NA	NA
Social	Transitivity	0.119	NA	NA	NA
Info Exchange	Edge Density	0.009	NA	NA	NA
Info Exchange	Degree Distribution	NA	-0.348	-0.391	6.000
Info Exchange	Centrality	0.001	0.880	1.930	10.000
Info Exchange	Reciprocity	0.157	NA	NA	NA
Info Exchange	Transitivity	0.135	NA	NA	NA

[John]: [One paragraph reviewing the independent variables in the analysis, perhaps include a table of descriptive data]

3.2 Hypotheses

The main hypothesis of the paper is that political parties actively work to form ties among members and thus counter the individualistic tendencies of politicians to increase their own electoral share/power at the expense of the party. The network then creates trust and cohesion among the actors, while also facilitating

Table 2. Summary Statistics: Independent Variables

Variable	Proportion		
	Communication	Social	Info Exchange
Same Education Level	71%	70%	74%
Same Floor	10%	10%	12%
Both Leadership	6%	7%	6%
Both non-Leadership	10%	10%	14%
Same State	10%	9%	12%
Same Party	46%	50%	43%
Average			
Δ Age	12	12	13

monitoring. The author expects that this will be the stronger driver of the tie formation, rather than individual traits.

The secondary hypothesis of the paper is that similar individual traits drive network ties, even if less so than party membership. The author includes five different traits. The first is that when politicians are both in leadership positions they will need to interact more, but when neither are in leadership positions they will not need to interact with one another and can rely on party leaders. The second is that when politicians have similar ages they will form connections based on shared experiences. The third is that when politicians are from the same state they will form connections since there is a proportional voting system in Brazil and that means that coordination among members in the same states is necessary to exchange information about what the local needs are. The fourth is that politicians with the same education level are likely to have similar knowledge and social class, which increases the chance of network ties. The fifth is that when politicians work on the same floor this creates opportunity for interaction.

3.3 Replication design

[John]: [One paragraph reviewing the model specification in the original paper, explaining the transitive ties term]

The ERGMs are specified in the article to include a intercept term, the exogenous covariates, and a term for transitive ties. Transitive ties is a network statistic, specific to network models and is defined as the change in the likelihood of a tie between two nodes given the presence of a two-path or a triad closure. A two-path between nodes i and j is defined as node i having a directed edge path to node k and node k having a directed edge to node j , or put differently, there exists a connection between the nodes with only two edges.

[John]: [One paragraph describing how we extend the application to use both the LSM and the SBM]

For the LSMs, the specification is the same except it does not include the include a transitive ties term. This is because the model does not have network statistics as an ERGM model does, but instead we specify the latent space that models higher order dependences, which would include transitivity. For our extensions we use the euclidean distance model term that is equal to the negative euclidean distance between actors in the unobserved social space. We estimated the term along two dimensions and across two or three clusters, depending on best model fit using the BIC. The SBMs do not include a constant term as the block or group effect takes the place of the intercept term. It also does not have network terms

like transitive ties as it accounts for dependencies using the latent space along a multinomial distribution to determine block membership for the nodes.

3.4 Results

[John]: [1–2 paragraphs discussing covariate results that are consistent and/or different across the ERGM LSM, and SBM.]

The results from the LSMs indicate that the latent space explains less of the variation in the ties than the network statistics in ERGMs, as the coefficients for exogenous covariates, while signed the same, are mostly larger in magnitude. One exception is the node-match variables for leadership, where in model 1 the effect size is slightly larger, but in model 3 the effect size is nearly one-third smaller. Also, the changes in magnitude are generally smaller for the leadership variables than the other variables.

The results from the SBMs indicate that the latent space and block membership explains more of the variation in the ties than the network statistics in ERGMs, as the coefficients for exogenous covariates, while signed the same, are mostly smaller in magnitude. Also, the coefficient estimates have smaller standard errors relative to the size of the estimate.

