

# Limited Dependent Variable Models

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# Overview

# Packages and Files

## Required packages:

- AER
- censReg
- foreign
- MASS
- pscl
- stargazer
- survival
- survminer
- truncreg

## Required files:

```
data("NMES1988", package="AER")
```

Regression models in which the dependent variable is somehow limited:

- Truncated data: Values above and/or below particular points are not reported
- Censored data: Values above and/or below particular points are reported at those points
- Count data: Discrete, integer count value
- Survival/duration data: Time to a certain event

# Truncation

## Concept

- Value above and/or below a certain point are not part of the data

## Examples

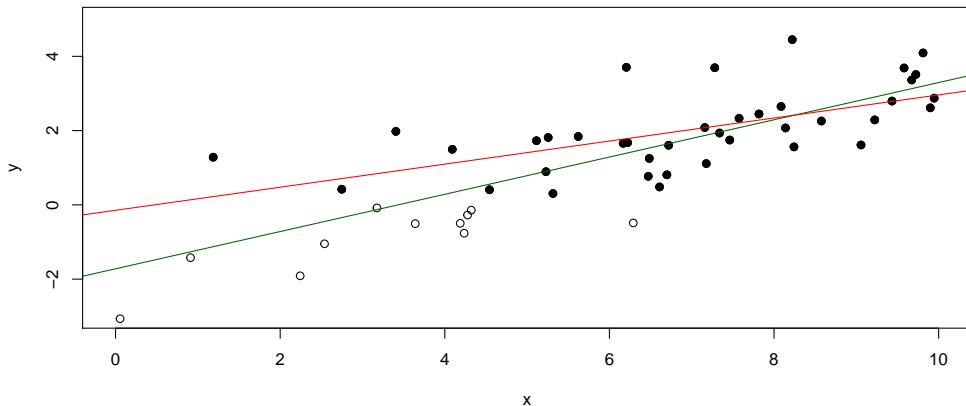
- Low income household studies
- On-site visitation data (unobserved non-visitors)
- Employment data on hours worked (excludes unemployed)

## Simulated data

- “True” Coefficients:  $\beta_0 = -2$  and  $\beta_1 = 0.5$
- Values  $y < 0$  are not reported in the data

Next slide: The green regression line is “correct” whereas the “red” is the line obtained from a regression model which ignores the truncation.

# Graphical Illustration



## Setup for truncation Data

```
truncation1      = truncation[c("y_real","x")]
truncation2      = subset(truncation,y_obs>0,select=c("y_obs","x"))
bhat_real        = lm(y_real~x,data=truncation1)
bhat_truncated   = lm(y_obs~x,data=truncation2)
```

Required package to estimate a truncated model

- `truncreg`

Additional variable output *sigma*:

- Related to the truncated normal distribution





# Results: Correcting for Truncation

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```
##
## Call:
## truncreg(formula = y_obs ~ x, data = truncation2, point = 0,
##          direction = "left")
##
## BFGS maximization method
## 22 iterations, 0h:0m:0s
## g'(-H)^-1g = 7.28E-09
##
##
##
## Coefficients :
##              Estimate Std. Error t-value Pr(>|t|)
## (Intercept) -0.716370   0.668897  -1.0710   0.2842
## x            0.378806   0.086811   4.3636 1.279e-05 ***
## sigma        0.889160   0.119117   7.4646 8.349e-14 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Log-Likelihood: -46.315 on 3 Df
```

# Achievement Scores: Data Load and Description

Loading the data using the package `foreign`

```
url          = "https://stats.idre.ucla.edu/stat/data/truncreg.dta"  
achievement = read.dta(url)
```

Description of the data from [UCLA Source](#):

*"A study of students in a special GATE (gifted and talented education) program wishes to model achievement as a function of language skills and the type of program in which the student is currently enrolled. A major concern is that students are required to have a minimum achievement score of 40 to enter the special program. Thus, the sample is truncated at an achievement score of 40."*

# Achievement Scores: Regular OLS Estimation

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```
##
## Call:
## lm(formula = achiv ~ langscore + prog, data = achievement)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -16.9413  -5.7033  -0.8462   5.2205  21.3010
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  27.63965    3.70639   7.457 4.01e-12 ***
## langscore     0.46319    0.06792   6.820 1.45e-10 ***
## progacademic  2.97343    1.44889   2.052  0.0416 *
## progvocacion -0.52118    1.72739  -0.302  0.7632
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7.534 on 174 degrees of freedom
## Multiple R-squared:  0.3054, Adjusted R-squared:  0.2934
## F-statistic: 25.5 on 3 and 174 DF, p-value: 1.01e-13
```

