

Introduction to Term Structure Models

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Introduction to Term Structure Models

Modeling dynamic term structures serves as a practical and indispensable tool in the realm of finance. It enables investors, institutions, and policymakers to make informed decisions, manage risk effectively, and allocate resources wisely. By understanding how interest rates and yields evolve over time, these models offer a clear lens through which to assess market trends and price financial instruments accurately.

This course has been developed by Jean-Paul Renne and Alain Monfort. It is illustrated by R codes using various packages that can be obtained from CRAN. This `TSMmodels` package is available on GitHub. To install it, one needs to employ the `devtools` library:

```
install.packages("devtools") # in case this library has not been loaded yet
library(devtools)
install_github("jrenne/TSMmodels")
library(AEC)
```

Useful (R) links:

- Download R:
 - R software: <https://cran.r-project.org> (the basic R software)
 - RStudio: <https://www.rstudio.com> (a convenient R editor)
- Tutorials:

- Rstudio: <https://dss.princeton.edu/training/RStudio101.pdf> (by Oscar Torres-Reyna)
- R: https://cran.r-project.org/doc/contrib/Paradis-rdebuts_en.pdf (by Emmanuel Paradis)
- My own tutorial: https://jrenne.shinyapps.io/Rtuto_publiShiny/

Chapter 1

Affine processes

1.1 Information in the Economy: The “factors”

On each date $t = 1, 2, \dots, T$, agents receive new information by observing *factors*, also called *states*. We denote the (K -dimensional) vector of factors by w_t . Vector w_t is usually random. On date t , vector w_t is supposed to be perfectly observed by the agents (investors), but can be only partially observed, or unobserved by the econometrician.

Naturally, w_t can be decomposed into different sub-vectors of different natures. For instance, we can have $w_t = (y'_t, z'_t)'$ with y_t : observable vector of (geometric) returns, z_t : regime, unobserved by the econometrician.

Some of the components of w_t can be prices. For instance, one component could be a short-term rate, a stock return, or an exchange rate. It can also include macroeconomic variables (inflation, GDP growth), or agent-specific variables.

1.2 Dynamic models and Laplace transform (L.T.)

The objective of a dynamic model is to describe the random changes in w_t . The dynamics can be historical or risk-neutral (see Chapter ??). The dynamics we consider are parametric, in the sense that the conditional distribution $w_{t+1}|\underline{w}_t$ (with $\underline{w}_t = \{w_t, w_{t-1}, \dots\}$) depends on a vector of parameters θ . In practice, it may be the case that θ is unknown by the econometrician (see Chapter ??). The choice (or estimation) of a conditional distribution is equivalent to the choice (or estimation) of a conditional *Laplace transforms*:

$$\varphi(u|\underline{w}_t, \theta) = \mathbb{E}_\theta[\exp(u'w_{t+1})|\underline{w}_t], \quad u \in \mathbb{R}^K, \quad (1.1)$$

or a conditional *log Laplace transforms*:

$$\psi(u|\underline{w}_t, \theta) = \log\{\mathbb{E}_\theta[\exp(u'w_{t+1})|\underline{w}_t]\}, \quad u \in \mathbb{R}^K.$$

Example 1.1 (Conditionally Bernoulli process). If $w_{t+1}|\underline{w}_t \sim \mathcal{I}[p(\underline{w}_t, \theta)]$, then:

$$\varphi(u|w_t) = \mathbb{E}[\exp(uw_{t+1}) | \underline{w}_t] = p_t \exp(u) + 1 - p_t$$

with $p_t = p(\underline{w}_t, \theta)$.

Example 1.2 (Conditionally Binomial process). If $w_{t+1}|\underline{w}_t \in \mathcal{B}(n, p_t)$, then:

$$\varphi(u|w_t) = [p_t \exp(u) + 1 - p_t]^n.$$

Example 1.3 (Conditionally Poisson process). If $w_{t+1}|\underline{w}_t \sim \mathcal{P}(\lambda_t)$, then:

$$\begin{aligned} \varphi(u|w_t) &= \sum_{j=0}^{\infty} \frac{1}{j!} \exp(-\lambda_t) \lambda_t^j \exp(uj) = \exp(-\lambda_t) \exp[\lambda_t \exp(u)] \\ &= \exp\{\lambda_t [\exp(u) - 1]\}. \end{aligned}$$

Example 1.4 (Conditionally normal (or Gaussian) process). If $w_{t+1}|\underline{w}_t \sim \mathcal{N}(m(\underline{w}_t, \theta), \Sigma(\underline{w}_t, \theta))$, then:

$$u'w_{t+1}|\underline{w}_t \sim \mathcal{N}(u'm(\underline{w}_t, \theta), u'\Sigma(\underline{w}_t, \theta)u).$$

$$\Rightarrow \begin{cases} \varphi(u|\underline{w}_t, \theta) &= \exp \left[u'm(\underline{w}_t, \theta) + \frac{1}{2}u'\Sigma(\underline{w}_t, \theta)u \right] \\ \psi(u|\underline{w}_t, \theta) &= u'm(\underline{w}_t, \theta) + \frac{1}{2}u'\Sigma(\underline{w}_t, \theta)u. \end{cases}$$

1.3 Laplace Transform and moments/cumulants

Here are some properties of the Laplace transform (Eq. (1.1)):

- $\varphi(0|\underline{w}_t, \theta) = 1$ and $\psi(0|\underline{w}_t, \theta) = 0$.
- It is defined in a convex set E (containing 0).
- If the interior of E is non empty, all the (conditional) moments exist.

As mentioned above, knowing the (conditional) Laplace transform is equivalent to knowing the (conditional) moments—if they exist. In the scalar case, we have that:

- the moment of order n closely relates to the n^{th} derivatives of φ :

$$\left[\frac{\partial^n \varphi(u|\underline{w}_t, \theta)}{\partial u^n} \right]_{u=0} = \mathbb{E}_\theta[w_{t+1}^n | \underline{w}_t],$$

- the cumulant of order n closely relates to the n^{th} derivatives of ψ :

$$\left[\frac{\partial^n \psi(u|\underline{w}_t, \theta)}{\partial u^n} \right]_{u=0} = K_n(\underline{w}_t, \theta).$$

In particular, what precedes implies that:

$$\begin{cases} K_1(\underline{w}_t, \theta) &= \mathbb{E}_\theta[w_{t+1} | \underline{w}_t] \\ K_2(\underline{w}_t, \theta) &= \mathbb{V}ar_\theta[w_{t+1} | \underline{w}_t]. \end{cases}$$

Accordingly, φ and ψ are respectively called conditional **moment** and **cumulant** generating function.

In the multivariate case, we have:

$$\begin{aligned} \left[\frac{\partial \psi}{\partial u}(u|\underline{w}_t, \theta) \right]_{u=0} &= \mathbb{E}_\theta[w_{t+1} | \underline{w}_t] \\ \left[\frac{\partial^2 \psi}{\partial u \partial u'}(u|\underline{w}_t, \theta) \right]_{u=0} &= \mathbb{V}ar_\theta[w_{t+1} | \underline{w}_t]. \end{aligned}$$

Example 1.5 (Conditionally normal (or Gaussian) process). Consider Example 1.5. Applying the previous formula, we have, in the scalar case:

- $\psi(u|\underline{w}_t, \theta) = um(\underline{w}_t, \theta) + \frac{1}{2}u^2\sigma^2(\underline{w}_t, \theta).$
- $\left[\frac{\partial \psi}{\partial u}(u|\underline{w}_t, \theta) \right]_{u=0} = m(\underline{w}_t, \theta).$
- $\left[\frac{\partial^2 \psi}{\partial u^2}(u|\underline{w}_t, \theta) \right]_{u=0} = \sigma^2(\underline{w}_t, \theta).$

and, in the multidimensional normal case:

- $\psi(u|\underline{w}_t, \theta) = u'm(\underline{w}_t, \theta) + \frac{1}{2}u'\Sigma(\underline{w}_t, \theta)u.$
- $\left[\frac{\partial \psi}{\partial u}(u|\underline{w}_t, \theta) \right]_{u=0} = m(\underline{w}_t, \theta).$
- $\left[\frac{\partial^2 \psi}{\partial u \partial u'}(u|\underline{w}_t, \theta) \right]_{u=0} = \Sigma(\underline{w}_t, \theta).$

In both cases, cumulants of order > 2 equal to 0.

1.4 Additional properties of the Laplace transform

Here are additional properties of multivariate Laplace transform:

- If $w_t = (w'_{1t}, w'_{2t})'$, $u = (u'_1, u'_2)'$:

$$\begin{aligned} \mathbb{E}_\theta[\exp(u'_1 w_{1,t+1} | \underline{w}_t)] &= \varphi(u_1, 0 | \underline{w}_t, \theta) \\ \mathbb{E}_\theta[\exp(u'_2 w_{2,t+1} | \underline{w}_t)] &= \varphi(0, u_2 | \underline{w}_t, \theta). \end{aligned}$$

- If $w_t = (w'_{1t}, w'_{2t})'$, and if w_{1t} and w_{2t} are conditionally independent:

$$\begin{aligned} \varphi(u | \underline{w}_t, \theta) &= \varphi(u_1, 0 | \underline{w}_t, \theta) \times \varphi(0, u_2 | \underline{w}_t, \theta) \\ \psi(u | \underline{w}_t, \theta) &= \psi(u_1, 0 | \underline{w}_t, \theta) + \psi(0, u_2 | \underline{w}_t, \theta). \end{aligned}$$

- If w_{1t} and w_{2t} have the same size and if

$$\varphi(u_1, u_2 | \underline{w}_t, \theta) = \mathbb{E}_\theta[\exp(u'_1 w_{1,t+1} + u'_2 w_{2,t+1} | \underline{w}_t),$$

1.4. ADDITIONAL PROPERTIES OF THE LAPLACE TRANSFORM 13

then the conditional Laplace transform of $w_{1,t+1} + w_{2,t+1}$ given \underline{w}_t is $\varphi(u, u|\underline{w}_t, \theta)$. In particular, if w_{1t} and w_{2t} are conditionally independent and have the same size, the conditional Laplace transform and Log-Laplace transform of $w_{1,t+1} + w_{2,t+1}$ are respectively:

$$\varphi(u, 0|\underline{w}_t, \theta) \times \varphi(0, u|\underline{w}_t, \theta), \quad \text{and} \quad \psi(u, 0|\underline{w}_t, \theta) + \psi(0, u|\underline{w}_t, \theta).$$

Lemma 1.1 (Conditional zero probability for non-negative processes). *If w_t is univariate and nonnegative its (conditional) Laplace transform $\varphi_t(u) = \mathbb{E}_t[\exp(uw_{t+1})]$ is defined for $u \leq 0$ and*

$$\mathbb{P}_t(w_{t+1} = 0) = \lim_{u \rightarrow -\infty} \varphi_t(u).$$

Proof. We have $\varphi_t(u) = \mathbb{P}_t(w_{t+1} = 0) + \int_{w_{t+1} > 0} \exp(uw_{t+1}) d\mathbb{P}_t(w_{t+1})$. The Lebesgue theorem ensures that the last integral converges to zero when u goes to $-\infty$. \square

Lemma 1.2 (Conditional zero probability for non-negative multivariate processes). *Assume that:*

- $w_{1,t}$ is valued in \mathbb{R}^d ($d \geq 1$),
- $w_{2,t}$ is valued in $\mathbb{R}^+ = [0, +\infty)$,
- $\mathbb{E}_t[\exp(u'_1 w_{1,t+1} + u_2 w_{2,t+1})]$ exists for a given u_1 and $u_2 \leq 0$.

Then, we have:

$$\mathbb{E}_t[\exp(u'_1 w_{1,t+1}) \mathbf{1}_{\{w_{2,t+1}=0\}}] = \lim_{u_2 \rightarrow -\infty} \mathbb{E}_t[\exp(u'_1 w_{1,t+1} + u_2 w_{2,t+1})]. \quad (1.2)$$

Proof. We have that

$$\begin{aligned} & \lim_{u_2 \rightarrow -\infty} \mathbb{E}_t[\exp(u'_1 w_{1,t+1} + u_2 w_{2,t+1})] \\ &= \mathbb{E}_t[\exp(u'_1 w_{1,t+1}) \mathbf{1}_{\{w_{2,t+1}=0\}}] + \\ & \quad \lim_{u_2 \rightarrow -\infty} \mathbb{E}_t[\exp(u'_1 w_{1,t+1} + u_2 w_{2,t+1}) \mathbf{1}_{\{w_{2,t+1}>0\}}], \end{aligned}$$

and since in the second term on the right-hand side $\exp(u_2 w_{2,t+1}) \mathbf{1}_{\{w_{2,t+1}>0\}} \rightarrow 0$ when $u_2 \rightarrow -\infty$, Eq. (1.2) is a consequence of the Lebesgue theorem. \square

1.5 Affine (or Car) processes

In term structure applications, we will often consider *affine* processes (Definitions 1.1 and 1.2). These processes are indeed such that their multi-horizon Laplace transform are simple to compute (Lemma 1.5 and Proposition 1.5), which is key to compute bond prices.

1.5.1 Car processes of order one

Here is the definition of a *compound auto-regressive (Car)* process of order one:

Definition 1.1 (Affine process of order 1). A multivariate process w_{t+1} is Affine of order 1 [or $Car(1)$] if

$$\varphi_t(u) = \mathbb{E}_t[\exp(u'w_{t+1})] = \exp[a(u)'w_t + b(u)]$$

for some functions $a(\cdot)$ and $b(\cdot)$. These functions are univariate if w_{t+1} (and therefore u) is scalar.

Note that $a(\cdot)$ and $b(\cdot)$ may be deterministic functions of time (e.g., Chikhani and Renne (2023)).

A first key example is that of the Gaussian auto-regressive processes:

Example 1.6 (Univariate AR(1) Gaussian process). If $w_{t+1}|\underline{w}_t \sim \mathcal{N}(\nu + \rho w_t, \sigma^2)$, then:

$$\varphi_t(u) = \exp\left(u\rho w_t + u\nu + u^2\frac{\sigma^2}{2}\right) = \exp[a(u)'w_t + b(u)],$$

$$\text{with } \begin{cases} a(u) &= u\rho \\ b(u) &= u\nu + u^2\frac{\sigma^2}{2}. \end{cases}$$

Example 1.7 (Gaussian VAR). If $w_{t+1}|\underline{w}_t \sim \mathcal{N}(\mu + \Phi w_t, \Sigma)$, then:

$$\varphi_t(u) = \exp\left(u'(\mu + \Phi w_t) + \frac{1}{2}u'\Sigma u\right) = \exp[a(u)'w_t + b(u)],$$

$$\text{with } \begin{cases} a(u) &= \Phi'u \\ b(u) &= u'\mu + \frac{1}{2}u'\Sigma u = u'\mu + \frac{1}{2}(u \otimes u)'vec(\Sigma). \end{cases}$$

Example 1.8 (Quadratic Gaussian process). Consider vector $w_t = (x'_t, \text{vec}(x_t x'_t))'$, where x_t is a n -dimensional vector following a Gaussian $\text{VAR}(1)$, i.e.

$$x_{t+1} | w_t \sim \mathcal{N}(\mu + \Phi x_t, \Sigma).$$

Proposition 1.2 shows that if $u = (v, V)$ where $v \in \mathbb{R}^n$ and V a square symmetric matrix of size n , we have:

$$\begin{aligned} \varphi_t(u) &= \mathbb{E}_t \{ \exp [(v', \text{vec}(V)') \times w_{t+1}] \} \\ &= \exp \{ a_1(v, V)' x_t + \text{vec}(a_2(v, V))' \text{vec}(x'_t x_t) + b(v, V) \}, \end{aligned}$$

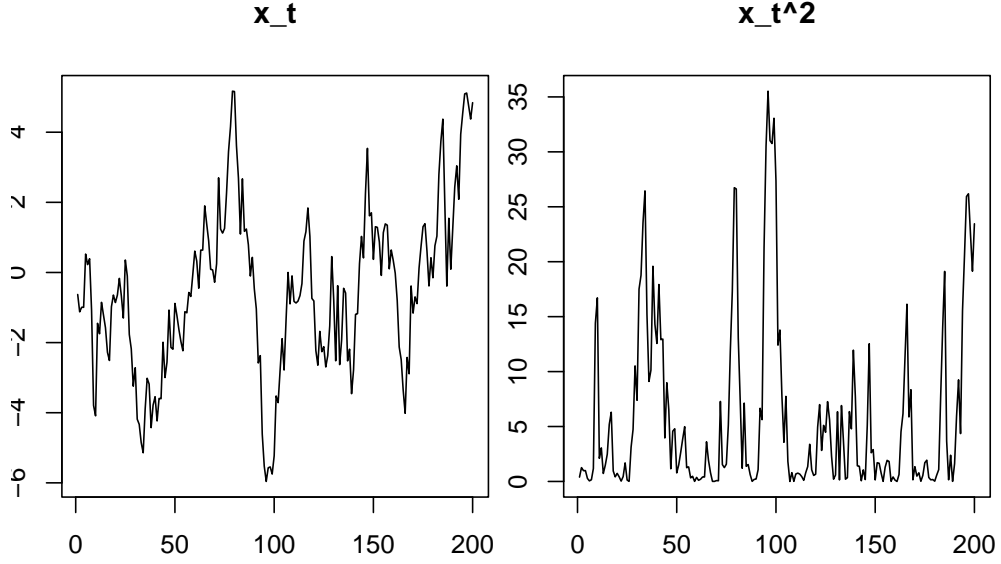
where:

$$\begin{aligned} a_2(u) &= \Phi' V (I_n - 2\Sigma V)^{-1} \Phi \\ a_1(u) &= \Phi' [(I_n - 2V\Sigma)^{-1} (v + 2V\mu)] \\ b(u) &= u' (I_n - 2\Sigma V)^{-1} \left(\mu + \frac{1}{2} \Sigma v \right) + \\ &\quad \mu' V (I_n - 2\Sigma V)^{-1} \mu - \frac{1}{2} \log |I_n - 2\Sigma V|. \end{aligned}$$

Quadratic processes can be used to construct positive process. Indeed, one can determine linear combinations of the components of w_t ($\alpha' w_t$, say) that are such that $\alpha' w_t \geq 0$. For instance, if x_t is scalar, $\alpha' w_t = x_t^2$ if $\alpha = (0, 1)'$. This is illustrated by Figure 1.1.

```
T <- 200
phi <- .9; sigma <- 1
x.t <- 0; x <- NULL
for(t in 1:T){
  x.t <- phi*x.t + sigma*rnorm(1)
  x <- c(x, x.t)}
par(mfrow=c(1,2), plt=c(.1, .95, .15, .85))
plot(x, type="l", xlab="", ylab="", main="x_t")
plot(x^2, type="l", xlab="", ylab="", main="x_t^2")
```

Another example of nonnegative process is that of the auto-regressive Gamma process (Gourieroux and Jasiak, 2006) and its extension (Monfort et al., 2017).

Figure 1.1: Simulation of a quadratic processes x_t .

Example 1.9 (Autoregressive gamma process, ARG(1)). An ARG process is defined as follows:

$$\frac{w_{t+1}}{\mu} \sim \gamma(\nu + z_t) \quad \text{where} \quad z_t \sim \mathcal{P}\left(\frac{\rho w_t}{\mu}\right),$$

with $\nu, \mu, \rho > 0$. (Alternatively $z_t \sim \mathcal{P}(\beta w_t)$, with $\rho = \beta\mu$.)

Proposition 1.3 shows that we have $\varphi_t(u) = \exp[a(u)'w_t + b(u)]$ with

$$\begin{cases} a(u) &= \frac{\rho u}{1 - u\mu} \\ b(u) &= -\nu \log(1 - u\mu). \end{cases}$$

One can simulate ARG processes by using this web-interface (select the “ARG” panel).

It can be shown that:

$$\begin{cases} \mathbb{E}(w_{t+1} | w_t) &= \nu\mu + \rho w_t \\ \text{Var}(w_{t+1} | w_t) &= \nu\mu^2 + 2\mu\rho w_t. \end{cases}$$

and that:

$$w_{t+1} = \nu\mu + \rho w_t + \varepsilon_{t+1},$$

where ε_{t+1} is a martingale difference $\Rightarrow w_{t+1}$ is a weak $AR(1)$.

Monfort et al. (2017) propose the extended ARG process and the ARG_0 process. The latter is such that $\nu = 0$ and βw_t is replaced with $\alpha + \beta w_t$, i.e.:

$$\frac{w_{t+1}}{\mu} \sim \gamma(z_t), \quad z_t \sim \mathcal{P}(\alpha + \beta w_t). \quad (1.3)$$

It is easily seen that we then have:

$$\varphi_t(u) = \exp \left[\frac{\beta \mu u}{1 - u \mu} w_t + \frac{\alpha \mu u}{1 - u \mu} \right].$$

The ARG_0 process features a point mass at zero, with conditional probability $\exp(-\alpha - \beta w_t)$. Note that 0 is absorbing if $\alpha = 0$.

Figure 1.2 displays the simulated path of an ARG_0 process (since we set ν to zero). Note that function `simul.ARG` is included in the `TSMODELS` package.

```
library(TSMODELS)
W <- simul.ARG(300,mu=.5,nu=0,rho=.9,alpha=.1)
plot(W,type="l")
```

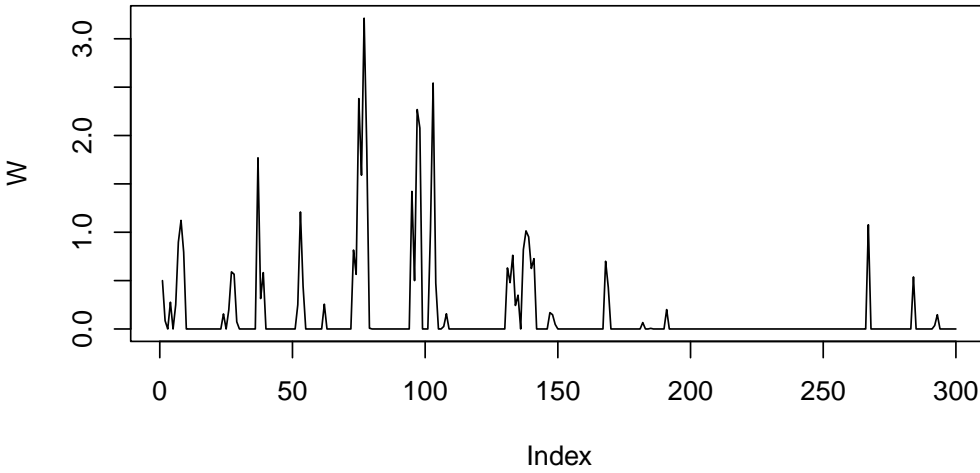


Figure 1.2: Simulation of an ARG_0 processes.

Certain affine processes are valued in specific sets (e.g., integers). It is the case of compound Poisson processes:

Example 1.10 (Compound Poisson process). A compound Poisson process is defined as follows (with $\gamma > 0$, $0 < \pi < 1$, and $\lambda > 0$):

$$\frac{w_{t+1}}{\gamma} = z_{t+1} + \varepsilon_{t+1},$$

where z_{t+1} and ε_{t+1} conditionally independent, and where $z_{t+1} \sim \mathcal{B}\left(\frac{w_t}{\gamma}, \pi\right)$, with $\varepsilon_{t+1} \sim \mathcal{P}(\lambda)$.

This process is valued in $\{j\gamma, j \in \mathbb{N}, \gamma \in \mathbb{R}^+\}$ and we have:

$$\varphi_t(u) = \exp \left\{ \frac{w_t}{\gamma} \log[\pi \exp(u\gamma) + 1 - \pi] - \lambda[1 - \exp(u\gamma)] \right\},$$

i.e., $\varphi_t(u) = \exp(a(u)w_t + b(u))$ with

$$\begin{cases} a(u) &= \frac{1}{\gamma} \log[\pi \exp(u\gamma) + 1 - \pi], \\ b(u) &= -\lambda[1 - \exp(u\gamma)]. \end{cases}$$

We also have: $w_{t+1} = \pi w_t + \lambda\gamma + \eta_{t+1}$, where η_{t+1} is a martingale difference.

One can simulate such processes by using this web-interface (select the “Compound Poisson” panel). Figure 1.3 makes use of function `simul.compound.poisson` (in package `TSMODELS`) to simulate a compound Poisson process.

```
library(TSMODELS)
W <- simul.compound.poisson(100, Gamma=.5, Pi=0.5, lambda=.9)
plot(W, type="l")
```

1.5.2 Car processes of order p

Let us now define Car processes of order p :

Definition 1.2 (Affine process of order p). A multivariate process w_{t+1} is affine of order p [or $Car(p)$] if there exist functions $a_1(\cdot), \dots, a_p(\cdot)$, and $b(\cdot)$ such that:

$$\varphi_t(u) = \mathbb{E}_t[\exp(u'w_{t+1})] = \exp[a_1(u)'w_t + \dots + a_p(u)'w_{t+1-p} + b(u)].$$

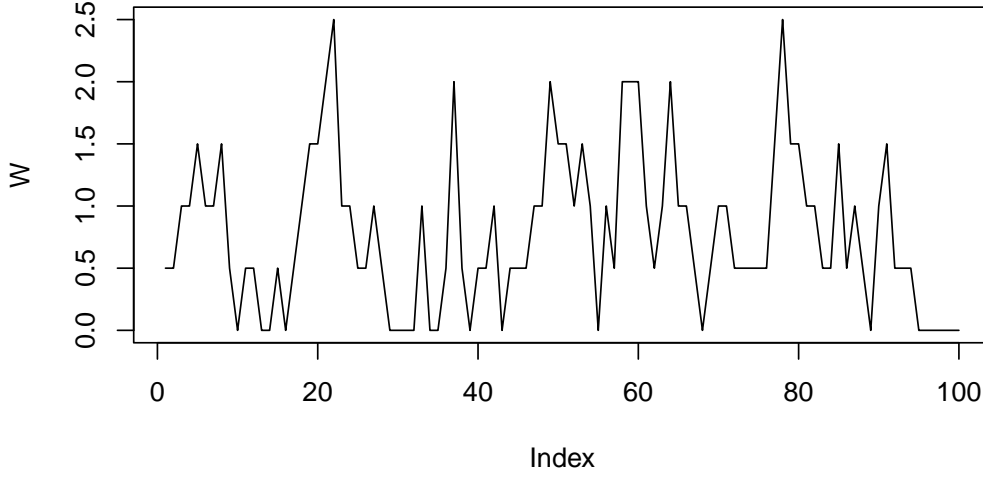


Figure 1.3: Simulation of a Compound Poisson process.

It can be seen that if w_t is $Car(p)$, then $W_t = [w'_t, w'_{t-1}, \dots, w'_{t-p+1}]'$ is $Car(1)$.¹ Therefore, without loss of generality we can assume $p = 1$.

The standard $Car(p)$ processes are auto-regressive processes of order p . These processes satisfy the definition of index affine processes:

Definition 1.3 (Univariate index affine process of order p). Let $\exp[a(u)w_t + b(u)]$ be the conditional Laplace transform of a univariate affine process of order 1, the process w_{t+1} is an *index-affine* process of order p if:

$$\varphi_t(u) = \mathbb{E}_t[\exp(uw_{t+1})] = \exp[a(u)(\beta_1 w_t + \dots + \beta_p w_{t+1-p}) + b(u)].$$

Examples 1.11 and 1.14 are two examples of index affine processes.

Example 1.11 (Gaussian AR(p) process). This example extends Example 1.6. Consider a Gaussian AR(p) process w_t ; that is:

$$w_{t+1} = \nu + \varphi_1 w_t + \dots + \varphi_p w_{t+1-p} + \sigma \varepsilon_{t+1}, \quad \varepsilon_{t+1} \sim i.i.d. \mathcal{N}(0, 1).$$

¹We have:

$$\begin{aligned} \mathbb{E}_t[\exp(u'W_{t+1})] &= \mathbb{E}_t[\exp(u'_1 w_{t+1} + u'_2 w_t + \dots + u'_p w_{t-p+2})] \\ &= \exp(u'_2 w_t + \dots + u'_p w_{t-p+2}) \mathbb{E}_t[\exp(u'_1 w_{t+1})] \\ &= \exp(u'_2 w_t + \dots + u'_p w_{t-p+2} + a'_1(u_1)w_t + \dots + a'_p(u_1)w_{t+1-p} + b(u_1)) \\ &= \exp[A(u)'W_t + B(u)], \end{aligned}$$

with $A(u)' = [u'_2 + a'_1(u_1), \dots, u'_p + a'_{p-1}(u_1), a'_p(u_1)]$ and $B(u) = b(u_1)$.

We have:

$$\varphi_t(u) = \exp \left[u\rho(\beta_1 w_t + \cdots + \beta_p w_{t+1-p}) + u\nu + u^2 \frac{\sigma^2}{2} \right]$$

with $\varphi_i = \rho\beta_i$.

Example 1.12 (ARG(p) process (positive)). This example extends Example 1.9. An ARG process of order p is defined as follows:

$$\frac{w_{t+1}}{\mu} \sim \gamma(\nu + z_t) \quad \text{where} \quad z_t \sim \mathcal{P}(\beta_1 w_t + \cdots + \beta_p w_{t+1-p}),$$

with $\nu, \mu, \beta_i > 0$. We have:

$$\varphi_t(u) = \exp \left[\frac{\rho u}{1 - u\mu} (\beta_1 w_t + \cdots + \beta_p w_{t+1-p}) - \nu \log(1 - u\mu) \right],$$

Process w_t admits the following AR(p) representation:

$$w_{t+1} = \nu\mu + \varphi_1 w_t + \cdots + \varphi_p w_{t+1-p} + \varepsilon_{t+1},$$

with $\varphi_i = \beta_i \mu$ and where ε_{t+1} is a martingale difference.

1.6 Markov chains

In this subsection, we show that the family of affine processes includes (some) regime-switching models. We consider a time-homogeneous Markov chain z_t , valued in the set of columns of Id_J , the identity matrix of dimension $J \times J$. The transition probabilities are denoted by $\pi(e_i, e_j)$, with $\pi(e_i, e_j) = \mathbb{P}(z_{t+1} = e_j | z_t = e_i)$. With these notations:

$$\mathbb{E}[\exp(v' z_{t+1}) | z_t = e_i, \underline{z_{t-1}}] = \sum_{j=1}^J \exp(v' e_j) \pi(e_i, e_j).$$

Hence, we have:

$$\varphi_t(v) = \exp[a_z(v)' z_t],$$

with

$$a_z(v) = \begin{bmatrix} \log \left(\sum_{j=1}^J \exp(v' e_j) \pi(e_1, e_j) \right) \\ \vdots \\ \log \left(\sum_{j=1}^J \exp(v' e_j) \pi(e_J, e_j) \right) \end{bmatrix}.$$

This proves that z_t is an affine process.

One can simulate a two-regime Markov chain by using this web-interface (select the “Markov-Switching” panel).

1.7 Wishart autoregressive (WAR) processes

WAR are *matrix processes*, valued in the space of $(L \times L)$ symmetric positive definite matrices.

Definition 1.4 (Wishart autoregressive (WAR) processes). Let W_{t+1} be a $WAR_L(K, M, \Omega)$ process. It is defined by:

$$\begin{aligned} & \mathbb{E}[\exp \text{Tr}(\Gamma W_{t+1}) | \underline{W}_t] \\ &= \exp \left\{ \text{Tr}[M' \Gamma (Id - 2\Omega \Gamma)^{-1} M W_t] - \frac{K}{2} \log[\det(Id - 2\Omega \Gamma)] \right\}, \end{aligned} \quad (1.4)$$

where Γ is a symmetric matrix,² K is a positive scalar, M is a $(L \times L)$ matrix, and Ω is a $(L \times L)$ symmetric positive definite matrix.

If K is an integer, Proposition 1.4 (in the appendix) shows that W_{t+1} can be obtained from:

$$\begin{cases} W_{t+1} &= \sum_{k=1}^K x_{k,t+1} x'_{k,t+1} \\ x_{k,t+1} &= M x_{k,t} + \varepsilon_{k,t+1}, \quad k \in \{1, \dots, K\}, \end{cases}$$

where $\varepsilon_{k,t+1} \sim i.i.d. \mathcal{N}(0, \Omega)$ (independent across k 's). The proposition also shows that we have:

$$\mathbb{E}(W_{t+1} | \underline{W}_t) = M W_t M' + K \Omega,$$

i.e. W_t follows a matrix weak AR(1) process.

²Indeed, $\text{Tr}(\Gamma W_{t+1})$ is equal to:

$$\sum_{i=1}^L (\Gamma W_{t+1})_{ii} = \sum_{i=1}^L \sum_{j=1}^L \gamma_{ij} W_{t+1,ij} = \sum_{i=1}^L \gamma_{ii} W_{t+1,ii} + \sum_{i < j}^L (\gamma_{ij} + \gamma_{ji}) W_{t+1,ij}.$$

In particular case, where $L = 1$ (univariate case), we have that:

$$\mathbb{E}[\exp(uW_{t+1})|\underline{W}_t] = \exp\left[\frac{um^2}{1-2\omega u}W_t - \frac{K}{2}\log(1-2\omega u)\right].$$

Hence, when $L = 1$, the Wishart process boils down to an $ARG(1)$ process (Example 1.9) with $\rho = m^2$, $\mu = 2\omega$, $\nu = \frac{K}{2}$.

1.8 Building affine processes

1.8.1 Univariate affine processes with stochastic parameters

Some univariate affine processes can be extended if they satisfy certain conditions. Specifically, consider a univariate affine process whose conditional L.T. is of the form:

$$\mathbb{E}_t \exp(uy_{t+1}) = \exp[a_0(u)y_t + b_0(u)\delta], \quad (1.5)$$

where $\delta = (\delta_1, \dots, \delta_m)' \in \mathcal{D}$. This process can be generalized by making δ stochastic (while staying in the affine family). More precisely assume that:

$$\mathbb{E}[\exp(uy_{t+1})|\underline{y}_t, \underline{z}_{t+1}] = \exp[a_0(u)y_t + b_0(u)'\Lambda z_{t+1}],$$

where Λ is a $(m \times k)$ matrix, with $\Lambda z_{t+1} \in \mathcal{D}$. In this case, if:

$$\mathbb{E}[\exp(v'z_{t+1})|\underline{y}_t, \underline{z}_t] = \exp[a_1(v)'z_t + b_1(v)],$$

then $w_{t+1} = (y_{t+1}, z'_{t+1})'$ is affine.³

³Indeed, we have:

$$\begin{aligned} & \mathbb{E}[\exp(uy_{t+1} + v'z_{t+1})|\underline{y}_t, \underline{z}_t] \\ &= \mathbb{E}\{\exp(v'z_{t+1})\mathbb{E}[\exp(uy_{t+1})|\underline{y}_t, \underline{z}_{t+1}]\underline{y}_t, \underline{z}_t\} \\ &= \mathbb{E}\{\exp[a_0(u)y_t + b_0(u)'\Lambda z_{t+1} + v'z_{t+1}]\underline{y}_t, \underline{z}_t\} \\ &= \exp\{a_0(u)y_t + a_1[\Lambda'b_0(u) + v]'z_t + b_1[\Lambda'b_0(u) + v]\}. \end{aligned}$$

Example 1.13 (Gaussian AR(p)). Using the notation of Example 1.11, it comes that an AR(p) processes satisfies Eq. (1.5) with $b_0(u) = \left(u, \frac{u^2}{2}\right)'$ and $\delta = (\nu, \sigma^2)' \in \mathcal{D} = \mathbb{R} \times \mathbb{R}^+$. In that case, δ (the vector of conditional mean and variance) can be replaced by ...

- $\begin{pmatrix} z_{1,t+1} \\ z_{2,t+1} \end{pmatrix}$, where $z_{1,t+1}$ and $z_{2,t+1}$ are independent AR(1) (see Example 1.6) and ARG(1) (see Example 1.9) processes, respectively.
- $\begin{pmatrix} \lambda'_1 & 0 \\ 0 & \lambda'_2 \end{pmatrix} \begin{pmatrix} z_{1,t+1} \\ z_{2,t+1} \end{pmatrix}$, where $z_{1,t+1}$ and $z_{2,t+1}$ are independent Markov chains.
- $\begin{pmatrix} \lambda'_1 \\ \lambda'_2 \end{pmatrix} z_{t+1}$, where z_{t+1} is a Markov chain.

Example 1.14 (ARG(p) model).

- $b_0(u) = -\nu \log(1 - u\mu)$, $\delta = \nu$.
- $\nu (\geq 0)$ can be specified for instance as a Markov chain or an ARG.

1.8.2 Multivariate affine processes

One can construct multivariate affine processes by employing the so-called recursive approach. Let us illustrate this by considering the bivariate case. (The multivariate generalization is straightforward.) Consider

$w_t = \begin{pmatrix} w_{1,t} \\ w_{2,t} \end{pmatrix}$, and assume that we have:

$$\begin{aligned} & \mathbb{E}[\exp(u_1 w_{1,t+1} | \underline{w_{1,t}}, \underline{w_{2,t}})] \\ &= \exp[a_{11}(u_1) w_{1,\textcolor{red}{t}} + a_{12}(u_1) w_{2,\textcolor{red}{t}} + b_1(u_1)], \end{aligned}$$

and:

$$\begin{aligned} & \mathbb{E}[\exp(u_2 w_{2,t+1} | \underline{w_{1,t+1}}, \underline{w_{2,t}})] \\ &= \exp[a_0(u_2) w_{1,\textcolor{red}{t+1}} + a_{21}(u_2) w_{1,\textcolor{red}{t}} + a_{22}(u_2) w_{2,\textcolor{red}{t}} + b_2(u_2)]. \end{aligned}$$

Then w_t is an affine process.⁴ The dynamics of the two components of w_t are of the form:

$$\begin{aligned} w_{1,t+1} &= \alpha_1 + \alpha_{11}w_{1,t} + \alpha_{12}w_{2,t} + \varepsilon_{1,t+1} \\ w_{2,t+1} &= \alpha_2 + \alpha_0w_{1,t+1} + \alpha_{21}w_{1,t} + \alpha_{22}w_{2,t} + \varepsilon_{2,t+1} \end{aligned}$$

Note that $\varepsilon_{1,t+1}$ and $\varepsilon_{2,t+1}$ are non-correlated martingale differences. In the general case, they are conditionally heteroskedastic. What precedes is at play in *VAR* model; Monfort et al. (2017) employ this approach to build vector auto-regressive gamma (VARG) processes.

1.8.3 Extending multivariate stochastic processes

Consider the same framework as in Section 1.8.1 when y_t is a n -dimensional vector. That is, replace Eq. (1.5) with:

$$\mathbb{E}_t \exp(u' y_{t+1}) = \exp[a_0(u)' y_t + b_0(u) \delta], \quad (1.6)$$

and further assume that δ is stochastic and depends on z_t , such that:

$$\mathbb{E}[\exp(uy_{t+1}) | \underline{y}_t, \underline{z}_{t+1}] = \exp[a_0(u)y_t + b_0(u)' \Lambda z_{t+1}],$$

where Λ is a $(m \times k)$ matrix, with $\Lambda z_{t+1} \in \mathcal{D}$. In this case, if:

$$\mathbb{E}[\exp(v' z_{t+1}) | \underline{y}_t, \underline{z}_t] = \exp[a_1(v)' z_t + b_1(v)],$$

then $w_{t+1} = (y_{t+1}, z'_{t+1})'$ is affine.

