

Predicting Seasonal Influenza Hospitalizations

Kevin W. McConeghy, Jason R. Gantenberg, Andrew R. Zullo, Chanelle J. Howe, ... (author list tentative)

Compiled: 2019-08-05

Purpose

To determine which of a set of simple candidate statistical models (CSM) most closely fits a series of hypothetical influenza hospitalization curves (HIHC), stratified by season severity ([Centers for Disease Control and Prevention, 2016](#)).

Which of these simple CSMs would provide the best severity forecast at the beginning of the flu season (i.e., epiweek 40) based purely on fit to the HIHCs described below?

Subquestion: Which model best predicts the proportion of flu-attributable hospitalizations?

General approach

1. Fit a quadratic trend filter model ([Brooks et al., 2015](#)) to FluSurv-NET hospitalization data for the seasons 2003–2004 through 2017–2018, based in part on the approach described by Brooks et al. ([Brooks et al., 2015](#)).
2. Simulate 3,000 HIHCs using the empirical Bayes model as a “generative model”.
3. Stratify these HIHCs into groups representing High/Moderate and Mild severity HIHCs, based on the CDC’s categorization ([Biggerstaff et al., 2018](#); [Centers for Disease Control and Prevention, 2018](#)).
4. Fit each CSM within each severity stratum and test for goodness of fit. Systematically alter the functional form of epiweek indicators as in Wang et al.’s study of influenza mortality ([Wang et al., 2012](#)).

Forecasting targets

All weeks will be specified using the *MMWR* Week convention ([Centers for Disease Control and Prevention, a](#)):

1. Peak week
2. Peak number of hospitalizations
3. Total hospitalizations

These targets follow in part from ([Brooks et al., 2015](#)).

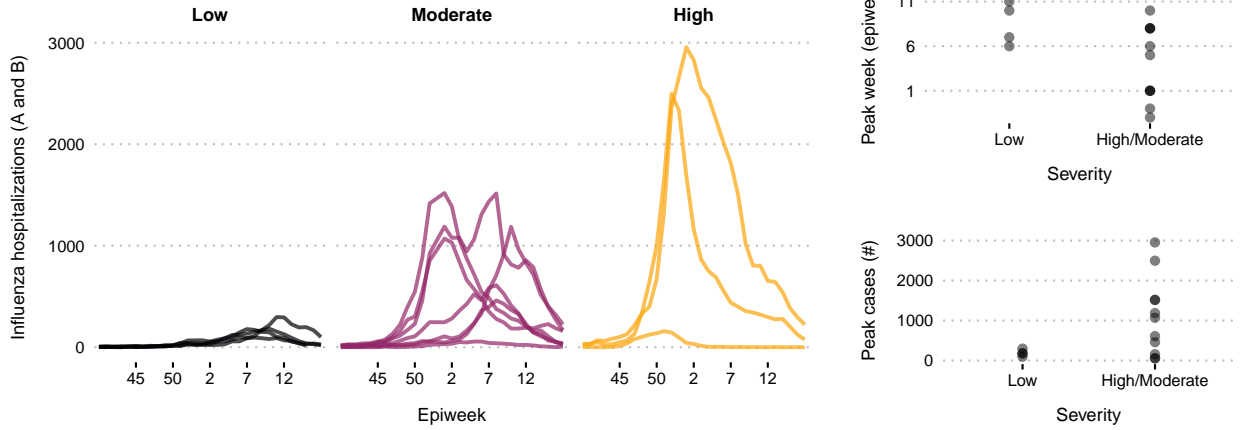


Figure 1: Empirical hospitalization curves, peak weeks, and peak number of cases – 2003–2017 (Source: FluSurv-NET, CDC). Data excludes 2009–2010 pandemic influenza season and 2017–2018 due to no official severity designation.

Curve fitting and simulation

We will simulate HHCs using a modified version of the curve-fitting approach described by Brooks et al. (Brooks et al., 2015).

First, we fit a quadratic trend filter to historical hospitalization curves released by the CDC (beginning with the 2003–2004 season), using the R package `glmgen` (Brooks et al., 2015; Centers for Disease Control and Prevention, 2016). This model is fit within flu severity group (Mild, High/Moderate).

For each season s and each week i , Brooks et al. conceptualize a seasonal influenza curve as some function plus noise¹:

$$y_i^s = f^s(i) + \epsilon_i^s, \epsilon \sim N(0, \tau^s),$$

where

$$f^s(i) = \frac{\theta}{\max_j f(j)} \left[f \left(\frac{i - \mu}{v} + \arg \max_j f(j) \right) \right]$$

Based on fitting the quadratic trend filtering model to empirical data (i.e., historical CDC flu hospitalization data), we estimate error τ^s :

$$(\hat{\tau}^s)^2 = \text{avg}_i \left[y_i^s - \hat{f}^s(i) \right]^2$$

and then sample from the model, introducing noise for each weekly observation based this estimate of τ^2 .