# Achievement Scores: Truncated Model

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```
##
## Call:
## truncreg(formula = achiv ~ langscore + prog, data = achievement,
##          point = 40, direction = "left")
##
## BFGS maximization method
## 57 iterations, 0h:0m:0s
## g'(-H)^-1g = 2.5E-05
##
##
## Coefficients :
##              Estimate Std. Error t-value Pr(>|t|)
## (Intercept)  11.29942     6.77173   1.6686  0.09519 .
## langscore     0.71267     0.11446   6.2264 4.773e-10 ***
## progacademic  4.06267     2.05432   1.9776  0.04797 *
## progvocaton -1.14422     2.66958  -0.4286  0.66821
## sigma         8.75368     0.66647  13.1343 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Log-Likelihood: -591.31 on 5 Df
```

# Censoring

## Concept

- Value above and/or below a certain point are not part of the data

## Examples

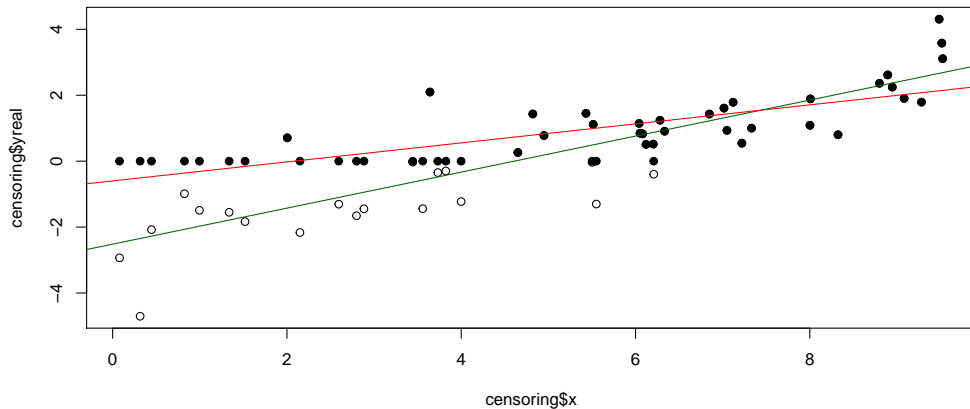
- Capacity constrained data, e.g., class enrollments or ticket sales
- Hours worked (or leisure demand), which is essentially capacity constrained
- Commodity purchases (non-negative)

## Simulated data

- “True” Coefficients:  $\beta_0 = -2$  and  $\beta_1 = 0.5$
- Values  $y < 0$  are reported at 0

R package [censReg](#) to reduce bias

## Graphical Illustration





```
## =====  
##                                     Dependent variable:  
## -----  
##                                yreal           y  
##                               (1)            (2)  
## -----  
## x                            0.546***      0.288***  
##                              (0.047)       (0.035)  
## Constant                     -2.518***    -0.599***  
##                              (0.282)       (0.212)  
## -----  
## Observations                  50             50  
## R2                           0.737          0.579  
## Adjusted R2                   0.732         0.570  
## Residual Std. Error (df = 48) 0.907        0.685  
## F Statistic (df = 1; 48)     134.839***   65.991***  
## =====
```

\*\*p<0.1; \*\*\*p<0.05; \*\*\*\*p<0.01

# Estimation of a Censored Model

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```
##
## Call:
## censReg(formula = y ~ x, data = censoring)
##
## Observations:
##           Total  Left-censored  Uncensored Right-censored
##           50      19           31           0
##
## Coefficients:
##           Estimate Std. error t value Pr(> t)
## (Intercept) -2.10846    0.44764  -4.710 2.48e-06 ***
## x           0.48898    0.06582   7.429 1.09e-13 ***
## logSigma    -0.18013    0.12993  -1.386  0.166
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Newton-Raphson maximisation, 6 iterations
## Return code 1: gradient close to zero (gradtol)
## Log-likelihood: -46.05041 on 3 Df
```

# Count Models

## Dependent variable

- Discrete, integer count data

## Examples

- What are the number of arrests for a person?
- What determines the number of credit cards a person owns?