⁴Indeed, we have:

$$\begin{aligned} & \mathbb{E}[\exp(u_1 w_{1,t+1} + u_2 w_{2,t+1} | \underline{w}_{1,t}, \underline{w}_{2,t})] \\ &= \mathbb{E}\{\exp(u_1 w_{1,t+1}) \mathbb{E}[\exp(u_2 w_{2,t+1}) | \underline{w}_{1,t+1}, \underline{w}_{2,t}] | \underline{w}_{1,t}, \underline{w}_{2,t}\} \\ &= \mathbb{E}\{\exp[u_1 + a_0(u_2)w_{1,t+1} + a_{21}(u_2)w_{1,t} + a_{22}(u_2)w_{2,t} + b_2(u_2)] | \underline{w}_{1,t}, \underline{w}_{2,t}\} \\ &= \exp\{a_{11}[u_1 + a_0(u_2)]w_{1,t} + a_{12}[u_1 + a_0(u_2)]w_{2,t} + b_1[u_1 + a_0(u_2)] \\ & \quad + a_{21}(u_2)w_{1,t} + a_{22}(u_2)w_{2,t} + b_2(u_2)\}. \end{aligned}$$

Example 1.15 (Stochastic parameters Gaussian VAR(1)). This example extends Example 1.7. Using the same notations as in the latter example 1.7, we have

$$b_0(u) = \left(u', \frac{1}{2}(u \otimes u)' \right)' \quad \text{and} \quad \delta = (\mu', \text{vec}(\Sigma)')' \in \mathbb{R}^n \times \text{vec}(\mathcal{S}),$$

where \mathcal{S} is the set of symmetric positive semi-definite matrices. Vector δ can be replaced by:

$$\begin{pmatrix} z_{1,t+1} \\ z_{2,t+1} \end{pmatrix},$$

where

- $z_{1,t+1}$ is, for instance, a Gaussian VAR process.
- $z_{2,t+1}$ is
- obtained by applying the vec operator to a Wishart process,
- replaced by $\Lambda_2 z_{2,t+1}$, where Λ_2 is a $(n^2 \times J)$ matrix whose columns are $\text{vec}(\Sigma_j)$, $j \in \{1, \dots, J\}$, the Σ_j being $(n \times n)$ positive semi-definite,
- a standardized J -dimensional VARG process (multivariate extension of Example 1.9).

Example 1.16 (Regime-switching VAR(1)). One can also use this approach to construct (affine) regime-switching VAR processes (which is another extension of Example 1.7 (see, e.g., [Gourieroux et al. \(2014a\)](#)). For that, replace δ with

- $\begin{pmatrix} \Lambda_1 & 0 \\ 0 & \Lambda_2 \end{pmatrix} \begin{pmatrix} z_{1,t+1} \\ z_{2,t+1} \end{pmatrix}$, where Λ_1 is a $(n \times J_1)$ matrix and $z_{1,t+1}$ is a Markov chain valued in the set of selection vectors of size J_1 (see Subsection 1.6), Λ_2 is the same matrix as in Example 1.15 and $z_{2,t+1}$ is a Markov chain valued in the set of selection vectors of size J_2 .
- or $\begin{pmatrix} \Lambda_1 \\ \Lambda_2 \end{pmatrix} z_{t+1}$, where Λ_1 and Λ_2 are the same matrices as above with $J_1 = J_2 = J$, and z_{t+1} is a Markov chain valued in the set of selection vectors of size J .

1.8.4 Extended affine processes

Some processes are not affine, but may be sub-components of an affine process. This can be useful to compute their conditional moments and multi-horizon Laplace transform (as one can use the formulas presented above for that, using the enlarged—affine—vector).

Let us formally define an extended affine process:

Definition 1.5 (Extended Affine Processes). A process $w_{1,t}$ is extended affine if there exists a process $w_{2,t} = g(\underline{w_{1,t}})$ such that $(w'_{1,t}, w'_{2,t})'$ is affine (of order 1).

For an extended affine processes, $\varphi_{1,t}(u) = \mathbb{E}[\exp(u'w_{1,t+1})|\underline{w_{1,t}}]$ can be obtained from:

$$\begin{aligned}\varphi_t(u_1, u_2) &= \mathbb{E}[\exp(u'_1 w_{1,t+1} + u'_2 w_{2,t+1})|\underline{w_{1,t}}, \underline{w_{2,t}}] \\ &= \exp[a'_1(u_1, u_2)w_{1,t} + a'_2(u_1, u_2)w_{2,t} + b(u_1, u_2)]\end{aligned}$$

by:

$$\varphi_{1,t}(u) = \varphi_t(u, 0) = \exp[a'_1(u, 0)w_{1,t} + a'_2(u, 0)g(\underline{w_{1,t}}) + b(u, 0)].$$

In particular $w_{1,t}$ may be non-Markovian.

Similarly the multi-horizon Laplace transform (see Section 1.9)

$$\mathbb{E}[\exp(\gamma'_1 w_{1,t+1} + \dots + \gamma'_h w_{1,t+h})|\underline{w_{1,t}}]$$

can be obtained from the knowledge of the extended multi-horizon Laplace transform:

$$\begin{aligned}&\mathbb{E}_t[\exp(\{\gamma'_{1,1} w_{1,t+1} + \gamma'_{2,1} w_{2,t+1}\} + \dots + \{\gamma'_{1,h} w_{1,t+h} + \gamma'_{2,h} w_{2,t+h}\})] \\ &= \exp[A'_{1,t,h}(\gamma_1^h, \gamma_2^h)w_{1,t} + A'_{2,t,h}(\gamma_1^h, \gamma_2^h)w_{2,t} + B_{t,h}(\gamma_1^h, \gamma_2^h)],\end{aligned}$$

(with $\gamma_1^h = (\gamma'_{1,1}, \dots, \gamma'_{1,h})'$, and $\gamma_2^h = (\gamma'_{2,1}, \dots, \gamma'_{2,h})'$). We indeed have:

$$\begin{aligned}&\mathbb{E}[\exp(\gamma'_1 w_{1,t+1} + \dots + \gamma'_h w_{1,t+h})|\underline{w_{1,t}}] \\ &= \exp[A'_{1,t,h}(\gamma^h, 0)w_{1,t} + A'_{2,t,h}(\gamma^h, 0)g(\underline{w_{1,t}}) + B_{t,h}(\gamma^h, 0)],\end{aligned}$$

with $\gamma^h = (\gamma'_1, \dots, \gamma'_h)'$.

Example 1.17 (Affine process of order p). If $\{w_{1,t}\}$ is affine of order $p > 1$, then $(w_{1,t}, \dots, w_{1,t-p+1})$ is affine of order 1, but $\{w_{1,t}\}$ is not affine. That is, in that case, $w_{2,t} = (w'_{1,t-1}, \dots, w'_{1,t-p+1})'$.

This is a kind of extreme case since $w_{2,t}$ belongs to the information at $t-1$, which implies $a_2(u_1, u_2) = u_2$.

Example 1.18 (Gaussian ARMA process). Consider an $ARMA(1, 1)$ process

$$w_{1,t} - \varphi w_{1,t-1} = \varepsilon_t - \theta \varepsilon_{t-1},$$

with $|\varphi| < 1$, $|\theta| < 1$, and $\varepsilon_t \sim i.i.d. \mathcal{N}(0, \sigma^2)$.

$w_{1,t}$ is not Markovian. Now, take $w_{2,t} = \varepsilon_t = (1 - \theta L)^{-1}(1 - \varphi L)w_{1,t}$. We have:

$$\begin{pmatrix} w_{1,t+1} \\ w_{2,t+1} \end{pmatrix} = \begin{pmatrix} \varphi & -\theta \\ 0 & 0 \end{pmatrix} \begin{pmatrix} w_{1,t} \\ w_{2,t} \end{pmatrix} + \begin{pmatrix} 1 \\ 1 \end{pmatrix} \varepsilon_{t+1}.$$

Hence $(w_{1,t}, w_{2,t})'$ is Gaussian $VAR(1)$, and, therefore, it is affine of order 1.

This is easily extended to $ARMA(p, q)$ and $VARMA(p, q)$ processes.

Example 1.19 (GARCH type process). Consider process $w_{1,t}$, defined by:

$$w_{1,t+1} = \mu + \varphi w_{1,t} + \sigma_{t+1} \varepsilon_{t+1},$$

where $|\varphi| < 1$ and $\varepsilon_t \sim i.i.d. \mathcal{N}(0, 1)$, and

$$\sigma_{t+1}^2 = \omega + \alpha \varepsilon_t^2 + \beta \sigma_t^2,$$

where $0 < \beta < 1$.

Consider $w_{2,t} = \sigma_{t+1}^2$ (which is a non-linear function of $w_{1,t}$). Proposition 1.7 shows that:

$$\begin{aligned} & \mathbb{E} \left[\exp(u_1 w_{1,t+1} + u_2 w_{2,t+1}) | \underline{w_{1,t}} \right] \\ &= \exp \left[u_1 \mu + u_2 \omega - \frac{1}{2} \log(1 - 2u_2 \alpha) \right. \\ & \quad \left. + u_1 \varphi w_{1,t} + (u_2 \beta + \frac{u_1^2}{2(1 - 2u_2 \alpha)}) w_{2,t} \right], \end{aligned}$$

which is exponential affine in $(w_{1,t}, w_{2,t})$.

1.9 Multi-horizon Laplace transform

1.9.1 Recursive computation and direct pricing implications

In this subsection, we show that multi-horizon Laplace transforms of affine processes can be calculated recursively. Various examples will show how this can be exploited to price long-dated financial instruments.

Let us consider a multivariate process w_t , affine of order one. (As explained in Subsection 1.5.2, this includes the case of the order p case.) For the sake of generality, we consider the case where functions $a(\cdot)$, $b(\cdot)$ are possibly deterministic functions of time, denoted in this case $a_{t+1}(\cdot)$ and $b_{t+1}(\cdot)$:

$$\mathbb{E}_t \exp[(u' w_{t+1})] = \exp[a'_{t+1}(u) w_t + b_{t+1}(u)].$$

The multi-horizon Laplace transform associated with date t and horizon h is defined by:

$$\varphi_{t,h}(\gamma_1, \dots, \gamma_h) = \mathbb{E}_t[\exp(\gamma'_1 w_{t+1} + \dots + \gamma'_h w_{t+h})]. \quad (1.7)$$

Lemma 1.5 (in the appendix) shows that we have:

$$\varphi_{t,h}(\gamma_1, \dots, \gamma_h) = \exp(A'_{t,h} w_t + B_{t,h}),$$

where $A_{t,h} = A_{t,h}^h$ and $B_{t,h} = B_{t,h}^h$, the $A_{t,i}^h, B_{t,i}^h$ $i = 1, \dots, h$, being given recursively by:

$$\begin{cases} A_{t,i}^h &= a_{t+h+1-i}(\gamma_{h+1-i} + A_{t,i-1}^h), \\ B_{t,i}^h &= b_{t+h+1-i}(\gamma_{h+1-i} + A_{t,i-1}^h) + B_{t,i-1}^h, \\ A_{t,0}^h &= 0, B_{t,0}^h = 0. \end{cases}$$

If the functions a_t and b_t do not depend on t , these recursive formulas do not depend on t , and we get $\varphi_{t,h}(\gamma_1, \dots, \gamma_h)$, for any t , with only one recursion for each h .

Moreover, if the functions a_t and b_t do not depend on t , and if different sequences $(\gamma_1^h, \dots, \gamma_h^h)$, $h = 1, \dots, H$ (say) satisfy $\gamma_{h+1-i}^h = u_i$, for $i = 1, \dots, h$, and for any $h \leq H$, that is if we want to compute (*reverse-order* case):

$$\varphi_{t,h}(u_h, \dots, u_1) = \mathbb{E}_t[\exp(u'_h w_{\textcolor{red}{t}+1} + \dots + u'_1 w_{\textcolor{red}{t}+h})], \quad h = 1, \dots, H, \quad (1.8)$$

then Proposition 1.5 (in the appendix) shows that we can compute the $\varphi_{t,h}(u_h, \dots, u_1)$ for any t and any $h \leq H$ with only one recursion. That is $\varphi_{t,h}(u_h, \dots, u_1) = \exp(A'_h w_t + B_h)$ with:

$$\begin{cases} A_h &= a(u_h + A_{h-1}), \\ B_h &= b(u_h + A_{h-1}) + B_{h-1}, \\ A_0 &= 0, \quad B_0 = 0. \end{cases}$$

As mentioned above, what precedes has useful implications to price long-dated financial instruments such as nominal and real bonds (Examples 1.20 and 1.22, respectively), or futures (Example 1.23).

Example 1.20 (Nominal interest rates). Let $B(t, h)$ denote the date- t price of a nominal zero-coupon bond of maturity h . We Have:

$$B(t, h) = \mathbb{E}_t^{\mathbb{Q}} \exp(-r_t - \dots - r_{t+h-1}), \quad (1.9)$$

where r_t is the nominal short rate between t and $t+1$ (observed at t), and the associated (continuously-compounded) yield-to-maturity is given by:

$$R(t, h) = -\frac{1}{h} \log B(t, h), \quad h = 1, \dots, H. \quad (1.10)$$

If $r_t = \omega' w_t$ (say), then:

$$B(t, h) = \exp(-r_t) \mathbb{E}_t^{\mathbb{Q}} \exp(-\omega' w_{t+1} - \dots - \omega' w_{t+h-1}).$$

One can then price this bond by directly employing Eq. (1.8), with $u_1 = 0$ and $u_i = -\omega$, $i = 2, \dots, H$. The price $B(t, h)$ is exponential affine in w_t , the associated yield-to-maturity $R(t, h) = -1/h \log B(t, h)$ is affine in w_t .

Example 1.21 (No-arbitrage Nelson-Siegel model). In this example, we employ the results of Example 1.20 in the context described by Christensen et al. (2009). Specifically, we consider a three factor model following a Gaussian VAR (see Example 1.7):

$$w_t = \begin{bmatrix} X_{1,t} \\ X_{2,t} \\ X_{3,t} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 - \lambda & \lambda \\ 0 & 0 & 1 - \lambda \end{bmatrix} \begin{bmatrix} X_{1,t-1} \\ X_{2,t-1} \\ X_{3,t-1} \end{bmatrix} + \begin{bmatrix} \sigma_{11} & 0 & 0 \\ \sigma_{21} & \sigma_{22} & 0 \\ \sigma_{31} & \sigma_{32} & \sigma_{33} \end{bmatrix} \begin{bmatrix} \varepsilon_{1,t} \\ \varepsilon_{2,t} \\ \varepsilon_{3,t} \end{bmatrix},$$

where $\begin{bmatrix} \varepsilon_{1,t} \\ \varepsilon_{2,t} \\ \varepsilon_{3,t} \end{bmatrix} \sim i.i.d. \mathcal{N}(0, Id).$

The nominal short-term rate is given by $r_t = X_{1,t} + X_{2,t}$. In that case, we can use the results of Example 1.20 with $\omega = (-1, -1, 0)'$. The following lines of code do that:

```
library(TSModels)
lambda <- .05
Phi <- diag(c(1,1-lambda,1-lambda));Phi[2,3] <- lambda
Sigma <- .0005 * diag(3)
psi.parameterization=list(mu=matrix(0,3,1),Phi=Phi,Sigma=Sigma)
u1 <- matrix(0,3,1)
u2 <- matrix(c(-1,-1,0),ncol=1)
H <- 20
AB <- reverse.MHLT(psi.GaussianVAR,u1 = u1,u2 = u2,H = H,
                    psi.parameterization = psi.parameterization)
AB$A[1:2,,] <- AB$A[1:2,,] - 1 # add terms corresponding to exp(-r_t)
a.yield <- - AB$A / array((1:H) %x% rep(1,3),c(3,1,H))
b.yield <- - AB$B / array((1:H) %x% rep(1,3),c(1,1,H))
plot(a.yield[1,,],type="l",lwd=2,ylim=c(0,1),
     xlab="Maturity",ylab="Factor loadings")
lines(a.yield[2,,],col="red",lwd=2,lty=2)
lines(a.yield[3,,],col="blue",lwd=2,lty=3)
```

In the previous example, note the use of function `reverse.MHLT` (in package `TSModels`), that notably takes a L.T. as an argument (`psi`). In the previous example, we consider a Gaussian VAR, and we therefore assign `psi.GaussianVAR` to `psi`. We then need to provide function `reverse.MHLT` with the arguments of the `psi` function. These arguments are provided in the form of a list (input `psi.parameterization`).

Example 1.22 (Real interest rates). Denote by q_t the price index on date t

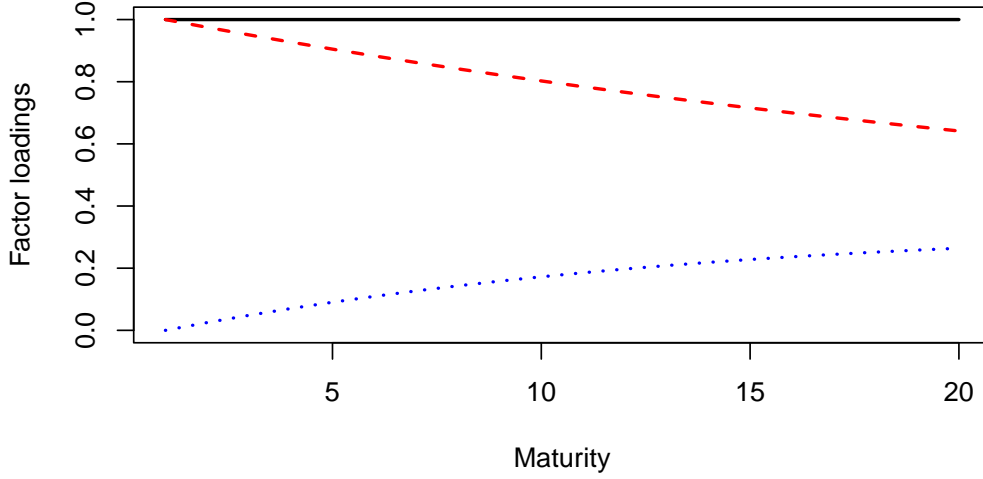


Figure 1.4: Factor loadings in the context of a no-arbitrage Nelson-Siegel model (Christensen, Diebold and Rudebusch, 2009). The first factor (black solid line) is a level factor. The second and third factors (red dashed line and blue dotted line, respectively) are slope factors.

and by $\pi_{t+1} = \log \frac{q_{t+1}}{q_t}$ the inflation rate on date $t + 1$. We have:

$$\bar{R}(t, h) = -\frac{1}{h} \log \bar{B}(t, h), \quad h = 1, \dots, H$$

$$\begin{aligned} \bar{B}(t, h) &= \mathbb{E}_t^{\mathbb{Q}} \exp(-r_t - \dots - r_{t+h-1} + \pi_{t+1} + \dots + \pi_{t+h}), \\ &= \exp(-r_t) \times \\ &\quad \mathbb{E}_t^{\mathbb{Q}} \exp(-r_{t+1} - \dots - r_{t+h-1} + \pi_{t+1} + \dots + \pi_{t+h}) \end{aligned}$$

If $r_t = \omega' w_t$ and $\pi_t = \bar{\omega}' w_t$, then $\bar{B}(t, h)$ is given by:

$$\exp(-r_t) \mathbb{E}_t^{\mathbb{Q}} \exp[(\bar{\omega} - \omega)' w_{t+1} + \dots + (\bar{\omega} - \omega)' w_{t+h-1} + \bar{\omega}' w_{t+h}]$$

One can then price this bond by directly employing Eq. (1.8), with $u_1 = \bar{\omega}$ and $u_i = \bar{\omega} - \omega$, $i = 2, \dots, H$.

Example 1.23 (Futures). Denote by $F(t, h)$ the date- t price of a future of maturity h (see Section XXX). That is $F(t, h) = \mathbb{E}_t^{\mathbb{Q}}(S_{t+h})$, $h = 1, \dots, H$, where S_t is the date- t price of the underlying asset.

- If $w_t = (\log S_t, x'_t)'$ then $F(t, h) = \mathbb{E}_t^{\mathbb{Q}} \exp(e'_1 w_{t+h})$. This can be calculated by using Eq. (1.8) with $u_1 = e_1$, and $u_i = 0$, for $i = 2, \dots, H$.
- If $w_t = (y_t, x'_t)'$ with $y_t = \log \frac{S_t}{S_{t-1}}$, then $F(t, h) = S_t \mathbb{E}_t^{\mathbb{Q}} \exp(e'_1 w_{t+1} + \dots + e'_1 w_{t+h})$. This can be calculated by using Eq. (1.8) with $u_i = e'_1$, $i = 1, \dots, H$.

1.9.2 Exponential payoff

Consider an asset providing the payoff $\exp(\nu' w_{t+h})$ on date $t + h$. Its price is given by:

$$P(t, h; \nu) = \mathbb{E}_t^{\mathbb{Q}} [\exp(-r_t - \dots - r_{t+h-1}) \exp(\nu' w_{t+h})].$$

If $r_t = \omega' w_t$, we have:

$$P(t, h; \nu) = \exp(-r_t) \mathbb{E}_t^{\mathbb{Q}} (\exp[-\omega' w_{t+1} - \dots - \omega' w_{t+h-1} + \nu' w_{t+h}]),$$

which can be calculated by Eq. (1.8), with $u_1 = \nu$ and $u_i = -\omega$ for $i = 2, \dots, H$.

What precedes can be extended to the case where the payoff (settled on date $t + h$) is of the form:

$$(\nu'_1 w_{t+h}) \exp(\nu'_2 w_{t+h}).$$

Indeed, we have

$$\left[\frac{\partial \exp[(s\nu_1 + \nu_2)' w_{t+h}]}{\partial s} \right]_{s=0} = (\nu'_1 w_{t+h}) \exp(\nu'_2 w_{t+h}).$$

Therefore:

$$\begin{aligned} & \mathbb{E}_t^{\mathbb{Q}} [\exp(-r_t - \dots - r_{t+h-1}) (\nu'_1 w_{t+h}) \exp(\nu'_2 w_{t+h})] \\ &= \left[\frac{\partial P(t, h; s\nu_1 + \nu_2)}{\partial s} \right]_{s=0}. \end{aligned} \tag{1.11}$$

This method is easily extended to price payoffs of the form $(\nu'_1 w_{t+h})^k \exp(\nu'_2 w_{t+h})$, with $k \in \mathbb{N}$.

1.10 VAR representation and conditional moments

An important property of affine processes is that their dynamics can be written as a vector-autoregressive process. This is useful to compute conditional moments of the process.

Proposition 1.1 (VAR representation of an affine process' dynamics). *If w_t is the affine process whose Laplace transform is defined in Def. 1.1, then its dynamics admits the following vectorial autoregressive representation:*

$$w_{t+1} = \mu + \Phi w_t + \Sigma^{\frac{1}{2}}(w_t) \varepsilon_{t+1}, \quad (1.12)$$

where ε_{t+1} is a difference of martingale sequence whose conditional covariance matrix is the identity matrix and where μ , Φ and $\Sigma(w_t) = \Sigma^{\frac{1}{2}}(w_t) \Sigma^{\frac{1}{2}}(w_t)'$ satisfy:

$$\mu = \left[\frac{\partial}{\partial u} b(u) \right]_{u=0}, \quad \Phi = \left[\frac{\partial}{\partial u} a(u)' \right]_{u=0} \quad (1.13)$$

$$\Sigma(w_t) = \left[\frac{\partial}{\partial u \partial u'} b(u) \right]_{u=0} + \left[\frac{\partial}{\partial u \partial u'} a(u)' w_t \right]_{u=0}. \quad (1.14)$$

Proof. When w_t is affine, its (conditional) cumulant generating function is of the form $\psi(u) = a(u)' w_t + b(u)$. The result directly follows from the formulas given in Section 1.3. \square

Proposition 1.6 (in the appendix) shows that the conditional means and variances of w_t are given by:

$$\mathbb{E}_t(w_{t+h}) = (I - \Phi)^{-1} (I - \Phi^h) \mu + \Phi^h w_t \quad (1.15)$$

$$\begin{aligned} \mathbb{V}ar_t(w_{t+h}) &= \Sigma(\mathbb{E}_t(w_{t+h-1})) + \Phi \Sigma(\mathbb{E}_t(w_{t+h-2})) \Phi' + \\ &\dots + \Phi^{h-1} \Sigma(w_t) \Phi^{h-1}'. \end{aligned} \quad (1.16)$$

Eq. (1.16) notably implies that $\mathbb{V}ar_t(w_{t+h})$ is an affine function of w_t . Indeed $\Sigma(\cdot)$ is an affine function, and the conditional expectations $\mathbb{E}_t(w_{t+h})$ are affine in w_t , as shown by Eq. (1.15).

The unconditional means and variances are given by:

$$\begin{cases} \mathbb{E}(w_t) &= (I - \Phi)^{-1} \mu \\ \text{vec}[\mathbb{V}ar(w_t)] &= (I_{n^2} - \Phi \otimes \Phi)^{-1} \text{vec}(\Sigma[(I - \Phi)^{-1} \mu]). \end{cases} \quad (1.17)$$

1.11 Truncated Laplace transforms of affine processes

In this section, we show how one can employ Fourier transforms to compute truncated conditional moments of affine processes. For that, let us introduce the following notation:

$$w_{t+1,T} = (w'_{t+1}, w'_{t+2}, \dots, w'_T)'$$

with w_t affine n -dimensional process.

We want to compute:

$$\tilde{\varphi}_t(u; v, \gamma) = \mathbb{E}_t[\exp(u' w_{t+1,T}) \mathbf{1}_{\{v' w_{t+1,T} < \gamma\}}].$$

Consider the complex untruncated conditional Laplace transform:

$$\varphi_t(z) = \mathbb{E}_t[\exp(z' w_{t+1,T})], \quad z \in \mathbb{C}^{nT},$$

computed using the same recursive algorithm as in the real case (see Section 1.9).

Duffie et al. (2000) have shown that we have (see also Proposition 1.8 in the appendix):

$$\tilde{\varphi}_t(u; v, \gamma) = \frac{\varphi_t(u)}{2} - \frac{1}{\pi} \int_0^\infty \frac{\text{Im}[\varphi_t(u + ivx) \exp(-i\gamma x)]}{x} dx. \quad (1.18)$$

where Im means imaginary part.

Note that the integral in Eq. (1.18) is one dimensional (whatever the dimension of w_t). As shown in the following example, this can be exploited to price options.

Example 1.24 (Option pricing). Pricing calls and puts amounts to conditional expectations of the type (with $k > 0$):

$$\begin{aligned} & \mathbb{E}_t([\exp(u'_1 w_{t+1,T}) - k \exp(u'_2 w_{t+1,T})]^+) \\ &= \mathbb{E}_t([\exp(u'_1 w_{t+1,T}) - k \exp(u'_2 w_{t+1,T})] \mathbf{1}_{\{[\exp(u_1 - u_2)' w_{t+1,T}] > k\}}) \\ &= \tilde{\varphi}_t(u_1; u_2 - u_1, -\log k) - k \tilde{\varphi}_t(u_2; u_2 - u_1, -\log k). \end{aligned}$$

Example 1.25 (Exogenous short rate). Consider an asset whose date- t price is p_t . Denote its geometric asset return by y_t , i.e., $y_t = \log(p_t/p_{t-1})$. Consider an option written on this asset, with a strike equal kp_t .

If interest rates are deterministic, the option price, for a maturity h , is given by:

$$p_t \exp(-r_t - \dots - r_{t+h-1}) \mathbb{E}_t^{\mathbb{Q}} [\exp u'_1 w_{t+1,t+h} - k]^+$$

with $u_1 = e \otimes e_1$, where e is the h -dimensional vector with components equal to 1, and e_1 is the n -vector selecting the 1st component (y_t being the 1st component of w_t , say).

Example 1.26 (Endogenous short rate). Consider the same context as in Example 1.25, but with a stochastic (endogenous) short-term rate. For instance, assume that $r_{t+1} = \omega_0 + \omega'_1 w_t$. The option price then is:

$$\begin{aligned} & p_t \mathbb{E}_t^{\mathbb{Q}} [\exp(-\omega_0 - \omega'_1 w_t - \dots - \omega_0 - \omega'_1 w_{t+h-1}) [\exp(u'_1 w_{t+1,t+h}) - k]^+] \\ &= p_t \exp(-h\omega_0 - \omega'_1 w_t) \mathbb{E}_t^{\mathbb{Q}} \left([\exp(\tilde{u}'_1 w_{t+1,t+h}) - k \exp(u_2 w_{t+1,t+h})]^+ \right), \end{aligned}$$

with $\tilde{u}'_1 = u_1 + u_2$, [$u_1 = e \otimes e_1$ as before], and $u_2 = (-\omega'_1, \dots, -\omega'_1, 0)'$.

Example 1.27 (Numerical example: Conditional cumulated distribution function (c.d.f.)). Let us use the model used in Example 1.21. Suppose we want to compute the conditional distribution of the average interest rate over the next H periods, i.e., $\frac{1}{H}(r_{t+1} + \dots + r_{t+H})$. Hence, we want to compute $\mathbb{E}_t[\mathbf{1}_{\{v'w_{t+1,T} < \gamma\}}]$ with $v'w_{t+1,T} = \frac{1}{H}(r_{t+1} + \dots + r_{t+H})$.

```
H <- 10
X <- matrix(c(0.01,.02,0),3,1)
x <- exp(seq(-10,10,length.out=1000))
u1 <- matrix(c(1/H,1/H,0),3,1) %*% matrix(1i*x,nrow=1); u2 <- u1
AB <- reverse.MHLT(psi.GaussianVAR,u1 = u1,u2 = u2,H = H,
                    psi.parameterization = psi.parameterization)
s1 <- matrix(exp(t(X) %*% AB$A[, ,H] + AB$B[, ,H]),ncol=1)
dx <- matrix(x-c(0,x[1:length(x)-1]),length(x),1)
gamma <- seq(-.2,.3,length.out=1000)
fx <- outer(x,gamma,function(r,c){Im(s1[,1]*exp(-1i*r*c))/r})*dx[,1]
f <- 1/2 - 1/pi * apply(fx,2,sum)
plot(gamma,f,type="l",xlab="",lwd=2)
```

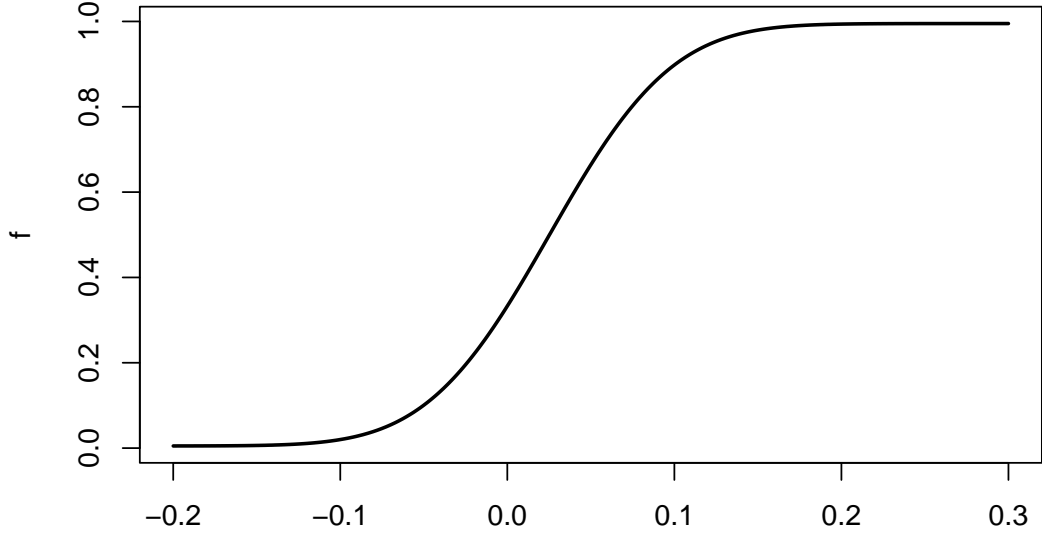


Figure 1.5: Conditional cumulated distribution function (c.d.f.) of $\frac{1}{H}(r_{t+1} + \dots + r_{t+H})$.

1.12 Appendices

Lemma 1.3. *If $\mu \in \mathbb{R}^L$ and Q is a $(L \times L)$ matrix symmetric positive definite, then:*

$$\int_{\mathbb{R}^L} \exp(-u'Qu + \mu'u) du = \frac{\pi^{L/2}}{(\det Q)^{1/2}} \exp\left(\frac{1}{4}\mu'Q^{-1}\mu\right).$$

Proof. The integral is:

$$\begin{aligned} & \int_{\mathbb{R}^L} \exp\left[-(u - \frac{1}{2}Q^{-1}\mu)'Q(u - \frac{1}{2}Q^{-1}\mu)\right] \exp\left(\frac{1}{4}\mu'Q^{-1}\mu\right) du \\ &= \frac{\pi^{L/2}}{(\det Q)^{1/2}} \exp\left(\frac{1}{4}\mu'Q^{-1}\mu\right) \end{aligned}$$

[using the formula for the unit mass of $\mathcal{N}(0.5Q^{-1}\mu, (2Q)^{-1})$]. \square

Lemma 1.4. *If $\varepsilon_{t+1} \sim \mathcal{N}(0, Id)$, we have*

$$\mathbb{E}_t(\exp[\lambda'\varepsilon_{t+1} + \varepsilon'_{t+1}V\varepsilon_{t+1}]) = \frac{1}{[\det(I - 2V)]^{1/2}} \exp\left[\frac{1}{2}\lambda'(I - 2V)^{-1}\lambda\right].$$

Proof. We have

$$\mathbb{E}_t \exp(\lambda' \varepsilon_{t+1} + \varepsilon'_{t+1} V \varepsilon_{t+1}) = \frac{1}{(2\pi)^{n/2}} \int_{\mathbb{R}^n} \exp \left[-u' \left(\frac{1}{2} I - V \right) u + \lambda' u \right] du$$

From Lemma 1.3, if $u \in \mathbb{R}^n$, then

$$\int_{\mathbb{R}^n} \exp(-u' Q u + \mu' u) du = \frac{\pi^{n/2}}{(\det Q)^{1/2}} \exp \left(\frac{1}{4} \mu' Q^{-1} \mu \right).$$

Therefore:

$$\begin{aligned} & \mathbb{E}_t \exp(\lambda' \varepsilon_{t+1} + \varepsilon'_{t+1} V \varepsilon_{t+1}) \\ &= \frac{1}{2^{n/2} \left[\det \left(\frac{1}{2} I - V \right) \right]^{1/2}} \exp \left[\frac{1}{4} \lambda' \left(\frac{1}{2} I - V \right)^{-1} \lambda \right]. \end{aligned}$$

□

Proposition 1.2 (Quadratic Gaussian process). *Consider vector $w_t = (x'_t, \text{vec}(x_t x'_t))'$, where x_t is a n -dimensional vector following a Gaussian VAR(1), i.e.*

$$x_{t+1} | \underline{w}_t \sim \mathcal{N}(\mu + \Phi x_t, \Sigma).$$

If $u = (v, V)$ where $v \in \mathbb{R}^n$ and V a square symmetric matrix of size n , we have:

$$\begin{aligned} \varphi_t(u) &= \mathbb{E}_t \{ \exp [(v', \text{vec}(V)') \times w_{t+1}] \} \\ &= \exp \{ a_1(v, V)' x_t + \text{vec}(a_2(v, V))' \text{vec}(x'_t x_t) + b(v, V) \}, \end{aligned}$$

where:

$$\begin{aligned} a_2(u) &= \Phi' V (I_n - 2\Sigma V)^{-1} \Phi \\ a_1(u) &= \Phi' [(I_n - 2V\Sigma)^{-1} (v + 2V\mu)] \\ b(u) &= u' (I_n - 2\Sigma V)^{-1} \left(\mu + \frac{1}{2} \Sigma v \right) + \\ &\quad \mu' V (I_n - 2\Sigma V)^{-1} \mu - \frac{1}{2} \log |I_n - 2\Sigma V|. \end{aligned}$$

Proof. We have:

$$\begin{aligned}
& \mathbb{E}_t(\exp(v'x_{t+1} + \text{vec}(V)' \text{vec}(x_{t+1}x'_{t+1}))) \\
&= \mathbb{E}_t[\exp(v'(\mu + \Phi x_t + \Sigma^{1/2}\varepsilon_{t+1}) + \\
&\quad \text{vec}(V)' \text{vec}((\mu + \Phi x_t + \Sigma^{1/2}\varepsilon_{t+1})(\mu + \Phi x_t + \Sigma^{1/2}\varepsilon_{t+1})'))] \\
&= \exp[v'(\mu + \Phi x_t) + \text{vec}(V)' \text{vec}\{(\mu + \Phi x_t)(\mu + \Phi x_t)'\}] \times \\
&\quad \mathbb{E}_t[\exp(v'\Sigma^{1/2}\varepsilon_{t+1} + \underbrace{2 \text{vec}(V)' \text{vec}\{(\mu + \Phi x_t)(\varepsilon'_{t+1}\Sigma^{1/2'})\}}_{=(\mu + \Phi x_t)'V\Sigma^{1/2}\varepsilon_{t+1}}) + \\
&\quad \underbrace{\text{vec}(V)' \text{vec}\{(\Sigma^{1/2}\varepsilon_{t+1})(\Sigma^{1/2}\varepsilon_{t+1})'\}}_{=\varepsilon'_{t+1}\Sigma^{1/2'}V\Sigma^{1/2}\varepsilon_{t+1}})]
\end{aligned}$$

Lemma 1.4 can be used to compute the previous conditional expectation, with $\lambda = \Sigma^{1/2'}(v + 2V'(\mu + \Phi x_t))$. Some algebra then leads to the result. \square

Proposition 1.3. *Consider the following auto-regressive gamma process:*

$$\frac{w_{t+1}}{\mu} \sim \gamma(\nu + z_t) \quad \text{where} \quad z_t \sim \mathcal{P}\left(\frac{\rho w_t}{\mu}\right),$$

with $\nu, \mu, \rho > 0$. (Alternatively $z_t \sim \mathcal{P}(\beta w_t)$, with $\rho = \beta\mu$.)

We have: $\varphi_t(u) = \exp\left[\frac{\rho u}{1 - u\mu}w_t - \nu \log(1 - u\mu)\right]$, for $u < \frac{1}{\mu}$.

Proof. Given \underline{w}_t , we have $z_t \sim \mathcal{P}\left(\frac{\rho w_t}{\mu}\right)$. We have:

$$\begin{aligned}
\mathbb{E}[\exp(uw_{t+1})|\underline{w}_t] &= \mathbb{E}\left\{\mathbb{E}\left[\exp\left(u\mu\frac{w_{t+1}}{\mu}\right)|\underline{w}_t, z_t\right]|\underline{w}_t\right\} \\
&= \mathbb{E}[(1 - u\mu)^{-(\nu + z_t)}|\underline{w}_t] \\
&= (1 - u\mu)^{-\nu} \mathbb{E}\{\exp[-z_t \log(1 - u\mu)]|\underline{w}_t\} \\
&= (1 - u\mu)^{-\nu} \exp\left\{\frac{\rho w_t}{\mu}[\exp(-\log(1 - u\mu))] - \frac{\rho w_t}{\mu}\right\} \\
&= \exp\left[\frac{\rho u w_t}{1 - u\mu} - \nu \log(1 - u\mu)\right],
\end{aligned}$$

using the fact that the L.T. of $\gamma(\nu)$ is $(1 - u)^{-\nu}$ and that the L.T. of $\mathcal{P}(\lambda)$ is $\exp[\lambda(\exp(u) - 1)]$. \square

Proposition 1.4 (Dynamics of a WAR process). *If K is an integer, W_{t+1} can be obtained from:*

$$\begin{cases} W_{t+1} &= \sum_{k=1}^K x_{k,t+1} x'_{k,t+1} \\ x_{k,t+1} &= Mx_{k,t} + \varepsilon_{k,t+1}, \quad k \in \{1, \dots, K\}, \end{cases}$$

where $\varepsilon_{k,t+1} \sim i.i.d.\mathcal{N}(0, \Omega)$ (independent across k 's). In particular, we have:

$$\mathbb{E}(W_{t+1} | \underline{W}_t) = MW_t M' + K\Omega,$$

i.e. W_t follows a matrix weak AR(1) process.

