# Yes, it matters who is spreading monkeypox: Supplemental Information

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### 1 Background

Using branching process theory, we assume the number of secondary cases produced from each infectious case is distributed according to a negative binomial distribution with mean  $R_0$  (the basic reproductive number) and dispersion parameter k. The parameter k quantifies the degree of individual heterogeneity by measuring overdispersion in the distribution (e.g., higher than expected variation). For a given  $R_0$ , smaller values of k (k << 1) suggest increased heterogeneity in the number of secondary cases between individual index cases. As k increases, transmission becomes more uniform and centered around the mean,  $R_0$ ; as k approaches infinity the negative binomial distribution converges to the Poisson distribution, wherein all differences in the number of secondary cases is attributed to demographic stochasticity.

Code to implement the included methods and recreate figure is available at the following GitHub repository:

https://github.com/jpsmithuga/Comment-Supplementary

## 2 Proportion of secondary cases attributed to the proportion of infectious cases

Lloyd-Smith et al. (in in supplementary information 2.2.5) propose a distribution for describing transmission based from the distribution of the individual reproduction number  $\nu$ :

$$F_{\rm trans}(x) = \frac{1}{R_0} \int_0^x u f_{\nu}(u) du$$

This gives the cumulative distribution function (CDF) in terms of the individual reproduction number density  $f_{\nu}$ , as the proportion of all transmission due to infectious individuals with reproduction number  $\nu < x$ . In the specific case when  $\nu$  is gamma distributed with shape parameter k > 0 and rate parameter  $k/R_0$ , where  $R_0$  is the mean (expected) individual reproduction number, we have

$$f_{\nu}(u) = \frac{k^k}{R_0^k \Gamma(k)} u^{k-1} e^{-ku/R_0}$$

It follows that

$$f_{\text{trans}}(x) = F'_{\text{trans}}(x) = \frac{1}{R_0} x f_{\nu}(x)$$

$$= \frac{k^k}{R_0^{k+1} \Gamma(k)} x^k e^{-kx/R_0}$$

$$= \frac{k^{k+1}}{R_0^{k+1} \Gamma(k+1)} x^k e^{-kx/R_0}$$

$$\sim \text{Gamma}(k+1, k/R_0)$$

The rate parameter remains the same and the shape parameter increases by 1. To calculate  $t_p$ , the expected proportion of transmission due to the most infectious 100p% of cases, we first find the (1-p)th centile of the individual reproduction number,  $x_p = F_{\nu}^{-1}(1-p)$ , then calculate  $t_p = 1 - F_{\text{trans}}(x_p)$ . As both random variables are gamma distributed, this is implemented in R via the qgamma  $(F^{-1})$  and pgamma (F) functions (see here).

# 3 The relationship between p and $t_p$

Parameterizing by  $x_p$ , and noting that  $p = 1 - F_{\nu}(x_p)$ , the first derivative is

$$\frac{dt_p}{dp} = \frac{dt_p/dx_p}{dp/dx_p} = \frac{-F'_{\text{trans}}(x_p)}{-F'_{\nu}(x_p)} = \frac{R_0^{-1}x_p f_{\nu}(x_p)}{f_{\nu}(x_p)} = \frac{x_p}{R_0} > 0$$

which confirms that  $t_p$  is an increasing function of p. When  $x_p = 0$ ,  $(p, t_p) = (1, 1)$  at which the point the derivative is zero (horizontal tangent). Similarly, when  $x_p = +\infty$ ,  $(p, t_p) = (0, 0)$  at which the point the derivative is infinite (vertical tangent).

The second derivative is negative implying concavity of the graph of  $t_p$  as a function of p.

$$\frac{d^2t_p}{dp^2} = \frac{d}{dx_p}\left(\frac{dt_p}{dp}\right) \times \frac{dx_p}{dp} = \frac{1}{R_0} \times \frac{1}{dp/dx_p} = -\frac{1}{R_0f_\nu(x_p)} < 0$$

These properties can be seen in Figure 1A.

#### 4 Final Outbreak Size

Let Y represent the sum of all cases in a transmission chain (including the index case and all secondary, tertiary, etc, cases resulting from that index case), resulting in a final outbreak size of Y. Using branching process theory, the distribution of the total number of secondary infections in a single transmission chain (including the index case at generation) can be defined by the probability generating function (pgf) yielding the implicit relationship (Yan 2008):

$$G_Y(s) = sG_Z((G_Y(s)))$$

We can treat this as  $G_Y(s) = \sum_{y=1}^{\infty} q_y s^y$ , where  $q_y = P(Y = y)$ , allowing us to extract the probability (Becker 1974; Blumberg 2014 (in Supplemental Text S1):

$$P(Y = y) = \frac{1}{y!} \frac{d_y G_Y(s)}{ds^y} \bigg|_{s=0}$$

Where Y = 1, 2, 3... Substituting the negative binomial distribution yields the following probability of a outbreak: (see Blumberg 2012 (R < 1), Nishiura 2012

 $(R \ge 1)$ , and Smith 2022 (multiple index cases, n, where  $n \ge 1$ ))

$$P(Y=y) = \left(\frac{n}{y}\right) \frac{\Gamma(ky+y-n)}{\Gamma(ky)(y-n)!} \frac{\left(\frac{R_0}{k}\right)^{y-n}}{\left(1 + \frac{R_0}{k}\right)^{ky+y-n}}$$

Since when R>1 we have  $P(Y=\infty|n)>0$  the distribution is improper for R>1, though becomes proper for an R>1 when we define  $P(Y=\infty|n)=1-q(\theta)^s$ , where  $q(\theta)^s$  is the probability of extinction (Nishiura 2012; Farrington 2003). The this is implemented in R software by exponentiation of the log-likelihood (for computational ease).

#### 5 References

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