[John]: [1–2 paragraphs discussing the latent structure in the LSM and SBM (i.e., visuals of the LSM estimates, distribution of nodes by blocks in the SBMs)]

The latent space plots (Tables 6 to 8) show that the clusters are not distinct from one another, but rather there are inside groups that form clusters and each subsequent group are nodes that are less central to the graph. In Table 9 for the SBM for communication ties, we see that the blocks are fairly evenly distributed, but that membership not predominant for each group and that many of the nodes have spread out likelihoods for membership assignment. In Table 10 we see that this shared membership is even greater for the social networks. However, for the information network in Table 11 membership likelihood is sharply separated and concentrated in one block for each node.

	Model 1	Model 2	Model 3
Edges (intercept)	−5.670*** (0.162)	−6.070*** (0.185)	−5.742*** (0.445)
Δ Age	−0.021*** (0.007)	−0.024*** (0.008)	−0.063*** (0.022)
Same Education Level	0.164 (0.125)	0.161 (0.144)	0.600* (0.334)
Same Floor	0.756*** (0.153)	1.373*** (0.141)	0.794** (0.390)
Both non-leadership	−0.655*** (0.125)	−0.241* (0.139)	−0.805** (0.337)
Both leadership	0.534*** (0.164)	0.473** (0.206)	0.837** (0.396)
Same State	2.121*** (0.124)	2.106*** (0.140)	0.989** (0.431)
Same Party	2.319*** (0.123)	2.051*** (0.135)	2.958*** (0.343)
Transitive ties	0.661*** (0.123)	0.901*** (0.139)	0.568 (0.399)
Akaike Inf. Crit.	3,056.153	2,281.901	466.732
Bayesian Inf. Crit.	3,133.387	2,358.395	526.326
N	39,402	36,290	5550

Note: *p<0.1; **p<0.05; ***p<0.01

Table 3. ERGM results

	Model 1	Model 2	Model 3
Edges (intercept)	−3.10*** (0.094)	−3.00*** (0.291)	−3.12*** (0.620)
Δ Age	−0.03*** (0.008)	−0.04*** (0.010)	−0.08*** (0.026)
Same Education Level	0.22*** (0.031)	0.35*** (0.084)	0.93** (0.433)
Same Floor	1.02*** (0.148)	1.53*** (0.010)	1.16** (0.490)
Both non-leadership	−0.82*** (0.054)	−0.28** (0.112)	−0.82** (0.418)
Both leadership	0.61*** (0.005)	0.44 (0.296)	0.64 (0.510)
Same State	2.75*** (0.105)	2.56*** (0.176)	1.31* (0.671)
Same Party	2.86*** (0.112)	2.30*** (0.033)	3.47*** (0.406)
BIC (Overall)	3589.51	2885.59	770.67
BIC (Likelihood)	2544.96	1879.84	376.27
BIC (Latent Positions)	1044.54	1005.75	394.40
d = Latent Dimensions / G = n Clusters	d=2, G=3	d=2, G=3	d=2, G=2
N	39,402	36,290	5,550

Note: *p<0.1; **p<0.05; ***p<0.01

Table 4. LSM Results

	Model 1	Model 2	Model 3
Δ Age	−0.014*** (0.003)	−0.013*** (0.003)	−0.035*** (0.009)
Same Education Level	0.143** (0.057)	0.081 (0.070)	0.293** (0.150)
Same Floor	0.316*** (0.067)	0.632*** (0.067)	0.430** (0.168)
Both non-leadership	−0.277*** (0.053)	−0.131** (0.064)	−0.407*** (0.140)
Both leadership	0.212*** (0.073)	0.196** (0.099)	0.340** (0.174)
Same State	1.049*** (0.058)	1.036*** (0.067)	0.456** (0.201)
Same Party	1.085*** (0.053)	0.9871*** (0.065)	1.360*** (0.141)
Deviance Inf. Crit.	2,818.370	2164.976	460.616
N blocks	3	2	3
N	39,402	36,290	5,550

Note: * $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Table 5. SBM results