Proof. For $K = 1$, $W_{t+1} = x_{t+1} x'_{t+1}$, $x_{t+1} = Mx_t + \Omega^{1/2} u_{t+1}$ and $u_{t+1} \sim i.i.d.\mathcal{N}(0, Id_L)$. We have:

$$\mathbb{E}[\exp(\text{Tr} \Gamma W_{t+1}) | \underline{w}_t] = \mathbb{E}\{\mathbb{E}[\exp(\text{Tr} \Gamma x_{t+1} x'_{t+1}) | \underline{x}_t] | \underline{w}_t\}$$

and:

$$\begin{aligned} & \mathbb{E}[\exp(\text{Tr} \Gamma x_{t+1} x'_{t+1}) | \underline{x}_t] = \mathbb{E}[\exp(x'_{t+1} \Gamma x_{t+1} | \underline{x}_t)] \\ &= \mathbb{E}[\exp((Mx_t + \Omega^{1/2} u_{t+1})' \Gamma (Mx_t + \Omega^{1/2} u_{t+1}) / x_t)] \\ &= \exp(x'_t M' \Gamma M x_t) \mathbb{E}[\exp(2x'_t M' \Gamma \Omega^{1/2} u_{t+1} + u'_{t+1} \Omega^{1/2} \Gamma \Omega^{1/2} u_{t+1}) / x_t] \\ &= \frac{\exp(x'_t M' \Gamma M x_t)}{(2\pi)^{L/2}} \times \\ & \quad \int_{\mathbb{R}^L} \exp \left[2x'_t M' \Gamma \Omega^{1/2} u_{t+1} - u'_{t+1} \left(\frac{1}{2} Id_L - \Omega^{1/2} \Gamma \Omega^{1/2} \right) u_{t+1} \right] du_{t+1}. \end{aligned}$$

Using Lemma 1.3 with $\mu' = 2x'_t M' \Gamma \Omega^{1/2}$, $Q = \frac{1}{2} Id_L - \Omega^{1/2} \Gamma \Omega^{1/2}$ and after some algebra, the RHS becomes:

$$\frac{\exp[x'_t M' \Gamma (Id_L - 2\Omega \Gamma)^{-1} M x_t]}{\det[Id_L - 2\Omega^{1/2} \Gamma \Omega^{1/2}]} = \frac{\exp \text{Tr}[M' \Gamma (Id_L - 2\Omega^{-1}) M W_t]}{\det[Id_L - 2\Omega \Gamma]^{1/2}},$$

which depends on x_t through W_t , and gives the result for $K = 1$; the result for any K integer follows. \square

Lemma 1.5. *We have:*

$$\varphi_{t,h}(\gamma_1, \dots, \gamma_h) = \exp(A'_{t,h} w_t + B_{t,h}),$$

where $A_{t,h} = A_{t,h}^h$ and $B_{t,h} = B_{t,h}^h$, the $A_{t,i}^h, B_{t,i}^h$ $i = 1, \dots, h$, being given recursively by:

$$(i) \begin{cases} A_{t,i}^h &= a_{t+h+1-i}(\gamma_{h+1-i} + A_{t,i-1}^h), \\ B_{t,i}^h &= b_{t+h+1-i}(\gamma_{h+1-i} + A_{t,i-1}^h) + B_{t,i-1}^h, \\ A_{t,0}^h &= 0, B_{t,0}^h = 0. \end{cases}$$

Proof. For any $j = 1, \dots, h$ we have:

$$\varphi_{t,h}(\gamma_1, \dots, \gamma_h) = \mathbb{E}_t[\exp(\gamma'_1 w_{t+1} + \dots \gamma'_j w_{t+j} + A_{t,h-j}^{h'} w_{t+j} + B_{t,h-j}^h)]$$

where:

$$(ii) \begin{cases} A_{t,h-j+1}^h &= a_{t+j}(\gamma_j + A_{t,h-j}^h), \\ B_{t,h-j+1}^h &= b_{t+j}(\gamma_j + A_{t,h-j}^h) + B_{t,h-j}^h, \\ A_{t,0}^h &= 0, B_{t,0}^h = 0. \end{cases}$$

Since this is true for $j = h$, and if this is true for j , we get:

$$\begin{aligned} \varphi_{t,h}(\gamma_1, \dots, \gamma_h) &= \mathbb{E}_t[\exp(\gamma'_1 w_{t+1} + \dots + \gamma'_{j-1} w_{t+j-1} + a'_{t+j}(\gamma_j + A_{t,h-j}^h) w_{t+j-1} \\ &\quad + b_{t+j}(\gamma_j + A_{t,h-j}^h) + B_{t,h-j}^h)], \end{aligned}$$

and, therefore, this is true for $j - 1$, with $A_{t,h-j+1}^h$ and $B_{t,h-j+1}^h$ given by formulas (ii) above.

For $j = 1$ we get:

$$\begin{aligned} \varphi_{t,h}(\gamma_1, \dots, \gamma_h) &= \mathbb{E}_t \exp(\gamma'_1 w_{t+1} + A_{t,h-1}^{h'} w_{t+1} + B_{t,h-1}^h) \\ &= \exp(A'_{t,h} w_t + B_{t,h}), \end{aligned}$$

Finally note that if we put $h - j + 1 = i$, formulas (ii) become (i). \square

Proposition 1.5 (Reverse-order multi-horizon Laplace transform). *If the functions a_t and b_t do not depend on t , and if different sequences $(\gamma_1^h, \dots, \gamma_h^h), h = 1, \dots, H$ (say) satisfy $\gamma_{h+1-i}^h = u_i$, for $i = 1, \dots, h$, and for any $h \leq H$, that is if we want to compute (“reverse order” case):*

$$\varphi_{t,h}(u_h, \dots, u_1) = \mathbb{E}_t[\exp(u'_h w_{t+1} + \dots + u'_1 w_{t+h})], \quad h = 1, \dots, H,$$

then we can compute the $\varphi_{t,h}(u_h, \dots, u_1)$ for any t and any $h \leq H$, with only one recursion, i.e. $\varphi_{t,h}(u_h, \dots, u_1) = \exp(A'_h w_t + B_h)$ with:

$$\begin{cases} A_h &= a(u_h + A_{h-1}), \\ B_h &= b(u_h + A_{h-1}) + B_{h-1}, \\ A_0 &= 0, \quad B_0 = 0. \end{cases} \quad (1.19)$$

Proof. According to Lemma 1.5, we have, in this case:

$$\begin{cases} A_i^h &= a(u_i + A_{i-1}^h), \\ B_i^h &= b(u_i + A_{i-1}^h) + B_{i-1}^h, \\ A_0^h &= 0, \quad B_0^h = 0. \end{cases}$$

The previous sequences do not depend on h and are given by Eq. (1.19). \square

Proposition 1.6 (Conditional means and variances of an affine process). *Consider an affine process w_t . Using the notation of Proposition 1.1, we have:*

$$\mathbb{E}_t(w_{t+h}) = (I - \Phi)^{-1}(I - \Phi^h)\mu + \Phi^h w_t \quad (1.20)$$

$$\begin{aligned} \text{Var}_t(w_{t+h}) &= \Sigma(\mathbb{E}_t(w_{t+h-1})) + \Phi \Sigma(\mathbb{E}_t(w_{t+h-2}))\Phi' + \\ &\quad \dots + \Phi^{h-1} \Sigma(w_t) \Phi^{h-1}'. \end{aligned} \quad (1.21)$$

Eq. (1.21) notably shows that $\text{Var}_t(w_{t+h})$ is an affine function of w_t . Indeed $\Sigma(\cdot)$ is an affine function, and the conditional expectations $\mathbb{E}_t(w_{t+h})$ are affine in w_t , as shown by Eq. (1.20).

The unconditional mean and variance of w_t are given by:

$$\begin{cases} \mathbb{E}(w_t) &= (I - \Phi)^{-1}\mu \\ \text{vec}[\text{Var}(w_t)] &= (I_{n^2} - \Phi \otimes \Phi)^{-1} \text{vec}(\Sigma[(I - \Phi)^{-1}\mu]). \end{cases} \quad (1.22)$$

Proof. Eq. (1.20) is easily deduced from Eq. (1.12), using that $\mathbb{E}_t(\varepsilon_{t+k}) = 0$ for $k > 0$.

As regards Eq. (1.21):

$$\text{Var}_t(w_{t+h}) = \text{Var}_t\left(\Sigma(w_{t+h-1})^{\frac{1}{2}}\varepsilon_{t+h} + \dots + \Phi^{h-1}\Sigma(w_t)^{\frac{1}{2}}\varepsilon_{t+1}\right).$$

The conditional expectation at t of all the terms of the sum is equal to zero since, for $i \geq 1$:

$$\begin{aligned} \mathbb{E}_t\left[\Sigma(w_{t+i-1})^{\frac{1}{2}}\varepsilon_{t+i}\right] &= \mathbb{E}_t[\underbrace{\mathbb{E}_{t+i-1}\{\Sigma(w_{t+i-1})^{\frac{1}{2}}\varepsilon_{t+i}\}}_{=\Sigma(w_{t+i-1})^{\frac{1}{2}}\mathbb{E}_{t+i-1}\{\varepsilon_{t+i}\}=0}], \end{aligned}$$

and $\forall i < j$,

$$\text{Cov}_t\left[\Sigma(w_{t+i-1})^{\frac{1}{2}}\varepsilon_{t+i}, \Sigma(w_{t+j-1})^{\frac{1}{2}}\varepsilon_{t+j}\right] = \mathbb{E}_t\left[\Sigma(w_{t+i-1})^{\frac{1}{2}}\varepsilon_{t+i}\varepsilon'_{t+j}\Sigma'(w_{t+j-1})^{\frac{1}{2}}\right],$$

which can be seen to be equal to zero by conditioning on the information available on date $t + j - 1$.

Using the same conditioning, we obtain that:

$$\begin{aligned}
\mathbb{V}ar_t \left[\Phi^{h-j} \Sigma(w_{t+j-1})^{\frac{1}{2}} \varepsilon_{t+j} \right] &= \mathbb{E}_t \left[\Phi^{h-j} \Sigma(w_{t+j-1})^{\frac{1}{2}} \varepsilon_{t+j} \varepsilon'_{t+j} \Sigma'(w_{t+j-1})^{\frac{1}{2}} \Phi^{h-j'} \right] \\
&= \mathbb{E}_t \left[\Phi^{h-j} \Sigma(w_{t+j-1})^{\frac{1}{2}} \mathbb{E}_{t+j-1}(\varepsilon_{t+j} \varepsilon'_{t+j}) \Sigma'(w_{t+j-1})^{\frac{1}{2}} \Phi^{h-j'} \right] \\
&= \Phi^{h-j} \mathbb{E}_t[\Sigma(w_{t+j-1})] \Phi^{h-j'} \\
&= \Phi^{h-j} \Sigma(\mathbb{E}_t[w_{t+j-1}]) \Phi^{h-j'},
\end{aligned}$$

where the last equality results from the fact that the fact that $\Sigma(\cdot)$ is affine (see Eq. (1.14)). \square

Proposition 1.7 (Affine property of the GARCH-type process). *The process $w_t = (w_{1,t}, w_{2,t})$ defined by:*

$$\begin{cases} w_{1,t+1} &= \mu + \varphi w_{1,t} + \sigma_{t+1} \varepsilon_{t+1} \mid \varphi < 1 \\ \sigma_{t+1}^2 &= \omega + \alpha \varepsilon_t^2 + \beta \sigma_t^2 \mid 0 < \beta < 1, \alpha > 0, \omega > 0 \\ w_{2,t} &= \sigma_{t+1}^2, \quad \varepsilon_t \sim i.i.d. \mathcal{N}(0, 1) \end{cases}$$

is affine.

Proof. Note that $w_{2,t}$ is function of $\underline{w_{1,t}}$

$$\begin{aligned}
&\mathbb{E}[\exp(u_1 w_{1,t+1} + u_2 w_{2,t+1}) \mid \underline{w_{1,t}}] \\
&= \exp(u_1 \mu + u_1 \varphi w_{1,t} + u_2 \omega + u_2 \beta w_{2,t}) \mathbb{E}[\exp(u_1 \sigma_{t+1} \varepsilon_{t+1} + u_2 \alpha \varepsilon_{t+1}^2) \mid \underline{w_{1,t}}]
\end{aligned}$$

and, using Lemma 1.4:

$$\begin{aligned}
&\mathbb{E}[\exp(u_1 w_{1,t+1} + u_2 w_{2,t+1}) \mid \underline{w_{1,t}}] \\
&= \exp(u_1 \mu + u_1 \varphi w_{1,t} + u_2 \omega + u_2 \beta w_{2,t}) \exp \left[-\frac{1}{2} \log(1 - 2u_2 \alpha) + \frac{u_1^2 w_{2,t}}{2(1 - 2u_2 \alpha)} \right] \\
&= \exp \left[u_1 \mu + u_2 \omega - \frac{1}{2} \log(1 - 2u_2 \alpha) + u_1 \varphi w_{1,t} + \left(u_2 \beta + \frac{u_1^2}{2(1 - 2u_2 \alpha)} \right) w_{2,t} \right],
\end{aligned}$$

which is exponential affine in $(w_{1,t}, w_{2,t})$. \square

Proposition 1.8 (Computation of truncated conditional moments). *If $\varphi(z) = \mathbb{E}[\exp(z'w)]$, we have:*

$$\mathbb{E}[\exp(u'w) \mathbf{1}_{(v'w < \gamma)}] = \frac{\varphi(u)}{2} - \frac{1}{\pi} \int_0^\infty \frac{\mathcal{I}m[\varphi(u + ivx) \exp(-i\gamma x)]}{x} dx. \quad (1.23)$$

Proof. We want to compute $\tilde{\varphi}_t(u; v, \gamma) = \mathbb{E}_t[\exp(u'w) \mathbf{1}_{(v'w < \gamma)}]$. Let us first note that, for given u and v , $\tilde{\varphi}_t(u; v, \gamma)$ is a positive increasing bounded function of γ and therefore can be seen as the c.d.f. of a positive finite measure on \mathbb{R} , the Fourier transform of which is:

$$\int_{\mathbb{R}} \exp(i\gamma x) d\tilde{\varphi}(u; v, \gamma) = \mathbb{E} \int_{\mathbb{R}} \exp(i\gamma x) d\tilde{\varphi}_w(u; v, \gamma),$$

where, for given w , $\tilde{\varphi}_w(u; v, \gamma)$ is the c.d.f. of the mass point $\exp(u'w)$ at $v'w$. We then get:

$$\begin{aligned} \int_{\mathbb{R}} \exp(i\gamma x) d\tilde{\varphi}(u; v, \gamma) &= \mathbb{E}[\exp(ixv'w) \exp(u'w)] \\ &= \mathbb{E}[\exp(u + ivx)'w] \\ &= \varphi(u + ivx). \end{aligned}$$

Let us now compute $A(x_0, \lambda)$ for any real number λ , with:

$$\begin{aligned} &A(x_0, \lambda) \\ &= \frac{1}{2\pi} \int_{-x_0}^{x_0} \frac{\exp(i\lambda x) \varphi(u - ivx) - \exp(-i\lambda) \varphi(u + ivx)}{ix} dx \\ &= \frac{1}{2\pi} \int_{-x_0}^{x_0} \left[\int_{\mathbb{R}} \frac{\exp[-ix(\gamma - \lambda)] - \exp[ix(\gamma - \lambda)]}{ix} d\tilde{\varphi}(u; v, \gamma) \right] dx \\ &= \frac{1}{2\pi} \int_{\mathbb{R}} \left[\int_{-x_0}^{x_0} \frac{\exp[-ix(\gamma - \lambda)] - \exp[ix(\gamma - \lambda)]}{ix} dx \right] d\tilde{\varphi}(u; v, \gamma). \end{aligned}$$

Now :

$$\begin{aligned} &\frac{1}{2\pi} \int_{-x_0}^{x_0} \frac{\exp[-ix(\gamma - \lambda)] - \exp[ix(\gamma - \lambda)]}{ix} dx \\ &= \frac{-\text{sign}(\gamma - \lambda)}{\pi} \int_{-x_0}^{x_0} \frac{\sin(x |\gamma - \lambda|)}{x} dx \end{aligned}$$

which tends to $-sign(\gamma - \lambda)$ when $x_0 \rightarrow \infty$ (where $sign(\omega) = 1$ if $\omega > 0$, $sign(\omega) = 0$ if $\omega = 0$, $sign(\omega) = -1$ if $\omega < 0$). Therefore:

$$A(\infty, \lambda) = - \int_{\mathbb{R}} sign(\gamma - \lambda) d\tilde{\varphi}(u; v, \gamma) = -\mathbb{E} \int_{\mathbb{R}} sign(\gamma - \lambda) d\tilde{\varphi}_w(u; \theta, \gamma),$$

where $\tilde{\varphi}_w(u; v, \gamma)$ is the c.d.f. of the mass point $\exp(u'w)$ at $v'w$ and

$$\int_{\mathbb{R}} sign(\gamma - \lambda) d\tilde{\varphi}_w(u; v, \gamma) = \begin{cases} \exp(u'w) & \text{if } \lambda < v'w \\ 0 & \text{if } \lambda = v'w \\ -\exp(u'w) & \text{if } \lambda > v'w. \end{cases}$$

Therefore, we have:

$$\begin{aligned} A(\infty, \lambda) &= -\mathbb{E}[\exp(u'w)(1 - \mathbf{1}_{(v'w < \lambda)}) - \exp(u'w)\mathbf{1}_{(v'w < \lambda)}] \\ &= -\varphi(u) + 2\tilde{\varphi}(u; v, \lambda) \end{aligned}$$

and, further,:

$$\tilde{\varphi}(u; , v, \gamma) = \frac{\varphi(u)}{2} + \frac{1}{2}A(\infty, \gamma),$$

where

$$\begin{aligned} \frac{1}{2}A(\infty, \gamma) &= \frac{1}{4\pi} \int_{-\infty}^{\infty} \frac{\exp(i\gamma x)\varphi(u - ivx) - \exp(-i\gamma x)\varphi(u + ivx)}{ix} dx \\ &= \frac{1}{2\pi} \int_o^{\infty} \frac{\exp(i\gamma x)\varphi(u - ivx) - \exp(-i\gamma x)\varphi(u + ivx)}{ix} dx \\ &= -\frac{1}{\pi} \int_o^{\infty} \frac{\mathcal{I}m[\exp(-i\gamma x)\varphi(u + ivx)]}{x} dx, \end{aligned}$$

which leads to Eq. (1.23). □

Chapter 2

Pricing and risk-neutral dynamics

In brief: bla bla bla bla.

2.1 Consumption-based Capital Asset Pricing Model (CCAPM) and stochastic discount factor (SDF)

Consider an economy featuring a single good, whose date- t price is q_t . There is a representative agent with an external income R_t at t , and a portfolio of assets, with an allocation vector α_{t-1} (decided at $t-1$). The vector of date- t prices is p_t . At any date $t+j$, $j = 0, 1, \dots$, the agent will face the budget constraint:

$$q_{t+j}C_{t+j} + \alpha'_{t+j}p_{t+j} = R_{t+j} + \alpha'_{t+j-1}p_{t+j},$$

where C_{t+j} is her consumption on date $t+j$. The representative agent maximizes her expected utility time-separable preferences $\mathbb{E}_t \sum_{j=0}^{\infty} \delta^j U(C_{t+j})$ subject to the budget constraints at $t+j$, $j = 0, 1, \dots$, where U is the utility function and δ is time discount factor or subjective discount factor. Note that the latter is different from SDF.

Replacing C_{t+j} with $[R_{t+j} - (\alpha'_{t+j} - \alpha'_{t+j-1})p_{t+j}]/q_{t+j}$, the objective function becomes:

$$\mathbb{E}_t \sum_{j=0}^{\infty} \delta^j U[R_{t+j}/q_{t+j} - (\alpha'_{t+j} - \alpha'_{t+j-1})p_{t+j}/q_{t+j}].$$

The vector of allocation α_t appears in the first two terms:

$$U[R_t/q_t - (\alpha'_t - \alpha'_{t-1})p_t/q_t] + \delta \mathbb{E}_t U[R_{t+1}/q_{t+1} - (\alpha'_{t+1} - \alpha'_t)p_{t+1}/q_{t+1}].$$

As a result, the first order condition associated with vector α_t reads

$$\frac{p_t}{q_t} \frac{dU(C_t)}{dC} = \delta \mathbb{E}_t \left[\frac{p_{t+1}}{q_{t+1}} \frac{dU(C_{t+1})}{dC} \right],$$

or

$$p_t = \mathbb{E}_t(p_{t+1} \mathcal{M}_{t,t+1}), \quad (2.1)$$

where $\mathcal{M}_{t,t+1}$ is a strictly positive scalar called **stochastic discount factor (SDF)** between dates t and $t+1$. We have:

$$\mathcal{M}_{t,t+1} \equiv \delta \frac{q_t}{q_{t+1}} \frac{\frac{dU(C_{t+1})}{dC}}{\frac{dU(C_t)}{dC}}.$$

Example 2.1 (Power utility function). A standard case is the one of the power utility $U(C) = \frac{C^{1-\gamma}-1}{1-\gamma}$, where $\gamma > 0$ is the coefficient of relative risk aversion.

We have $U'(C) = C^{-\gamma} > 0$ and $U''(C) = -\gamma C^{-\gamma-1} < 0$. As a result, the stochastic discount factor is

$$\begin{aligned} & \mathcal{M}_{t,t+1} \\ &= \frac{q_t}{q_{t+1}} \delta \left(\frac{C_{t+1}}{C_t} \right)^{-\gamma} \\ &= \exp(\log \delta + \log q_t + \gamma \log C_t - \log q_{t+1} - \gamma \log C_{t+1}). \end{aligned} \quad (2.2)$$

The vector of prices p_t satisfies:

$$p_t = \delta q_t C_t^\gamma \mathbb{E}_t \left(\frac{C_{t+1}^{-\gamma}}{q_{t+1}} p_{t+1} \right),$$

or (Euler equation):

$$\mathbb{E}_t \left[\delta \left(\frac{C_{t+1}}{C_t} \right)^{-\gamma} \frac{q_t}{q_{t+1}} \frac{p_{t+1}}{p_t} - 1 \right] = 0.$$

According to Eq. (2.1), for any asset j :

$$p_{j,t} = \mathbb{E}_t(\mathcal{M}_{t,t+1}p_{j,t+1}). \quad (2.3)$$

Using that $p_{j,t+1} = \mathbb{E}_t(\mathcal{M}_{t+1,t+2}p_{j,t+2})$, we get:

$$\begin{aligned} p_{j,t} &= \mathbb{E}_t[\mathbb{E}_{t+1}(p_{j,t+2}\mathcal{M}_{t+1,t+2})\mathcal{M}_{t,t+1}] \\ &= \mathbb{E}_t(\mathcal{M}_{t,t+1}\mathcal{M}_{t+1,t+2}p_{j,t+2}). \end{aligned}$$

This can be generalized as follows:

$$p_{j,t} = \mathbb{E}_t[\mathcal{M}_{t,t+1} \dots \mathcal{M}_{t+h-1,t+h}p_{j,t+h}], \quad \forall h.$$

2.2 Recursive utilities

XXXX

2.3 SDF: Absence of Arbitrage Approach

Consider a period of interest $\mathcal{T} = \{0, 1, 2, \dots, T^*\}$. As in Chapter 1, vector w_t constitutes the new information in the economy at t . The historical, or physical, dynamics of w_t , $f(w_t)$, is defined by $f(w_{t+1}|w_t)$. The physical probability is denoted by \mathbb{P} . $\bar{L}_{2t}, t \in \mathcal{T}$, is the (Hilbert) space of square integrate functions $g(\underline{w}_t)$, and we have $L_{2t} \subset L_{2s}, t < s$.

2.3.1 Existence and unicity of the SDF

Hypothesis 2.1 (Price existence and uniqueness). For any \underline{w}_t , there exists a unique $p_t[g(\underline{w}_s)]$, function of \underline{w}_t , price at t of a payoff $g(\underline{w}_s)$ delivered at $s, \forall t \leq s$.

Hypothesis 2.2 (Linearity and continuity). For all $t < s$, $\underline{w}_t, g_1, g_2$, we have

$$\bullet \quad p_t[\lambda_1 g_1(\underline{w}_s) + \lambda_2 g_2(\underline{w}_s)] = \lambda_1 p_t[g_1(\underline{w}_s)] + \lambda_2 p_t[g_2(\underline{w}_s)],$$

- If $g_n(\underline{w}_s) \xrightarrow[n \rightarrow \infty]{L_{2s}} 0$, then $p_t[g_n(\underline{w}_s)] \xrightarrow[n \rightarrow \infty]{} 0$.

Hypothesis 2.3 (Absence of Arbitrage Opportunity (AAO)). At any t , it is impossible to constitute a portfolio of future payoffs, possibly modified at subsequent dates, such that:

- the price of the portfolio at t is zero,
- payoffs at subsequent dates are ≥ 0 ,
- there is at least one subsequent date s such that the net payoff at s is strictly positive with a non zero conditional probability at t .

Theorem 2.1 (Riesz representation theorem). *Under Assumptions 2.1 and 2.2, for all \underline{w}_t , and $s > t$, there exists $\mathcal{M}_{t,s}(\underline{w}_s) \in L_{2s}$, unique such that, $\forall g(\underline{w}_s) \in L_{2s}$,*

$$p_t[g(\underline{w}_s)] = \mathbb{E}[\mathcal{M}_{t,s}(\underline{w}_s)g(\underline{w}_s)|\underline{w}_t].$$

In particular the price at t of a zero coupon bond maturing at s is $\mathbb{E}(\mathcal{M}_{t,s}|\underline{w}_t)$.

Proposition 2.1 (Positivity of \mathcal{M}). *If Assumption 2.3 is satisfied, then for all t and s , $\mathbb{P}(\mathcal{M}_{t,s} > 0|\underline{w}_t) = 1$.*

Proof. \Leftarrow is obvious. If \Rightarrow was not true, the payoff $\mathbf{1}_{\{\mathcal{M}_{t,s} \leq 0\}}$, at s , would be such that: $\mathbb{P}[\mathbf{1}_{\{\mathcal{M}_{t,s} \leq 0\}} = 1|\underline{w}_t] > 0$ and $p_t[\mathbf{1}_{\{\mathcal{M}_{t,s} \leq 0\}}] = \mathbb{E}_t[\mathcal{M}_{t,s}\mathbf{1}_{\{\mathcal{M}_{t,s} \leq 0\}}] \leq 0$. \square

Proposition 2.2 (Time consistency). *For all $t < r < s$, we have $\mathcal{M}_{t,s} = \mathcal{M}_{t,r}\mathcal{M}_{r,s}$, which implies:*

- $\mathcal{M}_{t,s} = \mathcal{M}_{t,t+1}\mathcal{M}_{t+1,t+2} \dots \mathcal{M}_{s-1,s}$
- $\mathcal{M}_{0,t} = \prod_{j=0}^{t-1} \mathcal{M}_{j,j+1}$ ($\mathcal{M}_{0,t}$ is called **pricing kernel**).

Proof. Using Lemma 2.1 we have:

$$\begin{aligned} p_t(g_s) &= \mathbb{E}(\mathcal{M}_{t,s}g_s|\underline{w}_t) = \mathbb{E}(\mathcal{M}_{t,r}p_r(g_s)|\underline{w}_t) \\ &= \mathbb{E}[\mathcal{M}_{t,r}\mathbb{E}(\mathcal{M}_{r,s}g_s|\underline{w}_r)|\underline{w}_t] = \mathbb{E}(\mathcal{M}_{t,r}\mathcal{M}_{r,s}g_s|\underline{w}_t), \forall g, \forall \underline{w}_t \end{aligned}$$

and, therefore, $\mathcal{M}_{t,s} = \mathcal{M}_{t,r}\mathcal{M}_{r,s}$. \square

Lemma 2.1. *For any payoff g_s at s , $p_t(g_s) = p_t[p_r(g_s)]$.*

Proof. If this was not true, we could construct a sequence of portfolios with a strictly positive payoff at s with zero payoff at any other future date and with price zero at t , contradicting Assumption 2.3. Indeed, assuming, for instance, $p_t(g_s) > p_t[p_r(g_s)]$, the payoff at s is defined by the following strategy: (i) at t : buy $p_r(g_s)$, (short) sell g_s , buy $\frac{p_t(g_s) - p_t[p_r(g_s)]}{\mathbb{E}(\mathcal{M}_{t,s}|\underline{w}_t)}$ zero-coupon bonds maturing at s , at global price zero, (ii) at r : buy g_s and sell $p_r(g_s)$, generating a zero net payoff, (iii) at s , the net payoff is: $g_s - g_s + \frac{p_t(g_s) - p_t[p_r(g_s)]}{\mathbb{E}(\mathcal{M}_{t,s}|\underline{w}_t)} > 0$. \square

Consider an asset whose payoff, on date s , is $g(\underline{w}_s)$. We have, $\forall t < s$:

$$\boxed{p_t[g(\underline{w}_s)] = \mathbb{E}_t[\mathcal{M}_{t,t+1} \dots \mathcal{M}_{s-1,s} g(\underline{w}_s)]}. \quad (2.4)$$

In particular, since $\mathcal{M}_{t,t+1}$ contains 1, the price at t of a zero-coupon with residual maturity one is given by:

$$B(t, 1) := \mathbb{E}_t[\mathcal{M}_{t,t+1}].$$

Denoting by r_t the continuously-compounded interest rate, defined through $B(t, 1) = \exp(-r_t)$, we get

$$r_t = -\log \mathbb{E}_t[\mathcal{M}_{t,t+1}].$$

Definition 2.1 (Bank account). The bank account process R_t is defined by $R_t \equiv \exp(r_0 + \dots + r_{t-1}) = \frac{1}{\mathbb{E}_0[\mathcal{M}_{0,1}] \times \dots \times \mathbb{E}_{t-1}[\mathcal{M}_{t-1,t}]}$.

R_t is the price of an investment initiated on date 0, when it was worth one dollar, and invested on each date at the risk-free rate (for one period).

For any price process p_t , we have $p_t = \mathbb{E}_t(\mathcal{M}_{t,s} p_s)$ (with $s > t$), or $\mathcal{M}_{0,t} p_t = \mathbb{E}_t(\mathcal{M}_{0,s} p_s)$. That is, $\mathcal{M}_{0,t} p_t$ is a martingale. In particular $\mathcal{M}_{0,t} R_t$ is a martingale.

2.3.2 Exponential affine SDF

A specific (tractable) case is that of exponential affine SDF. Assume that

$$\mathcal{M}_{t,t+1}(\underline{w}_{t+1}) = \exp[\alpha_t(\underline{w}_t)' \underline{w}_{t+1} + \beta_t(\underline{w}_t)]$$

where α_t defines the *prices of risk* or *sensitivity* vector. Using $\mathbb{E}_t[\mathcal{M}_{t,t+1}] = \exp(-r_t) = \exp[\psi_t(\alpha_t) + \beta_t]$, we get:

$$\mathcal{M}_{t,t+1} = \exp[-r_t + \alpha'_t w_{t+1} - \psi_t(\alpha_t)]. \quad (2.5)$$

Example 2.2 (CCAPM/Power utility case). In the CCAPM-power-utility case (see Example 2.1), we have (Eq. (2.2)):

$$\mathcal{M}_{t,t+1} = \exp(\log \delta + \log q_t + \gamma \log C_t - \log q_{t+1} - \gamma \log C_{t+1}),$$

where q_t is the price of the consumption good, C_t is the quantity consumed at t and δ is the intertemporal discount rate.

Hence, in that case, $\mathcal{M}_{t,t+1}$ is exponential affine in $w_{t+1} = (\log q_{t+1}, \log C_{t+1})'$ (and its first lag).

2.4 The risk-neutral (R.N.) dynamics

The historical Dynamics is characterized by $f(w_{T^*})$, or by the sequence of conditional p.d.f. $f_{t+1}(w_{t+1}|\underline{w}_t)$, or $f_{t+1}(w_{t+1})$, with respect to (w.r.t.) some measure μ .

We define the conditional risk-neutral p.d.f. w.r.t. the conditional historical probability. For that, we employ the Radon-Nikodym derivative $d_{t+1}^{\mathbb{Q}}(w_{t+1}|\underline{w}_t)$:¹

$$d_{t+1}^{\mathbb{Q}}(w_{t+1}|\underline{w}_t) = \frac{\mathcal{M}_{t,t+1}(\underline{w}_{t+1})}{\mathbb{E}[\mathcal{M}_{t,t+1}(\underline{w}_{t+1})|\underline{w}_t]},$$

or

$$d_{t+1}^{\mathbb{Q}}(w_{t+1}) = \frac{\mathcal{M}_{t,t+1}}{\mathbb{E}_t(\mathcal{M}_{t,t+1})} = \exp(r_t) \mathcal{M}_{t,t+1}.$$

In this context, the risk neutral conditional p.d.f. is:

$$\begin{aligned} f_{t+1}^{\mathbb{Q}}(w_{t+1}) &= f_{t+1}(w_{t+1}) d_{t+1}^{\mathbb{Q}}(w_{t+1}) \\ &= f_{t+1}(w_{t+1}) \mathcal{M}_{t,t+1}(\underline{w}_{t+1}) \exp[r_t(\underline{w}_t)]. \end{aligned} \quad (2.6)$$

¹Of course, the conditional historical p.d.f. with respect to the conditional risk-neutral (R.N.) p.d.f. is: $d_{t+1}(w_{t+1}) = \frac{1}{d_{t+1}^{\mathbb{Q}}(w_{t+1})}$ or $d_{t+1}(w_{t+1}) = \frac{\exp(-r_t)}{\mathcal{M}_{t,t+1}}$.

The p.d.f. of \mathbb{Q} w.r.t. the historical dynamics \mathbb{P} is:

$$\xi_{T^*} = \frac{d\mathbb{Q}}{d\mathbb{P}} = \Pi_{t=0}^{T^*-1} d_{t+1}^{\mathbb{Q}}(w_{t+1}) = \Pi_{t=0}^{T^*-1} \exp(r_t) \mathcal{M}_{t,t+1},$$

and the p.d.f. of the R.N. distribution of \underline{w}_t , w.r.t. the corresponding historical distribution is:

$$\xi_t = \Pi_{\tau=1}^{t-1} d_{\tau+1}^{\mathbb{Q}}(w_{\tau+1}) = \mathbb{E}_t \left(\frac{d\mathbb{Q}}{d\mathbb{P}} \right) = \mathbb{E}_t \xi_{T^*}.$$

Therefore, ξ_t is a \mathbb{P} -martingale.²

Consider the date- t price of a payoff $g(\underline{w}_s)$ at time $s > t$. An equivalent form of the pricing formula (2.4) is:

$$\begin{aligned} p_t[g(\underline{w}_s)] &= \mathbb{E}_t[\mathcal{M}_{t,t+1} \dots \mathcal{M}_{s-1,s} g(\underline{w}_s)] \\ &= \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{s-1}) g(\underline{w}_s)], \end{aligned}$$

or, with simpler notations:

$$p_t = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{s-1}) p_s] = \mathbb{E}_t^{\mathbb{Q}} \left(\frac{R_t}{R_s} p_s \right),$$

where R_t is the *bank account*.

We also have $p_t/R_t = \mathbb{E}_t^{\mathbb{Q}}(p_s/R_s)$, that is, p_t/R_t is a \mathbb{Q} -martingale. In particular $p_t = \exp(-r_t) \mathbb{E}_t^{\mathbb{Q}}(p_{t+1})$, or, using the arithmetic return of any payoff $(p_{t+1} - p_t)/p_t$, and the arithmetic return of the riskless asset $r_{A,t+1} = \exp(r_t) - 1$, we get:

$$\mathbb{E}_t^{\mathbb{Q}} \left(\frac{p_{t+1} - p_t}{p_t} \right) = r_{A,t}.$$

Moreover the excess arithmetic return process $(p_{t+1} - p_t)/p_t - r_{A,t}$ is a \mathbb{Q} -martingale difference and, therefore, \mathbb{Q} -serially uncorrelated.

²Indeed:

$$\mathbb{E}_t \left(\frac{d\mathbb{Q}}{d\mathbb{P}} \right) = \Pi_{\tau=1}^{t-1} d_{\tau+1}^{\mathbb{Q}}(w_{\tau+1}) \mathbb{E}_t \left(d_{t+1}^{\mathbb{Q}}(w_{t+1}) \dots d_{T^*}^{\mathbb{Q}}(w_{T^*}) \right).$$

Let us consider the case of an exponential affine SDF $\mathcal{M}_{t,t+1} = \exp(\alpha'_t w_{t+1} + \beta_t)$:

$$d_{t+1}^{\mathbb{Q}}(w_{t+1}) = \frac{\mathcal{M}_{t,t+1}}{\mathbb{E}_t(\mathcal{M}_{t,t+1})} = \frac{\exp(\alpha'_t w_{t+1} + \beta_t)}{\exp[\psi_t(\alpha_t) + \beta_t]} = \exp[\alpha'_t w_{t+1} - \psi_t(\alpha_t)]$$

We then have that $d_{t+1}^{\mathbb{Q}}(w_{t+1})$ is also exponential affine. Moreover:

$$f_{t+1}^{\mathbb{Q}}(w_{t+1}) = \frac{f_{t+1}(w_{t+1}) \exp(\alpha'_t w_{t+1})}{\varphi_t(\alpha_t)}.$$

The previous equation shows that $f_{t+1}^{\mathbb{Q}}$ is the Esscher transform of f_{t+1} evaluated at α_t .

Let us now consider the Laplace transform of the conditional R.N. probability, $\varphi_t^{\mathbb{Q}}(u|w_t)$, also denoted by $\varphi_t^{\mathbb{Q}}(u)$. We have:

$$\begin{aligned} \varphi_t^{\mathbb{Q}}(u) &= \mathbb{E}_t^{\mathbb{Q}} \exp(u' w_{t+1}) \\ &= \mathbb{E}_t \exp[(u + \alpha_t)' w_{t+1} - \psi_t(\alpha_t)] \\ &= \exp[\psi_t(u + \alpha_t) - \psi_t(\alpha_t)] = \frac{\varphi_t(u + \alpha_t)}{\varphi_t(\alpha_t)}. \end{aligned}$$

Hence:

$$\boxed{\psi_t^{\mathbb{Q}}(u) = \psi_t(u + \alpha_t) - \psi_t(\alpha_t)}. \quad (2.7)$$

We check that, if $\alpha_t = 0$, $\psi_t^{\mathbb{Q}} = \psi_t$ (since $\psi_t(0) = 0$).