## Three count data models

- ① Poisson regression
- ② Quasi-Poisson Regression Model
- ③ Negative Binomial Regression Model

## Choice criteria: Presence or absence of overdispersion

- Overdispersion Variance of the dependent variable is larger than its mean.
- Poisson model is not suitable for overdispersion

The main package used is [pscl](#). There is also an additional resource with more theoretical details on the topic: [Regression Models for Count Data in R](#). A more up-to-date version of the document may be found with the [pscl](#) package documentation.

# Poisson Regression Model

Recall Poisson distribution:

$$Pr(Y = k) = \frac{e^{-\lambda} \cdot \lambda^k}{k!}$$

Equidispersion as key characteristics:

- Mean and variance equal to  $\lambda$ , i.e.,  $E(Y) = \lambda$  and  $Var(Y) = \lambda$
- Poisson regression:  $\lambda = \exp(\beta_0 + \beta_1 \cdot x_1 + \dots + \beta_k \cdot x_k)$ .

# NHTS Example: Number of Vehicles (hhpub)

## Data source

- 2017 [National Household Travel Survey](#)
- Survey quantifying trip and travel habits across the United States
- Example use: Quantifying intra-day electricity demand from electric vehicles

## Outcome of interest

- Number of vehicles based on household income, home ownership, and urban/rural household location

## Data preparation

- Elimination of missing and unknown data value
- Conversion of income to 1,000 dollars

## Data Preparation

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```
hhpubdata = subset(hhpub, HHFAMINC %in% c(1:11) &
                    HOMEOWN %in% c(1,2) &
                    URBRUR %in% c(1,2) &
                    HHVEHCNT %in% c(0:12))

HHFAMINC = c(1:11)
INCOME = c(10, 12.5, 20, 30, 42.5, 57.5, 82.5, 112.5, 137.5,
           175, 200)

INCOME = data.frame(HHFAMINC, INCOME)
hhpubdata = merge(hhpubdata, INCOME)
hhpubdata$RURAL = hhpubdata$URBRUR-1
hhpubdata$RENT = hhpubdata$HOMEOWN-1
```



## Poisson Model Execution

Preliminary step: Calculation of mean and variance of dependent variable

```
mean(hhpubdata$HHVEHCNT)
```

```
## [1] 1.981142
```

```
var(hhpubdata$HHVEHCNT)
```

```
## [1] 1.386027
```

Similar values and thus, Poisson regression model as an appropriate first step.

```
bhat_pois = glm(HHVEHCNT~INCOME+RENT+RURAL,  
                data=hhpubdata,family=poisson)
```

```
##
## Call:
## glm(formula = HHVEHCNT ~ INCOME + RENT + RURAL, family = poisson,
##      data = hhpubdata)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.6889  -0.5568  -0.1558   0.3590   5.5063
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  4.654e-01  4.292e-03  108.43  <2e-16 ***
## INCOME       2.986e-03  3.601e-05   82.93  <2e-16 ***
## RENT        -3.733e-01  5.797e-03  -64.39  <2e-16 ***
## RURAL        2.224e-01  4.616e-03   48.19  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##      Null deviance: 86505  on 124400  degrees of freedom
## Residual deviance: 68533  on 124397  degrees of freedom
## AIC: 370161
##
## Number of Fisher Scoring iterations: 5
```

## Interpretation

Sign of coefficients as an indication of the direction of influence on the outcome variable, i.e., the number of cars.

- Association of higher income and rural living with a higher number of car
- Association of renting with lower number of vehicles.
- Possible correlation between income and renting

General coefficient interpretation using  $\exp(\beta)$ , i.e., every unit increase in  $X$  has a multiplicative effect of  $\exp(\beta)$  on the mean of  $Y$ , i.e.,  $\lambda$ :

- $\beta = 0 \Rightarrow \exp(\beta) = 1$ :  $Y$  and  $X$  are not related.
- $\beta > 0 \Rightarrow \exp(\beta) > 1$ : Expected count  $E(y)$  is  $\exp(\beta)$  times larger than when  $X = 0$
- $\beta < 0 \Rightarrow \exp(\beta) < 1$ : Expected count  $E(y)$  is  $\exp(\beta)$  times smaller than when  $X = 0$

# Testing for Overdispersion I

Function `overdispersion()` from the package [AER](#):

- Tests the null hypothesis of equidispersion (i.e., assuming no overdispersion)

Executed after the main regression using `glm(...,family=poisson)`

## Testing for Overdispersion II