Model 1

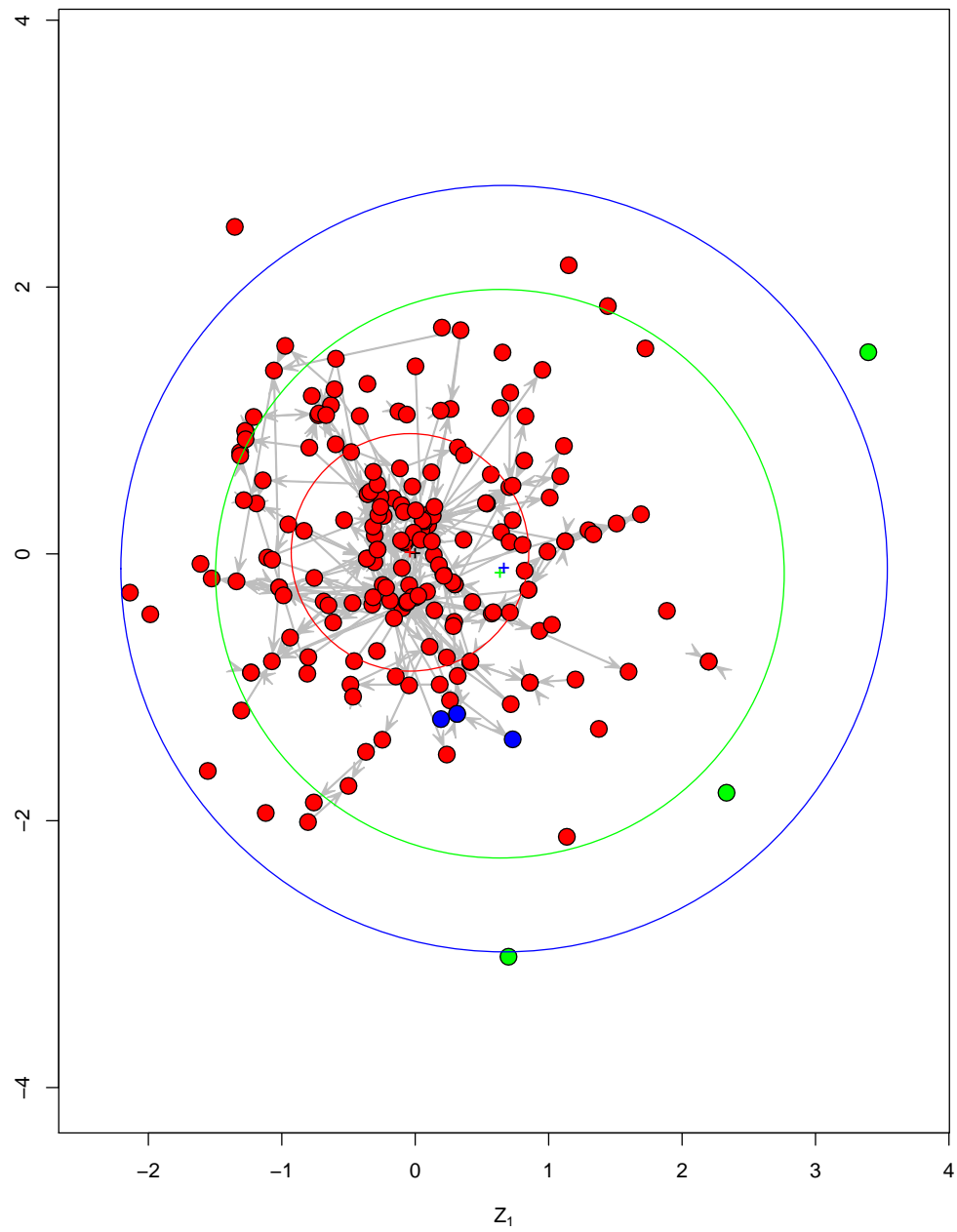


Table 6. Latent Space positions for LSM.