Moreover, putting $u = -\alpha_t$ in the expression of $\psi_t^{\mathbb{Q}}(u)$ we get $\psi_t^{\mathbb{Q}}(-\alpha_t) = -\psi_t(\alpha_t)$, and, replacing u by $u - \alpha_t$, we get:

$$\boxed{\psi_t(u) = \psi_t^{\mathbb{Q}}(u - \alpha_t) - \psi_t^{\mathbb{Q}}(-\alpha_t)}.$$

Also:

$$\begin{cases} d_{t+1}(w_{t+1}) &= \exp[-\alpha'_t(w_{t+1}) - \psi_t^{\mathbb{Q}}(-\alpha_t)] \\ d_{t+1}^{\mathbb{Q}}(w_{t+1}) &= \exp[\alpha'_t(w_{t+1}) + \psi_t^{\mathbb{Q}}(-\alpha_t)]. \end{cases}$$

2.5 Typology of econometric asset-pricing models

Definition 2.2 (Econometric Asset Pricing Model (EAPM)). An Econometric Asset Pricing Model (EAPM) is defined by the following functions:

- $r_t(\underline{w}_t)$,
- $f(w_{t+1}|\underline{w}_t)$ [or $\psi_t(u)$],
- $\mathcal{M}_{t,t+1}(\underline{w}_{t+1})$,
- $f^{\mathbb{Q}}(w_{t+1}|\underline{w}_t)$ [or $\psi_t^{\mathbb{Q}}(u)$].

The previous functions have to be specified and parameterized. They are linked by:

$$f^{\mathbb{Q}}(w_{t+1}|\underline{w}_t) = f(w_{t+1}|\underline{w}_t)\mathcal{M}_{t,t+1}(\underline{w}_{t+1})\exp[r_t(\underline{w}_t)].$$

In the following, we present three ways of specifying an EAPM:

1. the direct modelling,
2. the R.N.-constrained direct modelling (or mixed modelling),
3. the back modelling.

We focus on the case where $\mathcal{M}_{t,t+1}$ is exponential affine, as in Eq. (2.5):

$$\mathcal{M}_{t,t+1}(\underline{w}_{t+1}) = \exp \left\{ -r_t(\underline{w}_t) + \alpha'_t(\underline{w}_t)w_{t+1} - \psi_t[\alpha_t(w_t)] \right\}.$$

Once the short-term rate function $r_t(\underline{w}_t)$ is specified, we have to specify ψ_t , α_t , and $\psi_t^{\mathbb{Q}}$, that are linked by Eq. (2.7).

In all approaches, we have to specify the status of the short rate. The short rate r_t is a function of \underline{w}_t , this function may be known or unknown by the econometrician. It is known in two cases: (a) r_t is exogenous ($r_t(\underline{w}_t)$ does not depend on \underline{w}_t) or (b) r_t is a component of w_t . By contrast, if the function $r_t(\underline{w}_t)$ is unknown, it has to be specified parametrically:

$$\left\{ r_t(\underline{w}_t, \tilde{\theta}), \tilde{\theta} \in \tilde{\Theta} \right\},$$

where $r_t(\bullet, \bullet)$ is a known function.

Let us now detail the three steps on which each of the three ways of defining an EAPM is based.

2.5.1 The direct modelling

- **Step 1 – Specification of the historical dynamics.** We choose a parametric family for the conditional historical Log-Laplace transform $\psi_t(u|\underline{w}_t)$: $\{\psi_t(u|\underline{w}_t; \theta_1), \theta_1 \in \Theta_1\}$.
- **Step 2 – Specification of the SDF.** Considering the affine specification of as Eq. (2.5), that is:

$$\mathcal{M}_{t,t+1}(\underline{w}_{t+1}) = \exp \left\{ -r_t(\underline{w}_t, \tilde{\theta}) + \alpha'_t(\underline{w}_t)w_{t+1} - \psi_t[\alpha_t(\underline{w}_t)|\underline{w}_t; \theta_1] \right\},$$

we need to specify functions $r_t(\underline{w}_t, \tilde{\theta})$ and $\alpha_t(\underline{w}_t)$. Assume that $\alpha_t(\underline{w}_t)$ belongs to a parametric family: $\{\alpha_t(\underline{w}_t; \theta_2), \theta_2 \in \Theta_2\}$. We then have:

$$\mathcal{M}_{t,t+1}(\underline{w}_{t+1}, \theta) = \exp \left\{ -r_t(\underline{w}_t, \tilde{\theta}) + \alpha'_t(\underline{w}_t, \theta_2)w_{t+1} - \psi_t[\alpha_t(\underline{w}_t, \theta_2)|\underline{w}_t; \theta_1] \right\},$$

where $\theta = (\tilde{\theta}', \theta'_1, \theta'_2)' \in \tilde{\Theta} \times \Theta_1 \times \Theta_2 = \Theta$.

- **Step 3 – Internal consistency conditions (ICC).** For any payoff $g(\underline{w}_s)$ settled at $s > t$, with price $p(\underline{w}_t)$ at t which is a known function of \underline{w}_t , we must have:

$$p(\underline{w}_t) = \mathbb{E} \left\{ \mathcal{M}_{t,t+1}(\theta) \dots \mathcal{M}_{s-1,s}(\theta) g(\underline{w}_s) | \underline{w}_t, \theta_1 \right\} \forall \underline{w}_t, \theta.$$

These ICC pricing conditions may imply strong constraints on θ . For instance, when components of w_t are returns of some assets: if $w_{1,t} = \log(p_{1,t}/p_{1,t-1})$, then we must have $\mathbb{E}_t[\mathcal{M}_{t,t+1} \exp(e'_1 w_{t+1})] = 1$ (Euler equation). Or, in the case of interest rates with various maturities: if $w_{1,t} = -1/h \log B(t, h)$, then we must have $e'_1 w_t = -1/h \log \mathbb{E}_t(\mathcal{M}_{t,t+1} \times \dots \times \mathcal{M}_{t+h-1,t+h})$.

The previous three steps imply the specification of the R.N. dynamics (according to Eq. (2.7)):

$$\psi^{\mathbb{Q}}(u|\underline{w}_t, \theta_1, \theta_2) = \psi_t \left[u + \alpha_t(\underline{w}_t, \theta_2) | \underline{w}_t, \theta_1 \right] - \psi_t \left[\alpha_t(\underline{w}_t, \theta_2) | \underline{w}_t, \theta_1 \right].$$

2.5.2 The R.N.-constrained direct modelling (or mixed modelling)

- **Step 1 – Specification of the physical dynamics.** We select a family $\{\psi_t(u|\underline{w}_t, \theta_1), \theta_1 \in \Theta_1\}$.

- **Step 2 – Specification of the risk-neutral dynamics.** We select a family $\{\psi_t^{\mathbb{Q}}(u|\underline{w}_t, \theta^*), \theta^* \in \Theta^*\}$ and, possibly, $\{r_t(\underline{w}_t, \tilde{\theta}), \tilde{\theta} \in \tilde{\Theta}\}$.
- **Step 3 – Internal Consistency Conditions (ICC).** Once the parameterization $(\tilde{\theta}, \theta_1, \theta^*) \in \tilde{\Theta} \times \Theta_1^* \times \Theta^*$ is defined, ICCs may be imposed. For instance, if $w_{1,t} = \log(p_{1,t}/p_{1,t-1})$, then we must have $\exp(-r_t)\mathbb{E}_t^{\mathbb{Q}} \exp(e'_1 w_{t+1}) = 1$. Or if $w_{1,t} = B(t, h)$, then $e'_1 w_t = \mathbb{E}_t^{\mathbb{Q}} \exp(-r_t - \dots - r_{t+h-1})$.

The SDF is a by-product. If we want an exponential affine SDF, for any pair $(\psi_t^{\mathbb{Q}}, \psi_t)$ belonging to these families, there must exist a unique function $\alpha_t(\underline{w}_t)$ denoted by $\alpha_t(w_t; \theta_1, \theta^*)$, and satisfying:

$$\psi_t^{\mathbb{Q}}(u|\underline{w}_t) = \psi_t \left[u + \alpha_t(w_t)|\underline{w}_t \right] - \psi_t \left[\alpha_t(\underline{w}_t)|\underline{w}_t \right].$$

2.5.3 Back modelling (based on three steps)

- **Step 1 – Specification of the R.N. dynamics,** and possibly of $r_t(\underline{w}_t)$: $\psi_t^{\mathbb{Q}}(u|\underline{w}_t; \theta_1^*)$.
- **Step 2 – Internal consistency conditions (ICC),** if relevant, are taken into account:

$$p(\underline{w}_t) = \mathbb{E}_t^{\mathbb{Q}} \left[\exp(-r_t(\underline{w}_t, \tilde{\theta}) - \dots - r_{s-1}(\underline{w}_s, \tilde{\theta})) g(\underline{w}_s) | \underline{w}_t, \theta_1^* \right], \\ \forall \underline{w}_t, \tilde{\theta}, \theta_1^*.$$

- **Step 3 – Choice of the specification of the prices of risk.** One chooses function $\alpha_t(w_t)$ without any constraint; this amounts to defining the family $\{\alpha_t(\underline{w}_t, \theta_2^*), \theta_2^* \in \Theta_2^*\}$.

The historical dynamics is obtained as a by-product. Indeed:

$$\psi_t(u|\underline{w}_t; \theta_1^*, \theta_2^*) = \psi_t^{\mathbb{Q}} \left[u - \alpha_t(\underline{w}_t, \theta_2^*) | \underline{w}_t; \theta_1^* \right] - \psi_t^{\mathbb{Q}} \left[-\alpha_t(\underline{w}_t, \theta_2^*) | \underline{w}_t, \theta_1^* \right].$$

Chapter 3

The term structure or risk-free yields

3.1 Introduction

Risk-free yields are the yields-to-maturity associated with bonds that carry no default and/or liquidity risks. Bonds issued by sovereign entities with top credit quality are usually considered to be risk-free.

An important share of the term-structure literature pertains to the modelling of risk-free yields. Some models involve macro factors in w_t (Ang and Piazzesi, 2003); some do not (Duffie and Singleton, 1997). The latter are sometimes called *yield-only* models.

The basic pricing formula of a risk-free zero coupon bond is (see Eq. (1.9) in Example 1.20):

$$B(t, h) = \exp(-r_t) \mathbb{E}_t^{\mathbb{Q}} \exp(-r_{t+1} - \dots - r_{t+h-1}) \quad (3.1)$$

$$R(t, h) = -\frac{1}{h} \log B(t, h). \quad (3.2)$$

Term structure models are often used to extract **term premiums** from observed yields-to-maturity. Term premiums are those components of yields that would not exist if investors were not risk-averse.

If agents were not risk averse, i.e., under the **Expectation Hypothesis (EH)**, we would have $\mathcal{M}_{t,t+1} = \exp(-r_t)$. We would then have “ $\mathbb{P} \equiv \mathbb{Q}$ ”.

and $B(t, h)$ would become:

$$\exp(-r_t) \mathbb{E}_t \exp(-r_{t+1} - \dots - r_{t+h-1}).. \quad (3.3)$$

And the maturity- h yield-to-maturity would then be:

$$\begin{aligned} R^{EH}(t, h) &= -\frac{1}{h} \log (\mathbb{E}_t \exp(-r_t - \dots - r_{t+h-1})) \\ &\approx \frac{1}{h} \mathbb{E}_t (r_t + \dots + r_{t+h-1}). \end{aligned} \quad (3.4)$$

The term premium is given by:

$$\begin{aligned} TP_{t,h} &= \underbrace{-\frac{1}{h} \log \mathbb{E}_t^{\mathbb{Q}} \exp(-r_{t+1} - \dots - r_{t+h-1})}_{=R(t,h)} - \\ &\quad \underbrace{-\frac{1}{h} \log \mathbb{E}_t \exp(-r_{t+1} - \dots - r_{t+h-1})}_{=R^{EH}(t,h)}. \end{aligned} \quad (3.5)$$

What is the economic meaning of the term premium? Under EH, investors are willing to buy a maturity- h bond as long as its expected return is—up to Jensen’s inequality—equal to the average of future short-term rates. (Hence the definition of $R^{EH}(t, h)$, see Eq. (3.4).) When $TP_{t,h} > 0$, investors are willing to buy the maturity- h bond only if its return is, on average, higher than expected future short-term rates; this corresponds to a situation where investors consider that long-term bonds tend to lose value in *bad states of the world* (i.e., states of high marginal utility).

Risk premium: According to Eq. (2.3), the price of any asset j satisfies:

$$p_{jt} = \mathbb{E}_t(\mathcal{M}_{t,t+1} p_{j,t+1}).$$

The previous equation rewrites:

$$p_{jt} = \text{Cov}_t(\mathcal{M}_{t,t+1}, p_{j,t+1}) + \mathbb{E}_t(\mathcal{M}_{t,t+1}) \mathbb{E}_t(p_{j,t+1})$$

or

$$p_{jt} = \underbrace{\exp(-r_t) \mathbb{E}_t(p_{j,t+1})}_{=p_{jt}^{EH}} + \underbrace{\text{Cov}_t(\mathcal{M}_{t,t+1}, p_{j,t+1})}_{\text{Risk premium}}. \quad (3.6)$$

If investors were not risk-averse, then we would have $p_{jt} = p_{jt}^{EH}$. The S.D.F. is high (resp. low) in bad (resp. good) states of the world (states of high marginal utility in the equilibrium approach). Hence, we have $p_{jt} < p_{jt}^{EH}$ if asset j tends to pay less in bad states of the world (i.e., if $\text{Cov}_t(\mathcal{M}_{t,t+1}, p_{j,t+1}) < 0$).

3.2 Swap rates and forward rates

3.2.1 Swap rates

Bonds issued by top-rated (Aaa/AAA) countries are often considered to be risk-free. Because of call-margins mechanisms, swap rates are also used as risk-free benchmarks (Duffie and Stein, 2015).

Definition 3.1 (Interest Rate Swap (IRS)). In an Interest Rate Swap (IRS), a *fixed-rate payer* agrees to provide the *fixed-rate receiver* with a sequence of cash flows that are determined at the negotiation date of the swap, and at predetermined dates. These cash flows constitute the *fixed leg* of the swap. Conversely, the fixed-rate receiver provides the fixed-rate payer with cash-flows that depend on future values of a reference rate; this is the *floating rate* of the swap.

More precisely, the dates of payment are of the form $t + \tau, t + 2\tau, \dots, t + n\tau$, where τ is a period expressed in years (typically 1/2 or 1/4) and n is the number of payments. The maturity, or *tenor*, of the swap contract is $h = n\tau$.

The payoffs of the fixed leg are τS , where S is the annualized payment (or *swap rate*). On date $t + j\tau$, the payoff of the floating leg is $\tau L(t + (j-1)\tau, \tau)$, where L , the annualized linear rate is given by:

$$L(t + (j-1)\tau, \tau) = \frac{1 - B(t + (j-1)\tau, \tau)}{\tau B(t + (j-1)\tau, \tau)},$$

where $B(t + (j-1)\tau, \tau)$ is the price, at date $t + (j-1)\tau$ of a bond of maturity τ .

On the negotiation date, the values of the fixed and floating legs are identical, so that the value of the swap is zero.

At t , the price of the fixed leg is:

$$\sum_{j=1}^n \tau S B(t, j\tau).$$

Let us turn to the price of the floating leg. The payoff at date $t + j\tau$, that is $\frac{1-B(t+(j-1)\tau, \tau)}{B(t+(j-1)\tau, \tau)}$ is known on date $t + (j-1)\tau$, so its price at date $t + (j-1)\tau$ is $1 - B(t + (j-1)\tau, \tau)$, and its price at t therefore is $B(t, (j-1)\tau) - B(t, j\tau)$. Summing over $j = 1, \dots, n$, the date- t price of the floating leg is $1 - B(t, n\tau)$ (independent of the payment dates).

Since the price of the contract is zero at date t (by definition of the swap), we must have:

$$\sum_{j=1}^n \tau S B(t, j\tau) = 1 - B(t, n\tau) \Rightarrow S = \frac{1 - B(t, n\tau)}{\tau \sum_{j=1}^n B(t, j\tau)},$$

or

$$S(t, h) = \frac{1 - B(t, h)}{\tau \sum_{j=1}^{h/\tau} B(t, j\tau)}.$$

Note that all the terms appearing in the previous formula are available in closed-form in the context of an affine model (see Example 1.20).

3.2.2 Forward rates

Definition 3.2 (Forward Rate Agreement (FRA)). An interest rate forward contract is a contract in which the rate to be paid or received on a specific obligation for a set period, beginning in the future, is set at contract initiation.

Denote by $f(t, h_1, h_2)$ the forward interest rate, set on date t , for the period between $t + h_1$ and $t + h_2$. Let us relate $f(t, h_1, h_2)$ to bond rates. For that, consider two strategies (decided on date t):

1. Buy a zero-coupon bond of maturity h_2 (price $B(t, h_2)$) and sell zero-coupon bonds of maturity h_1 for the same amount (yielding a payoff of $B(t, h_2)/B(t, h_1)$ on date $t + h_1$).

2. Enter a forward rate agreement between dates $t+h_1$ and $t+h_2$, whereby you receive 1 on date $t+h_2$.

These two strategies deliver the same payoffs on date t (the payoff is zero) and on date $t+h_2$ (the payoff is 1). By absence of arbitrage, the payoffs on date $t+h_1$ have to be the same. Therefore

$$\begin{aligned} \exp(-(h_2 - h_1)f(t, h_1, h_2)) &= B(t, h_2)/B(t, h_1) \\ \Rightarrow f(t, h_1, h_2) &= \frac{1}{h_2 - h_1}(\log[B(t, h_1)] - \log[B(t, h_2)]), \end{aligned}$$

which gives:

$$\boxed{f(t, h_1, h_2) = \frac{h_2 R(t, h_2) - h_1 R(t, h_1)}{h_2 - h_1}.} \quad (3.7)$$

In an affine model (where $R(t, h)$ is an affine function of the state vector w_t), forward rates are linear in the state vector w_t .

3.3 The Affine Case

3.3.1 Affine yields

In this subsection, we consider the case where the state vector w_t is affine under both \mathbb{P} and \mathbb{Q} . If the nominal short-term rate is affine in w_t , i.e., if $r_t = \omega_0 + \omega'_1 w_t$, then:

$$\begin{aligned} B(t, h) &= \mathbb{E}_t^{\mathbb{Q}} \exp(-r_t - \dots - r_{t+h-1}) \\ &= \exp(-h\omega_0 - \omega'_1 w_t) \mathbb{E}_t^{\mathbb{Q}} \exp(-\omega'_1 w_{t+1} - \dots - \omega'_1 w_{t+h-1}). \end{aligned}$$

The (blue) expectation is easily computed using the recursive equations of Lemma 1.5 (see Example 1.20), leading to:

$$R(t, h) = -\frac{1}{h} \log B(t, h) = A'_h w_t + B_h. \quad (3.8)$$

It is easily seen that we can also get:

$$R^{EH}(t, h) = A_h^{EH'} w_t + B_h^{EH}. \quad (3.9)$$

Moreover, if inflation is also affine in w_t , i.e., if $\pi_t = \bar{\omega}_0 + \bar{\omega}'_1 w_t$, then real yields are given by:

$$\bar{B}(t, h) = \mathbb{E}_t^{\mathbb{Q}} \exp(-r_t - \dots - r_{t+h-1} + \pi_{t+1} + \dots + \pi_{t+h})$$

(see Example 1.22) which also leads to:

$$\bar{R}(t, h) = -\frac{1}{h} \log \bar{B}(t, h) = \bar{A}'_h w_t + \bar{B}_h. \quad (3.10)$$

Eqs. (3.8) and (3.9) imply that term premiums are affine in w_t (see Eq. (3.5)). Specifically:

$$TP(t, h) = R(t, h) - E^{EH}(t, h) = B_h - B_h^{EH} + (A_h - A_h^{EH})' w_t.$$

Expected excess returns resulting from holding zero-coupon bonds are also affine in w_t . Indeed, holding a maturity- h zero-coupon bond for one period provides the following expected gross return:

$$\mathbb{E}_t \left(\frac{B(t+1, h-1)}{B(t, h)} \right) = \mathbb{E}_t (\exp(B_{h-1} - B_h + A'_{h-1} w_{t+1} - A'_h w_t)),$$

which is clearly exponential affine in w_t if w_t is an affine process. Therefore, the expected excess return, that is:

$$\log \mathbb{E}_t \left(\frac{B(t+1, h-1)}{B(t, h)} \right) - r_t$$

is also affine in w_t in this context. The fact that excess returns are affine in this context is exploited in the estimation approach proposed by Adrian et al. (2013).

Moreover, *conditional expectations* of future interest rates (real or nominal) and of term premiums are also affine in w_t . In particular:

$$\mathbb{E}_t[R(t+k, h)] = \mathbb{E}_t[A'_h w_{t+k} + B_h] = A'_h \mathbb{E}_t(w_{t+k}) + B_h, \quad (3.11)$$

and $\mathbb{E}_t(w_{t+k})$ is affine in w_t (see Eq. (1.15)). This can notably be used at the estimation stage, if one wants to fit survey data (see Section XXX).

Similarly, *conditional variances* of future interest rates (real or nominal) and of term premiums are affine in w_t . In particular:

$$\mathbb{V}ar_t[R(t+k, h)] = \mathbb{V}ar_t[A'_h w_{t+k} + B_h] = A'_h \mathbb{V}ar_t(w_{t+k}) A_h, \quad (3.12)$$

where the components of $\mathbb{V}ar_t(w_{t+k})$ (and therefore $\mathbb{V}ar_t[R(t+k, h)]$) is affine in w_t (see Eq. (1.16)). This can also be used at the estimation stage, if one wants to fit (proxies of) conditional variances (Monfort et al., 2017).

3.3.2 Maximum Sharpe ratio

In an affine model, the maximum Sharpe ratio is easily computed. This has been noted early by Duffee (2010) for the Gaussian model; Gouriéroux et al. (2021) and Pallara and Renne (2023) use it in more sophisticated affine models.

Let us derive the maximum Sharpe ratio in the context of a genral affine framework. Eq. (3.6) implies that

$$\underbrace{\mathbb{E}_t \left(\frac{p_{j,t+1}}{p_{j,t}} - \exp(r_t) \right)}_{=xs_{j,t+1}, \text{ excess return}} = -\exp(r_t) \text{Cov}_t \left(\mathcal{M}_{t,t+1}, \frac{p_{j,t+1}}{p_{j,t}} \right),$$

and, using $|\text{Cov}(X, Y)| \leq \sqrt{\text{Var}(X)\text{Var}(Y)}$, we get the Hansen and Jagannathan (1991) bound:

$$\underbrace{\frac{\mathbb{E}_t(xs_{j,t+1})}{\sqrt{\text{Var}_t(xs_{j,t+1})}}}_{\text{Sharpe ratio}} \leq \underbrace{\frac{\sqrt{\text{Var}_t(\mathcal{M}_{t,t+1})}}{\mathbb{E}_t(\mathcal{M}_{t,t+1})}}_{\text{Maximum Sharpe ratio}}. \quad (3.13)$$

If the S.D.F. is given by $\mathcal{M}_{t,t+1} = \exp[-r_t + \alpha'_t w_{t+1} - \psi_t(\alpha_t)]$ (Eq. (2.5)), and using that $\mathbb{E}_t(\mathcal{M}_{t,t+1}^2) = \exp(-2r_t + \psi_t(2\alpha_t) - 2\psi_t(\alpha_t))$ we get:

$$\text{Maximum Sharpe ratio} = \sqrt{\exp(\psi_t(2\alpha_t) - 2\psi_t(\alpha_t)) - 1}.$$

3.4 Gaussian Affine Term Structure Model

The Gaussian Affine Term Structure Model (GATSM) is a *Workhorse* model, widely used in academic and economic-policy circles. In a GATSM, w_t follows a Gaussian vector autoregressive model, and is therefore affine under \mathbb{P} . The S.D.F. is exponential affine in w_t , which implies that it is also affine under \mathbb{Q} (see Subsection XXX). Since the components of w_t are valued in \mathbb{R} , one can easily introduce macro-factors among the state variables.

Let us be more specific. The state vector w_t follows:

$$w_{t+1} = \mu + \Phi w_t + \Sigma^{1/2} \varepsilon_{t+1}, \text{ where } \varepsilon_t \sim i.i.d. \mathcal{N}(0, Id). \quad (3.14)$$

(The fact that we consider a VAR(1) process is without loss of generality since a VAR(p) admits a VAR(1) companion representation.)

This implies the following Laplace transform for w_t (see Example ??):

$$\psi_t(u) = \log \mathbb{E}_t(\exp(u'w_{t+1})|\underline{w}_t) = u'\mu + u'\Phi w_t + \frac{1}{2}u'\Sigma' u.$$

Using the notations of Eq. (2.5), the s.d.f. is defined as:

$$\mathcal{M}_{t,t+1} = \exp(-r_t + \alpha_t' w_{t+1} - \psi_t(\alpha_t)), \text{ where } \alpha_t = \alpha_0 + \alpha_1' w_t.$$

In that case, using Eq. (2.7), we get:

$$\begin{aligned} \psi_t^{\mathbb{Q}}(u) &= \psi_t(u + \alpha_t) - \psi_t(\alpha_t) \\ &= (u + \alpha_t)'\mu + (u + \alpha_t)'\Phi w_t + \frac{1}{2}(u + \alpha_t)'\Sigma(u + \alpha_t) \\ &\quad - \left(\alpha_t'\mu + \alpha_t'\Phi w_t + \frac{1}{2}\alpha_t'\Sigma\alpha_t \right) \\ &= u'(\mu + \Sigma\alpha_0) + u'(\Phi + \Sigma\alpha_1')w_t + \frac{1}{2}u'\Sigma' u. \end{aligned}$$

The \mathbb{Q} -dynamics of w_t is (from Example ??):

$$w_{t+1} = \mu + \Sigma\alpha_0 + (\Phi + \Sigma\alpha_1')w_t + \Sigma^{1/2}\varepsilon_{t+1}^*, \text{ where } \varepsilon_t^* \sim i.i.d.\mathcal{N}^{\mathbb{Q}}(0, Id).$$

Hence, w_t also follows a VAR process under \mathbb{Q} since the previous equation rewrites:

$$w_{t+1} = \mu^{\mathbb{Q}} + \Phi^{\mathbb{Q}}w_t + \Sigma^{1/2}\varepsilon_{t+1}^*,$$

where $\mu^{\mathbb{Q}} = \mu + \Sigma\alpha_0$ and $\Phi^{\mathbb{Q}} = \Phi + \Sigma\alpha_1'$.

With affine specifications of the nominal short term rate ($r_t = \omega_0 + \omega_1'w_t$) and of the inflation rate ($\pi_t = \bar{\omega}_0 + \bar{\omega}_1'w_t$), we obtain affine formulas for nominal and real yields of any maturity (Eqs. (3.8) and (3.10)).

Example 3.1 (Kim and Wright (2005)). This model is a three-factor *yield-only model* (no macro variables, except inflation in one variant of the model), where the short-term rate reads $r_t = \omega_0 + \omega_{1,1}w_{1,t} + \omega_{1,2}w_{2,t} + \omega_{1,3}w_{3,t}$.

The model estimated by Kalman filter (see Subsection,4.2; the state-space model (Def. 4.1) includes survey-based variables (see Subsection 4.6).

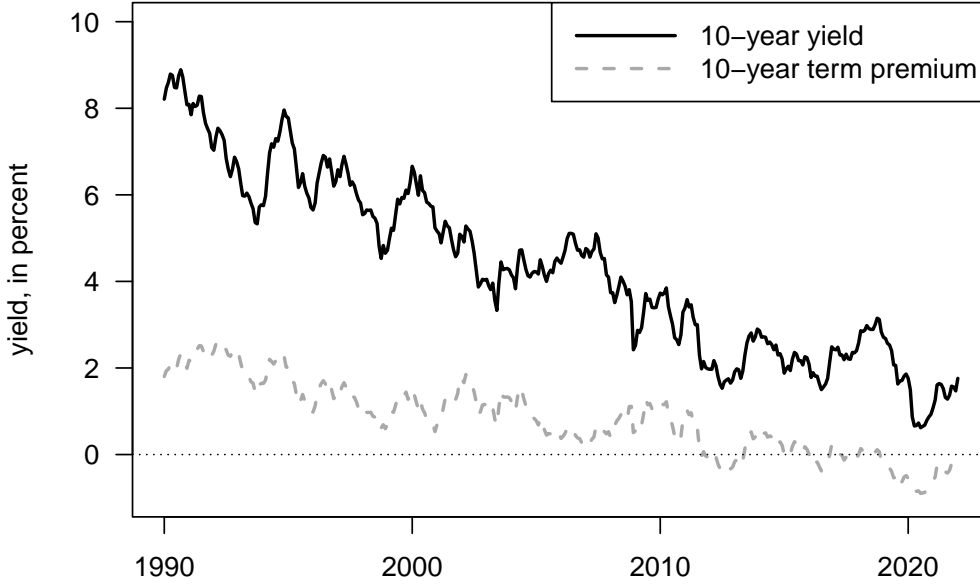


Figure 3.1: Kim and Wright (2005) outputs.

Outputs are regularly updated by the Federal Reserve Board.

Monthly data on the 6-month and 12-month-ahead forecasts of the three-month T-Bill yield from Blue Chip Financial Forecasts and semiannual data on the average expected three-month T-Bill yield from 6 to 11 years.

Example 3.2 (Ang and Piazzesi (2003)). Ang and Piazzesi (2003) propose one of the first paper mixing latent and macrovariables. The set up is also of the form Eq. (3.14), except that the VAR features several lags.¹ In their model, $w_t = [f_{1,t}^o, f_{2,t}^o, f_{1,t}^u, f_{2,t}^u, f_{3,t}^u]'$ where:

- $f_{1,t}^o$ is the first Principal Component of a set of 3 price indexes (growth rates)
- $f_{2,t}^o$ is the first Principal Component of a set of 4 real activity proxies (HELP, EMPLOY, IP, UE).
- $f_{i,t}^u$ are unobserved, or latent, factors.

The nominal short-term rate follows a Taylor rule. And latent factors are estimated via *inversion techniques* (Subsection 4.5).

¹Note that a VAR with p lags (i.e., a VAR(p)) admits a VAR(1) companion form.

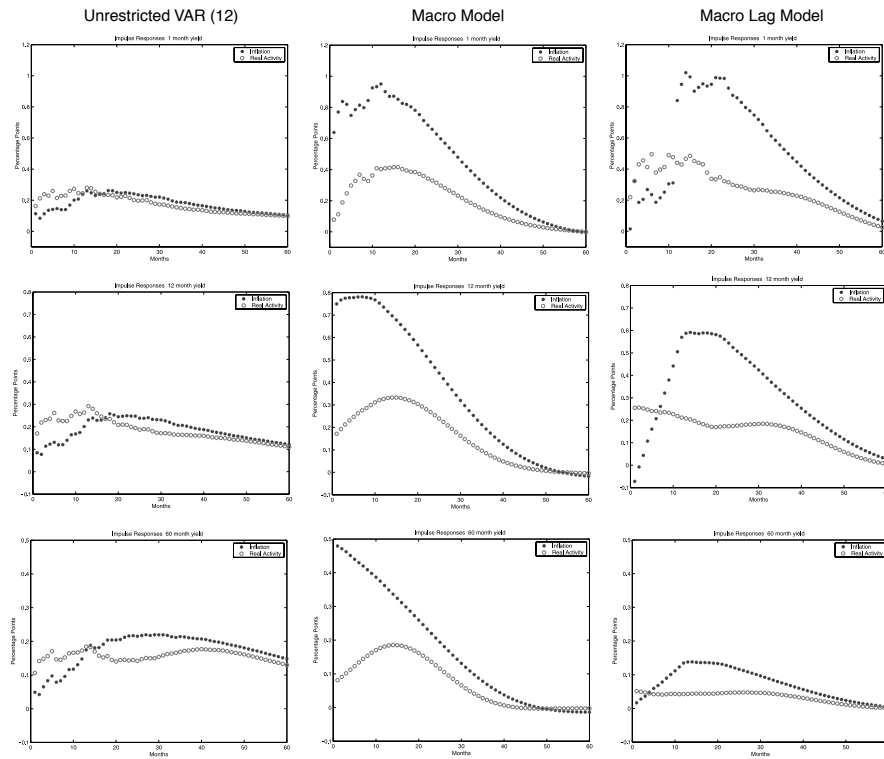


Figure 3.2: Source: Ang and Piazzesi (1998). Impulse response functions.

Example 3.3 (Joslin, Priebisch and Singleton (2014)). Joslin et al. (2014) first note that affine models stating that the short term rate is affine in macro factors imply that macro-factors are *spanned* by the yield curve: macro-factors should be perfectly explained by yields of different maturities. Further, they show that this is not the case in the data. (That is, regressing macro factors on yields provides R^2 that are far from one.)

They propose a model where macro factors are unspanned by the yield curve, but can still help predict yields. In their model, $w_t = [\mathcal{P}'_t, M'_t]'$, where \mathcal{P}_t are yield factors (\approx principal components) and M_t are macro factors. The model is as follows:

$$\begin{aligned} r_t &= \omega_0 + \omega'_\mathcal{P} \mathcal{P}_t \\ \begin{bmatrix} \mathcal{P}_t \\ M_t \end{bmatrix} &= \begin{bmatrix} \Phi_{\mathcal{P}\mathcal{P}} & \Phi_{\mathcal{P}M} \\ \Phi_{M\mathcal{P}} & \Phi_{MM} \end{bmatrix} \begin{bmatrix} \mathcal{P}_{t-1} \\ M_{t-1} \end{bmatrix} + \Sigma \varepsilon_t \\ \begin{bmatrix} \mathcal{P}_t \\ M_t \end{bmatrix} &= \mu + \begin{bmatrix} \Phi_{\mathcal{P}\mathcal{P}}^\mathbb{Q} & \textcolor{red}{0} \\ \Phi_{M\mathcal{P}}^\mathbb{Q} & \Phi_{MM}^\mathbb{Q} \end{bmatrix} \begin{bmatrix} \mathcal{P}_{t-1} \\ M_{t-1} \end{bmatrix} + \Sigma \varepsilon_t^\mathbb{Q}, \end{aligned}$$

where ε_t and $\varepsilon_t^\mathbb{Q}$ are $\mathcal{N}(0, Id)$ under \mathbb{P} and \mathbb{Q} , respectively.

M_t does Granger-cause \mathcal{P}_t under \mathbb{Q} and r_t is affine in \mathcal{P}_t (only).

In this context, yields $R(t, h)$ are affine in \mathcal{P}_t (only). However M_t does Granger-cause \mathcal{P}_t under \mathbb{P} , that is, macro-shocks affect the yield curve.

Example 3.4 (Ang, Boivin, Dong and Loo-Kung (2011)). Ang et al. (2011) propose a macro-finance model based on a quadratic framework. The short-term rate follows a Taylor rule with time-varying parameters:

$$r_t = \omega_0 + a_t g_t + b_t \pi_t,$$

where $x_t = (g_t, \pi_t, a_t, b_t)'$ follows a Gaussian VAR. This is the context described in Example 1.8. The previous equation shows that r_t is linear in $w_t = (x_t, \text{vec}(x_t x_t'))'$. Specifically:

$$r_t = \omega_0 + \omega'_1 w_t,$$

with $\omega_1 = [v, \text{vec}(V)]'$, where

$$v = \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \end{bmatrix} \quad \text{and} \quad V = \begin{bmatrix} 0 & 0 & 1/2 & 0 \\ 0 & 0 & 0 & 1/2 \\ 1/2 & 0 & 0 & 0 \\ 0 & 1/2 & 0 & 0 \end{bmatrix}.$$

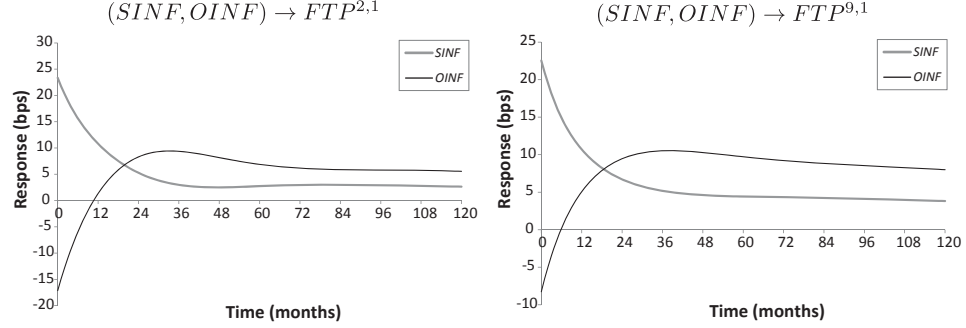


Figure 6. Response to inflation shock. Each panel plots the impulse responses of forward term premiums (FTPs) to shocks to spanned and unspanned inflation ($SINF$, $OINF$), implied by our preferred model with unspanned macro risks (\mathcal{M}_{us}). $SINF$ is the projection of INF onto the PCs; $OINF$ is the component of INF orthogonal to the PCs. $FTP^{p,1}$ is defined as the difference between the forward rate that one could lock in today for a one-year loan commencing in p years and the expectation for p years in the future of the one-year yield.

Figure 3.3: Source: Joslin, Priebsch, and Singleton (2014). Impulse response functions.

3.5 Non-Negative Affine Term Structure Model

In the presence of physical currency, absence of arbitrage opportunity and of storing cost of cash, nominal interest rates should be nonnegative. Many standard models (e.g. Gaussian ATSM) are non consistent with non-negative nominal yields. The period of extremely low interest rates challenged these models. Against this backdrop, approaches have been developed to accommodate zero (or effective) lower bounds. We provide two examples; only the second is an affine model.

3.5.1 The shadow-rate approach

The shadow-rate model is originally due to Black (1995). In this model, the short term rate is given by:

$$r_t = \max(s_t, \underline{r}), \quad (3.15)$$

where s_t is the shadow short-term interest rate and \underline{r} is the effective lower bound (≤ 0). While s_t can be real-valued, the short term rate is nonnegative

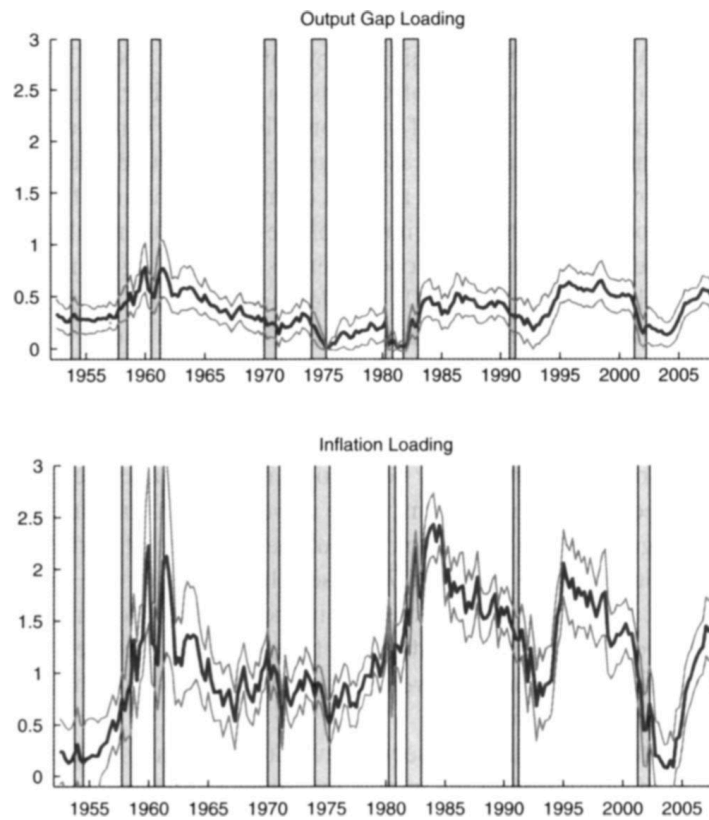


Figure 3.4: Source: Ang, Boivin, Dong, Loo-Kung (2011). Estimated factor loadings (a_t and b_t).

under (3.15). In shadow-rate models, the shadow rate s_t is usually a linear combination of a vector w_t that follows a Gaussian auto-regressive model. While s_t is a linear combination of components of an affine process, this is not the case for r_t . As a result, pricing formula are not available in closed-form. Approximation formula have been proposed by, e.g., Krippner (2013), Pribsch (2013), Wu and Xia (2016).