```
dispersiontest(bhat_pois)
```

```
##  
##   Overdispersion test  
##  
## data:  bhat_pois  
## z = -115.75, p-value = 1  
## alternative hypothesis: true dispersion is greater than 1  
## sample estimates:  
## dispersion  
##   0.5670593
```

Given the  $p$ -value, the null hypothesis cannot be rejected. If the data suggests overdispersion, two alternative regression models can be used: (1) Quasi-Poisson and (2) Negative Binomial.

# Quasi-Poisson Regression Model

Dataset `blm` from article [Black Lives Matter: Evidence that Police-Caused Deaths Predict Protest Activity](#).

- Dependent variable: Total number of protests in a city
- Note that the paper includes a significant number of supplementary materials which allows for the replication of the results and much more.

First step: Calculation of mean and variance of the variable *totalprotests*:

```
mean(blm$totprotests)
```

```
## [1] 0.4959529
```

```
var(blm$totprotests)
```

```
## [1] 6.35326
```

## Presence of Overdispersion

Likely overdispersion due to variance being significantly higher than mean. In a first step, a regular Poisson model is estimated.

```
eq1      = "totprotests~log(pop)+log(popdensity)+percentblack+
            blackpovertyrate+I(blackpovertyrate^2)+
            percentbachelor+collegeenrollpc+demshare"

eq2      = paste(eq1,"+deathsblackpc",sep="")
eq3      = paste(eq1,"+deathspc",sep="")
bhat1    = glm(eq1,data=blm,family=poisson)
bhat2    = glm(eq1,data=blm,family=quasipoisson)
```

# Estimation Results

```
##
## =====
##                               Dependent variable:
##                               -----
##                               totprotests
##                               Poisson      glm: quasipoisson
##                               link = log
##                               (1)          (2)
## -----
## log(pop)                1.129*** (0.040)  1.129*** (0.062)
## log(popdensity)         -0.183** (0.087)  -0.183 (0.135)
## percentblack            0.017*** (0.003)  0.017*** (0.005)
## blackpovertyrate        0.146*** (0.026)  0.146*** (0.041)
## I(blackpovertyrate2)    -0.002*** (0.0004) -0.002** (0.001)
## percentbachelor         0.039*** (0.004)  0.039*** (0.006)
## colleegenrollpc         0.009*** (0.002)  0.009** (0.004)
## demshare                0.043*** (0.005)  0.043*** (0.008)
## Constant                -20.009*** (0.633) -20.009*** (0.984)
## -----
## Observations              1,226          1,226
## Log Likelihood            -612.473
## Akaike Inf. Crit.        1,242.946
## =====
## Note:                      *p<0.1; **p<0.05; ***p<0.01
```

Note: Switch of statistical significance for population density



## Testing for Overdispersion

```
##  
##   Overdispersion test  
##  
## data:  bhat1  
## z = 1.4052, p-value = 0.07998  
## alternative hypothesis: true dispersion is greater than 1  
## sample estimates:  
## dispersion  
##    2.212733
```

Null hypothesis rejected at 10% but not 5% significance level. The Quasi-Poisson Regression Model handles overdispersion by adjusting standard errors but leaving the coefficient estimates the same.

# Negative Binomial Regression Model

The Negative Binomial Regression Model can be used in the presence of count data and overdispersion. Below, the results from the article [Black Lives Matter: Evidence that Police-Caused Deaths Predict Protest Activity](#) are recreated using the negative binomial models presented in the paper.

Three models:

- ① Resource mobilization and opportunity structure
- ② Adding black death
- ③ Adding all police-caused deaths instead (victims of any race)

# BLM Models