Model 2

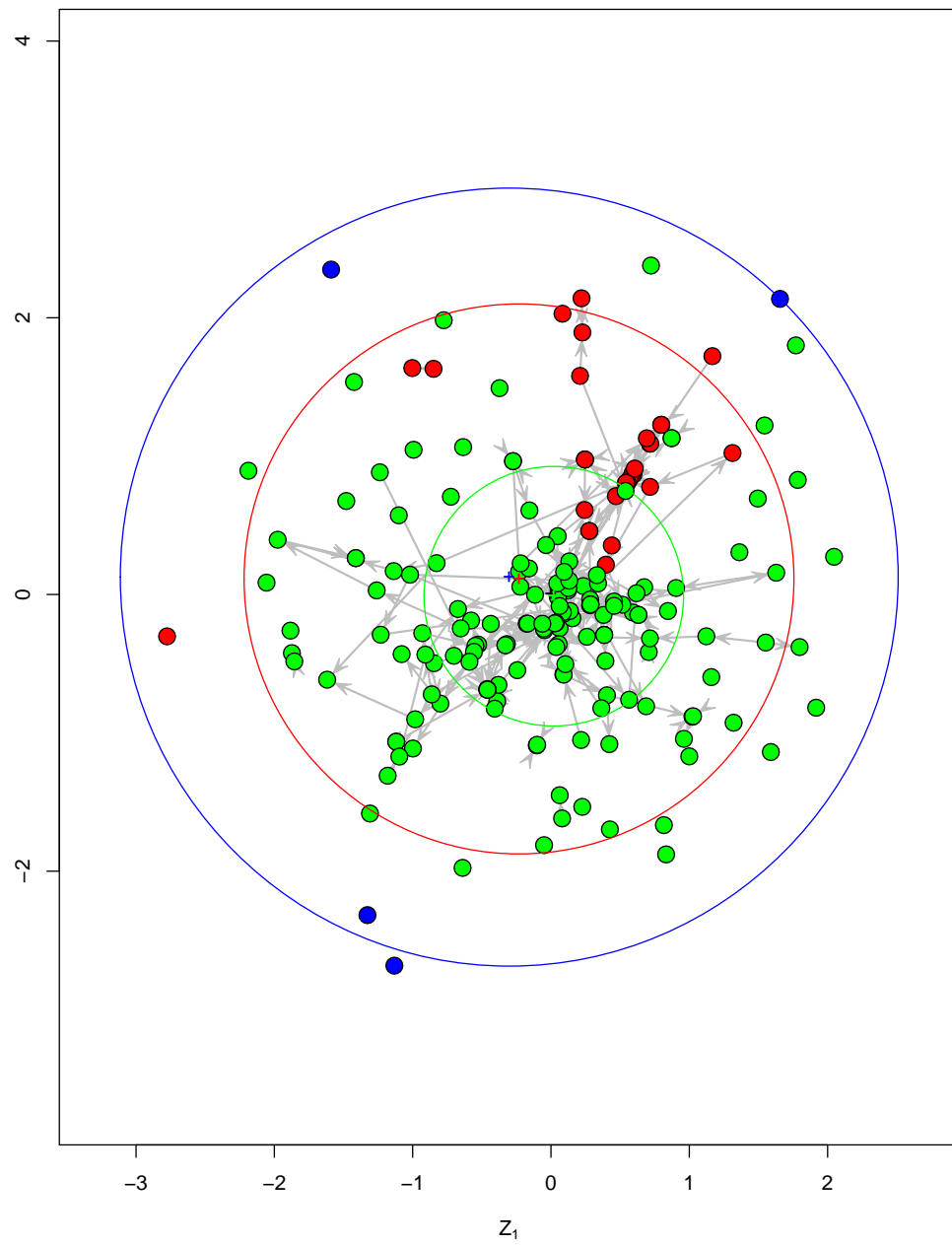


Table 7. Latent Space positions for LSM.