Let us describe the latter approach (Wu and Xia, 2016). As in Subsection 3.4, the S.D.F. is defined as:

$$\mathcal{M}_{t,t+1} = \exp(-r_t + \alpha'_t w_{t+1} - \psi_t(\alpha_t)), \text{ where } \alpha_t = \alpha_0 + \alpha'_1 w_t,$$

(this is Eq. (2.5)), but the short-term rate r_t is given by $r_t = \max(s_t, 0)$, with

$$s_t = \delta_0 + \delta'_1 w_t.$$

The approximation approach proposed by Wu and Xia (2016) is based on an approximation to the conditional expectations of forward rates. Using the results of Subsection 3.2.2, we have (Eq. (3.7)):

$$f_{n-1,n,t} = nR_{t,n} - (n-1)R_{t,n-1},$$

for $n > 0$ (and using $R_{t,0} = 0$). This implies that:

$$R_{t,h} = \frac{1}{h}(f_{t,0,1} + f_{t,1,2} + \dots + f_{t,h-1,h}).$$

The approximation of Wu and Xia (2016) consists in finding approximations of the forward rates $f_{t,n-1,n}$ (denoted by $\tilde{f}_{t,n-1,n}$, say) and to use them in the previous equation to get:

$$\boxed{R_{t,h} \approx \frac{1}{h}(\tilde{f}_{t,0,1} + \tilde{f}_{t,1,2} + \dots + \tilde{f}_{t,h-1,h})}. \quad (3.16)$$

Using that, for any random variable Z , we have $\log(\mathbb{E}[e^Z]) \approx \mathbb{E}[Z] + \frac{1}{2}\text{Var}[Z]$ (based on a second order Taylor expansion), Wu and Xia (2016) further show that:

$$f_{t,n,n+1} = -\log\left(\mathbb{E}_t^{\mathbb{Q}}\left(e^{-\sum_{j=0}^n r_{t+j}}\right)\right) + \log\left(\mathbb{E}_t^{\mathbb{Q}}\left(e^{-\sum_{j=0}^{n-1} r_{t+j}}\right)\right) \quad (3.17)$$

$$\approx \mathbb{E}_t^{\mathbb{Q}}[r_{t+n}] - \frac{1}{2}\left(\text{Var}_t^{\mathbb{Q}}\left(\sum_{j=0}^n r_{t+j}\right) - \text{Var}_t^{\mathbb{Q}}\left(\sum_{j=0}^{n-1} r_{t+j}\right)\right) \quad (3.18)$$

The expectation can be computed analytically:

$$\mathbb{E}_t^{\mathbb{Q}}[r_{t+n}] = \underline{r} + \sigma_n^{\mathbb{Q}} g\left(\frac{\bar{a}_n + b'_n X_t - \underline{r}}{\sigma_n^{\mathbb{Q}}}\right),$$

where $g(x) = x\Phi(x) - \phi(x)$, Φ and ϕ being the c.d.f. and p.d.f. of the standard normal distribution, respectively, and where

$$\begin{aligned}\bar{a}_n &= \delta_0 + \delta'_1 \left(\sum_{j=0}^{n-1} [\Phi^{\mathbb{Q}}]^j \right) \mu^{\mathbb{Q}} \\ b'_n &= \delta'_1 (\Phi^{\mathbb{Q}})^n.\end{aligned}$$

They also show that

$$\frac{1}{2} \left(\text{Var}_t^{\mathbb{Q}} \left(\sum_{j=0}^n r_{t+j} \right) - \text{Var}_t^{\mathbb{Q}} \left(\sum_{j=0}^{n-1} r_{t+j} \right) \right) \approx \Phi \left(\frac{\bar{a}_n + b'_n X_t - \underline{r}}{\sigma_n^{\mathbb{Q}}} \right) \times (\bar{a}_n - a_n),$$

where

$$a_n = \bar{a}_n - \frac{1}{2} \sigma_n^{\mathbb{Q}},$$

with

$$\sigma_n^{\mathbb{Q}} := \text{Var}_t^{\mathbb{Q}}(s_{t+n}) = \delta'_1 \left(\sum_{j=0}^{n-1} [\Phi^{\mathbb{Q}}]^j \right) \Sigma \Sigma' \left(\sum_{j=0}^{n-1} [\Phi^{\mathbb{Q}}]^j \right)' \delta_1.$$

They finally obtain:

$$\boxed{f_{t,n,n+1} \approx \tilde{f}_{t,n,n+1} = \underline{r} + \sigma_n^{\mathbb{Q}} g\left(\frac{a_n + b'_n X_t - \underline{r}}{\sigma_n^{\mathbb{Q}}}\right),}$$

which is used in (3.16) to obtain an approximation to $R_{t,h}$.

```
library(TSModels)
# Specify model:
n <- 2 # number of factors
rho <- matrix(0,n,n)
diag(rho) <- .97
mu <- matrix(0,n,1)
```

```

Sigma <- diag(n)
delta.0 <- 0; delta.1 <- rep(.01,n)
r.bar <- 0 # r = max(s,r.bar) [i.e., r.bar=0 in standard model]
Model <- list(rho = rho,mu = mu,Sigma = Sigma,
  delta.0 = delta.0,delta.1 = delta.1,r.bar = r.bar)
# Simulate model and compute shadow rate:
X <- simul.var(Model,nb.sim = 200) # simulated path
s <- delta.0 + X %*% delta.1
# Compute yields:
res <- compute.price.WX(Model,X,max.H=100)
# Prepare plots:
par(plt=c(.1,.95,.2,.75))
par(mfrow=c(2,1))
plot(s,type="l",xlab="time",ylab="",lwd=2,main="(a) Shadow rate")
t <- 50 #t <- which(s==min(s))
abline(v=t,col="dark grey",lwd=2,lty=3)
plot(res$vec.f[t,],type="l",xlab="maturity",ylab="",
  lwd=2,main="(b) yields and forward rates")
lines(res$vec.y[t,],col="red",lwd=2)
legend("topright",
  c("forward rates","yields to maturity"),lwd=c(2),lty=1,
  col=c("black","red"),bg = "white")

```

3.5.2 The auto-regressive gamma approach

Monfort et al. (2017) introduce an affine framework where the short-term rate can stay at zero for a prolonged period of time and with a stochastic lift-off probability.

Under \mathbb{P} and \mathbb{Q} , the state vector w_t follows a multi-variate auto-regressive gamma (VARG) process—a multivariate extension of Example 1.9. Conditionally on \underline{w}_t , the n components of w_{t+1} are independent and distributed as follows:

$$\frac{w_{i,t+1}}{\mu_i} \sim \gamma(\nu_i + z_{i,t}) \quad \text{where} \quad z_{i,t} \sim \mathcal{P}(\alpha_i + \beta_i' w_t). \quad (3.19)$$

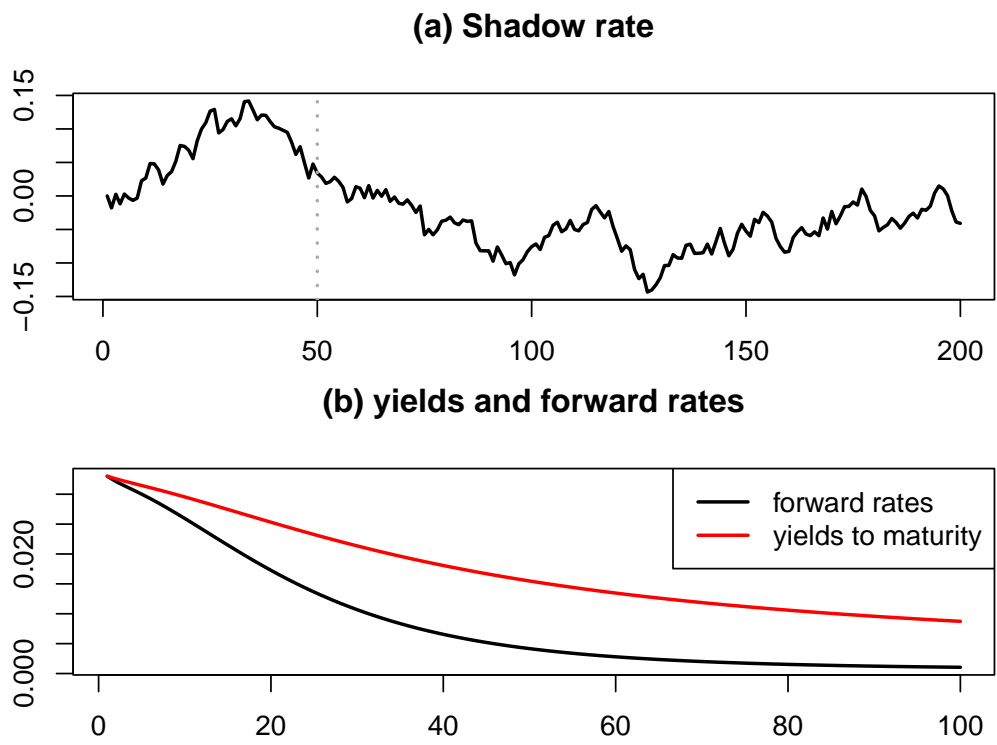


Figure 3.5: bla bla bla.

If $\mu = (\mu_1, \dots, \mu_n)'$, $\alpha = (\alpha_1, \dots, \alpha_n)'$, $\nu = (\nu_1, \dots, \nu_n)'$ and $\beta = (\beta_1, \dots, \beta_n)$, then

$$\begin{aligned} \varphi_t(u) = & \exp \left[\left(\frac{u \odot \mu}{1 - u \odot \mu} \right)' \beta' w_t \right. \\ & \left. + \alpha' \left(\frac{u \odot \mu}{1 - u \odot \mu} \right) - \nu' \log(1 - u \odot \mu) \right], \end{aligned}$$

where \odot denotes the element-by-element multiplication and, where, with abuse of notation, the division and log operators work element-by-element when applied to vectors.

In their baseline model, Monfort et al. (2017) use four factors. They set $\nu_1 = \nu_2 = 0$, implying that $w_{1,t}$ and $w_{2,t}$ can stay at zero (see Example 1.9). The short-term rate r_t is posited to be an affine combination of $w_{1,t}$ and $w_{2,t}$, that is:

$$r_t = \omega' w_t = \omega_1 w_{1,t} + \omega_2 w_{2,t},$$

hence, it can stay at zero.

Factors $w_{3,t}$ and $w_{4,t}$ Granger-cause $w_{1,t}$ and $w_{2,t}$, thereby causing r_t . As a result, for $h \geq 2$, $R(t, h)$ is a non-zero combination of the four components of w_t .

For the same reason, when $r_t = 0$, the lift-off probability depends on $w_{3,t}$ and $w_{4,t}$. The framework offers closed-form solutions for lift-off probabilities. Indeed, using Lemma ??:

$$\mathbb{P}_t(\alpha' w_{t+h} = 0) = \lim_{u \rightarrow -\infty} \varphi_{t,h}(0, \dots, 0, u\alpha),$$

where $\varphi_{t,h}$ is the multi-horizon Laplace transform defined in Eq. (1.7), which can be computed using Proposition 1.5. We have:

$$\begin{cases} \mathbb{P}_t(r_{t+h} > 0) = 1 - \lim_{u \rightarrow -\infty} \varphi_{t,h}(0, \dots, 0, u\omega) \\ \mathbb{P}_t(r_{t+1} = 0, \dots, r_{t+h} = 0) = \lim_{u \rightarrow -\infty} \varphi_{t,h}(u\omega, \dots, u\omega, u\omega) \equiv p_h \\ \mathbb{P}_t(r_{t+1} = 0, \dots, r_{t+h-1} = 0, r_{t+h} > 0) = p_{h-1} - p_h. \end{cases} \quad (3.20)$$

Other lift-off probabilities, of the type $\mathbb{P}_t[R(t+h, k) > \text{threshold}]$, can be derived from Eq. (1.18).

Monfort et al. (2017) estimate this model by means of Kalman filtering techniques (see Subsection ??). Observed variables include (levels of) yields, as well as survey-based forecasts of yields (see Subsection 4.6 and (e-GARCH-based) proxies of conditional variances (see Eq. (1.16)).

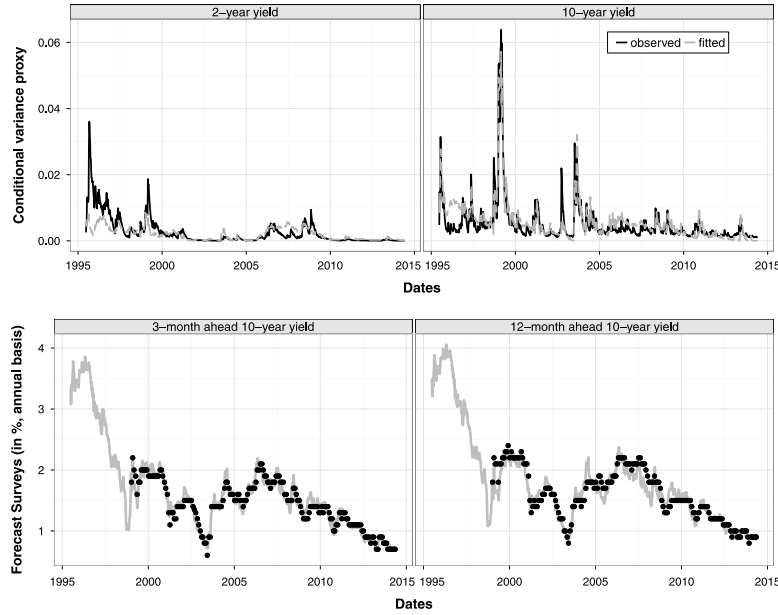


Fig. 7. Fitted conditional variance proxies and surveys. *Notes:* The top panel presents the two conditional variance proxies $V_t(h)$ estimated with an EGARCH(1,1) model on 2- and 10-year yields (left and right tiles) of weekly data from June 30, 1995 to May 30, 2014. The black solid lines are the observed variance proxies and the grey dotted lines are the model-implied (or fitted) equivalent. The bottom panel presents the survey of professional forecasters for the 10-year yield, 3- and 12-months ahead. Survey-based data are available at the monthly frequency from 1999 to the end of the sample. The black dots correspond to the observed data, and the grey solid lines are the fitted equivalent.

Figure 3.6: Source: Monfort et al. (2017). Model fit of conditional variances and surveys of professional forecasters.

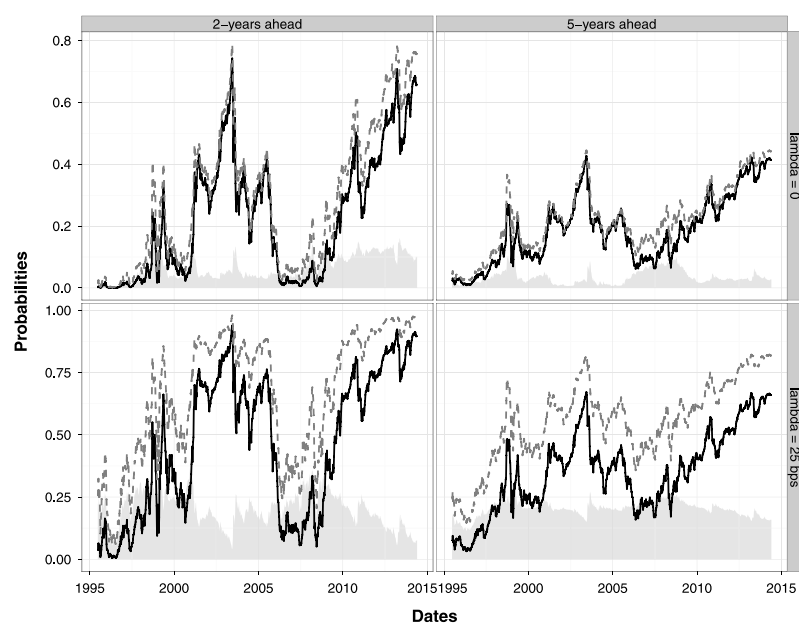


Fig. 9. Time-series of ZLB probabilities: $\mathbb{P}_t(r_{t+k} \leq \lambda)$ and $Q_t(r_{t+k} \leq \lambda)$. Notes: Probabilities are computed with weekly data from June 16, 1995 to May 30, 2014. The top panels present the probabilities of the short-rate hitting zero in two years (top-left panel) and 5 years (top-right panel). On bottom panels, we represent the probabilities of the short-rate being below 25 bps in 2 years (bottom-left panel) and 5 years (bottom-right panel). Black solid lines are the risk-neutral probabilities whereas grey dashed lines are the historical ones; grey-shaded areas are the difference between the two probabilities.

Figure 3.7: Source: Monfort et al. (2017). Lift-off probabilities.

Chapter 4

Estimation of affine asset-pricing models

4.1 State-space model

By nature, dynamic asset-pricing models are *state-space models*: The dynamics of all variables, gathered in vector y_t (yields, equity returns, macroeconomic variables, survey-based variables) are accounted for by state variables (w_t). The equations defining the relationship between the other variables and the state variables are called *measurement equations* (Eq. (4.1)). The equations defining the dynamics of the state variables are called *transition equations* (Eq. (4.2)).

In the case where w_t is an affine process (see Definition 1.1), the transition equations admit a VAR(1) representation (Proposition 1.1). In that case, the state-space model is said to be linear, as formalized by the following definition. This definition introduces, in particular, the notion of *measurement errors*.

Definition 4.1 (Linear State-Space Model). A linear state-space model writes as follows:

$$\begin{matrix} y_t \\ (m \times 1) \end{matrix} = A + Bw_t + \eta_t \quad \text{with } \eta_t \sim i.i.d. \mathcal{N}(0, \Omega) \quad (4.1)$$

$$\begin{matrix} w_t \\ (n \times 1) \end{matrix} = \mu + \Phi w_{t-1} + \Sigma^{\frac{1}{2}}(w_{t-1})\varepsilon_t, \quad (4.2)$$

where ε_t is a martingale difference sequence satisfying $\text{Var}_t(\varepsilon_{t+1}) = Id$. The components of η_t are measurement errors.

Note: Eq. (4.2) derives from Proposition 1.1 (Eq. (1.12)).

In practice, one can distinguish two situations: (a) all state variables (components of w_t) are observed and (b) some of these variables are latent. What do we mean by *model estimation* in case (b)? There are different situations:

- (b.i) We know the model parameters but we want to recover the latent factors—for instance to compute model-implied prices.
- (b.ii) We know neither the model parameters nor the latent variables, we want to estimate both of them.
- (b.iii) We know neither the model parameters nor the latent variables, we are just interested in the model parameters.

In case (a): One can resort to standard estimation techniques (GMM, Maximum Likelihood) to estimate model parameters. Take for instance the maximum-likelihood case, we have:

$$\begin{aligned} f(y_t, w_t | \underline{y_{t-1}}, \underline{w_{t-1}}) &= f(y_t | \underline{y_{t-1}}, \underline{w_t}) \times \underbrace{f(w_t | \underline{y_{t-1}}, \underline{w_{t-1}})}_{=f(w_t | \underline{w_{t-1}})} \\ &= \mathbf{n}(y_t; A + Bw_t, \Omega) f(w_t | \underline{w_{t-1}}), \end{aligned}$$

where $\mathbf{n}(x; \mu, \Omega)$ denotes the evaluation, at vector x , of the p.d.f. of the multivariate normal distribution $\mathcal{N}(\mu, \Omega)$. That is, if x is a m -dimensional vector:

$$\mathbf{n}(x; \mu, \Omega) = \frac{1}{\sqrt{(2\pi)^m |\Omega|}} \exp \left(-\frac{1}{2} \{x - \mu\}' \Omega^{-1} \{x - \mu\} \right). \quad (4.3)$$

Once this conditional p.d.f. is known, the total likelihood is given by (conditional on y_0 and w_0):

$$\prod_{t=1}^T f(y_t, w_t | \underline{y_{t-1}}, \underline{w_{t-1}}).$$

Of course, $f(w_t|\underline{w_{t-1}})$ depends on the process chosen for w_t . If it is complicated to compute, one can employ Pseudo Maximum Likelihood (Gourieroux et al., 1984). The p.d.f. may involve, for instance, an infinite sum, which is the case in the ARG case of Example 1.9. When w_t is affine, the Pseudo Maximum Likelihood approach consists in replacing $f(w_t|\underline{w_{t-1}})$ by:

$$\mathbf{n}(w_t; \mu + \Phi w_{t-1}, \Sigma(w_t)),$$

where μ , Φ and $\Sigma(w_t)$ are introduced in Eqs. (1.13) and (1.14) in Proposition 1.1.

In case (b.iii), one can estimate the model by Generalized Method of Moments (GMM), fitting sample moments computed using observed variables (prices, yields, returns). In the context of affine processes, conditional and unconditional moments of the state vector w_t are available in closed form, as shown by Eqs. (1.15), (1.16) and (1.17). If the model-implied moments are not available in closed-form, one may have to resort to the *Simulated Method of Moments (SMM)* (see, e.g., Gouriéroux and Monfort (1997) or Duffie and Singleton (1993)).

In cases (b.i) and (b.ii), one has to implement *filtering methods*, on which we focus on in the following.

4.2 Kalman-filter-based approach

4.2.1 The Gaussian linear state-space case

Let us start with a particular case of state-space model (Def. 4.1) where ε_t is Gaussian and where $\Sigma^{\frac{1}{2}}$ does not depend on w_t , i.e. with a homoskedastic linear Gaussian state-space model.

Let's denote by θ the vector of parameters that defines the model. For a given θ and a sequence of observations $\{y_1, \dots, y_T\}$, the Kalman filter computes the distribution of w_t given $\{y_1, \dots, y_t\}$ (see Def. 4.2). This distribution is Gaussian, and obtained by a recursive algorithm. A byproduct of this algorithm is the likelihood function associated with θ and $\{y_1, \dots, y_T\}$. This opens the door to the estimation of θ by MLE, maximizing this function. In this sense, Kalman-filter techniques can address Objective (b.ii).

Let us first introduce the notion of *filtered* and *smoothed* estimates of a latent variable (or vector of variables) w_t :

Definition 4.2 (Filtered versus smoothed estimates). The filtering and smoothing problems consist in computing the following conditional moments:

$$\begin{aligned} \text{Filtering: } w_{t|t} &= \mathbb{E}(w_t | \underline{y}_t) \quad \text{and} \quad P_{t|t} = \mathbb{V}ar(w_t | \underline{y}_t) \\ \text{Smoothing: } w_{t|T} &= \mathbb{E}(w_t | \underline{y}_T) \quad \text{and} \quad P_{t|T} = \mathbb{V}ar(w_t | \underline{y}_T). \end{aligned}$$

The following proposition outlines the Kalman algorithm (see, e.g. Nelson and Kim (1999)).

Proposition 4.1 (Kalman filter and smoother). *If $\varepsilon_t \sim \mathcal{N}(0, I)$ in the state-space defined in Def. 4.1, then we have (filtering):*

$$w_t | y_1, \dots, y_t \sim \mathcal{N}(w_{t|t} | P_{t|t}),$$

where $w_{t|t}$ and $P_{t|t}$ result from the following recursive equations:

| |
|--|
| $\begin{aligned} w_{t t} &= w_{t t-1} + K_t \lambda_t \\ P_{t t} &= (I - K_t B) P_{t t-1} \end{aligned}$ <p style="text-align: center; margin: 10px 0;"><i>where (updating step)</i></p> $\begin{aligned} \lambda_t &= y_t - A - B w_{t t-1} \quad (\text{forecast error}) \\ S_{t t-1} &= \mathbb{V}ar(y_t y_{t-1}) = B P_{t t-1} B' + \Omega \\ K_t &= P_{t t-1} B' S_{t t-1}^{-1} \quad (\text{Kalman gain}) \end{aligned}$ <p style="text-align: center; margin: 10px 0;"><i>and where (forecasting step)</i></p> $\begin{aligned} w_{t t-1} &= \mu + \Phi w_{t-1 t-1} \\ P_{t t-1} &= \Sigma + \Phi P_{t-1 t-1} \Phi' \quad (\Sigma = \Sigma^{\frac{1}{2}} \Sigma^{\frac{1}{2}}'). \end{aligned}$ |
|--|

The log likelihood is (recursively) computed as follows:

$$\begin{aligned} \log \mathcal{L}(\theta; \underline{y}_T) &= \frac{mT}{2} \log(2\pi) \\ &\quad - \frac{1}{2} \sum_{t=1}^T \left(\log |S_{t|t-1}(\theta)| + \lambda_t'(\theta) S_{t|t-1}^{-1}(\theta) \lambda_t(\theta) \right). \end{aligned} \tag{4.4}$$

Moreover, we have (smoothing):

$$w_t|y_1, \dots, y_T \sim \mathcal{N}(w_{t|T}|P_{t|T}),$$

where $w_{t|T}$ and $P_{t|T}$ result from the following recursive equations:

$$\begin{aligned} w_{t|T} &= w_{t|t} + F_t(w_{t+1|T} - w_{t+1|t}) \\ P_{t|T} &= P_{t|t} + F_t(P_{t+1|T} - P_{t+1|t})F_t' \\ \text{where} \\ F_t &= P_{t|t}\Phi_{t+1}'P_{t+1|t}^{-1}. \end{aligned}$$

The following figure illustrates the updating step of the Kalman algorithm:

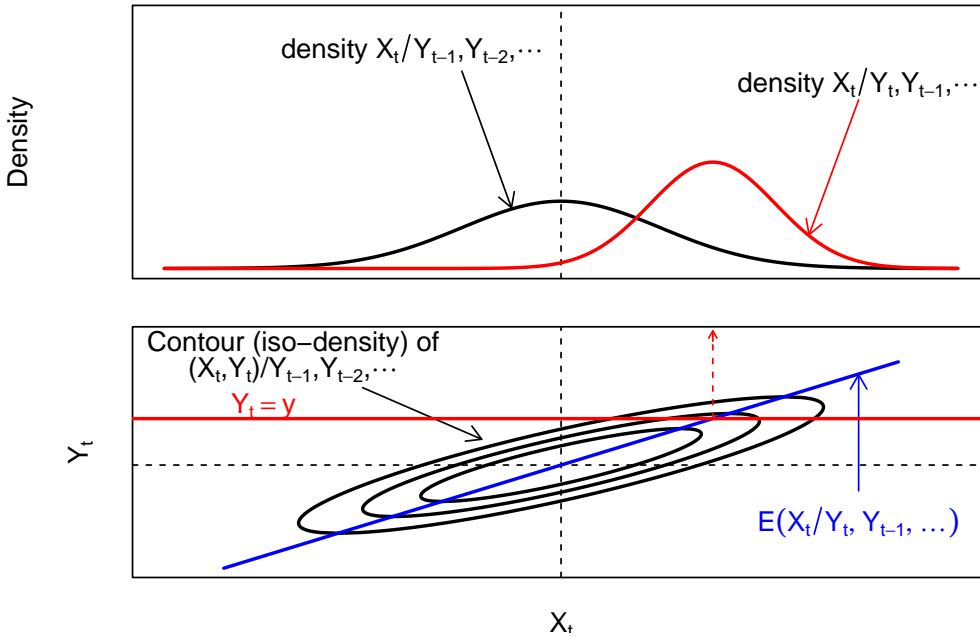


Figure 4.1: Updating in the Kalman filter.

The recursive equations of the Kalman filter need to be initialized. That is, one needs to provide initial values for $w_{0|0}$, $P_{0|0}$. Different possibilities have been proposed. One can for instance:

- Include the elements of $(w_{0|0}, P_{0|0})$ among the parameters to estimate;

- Set $w_{0|0}$ and $P_{0|0}$ to their unconditional values (using, e.g., Eq. (1.17));
- Set $w_{0|0}$ to a prior value and take either an arbitrary large value for $P_{0|0}$ if the prior value is uncertain (which depicts a situation of *diffuse prior*) or a small value for $P_{0|0}$ if we are confident in this prior value $w_{0|0}$.

Example 4.1 (Kalman filtering and smoothing). To illustrate, consider the following model:

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \begin{bmatrix} \alpha_1 & 0 \\ 0 & \alpha_2 \end{bmatrix} \begin{bmatrix} y_{1,t-1} \\ y_{2,t-2} \end{bmatrix} + \begin{bmatrix} \gamma_1 \\ \gamma_2 \end{bmatrix} w_t + D\eta_t \quad (4.5)$$

$$w_t = \phi w_{t-1} + \varepsilon_t, \quad (4.6)$$

where $\eta_t \sim \mathcal{N}(0, Id)$.

In the following lines, we specify one version of the previous model. We simulate trajectories of $y_{1,t}$, $y_{2,t}$ and w_t over 100 periods (Figure 4.2) and we further call `Kalman_filter` and `Kalman_filter` (from the `TSMODELS` package) to compute the filtered and smoothed estimates of w_t (Figure 4.3).

```
library(TSMODELS) # Kalman filter procedure in there.
# Set model specifications:
alpha1 <- .5; alpha2 <- .95; Alpha <- diag(c(alpha1, alpha2))
d_11 <- 1; d_12 <- .5; d_21 <- .5; d_22 <- 2
D <- matrix(c(d_11, d_21, d_12, d_22), 2, 2)
gamma1 <- 1; gamma2 <- 2; Gamma <- matrix(c(gamma1, gamma2), 2, 1)
phi <- .8
# Simulate model:
T <- 100
Y <- NULL; X <- NULL
Alpha.Y_1 <- NULL
y <- c(0, 0); x <- 0
for(i in 1:T){
  Alpha.Y_1 <- rbind(Alpha.Y_1, c(Alpha %*% y))
  y <- Alpha %*% y + Gamma * x + D %*% rnorm(2)
  x <- phi * x + rnorm(1)
  Y <- rbind(Y, t(y)); X <- rbind(X, x)}
# Define matrices needed in the Kalman_filter procedures:
nu_t <- matrix(0, T, 1)
```

```

H <- phi; G <- Gamma
mu_t <- Alpha.Y_1
N <- 1; M <- D
Sigma_0 <- 1/(1-phi^2) # unconditional variance of w
rho_0 <- 0 # unconditional mean of w
filter.res <- Kalman_filter(Y,nu_t,H,N,mu_t,G,M,Sigma_0,rho_0,indic_pos=0)
smoother.res <- Kalman_smoother(Y,nu_t,H,N,mu_t,G,M,Sigma_0,rho_0,indic_pos=0)

```

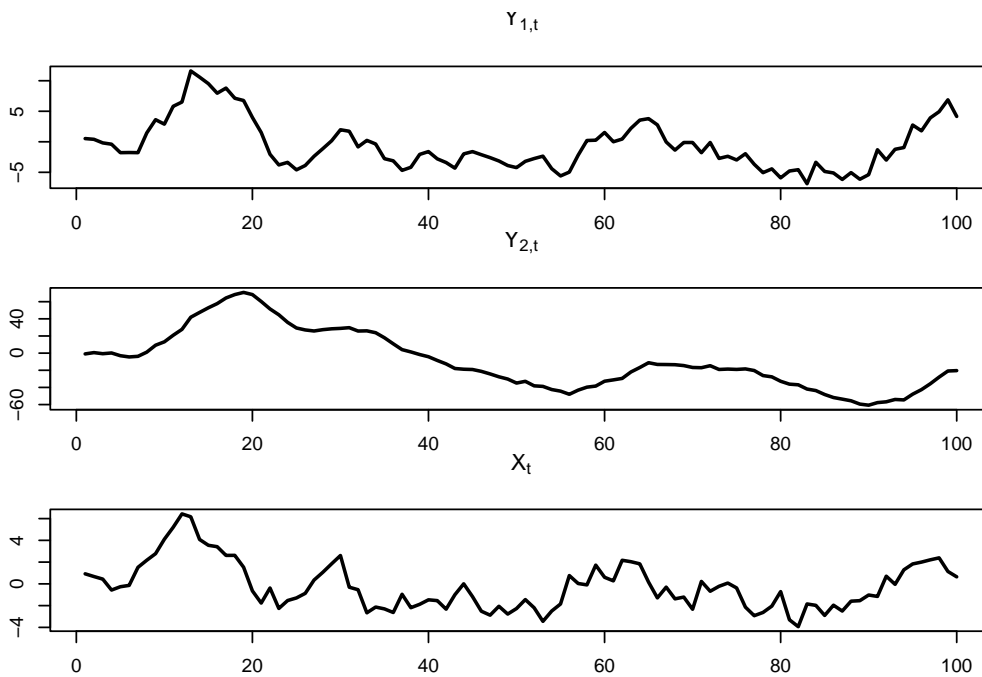
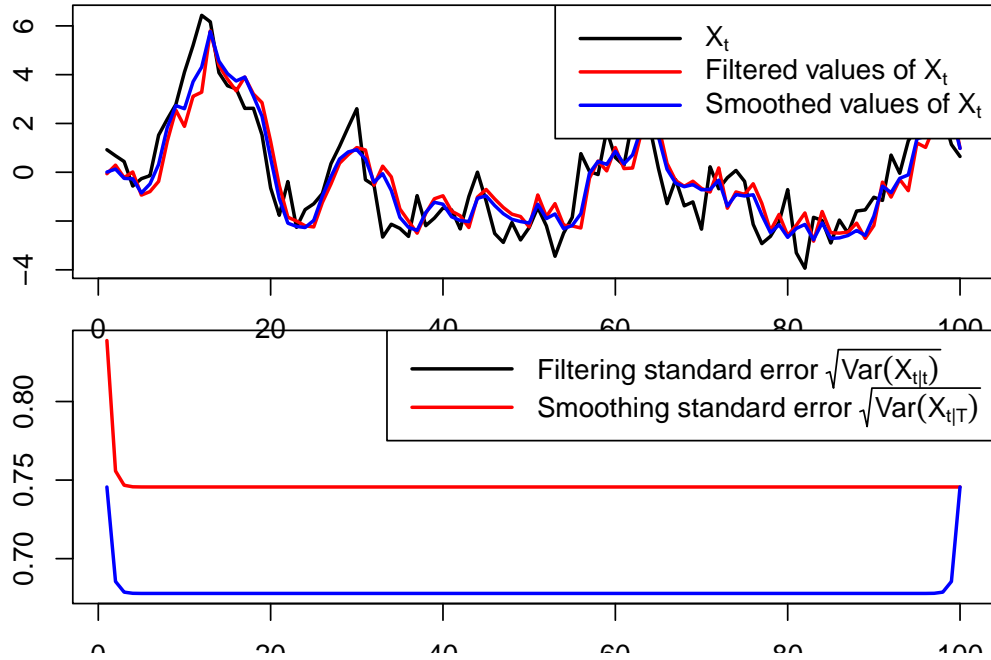


Figure 4.2: Simulated trajectories of $y_{1,t}$, $y_{2,t}$, and w_t .

4.2.2 Missing observations

In many application, one does not observe all the entries of y_t at every date. This arises for instance in situations where (i) measurement variables feature different frequencies, (ii) we have unreliable data for some period (that we prefer not to include among observations), (iii) some of the measurement variables are observed over a shorter time span.

Figure 4.3: Filtered and smoothed estimates of w_t .

These situations are easily addressed by Kalman filtering/smoothing (e.g., Chow and Lin (1971) or Harvey and Pierse (1984)). To accommodate missing observations in some of the y_t 's, one simply has to change the size of this vector (and of λ_t , $S_{t|t-1}$, A and B) for the relevant dates. Of course, the accuracy of $w_{t|t}$ will tend to be lower during periods where one or several (or all) the entries of y_t are unobserved. (This will be apparent in the resulting covariance matrix of $w_{t|t}$, namely $P_{t|t}$.) The log-likelihood computation (Eq. (4.4)) is still valid in this case; one simply has to adjust the number of observed variables at each iteration; that is, m then depends on time.

Example 4.2 (Kalman filtering and smoothing). This example extends Example 4.2. We take the simulated path of the observed variables $y_{1,t}$ and $y_{2,t}$ and remove observations of $y_{1,t}$ (respectively of $y_{2,t}$) between periods $t = 30$ and $t = 50$ (resp. between periods $t = 40$ and $t = 70$), and then use the Kalman filter and smother to recover the states w_t in this situation with missing observations.

```

Y.modif <- Y
Y.modif[30:50,1] <- NaN
Y.modif[40:70,2] <- NaN
# Call of Kalman filter and smoother:
filter.res <- Kalman_filter(Y.modif,nu_t,H,N,mu_t,G,M,Sigma_0,rho_0,indic_pos=0)
smoother.res <- Kalman_smoother(Y.modif,nu_t,H,N,mu_t,G,M,Sigma_0,rho_0,indic_pos=0)

```

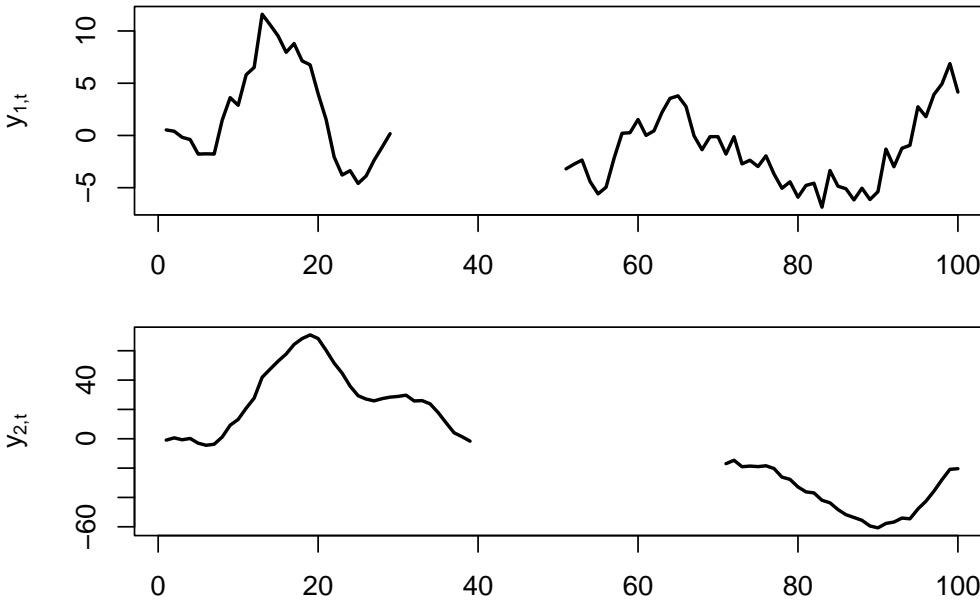


Figure 4.4: Situation of missing observations.