```
bhat3 = glm.nb(eq1,data=blm,link=log)
bhat4 = glm.nb(eq2,data=blm,link=log)
bhat5 = glm.nb(eq3,data=blm,link=log)
```

# BLM Model Results

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```
##
## =====
##                               Dependent variable:
##                               -----
##                               totprotests
##                               (1)         (2)         (3)
## -----
```

## log(pop)	1.292*** (0.072)	1.281*** (0.070)	1.277*** (0.071)
## log(popdensity)	-0.313** (0.133)	-0.305** (0.131)	-0.312** (0.132)
## percentblack	0.022*** (0.005)	0.018*** (0.005)	0.022*** (0.005)
## blackpovertyrate	0.132*** (0.031)	0.128*** (0.031)	0.129*** (0.031)
## I(blackpovertyrate2)	-0.001*** (0.0005)	-0.001*** (0.0005)	-0.001*** (0.0005)
## percentbachelor	0.045*** (0.005)	0.044*** (0.005)	0.045*** (0.005)
## colleegenrollpc	0.011** (0.004)	0.010** (0.004)	0.010** (0.004)
## demshare	0.041*** (0.007)	0.041*** (0.007)	0.041*** (0.007)
## deathsblackpc		2.825*** (0.931)	
## deathspc			0.956 (0.633)
## Constant	-20.905*** (1.117)	-20.734*** (1.101)	-20.801*** (1.108)
## -----			
## Observations	1,226	1,226	1,226
## Log Likelihood	-551.093	-546.677	-549.919
## theta	1.559*** (0.351)	1.686*** (0.404)	1.622*** (0.374)
## Akaike Inf. Crit.	1,120.187	1,113.353	1,119.839
## =====			
## Note:	*p<0.1; **p<0.05; ***p<0.01		

# Hurdle and Zero-Inflation Models

## Problem:

- Presence of many observations at 0 in count data
- Issues using Poisson or a Negative-Binomial Regression Model.

## Application of hurdle and zero-inflated models:

- Data NMES1988 from the package [AER](#)
- Data BLM protests

## NMES1988 Data:

- 4406 observations of people on Medicare who are 66 years or older.
- Outcome of interest: Number of doctor *visits*
- Independent variables: *hospital* (number of hospital visits), *health* (self-indicated health status), *chronic* (number of chronic conditions), *gender*, *school*, and *insurance*.

```
eq          = visits~hospital+health+chronic+gender+
              school+insurance
bhat_pois   = glm(eq,data=NMES1988,family=poisson)
bhat_nb     = glm(eq,data=NMES1988)
bhat_hurdle = hurdle(eq,data=NMES1988,dist="negbin")
bhat_zi     = zeroinfl(eq,data=NMES1988)
```





# Survival Models

Length of time until a certain event occurs and variables influencing time passed (also known as time-to-event data analysis). Examples:

- Time to failure of mechanical device
- Time to death after diagnosis with a certain disease
- Time to re-arrest after release from prison
- Time to defaulting on loan or mortgage

Data used for this topic

- `rossi`

## Theoretical Aspects

$T$  as a random variable representing survival time with the cumulative distribution function written as:

$$F(t) = \Pr(T \leq t)$$

where  $t$  is a realization of  $T$ . Survival function as the complement probability (at least  $t$ ) :

$$S(t) = 1 - F(t) = \Pr(T \geq t)$$

Hazard function or hazard rate  $h(t)$  as risk of failure at time  $t$ .

## Example Data *rossi*

Experimental recidivism study on 432 male prisoners over a period of one year after release from prison ([Rossi et al., 1980](#)):

- *week*: Week of first arrest after release
- *arrest*: Event indicator equal to 1 for rearrest during study period
- *fin*: Receipt of financial aid after release from prison (randomly assigned factor by the researchers)
- *age*: Age at the time of release
- *race*: Black and other
- *wexp*: Full-time work experience prior to incarceration
- *mar*: Married at the time of release
- *paro*: Released on parole
- *prio*: Number of prior convictions.
- *educ*: Education coded as 2 (grade 6 or less), 3 (grades 6-9), 4 (grades 10-11), 5 (grade 12), or 6 (some post-secondary).