Model 3

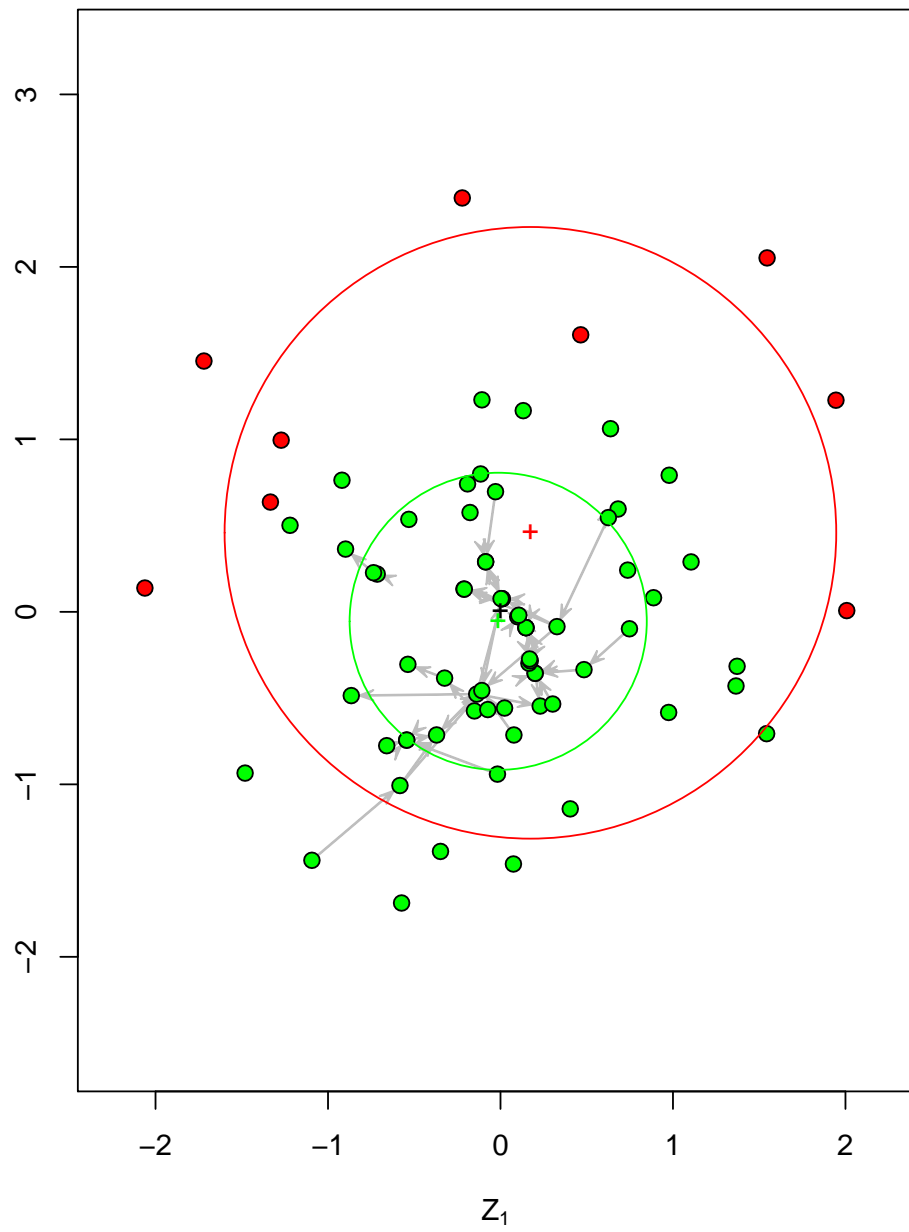


Table 8. Latent Space positions for LSM.