We impose a lower bound of 0 hospitalizations via the following transformation of \hat{y}_i^s , which we denoted below as \hat{z}_i^s :

$$\hat{z}_i^s = 0.5 \left(|\hat{y}_i^s| + \hat{y}_i^s \right)$$

¹ Brook et al. use their method to forecast a current flu season, where b represents the current season's epidemic threshold of weighted influenza-like illness percent. Because hospitalization curves have a lower bound of 0, we drop b from the original equation.

Note: In the Brooks paper, exactly what j stood for was a little unclear. I believe the only sensible interpretation is that it is the indicator for the week in the trend filter predictions $f(j)$ conducted as part of each curve simulation.

Parameters in quadratic trend filtering model

The HIHCs are simulated using the following sampling scheme for each parameter represented in the model for hospitalization count (y_i^s) at each week. Note that all equations are either adapted from or appear in (Brooks et al., 2015).

$$\langle f, o, \nu, \theta, \mu \rangle$$

Shape (f)

$$f \sim U\{\hat{f} : \text{historical season } s\}$$

Noise (σ)

$$\sigma \sim U\{\hat{\tau}^s : \text{historical season } s\}$$

Peak height (θ)

$$\theta \sim U[\theta_m, \theta_M]$$

Results in the following curve:

$$f_3(i) = f_2(i - \mu + \arg \max_j f_2(j))$$

Pacing (ν)

Curve-stretching around peak week:

$$\nu \sim U[0.75, 1.25]$$

Results in following curve:

$$f_4(i) = f_3\left(\frac{i - \arg \max_j f_3(j)}{\nu} + \arg \max_j f_3(j)\right)$$

Note: Brooks et al. say they get "unbiased estimators" for the minimum and maximum peak heights θ_m and θ_M , respectively. However, given the notation does not indicate these parameters are estimates, I am using the observed minimum and maximum heights based on CDC data.

Candidate models

Incidence-rate difference models

- Generate a set of predicted flu-related hospitalizations based on a standard incidence rate difference model typically used to model flu cases. Implement in flumodelr based on (Izurieta et al., 2000; Thompson et al., 2009).
- Thresholds: 0.1 and 0.15 as in (Thompson et al., 2009).

Question for Kevin: Does it make sense to use the rate-difference model in the context of hospitalizations? We essentially have a rate of 0 (or very close to it) at the beginning of each flu season, unlike with wILI.

Serfling model (least squares)

Modified from ([McConeghy et al.](#)):

$$Y = \beta_0\alpha + \beta_1t + \beta_rX_r + \dots + \beta_p\cos\left(\frac{2\pi t}{52}\right) + \beta_q\sin\left(\frac{2\pi t}{52}\right)$$

Where t = time (epiweek), subscript r denotes a vector of β coefficients and variables, and subscripts p and q take on particular numbers based on the length of r .

Modified Serfling model

Modified from ([McConeghy et al.](#)):

$$y = \alpha_0 + \beta_1t + \beta_2Flu_t + \beta_pX_p + \dots + \sin\left(\frac{2\pi t}{period}\right) + \cos\left(\frac{2\pi t}{period}\right) + u$$

Where t = time (epiweek) and subscript p denotes a vector of *beta* coefficients and variables.

Generalized additive model (Prophet)

This model will be implemented using the R package *prophet* ([Taylor and Letham, 2018](#)), which implements a generalized additive modeling approach developed by engineers at Facebook, Inc.

The general form of the equation that will be fit to the IHHCs:

$$y(t) = g(t) + \dots + h(t) + \epsilon_t,$$

where $g(t)$ models nonperiod trends and $h(t)$ stands for a vector of holidays or other events that are known to correlate with flu hospitalization ([Brooks et al., 2015](#)). As with the Serfling and modified Serfling models, the Prophet model will include additional model terms included to improve fit.

Model terms

The following model terms will be entered into the vector of covariates considered for each model:

- Cyclical terms (Serfling, Fourier, etc.)
- Historical (empirical) flu hospitalizations
- Historical data on viral activity (NREVSS), outpatient surveillance (ILI-Net)
- Average weekly temperature

Question for Kevin: Do you agree that we won't need seasonality terms here? We are essentially modeling one flu hospitalization season at a time, so it's unclear to me where seasonality would come into the picture. Consequently, does this fact obviate the need for cyclical terms in the Prophet model and perhaps any need for the Serfling models? The curves look as if they could be modeled simply as quadratic functions.