Number of prisoners rearrested during study period:

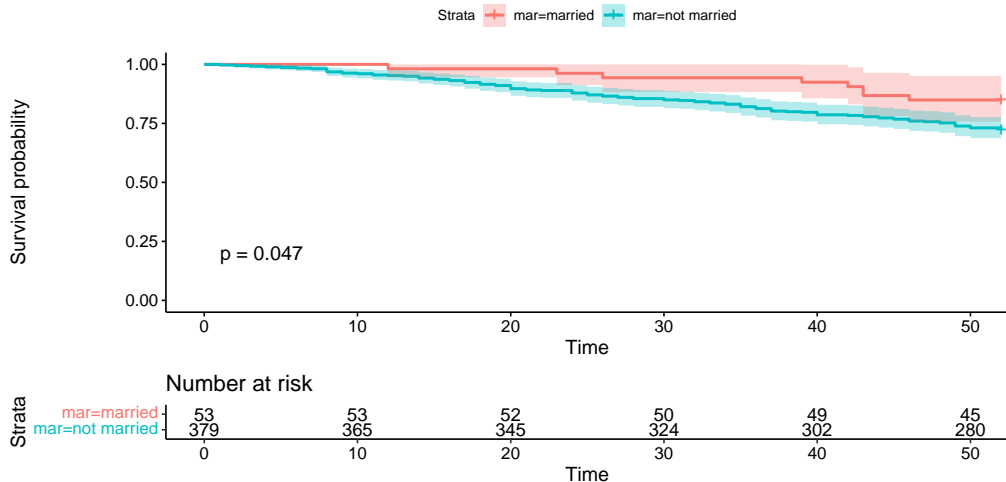
```
sum(rossi$arrest)
```

```
## [1] 114
```

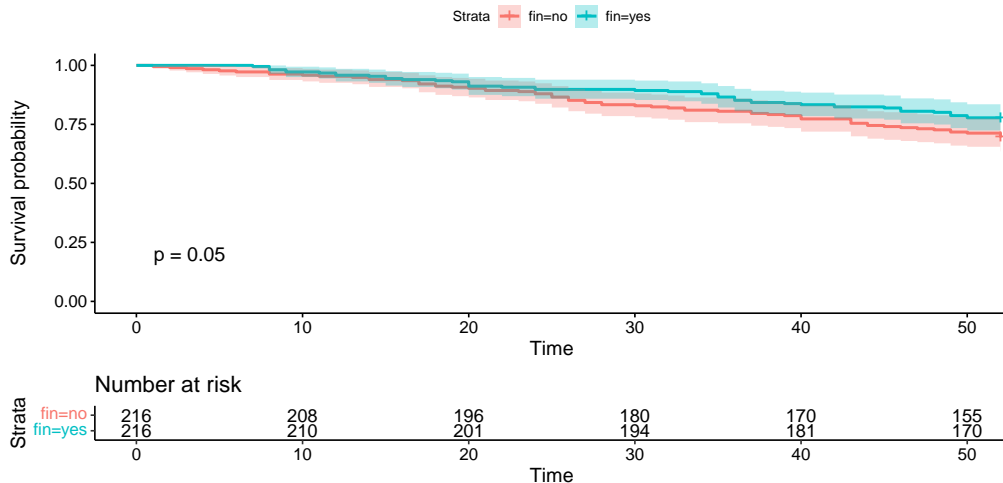
Survival object in R created by function `Surv()`:

```
bhatmar = survfit(Surv(week,arrest)~mar,data=rossi)
bhatfin = survfit(Surv(week,arrest)~fin,data=rossi)
ggsurvplot(bhatmar,pval=TRUE,risk.table=TRUE)
ggsurvplot(bhatfin,pval=TRUE,risk.table=TRUE)
```