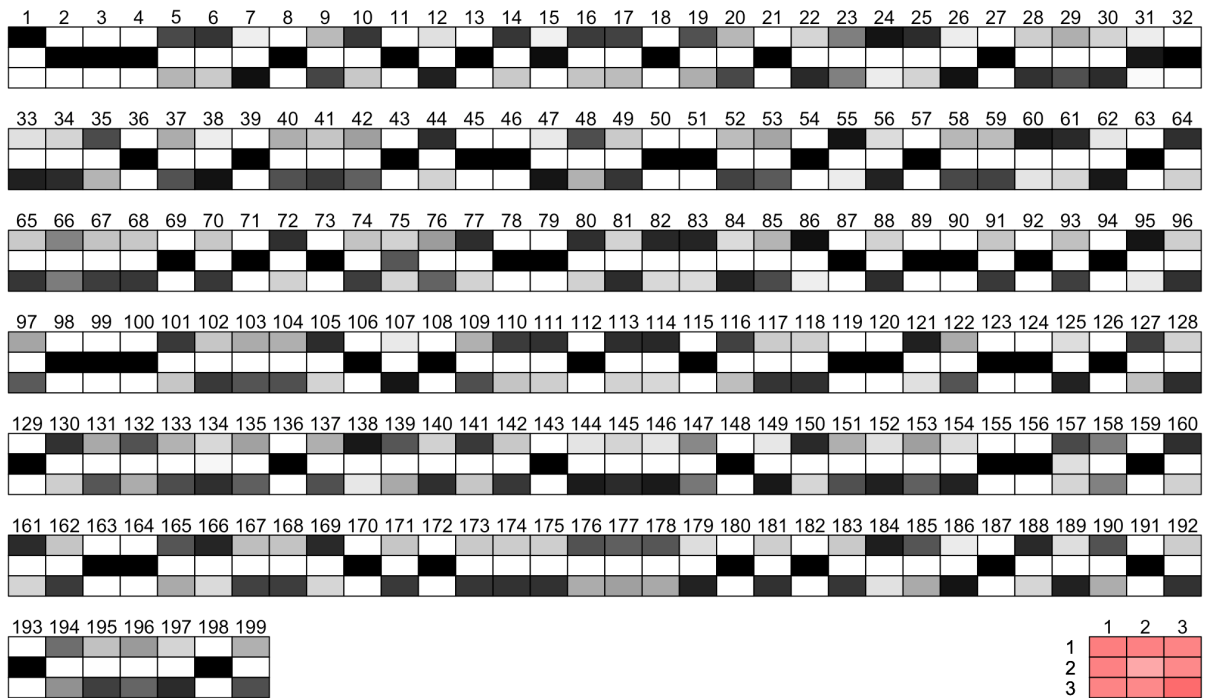


Table 9. Block positions for SBM 1.