4.3 About non-constant conditional matrix Σ

Proposition 4.1 is valid when ε_t is Gaussian and when $\Sigma^{\frac{1}{2}}$ does not depend on w_t . However, in the general case (but when w_t is an affine process), we have that $\Sigma(w_{t-1}) \equiv \text{Var}(w_{t+1}|w_t)$ is affine in w_t (see Prop. 1.1). In order to deal with this, the Kalman filter algorithm can be modified. Specifically, in the prediction step (see Prop. 4.1), $P_{t|t-1}$ can be approximated by:

$$P_{t|t-1} = \Sigma(w_{t-1|t-1}) + \Phi P_{t-1|t-1} \Phi',$$

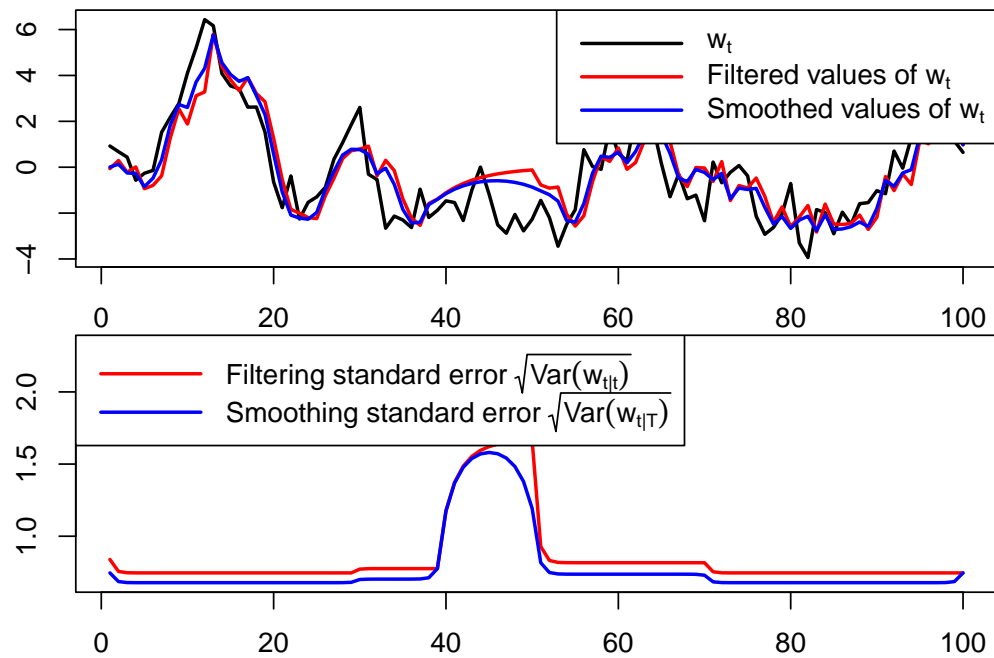


Figure 4.5: Filtered and smoothed estimates of w_t in a situation of missing observations. The lower plot shows the standard errors associated with filtered and smoothed estimates. As expected, uncertainty is larger for those dates where observations are missing.

i.e. we replace $\Sigma(w_{t-1})$ by $\Sigma(w_{t-1|t-1})$.

Though the approach is then not necessarily optimal, it shows good empirical properties (de Jong (2000) or Duan and Simonato (1999)). In order to test for the validity of the approach in a specific context, one can resort to Monte-Carlo simulations (Monfort et al., 2017).

4.4 Non-linear models

As soon as w_t follows an affine process, it admits the VAR dynamics presented in Prop. 1.1, i.e., it features a linear transition equation. However, measurement equations may be non-linear (affine) functions of w_t . This is in particular the case if observed variables include Swaps rates (see Remark SWAPS XXXX), CDS rates (see Subsection ??, in particular Eq. (6.41)) or prices of tranche products (see Example 6.5).

In a context of non-linear measurement equations, one can for instance resort to the Extended Kalman Filter (linearizing the measurement equations) or, to higher-order Taylor. Monfort et al. (2015) develop a Quadratic Kalman Filter (QKF), where measurement equations are quadratic functions of the state vector.

4.5 The inversion technique

The inversion technique has been introduced by Chen and Scott (1993). It is used, e.g., by Ang and Piazzesi (2003) and Liu et al. (2006). Contrary to Kalman-type approaches, this approach is not recursive. it can therefore be faster, especially for long sample.

This approach works under the assumption that some of the observed variables are *perfectly priced* (or modelled). (Recall that y_t and w_t are respectively of dimension m and n , see Eqs. (4.1) and (4.2) in Def. 4.1.) Formally:

Hypothesis 4.1 (Perfectly-modelled variables). n components of the m -dimensional vector y_t (with $n \leq m$) are perfectly modelled. That is, there is no measurement errors in associated measurement equations.

Without loss of generality, these perfectly-modelled variables are the first n components of y_t , that is:

$$y_t = \begin{pmatrix} \underbrace{y_{1,t}}_{(n \times 1)} \\ \underbrace{y_{2,t}}_{(m-n) \times 1} \end{pmatrix}.$$

Under Assumption 4.1, the measurement equation (Eq. (4.1)) becomes:

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \begin{bmatrix} A_1 \\ A_2 \end{bmatrix} + \begin{bmatrix} B_1 \\ B_2 \end{bmatrix} w_t + \begin{bmatrix} 0 \\ \eta_{2,t} \end{bmatrix},$$

where $\eta_{2,t} \sim \mathcal{N}(0, \Omega_2)$ (say). This notably implies that

$$w_t = B_1^{-1}(y_{1,t} - A_1). \quad (4.7)$$

Under this assumption, and if the conditional distribution of w_t is available in closed form, then Proposition 4.2 shows that the (exact) likelihood of the model can then be computed. This proposition shows in particular that the conditional p.d.f. $f_{Y_t|Y_{t-1}}(y_t; \underline{y_{t-1}})$ involves three terms:

- The first term (in blue in (4.8)) stems from the conditional distribution $w_t | \underline{w_{t-1}}$.
- The second term (in red in (4.8)) is associated with the measurement errors pertaining to $y_{2,t}$, that are the components of $\eta_{2,t}$.
- The third term (in brown in (4.8)) is the determinant of the Jacobian matrix associated with the linear transformation between w_t and $y_{1,t}$ (Eq. (4.7)), that is $|B_1|$.

Once one knows how to compute $f_{Y_t|Y_{t-1}}(y_t; \underline{y_{t-1}})$, the total likelihood is easily obtained since:

$$f_{Y_1, \dots, Y_T}(y_1, y_2, \dots, y_T) = f_{Y_1}(y_1) \prod_{t=2}^T f_{Y_t|Y_{t-1}}(y_t; \underline{y_{t-1}}).$$

Proposition 4.2 (Log-likelihood in the inversion context). *In the context of a linear state-space model as defined in Def. 4.1, under Assumption 4.1, and if w_t is a Markovian process, we have:*

$$f_{Y_t|Y_{t-1}}(y_t; \underline{y_{t-1}}) = f_{w_t|w_{t-1}}(w(y_{1,t}); w(y_{t-1})) \times \mathbf{n}(y_{2,t}; A_2 + B_2 w(y_{1,t}), \Omega_2) \times |B_1|^{-1}. \quad (4.8)$$

where $w(y_{1,t}) = B_1^{-1}(y_{1,t} - A_1)$ and where \mathbf{n} denotes the multivariate normal p.d.f. (Eq. (4.3)).

Proof. Since w_t is Markov, so is $y_{1,t}$ and since $y_{2,t} = A_2 + B_2 w(y_{1,t}) + \eta_{2,t}$, with $w(y_{1,t}) = B_1^{-1}(y_{1,t} - A_1)$, we have:

$$\begin{aligned} f(y_t|y_{t-1}) &= f_1(y_{1,t}|y_{1,t-1})f_2(y_{2,t}|y_{1,t}) \\ &= |B_1|^{-1}f_w(w(y_{1,t})|w(y_{1,t-1}))\mathbf{n}(y_{2,t}; A_2 + B_2 w(y_{1,t}), \Omega_2), \end{aligned}$$

where

- $f_1(y_{1,t}|y_{1,t-1}) = |B_1|^{-1}f_w(w(y_{1,t})|w(y_{1,t-1}))$ comes from the fact that, if U and V are two random variables such that $V = g(U)$, where g is a bijective and differentiable function, then $f_V(v) = \left| \frac{\partial g^{-1}(v)}{\partial v'} \right| f_U(g^{-1}(v))$,
- and $f_2(y_{2,t}|y_{1,t}) = \mathbf{n}(y_{2,t}; A_2 + B_2 w(y_{1,t}), \Omega_2)$ comes from the fact that $y_{2,t}|y_{1,t} \sim \mathcal{N}(A_2 + B_2 w(y_{1,t}), \Omega_2)$.

□

4.5.1 Dealing with unobserved regimes

Kalman and inversion techniques are not suited to the case where some of the components of w_t are valued in a discrete set. This is typically the case if w_t is of the form:

$$w_t = \begin{pmatrix} z_t \\ x_t \end{pmatrix},$$

where z_t is valued in $\{e_1, \dots, e_J\}$, e_j being the j^{th} column of Id_J , which is the case in the prsence of regime switching features (see Subsection 1.6).

Assume that z_t is an exogenous and homogenous Markov chain whose dynamics is defined by the $\pi(e_i, e_j)$'s that are such that:

$$\pi(e_i, e_j) = \mathbb{P}(z_{t+1} = e_j | z_t = e_i). \quad (4.9)$$

Denote by Π the matrix of transition probabilities, i.e., the (i, j) component of Π is $\pi(e_i, e_j)$. Assume further that we have:

$$x_t = m(z_t, x_{t-1}) + \varepsilon_t, \quad (4.10)$$

where $\mathbb{E}(\varepsilon_t | \underline{z}_t, \underline{x}_{t-1}) = 0$.

We denote the conditional distribution of ε_t w.r.t. (z_t, x_{t-1}) by $f_\varepsilon(\cdot; z_t, x_{t-1})$.

Hypothesis 4.2 (Measurement equations). The measurement equation is of the form:

$$y_t = Az_t + Bx_t + \eta_t, \quad \text{with } \eta_t \sim i.i.d. \mathcal{N}(0, \Omega). \quad (4.11)$$

Example 4.3 (Regime-Switching Gaussian VAR). Building on Example 1.15, we know that if

$$x_t = \mu z_t + \Phi x_{t-1} + \varepsilon_t, \quad (4.12)$$

where $\varepsilon_t | \underline{x}_t, z_t \sim \mathcal{N}(0, \Sigma(z_t))$ and if z_t is an exogenous independent Markov chain, then $w_t = (x'_t, z'_t)'$ is affine. Using the notations of Eq. (4.10), we have:

$$m(z_t, x_{t-1}) = \mu z_t + \Phi x_{t-1}.$$

If r_t and the s.d.f. are respectively affine and exponential in w_t , then, in particular, yields are also affine in w_t , i.e. of the form $R(t, h) = A'_h z_t + B'_h x_t$ (see Eq. (3.8)). Therefore, if the components of y_t are yields of different maturities, the measurement equations are consistent with Assumption 4.2.

How to estimate such a model when the regimes z_t are unobservable? We distinguish two distinct situations:

Case 1. The x_t factors are observable. The probabilities of being in the different regimes on each date can be estimated by employing the Kitagawa-Hamilton filter, with (using the notations of Proposition ??): $F_t = (y'_t, x'_t)'$ and

$$\begin{aligned} f(F_t | z_t = e_j, \underline{F}_{t-1}) &= f(y_t | x_t, z_t = e_j, \underline{F}_{t-1}) f(x_t | z_t = e_j, \underline{F}_{t-1}) \\ &= \mathbf{n}(y_t; Az_t + Bx_t, \Omega) \times \\ &\quad f_\varepsilon(x_t - m(z_t, x_{t-1}); z_t, x_{t-1}), \end{aligned}$$

where $\mathbf{n}(u; \mu, \Omega)$ denotes the evaluation, at vector u of the p.d.f. of the multivariate normal distribution $\mathcal{N}(\mu, \Omega)$ (see eq. (4.3)). A by-product of the Kitagawa-Hamilton filter is the likelihood function associated with the dataset. As a result, the model parameterization can be estimated by maximum likelihood approach.

Case 2. The x_t factors are not observable. There are two sub-cases:

- (2.i) the components of y_t are not perfectly modelled (i.e. $\Omega \neq 0$, where Ω defined in Eq. (4.11)). One has then to resort to filters dealing with two types of uncertainty: hidden discrete values (z_t) and continuously distributed latent variables (x_t). Kim (1994)'s filter can be employed when the state-space model is of the form (4.11)-(4.12) (see, e.g., Monfort and Renne (2014)), detailed in Example 6.2).
- (2.ii) n_x components of y_t are perfectly modelled (where n_x is the dimension of x_t). One can then resort to an inversion technique, complemented with the Kitagawa-Hamilton filter, to estimate the model (see Monfort and Renne (2013) and Renne (2017) for applications).

Proposition 4.3 (Kitagawa-Hamilton filter). *Consider a q -dimensional vector of variables F_t and an exogenous homogenous Markov chain z_t . We make use of the following notations:*

- η_t is a J -dimensional vector whose j^{th} component is the p.d.f. of F_t conditional on $(z_t = e_j, \underline{F}_{t-1})$, i.e. $f(F_t | z_t = e_j, \underline{F}_{t-1})$
- ξ_t is a J -dimensional vector whose j^{th} component is $\mathbb{P}(z_t = e_j | \underline{F}_t)$.

The sequence ξ_t can then be computed recursively as follows:

$$\xi_t = \frac{(\Pi' \xi_{t-1}) \odot \eta_t}{\mathbf{1}'(\Pi' \xi_{t-1} \odot \eta_t)}, \quad (4.13)$$

where \odot denotes the element-by-element (Hadamard) product and where $\mathbf{1}$ denotes a J -dimensional vector of ones.

Moreover, the previous formulas also show how to compute the likelihood of the model since:

$$f(F_t | \underline{F}_{t-1}) = \mathbf{1}'(\Pi' \xi_{t-1} \odot \eta_t). \quad (4.14)$$

4.5.2 Mixed use of Kitagawa-Hamilton and inversion techniques

Without loss of generality, assume that the n_x first components of y_t are observed without error, i.e. Assumption 4.2 becomes

$$\begin{bmatrix} y_{1,t} \\ y_{2,t} \end{bmatrix} = \begin{bmatrix} A_1 z_t \\ A_2 z_t \end{bmatrix} + \begin{bmatrix} B_1 \\ B_2 \end{bmatrix} x_t + \begin{bmatrix} 0 \\ \eta_{2,t} \end{bmatrix},$$

where $\varepsilon_2 \sim \mathcal{N}(0, \Omega_2)$.

Since $y_{1,t} = A_1 z_t + B_1 x_t$, we then have:

$$x_t \equiv x(y_t, z_t) = B_1^{-1}(y_{1,t} - A_1 z_t). \quad (4.15)$$

In order to employ the Kitagawa-Hamilton filter (Proposition 4.3), one need to define the extended Markov chain:

$$\mathcal{Z}_t = z_{t-1} \otimes z_t,$$

whose matrix of transition probabilities is detailed in Proposition 4.4.

Proposition 4.4 (Kitagawa-Hamilton and inversion techniques). *The matrix of transition probabilities of \mathcal{Z}_t is of the form $\mathbf{1}_{n \times 1} \otimes \tilde{\Pi}$, with*

$$\tilde{\Pi} = \begin{bmatrix} \pi_{1,\bullet} & 0_{1 \times n} & \dots & & 0_{1 \times n} \\ 0_{1 \times n} & \pi_{2,\bullet} & 0_{1 \times n} & \dots & 0_{1 \times n} \\ & & \ddots & & \\ & & 0_{1 \times n} & \pi_{n-1,\bullet} & 0_{1 \times n} \\ 0_{1 \times n} & \dots & & 0_{1 \times n} & \pi_{n,\bullet} \end{bmatrix},$$

where $\pi_{i,\bullet}$ denotes the i^{th} row of Π (Π is defined on Slide XXX).

The last term appearing in Eq. (4.16) can be computed as follows:

$$\begin{aligned} & f \left(\begin{bmatrix} x(y_t, z(\mathcal{Z}_t)) \\ y_{2,t} \end{bmatrix} \middle| \mathcal{Z}_t, \underline{y_{t-1}} \right) \\ &= f(y_{2,t} | x_t = x(y_t, z(\mathcal{Z}_t)), \mathcal{Z}_t, \underline{y_{t-1}}) \times \\ & \quad f(x(y_t, z(\mathcal{Z}_t)) | \mathcal{Z}_t, \underline{y_{t-1}}) \\ &= \mathbf{n}(y_{2,t}; A_2 z_t + B_2 x_t, \Omega_2) \times \\ & \quad f_\varepsilon(\varepsilon_t | z_t = z(\mathcal{Z}_t), x_{t-1} = x(y_{t-1}, z_{-1}(y_{t-1}, \mathcal{Z}_t))), \end{aligned}$$

where \mathbf{n} is the p.d.f. of a multivariate normal distri, (Prop. 4.2, where $\varepsilon_t = x_t - m[z(\mathcal{Z}_t), x(y_{t-1}, z_{-1}(y_{t-1}, \mathcal{Z}_t))]$ (m defined in Eq. (4.10)).

Note that we have:

$$\begin{cases} z_t & \equiv z(\mathcal{Z}_t) & = (\mathbf{1}' \otimes Id_n) \mathcal{Z}_t \\ z_{t-1} & \equiv z_{-1}(\mathcal{Z}_t) & = (Id_n \otimes \mathbf{1}') \mathcal{Z}_t. \end{cases}$$

The Kitagawa-Hamilton filter (Proposition 4.3) can then be employed, with $F_t = y_t$ and:

$$f(y_t | \mathcal{Z}_t, \underline{y_{t-1}}) = |B_1^{-1}| \times f\left(\begin{bmatrix} x(y_t, z(\mathcal{Z}_t)) \\ y_{2,t} \end{bmatrix} | \mathcal{Z}_t, \underline{y_{t-1}}\right), \quad (4.16)$$

where the computation of the last term is detailed in Proposition 4.4.

Example 4.4 (Interest-rate model with monetary policy-related regimes). Renne (2017) proposes a model where regimes have monetary-policy interpretations. The model is estimated on daily data. The short-term rate is the euro-area overnight interbank rate (EONIA). It is modeled as follows:

$$r_t = \underbrace{\bar{r}_t}_{\text{Target}} + \underbrace{x_t}_{\text{EONIA spread}}. \quad (4.17)$$

The target rate \bar{r}_t has a step-like path $\bar{r}_t = \Delta' z_{r,t}$, where

- $\Delta = [0 \quad 0.25 \quad 0.50 \quad \dots \quad \bar{r}_{max}]'$ and
- $z_{r,t} = [0 \quad \dots \quad 0 \quad 1 \quad 0 \quad \dots \quad 0]'$
- EONIA spread (x_t) persistent and mean-reverting fluctuations (AR process).
- $z_t = z_{r,t} \otimes z_{m,t}$ where $z_{m,t}$ is the monetary-policy regime:
- Easing ($z_{m,t} = [1 \quad 0 \quad 0]$),
- Status Quo ($z_{m,t} = [0 \quad 1 \quad 0]$),
- Tightening ($z_{m,t} = [0 \quad 0 \quad 1]$).

One can simulate this model by using this web-interface.

In terms of observability, $z_{r,t}$ is observed (since the policy rate is observed), but not $z_{m,t}$. Hence, a filtering procedure is needed. Renne (2017) adapts the approach presented in Subsection 4.5.2 to the case where z_t is partially observed.

In this model, the results of Example 4.3 imply that we have:

$$R(t, h) = A'_h z_t + B_h x_t.$$

Denote by \mathcal{A}_h the $(3 \times n_r)$ matrix such that $A_h = \text{vec}(\mathcal{A}_h)$. We then have $A'_h z_t = (\mathcal{A}_h z_{r,t})' z_{m,t}$ and, therefore:

$$R(t, h) = A'_{t,h} z_t + B_h x_t, \quad \text{where } A_{t,h} = \mathcal{A}_h z_{r,t}.$$

The model is estimated by a combination of Kitagawa-Hamilton and inversion techniques, assuming that a linear combination of yields is modeled without errors (which gives $x_t = x(y_t, z_{m,t}, z_{r,t})$).

4.6 A typical small-sample issue

Interest rates are particularly persistent variables. Since affine models eventually lead to linear relationships between state variables and interest rates (see Eq. (3.8)), some state variables are also necessarily highly persistent. Accordingly, in small sample, maximum-likelihood estimates of the model parameters are likely to suffer from a downward bias (see, e.g., Michael D. Bauer and Wu (2012) or Jardet et al. (2013)). This relates to a well-known econometric problem illustrated by Figure 4.7.

This small-sample downward bias has dramatic consequences in terms of term premium estimates. For the sake of illustration, consider the following process for the short-term interest rate under the physical measure (monthly frequency):

$$i_{t+1} = \bar{i} + \phi(i_t - \bar{i}) + \sigma \varepsilon_{t+1}, \quad \varepsilon_t \sim \mathcal{N}(0, 1).$$

and the following under the risk-neutral measure:

$$i_{t+1} = \bar{i}^* + \phi^*(i_t - \bar{i}^*) + \sigma \varepsilon_{t+1}, \quad \varepsilon_t \sim \mathcal{N}(0, 1).$$

with $\bar{i} = 3\%/12$, $\bar{i}^* = 5\%/12$, $\phi = 0.98$, $\phi^* = 0.99$, $\sigma = 0.2\%$. Assume the estimate of ϕ is downward biased ($\hat{\phi} = 0.9$). The influence of that bias on the 10-year term premium is illustrated by Figure 4.8.

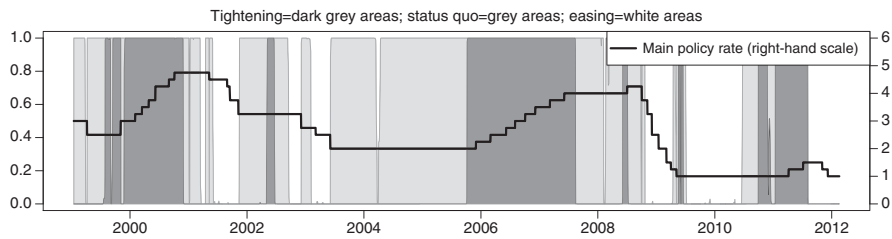


Figure 3: Regimes' estimates.

This figure displays the (smoothed) probabilities of being in the different monetary-policy regimes (left-hand scale). The main policy rate is also shown (right-hand scale).

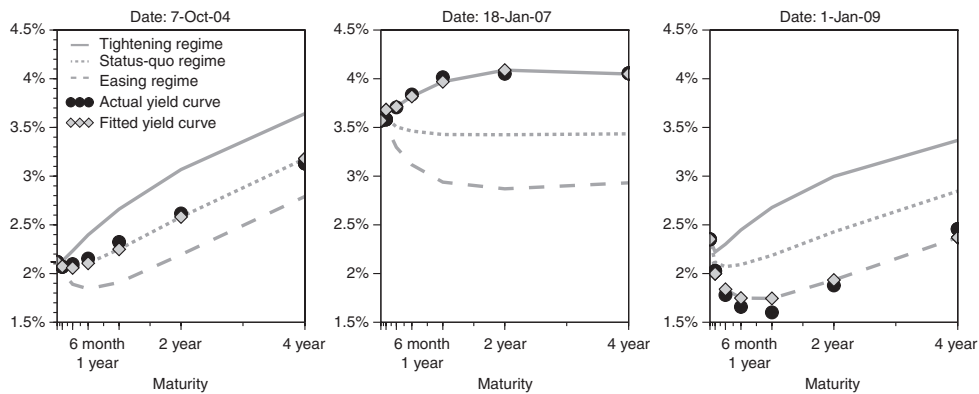


Figure 4: Fitted yield curves and influence of monetary-policy regimes.

Figure 4.6: Source: Renne (2017).

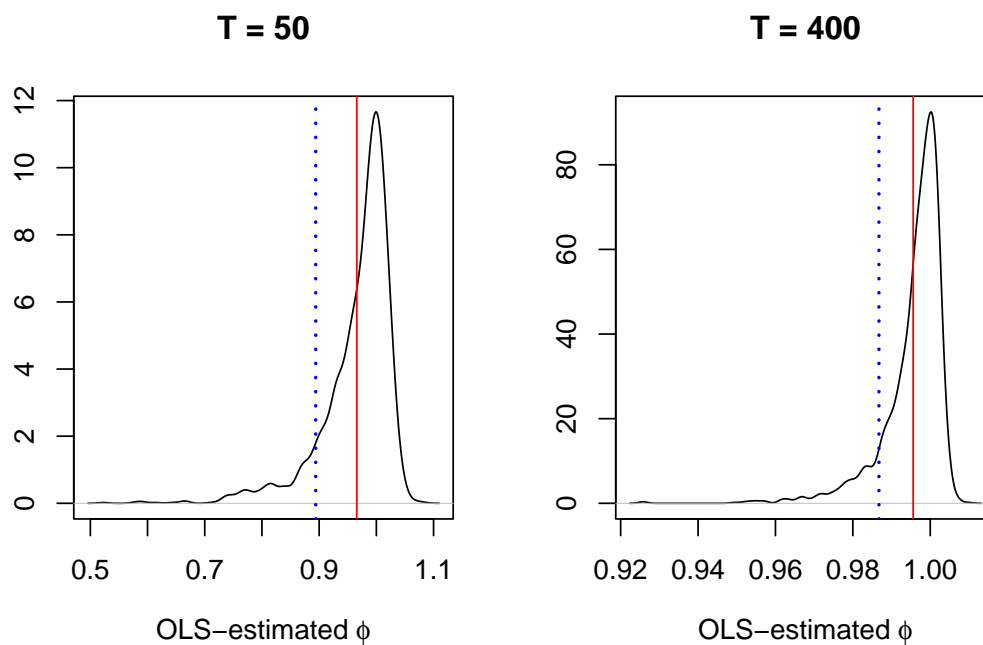


Figure 4.7: 1000 random walk samples of size T ($T = 50$ for the left plot and $T = 400$ for the right plot) have been simulated. For each sample, we run the OLS regression $y_t = \phi y_{t-1} + \varepsilon_t$ to estimate ϕ (whose true value is 1). The plots show the distributions (kernel-based estimation) of the estimated ϕ . The vertical red bar indicate the means of the distributions; the vertical blue line shows the usual bias approximation $(1 - 5.3/T)$.

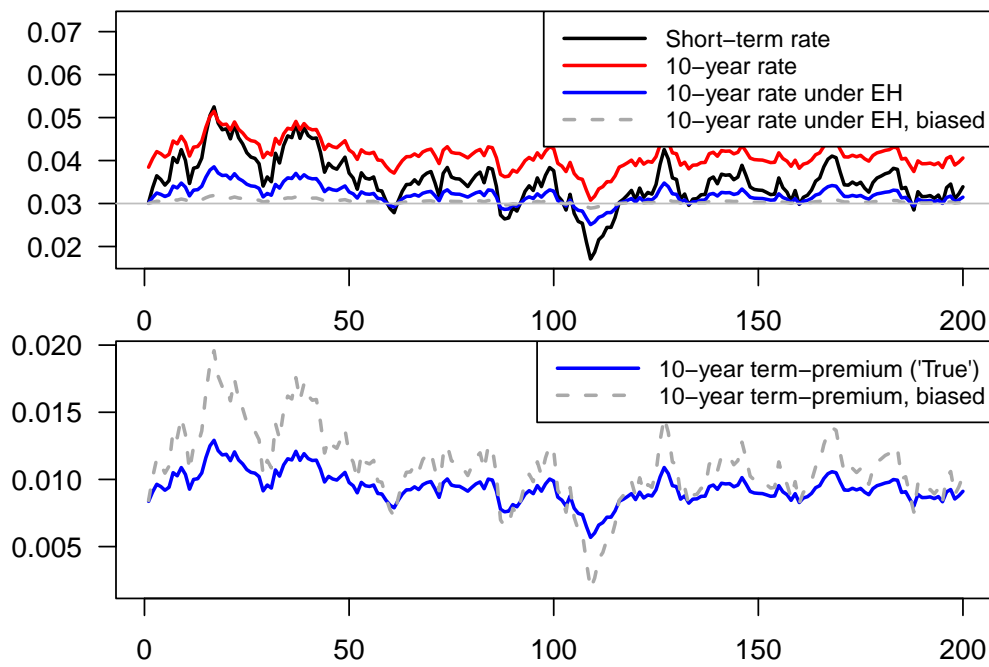


Figure 4.8: This figure illustrates the influence of the downward bias on on estimated term premiums, see the text for more details.

Kim and Orphanides (2005) have proposed a simple approach to deal with this problem. Their approach consists in adding measurement equations to impose that the model-implied forecasts are—up to some measurement errors—equal to survey-based ones. In their empirical exercise, they use the Blue Chip Financial Forecasts. Alternative (publicly available) surveys are: the Philly Fed Survey of Professional Forecasters and the ECB SPF.

Kim and Orphanides (2005) exploit the fact that, in the context of affine models, model-implied forecasts of yields are affine in the state vector (see see Eq. (3.11)); this implies that their measurement equations are affine in the state vector, which facilitates the estimation of the latent factors. As shown by Figure 4.9, their model is able to satisfyingly fit both survey-based forecasts and market yields.

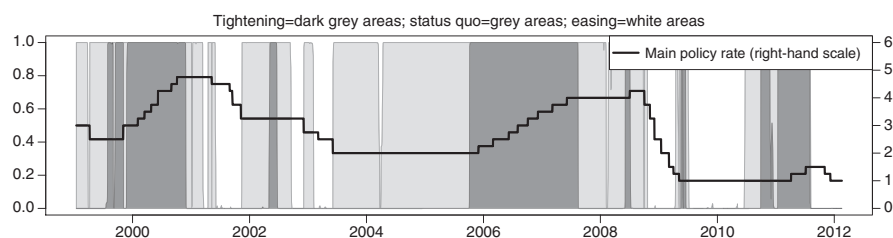


Figure 3: Regimes' estimates.

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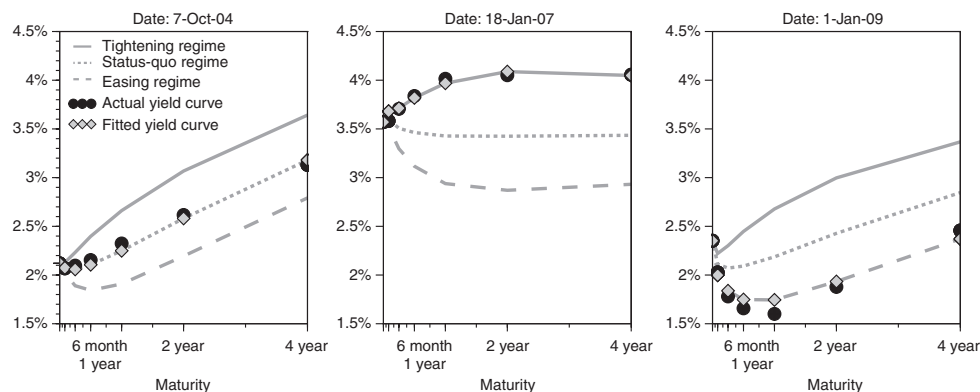


Figure 4: Fitted yield curves and influence of monetary-policy regimes.

Figure 4.9: Source: Kim and Orphanides (2005).

Chapter 5

Forward, futures, dividends, commodity pricing, and convenience yields

5.1 Forward contracts

Let us denote by \underline{w}_t the information available at t , i.e., $\underline{w}_t = \{w_t, w_{t-1}, \dots\}$.

A Forward contract is defined as an agreement, signed at t , to buy/sell an asset (a commodity) at a given delivery date $T > t$ at a price $\Phi_{t,T}$, called *delivery price* or *forward price*, decided at t . We denote by S_T the value of the asset (commodity) at T ; this is a function of \underline{w}_T . For the buyer, the payoff of the contract at T is $S_T - \Phi_{t,T}$.

In this chapter, we denote by r_t the interest rate between t and $t+1$ (known at t), it is assumed to be a function of \underline{w}_t . We also denote by $B(t_1, t_2 - t_1)$ the date- t_1 price of a zero-coupon bond whose value is 1 on t_2 .

The following proposition provides an expression for the forward price $\Phi_{t,T}$:

Proposition 5.1 (Forward price). *We have:*

$$\begin{aligned}\Phi_{t,T} &= \frac{\mathbb{E}_t(\mathcal{M}_{t,t+1}, \dots, \mathcal{M}_{t-1,T} S_T)}{B(t, T-t)} \\ &= \frac{\mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{T-1}) S_T]}{B(t, T-t)}.\end{aligned}$$

Proof. The price at t of $S_T - \Phi_{t,T}$ is zero. Therefore: $\mathbb{E}_t(\mathcal{M}_{t,t+1}, \dots, \mathcal{M}_{t-1,T} S_T) - \Phi_{t,T} \mathbb{E}_t(\mathcal{M}_{t,t+1} \dots \mathcal{M}_{t-1,T}) = 0$, which gives the result. \square

If processes (r_t) and (S_t) are \mathbb{Q} -independent, then we have

$$\Phi_{t,T} = \mathbb{E}_t^{\mathbb{Q}}(S_T),$$

and $\Phi_{t,T}$ therefore is a \mathbb{Q} -martingale.

If the asset does not generate any payoff before T (in the case of shares, there is no dividends), then

$$\Phi_{t,T} = \frac{S_t}{B(t, T - t)}.$$

The price at s , with $t < s < T$, of a forward contract signed at t is:

$$(\Phi_{s,T} - \Phi_{t,T})B(s, T - s),$$

since the payoff $S_T - \Phi_{t,T}$ at T has this price at s .

How are these formulas affected by the presence of dividends? To investigate that, let us introduce additional notations. Denote by \tilde{S}_t the ex-dividend price at t , and by S_t : the cum-dividend price at t . We have $S_t = \tilde{S}_t \exp(\delta_t)$, where δ_t is the dividend yield (or rate), observed on date t . We get:

$$\begin{aligned} \tilde{S}_{T-1} &= \mathbb{E}_{T-1}(\mathcal{M}_{t-1,T} S_T) \\ \text{or} \\ S_{T-1} &= \mathbb{E}_{T-1}[\mathcal{M}_{t-1,T} \exp(\delta_{T-1}) S_T], \end{aligned}$$

and, recursively:

$$\begin{aligned} S_t &= \mathbb{E}_t[\mathcal{M}_{t,t+1} \dots \mathcal{M}_{t-1,T} \exp(\delta_t + \dots + \delta_{T-1}) S_T] \\ &= \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t \dots - r_{T-1} + \delta_t + \dots + \delta_{T-1}) S_T]. \end{aligned}$$

This shows that the formulas are as in the non-dividend case, except that r_t is replaced by $r_t - \delta_t$.

Replacing S_t by $\tilde{S}_t \exp(\delta_t)$ and S_T by $\tilde{S}_T \exp(\delta_T)$, we get

$$\tilde{S}_t = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{T-1} + \delta_{t+1} + \dots + \delta_T) \tilde{S}_T].$$

When the dividend rate δ_t is deterministic, we have:

$$\Phi_{t,T} = \frac{S_t \exp(-\delta_t - \dots - \delta_{T-1})}{B(t, T-t)}.$$

And if r_t is deterministic too, we have:

$$\Phi_{t,T} = S_t \exp(r_t + \dots + r_{T-1} - \delta_t - \dots - \delta_{T-1}).$$

5.2 Futures contracts, futures prices

A futures contract is an agreement, signed at t , to buy/sell an asset (a commodity) at given delivery date $T > t$ at a price $F_{t,T}$, called *futures price*, decided at t . The difference with forward contract is that both counterparties are required to deposit into a *margin account*, at every trading day $s > t$, the resettlement payment (*margin call*). The latter is equal to:

$$\Delta_{s,T} = F_{s,T} - F_{s-1,T} \quad (\text{for the buyer}).$$

Hence, a futures contract is actually closed out after every day, starts afresh the next day. Its value therefore is zero.

The following proposition values a futures contract.

Proposition 5.2 (Pricing futures). *We have:*

$$F_{t,T} = \mathbb{E}_t^{\mathbb{Q}}(S_T),$$

that is $F_{t,T}$ is a \mathbb{Q} -martingale, and $\Delta_{t,T}$ is a \mathbb{Q} -martingale difference.

Proof. At each date $s \geq t$, after the deposit of the resettlement payment, there is a new contract valued zero and paying $F_{s+1,T} - F_{s,T}$ at $s+1$ (and providing another zero valued contract at $s+1$). Therefore $0 = \mathbb{E}_s^{\mathbb{Q}}[\exp(-r_s)(F_{s+1,T} - F_{s,T})]$, and $0 = \mathbb{E}_s^{\mathbb{Q}}(F_{s+1,T} - F_{s,T})$ since $\exp(-r_s)$ is known at s . Hence $F_{s,T} = \mathbb{E}_s^{\mathbb{Q}}(F_{s+1,T})$, and the results follows from $F_{T,T} = S_T$. \square

Proposition 5.3 (Forward-Futures deviation). *We have:*

$$\Phi_{t,T} - F_{t,T} = \frac{\text{Cov}_t^{\mathbb{Q}} \left[\prod_{s=t}^{T-1} \exp(-r_s), S_T \right]}{B(t, T-t)}.$$

Proof. We have:

$$\begin{aligned}
\Phi_{t,T} - F_{t,T} &= \mathbb{E}_t^{\mathbb{Q}} \left[\frac{\prod_{s=t}^{T-1} \exp(-r_s) S_T}{B(t, T-t)} - S_T \right] \\
&= \frac{\mathbb{E}_t^{\mathbb{Q}} \left[\prod_{s=t}^{T-1} \exp(-r_s) S_T \right] - \mathbb{E}_t^{\mathbb{Q}} \left[\prod_{s=t}^{T-1} \exp(-r_s) \right] \mathbb{E}_t^{\mathbb{Q}}(S_T)}{B(t, T-t)} \\
&= \frac{\text{cov}_t^{\mathbb{Q}} \left[\prod_{s=t}^{T-1} \exp(-r_s), S_T \right]}{B(t, T-t)}.
\end{aligned}$$

□

Hence, $\Phi_{t,T} = F_{t,T}$ if, and only if, $\prod_{s=t}^{T-1} \exp(-r_s)$ and S_T are conditionally uncorrelated under \mathbb{Q} . It is true, in particular, in the case of deterministic short rates.

5.3 Convenience yields

A convenience yield is defined as the net benefit associated with holding a physical asset (rather than a forward or futures contract). It is *net* in the sense that it is equal to the positive gain of holding minus potential storage costs.

The notion of convenience yield is relevant only for storable commodities (not, e.g., for electricity). It can be positive or negative. The notion of convenience yield is mathematically similar to that of a dividend yield (but can be < 0 , and is often latent).

In the following, we denote the convenience yield by c_t . The price is here the cum-convenience yield price.

Let us first consider the case where r_t and c_t are deterministic. We have:

$$\Phi_{t,T} = F_{t,T} = S_t \exp(r_t + \dots + r_{T-1} - c_t - \dots - c_{T-1}).$$

Moreover, if r_t and c_t are time independent, we get:

$$\Phi_{t,T} = F_{t,T} = S_t \exp[(T-t)(r-c)].$$

If $r > c$, the forward curve is an increasing function of maturity T ; this is known as a situation of *contango*. If $r < c$, the forward curve is a decreasing function of maturity T , known as a situation of *backwardation*.