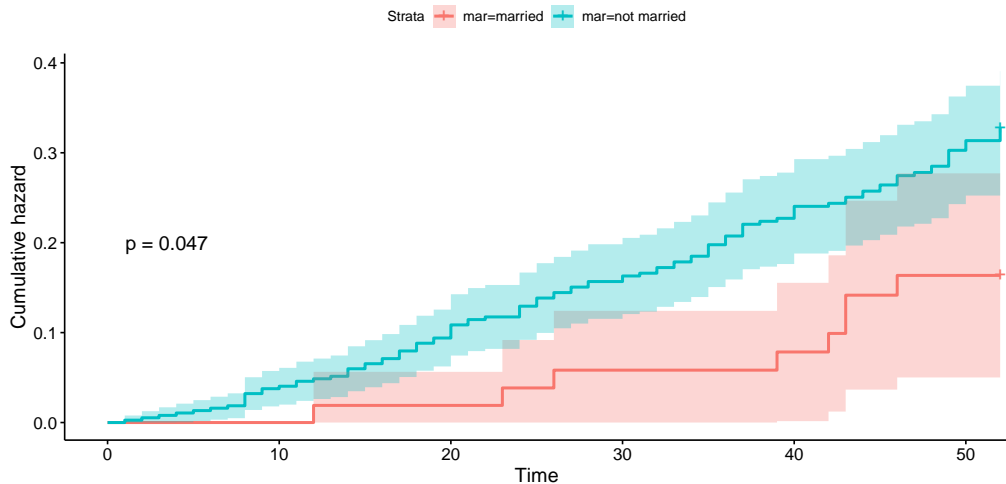
# Survival Curve: Marriage



# Survival Curve: Financial Aid

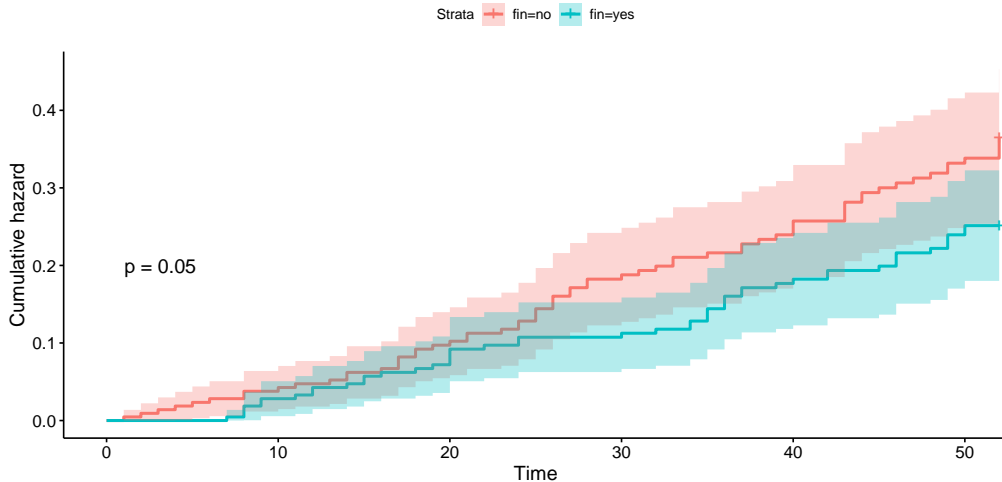


# Cumulative Hazard Function: Marriage





# Cumulative Hazard Function: Financial Aid



## Cox Regression in R: Setup

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```
bhat1 = coxph(Surv(week,arrest)~mar,data=rossi)
bhat2 = coxph(Surv(week,arrest)~mar+fin,data=rossi)
bhat3 = coxph(Surv(week,arrest)~fin+age+race+wexp+mar+paro+prio,
              data=rossi)
```

Statistically insignificant variables excluded from regression output on next slide due to space constraints: *paroyes*, *raceother*, and *wexpyes*

- In general, all variables must be reported!

# Cox Regression in R: Results

```
##
## =====
##                               Dependent variable:
##                               -----
##                               week
##                               (1)      (2)      (3)
## -----
## marnot married      0.712*      0.738**      0.434
##                      (0.367)      (0.367)      (0.382)
## prio                0.091***
##                      (0.029)
## finyes              -0.387**      -0.379**
##                      (0.190)      (0.191)
## age                 -0.057***
##                      (0.022)
## -----
## Observations        432          432          432
## R2                   0.011        0.020        0.074
## Max. Possible R2    0.956        0.956        0.956
## Log Likelihood      -673.060      -670.955      -658.748
## Wald Test           3.770* (df = 1) 7.930** (df = 2) 32.110*** (df = 7)
## LR Test             4.642** (df = 1) 8.852** (df = 2) 33.266*** (df = 7)
## Score (Logrank) Test 3.935** (df = 1) 8.139** (df = 2) 33.529*** (df = 7)
## =====
## Note:                *p<0.1; **p<0.05; ***p<0.01
```