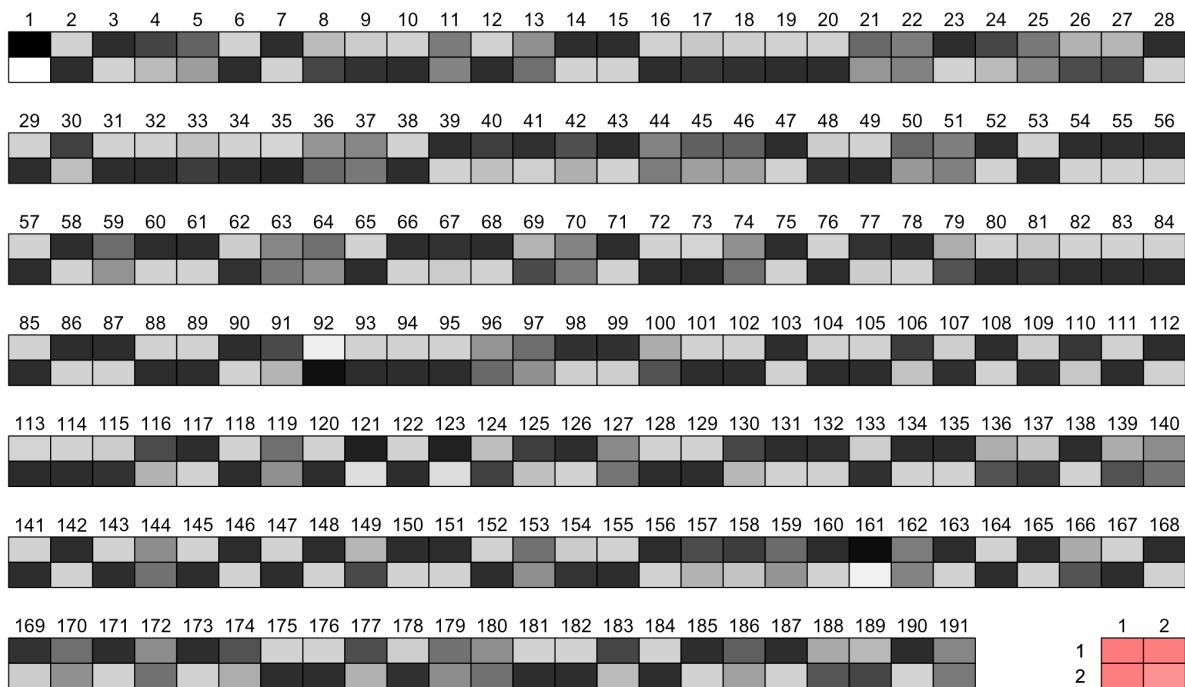


Table 10. Block positions for SBM 2.

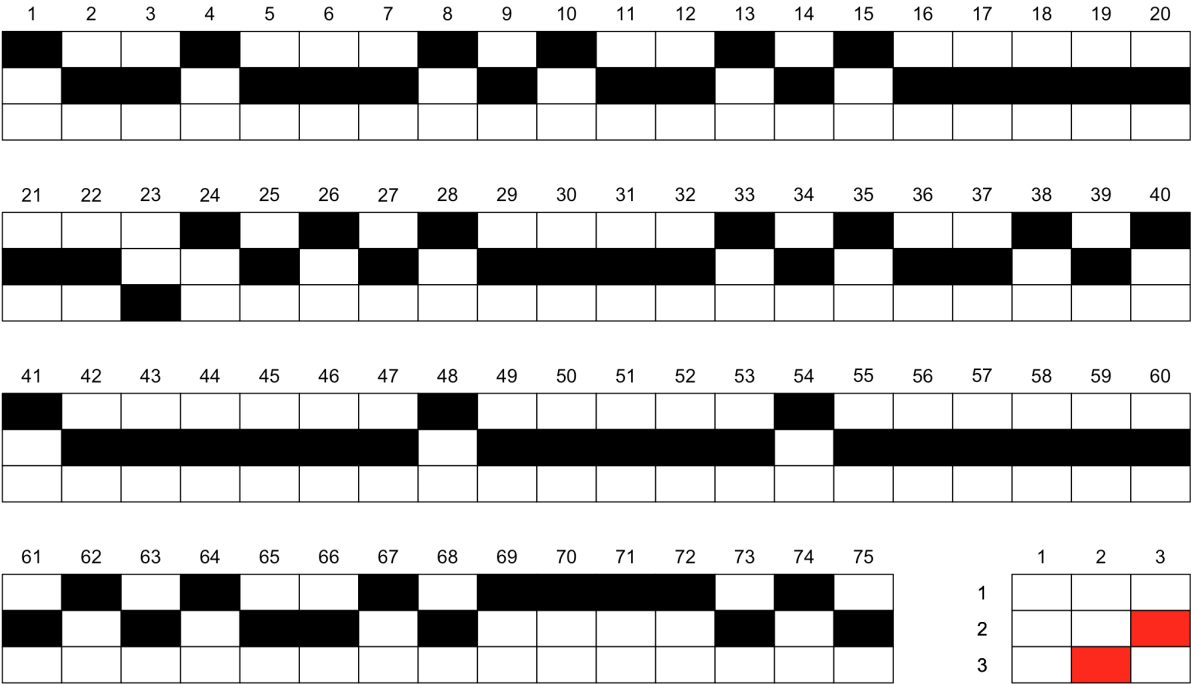


Table 11. Block positions for SBM 3.

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