- e) Climate factors (e.g., prior summer temperatures)
- f) Lags and leads of c) or d)
- g) Indicators for weeks of Thanksgiving and/or Christmas ([Brooks et al., 2015](#); [Taylor and Letham, 2018](#))

Goodness of fit

- Root mean squared error (RMSE)
- Bayesian information criterion (BIC)
- Relative bias

Sensitivity analysis

Challenge

The composition of institutions reflected in the FluSurv-NET has changed over time ([Centers for Disease Control and Prevention, b](#); [Kandula et al., 2019](#)), meaning the FluSurv-NET estimates for influenza-related hospitalizations may not be comparable across time.

Response

Redo the analysis three times: one for each set of years in which the same participating institutions reported flu data to CDC. See ([Centers for Disease Control and Prevention, b](#)) for more information.

Limitations

Brooks et al. showed that their empirical Bayes model improved upon standard lagged CDC predictions for several forecasting targets of the overall influenza curves (season onset, peak week, peak rate/count, duration of season). In adapting their approach to hospitalizations, we should ensure we can achieve similar accuracy.

References

- Matthew Biggerstaff, Krista Kniss, Daniel B Jernigan, Lynnette Brammer, Joseph Bresee, Shikha Garg, Erin Burns, and Carrie Reed. Systematic assessment of multiple routine and near Real-Time indicators to classify the severity of influenza seasons and pandemics in the united states, 2003-2004 through 2015-2016. *Am. J. Epidemiol.*, 187(5):1040–1050, May 2018. ISSN 0002-9262, 1476-6256. doi: 10.1093/aje/kwx334.
- Logan C Brooks, David C Farrow, Sangwon Hyun, Ryan J Tibshirani, and Roni Rosenfeld. Flexible modeling of epidemics with an empirical bayes framework. *PLoS Comput. Biol.*, 11(8):e1004382, August 2015. ISSN 1553-734X, 1553-7358. doi: 10.1371/journal.pcbi.1004382.

Question for Dr. Naimi: Should we be doing cross-validation or sample-splitting given the aims of this analysis? We can simulate an arbitrary number of hypothetical hospitalization curves, so the limited number of historical flu season available may not be an issue.

Centers for Disease Control and Prevention. MMWR week overview. https://www.cdc.gov/nndss/document/MMWR_Week_overview.pdf, a. Accessed: 2019-7-9.

Centers for Disease Control and Prevention. Flu view phase 3 quick reference guide, b.

Centers for Disease Control and Prevention. Laboratory-confirmed influenza hospitalizations. <https://gis.cdc.gov/GRASP/Fluview/FluHospRates.html>, 2016. Accessed: 2019-7-1.

Centers for Disease Control and Prevention. How CDC classifies flu severity. <https://www.cdc.gov/flu/about/classifies-flu-severity.htm>, September 2018. Accessed: 2019-7-9.

H S Izurieta, W W Thompson, P Kramarz, D K Shay, R L Davis, F DeStefano, S Black, H Shinefield, and K Fukuda. Influenza and the rates of hospitalization for respiratory disease among infants and young children. *N. Engl. J. Med.*, 342(4):232–239, January 2000. ISSN 0028-4793. doi: 10.1056/NEJM200001273420402.

Sasikiran Kandula, Sen Pei, and Jeffrey Shaman. Improved forecasts of influenza-associated hospitalization rates with google search trends. *J. R. Soc. Interface*, 16(155):20190080, June 2019. ISSN 1742-5689, 1742-5662. doi: 10.1098/rsif.2019.0080.

Kevin W McConeghy, Rob van Aalst, Andrew R Zullo, and Nina Joyce. An R package for estimating attributable influenza morbidity and mortality. <https://kmcconeghy.github.io/flumodelr/>. Accessed: 2019-7-1.

Sean J Taylor and Benjamin Letham. Forecasting at scale. *Am. Stat.*, 72(1):37–45, January 2018. ISSN 0003-1305. doi: 10.1080/00031305.2017.1380080.

William W Thompson, Eric Weintraub, Praveen Dhankhar, Po-Yung Cheng, Lynnette Brammer, Martin I Meltzer, Joseph S Bresee, and David K Shay. Estimates of US influenza-associated deaths made using four different methods. *Influenza Other Respi. Viruses*, 3(1):37–49, January 2009. ISSN 1750-2640, 1750-2659. doi: 10.1111/j.1750-2659.2009.00073.x.

Xi-Ling Wang, Lin Yang, King-Pan Chan, Susan S Chiu, Kwok-Hung Chan, J S Malik Peiris, and Chit-Ming Wong. Model selection in time series studies of influenza-associated mortality. *PLoS One*, 7(6):e39423, June 2012. ISSN 1932-6203. doi: 10.1371/journal.pone.0039423.