Let us turn to the situation where c_t is stochastic. We have:

$$\Phi_{t,T} = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{T-1})S_T]/B(t, T-t) \quad (5.1)$$

$$F_{t,T} = \mathbb{E}_t^{\mathbb{Q}}(S_T) \quad (5.2)$$

$$S_t = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{T-1} + c_t + \dots + c_{T-1})S_T]. \quad (5.3)$$

Going further necessitates a joint modelling—at least under \mathbb{Q} —of S_t (or $s_t = \log S_t$), c_t (convenience yield), and r_t (short rate).

In the case of a **nonstorable commodity**, there is no convenience yield. If, moreover r_t deterministic, $\Phi_{t,T} = F_{t,T} = \mathbb{E}_t^{\mathbb{Q}}(S_T)$, but since Eq. (5.3) is not valid we do not have $S_t = B(t, T)\mathbb{E}_t^{\mathbb{Q}}(S_T)$ and therefore we do not have $\Phi_{t,T} = S_t/B(t, T)$.

In the case of a **storable commodity**, the price ex-convenience yield at t , i.e. $S_t \exp(-c_t)$, is the price at t of S_{t+1} :

$$\begin{aligned} S_t \exp(-c_t) &= \exp(-r_t) \mathbb{E}_t^{\mathbb{Q}}(S_{t+1}) \\ \mathbb{E}_t^{\mathbb{Q}}(S_{t+1}) &= S_t \exp(r_t - c_t) \\ \mathbb{E}_t^{\mathbb{Q}} \exp(s_{t+1}) &= \exp(s_t + r_t - c_t). \end{aligned} \quad (5.4)$$

Or if $y_{t+1} = \log \frac{S_{t+1}}{S_t}$ denotes the (geometric) return, then:

$$\mathbb{E}_t^{\mathbb{Q}} \exp(y_{t+1}) = \exp(r_t - c_t).$$

Eq. (5.3) is automatically satisfied (using Eq. (5.4) recursively).

5.4 Pricing with affine models

In this section, we consider a storable commodity. The state vector is as follows:

$$w_t = (s_t, c_t, r_t, x_t')',$$

where $s_t = \log S_t$, and where x_t is a vector of additional factors. We further assume that w_t is a \mathbb{Q} -affine process, that is:

$$\mathbb{E}_t^{\mathbb{Q}} \exp(u' w_{t+1}) = \exp[a'(u)w_t + b(u)].$$

Since $S_{t+1} = \exp(e_1' w_{t+1})$, Internal Consistency Constraints (ICCs) apply (see Subsection 2.5.1), and we have:

$$\begin{aligned} \mathbb{E}_t^{\mathbb{Q}} \exp(s_{t+1}) &= \exp(s_t - c_t + r_t) \\ \Leftrightarrow \mathbb{E}_t^{\mathbb{Q}} \exp(e_1' w_{t+1}) &= \exp[(e_1 - e_2 + e_3)' w_t] \\ \Leftrightarrow \begin{cases} a(e_1) &= e_1 - e_2 + e_3 \\ b(e_1) &= 0. \end{cases} \end{aligned}$$

If the commodity is non-storable, there are no c_t , and no ICCs.

Eqs. (5.1) and (5.2) give:

$$\begin{aligned} \Phi_{t,T} &= \frac{\mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{T-1} + s_T)]}{\mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{T-1})]} \\ F_{t,T} &= \mathbb{E}_t^{\mathbb{Q}}[\exp(s_T)]. \end{aligned}$$

Using the multi-horizon Laplace transforms of $w_t = (s_t, c_t, r_t, x_t')'$, we get quasi explicit formulas for these prices (as they are exponential affine in future values of w_t , which is an affine process).

5.5 Historical dynamics

Using the notation of Subsection 2.3.2, assume the s.d.f. is of the form:

$$\mathcal{M}_{t,t+1}(\underline{w}_{t+1}) = \exp \left[-r_t + \alpha_t' w_{t+1} + \psi_t^{\mathbb{Q}}(-\alpha_t) \right],$$

such that

$$\mathbb{E}_t^{\mathbb{Q}}(\mathcal{M}_{t,t+1}^{-1}) = \exp(r_t).$$

The historical dynamics of w_t is defined through:

$$\psi_t(u) = \psi_t^{\mathbb{Q}}(u - \alpha_t) - \psi_t^{\mathbb{Q}}(-\alpha_t).$$

There is no constraint on α_t , which can be used, e.g., to specify some forms of seasonality (see Example 5.1).

Example 5.1 (A Gaussian VAR model). Consider the state vector $w_t = (s_t, c_t, r_t)'$, whose R.N. dynamics reads:

$$w_{t+1} = A_0 + A_1 w_t + \varepsilon_{t+1}, \quad \varepsilon_{t+1} \sim i.i.d. \mathcal{N}(0, \Sigma) \text{ under } \mathbb{Q}.$$

We then have (see Example 1.6):

$$\begin{aligned} \mathbb{E}_t \exp(u' w_{t+1}) &= \exp \left[u'(A_0 + A_1 w_t) + \frac{1}{2} u' \Sigma u \right] \\ \Rightarrow \quad \begin{cases} a(u) &= A_1' u \\ b(u) &= A_0' u + \frac{1}{2} u' \Sigma u. \end{cases} \end{aligned}$$

Decompose A_0 and A_1 as follows:

$$A_0 = \begin{pmatrix} A_{01} \\ \tilde{A}_0 \end{pmatrix}, \quad A_1 = \begin{pmatrix} A_{11} \\ \tilde{A}_1 \end{pmatrix}.$$

If the considered commodity is storable, we have the following ICCs (no constraint otherwise):

$$\begin{cases} a(e_1) &= e_1 - e_2 + e_3 \\ b(e_1) &= 0. \end{cases}$$

Therefore, the first row of A_1 is $A_{11} = e_1' - e_2' + e_3'$, the first element of A_0 is $A_{01} = -\frac{1}{2}\sigma_1^2$ (σ_1^2 being the conditional variance of s_{t+1}).

In other words the \mathbb{Q} -VAR is:

$$\begin{cases} \begin{pmatrix} s_{t+1} \\ c_{t+1} \\ r_{t+1} \end{pmatrix} &= \begin{pmatrix} -\frac{1}{2}\sigma_1^2 + s_t - c_t + r_t \\ \tilde{A}_0 + \tilde{A}_1 w_t \end{pmatrix} + \begin{pmatrix} \varepsilon_{1,t+1} \\ \varepsilon_{2,t+1} \\ \varepsilon_{3,t+1} \end{pmatrix}, \end{cases}$$

where \tilde{A}_0 and \tilde{A}_1 are not constrained.

Noting $y_{t+1} = \log(S_{t+1}/S_t) = s_{t+1} - s_t$, this implies that

$$\mathbb{E}_t^{\mathbb{Q}} y_{t+1} = r_t - c_t - \frac{1}{2}\sigma_1^2.$$

Hence, $y_{t+1} - r_t + c_t + \frac{1}{2}\sigma_1^2$ is a \mathbb{Q} -martingale difference.

The historical dynamics of the state vector is as follows:

$$\begin{aligned}\psi_t(u) &= \psi_t^{\mathbb{Q}}(u - \alpha_t) - \psi_t^{\mathbb{Q}}(-\alpha_t) \\ &= u'(A_0 + A_1 w_t) + \frac{1}{2}(u - \alpha_t)' \Sigma (u - \alpha_t) - \frac{1}{2} \alpha_t' \Sigma \alpha_t \\ &= u'(A_0 + A_1 w_t) - u' \Sigma \alpha_t + \frac{1}{2} u' \Sigma u.\end{aligned}$$

If we take $\alpha_t = \alpha_0 + \alpha_1 w_t$, we get:

$$\begin{aligned}\psi_t(u) &= u'[A_0 - \Sigma \alpha_0 + (A_1 - \Sigma \alpha_1) w_t] + \frac{1}{2} u' \Sigma u \\ \Rightarrow w_{t+1} &= A_0 - \Sigma \alpha_0 + (A_1 - \Sigma \alpha_1) w_t + \xi_t, \\ &\quad \xi_t \sim i.i.d. \mathcal{N}(0, \Sigma) \text{ under } \mathbb{P}.\end{aligned}$$

Hence, any VAR(1), with same Σ , can be reached.

We can also take a time-varying specification for α_t , of the form $\alpha_t = \alpha_{0t} + \alpha_1 w_t$. We then get the historical dynamics:

$$w_{t+1} = A_0 - \Sigma \alpha_{0t} + (A_1 - \Sigma \alpha_1) w_t + \xi_{t+1}$$

We can for instance choose α_{0t} such that: $A_0 - \Sigma \alpha_{0t} = \begin{pmatrix} \mu_{1t} \\ \mu_2 \\ \mu_3 \end{pmatrix}$, where μ_{1t} , μ_2 , and μ_3 are given. In particular:

$$\begin{aligned}s_{t+1} &= \mu_{1t} + (A_{11} - \Sigma_1 \alpha_1) w_t + \xi_{1,t+1} \\ &= \mu_{1t} + \bar{A}_{11} w_t + \xi_{1,t+1} \text{ (say),}\end{aligned}$$

where \bar{A}_{11} is not constrained, or:

$$s_{t+1} - \nu_{t+1} = \mu_1 + \bar{A}_{11,1}(s_t - \nu_t) + \bar{A}_{11,2}c_t + \bar{A}_{11,3}r_t + \xi_{1,t+1}$$

with

$$\mu_{1t} = \mu_1 + \nu_{t+1} - \bar{A}_{11,1}\nu_t.$$

In other words any historical seasonal function ν_t can be reached by choosing μ_{1t} , i.e., α_{0t} .

Example 5.2 (Schwartz (1997)). Schwartz (1997) proposes three models whose discrete-time versions are:

- **Model 1:**

$$\begin{aligned}\mathbb{P} &: s_{t+1} = a_0 + a_1 s_t + \sigma \varepsilon_{t+1}, \quad \varepsilon_t \stackrel{\mathbb{P}}{\sim} i.i.d.\mathcal{N}(0, 1) \\ \mathbb{Q} &: s_{t+1} = a_0^* + a_1 s_t + \sigma \xi_{t+1}, \quad \xi_t \stackrel{\mathbb{Q}}{\sim} i.i.d.\mathcal{N}(0, 1).\end{aligned}$$

- **Model 2:**

$$\begin{aligned}\mathbb{P} &: s_{t+1} = a_{10} + s_t - c_t + \sigma_1 \varepsilon_{1,t+1} \\ &\quad c_{t+1} = a_{20} + a_{21} c_t + \sigma_2 \varepsilon_{2,t+1} \quad \text{corr}(\varepsilon_{1t}, \varepsilon_{2t}) = \rho \\ \mathbb{Q} &: s_{t+1} = r - \frac{\sigma_1^2}{2} + s_t - c_t + \sigma_1 \xi_{1,t+1} \\ &\quad c_{t+1} = a_{20}^* + a_{21} c_t + \sigma_2 \xi_{2,t+1} \quad \text{corr}(\xi_{1t}, \xi_{2t}) = \rho.\end{aligned}$$

- **Model 3:**

$$\begin{aligned}\mathbb{P} &: s_{t+1} = a_{10} + s_t - c_t + r_t + \sigma_1 \varepsilon_{1,t+1} \\ &\quad c_{t+1} = a_{20} + a_{21} c_t + \sigma_2 \varepsilon_{2,t+1} \\ &\quad r_{t+1} = a_{30} + a_{31} r_t + \sigma_3 \varepsilon_{3,t+1} \\ \mathbb{Q} &: s_{t+1} = -\frac{\sigma_1^2}{2} + s_t - c_t + r_t + \sigma_1 \xi_{1,t+1} \\ &\quad c_{t+1} = a_{20}^* + a_{21} c_t + \sigma_2 \xi_{2,t+1} \\ &\quad r_{t+1} = a_{30} + a_{31} r_t + \sigma_3 \xi_{3,t+1},\end{aligned}$$

with $\text{Corr}(\varepsilon_{1t}, \varepsilon_{2t}) = \rho_1$, $\text{Corr}(\varepsilon_{2t}, \varepsilon_{3t}) = \rho_2$, $\text{Corr}(\varepsilon_{1t}, \varepsilon_{3t}) = \rho_3$, and idem for the ξ_{it} 's. (All shocks are $\mathcal{N}(0, 1)$.)

In these models, $\log F_{t,T}$ is an affine function of $\{s_t, c_t, r_t\}$.

From an estimation point of view, s_t and c_t are latent. The parameters are estimated by the ML method, the likelihood functions being computed by the Kalman filter (see Subsection 4.2). The state variable is s_t in Model 1, and $\{s_t, c_t\}$ in Models 2 and 3. (The \mathbb{P} and \mathbb{Q} dynamics of r_t are estimated separately.)

As can be seen in the previous equations, constraints are imposed between the \mathbb{P} and \mathbb{Q} parameters; this is to help the identification of the model parameters.

Table XIII
Cross-Section Comparison Between Models 1, 2, and 3 Out of Sample
Copper Data 7/29/88 to 6/13/95

| Model Contract | RMSE | | | Mean Error | | |
|------------------------|-------|-------|-------|------------|--------|--------|
| | 1 | 2 | 3 | 1 | 2 | 3 |
| Panel A: In Cents | | | | | | |
| F2 | 4.617 | 1.311 | 1.27 | 1.75 | -0.155 | -0.142 |
| F4 | 1.17 | 0.548 | 0.539 | 0.327 | 0.116 | 0.119 |
| F6 | 1.133 | 0.616 | 0.621 | -0.194 | -0.105 | -0.098 |
| F8 | 1.824 | 0.534 | 0.482 | 0.257 | -0.042 | -0.064 |
| All | 2.612 | 0.819 | 0.794 | 0.535 | -0.046 | -0.046 |
| Panel B: In Percentage | | | | | | |
| F2 | 3.929 | 1.103 | 1.073 | 1.404 | -0.151 | -0.135 |
| F4 | 1.063 | 0.483 | 0.474 | 0.284 | 0.093 | 0.094 |
| F6 | 1.064 | 0.568 | 0.572 | -0.188 | -0.102 | -0.097 |
| F8 | 1.78 | 0.512 | 0.462 | 0.213 | -0.036 | -0.058 |
| All | 2.284 | 0.713 | 0.692 | 0.428 | -0.049 | -0.049 |

Figure 5.1: Out of sample means for maturities not used at the estimation stage. F1 contract : the closest to maturity F2 : the second contract to maturity and so one. Source: Schwartz (1997).

Chapter 6

Credit and liquidity risks

6.1 Notations

The information of the investor at date t is still denoted by $w_t = (w'_t, w'_{t-1}, \dots, w'_1)'$, but w_t now comprises a new subsector, namely d_t , that keeps track of the “default” status of some creditors. Formally, $w_t = (y'_t, d'_t)'$, where y_t is a n_y -dimensional vector of common factors, and $d_t = (d_{1,t}, \dots, d_{E,t})'$ a E -dimensional vector of binary variables representing the possible default of entities $e \in \{1, \dots, E\}$. Vector w_t is K -dimensional, that is $K = n_y + E$.

By convention:

- $d_{e,t} = 1$ if entity e is in default status at time t ,
- $d_{e,t} = 0$ if entity e is not in default status at time t .

We assume that the process $\{w_t\}$ is Markov. Formally, its historical (\mathbb{P}) dynamics is defined by the conditional densities:

$$f(y_t|w_{t-1}) \text{ and } p(d_t|y_t, w_{t-1}).$$

In what follows, we assume that the default state is absorbing:

Hypothesis 6.1 (absorbing defaults). The default event is an absorbing state. Formally: $p_e(0|y_t, w_{t-1}, d_{e,t-1} = 1) = 0$.

The *survival probability* is defined as follows:

$$p_e(0|y_t, w_{t-1}, d_{e,t-1} = 0) = \exp[- \underbrace{\lambda_e(y_t, w_{t-1})}_{\text{default intensity}}]. \quad (6.1)$$

It is easily checked that $\lambda_e(y_t, w_{t-1})$ is close to the conditional probability of default when it is small.

6.2 Pricing under standard assumptions

To start with, we will consider the pricing of defaultable bonds under assumptions that are standard in the literature, a framework that we call the *classical credit-risk framework*. These assumptions—namely 6.2 and 6.3—lead to simple formulas. We will investigate further what happens when these assumptions are relaxed (Subsection 6.4).

Hypothesis 6.2 (Non-systemic entity). $\{d_t\}$ does not Granger-cause $\{y_t\}$, in the sense that:

$$f(y_t|w_{t-1}) = f(y_t|y_{t-1}).$$

In other words, under Hypothesis 6.2, y_t is exogenous to d_t .

Hypothesis 6.3 (Unpriced credit events). The default events (or credit events) are not priced, in the sense that:

$$\mathcal{M}_{t-1,t}(w_t, w_{t-1}) = \mathcal{M}_{t-1,t}(y_t, y_{t-1}).$$

In other words, under Hypothesis 6.3, defaults do not affect the SDF (conditionally on y_t).

Hypothesis 6.3 also implies that the short-term risk-free rate r_t ($= -\log \mathbb{E}_t(\mathcal{M}_{t,t+1})$) depends on y_t (but not on d_t).

Figure 6.1 provides a graphical representation of the causality scheme in the classical credit-risk framework.

Proposition 6.1 (Risk-neutral exogeneity of factors). *Under Hypotheses 6.2 and 6.3, we have:*

$$f^{\mathbb{Q}}(y_t|w_{t-1}) = f^{\mathbb{Q}}(y_t|y_{t-1}),$$

that is, $\{d_t\}$ does not Granger-cause $\{y_t\}$ under the risk-neutral measure.

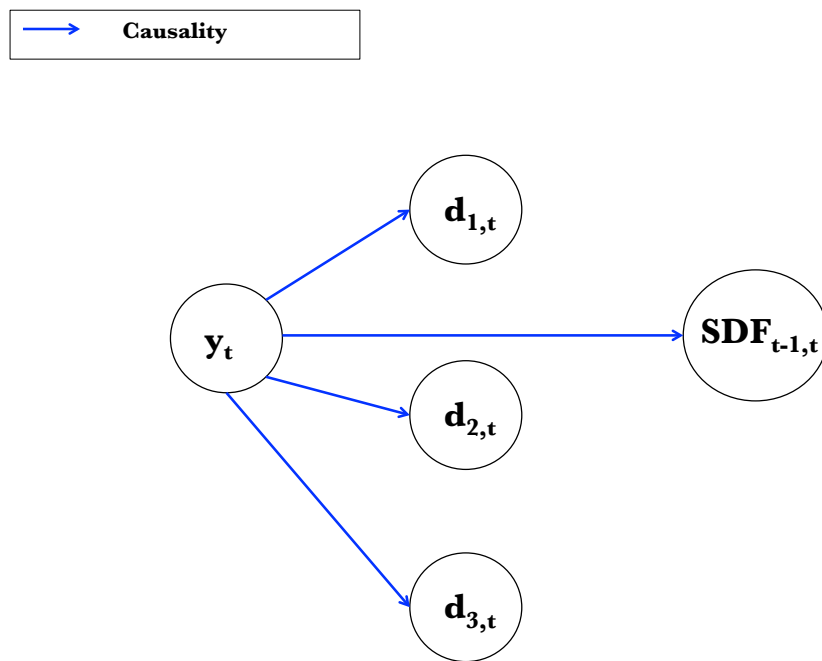


Figure 6.1: Graphical representation of the causality scheme in the classical credit-risk framework.

Proof. We have:

$$f^{\mathbb{Q}}(y_t|w_{t-1}) = f(y_t|w_{t-1})\mathcal{M}_{t-1,t}\exp(r_{t-1}) \quad (6.2)$$

$$= f(y_t|y_{t-1})\mathcal{M}_{t-1,t}\exp(r_{t-1}), \quad (6.3)$$

which gives the results, using that $\mathcal{M}_{t-1,t}$ and $\exp(r_{t-1})$ depend on y_{t-1} only (and not on d_t). \square

To simplify, let us focus on a single entity ($d_t \equiv d_{1,t}$), alive on date t ($d_t = 0$). Consider the date- t price of a bond issued by the defaultable entity. The residual maturity of this bond is denoted by h , and we consider a zero recovery rate (in case of default, the payoff is zero). The price of this bond writes:

$$B(t, h) = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{t+h-1})(1 - d_{t+h})].$$

The following proposition is at the core of the *classical credit-risk framework* (Duffie and Singleton, 1999):

Proposition 6.2 (Defaultable zero-coupon bond pricing in the classical credit-risk framework). *Under Hypotheses 6.2 and 6.3, and with a default intensity defined—as in (6.1)—through:*

$$\exp[-\lambda_t^{\mathbb{P}}] := \mathbb{P}(d_t = 0|\underline{y}_t, d_{t-1} = 0), \quad (6.4)$$

we have:

$$\lambda_t^{\mathbb{Q}} = \lambda_t^{\mathbb{P}} (= \lambda_t), \quad (6.5)$$

and

$$\boxed{B(t, h) = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{t+h-1} - \lambda_{t+1} - \dots - \lambda_{t+h})]}. \quad (6.6)$$

which rewrites, using $\tilde{r}_t = r_t + \lambda_{t+1}$:

$$B(t, h) = \mathbb{E}_t^{\mathbb{Q}}[\exp(-\tilde{r}_t - \dots - \tilde{r}_{t+h-1})], \quad (6.7)$$

Proof. We have (Eq. (2.6)):

$$f^{\mathbb{Q}}(d_t, y_t|d_{t-1}, y_{t-1}) = \exp(r_{t-1})\mathcal{M}_{t-1,t}f^{\mathbb{P}}(d_t, y_t|d_{t-1}, y_{t-1}). \quad (6.8)$$

Using Hypothesis 6.3, we obtain, by integrating both sides w.r.t. d_t :

$$\begin{aligned} f^{\mathbb{Q}}(y_t|d_{t-1}, y_{t-1}) &= \exp(r_{t-1})\mathcal{M}_{t-1,t}f^{\mathbb{P}}(y_t|d_{t-1}, y_{t-1}) \\ &= \exp(r_{t-1})\mathcal{M}_{t-1,t} \underbrace{f^{\mathbb{P}}(y_t|y_{t-1})}_{\text{by hypothesis}}. \end{aligned} \quad (6.9)$$

By Bayes, we have:

$$f^{\mathbb{Q}}(d_t|d_{t-1}, y_t, y_{t-1}) = \frac{f^{\mathbb{Q}}(d_t, y_t|d_{t-1}, y_{t-1})}{f^{\mathbb{Q}}(y_t|d_{t-1}, y_{t-1})}.$$

Using Eq. (6.8) (numerator) and Eq. (6.9) (denominator), we get:

$$f^{\mathbb{Q}}(d_t|d_{t-1}, y_t, y_{t-1}) = f^{\mathbb{P}}(d_t|d_{t-1}, y_t, y_{t-1}),$$

which gives (6.5).

Conditioning w.r.t. \underline{y}_{t+h} , the bond price is given by:

$$\begin{aligned} &\mathbb{E}_t^{\mathbb{Q}}[\mathbb{E}_t^{\mathbb{Q}}\{\exp(-r_t - \dots - r_{t+h-1})(1 - d_{t+h})|\underline{y}_{t+h}\}] \\ &= \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{t+h-1})\mathbb{E}_t^{\mathbb{Q}}\{(1 - d_{t+h})|\underline{y}_{t+h}\}] \\ &= \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \dots - r_{t+h-1})\mathbb{Q}\{d_{t+h} = 0|\underline{y}_{t+h}\}] \\ &= \mathbb{E}_t^{\mathbb{Q}}\left[\exp(-r_t - \dots - r_{t+h-1})\prod_{j=1}^h \mathbb{Q}\{d_{t+j} = 0|d_{t+j-1} = 0, \underline{y}_{t+h}\}\right], \end{aligned}$$

Because Granger and Sims non-causalities are equivalent, and since d_t does not Granger-cause y_t under the risk-neutral measure (see Proposition 6.1), we have:

$$\mathbb{Q}\{d_{t+j} = 0|d_{t+j-1} = 0, \underline{y}_{t+h}\} = \underbrace{\mathbb{Q}\{d_{t+j} = 0|d_{t+j-1} = 0, \underline{y}_{t+j}\}}_{=\exp(-\lambda_{t+j}^{\mathbb{Q}})}.$$

As a result:

$$B(t, h) = \mathbb{E}_t^{\mathbb{Q}}\left[\exp\left(-r_t - \dots - r_{t+h-1} - \lambda_{t+1}^{\mathbb{Q}} - \dots - \lambda_{t+h}^{\mathbb{Q}}\right)\right]. \quad (6.10)$$

Using (6.5) in (6.10), we obtain (6.6). \square

Although \mathbb{P} and \mathbb{Q} intensities are the same functions of y_t and y_{t-1} , their historical and risk-neutral dynamics are in general different since y_t 's \mathbb{P} and \mathbb{Q} dynamics are different.

Eq. (6.7) is reminiscent of Eq. (1.9), where the risk-free short-term rate r_t is replaced by the *credit-adjusted short-term rate* \tilde{r}_t .

Importantly, if both r_t and λ_t are affine in y_t , where the latter is an affine process, then bond prices are easily computed by means of recursive formulas (as was the case in the context of Eq. (1.9), see Example 1.20).

In this case, $B(t, h)$ is exponential affine in y_t , that is, it is of the form:

$$B(t, h) = \exp(A_h^{d'} y_t + B_h^d) \quad (\text{say}). \quad (6.11)$$

The previous framework can be adapted in order to accommodate non-zero recovery rates. In this framework, the default intensity will be replaced by a *pseudo*, or *recovery-adjusted default intensity* denoted by $\tilde{\lambda}_t$. To go further, we need to specify the payoff taking place upon default. There exist different modelling conventions for that; one is the so-called *Recovery of Market Value (RMV)* (Duffie and Singleton, 1999). Loosely speaking, under the RMV, the recovery payoff is a fraction $\zeta \in [0, 1]$, called *Recovery Rate (RR)* of the price that would have prevailed, absent the default. Proposition 6.3 makes this definition more precise and gives the bond price stemming from it.

Proposition 6.3 (Bond pricing in the RMV classical framework). *Under Hypotheses 6.2 and 6.3, and if, in case of default on date $t + i$, the recovery payment is $\zeta \tilde{B}(t + i, h - i)$, with $\zeta \in [0, 1]$, then*

$$B(t, h) = (1 - d_t) \tilde{B}(t, h), \quad (6.12)$$

where $\tilde{B}(t, h)$ is a pseudo-price given by:

$$\tilde{B}(t, h) = \mathbb{E}_t^{\mathbb{Q}}[\exp(-r_t - \tilde{\lambda}_{t+1} - \dots - r_{t+h-1} - \tilde{\lambda}_{t+h})], \quad (6.13)$$

where the pseudo-intensity $\tilde{\lambda}_t$ is defined by:

$$\begin{aligned} \exp(-\tilde{\lambda}_{t+1}) &= \mathbb{E}[\{1 - d_{t+1}(1 - \zeta)\} | \underline{y}_{t+1}, d_t = 0] \\ &= 1 - (1 - \zeta)(1 - \exp(-\lambda_{t+1})). \end{aligned} \quad (6.14)$$

Proof. Relation (6.12) is true for $B(t+h, 0)$ since $\widetilde{B}(t+h, 0) = 1$ and $B(t+h, 0) = 1 - d_{t+h}$. Assuming that $B(t+i+1, h-i-1) = (1 - d_{t+i+1})\widetilde{B}(t+i+1, h-i-1)$ (which is valid for $i = h-1$), we get:

$$\begin{aligned} B(t+i, h-i) &= (1 - d_{t+i}) \mathbb{E}_{t+i}^{\mathbb{Q}} \left\{ \exp(-r_{t+i}) \left[(1 - d_{t+i+1}) \widetilde{B}(t+i+1, h-i-1) \right. \right. \\ &\quad \left. \left. + d_{t+i+1} \zeta \widetilde{B}(t+i+1, h-i-1) \right] \right\}, \end{aligned}$$

since, in case of no default at date $t+i$, the value of the bond at date $t+i+1$ is either $\zeta B(t+i+1, h-i-1) = \zeta \widetilde{B}(t+i+1, h-i-1)$ if default happens, and $B(t+i+1, h-i-1) = \widetilde{B}(t+i+1, h-i-1)$ otherwise. Hypotheses 6.2 and 6.3 imply that $\widetilde{B}(t+i+1, h-i-1)$ does not depend on d_{t+i+1} and taking first the conditional expectation given w_{t+i} and y_{t+i+1} we obtain:

$$\begin{aligned} B(t+i, h-i) &= (1 - d_{t+i}) \mathbb{E}_{t+i}^{\mathbb{Q}} \left\{ \exp(-r_{t+i}) \widetilde{B}(t+i+1, h-i-1) \times \right. \\ &\quad \left. \mathbb{E}^{\mathbb{Q}} [(1 - d_{t+i+1}) + \zeta d_{t+i+1} | y_{t+i+1}, w_{t+i}] \right\}. \end{aligned} \quad (6.15)$$

Using the definition of $\tilde{\lambda}_t$ (Eq. (6.14)) we get:

$$\begin{aligned} B(t+i, h-i) &= (1 - d_{t+i}) \mathbb{E}_{t+i}^{\mathbb{Q}} \left\{ \exp(-r_{t+i} - \tilde{\lambda}_{t+i+1}) \widetilde{B}(t+i+1, h-i-1) \right\} \\ &= (1 - d_{t+i}) \widetilde{B}(t+i, h-i). \end{aligned}$$

Therefore it is also true for $i = 0$. □

Again, if r_t and $\tilde{\lambda}_t$ are affine combinations of an affine process y_t , then (exponential affine) bond prices are easily computed by means of recursive formulas. This is formalized below.

Hypothesis 6.4 (Affine process). The process $\{y_t\}$ is affine under the \mathbb{Q} measure:

$$\varphi_{y,t-1}^{\mathbb{Q}}(u_y) = \mathbb{E}^{\mathbb{Q}} \left[\exp(u_y' y_t) | \underline{y}_{t-1} \right] = \exp \left[a_y^{\mathbb{Q}}(u_y)' y_{t-1} + b_y^{\mathbb{Q}}(u_y) \right], \quad (6.16)$$

and r_t and $\tilde{\lambda}_{e,t}$ are affine in (y_t, y_{t-1}) , that is (say):

$$r_t = \omega_0 + \omega_1' y_t \quad (6.17)$$

$$\tilde{\lambda}_{e,t} = \kappa_0 + \kappa_1' y_t. \quad (6.18)$$

Proposition 6.4 (Standard affine bond pricing, under RMV convention). *Under the Assumptions of Proposition 6.3 and under Hypothesis 6.4, we have, for $h \geq 1$:*

$$B(t, h) = (1 - d_t) \exp(-h[\omega_0 + \kappa_0] - \omega'_1 y_t + A'_h y_t + B_h), \quad (6.19)$$

where A_h and B_h are given by the following recursive equations:

$$\begin{cases} A_h &= a(u_h + A_{h-1}), \\ B_h &= b(u_h + A_{h-1}) + B_{h-1}, \\ A_0 &= 0, \quad B_0 = 0, \end{cases} \quad (6.20)$$

with $u_1 = -\kappa_1$ and $u_i = -(\kappa_1 + \omega_1)$ for $i > 1$.

Proof. The results directly follows from Proposition 6.3. \square

What precedes has been established under the assumption that the state vector w_t comprises a single defaultable entity ($E = 1$). What about when different entities are considered ($E = 2$)? We can demonstrate that similar pricing formulas are obtained under the assumption of no-contagion (Hypothesis 6.5):

Hypothesis 6.5 (no contagion). There is no instantaneous or lagged contagion between entities, i.e.:

$$p(d_t | y_t, w_{t-1}) = p_1(d_{1,t} | y_t, y_{t-1}, d_{1,t-1}) \times p_2(d_{2,t} | y_t, y_{t-1}, d_{2,t-1}).$$

If Hypothesis 6.5 is not satisfied, for instance, if we have contagion effect from $e = 1$ towards $e = 2$:

$$\begin{aligned} p(d_t | y_t, w_{t-1}) &= p_1(d_{1,t} | y_t, y_{t-1}, d_{1,t-1}) \times \\ &\quad p_2(d_{2,t} | y_t, y_{t-1}, d_{1,t}, d_{1,t-1}, d_{2,t-1}), \end{aligned}$$

then:

1. Proposition 6.3 is still valid for both entities;
2. In an affine framework, the computation of $\widetilde{B}_1(t, h)$ (and $B_1(t, h)$) is straightforward;

3. but formulas for $\widetilde{B}_2(t, h)$ and $B_2(t, h)$ are not explicit anymore even if $\widetilde{\lambda}_{2,t}$ is affine in $(y_t, y_{t-1}, d_{1,t}, d_{1,t-1})$.

Proposition 6.5 (Absence of contagion). *Under Hypotheses 6.1, 6.2, 6.3 and 6.5, the risk-neutral (\mathbb{Q}) dynamics is such that:*

$$f^{\mathbb{Q}}(y_t|w_{t-1}) = f^{\mathbb{Q}}(y_t|y_{t-1}) \propto \mathcal{M}_{t-1,t}(y_t, y_{t-1})f(y_t|y_{t-1})$$

(exogeneity of y_t preserved under \mathbb{Q}) and

$$p_t^{\mathbb{Q}}(d_t|y_t, w_{t-1}) = p_1(d_{1,t}|y_t, y_{t-1}, d_{1,t-1}) \times p_2(d_{2,t}|y_t, y_{t-1}, d_{2,t-1}),$$

(absence of contagion preserved under \mathbb{Q}).

Denoting by $\lambda_e^{\mathbb{Q}}(y_t, y_{t-1}) \equiv -\log[p_e^{\mathbb{Q}}(0|y_t, y_{t-1}, 0)]$ the risk-neutral default intensity of entity e , it comes that:

$$\lambda_e^{\mathbb{Q}}(y_t, y_{t-1}) = \lambda_e(y_t, y_{t-1})$$

(default intensities are the same under \mathbb{P} and \mathbb{Q}).

Proof. The \mathbb{Q} conditional density of w_t given w_{t-1} , namely $f^{\mathbb{Q}}(y_t|w_{t-1})p^{\mathbb{Q}}(d_t|y_t, w_{t-1})$, is proportional to $\mathcal{M}_{t-1,t}(w_t, w_{t-1})f(y_t|w_{t-1})p(d_t|y_t, w_{t-1})$, and thus proportional to $\mathcal{M}_{t-1,t}(y_t, y_{t-1})f(y_t|y_{t-1})p_1(d_{1,t}|y_t, y_{t-1}, d_{1,t-1})p_2(d_{2,t}|y_t, y_{t-1}, d_{2,t-1})$. The result follows immediately. \square

6.3 Pricing illiquid bonds

This subsection shows that a similar approach can be used to price illiquid assets. The idea consists in replacing the Loss Given Default ($LGD = 1 - RR^{(e)}$) by a Loss Given Liquidity Shock ($1 - \theta_\ell^{(e)}$). A structural interpretation of this approach is provided by Ericsson and Renault (2006) (see Example 6.1).

Let us introduce a novel binary variable: ℓ_t , that represents a *liquidity shocks*. When a bondholder is hit by the liquidity shock, she needs to liquidate her portfolio (reflecting cash constraints or needs to rebalance portfolio). But in such distressed conditions, she will sell her bonds at a discount (and will

therefore face losses, akin to credit losses). Formally: $\ell_t = 1$ when the bondholder is hit, and $\ell_t = 0$ otherwise.

The *liquidity-shock intensity* $\lambda_\ell(y_t)$ —that defines the probability of occurrence of liquidity shocks—is defined through:

$$\underbrace{1 - \exp[-\lambda_\ell(y_t)]}_{\approx \lambda_\ell(y_t) \text{ if small}} = \mathbb{P}(\ell_t = 1 | \underline{w}_{t-1}, y_t).$$

The binary variable ℓ_t is assumed not to Granger-cause y_t (as in Hypothesis 6.2) and not to be priced (as in Hypothesis 6.3). As in Proposition 6.2 (and more precisely Eq. (6.5)), this implies that

$$\lambda_\ell^{\mathbb{Q}}(y_t) = \lambda_\ell(y_t),$$

where

$$1 - \exp[-\lambda_\ell^{\mathbb{Q}}(y_t)] = \mathbb{Q}(\ell_t = 1 | \underline{w}_{t-1}, y_t).$$

Upon the arrival of the liquidity shock ($\ell_t = 1$), the investor has to exit by selling her bond holdings. She is then limited in the number of traders she can call and, as a result, get only a fraction ($\theta_\ell^{(e)}$, say) of the price she would get in normal-liquidity times, i.e., when $\ell_t = 0$ (see the structural interpretation in Example 6.1).

In that context, following the approach used in Proposition 6.3, one can define a *pseudo-illiquidity intensity* as follows (mimicking Eq. (6.14)):

$$\begin{aligned} \tilde{\lambda}_{e,\ell,t}(y_t, y_{t-1}) &= -\log \left\{ \exp \left[-\lambda_{e,\ell}(y_t) \right] + \right. \\ &\quad \left. \left(1 - \exp \left[-\lambda_{e,\ell}(y_t) \right] \right) \theta_\ell^{(e)} \right\} \end{aligned} \quad (6.21)$$

Still consiering credit risk (on top of liquidity risk), we then have the following payoff on date t (assuming no default on date $t-1$):

$$\tilde{B}_e(t, h) = \mathbb{E}_t^{\mathbb{Q}} \left\{ \exp \left[-\sum_{i=1}^h (r_{t+i-1} + \tilde{\lambda}_{e,c,t+i} + \tilde{\lambda}_{e,\ell,t+i}) \right] \right\}. \quad (6.22)$$

Example 6.1 (Structural interpretation of the illiquidity intensity). This structural interpretation of the illiquidity intensity is due to Ericsson and Renault (2006).

When hit by a liquidity shock, a bondholder has to liquidate her bond in a small amount of time (between t and $t + \varepsilon$, with $\varepsilon \ll 1$). She collects offers on the market; each offer is a random fraction ω_i ($i \in \{1, \dots, \kappa\}$) of $\widetilde{B}_e(t, h)$ (see Eq. (6.22)), where the ω_i 's are uniformly distributed in $[0, 1]$.

At $t + \varepsilon$, the bond is sold to the trader that has offered the highest price. Formally, when $\ell_t = 1$, the selling price is:

$$\left(\max_{i \in \{1, \dots, \kappa\}} \omega_i \right) \tilde{P}_{t,h}^{(e)}.$$

Conditional on κ , the expectation of $\max_i(\omega_i)$ is $\kappa/(\kappa + 1)$. (Because the ω_i are i.i.d., the c.d.f. of $\max(\omega_1, \dots, \omega_\kappa)$ is $x \mapsto F(x)^\kappa$, where F is the c.d.f. of ω_i .) Hence, $\theta_\ell^{(e)}$ —the unconditional expectation of $\max_i(\omega_i)$ —is given by:

$$\begin{aligned} \sum_{k=1}^{\infty} \frac{e^{-\varepsilon\gamma^{(e)}} k (\varepsilon\gamma^{(e)})^k}{k+1} &= \left[\left(1 - e^{-\varepsilon\gamma^{(e)}}\right) \frac{\varepsilon\gamma^{(e)} - 1}{\varepsilon\gamma^{(e)}} + e^{-\varepsilon\gamma^{(e)}} \right] \\ &= g(\varepsilon\gamma^{(e)}) \quad (\text{say}). \end{aligned}$$

Function g is increasing and valued in $[0, 1]$. In addition $\lim_{x \rightarrow \infty} g(x) \rightarrow 1$.

Example 6.2 (Euro area sovereign credit spreads). Monfort and Renne (2011) and Monfort and Renne (2014) propose credit/liquidity decompositions of euro-area sovereign spreads. They employ different estimation approaches: Kim (1994)'s filter for MR2014, and Kitagawa-Hamilton filter for Monfort and Renne (2011) (see Subsection 4.5.1).

In both papers, the y_t part of the state vector w_t is of the form $y_t = (x'_t, z'_t)'$; it follows a Regime-Switching Gaussian VAR (see Example 4.3). The time-homogenous Markov chain z_t tracks the *crisis status* (crisis versus normal times).

The total intensity is of the form:

$$\tilde{\lambda}_{e,t} = \underbrace{\alpha'_{e,c} z_t + \beta'_{e,c} y_t}_{\text{credit intensity } \tilde{\lambda}_{e,c,t}} + \underbrace{\alpha'_{e,\ell} z_t + \beta'_{e,\ell} y_t}_{\text{liquidity intensity } \tilde{\lambda}_{e,\ell,t}}.$$

In both papers, the identification of credit versus liquidity components is based on:

$$\text{KfW-Bund spreads} = \text{liquidity spreads}.$$

KfW (*Kreditanstalt für Wiederaufbau*) is a German agency whose bonds are guaranteed by the Federal Republic of Germany. Therefore, they benefit from the same credit quality than their sovereign counterparts—the Bunds—but are less liquid; accordingly, the KfW-Bund spread should be essentially liquidity-driven.

In Monfort and Renne (2014), we have

$$\tilde{\lambda}_{e,t} = \underbrace{x_{e,t}}_{=\tilde{\lambda}_{e,c,t}} + \underbrace{(1 - \theta_\ell^{(e)})x_{E+1,t}}_{=\tilde{\lambda}_{e,\ell,t}},$$

hence x_t of dimension $E + 1 = 12$ (11 countries are involved).

There are 9 Markovian regimes: $z_t = z_{c,t} \otimes z_{\ell,t}$, where $z_{i,t}$ ($i \in \{c, \ell\}$) is 3-dimensional:

- $z_{i,t} = [1, 0, 0]'$: low stress,
- $z_{i,t} = [0, 1, 0]'$: medium stress,
- $z_{i,t} = [0, 0, 1]'$: high stress.
- $z_{c,t}$ and $z_{\ell,t}$ are not independent: for instance, the probability to transit from the low-credit-stress regime to the medium/high-credit-stress regimes depends on the liquidity-stress regime.

For each intensity $x_{i,t}$ (simulation below XXX):

$$x_{i,t} = \mu'_i z_t + \phi x_{i,t-1} + \sigma_i \varepsilon_{i,t},$$

with $\varepsilon_t \sim i.i.d. \mathcal{N}(0, \Sigma)$, where Σ is diagonal, and $\mu_{i,1} < \mu_{i,2} < \mu_{i,3}$. The Markov chain z_t is the only source of correlation across the $x_{i,t}$'s.

Figure 6.2 shows a simulation of such a process, with:

$$\mu = \begin{bmatrix} 0.01 \\ 0.03 \\ 0.10 \end{bmatrix}, \quad \phi = .8, \quad P = \begin{bmatrix} 0.98 & 0.02 & 0 \\ 0.05 & 0.90 & 0.05 \\ 0 & 0.20 & 0.80 \end{bmatrix}, \quad \sigma = 0.002.$$

The SDF exponential affine in $(z'_t, x'_t)'$. We have the same type of dynamics under \mathbb{Q} (with risk-adjusted parameters, in particular risk-adjusted transition probabilities).

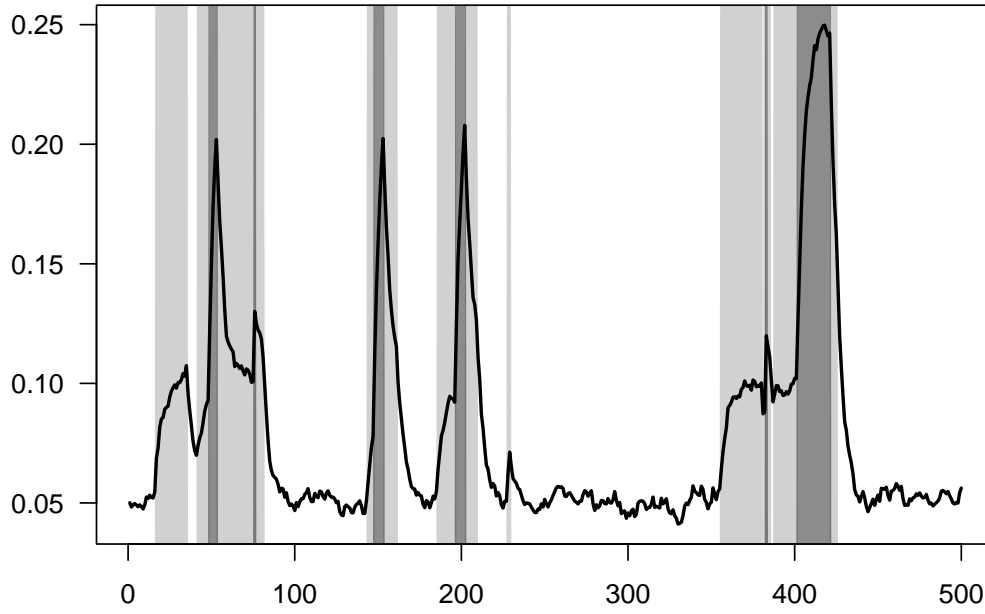


Figure 6.2: Simulation of an process following $x_t = \mu' z_t + \phi x_{t-1} + \sigma \varepsilon_t$, where z_t follows a three-state time-homogenous Markov process. See the text for the exact parameterization. The light-grey shaded area corresponds to the second regime ($z_t = [0, 1, 0]'$); the dark-grey shaded area corresponds to the third regime ($z_t = [0, 0, 1]'$)

Example 6.3 (IBOR-OIS spreads). Dubecq et al. (2016) decompose the EURIBOR versus Overnight Index Swap (OIS) spreads into a credit and a liquidity component. (See Subsection ?? for the definition of a swap contract.)

From the point of view of bank j , the EURIBOR rate is the rate of an unsecured loan to one bank (e , say) of the panel. Assume that the panel of the N constituting the EURIBOR are homogenous. We have:

$$R_{t,h}^{IBOR} = -\frac{1}{h} \log \mathbb{E}_t^{\mathbb{Q}} \left\{ \exp \left[-\sum_{i=1}^h (r_{t+i-1} + \tilde{\lambda}_{e,c,t+i} + \tilde{\lambda}_{j,\ell,t+i}) \right] \right\}. \quad (6.23)$$

Note that the liquidity intensity refers to the lending bank (j) and the default intensity relates to the borrowing bank (e).

By contrast, the Overnight-Index Swap (OIS) rate satisfies:¹

$$R_{t,h}^{OIS} = -\frac{1}{h} \log \mathbb{E}_t^{\mathbb{Q}} \left\{ \exp \left(-\sum_{i=1}^h r_{t+i-1} \right) \right\}. \quad (6.24)$$

Assuming that r_t , the short-term risk-free rate (i.e., the reference rate of the OIS, EONIA here) is independent from the intensities and under the homogeneity assumption, we obtain:²

$$R_{t,h}^{IBOR} - R_{t,h}^{OIS} = -\frac{1}{h} \log \mathbb{E}_t^{\mathbb{Q}} \left\{ \exp \left[-\sum_{i=1}^h (\tilde{\lambda}_{c,t+i} + \tilde{\lambda}_{\ell,t+i}) \right] \right\}.$$

Dubecq et al. (2016) employ a quadratic specification for $\tilde{\lambda}_{c,t}$ and $\tilde{\lambda}_{\ell,t}$:

$$\tilde{\lambda}_{c,t} = x_{c,t}^2 \quad \text{and} \quad \tilde{\lambda}_{\ell,t} = x_{\ell,t}^2,$$

¹The OIS is an interest-rate swap where the floating rate is an overnight-rate reference (EONIA in the euro area). At maturity, the payoff received by the fixed-rate payer is the difference between: (a) the notional (W , say) inflated with the date- t OIS (fixed) rate (i.e. $W \exp\{h R_{t,h}^{OIS}\}$) and (b) the same notional capitalized with the realized short-term rates (i.e. $W \exp\{r_t + \dots + r_{t+h-1}\}$).

²It is assumed here that the overnight interbank market preserves the lending bank from (i) liquidity and (ii) credit risk. The rationale behind that is twofold: (a) By rolling its cash on the overnight market (at the EONIA rate), a bank is not exposed to the risk of having to liquidate longer-term investments upon the realization of the liquidity shock; (b) while the EONIA is an unsecured-transaction rate, the extremely-short maturity of these transactions substantially reduces the credit-risk exposure of the lending bank.

where $X_t = (x_{c,t}, x_{\ell,t})'$ follows a Gaussian VAR(1) process. (This is the quadratic Gaussian framework presented in Example 1.8.)

The state vector X_t is latent. The estimation involves the Quadratic Kalman Filter (Monfort et al., 2015). The measurement equations include equations stating that observed proxies of credit and liquidity risks relate to quadratic functions of $x_{c,t}$ and $x_{\ell,t}$, respectively.

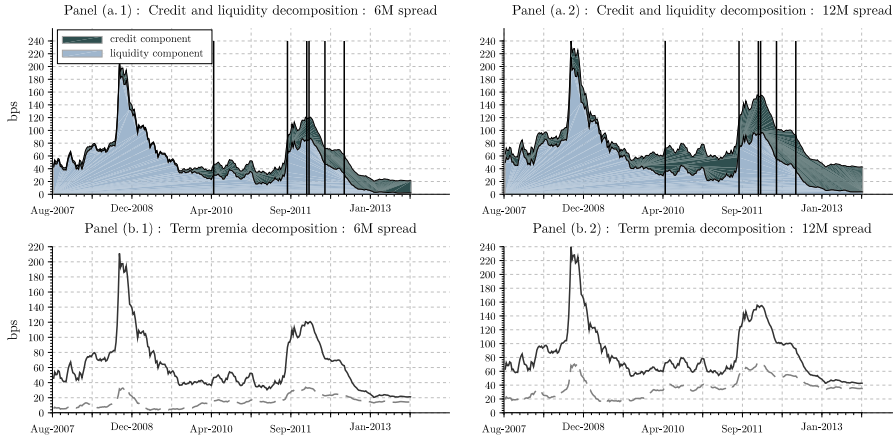


Fig. 4. 6M EURIBOR-OIS spreads decomposition. Notes: Date ranges from August 31 2007 to September 13 2013. Units are in basis points. Panels (a) represent the stacked components of the spread: light gray component is the liquidity component and the dark gray corresponds to the credit component. Panels (b) represent the modeled spread (black) and its term premia (gray dashed). The black vertical axes stand from left to right for: SMP program announcements (first two axes), VLTRO announcement and allotments (next three axes), and Mario Draghi's London speech (last axis).

Figure 6.3: Source: Dubecq et al. (2016).

6.4 Relaxing the classical framework assumptions

Formula (6.19) and its tractability is a key feature of numerous credit-risk term-structure models, e.g., Duffie and Singleton (2003), Pan and Singleton (2008), and Longstaff et al. (2011) among many others. This formula is valid under Assumptions 6.2 of non-systemic entities, also called *no-jump condition*, 6.3 of unpriced credit risk events, and 6.5 absence of contagion. These hypotheses define what we have called the *classical credit-risk models*. We lose the high degree of tractability of Formula (6.19) when some of the previous assumptions are relaxed:

- If $f(y_t|w_{t-1}) = f(y_t|y_{t-1}, d_{1,t-1})$ (entity $e = 1$ is *systemic*) then:
 - The pseudo-price $\widetilde{B}_e(t, h)$ now depends on y_t and $d_{1,t}$;
 - Eq. (6.6) is valid for entity $e = 2$ only;
 - The computation of $\widetilde{B}_2(t, h)$ is not straightforward: $\{y_t\}$ is not autonomous and the autonomous process $\{y_t, d_{1,t}\}$ is not affine.

- If there is a contagion effect from $e = 1$ towards $e = 2$, i.e., if

$$\begin{aligned} p(d_t|y_t, w_{t-1}) &= p_1(d_{1,t}|y_t, y_{t-1}, d_{1,t-1}) \\ &\quad p_2(d_{2,t}|y_t, y_{t-1}, d_{1,t}, d_{1,t-1}, d_{2,t-1}), \end{aligned}$$

then:

- Eq. (6.6) is still valid for both entities;
- In an affine framework, the computation of $\widetilde{B}_1(t, h)$ (and $B_1(t, h)$) is straightforward; but formulas for $\widetilde{B}_2(t, h)$ and $B_2(t, h)$ are not explicit anymore even if $\widetilde{\lambda}_{2,t}$ is affine in $(y_t, y_{t-1}, d_{1,t}, d_{1,t-1})$.
- If the default event of the first entity only ($d_{1,t}$) is a source of risk that is priced: $\mathcal{M}_{t-1,t}(w_t, w_{t-1}) = \mathcal{M}_{t-1,t}(y_t, y_{t-1}, d_{1,t}, d_{1,t-1})$, then
 - $\lambda_{1,t}^{\mathbb{Q}} \neq \lambda_{1,t}$ and $\lambda_{2,t}^{\mathbb{Q}} = \lambda_{2,t}$;
 - the exogeneity of $\{y_t\}$ is no longer preserved under \mathbb{Q} ;
 - Eq. (6.6) is no longer valid for entity $e = 1$. It remains valid for $e = 2$ but the computation of $\widetilde{B}_2(t, h)$ is not straightforward: even if y_t is \mathbb{P} -autonomous, this is not true under \mathbb{Q} .

6.4.1 General affine credit-risk framework

Exploiting Vector Auto-Regressive Gamma (VARG) processes, Monfort et al. (2021) propose a general affine credit-risk pricing model jointly allowing for:

1. *systemic entities* (breaking down the no-jump condition; see Collin-Dufresne et al. (2004));
2. *contagion effects between entities* (economic/financial linkages; see Ait-Sahalia et al. (2014));

3. *pricing of credit events* (credit spread puzzle; see Gouriou et al. (2014b));
4. and *stochastic recovery rates (RR)* (Altman et al., 2005).

In this general framework, the state vector w_t is of the form $[y'_t, \delta'_t]'$, where δ_t is a E -dimensional vector of *credit-event variables* (see Hypothesis 6.6).

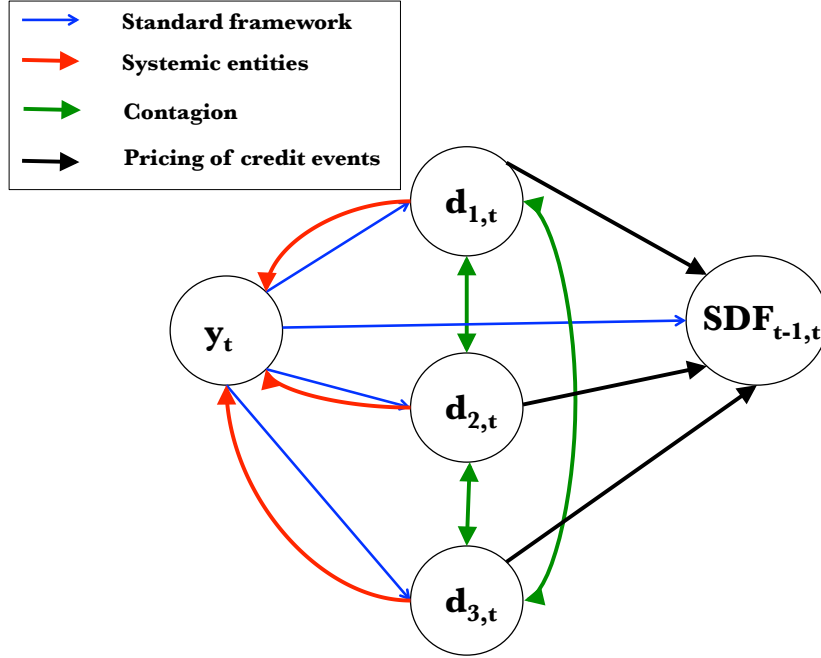


Figure 6.4: Schematic comparison of the classical and general credit-risk frameworks.

Hypothesis 6.6 (Credit events). The default date $\tau^{(e)}$ (say) of entity e is defined as:

$$\tau^{(e)} = \inf \{t > 0 : \delta_t^{(e)} > 0\}, \quad (6.25)$$

where $\delta_t^{(e)}$ is a non-negative variable called *credit-event variable*. The default indicator function can be equivalently written as $d_t^{(e)} = \mathbf{1}_{\{\tau^{(e)} \leq t\}}$ or $d_t^{(e)} = 1 - \mathbf{1}_{\{\delta_t^{(e)'} \mathbf{1} = 0\}}$, with $\delta_t^{(e)} = (\delta_t^{(e)}, \dots, \delta_1^{(e)})$ and where $\mathbf{1} = (1, \dots, 1)'$ with conformable dimension.

Hypothesis 6.7 (State-vector dynamics). The stochastic process $\{w_t\}$ is affine under the historical probability measure \mathbb{P} . The historical Laplace transform of w_t , conditional \underline{w}_{t-1} , is denoted by:

$$\begin{aligned}\varphi_{w,t-1}^{\mathbb{P}}(u_w) &= \mathbb{E} \left[\exp(u'_w w_t) | \underline{w}_{t-1} \right] \\ &= \exp [a_w(u_w)' w_{t-1} + b_w(u_w)],\end{aligned}\quad (6.26)$$

with $u_w = (u'_y, u'_\delta)'$.

Example 6.4 (Vector Autoregressive Gamma process). In Monfort et al. (2021), w_t follows a VARG (positive affine) process, which satisfies Hypthesis 6.7 (Example 1.9 presents the univariate version of this process.):

$$\begin{aligned}y_t | \underline{w}_{t-1} &\stackrel{\mathbb{P}}{\sim} \otimes_{j=1}^{N_y} \gamma_{\nu_j^{(y)}} \left(\alpha_j^{(y)} + \overbrace{\beta_{j,y}^{(y)} y_{t-1}}^{\text{factors}} + \underbrace{\beta_{j,\delta}^{(y)} \delta_{t-1}}_{\substack{\text{kills no-jump} \\ \text{condit.}}}, \mu_j^{(y)} \right) \\ \delta_t | y_t, \underline{w}_{t-1} &\stackrel{\mathbb{P}}{\sim} \otimes_{e=1}^E \gamma_0 \left(\alpha_e^{(\delta)} + \beta_{e,y}^{(\delta)} y_t + \underbrace{\beta_{e,\delta}^{(\delta)} \delta_{t-1}}_{\text{contagion}}, \mu_e^{(\delta)} \right),\end{aligned}\quad (6.27)$$

where $\delta_t = (\delta_t^{(1)}, \dots, \delta_t^{(E)})'$ and $\gamma_\nu(\lambda, \mu)$: non-central Gamma distribution.

Hypothesis 6.8 (Stochastic discount factor). The one-period positive SDF $\mathcal{M}_{t-1,t}$ is given by (this is Eq. (2.5), with $\alpha_t \equiv \alpha_w$):

$$\mathcal{M}_{t,t+1} = \exp[-r_t + \alpha'_w w_{t+1} - \psi_t(\alpha_w)], \quad (6.28)$$

and the risk-free short rate (between t and $t+1$) is given by the following affine function of the factors:

$$(\#eq : short_rate) \ r_t(\underline{w}_t) = \xi_0 + \xi'_1 w_t. \quad (6.29)$$

$\alpha_w = (\alpha'_y, \alpha'_x, \alpha'_\delta)'$ is the vector of *prices of risk*.

Under Hypotheses 6.7 and 6.8, it comes that:

$$\begin{aligned}\varphi_{w,t-1}^{\mathbb{Q}}(u_w) &= \mathbb{E}^{\mathbb{Q}} \left[\exp(u'_w w_t) | \underline{w}_{t-1} \right] \\ &= \exp [a_w^{\mathbb{Q}}(u_w)' w_{t-1} + b_w^{\mathbb{Q}}(u_w)],\end{aligned}\quad (6.30)$$

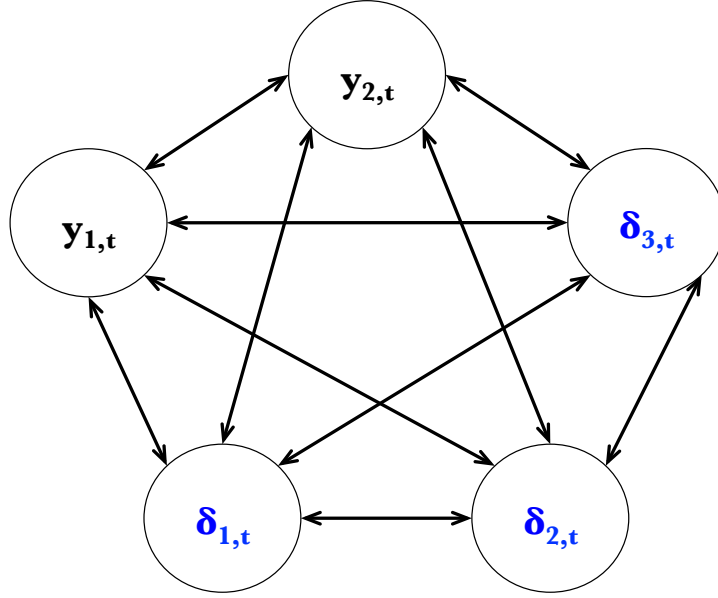


Figure 6.5: A potential causality scheme in Monfort et al. (2021).

where, using Eq. (2.7):

$$\begin{cases} a_w^Q(u_w) &= a_w(u_w + \alpha_w) - a_w(\alpha_w) \\ b_w^Q(u_w) &= b_w(u_w + \alpha_w) - b_w(\alpha_w). \end{cases} \quad (6.31)$$

Hypothesis 6.9 (General recovery payment). The *Recovery Payment* of a defaultable ZCB, in the case of default at date $t + i = \tau^{(e)}$, is given by:

$$RR_{t+i}^{(e)} \times \mathcal{V}_{t+i,h-i}^{(e)} \quad (6.32)$$

where the *Recovery Rate* $RR_{t+i}^{(e)}$ is

$$RR_{t+i}^{(e)} = \exp(-a_e - a'_{w,e} w_{t+i}), \quad (6.33)$$

and where $\mathcal{V}_{t+i,h-i}^{(e)}$ denotes the *Recovery Value* (Exposure-at-Default) at $t + i$.

The general framework allows for flexible specifications of the recovery payment. It can be:

- the fraction of the pre-default value of the claim (Recovery of Market Value, or RMV) [see Prop. 6.7];
- the fraction of par (Recovery of Face Value, or RFV) [see Proposition 6.6];
- the fraction of a no-default version of the same claim (Recovery of Treasury, or RT). In that case, the recovery payment is a fraction of the (risk-free) present value of the principal.

Proposition 6.6 (General pricing under the RFV convention). *The price, at date $t < \tau^{(e)}$, of a ZCB issued by entity e and maturing in h periods is given by:*

$$B_e(t, h) = \sum_{i=1}^h \left(\Lambda_{(1,t,i)}^{\mathbb{Q}} - \Lambda_{(2,t,i)}^{\mathbb{Q}} \right) + \Lambda_{(3,t,h)}^{\mathbb{Q}}, \quad (6.34)$$

where:

$$\begin{aligned} \Lambda_{(1,t,i)}^{\mathbb{Q}} &:= \lim_{u \rightarrow -\infty} \Psi_{(t,i)}^{\mathbb{Q}}(a_e, u\tilde{e}_\delta - \xi_1, -a_{w,e}) \\ \Lambda_{(2,t,i)}^{\mathbb{Q}} &:= \lim_{u \rightarrow -\infty} \Psi_{(t,i)}^{\mathbb{Q}}(a_e, u\tilde{e}_\delta - \xi_1, u\tilde{e}_\delta - a_{w,e}) \\ \Lambda_{(3,t,i)}^{\mathbb{Q}} &:= \lim_{u \rightarrow -\infty} \Psi_{(t,i)}^{\mathbb{Q}}(0, u\tilde{e}_\delta - \xi_1, u\tilde{e}_\delta) \end{aligned} \quad (6.35)$$

with $u \in \mathbb{R}$ and where [denoting $\varphi_{w,t,i}^{\mathbb{Q}}(u_2, \dots, u_2, u_1) = \varphi_{w,t,i}^{\mathbb{Q}}(u_2, u_1)$]:

$$\Psi_{(t,i)}^{\mathbb{Q}}(\kappa, u_1, u_2) := \exp[-i\xi_0 + \kappa + u'_2 w_t] \varphi_{w,t,i}^{\mathbb{Q}}(u_2, u_1) \quad (6.36)$$

Proof. See Monfort et al. (2021). \square

Proposition 6.7 (General pricing under the RMV convention). *If the recovery value at date $t + i$ (defined in 6.9) is of the form:*

$$\mathcal{V}_{t+i,h-i}^{(e)} = \mathbb{E}^{\mathbb{Q}} \left\{ \exp \left[- \sum_{j=i}^{h-1} (r_{t+j} + \delta_{t+j+1}^{(e)}) \right] \middle| \underline{w}_{t+i} \right\}, \quad (6.37)$$

with $RR_t^{(e)} = \exp(-\delta_t^{(e)})$, and under Hypothesis 6.6 to 6.8, the price $B_e(t, h)$ at date $t < \tau^{(e)}$ is given by:

$$B_e(t, h) = \mathcal{V}_{t,h}^{(e)}, \quad (6.38)$$

where

$$\mathcal{V}_{t,h}^{(e)} = \exp \left[(\mathcal{A}_h - \xi_1)' w_t + (\mathcal{B}_h - h\xi_0) \right]. \quad (6.39)$$

where \mathcal{A}_h and \mathcal{B}_h are obtained recursively by employing Eqs. (1.19) of Proposition 1.5, replacing functions a and b with $a^{\mathbb{Q}}$ and $b^{\mathbb{Q}}$ (see Eq. (6.31)) with $u_1 = -\tilde{e}_\delta$ and, for $i > 1$, $u_i = -(\tilde{e}_\delta + \xi_1)$ where $\tilde{e}_\delta = (0', e'_\delta)'$ is a N -dimensional vector, and where e_δ is the e^{th} column of the (N_δ, N_δ) -dimensional identity matrix.

Proof. See Appendix A.4 of Monfort et al. (2021). \square

Eq. (6.38) is a key result. It reads:

$$B_e(t, h) = \mathbb{E}^{\mathbb{Q}} \left\{ \exp \left[- \left(\sum_{i=1}^h r_{t+i-1} + \delta_{t+i}^{(e)} \right) \right] \middle| \underline{w}_t \right\}. \quad (6.40)$$

It is reminiscent of Eq. (6.6), with λ (default intensity) replaced with δ (credit-event variable). However, contrary to the classical framework, credit events are priced sources of risk, the no-jump condition is relaxed, contagion is allowed, and the recovery rate is stochastic. This formula can therefore be seen as a generalization of the setting of Duffie and Singleton (1999).

6.4.2 CDS pricing in the general framework

This subsection shows how CDS can be priced in the context of the general credit-risk model. We consider two types of CDS. The first is the standard one; the second is a CDS whose payoffs are expressed in a foreign currency (which is typical to sovereign CDSs). Figures 6.6 and 6.7 show the payoffs associated with these two types of CDSs.

Since the domestic-currency case is a special case of the multi-currency one, we focus on the latter case in the following.

Pricing a CDS amounts to determining $\mathcal{S}_{t,t+h}^{(e)}$, that is the *CDS spread*, for $t < \tau^{(e)}$. The spread $\mathcal{S}_{t,t+h}^{(e)}$ is such that the date- t value of the *fixed leg*'s payoffs is equal to the date- t value of the *floating leg*'s payoffs. The payoffs of the assets issued by the reference entity are assumed to be expressed in the domestic currency.

The (logarithmic) exchange rate is denoted by $s_t = \ln(FX_t)$. By convention, an increase in s_t corresponds to a depreciation of the domestic currency. (In

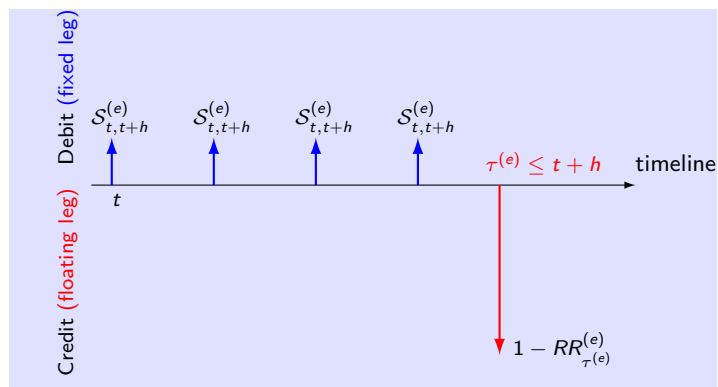


Figure 6.6: CDS payoffs.

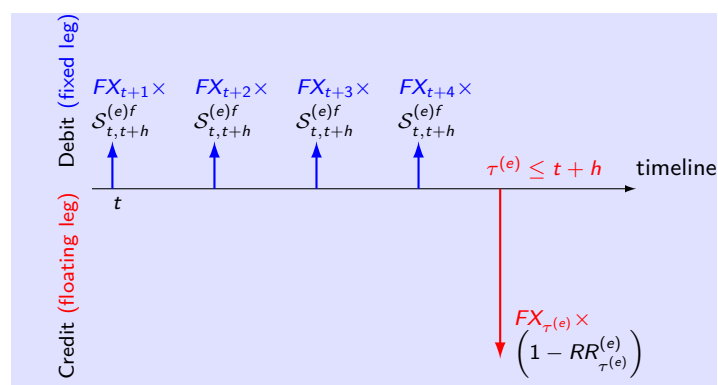


Figure 6.7: CDS payoffs when the payoffs are expressed in a foreign currency.

other words, FX_t is the value of one unit of foreign currency expressed in the domestic currency.) We consider a CDS whose notional is equal to one unit of the foreign currency (i.e. to $\exp(s_t)$ units of the domestic currency).

Let us determine the value of the *fixed leg*. If entity e has not defaulted at date $t + i$ ($\leq t + h$), the cash flow on this date, expressed in the domestic currency, is:

$$\mathcal{S}_{t,t+h}^{(e)f} \exp(s_{t+i}).$$

Hence, the present values of the fixed-leg payments, expressed in the domestic currency are:

$$\mathcal{S}_{t,t+h}^{(e)f} \sum_{i=1}^h \mathbb{E}^{\mathbb{Q}} \left[\exp \left(s_{t+i} - \sum_{j=1}^i r_{t+j-1} \right) \mathbf{1}_{\{\delta_{t:t+i}^{(e)'} \mathbf{1}=0\}} \middle| \underline{w}_t \right],$$

Let us turn to the *floating leg*. Under the RFV convention, the protection seller will make a payment of $(1 - RR_{t+i}^{(e)}) \exp(s_{t+i})$ (this is the loss-given-default. LGD) at date $t + i$ in case of default over the time interval $]t + i - 1, t + i]$. The present values of the floating leg, expressed in the domestic currency, is:

$$\sum_{i=1}^h \mathbb{E}^{\mathbb{Q}} \left[\exp \left(s_{t+i} - \sum_{j=1}^i r_{t+j-1} \right) (1 - RR_{t+i}^{(e)}) \left(\mathbf{1}_{\{\delta_{t:t+i-1}^{(e)'} \mathbf{1}=0\}} - \mathbf{1}_{\{\delta_{t:t+i}^{(e)'} \mathbf{1}=0\}} \right) \middle| \underline{w}_t \right],$$

We assume that $RR_t^{(e)} = \exp(-a_e - a'_{w,e} w_t)$ and that $s_t = \chi + u'_s w_t$.

Proposition 6.8 (Price of a multi-currency CDS). *In the context described above, we have:*

$$\boxed{\mathcal{S}_{t,t+h}^{(e)f} = \frac{\sum_{i=1}^h \Lambda_{(t,i)}^{\mathbb{Q}}}{\sum_{i=1}^h \lim_{u \rightarrow -\infty} \Psi_{(t,i)}^{\mathbb{Q}}(\chi, u\tilde{e}_\delta - \xi_1, u\tilde{e}_\delta + u_s)},} \quad (6.41)$$

where:

$$\begin{aligned} \Lambda_{(t,i)}^{\mathbb{Q}} = & \lim_{u \rightarrow -\infty} \left[\Psi_{(t,i)}^{\mathbb{Q}}(\chi, u\tilde{e}_\delta - \xi_1, u_s) - \Psi_{(t,i)}^{\mathbb{Q}}(\chi, u\tilde{e}_\delta - \xi_1, u\tilde{e}_\delta + u_s) \right. \\ & - \Psi_{(t,i)}^{\mathbb{Q}}(\chi - a_e, u\tilde{e}_\delta - \xi_1, u_s - a_w) \\ & \left. + \Psi_{(t,i)}^{\mathbb{Q}}(\chi - a_e, u\tilde{e}_\delta - \xi_1, u\tilde{e}_\delta + u_s - a_w) \right], \end{aligned}$$

and where $\Psi_{(t,i)}^{\mathbb{Q}}(\kappa, u_2, u_1)$ is given in Eq. (6.36).

Proof. See Online Appendix A.4 of Monfort et al. (2021). The proof notably makes use of Lemma ?? (Subsection 1.3). \square

To price a standard CDS (with payments in domestic currency), one simply has to set $\chi = 0$ and $u_s = 0$, which gives $s_t = 0$ (since $s_t = \chi + u'_s w_t$).

Monfort et al. (2021) exploit these formulas to model quanto CDSs in the euro area. Quanto CDSs are spread differentials between the two types of CDS. They exploit this framework to estimate “depreciations upon default”. Indeed, in their model, s_t (the log EURUSD exchange rate) is affected by the credit event variables δ_t .

6.5 Top-down approach

- Relaxing the assumptions underlying the classical framework can also be done in the context of *top-down* approaches. Top-down models focus on default counting (or loss) processes (see, e.g., Azizpour et al. (2011) and Giesecke et al. (2011)), contrary to the (*bottom-up*) approaches presented above. The latter consider default processes of individual firms as the model primitives (e.g., Lando (1998), Duffie and Singleton (1999), Duffie and Garleanu (2001)).

The top-down approach has been shown to satisfactorily capture the existence of default clustering (e.g., Brigo et al. (2007), Errais et al. (2010)).

Building on Gourioux et al. (2014b), Gourioux et al. (2021) propose an affine top-down model consistent with: the presence of systemic entities, contagion, and the pricing of default events (see Example 6.5).

Example 6.5 (Disastrous Defaults). Gourioux et al. (2021) propose an equilibrium model where different credit derivatives, including CDSs and tranche products (iTraxx), can be priced. In their framework, the default of large firms (those included in the iTraxx index) can have systemic consequences in the sense that: (a) they result in decrease in consumption and (b) they increase the probability of default of the other systemic firms.

Let n_t^s denote the number of systemic defaults on date t , and let N_t^s denote number of systemic entities in default at date t ($N_t^s = n_t^s + N_{t-1}^s$).

The conditional distribution of the number of systemic defaults is given by:

$$n_{t+1}^s | \underline{x}_{t+1}, \underline{y}_{t+1}, \underline{N}_t^s \sim \text{Poisson}(\beta y_{t+1} + c n_t^s), \quad (6.42)$$

where y_t is a nonnegative factor fluctuating around a low-frequency component x_t . Vector $(y_t, x_t)'$ follows a VARG dynamics (Eq. (3.19)) that admits the following VAR representation (using Prop. 1.1):

$$\begin{cases} y_t - x_t &= \rho_y(y_{t-1} - x_{t-1}) + \sigma_{y,t}\varepsilon_{y,t} \\ x_t - \mu_x &= \rho_x(x_{t-1} - \mu_x) + \sigma_{x,t}\varepsilon_{x,t}, \end{cases} \quad (6.43)$$

with $0 < \rho_y < \rho_x < 1$.

If $c > 0$, defaults on date t increases the conditional probability of having additional defaults on the next date \Rightarrow Systemic defaults are infectious (Davis and Lo, 2001), or contagious.

The log growth rate of per capita consumption ($\Delta c_t = \log(C_t/C_{t-1})$) is given by:

$$\Delta c_t = \mu_{c,0} + \mu_{c,x}x_t + \mu_{c,y}y_t + \mu_{c,z}z_t, \quad (6.44)$$

where z_t depends on systemic defaults:

$$z_t | \underline{x}_t, \underline{y}_t, \underline{N}_t^s \sim \gamma_0(\xi_z n_{t-1}^s, \mu_z). \quad (6.45)$$

γ_0 being a distribution featuring a point mass at zero (see Eq. (1.3) in Example 1.9). In that context, the conditional probability that $z_t = 0$ is $\exp(-\xi_z n_{t-1}^s)$. $z_t = 0$ as long as there has been no systemic defaults in the previous period, which is rather frequent.

If $\mu_{c,z} < 0$ and $|\mu_{c,z}|$ is large or if c (contamination) is large, then systemic defaults can give rise to disastrous decreases in C_t .

Gourieroux et al. (2021) consider agents featuring Epstein-Zin preferences, with a unit elasticity of intertemporal substitution (EIS). In that context, the SDF is exponential affine (as in Eq. (2.5)), and $w_t = (x_t, y_t, z_t, n_t^s)'$ is affine under \mathbb{P} and \mathbb{Q} .

In this context, a wide range of credit derivatives can be priced. In particular those whose payoffs depend on the default status of the constituents of a reference portfolio: Credit Index swaps (CIS) and synthetic CDO.

- CDS: protection payoff > 0 , when the entity on which the CDS is written defaults.

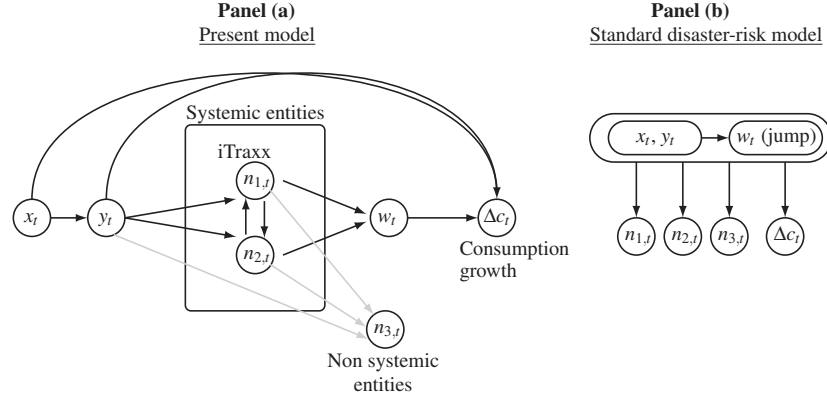


Figure 6.8: Causality scheme in Gourieroux et al. (2021).

- CIS: protection payoff > 0 , when one entity of the underlying portfolio defaults.
- CDO: protection payoff > 0 , when one entity of the underlying portfolio defaults, given that losses are in a given interval $[a, b]$ (e.g. $[a, b] = [3\%, 6\%]$). See Figure ??.

Typical credit indices are the iTraxx (Europe) and CDX (U.S.); these indices track the default status of 125 large firms.

The pricing CDO formula make an intensive use of the truncated Laplace transform (see Eq. (1.18), Duffie et al. (2000)). The model estimation would be infeasible without the tractability provided by affine processes.

Chapter 7

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